



# Properties of Joint Distributions II

Chris Piech  
CS109, Stanford University



# Titanic Probability

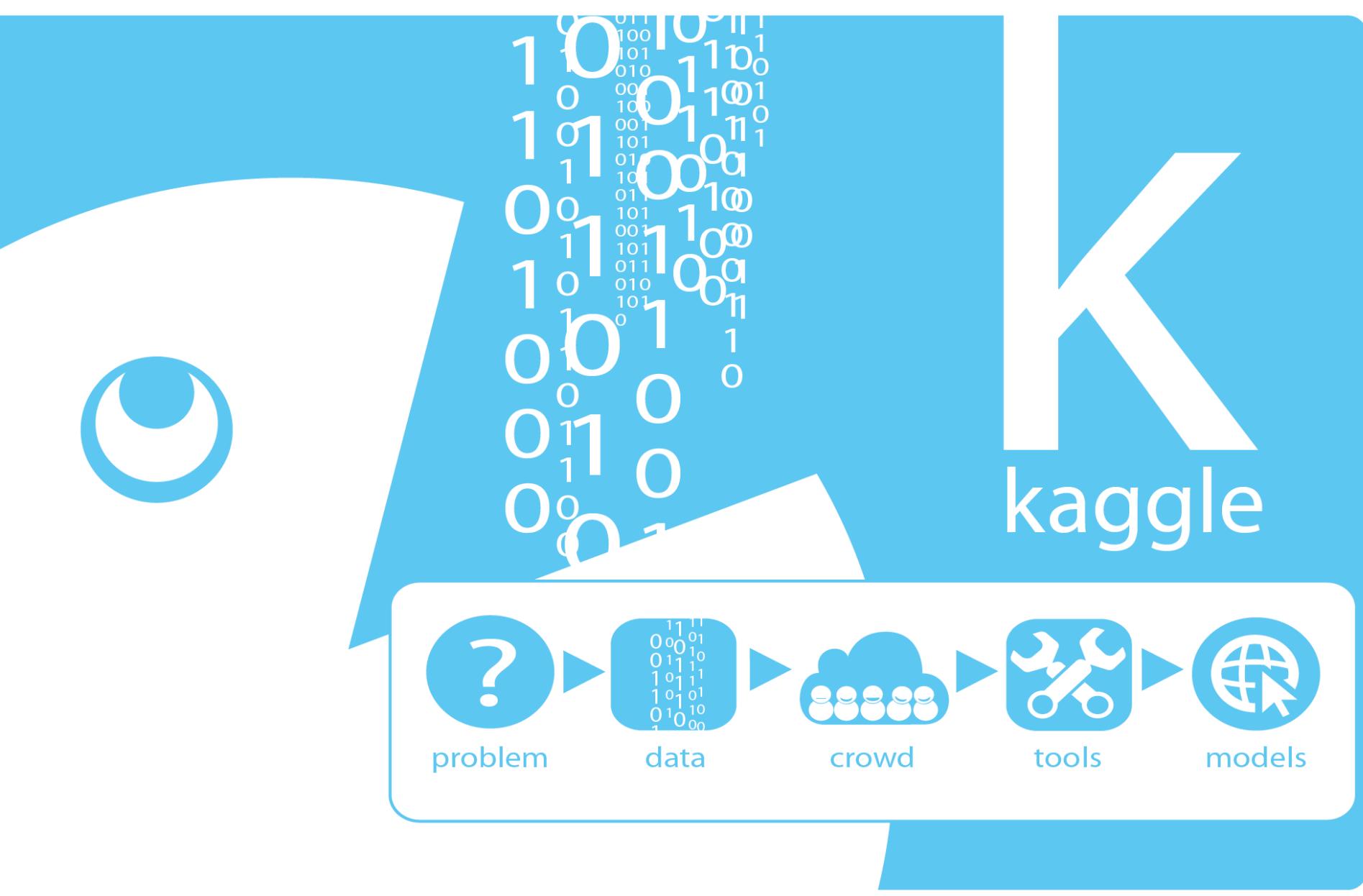
Screenshot of a web browser showing the Titanic dataset in CSV format, and an adjacent Microsoft Excel spreadsheet displaying the same data.

**CSV Data (titanic.csv):**

Survived	Pclass	Name	Sex	Age	Siblings/Spouses Aboard	Parents/Children Aboard	Fare	
0	3	Braund, Mr. Owen Harris	male	22	1	0	7.25	
1	1	Cumings, Mrs. John Bradley (Florence Thayer)	female	38	1	0	71.2833	
1	3	Heikkinen, Miss. Laina	female	26	0	0	7.925	
1	1	Futrelle, Mrs. Jacques Heath (Lily May Peel)	female	35	1	0	53.1	
0	3	Allen, Mr. William Henry	male	35	0	0	8.05	
0	3	Moran, Mr. James	male	27	0	0	8.4583	
0	1	McCarthy, Mr. Timothy J	male	54	0	0	51.8625	
0	3	Palsson, Master. Gosta Leonard	male	2	3	1	21.075	
1	3	Johnson, Mrs. Oscar W (Elisabeth Vilhelmina Berg)	female	27	0	0	8.4583	
1	2	Nasser, Mrs. Nicholas (Adele Achem)	female	14	1	0	30.0708	
1	3	Sandstrom, Miss. Marguerite Rut	female	4	1	0	30.0708	
1	1	Bonnell, Miss. Elizabeth	female	58	0	0	16.7	
0	3	Saunderscock, Mr. William Henry	male	20	0	0	8.05	
0	3	Andersson, Mr. Anders Johan	male	39	1	5	31.275	
0	3	Vestrom, Miss. Hulda Amanda Adolfina	female	14	0	0	7.8542	
1	2	Hewlett, Mrs. (Mary D Kingcome)	female	55	0	0	16	
0	3	Rice, Master. Eugene	male	2	4	1	29.125	
1	2	Williams, Mr. Charles Eugene	male	23	0	0	13	
0	3	Vander Planke, Mrs. Julius (Emelia Maria Vandemoortele)	female	31	1	0	18	
1	3	Masselmani, Mrs. Fatima	female	22	0	0	7.225	
0	2	Fynney, Mr. Joseph J	male	35	0	0	26	
1	2	Beesley, Mr. Lawrence	male	34	1	0	13	
24	1	McGowan, Miss. Anna "Annie"	female	15	0	0	8.0292	
25	1	Sloper, Mr. William Thompson	male	28	0	0	35.5	
26	0	3	Palsson, Miss. Torborg Danira	female	8	3	21.075	
27	1	Asplund, Mrs. Carl Oscar (Selma Augusta Emilia Johansson)	female	38	1	5	31.3875	
28	0	3	Emir, Mr. Farred Chehab	male	26	0	0	7.225
29	0	1	Fortune, Mr. Charles Alexander	male	19	3	2	263
30	1	O'Dwyer, Miss. Ellen "Nellie"	female	24	0	0	7.8792	
31	0	3	Todoroff, Mr. Lalio	male	23	0	0	7.8958
32	0	1	Uruchurtu, Don Manuel E	male	40	0	0	27.7208
33	1	Spencer, Mrs. William Augustus (Marie Eugenie)	female	48	1	0	146.5208	
34	1	Glynn, Miss. Mary Agatha	female	18	0	0	7.75	
35	0	2	Wheaton, Mr. Edward H	male	66	0	0	10.5
0	0	1	Meyer, Mr. Edgar Joseph	male	28	1	0	82.1708
0	2	Wheadon, Mr. Edward H	male	66	0	0	10.5	

**Excel Data:**

Survived	Pclass	Name	Sex	Age	Siblings/Spouses Aboard	Parents/Children Aboard	Fare
0	3	Braund, Mr. Owen Harris	male	22	1	0	7.25
1	1	Cumings, Mrs. John Bradley (Florence Thayer)	female	38	1	0	71.2833
1	3	Heikkinen, Miss. Laina	female	26	0	0	7.925
1	1	Futrelle, Mrs. Jacques Heath (Lily May Peel)	female	35	1	0	53.1
0	3	Allen, Mr. William Henry	male	35	0	0	8.05
0	3	Moran, Mr. James	male	27	0	0	8.4583
0	1	McCarthy, Mr. Timothy J	male	54	0	0	51.8625
0	3	Palsson, Master. Gosta Leonard	male	2	3	1	21.075



