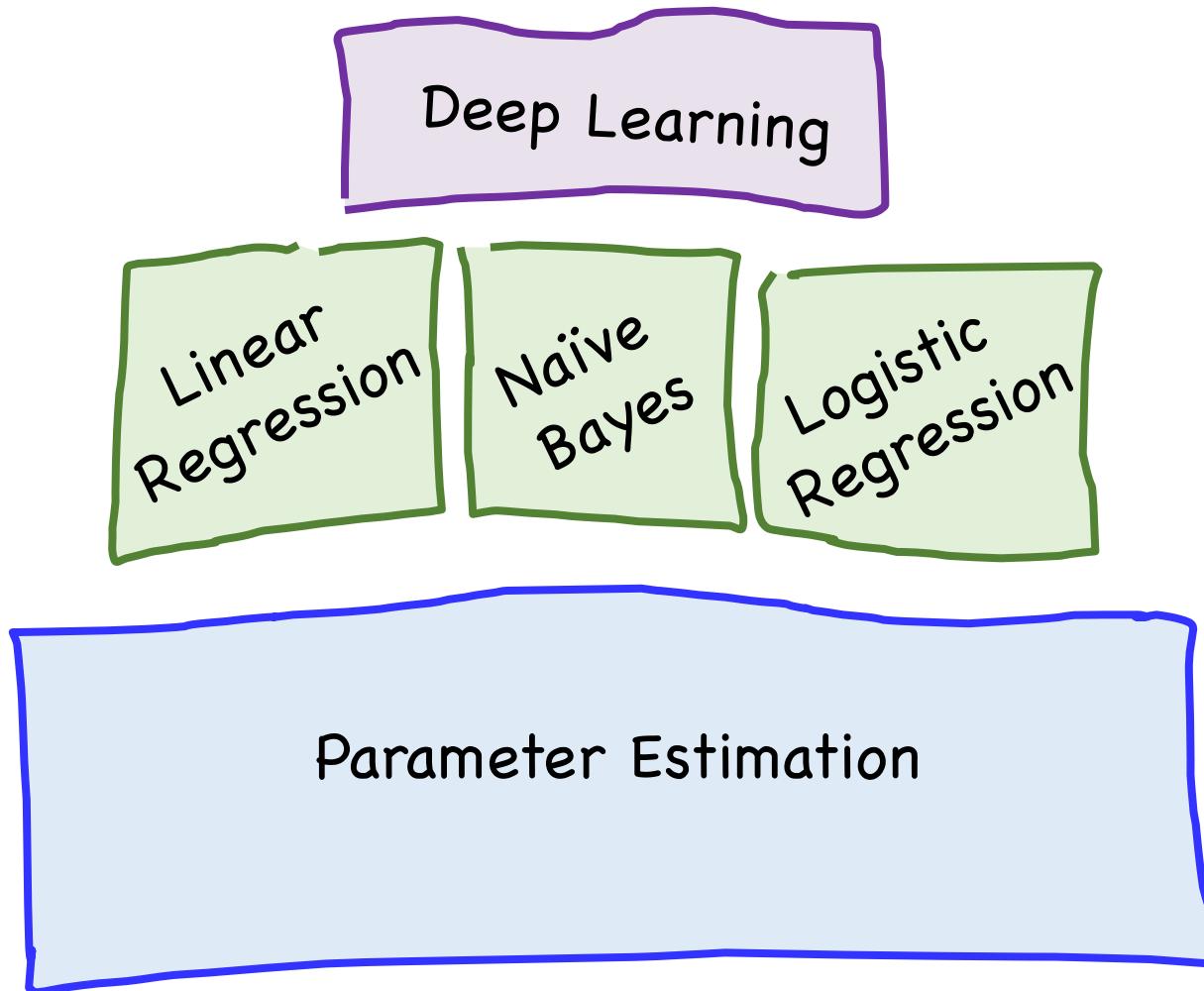




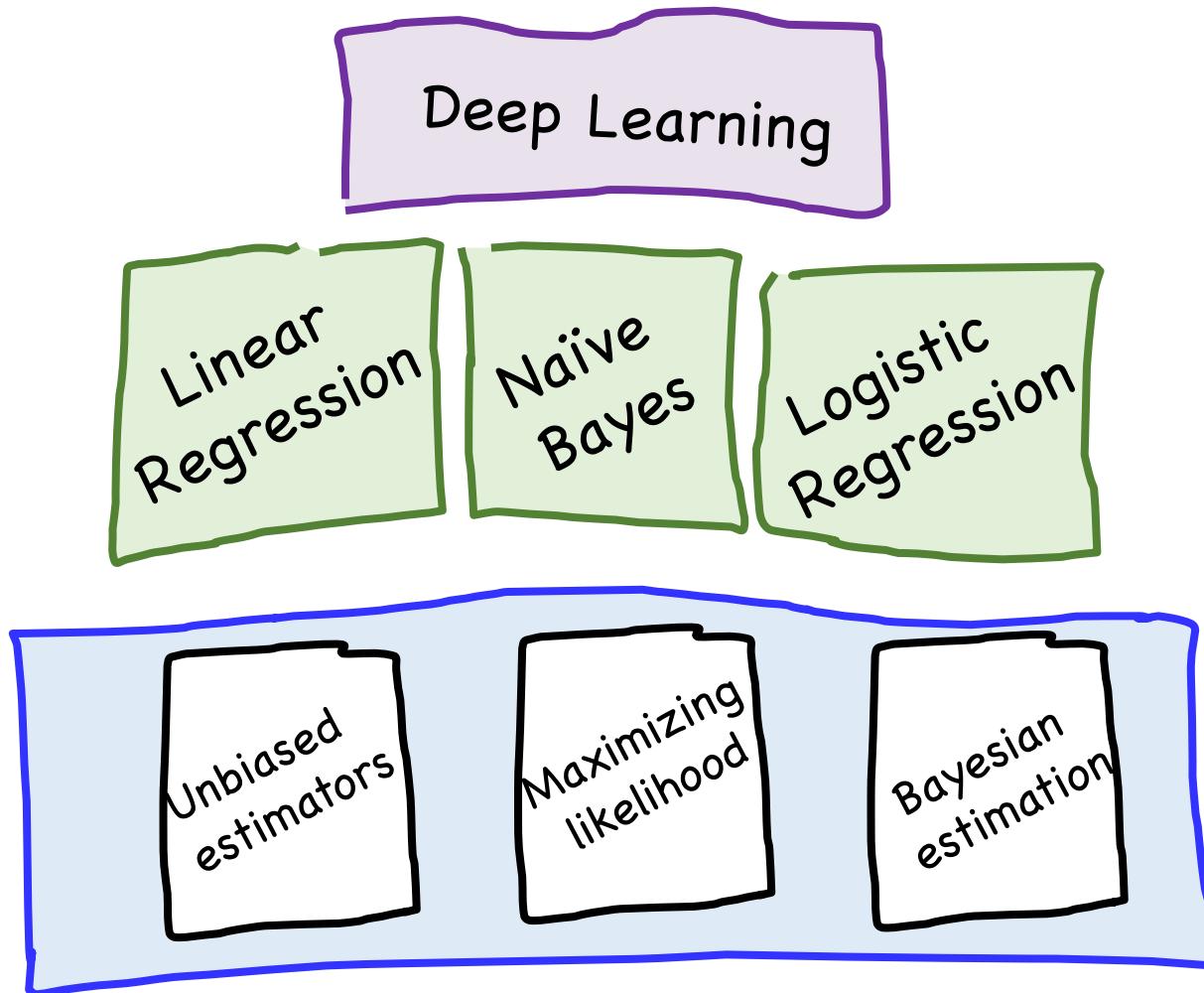
Gradient Ascent

Chris Piech
CS109, Stanford University

Our Path



Our Path





Parameter Learning

- Consider n I.I.D. random variables X_1, X_2, \dots, X_n
 - X_i is a sample from density function $f(X_i | \theta)$
 - What are the best choice of parameters θ ?



Likelihood (of data given parameters):

$$L(\theta) = \prod_{i=1}^n f(X_i \mid \theta)$$



Maximum Likelihood



$$L(\theta) = \prod_{i=1}^n f(X_i | \theta)$$

$$LL(\theta) = \sum_{i=1}^n \log f(X_i | \theta)$$

$$\hat{\theta} = \operatorname{argmax}_{\theta} LL(\theta)$$



Argmax?



Option #1: Straight optimization

Computing the MLE

- General approach for finding MLE of θ
 - Determine formula for $LL(\theta)$
 - Differentiate $LL(\theta)$ w.r.t. (each) θ : $\frac{\partial LL(\theta)}{\partial \theta}$
 - To maximize, set $\frac{\partial LL(\theta)}{\partial \theta} = 0$
 - Solve resulting (simultaneous) equations to get θ_{MLE}
 - Make sure that derived $\hat{\theta}_{MLE}$ is actually a maximum (and not a minimum or saddle point). E.g., check $LL(\theta_{MLE} \pm \varepsilon) < LL(\theta_{MLE})$
 - This step often ignored in expository derivations
 - So, we'll ignore it here too (and won't require it in this class)

