```
In [1]: %pylab inline
    import numpy as np
    import scipy
    import io
    import base64
    from IPython.display import HTML, Image
    from scipy.stats import binom, poisson
    from matplotlib.collections import PatchCollection
    from matplotlib.patches import Circle, Rectangle

## Customising the font size
    plt.rcParams.update({'font.size': 14})
    plt.rcParams['figure.figsize'] = [16,8]
```

Populating the interactive namespace from numpy and matplotlib

Lecture Notes Week 8 ¶

Syllabus

- 1. Errors, Probabilities, and Interpretations; basics of presentation of data, mean, spread
- 2. Probability distributions
- 3. Monte Carlo basics
- 4. Parameter Estimation
- 5. Maximum Likelihood + extended maximum likelihood
- 6. Least Square, chi2, correlations, Best Linear Unbiased Estimator
- 7. Probability and Confidence level
- 8. Hypothesis testing
- 9. Goodness of fit tests
- 10. Limit setting
- 11. Introduction to Multivariate Analysis Techniques
- 12. Basics of Unfolding

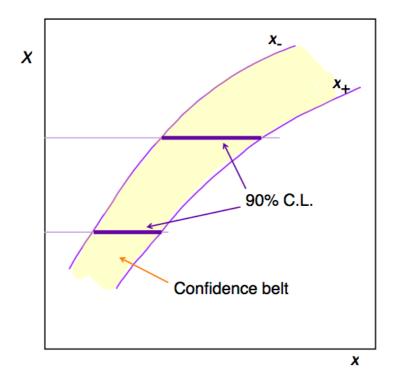
Comments up front

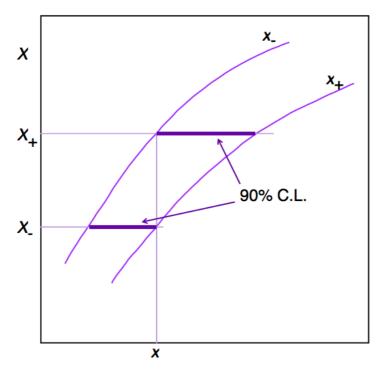
There has been a change in the rhythm for weekly assignments as Rene provided the assignments a bit later and was giving more relaxed timescales. I hope to get feedback from the assignments before the next set of lectures, hence the shorter period. This week's assignment should not have been overly onerous and of course all following assignments will be spaced out by a week. The next assignment is due on 18 April.

The strike period is over (at least for now) and while this course was affected, rest assured that there is a detailed assessment of the impact ongoing to ensure that no student is disadvantaged by the action taken. Last week's lectures will not be rescheduled, but you will

see the main concepts introduced there reappear in today's and other lectures in the future.

Recap





Confidence level

Binomial confidence intervals

For Binomial distributions, events belong to exactly one of two classes, e.g. true or false, greater or smaller than a threshold, male or female, etc. This applies to samples of finite size and so the observed events are discretely distributed. Contrary to that the true variable is continuous, e.g. the probability for an event to be true can take any value and even the expectation value, i.e. the probability multiplied by the sample size is not necessarily an integer.

Recall the Binomial distribution

$$P(r; p, n) = p^{r} (1 - p)^{n - r} \frac{n!}{r!(n - r)!},$$

where p is the probability of success, n is the sample size, and r is the number of successes for this sample.

The expectation value is

$$\langle r \rangle = np.$$

Construction of a Binomial confidence belt

Given that the distribution of events is discrete, the integrals used in the construction of confidence intervals have to be replaced by sums. Recall that for a central interval, for a given confidence level C, we have to determine

$$\int_{-\infty}^{x_{-}} P(x)dx = \int_{x_{+}}^{\infty} P(x)dx = (1 - C)/2.$$

The direct replacement would lead, for a given confidence level C, to

$$\sum_{r=0}^{r_{-}} P(r; p, n) = \sum_{r=r_{+}}^{n} P(r; p, n) = (1 - C)/2.$$

In general the discrete nature of r will prevent these equalities to be satisfied exactly. Therefore they have to be replaced by inequalities that guarantee that the confidence interval covered by the range r_- to r_+ is at least C. This is given by the following constructions

