# **Aggregate Data**

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Grad IO

#### **Aggregate Data**

- ▶ Now we want to have both price endogeneity and flexible substitution in the same model.
- We are ultimately going with the random coefficients logit model, but we will start with the logit and nested logit.
- ▶ We will explore a technique that works with aggregate data.

### **Multinomial: Aggregation Property**

Each individual's choice  $y_{ij} \in \{0,1\}$  and  $\sum_{j \in \mathcal{J}} d_{ij} = 1$ .

Choices follow a Multinomial distribution with m=1:

$$(d_{i1},\ldots,d_{iJ},d_{i0}) \sim \operatorname{Mult}(1,s_{i1},\ldots,s_{iJ},s_{i0})$$

If each individual faces the same  $s_{ij}=s_j$  the the sum of Multinomials is itself Multinomial:

$$(q_1^*,\ldots,q_J^*,q_0^*) \sim \operatorname{Mult}(M,s_1,\ldots,s_J,s_0)$$

where  $q_j^* = \sum_{i=1}^M d_{ij}$  is a sufficient statistic.

# **Multinomial: Aggregation Property (Likelihood)**

We can write the likelihood as  $L\left((y_{i1},\ldots,y_{iJ},y_{i0})\mid\mathbf{x_i},\theta\right)$  where  $\mathbf{x_i}$  is a J vector that includes all relevant product characteristics interacted with all relevant individual characteristics.

$$\begin{split} &= \left( \begin{array}{c} M \\ q_{i1}, \dots, q_{iJ}, q_{i0} \end{array} \right) \prod_{i=1}^M s_{i1}(\mathbf{x_i}, \theta)^{d_{i1}} \cdots s_{iJ}(\mathbf{x_i}, \theta)^{d_{iJ}} s_{i0}(\mathbf{x_i}, \theta)^{d_{i0}} \\ &\rightarrow \ell(\mathbf{x_i}, \theta) = \sum_{i=1}^M \sum_{j \in \mathcal{J}} d_{ij} \log s_{ij}(\mathbf{x_i}, \theta) + \log C(\mathbf{q}) \end{split}$$

If all individuals face the same  $(\mathbf{x_i})$  and  $\mathcal J$  they will have the same  $s_{ij}(\mathbf{x_i},\theta)$  and we can aggregate outcomes into sufficient statistics.

$$\to \ell(\theta) = \sum_{j \in \mathcal{J}} q_j^* \log s_j(\theta)$$

### **Multinomial Logit: Estimation with Aggregate Data**

#### Aggregation is probably the most important property of discrete choice:

- ▶ Instead of individual data, or a single group we might have multiple groups: if prices only change once per week, we can aggregate all of the week's sales into one "observation".
- ▶ Likewise if we only observe that an individual is within one of five income buckets there is no loss from aggregating our data into these five buckets.
- ▶ All of this depends on the precise form of  $s_{ij}(\mathbf{x_i}, \theta)$ . When it doesn't change across observations: we can aggregate.
- ▶ Notice I didn't need anything to follow a logit/probit.

# **Aggregation with Unobserved Heterogeneity**

$$s_{ij}(\mathbf{x_i}, \theta) \quad = \quad \int \frac{\exp[x_{ij}\beta_\iota]}{1 + \sum_k \exp[x_{ik}\beta_\iota]} f(\beta_\iota|\theta) \partial \beta_\iota = \sum_{\iota=1}^S w_\iota \frac{\exp[x_{ij}\beta_\iota]}{1 + \sum_k \exp[x_{ik}\beta_\iota]}$$

- lacksquare Notice that while i subscripts "individuals" with different characteristics  $\mathbf{x_i}$
- $ightharpoonup \iota$  is the dummy index of integration/summation.
  - Even though we sometimes call these "simulated individuals"
  - $\bullet$  Everyone with the same  $\mathbf{x_i}$  still has the same  $s_{ij}(\mathbf{x_i},\theta)$
- lacktriangle Most papers will abuse notation and i will serve double duty!

### **Multinomial Logit: Estimation with Aggregate Data**

Now suppose we have aggregate data:  $(q_1,\ldots,q_J,q_0)$  where  $M=\sum_{j\in\mathcal{J}}q_j$ .

- $\blacktriangleright \ \text{ If } M \text{ gets large enough then } (\tfrac{q_1}{M}, \ldots, \tfrac{q_J}{M}, \tfrac{q_0}{M}) \to (\mathfrak{s}_1, \ldots, \mathfrak{s}_J, \mathfrak{s}_0)$ 
  - Idea: Observe  $(\mathfrak{s}_1(\mathbf{x_i}), \dots, \mathfrak{s}_J(\mathbf{x_i}), \mathfrak{s}_0(\mathbf{x_i}))$  without sampling variance.
  - ullet Challenges: We probably don't really observe  $q_0$  and hence  ${\cal M}.$
- ▶ Idea: Equate observed market shares to the conditional choice probabilities  $(s_1(\mathbf{x_i}, \theta), \dots, s_J(\mathbf{x_i}, \theta), s_0(\mathbf{x_i}, \theta)).$
- lacktriangle Choose heta that minimizes distance: MLE? MSM? Least Squares? etc.