7% of passengers were  
from the Ottoman Empire

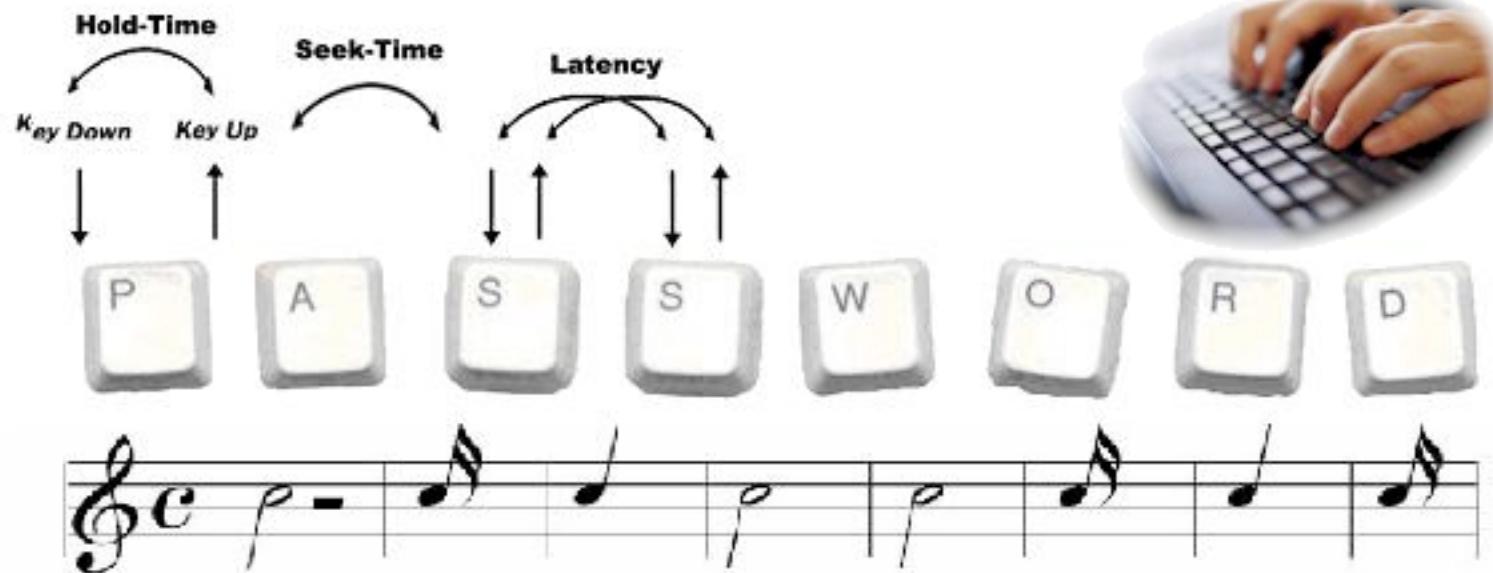
Ottoman Empire

Persian Empire

1912



# Biometric Keystrokes

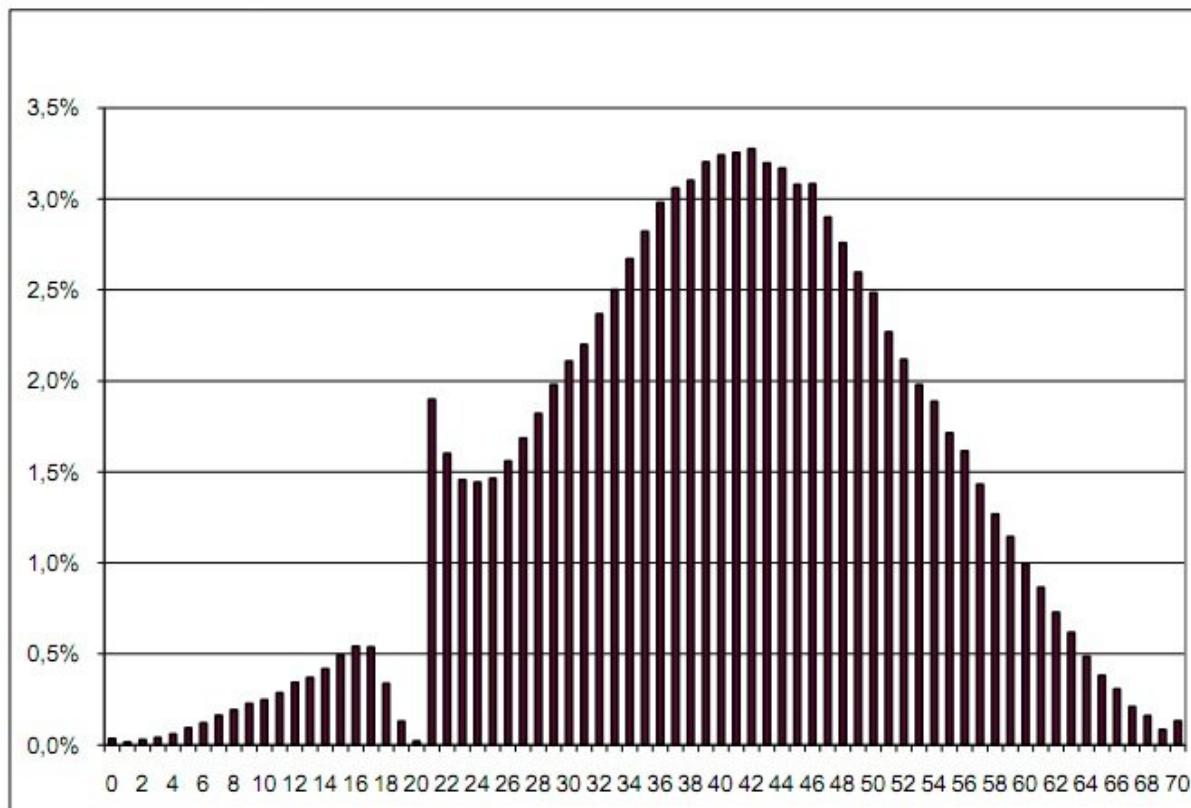


# Altruism?

Scores for a standardized test that students in Poland are required to pass before moving on in school

See if you can guess the minimum score to pass the test.

## 2.1. Poziom podstawowy



Wykres 1. Rozkład wyników na poziomie podstawowym

# Boolean Operation on Variable = Event

Recall: any boolean question about a random variable makes for an event. For example:



$$P(X \leq 5)$$

$$P(Y = 6)$$

$$P(5 \leq Z \leq 10)$$

# Joint Random Variables



Use a joint table, density function or CDF to solve probability question



Use and find **independence** of random variables



Think about **conditional** probabilities with joint variables (which might be continuous)



What happens when you **add** random variables?

# Independence and Random Variables

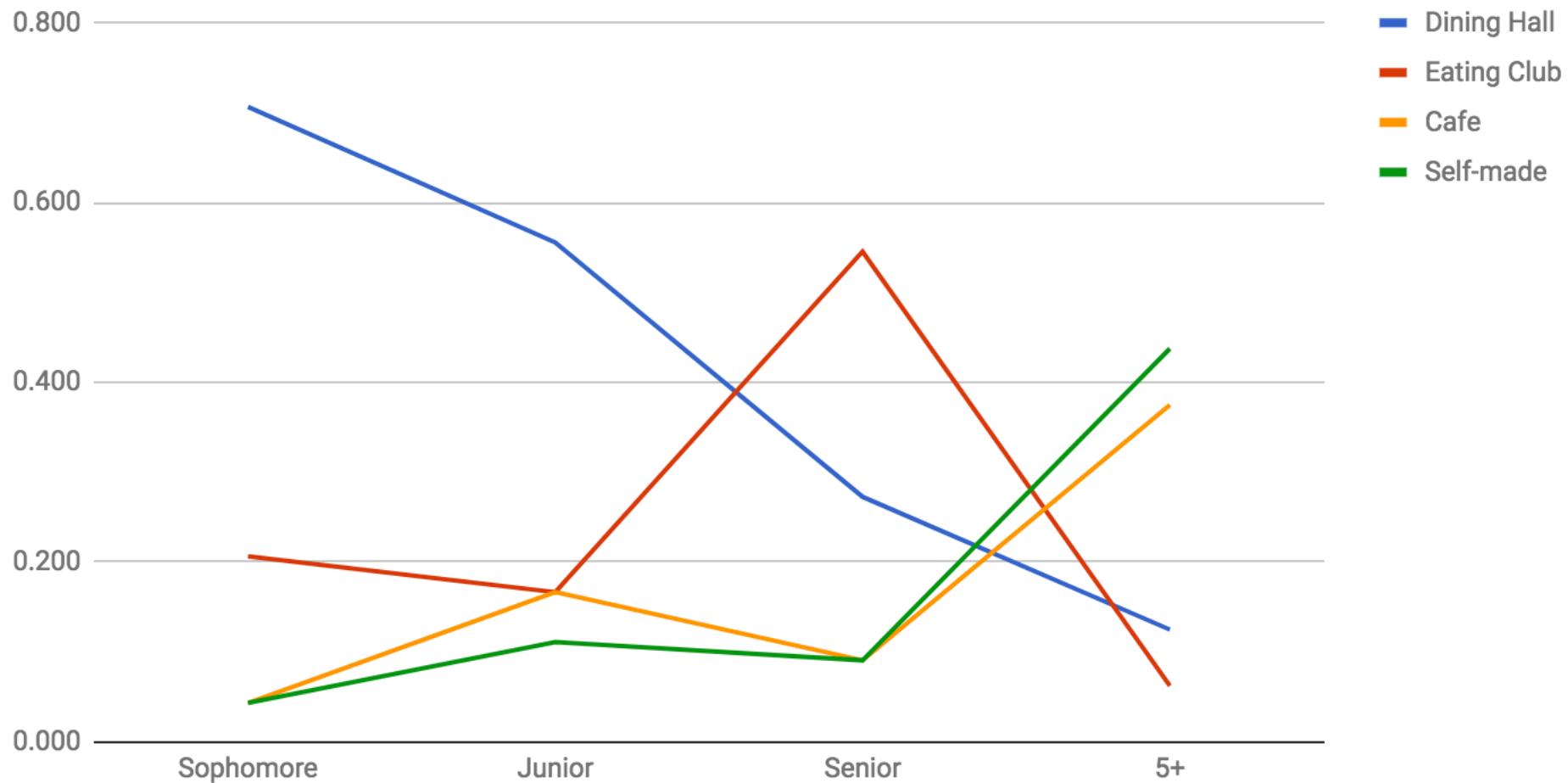
# Independent Continuous Variables

- Two continuous random variables X and Y are called independent if:  
$$P(X \leq a, Y \leq b) = P(X \leq a) P(Y \leq b) \text{ for any } a, b$$
- Equivalently:  
$$F_{X,Y}(a,b) = F_X(a)F_Y(b) \text{ for all } a,b$$
  
$$f_{X,Y}(a,b) = f_X(a)f_Y(b) \text{ for all } a,b$$
- More generally, joint density factors separately:  
$$f_{X,Y}(x,y) = h(x)g(y) \text{ where } -\infty < x, y < \infty$$