$$\sum_{r=0}^{r_{+}} P(r; p, n) \ge 1 - (1 - C)/2.$$

and

$$\sum_{r=r}^{n} P(r; p, n) \ge 1 - (1 - C)/2.$$

Finally, if m successes are observed, the limits on the true probability can be assigned by

$$\sum_{r=m+1}^{n} P(r; p_+, n) = 1 - (1 - C)/2,$$

and

$$\sum_{r=0}^{m-1} P(r; p_-, n) = 1 - (1 - C)/2.$$

In practice, these are the outward-facing corners of the confidence belt at a position r=m. These are also known as the *Clopper-Pearson confidence limits*.

Example: Fraction of female students

In order for this excercise to be applicable to Binomial distributions, we need to work with exclusively two genders. Apologies to everyone who does not consider themselves falling into any of these categories.

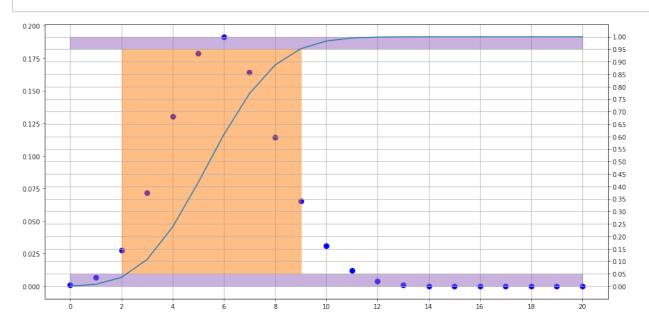
Let's first establish the number of students in the course and those who are female.

• Over to menti.com (https://www.menti.com) with code 94 62 50

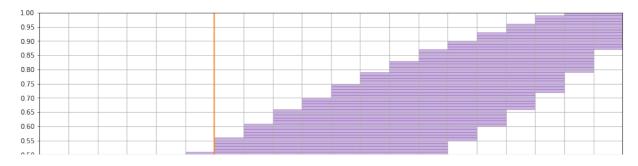
Construction of a confidence belt based on these data