### **Inversion: IIA Logit**

Add unobservable error for each  $\mathfrak{s}_{jt}$  labeled  $\xi_{jt}.$ 

$$u_{ijt} = x_{jt}\beta - \alpha p_{jt} + \xi_{jt} + \varepsilon_{ijt}, \quad s_{jt} = \frac{\exp[x_{jt}\beta - \alpha p_{jt} + \xi_{jt}]}{1 + \sum_k \exp[x_{kt}\beta - \alpha p_{kt} + \xi_{kt}]}$$

- lacktriangle The idea is that  $\xi_{jt}$  is observed to the firm when prices are set, but not to us the econometricians.
- $\blacktriangleright$  Potentially correlated with price  $\operatorname{Corr}(\xi_{jt},p_{jt})\neq 0$
- $\blacktriangleright$  But not characteristics  $E[\xi_{jt}|x_{jt}]=0.$ 
  - This allows for products j to better than some other product in a way that is not fully explained by differences in  $x_j$  and  $x_k$ .
  - Something about a BMW makes it better than a Peugeot but is not fully captured by characteristics that leads higher sales and/or higher prices.
  - Consumers agree on its value (vertical component).

Taking logs:

$$\begin{split} \ln s_{0t} &= -\log \left(1 + \sum_k \exp[x_{kt}\beta + \xi_{kt}]\right) \\ \ln s_{jt} &= [x_{jt}\beta - \alpha p_{jt} + \xi_{jt}] - \log \left(1 + \sum_k \exp[x_{kt}\beta + \xi_{kt}]\right) \\ \underbrace{\ln s_{jt} - \ln s_{0t}}_{Data!} &= x_{jt}\beta - \alpha p_{jt} + \xi_{jt} \end{split}$$

#### Exploit the fact that:

- 1.  $\ln s_{jt} \ln s_{0t} = \ln \mathfrak{s}_{jt} \ln \mathfrak{s}_{0t}$  (with no sampling error)
- 2. We have one  $\xi_{jt}$  for every share  $\boldsymbol{s}_{jt}$  (one to one mapping)

### **IV Logit Estimation (Berry 1994)**

- 1. Transform the data:  $\ln \mathfrak{s}_{jt} \ln \mathfrak{s}_{0t}$ .
- 2. IV Regression of:  $\ln\mathfrak{s}_{jt}-\ln\mathfrak{s}_{0t}$  on  $x_{jt}\beta-\alpha p_{jt}+\xi_{jt}$  with IV  $z_{jt}$

#### Was it magic?

- $lackbox{ No. It was just a nonlinear change of variables from } s_{jt} 
  ightarrow \xi_{jt}.$
- $\,\blacktriangleright\,$  Our (conidtional) moment condition is just that  $E[\xi_{jt}|x_{jt},z_{jt}]=0.$
- We moved from the space of shares and MLE for the logit to the space of utilities and an IV model.
  - We are losing some efficiency but now we are able to estimate under weaker conditions.
  - But we need aggregate data and shares without sampling variance.

### **Naive Approach**

Did we need to do change of variables? Imagine we work with:

$$\begin{split} s_{jt} &= \frac{\exp[x_{jt}\beta - \alpha p_{jt}]}{1 + \sum_k \exp[x_{kt}\beta - \alpha p_{kt}]} \\ \eta_{jt} &\equiv (s_{jt}(\theta) - \mathfrak{s}_{jt}) \end{split}$$

- $\blacktriangleright$  Each share depends on all prices  $(p_{1t},\ldots,p_{Jt})$  and characteristics  $\mathbf{x_t}.$
- ► Harder to come up with IV here.

### Inversion: Nested Logit (Berry 1994 / Cardell 1991)

This takes a bit more algebra but not much

$$\underbrace{\frac{\ln s_{jt} - \ln s_{0t} - \rho \log(s_{j|gt})}_{\text{data!}}}_{\text{data!}} = x_{jt}\beta - \alpha p_{jt} + \xi_{jt}$$
 
$$\ln \mathfrak{s}_{jt} - \ln \mathfrak{s}_{0t} = x_{jt}\beta - \alpha p_{jt} + \rho \log(\mathfrak{s}_{j|gt}) + \xi_{jt}$$

- lacktriangle Same as logit plus an extra term  $\log(s_{j|q})$  the within group share.
  - We now have a second endogenous regressor.
  - ullet If you don't see it realize we are regressing Y on a function of Y. This should always make you nervous.
- lacksquare If you forget to instrument for ho you will get ho o 1 because of attenuation bias.
- $\blacktriangleright\,$  A common instrument for  $\rho$  is the number of products within the nest. Why?