# Conditionals with multiple variables

# Lunch | Year

Lunch Type | Year



# Continuous Conditional Distributions

Let X and Y be continuous random variables

$$P(X = x|Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

$$f_{X|Y}(x|y) \cdot \epsilon_x = \frac{f_{X,Y}(x, y) \cdot \epsilon_x \cdot \epsilon_y}{f_Y(y) \cdot \epsilon_y}$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$$

# Bayes with a mix

Let  $X$  be a continuous random variable

Let  $N$  be a discrete random variable

$$P(X = x|N = n) = \frac{P(N = n|X = x)P(X = x)}{P(N = n)}$$

$$P_{x|N}(x|n) = \frac{P_{N|X}(n|x)P_X(x)}{P_N(n)}$$

$$f_{x|N}(x|n) \cdot \epsilon_x = \frac{P_{N|X}(n|x)f_X(x) \cdot \epsilon_x}{P_N(n)}$$

$$f_{x|N}(x|n) = \frac{P_{N|X}(n|x)f_X(x)}{P_N(n)}$$

# All the Bayes Belong to Us

M,N are discrete. X, Y are continuous

OG Bayes

$$p_{M|N}(m|n) = \frac{P_{N|M}(n|m)p_M(m)}{p_N(n)}$$

Mix Bayes #1

$$f_{X|N}(x|n) = \frac{P_{N|X}(n|x)f_X(x)}{P_N(n)}$$

Mix Bayes #2

$$p_{N|X}(n|x) = \frac{f_{X|N}(x|n)p_N(n)}{f_X(x)}$$

All Continuous

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$$

# Let's Do an Example

- X and Y are continuous RVs with PDF:

$$f(x, y) = \begin{cases} \frac{12}{5} x(2 - x - y) & \text{where } 0 < x, y < 1 \\ 0 & \text{otherwise} \end{cases}$$



- Compute conditional density:  $f_{X|Y}(x | y)$

$$\begin{aligned} f_{X|Y}(x | y) &= \frac{f_{X,Y}(x, y)}{f_Y(y)} = \frac{f_{X,Y}(x, y)}{\int_0^1 f_{X,Y}(x, y) dx} \\ &= \frac{\frac{12}{5} x(2 - x - y)}{\int_0^1 \frac{12}{5} x(2 - x - y) dx} = \frac{x(2 - x - y)}{\int_0^1 x(2 - x - y) dx} = \frac{x(2 - x - y)}{\left[ x^2 - \frac{x^3}{3} - \frac{x^2 y}{2} \right]_0^1} \\ &= \frac{x(2 - x - y)}{\frac{2}{3} - \frac{y}{2}} = \frac{6x(2 - x - y)}{4 - 3y} \end{aligned}$$



# Warmup: Bayes Revisited

$$P(B|E) = \frac{P(E|B) P(B)}{P(E)}$$

Posterior belief

Likelihood of evidence

Prior belief

Normalization constant

The diagram illustrates the Bayes' Rule formula with handwritten-style annotations. The formula is  $P(B|E) = \frac{P(E|B) P(B)}{P(E)}$ . A purple arrow labeled "Posterior belief" points to the term  $P(B|E)$ . A green arrow labeled "Likelihood of evidence" points to the term  $P(E|B)$ . A blue arrow labeled "Prior belief" points to the term  $P(B)$ . A red arrow labeled "Normalization constant" points to the term  $P(E)$ .

# Warmup: Bivariate Normal

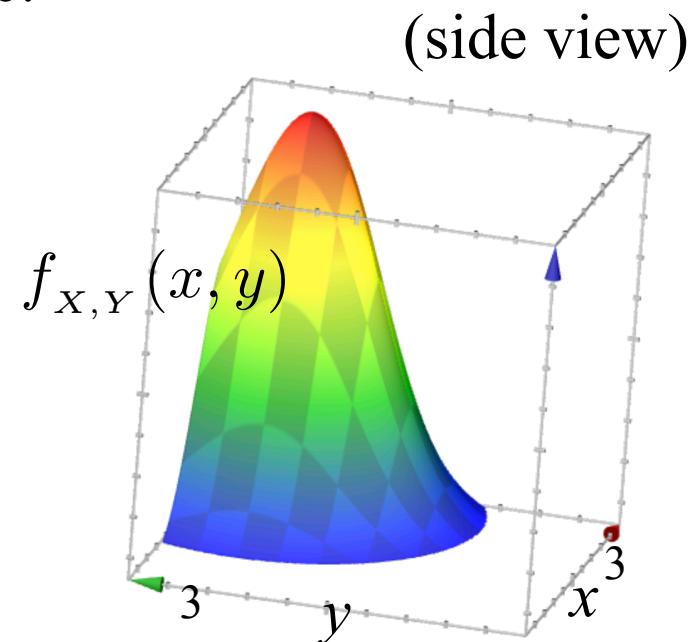
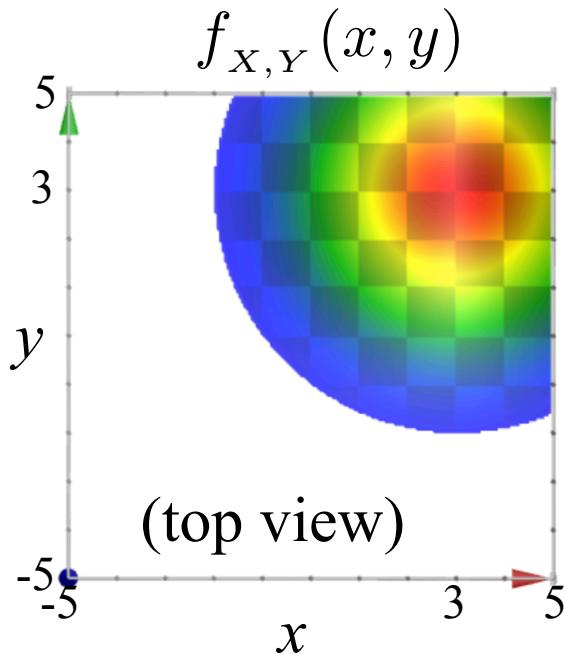
- $X, Y$  follow a symmetric bivariate normal distribution if they have joint PDF:

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma^2} \cdot e^{-\frac{[(x-\mu_x)^2 + (y-\mu_y)^2]}{2\cdot\sigma^2}}$$