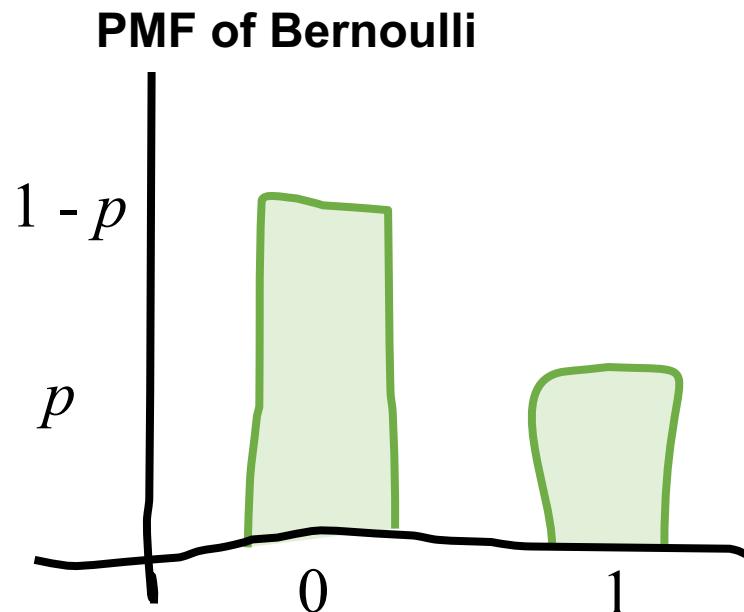
Maximizing Likelihood with Bernoulli

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Ber}(p)$
 - Probability mass function, $f(X_i | p)$:

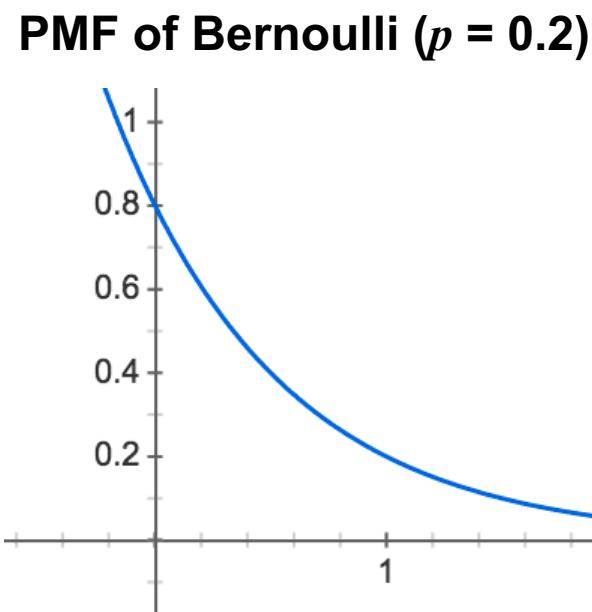


Maximizing Likelihood with Bernoulli

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Ber}(p)$
 - Probability mass function, $f(X_i | p)$:



$$f(X_i | p) = p^{x_i} (1-p)^{1-x_i}$$



$$f(x) = 0.2^x (1 - 0.2)^{1-x}$$

Bernoulli PMF

$$X \sim \text{Ber}(p)$$



$$f(X = x|p) = p^x(1 - p)^{1-x}$$

Maximizing Likelihood with Bernoulli

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Ber}(p)$
 - Probability mass function, $f(X_i | p)$, can be written as:

$$f(X_i | p) = p^{x_i} (1-p)^{1-x_i} \quad \text{where } x_i = 0 \text{ or } 1$$

- Likelihood: $L(\theta) = \prod_{i=1}^n p^{X_i} (1-p)^{1-X_i}$

- Log-likelihood:

$$\begin{aligned} LL(\theta) &= \sum_{i=1}^n \log(p^{X_i} (1-p)^{1-X_i}) = \sum_{i=1}^n [X_i (\log p) + (1-X_i) \log(1-p)] \\ &= Y(\log p) + (n-Y)\log(1-p) \quad \text{where } Y = \sum_{i=1}^n X_i \end{aligned}$$

- Differentiate w.r.t. p , and set to 0:

$$\frac{\partial LL(p)}{\partial p} = Y \frac{1}{p} + (n-Y) \frac{-1}{1-p} = 0 \quad \Rightarrow \quad p_{MLE} = \frac{Y}{n} = \frac{1}{n} \sum_{i=1}^n X_i$$

Isn't that the same as
unbiased estimator?

Yes. For Bernoulli.



Maximum Likelihood Algorithm

1. Decide on a model for the distribution of your samples. Define the PMF / PDF for your sample.

2. Write out the log likelihood function.

3. State that the optimal parameters are the argmax of the log likelihood function.

4. Use an optimization algorithm to calculate argmax



Maximizing Likelihood with Poisson

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Poi}(\lambda)$
 - PMF: $f(X_i | \lambda) = \frac{e^{-\lambda} \lambda^{x_i}}{x_i!}$ Likelihood: $L(\theta) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{X_i}}{X_i!}$
 - Log-likelihood:
$$\begin{aligned} LL(\theta) &= \sum_{i=1}^n \log\left(\frac{e^{-\lambda} \lambda^{X_i}}{X_i!}\right) = \sum_{i=1}^n [-\lambda \log(e) + X_i \log(\lambda) - \log(X_i!)] \\ &= -n\lambda + \log(\lambda) \sum_{i=1}^n X_i - \sum_{i=1}^n \log(X_i!) \end{aligned}$$
 - Differentiate w.r.t. λ , and set to 0:

$$\frac{\partial LL(\lambda)}{\partial \lambda} = -n + \frac{1}{\lambda} \sum_{i=1}^n X_i = 0 \quad \Rightarrow \quad \lambda_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i$$

Its so general!

