Given probe model with sample space Ω , random variable (v.v.) assigns real number to each ontenne $\omega \in \Omega$, i.e. v.v. is a function from Ω to R.

r.v. X

For $w \in \mathcal{R}$, $X(w) = x \in \mathbb{R}$ event [X = x] $= [w \in \mathcal{R} | X(w) = x]$

Discrete r.v. - range of r.v. is further countable $S_X = \{x_1, x_2, \dots \} \text{ or } S_X = \{x_1, \dots, x_n\}$

Continuous r.v. - vange is uncountable 5x = [-1, 1]

Discrete r.v.'s pubability mass function (PMF) $P_X: S_X \rightarrow [0,1]$ $P_X(x) = P(X=x) = P(\{w \in x \mid x(w) = x\})$

(a) $P_X(x) \ge 0 \quad \forall x \in S_X$

(b) $\sum_{x \in S_X} p_x(x) = 1$

(e) Frang $B \subset S_X$, $P(B) = \sum_{x \in B} P_X(x)$

e.g. 2 indep. tosses of fair coin. $\mathcal{E} = \{hh, ht, th, tt\}$ Let X = # heads, $S_X = \{0, 1, 2\}$ $P_X(x) = \{1/2, x = 1\}$ 0, 0/W

enter of mass of PMF

Center of mass of PMF

Center of mass of PMF

Center of mass of PMF

Function of r.v.

X: A > R

g: R > R

Let Y = g(X): $\Omega \to R \to R$ is another r.v.

e.g. X = temp in celsins

Y = g(x) = 1.8x + 32 = temp. in Fahrenheit

 $P_{Y}(y) = P(Y=y) =$ $\begin{cases} \sum_{\{x \mid g(x)=y\}} P_{X}(x) \end{cases}$

 $E[g(x)] = E[Y] = \sum_{y \in S_Y} y P_Y(y) = \sum_{x \in S_X} g(x) P_X(x)$

Momente: X ~ Px, Y= X2

"E(X2] = 2nd moment of X

E[xn] = n+h " "

- $Y = (X - E[X])^L$ is 2nd central moment $J \times D$ E[Y] = D, $Y \ge D$ $E[Y] = E[(X - E[X])^L] = Variance <math>J \times D$

 $\sigma_{X} = \sqrt{Var(x)} = standard deviation <math>\eta X$ $Var(x) = F[(x - F[x])^{2}]$

 $Var(x) = E[(x - E[x])^{2}]$ $= \sum_{x \in S_{X}} (x - E[x])^{2} P_{x}(x)$

- E(aX+b] = aE(X]+b $Var(aX+b] = a^{2}Var(X)$ $Var(X) = E(X^{2}] - E(X)^{2}$

Multiple v.v.'s (discrete)

Given v.v.'s X and Y assoc. w/same experiment

Tout $PMF \neq X$, Y is P(X = X, Y = Y) P(X, Y) = P(X = X, Y = Y) = P(X = X, M = Y)

Marginal PMF's $P_{X}(x) = \sum_{y} P_{X,Y}(x,y)$ $P_{Y}(y) = \sum_{x} P_{X,Y}(x,y)$

Covariance of 2 v.v.'s and Y Cov(X,Y) = E[(X-E[X])(Y-F[Y])

If Cov(X, Y) = 0, then say X, Y uncorrelated Cov(X, Y) > 0 values $q \times -E[X]$ and Y - E[Y]tend to have same sign.

co: opposite sign.

Cov (X, Y) = E[XY] - ECX]E[Y] Cov (X, X) = Vou (X).

Var(X+Y) = Var(X) + Var(Y) + 2 cor(X, Y)

Conditional PMF of X, conditioned on event B, P(B)>0, P(X)=P(X=X|B)=P(X=X|B) P(B)

 $\sum_{x} P_{X|B}(x) = 1$

If B_1, \dots, B_m is an event space, then $P_X(x) = \sum_{i=1}^m P_{X|B_i}(x) P(B_i)$

Conditioning one v.v. on another

Conditional PMF of X given Y

(specialize B to form $\Sigma Y = y$), $P_Y(y) > 0$) $P_{X|Y} (x|y) = P(X = x|Y = y)$ $= \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{P_{X,Y}(x,y)}{P_Y(y)}$

for f(x) of y, P(x) Y(x) is a PMF for x.