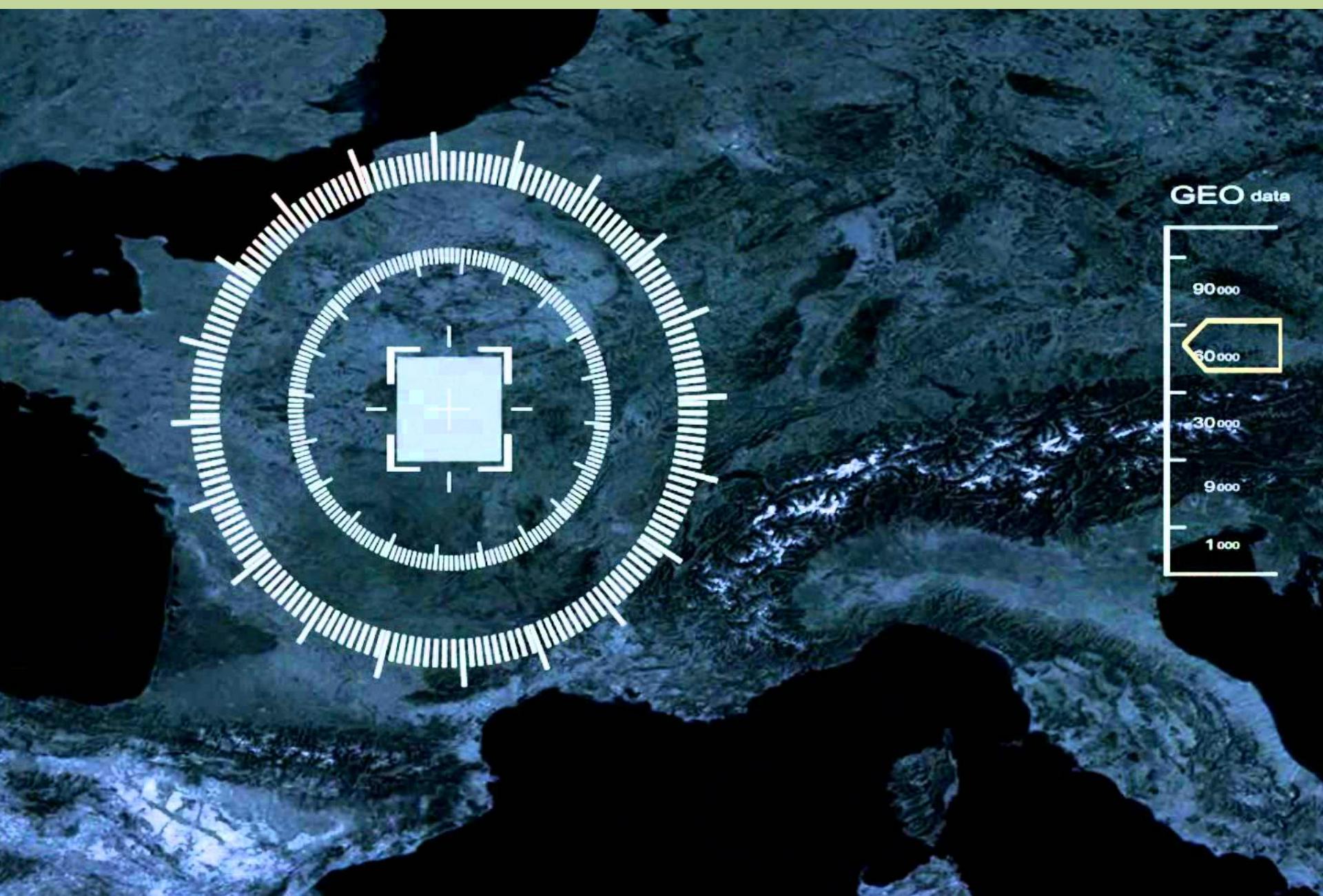
---

Here is an example where:

$$\begin{aligned}\mu_x &= 3 \\ \mu_y &= 3 \\ \sigma &= 2\end{aligned}$$

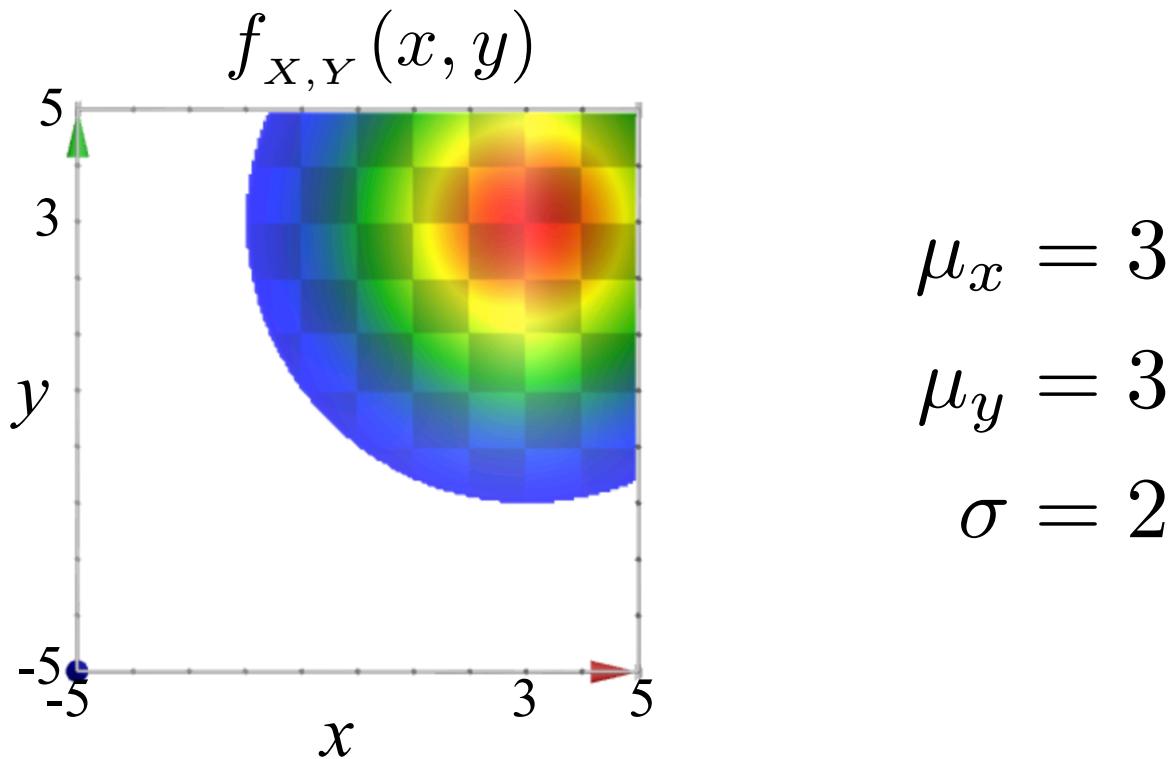


# Tracking in 2D Space?



# Tracking in 2D Space: Prior

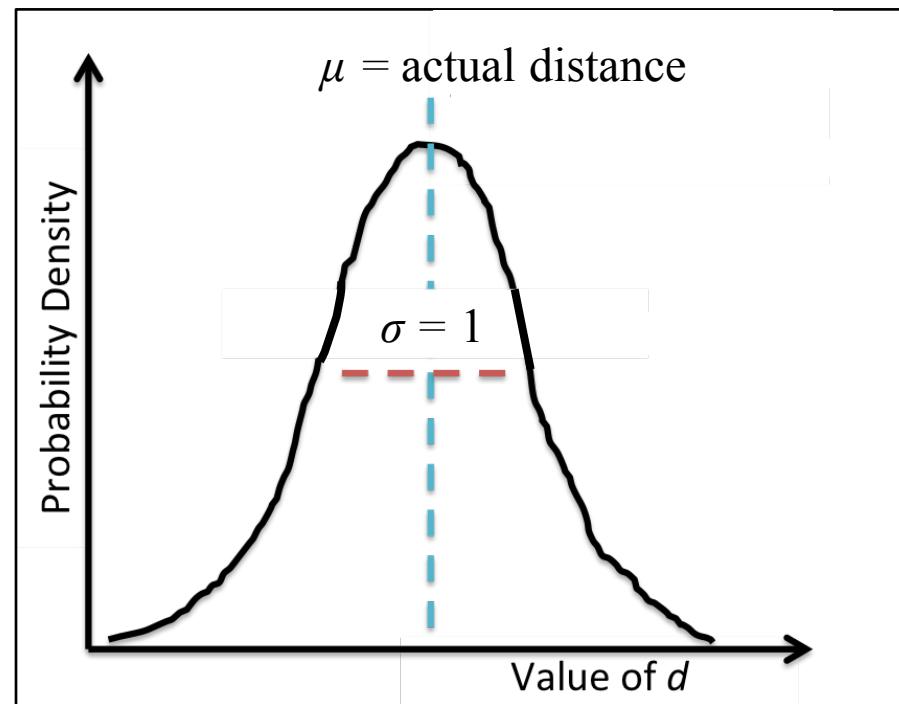
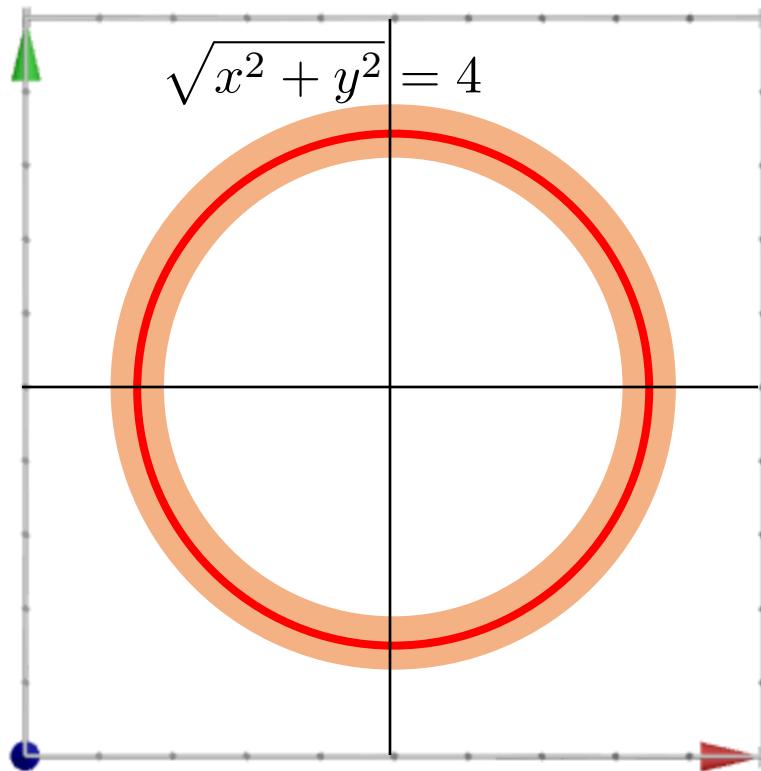
Prior belief:  $f_{X,Y}(x,y) = \frac{1}{2\pi\sigma^2} \cdot e^{-\frac{[(x-\mu_x)^2+(y-\mu_y)^2]}{2\cdot\sigma^2}}$



Prior belief with K:  $f_{X,Y}(x,y) = K \cdot e^{-\frac{[(x-3)^2+(y-3)^2]}{8}}$

# Tracking in 2D Space: Observation!

Observe a ping of the object that is distance  $D = 4$  away!