Maximizing Likelihood with Normal

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim N(\mu, \sigma^2)$
 - PDF: $f(X_i | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-(X_i - \mu)^2 / (2\sigma^2)}$
 - Log-likelihood:

$$LL(\theta) = \sum_{i=1}^n \log\left(\frac{1}{\sqrt{2\pi}\sigma} e^{-(X_i - \mu)^2 / (2\sigma^2)}\right) = \sum_{i=1}^n \left[-\log(\sqrt{2\pi}\sigma) - (X_i - \mu)^2 / (2\sigma^2) \right]$$

- First, differentiate w.r.t. μ , and set to 0:

$$\frac{\partial LL(\mu, \sigma^2)}{\partial \mu} = \sum_{i=1}^n 2(X_i - \mu) / (2\sigma^2) = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0$$

- Then, differentiate w.r.t. σ , and set to 0:

$$\frac{\partial LL(\mu, \sigma^2)}{\partial \sigma} = \sum_{i=1}^n -\frac{1}{\sigma} + 2(X_i - \mu)^2 / (2\sigma^3) = -\frac{n}{\sigma} + \sum_{i=1}^n (X_i - \mu)^2 / (\sigma^3) = 0$$

Being Normal, Simultaneously

- Now have two equations, two unknowns:

$$\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0 \quad -\frac{n}{\sigma} + \sum_{i=1}^n (X_i - \mu)^2 / (\sigma^3) = 0$$

- First, solve for μ_{MLE} :

$$\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu) = 0 \Rightarrow \sum_{i=1}^n X_i = n\mu \Rightarrow \mu_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i$$

- Then, solve for σ^2_{MLE} :

$$-\frac{n}{\sigma} + \sum_{i=1}^n (X_i - \mu)^2 / (\sigma^3) = 0 \Rightarrow n\sigma^2 = \sum_{i=1}^n (X_i - \mu)^2$$

$$\sigma_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu_{MLE})^2$$

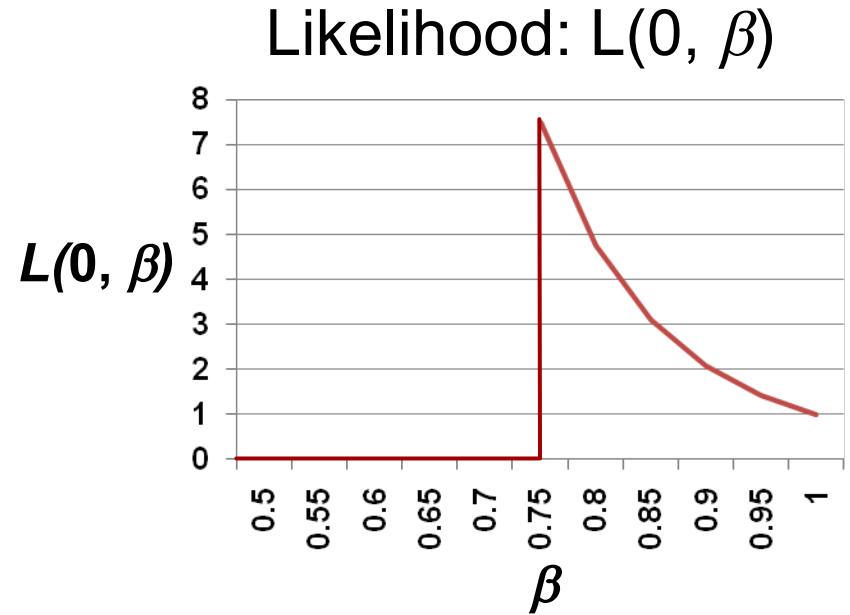
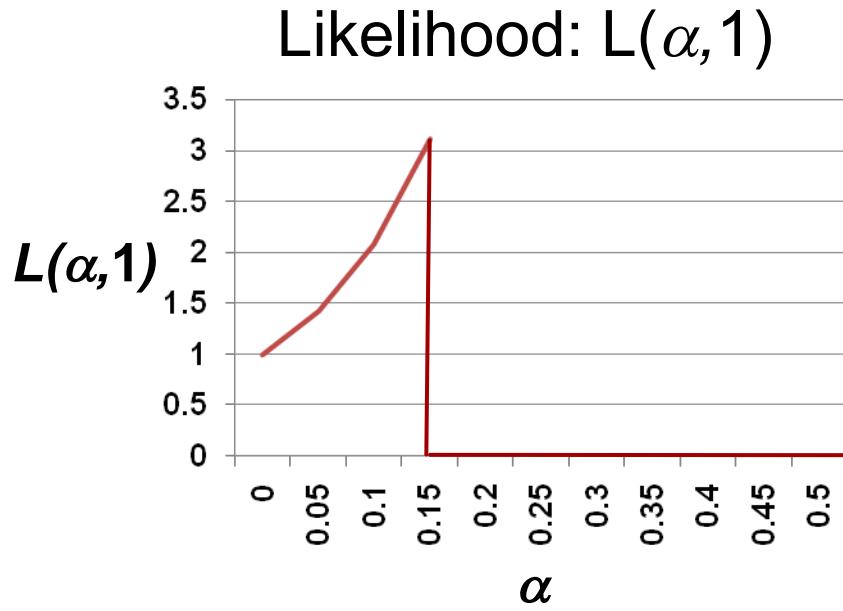
- Note: μ_{MLE} unbiased, but σ^2_{MLE} biased

