

1 Additions to the first observation of ionization cooling

1.1 Amplitude quantiles

Transverse 4D amplitudes distributed as

$$A_{\perp} \sim \epsilon_{\perp} \chi_4^2 = \frac{1}{4\epsilon_{\perp}^2} x e^{-x/2\epsilon_{\perp}}. \quad (1)$$

The α -quantile of the distribution, or α -amplitude, corresponds to

$$A_{\alpha} = \epsilon_{\perp} \chi_4^2(\alpha), \quad (2)$$

with $\chi_4^2(\alpha)$ the quantiles of the 4DOF chi-squared distribution. At $\alpha \simeq 9\%$, the probability content corresponds to that of the RMS ellipse and we naturally have $\chi_4^2(\alpha) = 1$ and $A_{\alpha} = \epsilon_{\perp}$.

For a distribution of covariance Σ , the volume of its RMS ellipse is simply $\mathcal{V} = \frac{1}{2}\pi^2|\Sigma|^{\frac{1}{2}}$. For a beam of emittance $\epsilon_{\perp} = |\Sigma|^{\frac{1}{4}}/m_{\mu}c$, the volume of the RMS ellipse thus reads

$$\mathcal{V} = \frac{1}{2}\pi^2 m^2 c^2 \epsilon_{\perp}^2. \quad (3)$$

The α -amplitude corresponds to an ellipse similar to the RMS ellipse of squared radius $A_{\alpha}/\epsilon_{\perp}$, so that the fractional emittance may be reconstructed as

$$\epsilon_{\alpha} = \frac{1}{2}\pi^2 m^2 c^2 A_{\alpha}^2. \quad (4)$$

The relative statistical uncertainty on the α -amplitude is identical to the chi-squared distribution quantile uncertainty, i.e.

$$\frac{\sigma_{A_{\alpha}}}{A_{\alpha}} = \epsilon_{\perp} \frac{\sigma_{\chi_4^2(\alpha)}}{\chi_4^2(\alpha)} = \sqrt{\frac{\alpha(1-\alpha)}{n}} [\chi_4^2(\alpha) f(4; \chi_4^2(\alpha))]^{-1}, \quad (5)$$

with $f(4; x)$ the 4DOF chi square distribution PDF in x . As the fractional emittance is proportional to the square of the α -amplitude, its statistical uncertainty satisfies

$$\frac{\sigma_{\epsilon_{\alpha}}}{\epsilon_{\alpha}} = 2 \frac{\sigma_{A_{\alpha}}}{A_{\alpha}}. \quad (6)$$

1.2 kNN density estimator

Take a set of n points, $\{\mathbf{x}\}_{i=1}^n$, in 4 dimensions. To evaluate the density in a point \mathbf{x} , find the squared distances to the n points, i.e.

$$R_i^2 = (\mathbf{x} - \mathbf{x}_i)^T \Sigma^{-1} (\mathbf{x} - \mathbf{x}_i), \quad (7)$$

with Σ the covariance matrix. If Σ is set to unity, it is the Euclidean distance. Find the k closest points and the distance R_k to the farthest one of them. The density in \mathbf{x} is evaluated as

$$\rho(\mathbf{x}) = \frac{k}{nV_k} = \frac{2k}{n\pi^2|\Sigma|^{\frac{1}{2}}R_k^4}, \quad (8)$$

with V_k the volume of the 4-ellipse of radius R_k . The optimal k in 4D is \sqrt{n} (see dissertation).

Cannot compare 4D profiles, must find 1D aggregate, I propose the “density profiles”, i.e. the quantiles of the inverse PDF. For a fraction α of the beam, find the density, ρ_{α} , above which lies a fraction α of the beam. This can be represented as a function of α to create the profiles.

The statistical uncertainty on the contours, ρ_α , have been showed to closely follow

$$\frac{\sigma_{\rho_\alpha}}{\rho_\alpha} = \left(\frac{4}{\alpha(1-\alpha)} \right)^{\frac{1}{4}} n^{-\frac{1}{2}}, \quad (9)$$

but a bootstrapped estimate is also provided in the code