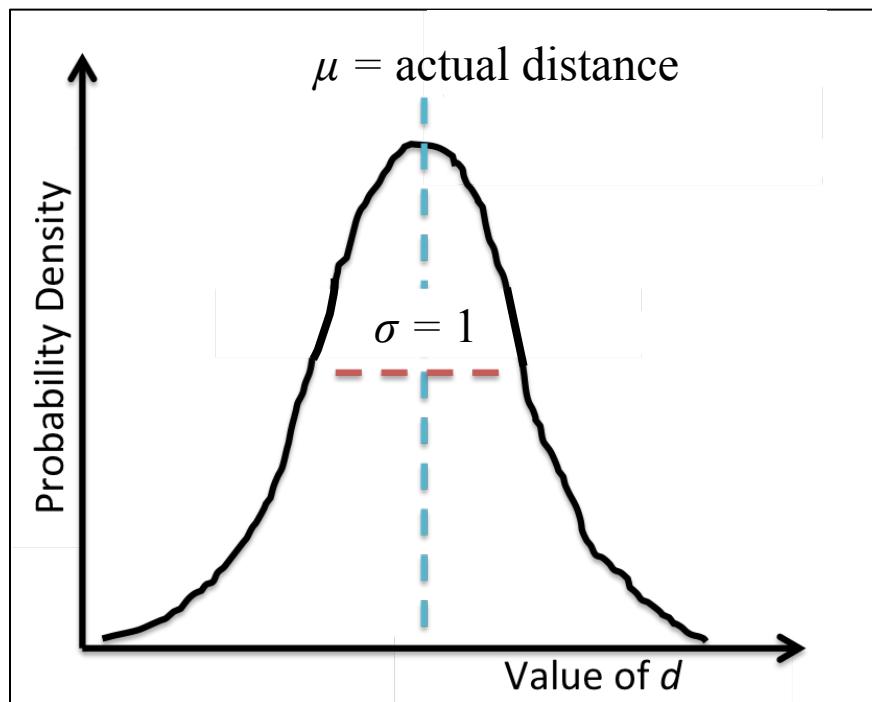


Know that the distance of a ping is normal with respect to the true distance

# Tracking in 2D Space: Observation!

Observe a ping of the object that is distance  $D = 4$  away!

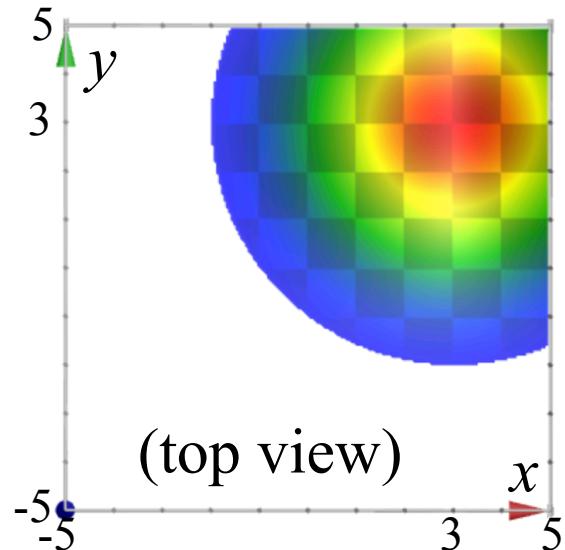
$$D|X, Y \sim N(\mu = \sqrt{x^2 + y^2}, \sigma^2 = 1)$$



Know that the distance of a ping is normal with respect to the true distance. What is the PDF of  $D$ ?

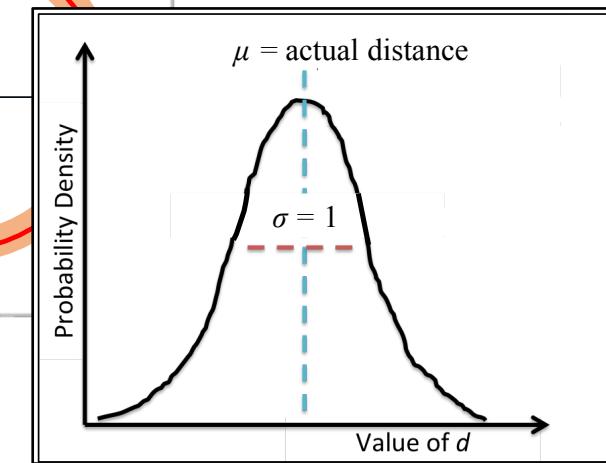
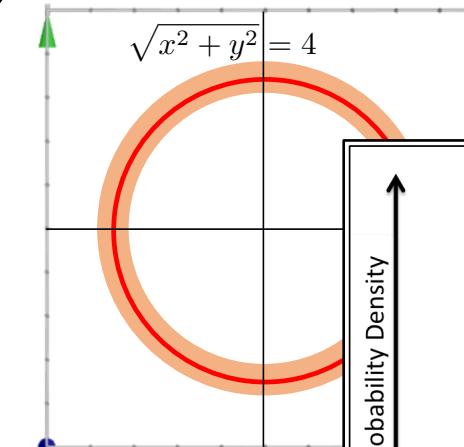
# Tracking in 2D Space: New Belief

$$f(X = x, Y = y) = K_1 \cdot e^{-\frac{[(x-3)^2 + (y-3)^2]}{8}}$$



Prior

Observation



$$f(D = d | X = x, Y = y) = K_2 \cdot e^{-\frac{[d - \sqrt{x^2 + y^2}]^2}{2}}$$

What is your *new* belief for the location of the object being tracked?  
Your joint probability density function can be expressed with a constant

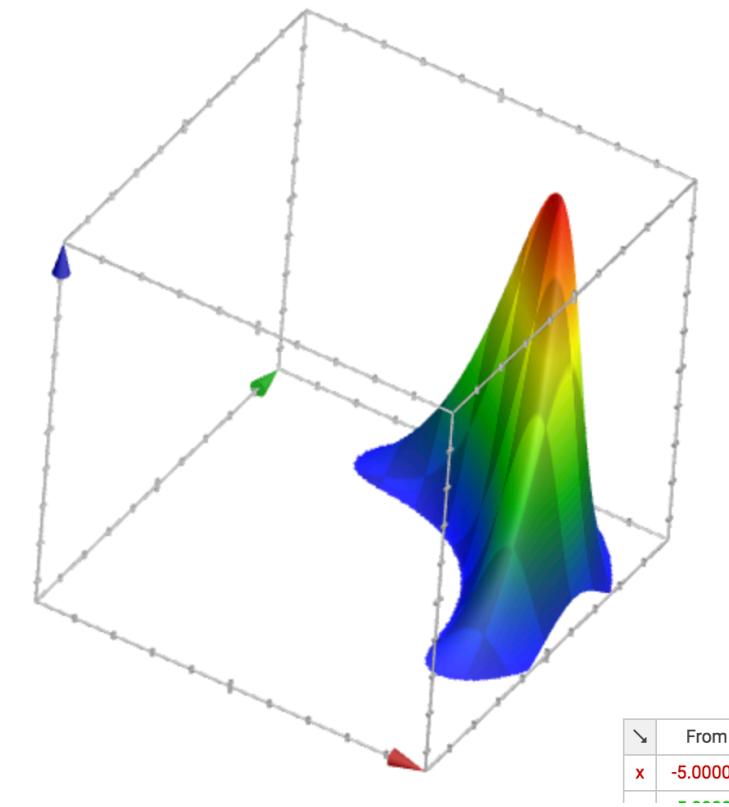
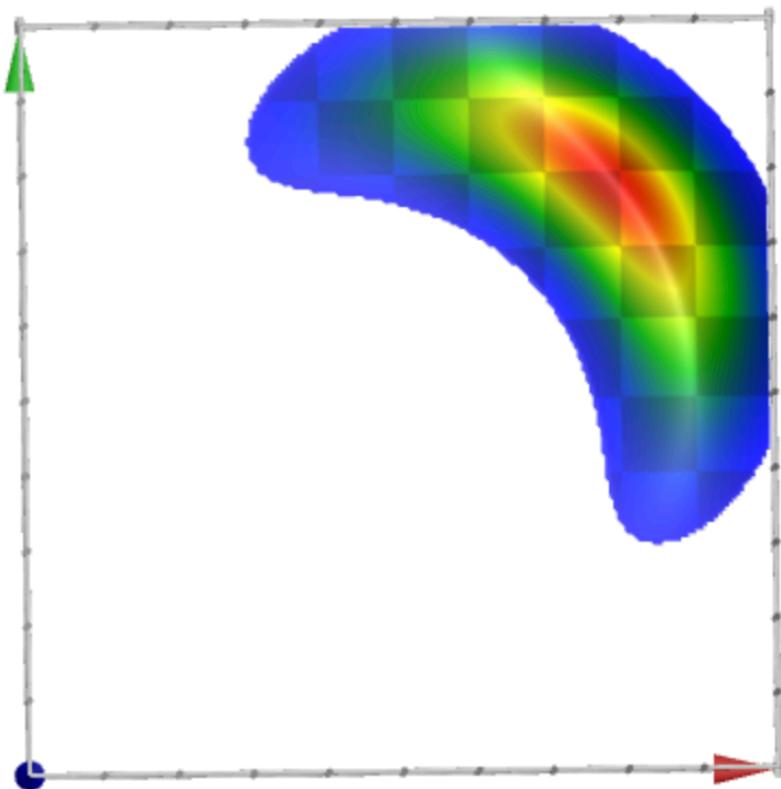
# Tracking in 2D Space: New Belief