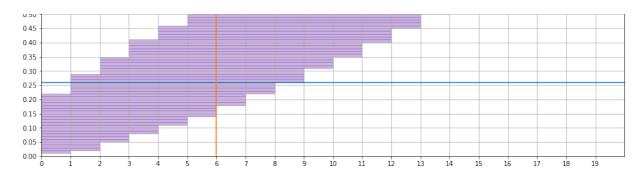
```
x = range(nStudents+1) # list of integers from 0 to nStudents
In [14]:
           n = nStudents
                                     # example probability
           p = 0.3
           limlow = (1-cl)/2  # lower limit of central confidence interval
limhigh = 1-limlow  # upper limit of central confidence interval
rv = binom(n, p)  # initialise binomial distribution
           probs = rv.pmf(x)
                                     # calculate binomial probabilities for all value
           ints = np.cumsum(probs, dtype=float) # calculate cumulative sum of pro
           # work out limits of confidence interval based on cumulative sum of pr
           x low = 0
           x high = nStudents
           for i in x:
                if ints[i] < limlow: x low = i</pre>
                if ints[n-i]>limhigh: x high = n-i
           # prepare two plots with common x axis
           fig,ax1 = plt.subplots()
           ax1.set xticks(x[0::2])
           ax2 = ax1.twinx()
```

```
# draw two bands indicating the excluded part of the confidence level
pp = PatchCollection([Rectangle((0,0),n,limlow),Rectangle((0,limhigh),))
# draw a rectangle indicating the selected confidence interval
pp2 = PatchCollection([Rectangle((x low,limlow),x high-x low,cl)],alpha
# add these to the plot
ax2.add collection(pp)
ax2.add collection(pp2)
# plot the probability distribution
ax1.plot(x,probs, 'bo', ms=8, label='binom pmf')
# plot the cumulative integral
ax2.plot(x,ints, '-', ms=8, label='binom pmf')
# plot optics
ax2.set yticks([x/20. for x in range(21)])
ax1.grid(which='both',axis='x')
ax2.grid(which='major',axis='y')
plt.show()
```



```
sump = 0.
    summ = 0.
    foundp = False
    foundm = False
    rv = binom(n, p) # initialise binomial distribution
    for x in range(0,n+1):
        p = rv.pmf(x) # obtain probability
        sump += p
        if not foundp and 0.5*(1-cl) + cl <= sump: # check for first t
            rp = x
            foundp = True
        p = rv.pmf(n-x) \# obtain probability, summing up from maximum
        summ += p
        if not foundm and 0.5*(1-cl) + cl <= summ: # check for first t
            rm = n-x
            foundm = True
    return rm, rp
nSamples = 100 # defines granularity in y (=probability)
binom patches = []
for i in range(1,nSamples):
    pFemale = 1. * i / nSamples # translate into probability
    #rm, rp = get central interval details(0.95, nStudents, pFemale) #
    rm, rp = get central interval(cl, nStudents, pFemale) # use fast i
    binom_patches.append( Rectangle((rm,pFemale),rp-rm,1./nSamples) )
binoms = PatchCollection(binom patches, alpha=0.5, color='tab:purple')
fig, ax = plt.subplots()
ax.set xlim(0,nStudents) # set x axis range
ax.set ylim(0,1)
                         # set y axis range
ax.add collection(binoms) # draw belt
# plot optics
ax.set yticks([x/20. for x in range(21)])
ax.set xticks([x for x in range(nStudents)])
ax.grid(which='both',axis='x')
ax.grid(which='major',axis='y')
# reality check
ax.plot([0,nStudents],[0.26]*2)
ax.plot([nFemale]*2,[0,1])
plt.show()
```





```
In [5]: from scipy import special
    print(cl,n,p)
    vals = ceil(special.bdtrik(cl, n, p))
    special.bdtr(vals-1, n, p)
    binom.ppf(0.5*(1-cl),n,p)
    print(binom.cdf(9,n,p)-binom.cdf(2,n,p))
    binom.interval(cl,n,p)
```

0.9 20 0.3 0.9165549703701878

Out[5]: (3.0, 9.0)

In []:

Poisson confidence intervals

The procedure for a Poisson process is rather similar to that of a Binomial distribution with the main difference that the range of possible values is not limited by a total number but goes to infinity (this is exactly how the distribution is defined, by taking the $n \to \infty$ limit of the Binomial distribution function).

Recall the Poisson distribution

For a given number of observed events r and an expectation value of λ , the Poisson probability is given by

$$P(r;\lambda) = \frac{e^{-\lambda}\lambda^r}{r!}.$$

Construction of Poisson intervals

A 90% upper limit is therefore given by the value N_{\pm} for which

$$\sum_{r=n+1}^{\infty} P(r; N_+) = 0.90.$$

This is equivalent to

$$\sum_{r=0}^{n} P(r; N_{+}) = 1 - 0.90 = 0.10,$$

which may be easier to calculate.

One useful number to remember is the 90% upper limit for the case that the number of observed events is 0; this limit is $N_+=2.3$. Hence, in any counting experiment that yields an obervation of 0, we can be 90% sure that the true number of events is no greater than 2.3.

Accordingly, the 90% lower limit is given by

$$\sum_{r=0}^{n-1} P(r; N_{-}) = 0.90.$$

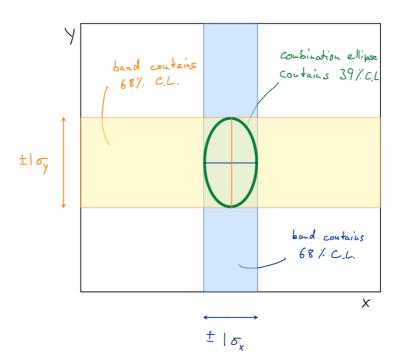
Confidence intervals for several variables: confidence regions

In the case of several variables the confidence intervals does not apply directly, e.g. in the case of two variables, the parameter space corresponding to a certain confidence level is not given by independent intervals in the individual variables, which would result in a rectangular area in parameter space, but rather by a more complex confidence region.