Maximizing Likelihood with Uniform

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Uni}(\alpha, \beta)$
 - PDF: $f(X_i | \alpha, \beta) = \begin{cases} \frac{1}{\beta - \alpha} & \alpha \leq x_i \leq \beta \\ 0 & \text{otherwise} \end{cases}$
 - Likelihood: $L(\theta) = \begin{cases} \left(\frac{1}{\beta - \alpha}\right)^n & \alpha \leq x_1, x_2, \dots, x_n \leq \beta \\ 0 & \text{otherwise} \end{cases}$
 - Constraint $\alpha \leq x_1, x_2, \dots, x_n \leq \beta$ makes differentiation tricky
 - Intuition: want interval size $(\beta - \alpha)$ to be as small as possible to maximize likelihood function for each data point
 - But need to make sure all observed data contained in interval
 - If all observed data not in interval, then $L(\theta) = 0$
 - Solution: $\alpha_{MLE} = \min(x_1, \dots, x_n) \quad \beta_{MLE} = \max(x_1, \dots, x_n)$

Understanding MLE with Uniform

- Consider I.I.D. random variables X_1, X_2, \dots, X_n
 - $X_i \sim \text{Uni}(0, 1)$
 - Observe data:
 - 0.15, 0.20, 0.30, 0.40, 0.65, 0.70, 0.75



Small Samples = Problems

- How do small samples affect MLE?
 - In many cases, $\mu_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i$ = sample mean
 - Unbiased. Not too shabby...
 - As seen with Normal, $\sigma_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu_{MLE})^2$
 - Biased. Underestimates for small n (e.g., 0 for $n = 1$)
 - As seen with Uniform, $\alpha_{MLE} \geq \alpha$ and $\beta_{MLE} \leq \beta$
 - Biased. Problematic for small n (e.g., $\alpha = \beta$ when $n = 1$)
 - Small sample phenomena intuitively make sense:
 - Maximum likelihood \Rightarrow best explain data we've seen
 - Does not attempt to generalize to unseen data

Properties of MLE

- Maximum Likelihood Estimators are generally:
 - Consistent: $\lim_{n \rightarrow \infty} P(|\hat{\theta} - \theta| < \varepsilon) = 1$ for $\varepsilon > 0$
 - Potentially biased (though asymptotically less so)
 - Asymptotically optimal
 - Has smallest variance of “good” estimators for large samples
 - Often used in practice where sample size is large relative to parameter space
 - But be careful, there are some very large parameter spaces