$$\begin{aligned} f(X = x, Y = y | D = 4) &= \frac{f(D = 4 | X = x, Y = y) \cdot f(X = x, Y = y)}{f(D = 4)} \\ &= \frac{K_1 \cdot e^{-\frac{[4 - \sqrt{x^2 + y^2})^2]}{2}} \cdot K_2 \cdot e^{-\frac{[(x-3)^2 + (y-3)^2]}{8}}}{f(D = 4)} \\ &= \frac{K_3 \cdot e^{-\left[\frac{[4 - \sqrt{x^2 + y^2})^2]}{2} + \frac{[(x-3)^2 + (y-3)^2]}{8}\right]}}{f(D = 4)} \\ &= K_4 \cdot e^{-\left[\frac{(4 - \sqrt{x^2 + y^2})^2}{2} + \frac{[(x-3)^2 + (y-3)^2]}{8}\right]} \end{aligned}$$

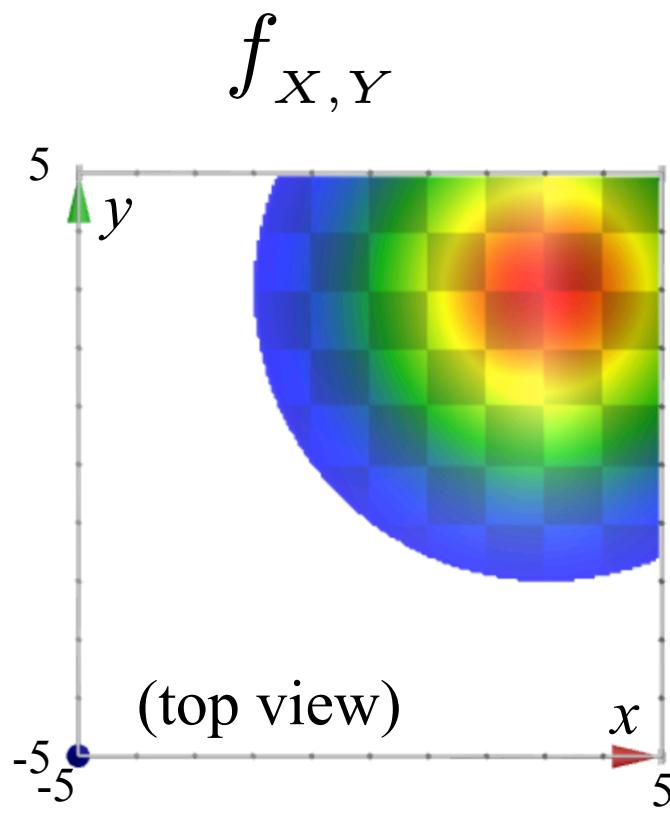
For your notes...

# Tracking in 2D Space: Posterior

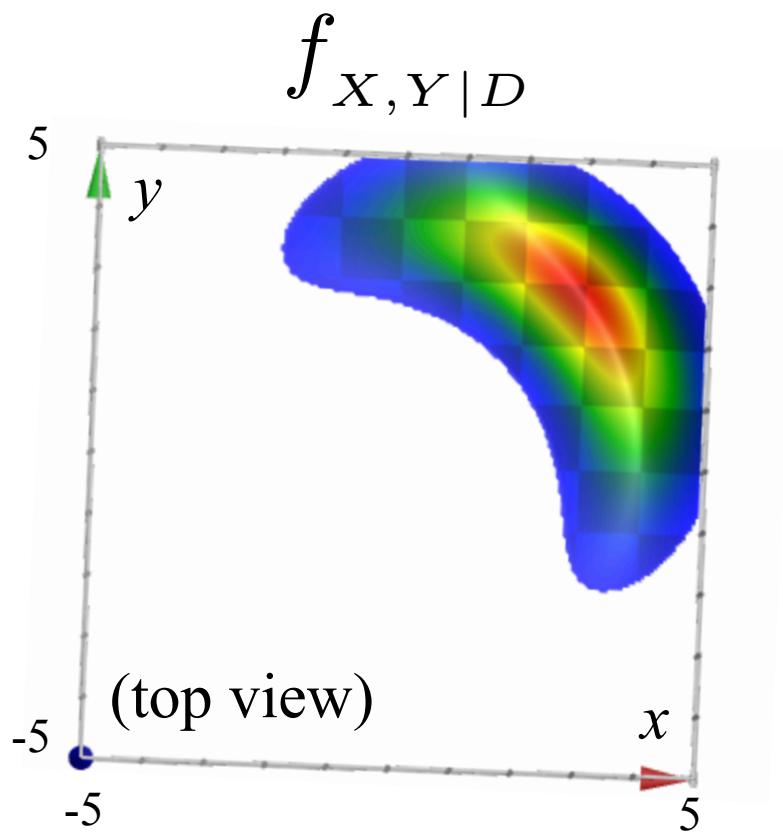
$$f_{x,y|D}(x, y|4) = K_4 \cdot e^{-\left[\frac{(4-\sqrt{x^2+y^2})^2}{2} + \frac{[(x-3)^2+(y-3)^2]}{8}\right]}$$



# Tracking in 2D Space: Posterior

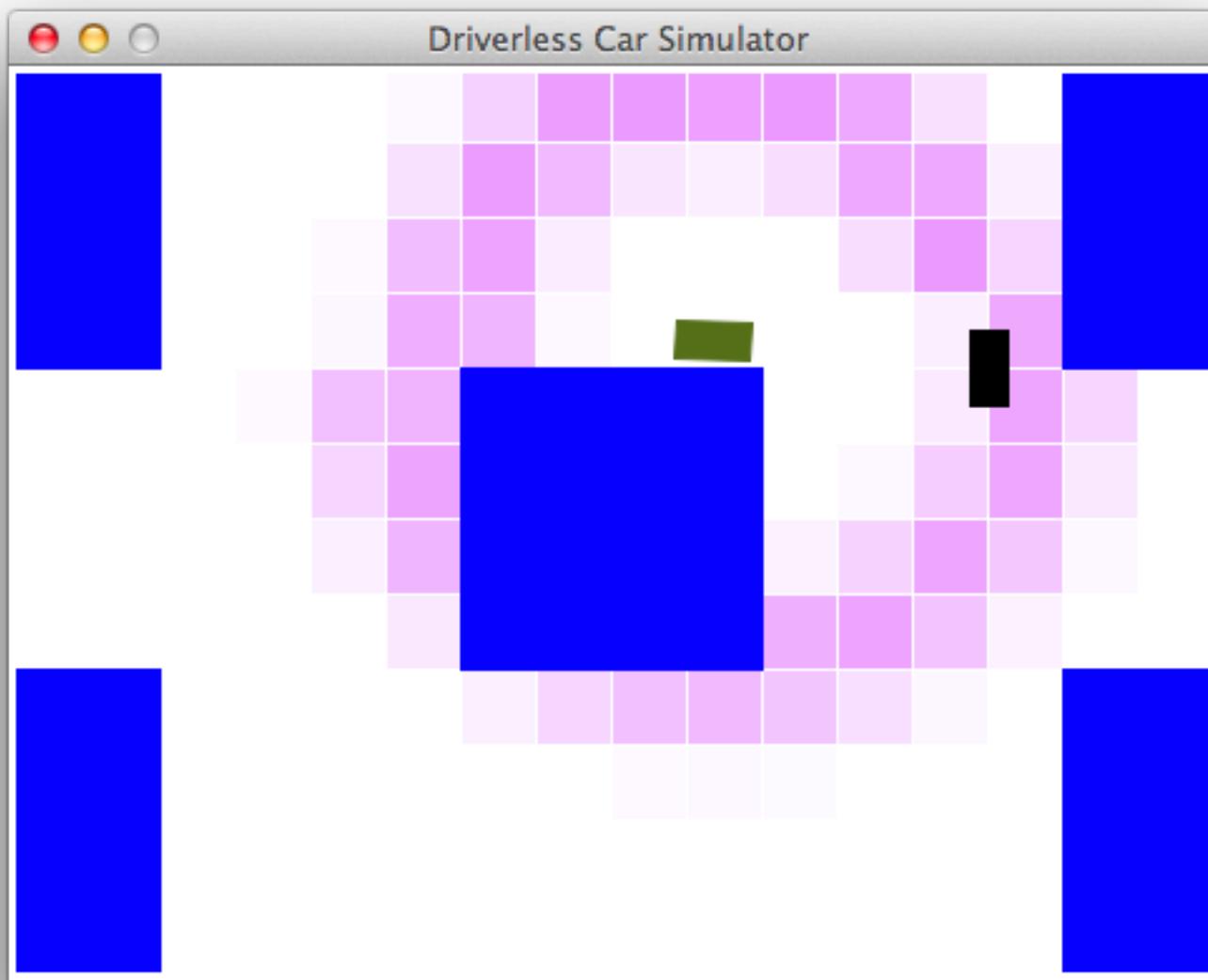


Prior



Posterior

# Tracking in 2D Space: CS221



What happens when you add random variables?