Thus, $\sum_{y \in P(x)} P(x) = 1$ for each $y \in S$. P(y) > 0

Also, Pylx (ylx) = Pxy (x,y)
Px(x)

 $P_{X,Y}(x,y) = P_Y(y) P_{X|Y}(x|y) = P_X(x) P_{Y|X}(y|x)$ for $P_X(x) > 0$, $P_Y(y) > 0$

Conditional Expectation

 $E[XIB] = \sum_{x} P_{xIB}(x)$ conditional expectation of X given event B P(B) > 0

 $E[g(x)|B] = \sum_{x} g(x) P_{X|B}(x)$.

For event space $B_1, \dots, B_m, P(B_i)>0 \ \forall i$ $E[X] = \sum_{i=1}^{m} E[X]B_i]P(B_i)$

Specialize to B = { Y=y}

 $E[X|Y=y] = \sum_{x} XP_{X|Y}(x|y)$ Conditional expectation of X given Y=y Y=y

Independence of r.v.'s

X, Y indep if

 $P_{X,Y}(x,y) = P_X(x)P_Y(y) \quad \forall (x,y)$

i.e. events $\{x=x\}$ and $\{y=y\}$ are indep. $\{x,y\}$.

(=) $P_{X|Y}(x|y) = P_{X}(x)$ $\forall y \ xt \ P_{Y}(y) > 0, \ \forall x$ $P_{Y|X}(y|x) = P_{Y}(y)$ $\forall x \ s.t. \ P_{X}(x) > 0, \ \forall y$ value of y provides no info. on value $g \times g$ and vie versa.

· 18 X, Y indep., then ECXY] = E[X]· E[Y].

* X, Y indep: corr(x, y) = E[xy] - E[x]E[y] = 0* X, Y

=) uncorrelated, but reverse not true

i.e. X, Y uncorrelated \neq X, Y indep.

sums of indep. r.v.'s

- $16 \times 1, \dots, \times n$ are indep, $Var(X_1 + \dots + X_n) = Var(X_1) + \dots + Var(X_n)$

- Let $X_1, ..., X_n$ indep, ident, dish. (iiid)

where $E[X_i] = \mu$, $Var(X_i) = \sigma^2$ $S_n = \frac{1}{n}(X_1 + ... + X_n)$.

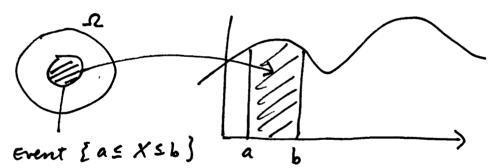
$$E[S_n] = M, \quad Var(S_n) = \sum_{i=1}^n \frac{1}{n^i} \log \sigma^2 = \frac{\sigma^2}{n} \rightarrow 0 \text{ as}$$

Continuous r. v. 's

- main difference: cannot specify prob. of each outcome in S2. In fact, prob. of each outcome = 0!
- Instead, specify probs for intervals
- Probability density function (PDF) f_X $P(X \in B) = \int_B f_X(x) dx \qquad \text{(Piemann Integral)}$

for every subset B of R.

$$P(a \le x \le b) = \int_a^b f_x(x) dx$$



-
$$f_X(x) > 0$$
 $\forall x$ and $\int_{-\infty}^{\infty} f_X(x) dx = 1$

- For δ small, $P(X \in [x, x+\delta]) = \int_{x}^{x+\delta} f_{x}(t)dt \sim f_{x}(x) \delta$
 - fx(x) as "prob. mass per unit length"

 fx(x) is not prob. of any event.