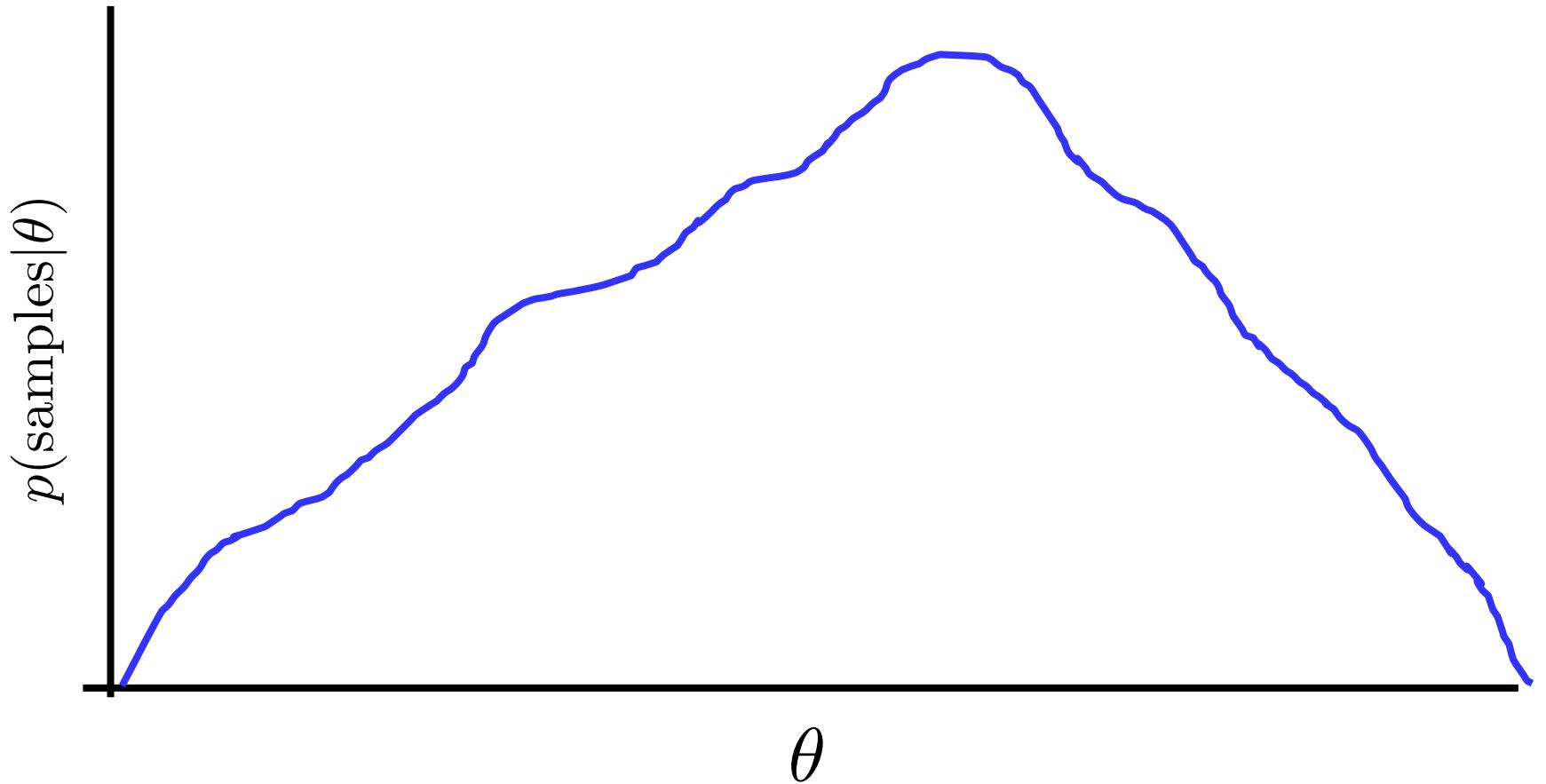


Argmax?