# Sum of Independent Binomials

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Bin}(n_1, p)$  and  $Y \sim \text{Bin}(n_2, p)$
  - $X + Y \sim \text{Bin}(n_1 + n_2, p)$
- Intuition:
  - $X$  has  $n_1$  trials and  $Y$  has  $n_2$  trials
    - Each trial has same “success” probability  $p$
  - Define  $Z$  to be  $n_1 + n_2$  trials, each with success prob.  $p$
  - $Z \sim \text{Bin}(n_1 + n_2, p)$ , and also  $Z = X + Y$

# Sum of Independent Poissons

First! Recall the Binomial Theorem

$$(a + b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$$

# Sum of Independent Poissons

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Poi}(\lambda_1)$  and  $Y \sim \text{Poi}(\lambda_2)$
  - $X + Y \sim \text{Poi}(\lambda_1 + \lambda_2)$
- Proof: (just for reference)
  - Rewrite  $(X + Y = n)$  as  $(X = k, Y = n - k)$  where  $0 \leq k \leq n$

$$P(X + Y = n) = \sum_{k=0}^n P(X = k, Y = n - k) = \sum_{k=0}^n P(X = k)P(Y = n - k)$$
$$= \sum_{k=0}^n e^{-\lambda_1} \frac{\lambda_1^k}{k!} e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!} = e^{-(\lambda_1 + \lambda_2)} \sum_{k=0}^n \frac{\lambda_1^k \lambda_2^{n-k}}{k!(n-k)!} = \frac{e^{-(\lambda_1 + \lambda_2)}}{n!} \sum_{k=0}^n \frac{n!}{k!(n-k)!} \lambda_1^k \lambda_2^{n-k}$$

- Noting Binomial theorem:  $(\lambda_1 + \lambda_2)^n = \sum_{k=0}^n \frac{n!}{k!(n-k)!} \lambda_1^k \lambda_2^{n-k}$
- $P(X + Y = n) = \frac{e^{-(\lambda_1 + \lambda_2)}}{n!} (\lambda_1 + \lambda_2)^n$  so,  $X + Y = n \sim \text{Poi}(\lambda_1 + \lambda_2)$

# Reference: Sum of Independent RVs

- Let  $X$  and  $Y$  be independent Binomial RVs
  - $X \sim \text{Bin}(n_1, p)$  and  $Y \sim \text{Bin}(n_2, p)$
  - $X + Y \sim \text{Bin}(n_1 + n_2, p)$
  - More generally, let  $X_i \sim \text{Bin}(n_i, p)$  for  $1 \leq i \leq N$ , then

$$\left( \sum_{i=1}^N X_i \right) \sim \text{Bin}\left( \sum_{i=1}^N n_i, p \right)$$

- Let  $X$  and  $Y$  be independent Poisson RVs
  - $X \sim \text{Poi}(\lambda_1)$  and  $Y \sim \text{Poi}(\lambda_2)$
  - $X + Y \sim \text{Poi}(\lambda_1 + \lambda_2)$
  - More generally, let  $X_i \sim \text{Poi}(\lambda_i)$  for  $1 \leq i \leq N$ , then

$$\left( \sum_{i=1}^N X_i \right) \sim \text{Poi}\left( \sum_{i=1}^N \lambda_i \right)$$

If only it were always that simple

# Convolution of Probability Distributions



We talked about sum of Binomial and Poisson...who's missing from this party?  
Uniform.

Summation: not just for the 1%

# Dance, Dance Convolution

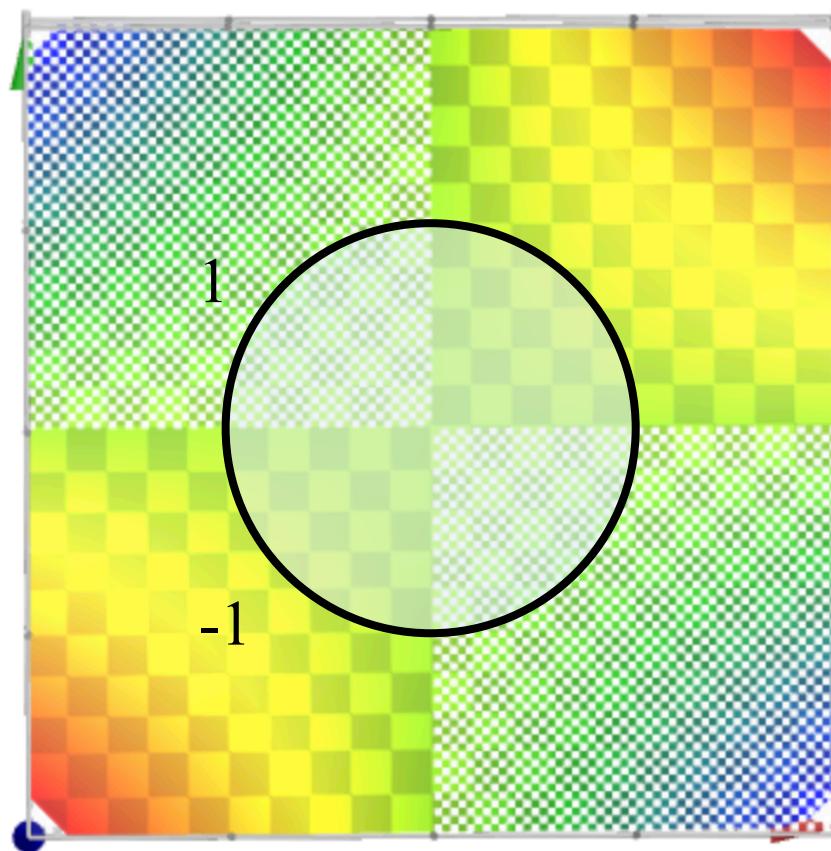
- Let  $X$  and  $Y$  be independent random variables
  - Probability Density Function (PDF) of  $X + Y$ , analogous:

$$f_{X+Y}(a) = \int_{y=-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$

- In discrete case, replace  $\int_{y=-\infty}^{\infty}$  with  $\sum_y$ , and  $f(y)$  with  $p(y)$

# Integration with Constraint

$$\iint_{x^2+y^2<1} f_{x,y} \ dy \ dx = \int_{-1}^1 \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} f_{x,y} \ dy \ dx$$



# Dance, Dance Convolution

- Let  $X$  and  $Y$  be independent random variables
  - Cumulative Distribution Function (CDF) of  $X + Y$ :

$$\begin{aligned} F_{X+Y}(a) &= P(X + Y \leq a) \\ &= \iint_{x+y \leq a} f_X(x) f_Y(y) dx dy = \int_{y=-\infty}^{\infty} \int_{x=-\infty}^{a-y} f_X(x) dx f_Y(y) dy \\ &= \int_{y=-\infty}^{\infty} F_X(a - y) f_Y(y) dy \end{aligned}$$

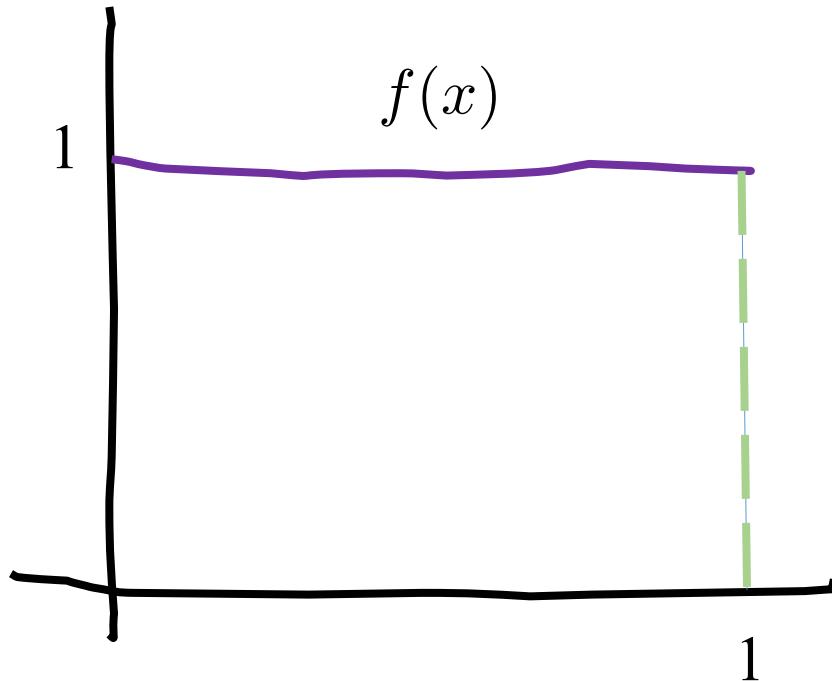
*CDF of  $X + Y$*

*PDF of  $Y$*

- In discrete case, replace  $\int$  with  $\sum$ , and  $f(y)$  with  $p(y)$

# Sum of Independent Uniforms

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Uni}(0, 1)$  and  $Y \sim \text{Uni}(0, 1) \rightarrow f(x) = 1$  for  $0 \leq x \leq 1$