# BLP 1995/1999 and Berry Haile (2014)

Think about a generalized inverse for  $\sigma_j(\mathbf{x}_t,\theta_2)=\mathfrak{s}_{jt}$  so that

$$\sigma_{jt}^{-1}(\mathcal{S}_{\cdot t}, \tilde{\theta}_2) = x_{jt}\beta - \alpha p_{jt} + \xi_{jt}$$

- lacktriangle After some transformation of data (shares  $\mathcal{S}_{\cdot t}$ ) we get mean utilities  $\delta_{jt}$ .
  - We assume  $\delta_{jt}=h(x_{jt},v_{jt},\theta_1)-\alpha p_{jt}+\xi_{jt}$  follows some parametric form (often linear).
- ► Same IV-GMM approach after transformation
- Examples:
  - $\bullet$  Plain Logit:  $\sigma_j^{-1}(\mathcal{S}_{\cdot t}) = \ln \mathfrak{s}_{jt} \ln \mathfrak{s}_{0t}$
  - $\bullet \ \ \text{Nested Logit:} \ \sigma_j^{-1}(\mathcal{S}_{\cdot t},\rho) = \ln \mathfrak{s}_{jt} \ln \mathfrak{s}_{0t} + \rho \ln \mathfrak{s}_{j|gt}$

# **Inversion: BLP (Random Coefficients)**

We can't solve for  $\delta_{it}$  directly this time.

$$\sigma_j(\pmb{\delta_{\rm t}}, \tilde{\theta}_2) = \int \frac{\exp[\delta_{jt} + \mu_{ijt}]}{1 + \sum_k \exp[\delta_{kt} + \mu_{ikt}]} f(\pmb{\mu_{\rm it}} | \tilde{\theta}_2)$$

- lacktriangleright This is a  $J \times J$  system of equations for each t.
- ▶ It is diagonally dominant (with outside good).
- $\blacktriangleright$  There is a unique vector  $\pmb{\delta}_{\mathbf{t}}$  that solves it for each market t.
- lackbox If you can work out  $rac{\partial s_{jt}}{\partial \delta_{kt}}$  (easy) you can solve this using Newton's Method.

#### **Contraction: BLP**

BLP actually propose an easy solution to find  $\delta_t$ . Fix  $\widetilde{\theta_2}$  and solve for  $\delta_t$ . Think about doing this one market at a time:

$$\pmb{\delta}_{\mathbf{t}}^{(k)}(\tilde{\theta}_2) = \pmb{\delta}_{\mathbf{t}}^{(k-1)}(\tilde{\theta}_2) + \left[\log(\mathfrak{s}_j) - \log(\mathbf{s}_j(\pmb{\delta}_{\mathbf{t}}^{(k-1)}, \tilde{\theta}_2)\right]$$

- ▶ They prove (not easy) that this is a contraction mapping.
- $\begin{tabular}{l} \blacksquare \begin{tabular}{l} \blacksquare \begin$
- lacktriangle Practical tip:  $\epsilon_{tol}$  needs to be as small as possible. ( $pprox 10^{-13}$ ).
- lacktriangledown Practical tip: Contraction isn't as easy as it looks:  $\mathbf{s}_j(\pmb{\delta}_{\mathbf{t}}^{(k-1)}, \widetilde{ heta}_2)$  requires computing the numerical integral each time (either via quadrature or monte carlo).

#### **BLP Pseudocode**

From the outside, in:

 $\blacktriangleright$  Outer loop: search over nonlinear parameters  $\theta$  to minimize GMM objective:

$$\widehat{\theta_{BLP}} = \arg\min_{\theta_2}(Z'\hat{\xi}(\theta_2))W(Z'\hat{\xi}(\theta_2))'$$

- ► Inner Loop:
  - Fix a guess of  $\tilde{\theta}_2$ .
  - $\bullet \ \ \text{Solve for } \pmb{\delta_{\mathbf{t}}}(\mathcal{S}_t, \tilde{\theta}_2) \text{ which satisfies } \sigma_{jt}(\pmb{\delta_{\mathbf{t}}}, \tilde{\theta}_2) = \mathfrak{s}_{jt}.$ 
    - lacksquare Computing  $s_{jt}(oldsymbol{\delta_t}, ilde{ heta}_2)$  requires numerical integration (quadrature or monte carlo).
  - We can do IV-GMM to recover  $\hat{\alpha}(\tilde{\theta}_2),\hat{\beta}(\tilde{\theta}_2),\hat{\xi}(\tilde{\theta}_2).$

$$\boldsymbol{\delta}_{\mathbf{t}}(\mathcal{S}_{t},\tilde{\boldsymbol{\theta}}_{2}) = \boldsymbol{x}_{jt}\boldsymbol{\beta} - \alpha \boldsymbol{p}_{jt} + \boldsymbol{\xi}_{jt}$$

- $\bullet \;\;$  Use  $\hat{\pmb{\xi}}(\theta)$  to construct sample moment conditions  $\frac{1}{N}\sum_{j,t}Z'_{jt}\xi_{jt}$
- $\blacktriangleright$  When we have found  $\hat{\theta}_{BLP}$  we can use this to update  $W\to W(\hat{\theta}_{BLP})$  and do 2-stage GMM.

# **Coming Up**

- ► Extensions and Variants
- ► Supply Side Restrictions
- ▶ Instruments
- ▶ Implementation Details