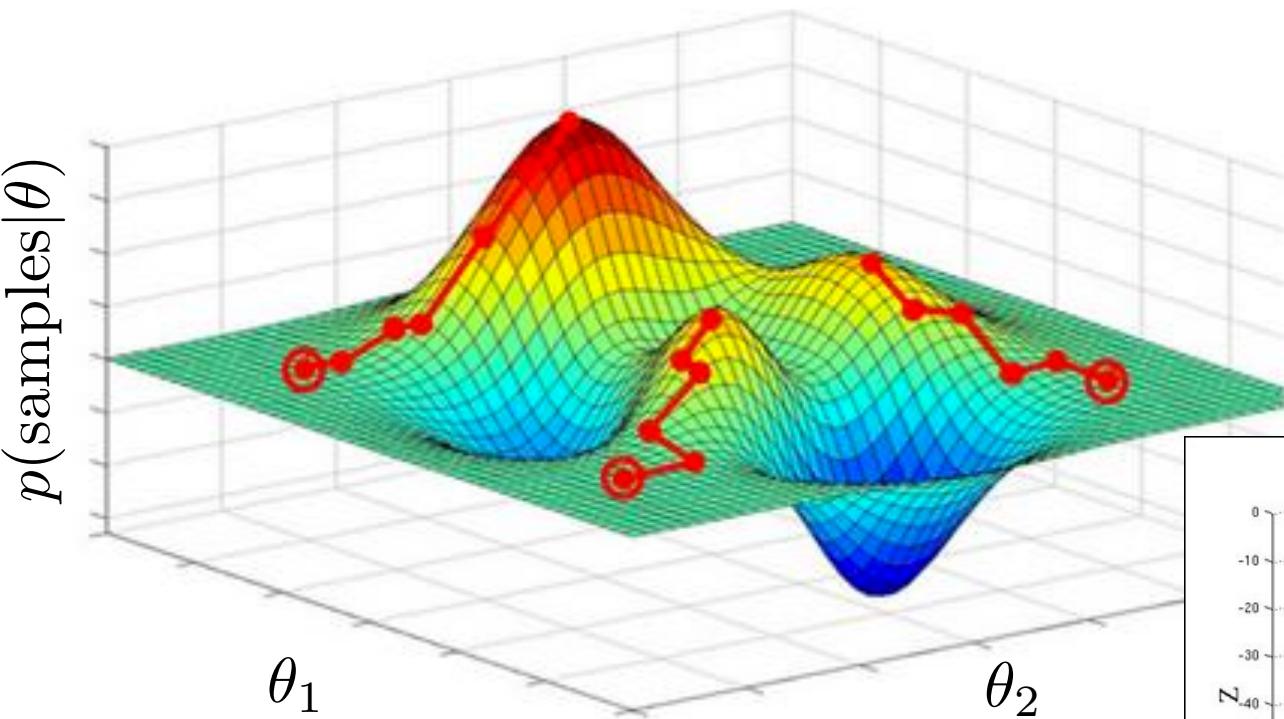
Option #2: Gradient Ascent

Gradient Ascent

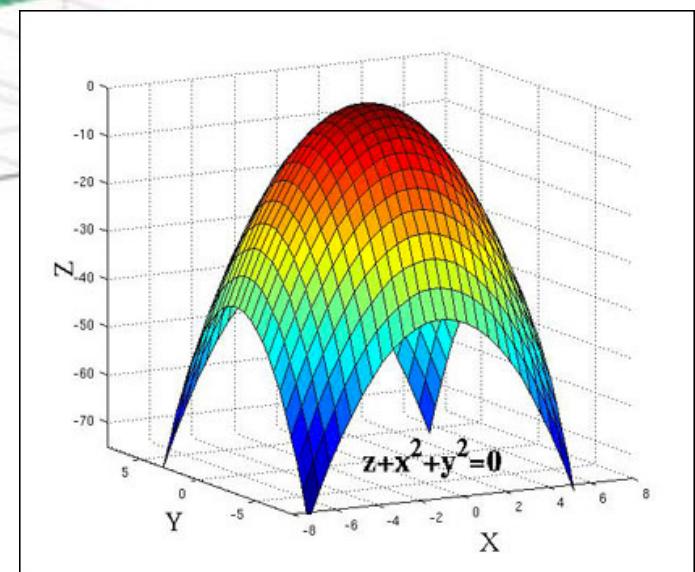


Walk uphill and you will find a local maxima
(if your step size is small enough)

Gradient Ascent



Especially good if
function is convex



Walk uphill and you will find a local maxima
(if your step size is small enough)



Gradient ascent is your
bread and butter
algorithm for optimization
(eg argmax)

Gradient Ascent

Initialize: $\theta_j = 0$ for all $0 \leq j \leq m$

Calculate all θ_j

Gradient Ascent

Initialize: $\theta_j = 0$ for all $0 \leq j \leq m$

Repeat many times:

gradient[j] = 0 for all $0 \leq j \leq m$

Calculate all gradient[j]'s based on data

$\theta_j += \eta * \text{gradient}[j]$ for all $0 \leq j \leq m$

Gradient Ascent

Initialize: $\theta_j = 0$ for all $0 \leq j \leq m$

Repeat many times:

gradient[j] = 0 for all $0 \leq j \leq m$

For each training example (x, y) :

For each parameter j :

Update gradient[j] for current training example

$\theta_j += \eta * \text{gradient}[j]$ for all $0 \leq j \leq m$

Linear Regression

Predicting Warriors

X_1 = Opposing team ELO

X_2 = Points in last game

X_3 = Curry playing?

X_4 = Playing at home?

Y = Warriors points

Predicting CO₂ (simple)

X = CO₂ level

Y = Average Global Temperature

N training datapoints

$$(\mathbf{x}^{(1)}, y^{(1)}), (\mathbf{x}^{(2)}, y^{(2)}), \dots (\mathbf{x}^{(n)}, y^{(n)})$$

Linear Regression Model

$$Y = \theta \cdot X + Z$$

$$Z \sim N(0, \sigma^2)$$

$$Y|X \sim N(\theta X, \sigma^2)$$

Linear Regression (simple)

Initialize: $\theta = 0$

Repeat many times:

gradient = 0

For each training example (x, y) :

Update gradient for current training example

$\theta += \eta * \text{gradient}$

Linear Regression (simple)

Initialize: $\theta = 0$

Repeat many times:

gradient = 0

For each training example (x, y) :

gradient += $(y - \theta x) (-x)$

$\theta += \eta * \text{gradient}$

Linear Regression

Predicting CO₂

X₁ = Temperature

X₂ = Elevation

X₃ = CO₂ level yesterday

X₄ = GDP of region

X₅ = Acres of forest growth

Y = CO₂ levels

Linear Regression

Problem: Predict real value Y based on observing variable X

Model: Linear weight every feature

$$\begin{aligned}\hat{Y} &= \theta_1 X_1 + \cdots + \theta_m X_m + \theta_{m+1} \\ &= \theta^T \mathbf{X}\end{aligned}$$

Training: Gradient ascent to chose the best thetas to describe your data

$$\hat{\theta}_{MLE} = \operatorname{argmax}_{\theta} - \sum_{i=1}^n (Y^{(i)} - \theta^T \mathbf{x}^{(i)})^2$$