For both  $X$  and  $Y$

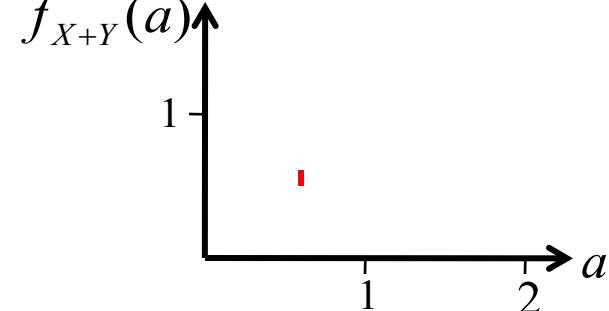
# Sum of Independent Uniforms

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Uni}(0, 1)$  and  $Y \sim \text{Uni}(0, 1) \rightarrow f(x) = 1$  for  $0 \leq x \leq 1$
  - What is PDF of  $X + Y$ ?

$$f_{X+Y}(a) = \int_{y=0}^1 f_X(a-y) f_Y(y) dy = \int_{y=0}^1 f_X(a-y) dy$$

When  $a = 0.5$ :

$$\begin{aligned} f_{X+Y}(0.5) &= \int_{y=?}^{y=?} f_X(0.5 - y) dy \\ &= \int_0^{0.5} f_X(0.5 - y) dy \\ &= \int_0^{0.5} 1 dy \\ &= 0.5 \end{aligned}$$



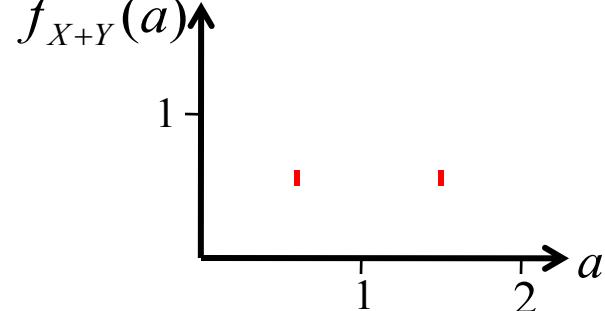
# Sum of Independent Uniforms

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Uni}(0, 1)$  and  $Y \sim \text{Uni}(0, 1) \rightarrow f(x) = 1$  for  $0 \leq x \leq 1$
  - What is PDF of  $X + Y$ ?

$$f_{X+Y}(a) = \int_{y=0}^1 f_X(a-y) f_Y(y) dy = \int_{y=0}^1 f_X(a-y) dy$$

When  $a = 1.5$ :

$$\begin{aligned} f_{X+Y}(1.5) &= \int_{y=?}^{y=?} f_X(1.5 - y) dy \\ &= \int_{0.5}^1 f_X(1.5 - y) dy \\ &= \int_{0.5}^1 1 dy \\ &= 0.5 \end{aligned}$$



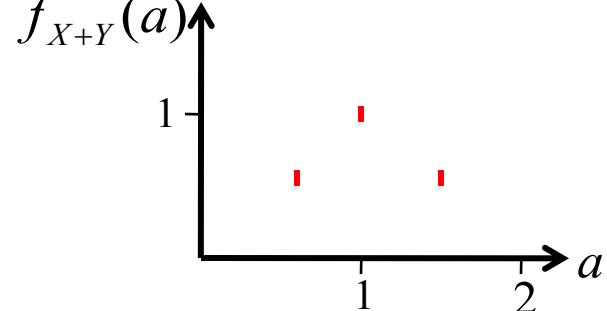
# Sum of Independent Uniforms

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Uni}(0, 1)$  and  $Y \sim \text{Uni}(0, 1) \rightarrow f(x) = 1$  for  $0 \leq x \leq 1$
  - What is PDF of  $X + Y$ ?

$$f_{X+Y}(a) = \int_{y=0}^1 f_X(a-y) f_Y(y) dy = \int_{y=0}^1 f_X(a-y) dy$$

When  $a = 1$ :

$$\begin{aligned} f_{X+Y}(1) &= \int_{y=?}^{y=?} f_X(1-y) dy \\ &= \int_0^1 f_X(1-y) dy \\ &= \int_0^1 1 dy \\ &= 1 \end{aligned}$$



# Sum of Independent Uniforms

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim \text{Uni}(0, 1)$  and  $Y \sim \text{Uni}(0, 1) \rightarrow f(x) = 1$  for  $0 \leq x \leq 1$
  - What is PDF of  $X + Y$ ?

$$f_{X+Y}(a) = \int_{y=0}^1 f_X(a-y) f_Y(y) dy = \int_{y=0}^1 f_X(a-y) dy$$

- When  $0 \leq a \leq 1$  and  $0 \leq y \leq a$ ,  $0 \leq a-y \leq 1 \rightarrow f_X(a-y) = 1$

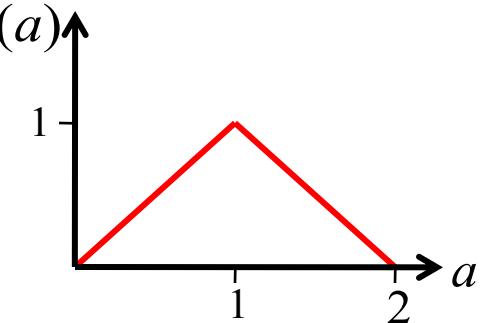
$$f_{X+Y}(a) = \int_{y=0}^a dy = a$$

- When  $1 \leq a \leq 2$  and  $a-1 \leq y \leq 1$ ,  $0 \leq a-y \leq 1 \rightarrow f_X(a-y) = 1$

$$f_{X+Y}(a) = \int_{y=a-1}^1 dy = 2-a$$

$$f_{X+Y}(a)$$

- Combining:  $f_{X+Y}(a) = \begin{cases} a & 0 \leq a \leq 1 \\ 2-a & 1 < a \leq 2 \\ 0 & \text{otherwise} \end{cases}$



# Sum of Independent Normals

- Let  $X$  and  $Y$  be independent random variables
  - $X \sim N(\mu_1, \sigma_1^2)$  and  $Y \sim N(\mu_2, \sigma_2^2)$
  - $X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$
- Generally, have  $n$  independent random variables  $X_i \sim N(\mu_i, \sigma_i^2)$  for  $i = 1, 2, \dots, n$ :

$$\left( \sum_{i=1}^n X_i \right) \sim N\left( \sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2 \right)$$

# Virus Infections

- Say you are working with the WHO to plan a response to the initial conditions of a virus:
  - Two exposed groups
  - P1: 50 people, each independently infected with  $p = 0.1$
  - P2: 100 people, each independently infected with  $p = 0.4$
  - Question: Probability of more than 40 infections?

**Sanity check:** Should we use the Binomial Sum-of-RVs shortcut?

- A. YES!
- B. NO!
- C. Other/none/more

# Virus Infections

- Say you are working with the WHO to plan a response to the initial conditions of a virus:
  - Two exposed groups
  - P1: 50 people, each independently infected with  $p = 0.1$
  - P2: 100 people, each independently infected with  $p = 0.4$
  - $A = \# \text{ infected in P1}$       $A \sim \text{Bin}(50, 0.1) \approx X \sim N(5, 4.5)$
  - $B = \# \text{ infected in P2}$       $B \sim \text{Bin}(100, 0.4) \approx Y \sim N(40, 24)$
  - What is  $P(\geq 40 \text{ people infected})?$
  - $P(A + B \geq 40) \approx P(X + Y \geq 39.5)$
  - $X + Y = W \sim N(5 + 40 = 45, 4.5 + 24 = 28.5)$

$$P(W \geq 39.5) = P\left(\frac{W - 45}{\sqrt{28.5}} > \frac{39.5 - 45}{\sqrt{28.5}}\right) = 1 - \Phi(-1.03) \approx 0.8485$$

# Linear Transform

$$X \sim N(\mu, \sigma^2)$$

$$Y = X + X = 2 \cdot X$$

$$Y \sim N(2\mu, 4\sigma^2)$$

---

$$Y = X + X = 2 \cdot X$$

$$X + X \sim N(\mu + \mu, \sigma^2 + \sigma^2)$$

$$Y \sim N(2\mu, 2\sigma^2)$$



$X$  is not  
independent of  $X$

End sum of independent vars