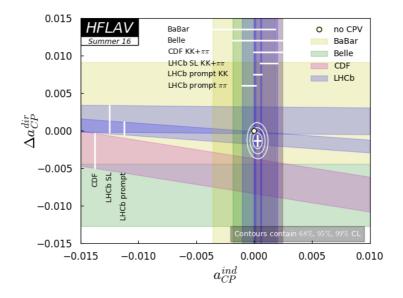
In the case of Gaussian uncertainties, the confidence region in two dimensions is an ellipse. However, this can take more complicated shapes depending on how the parameters are determined and on whether Gaussian uncertainties apply.

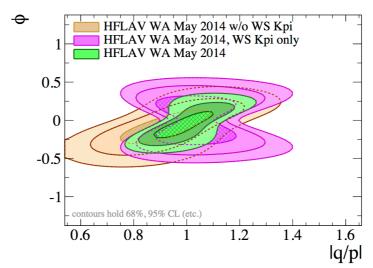
In general, the maximum likelihood method provides a straightforward way of constructing confidence regions as the likelihood is multiplicative and hence the likelihood defining the confidence region is the product of the likelihoods of the individual vairables, potentially accounting for correlations.

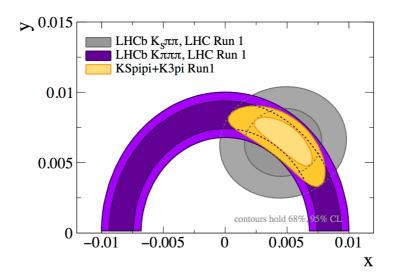
One caveat applies regarding the confidence level of a multi-dimensional region compared to a one-dimensional interval. As illustrated below, the ellispe spanned by the 1D 68% confidence-level intervals, i.e. the standard deviations of the measurements, contains only 39% confidence level as the remainder is covered in infinitely long bands limited only in one variable. It is therefore essential to state the coverage of any multi-dimensional confidence region to avoid confusion and mistakes.



Below are a few real-life examples of more comlicated confidence contours. These are taken from my work in the <u>Heavy Flavor Averaging Group (https://hflav.web.cern.ch)</u> and more details for those who want it can be found on the linked web page, where these particular examples refer to *charm physics*.







Hypothesis testing

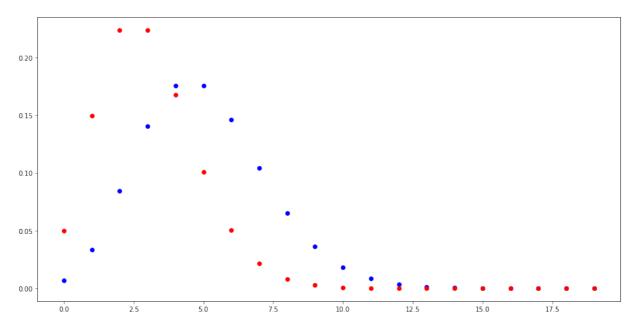
Regarding the lecture notes for this part, please refer to the visualiser copies on blackboard for the moment.

Below is an example of two Poisson distributions that were used in the lecture to illustrate composite hypotheses.

```
In [19]: p5 = poisson(5)
    p3 = poisson(3)
    x = range(20)
    pp5 = p5.pmf(x)
    pp3 = p3.pmf(x)

    plot(x,pp5,'bo')
    plot(x,pp3,'ro')
```

Out[19]: [<matplotlib.lines.Line2D at 0x1a1b0023c8>]



```
In [ ]:
```