Linear Regression

Initialize: $\theta_j = 0$ for all $0 \leq j \leq m$

Repeat many times:

gradient[j] = 0 for all $0 \leq j \leq m$

For each training example (x, y) :

For each parameter j :

gradient[j] += $(y - \theta^T x) (-x[j])$

$\theta_j += \eta * \text{gradient}[j]$ for all $0 \leq j \leq m$

Predicting Warriors

$Y = \text{Warriors points}$

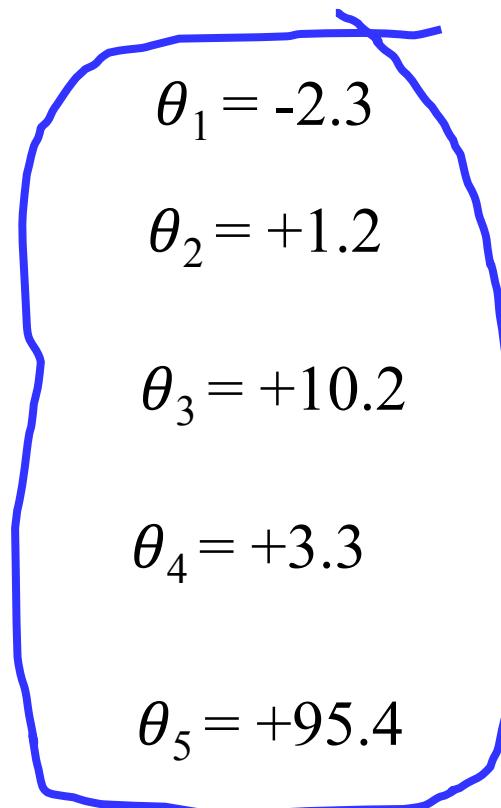
$$\begin{aligned}\hat{Y} &= \theta_1 X_1 + \cdots + \theta_m X_m + \theta_{m+1} \\ &= \theta^T \mathbf{X}\end{aligned}$$

$X_1 = \text{Opposing team ELO}$

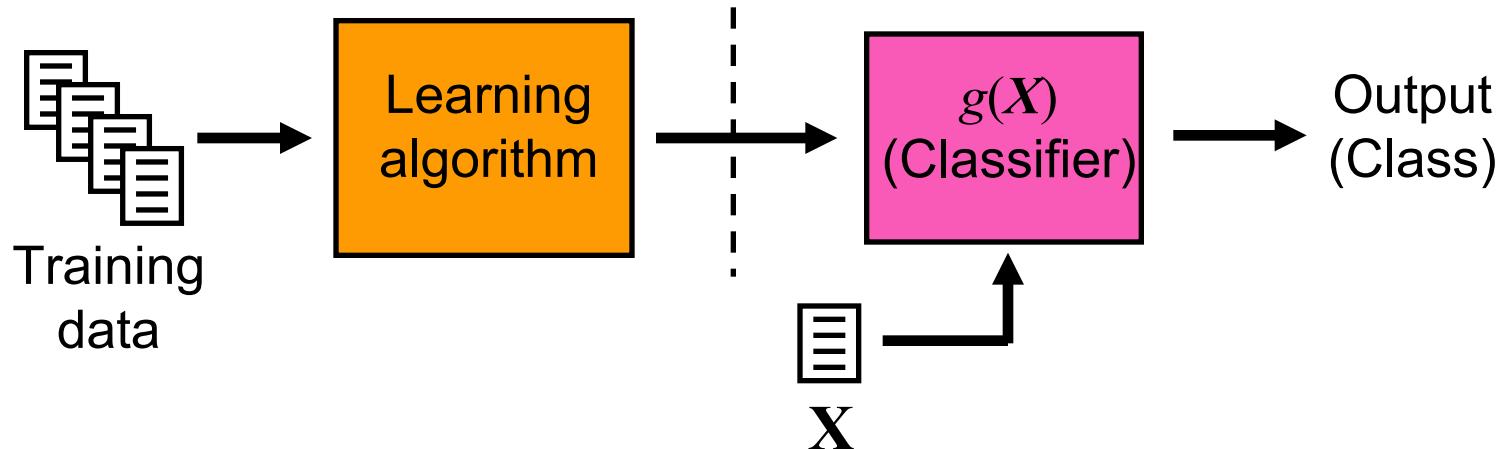
$X_2 = \text{Points in last game}$

$X_3 = \text{Curry playing?}$

$X_4 = \text{Playing at home?}$



The Machine Learning Process



- Training data: set of N pre-classified data instances
 - N training pairs: $(x^{(1)}, y^{(1)})$, $(x^{(2)}, y^{(2)})$, ..., $(x^{(n)}, y^{(n)})$
 - Use superscripts to denote i -th training instance
- Learning algorithm: method for determining $g(X)$
 - Given a new input observation of $x = x_1, x_2, \dots, x_m$
 - Use $g(x)$ to compute a corresponding output (prediction)