- Expectation
$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

(mid
$$\int_{\infty}^{\infty} |x| f_{x}(x) dx < \infty$$
)

$$- E[g(x)] = \int_{-\infty}^{\infty} g(x) f_{x}(x) dx$$

-
$$E[X^n] = \int_{-\infty}^{\infty} x^n f_X(x) dx$$
 with moment

$$-Vor(x) = E[(x-E[x])^{2}) = \int_{-\infty}^{\infty} (x-E[x])^{2} + \chi(x)dx$$

Cumulative Distribution Function (CDF)

- unity treatment for discrete and continuous r.v.'s

$$-F_{X}(x) = P(X \le x) = \begin{cases} \sum_{k \le x} P_{X}(k) & X \text{ discrete} \\ \int_{-\infty}^{X} f_{X}(x) dx & X \text{ continuous} \end{cases}$$

- By Fund. Th. of Calculus $f_X(x) = \frac{dF_X(x)}{dx}$

For continuous v.v.'s CDF is continuous.

$$-F_{X}(x) \nearrow F_{X}(x) \rightarrow 0 \text{ as } x \rightarrow -\infty$$

$$F_{X}(x) \rightarrow -1 \text{ as } x \rightarrow +\infty$$

- say X, Y are jointly continuous with joint PDF fx, Y if fx, Y is non-neg. function 5.t.

 $P((X,Y) \in B) = \iint_{(X,Y) \in B} f_{X,Y}(x,y) dx dy$

for every subset BCIR2

 $P(a \le x \le b, c \le y \le a) = \int_{c}^{d} \int_{a}^{b} f_{x,y}(x,y) dx dy$

 $-\int_{-\infty}^{\infty}\int_{-\infty}^{\infty}f_{x,y}(x,y)dxdy=1$

 $- p(a \le X \le a + \delta, c \le Y \le c + \delta) \sim f_{x,y}(a,c)\delta^{2}$

fx, y (a, c) is "prob. per unit area" near (a, c).

 $-f_{X}(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy \frac{margnual}{PDF's}$ $f_{Y}(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$

- Joint CDF $F_{X,Y}(x,y) = P(X \leq x, Y \leq y)$ $= \int_{-\infty}^{x} \int_{-\infty}^{y} f_{X,Y}(s,t) dt ds$ $f_{X,Y}(x,y) = \frac{\partial^{2} F_{X,Y}}{\partial x \partial y}(x,y)$

(x,y)~ fx,y

 $E[g(x,y)] = \int_{\infty}^{\infty} \int_{\infty}^{\infty} g(x,y) f_{x,y}(x,y) dx dy$

Conditional PPF of r.v. X, given went B, PCB) > 0 is non-neg. fx1B satisfying

 $P(X \in A \mid B) = \int_{A} f_{X \mid B}(x) dx$

 $\int_{-\infty}^{\infty} f_{XB}(x) dx = 1.$

- {B,,.., Bm} an event space. Then $f_{x}(x) = \sum_{i=1}^{\infty} f_{x|B_{i}}(x) P(B_{i})$

Conditionshy one r.v. on another

(x,y)~fx,y; conditional PDF of X given {Y=y}

is $f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}$ for $g \times f_{Y} f_{X}(x,y)$

- fxiy (xly) has same shape as fx,y(x,y) since fyly) does not depend on to.

 $-\int_{\infty}^{\infty} f_{X|Y}(x|y) dx = 1$

& Conditional Expectation

$$E[XIB] = \int_{-\infty}^{\infty} x f_{XIB}(x) dx$$

$$E[g(x)]BJ = \int_{-\infty}^{\infty} g(x) f_{x|B}(x) dx$$

$$- E[x|y=y] = \int_{-\infty}^{\infty} x f_{x|y}(x|y) dx$$

$$E[X] = \int_{-\infty}^{\infty} E[X|Y=y] f_{y}(y) dy$$
.

Independence

$$f_{X,Y}(x,y) = f_{X}(x) f_{Y}(y) \quad \forall x,y$$

$$f_{X|Y}(x|y) = f_{X}(x) + y$$
 with $f_{y}(y) > 0$

all y.

x ~ fx(x) unobserved phenomenon

Make noisy measurement Y = fylx cond. PDF

$$f_{X|Y}(x|y) = \frac{f_{X}(x) f_{Y|X}(y|x)}{f_{Y}(y)}$$

$$= \frac{f_{X}(x) f_{Y|X}(y|x)}{f_{X}(x) f_{Y|X}(y|x)}$$

$$= \frac{\int_{-\infty}^{\infty} f_{X}(t) f_{Y|X}(y|t) dt}{\int_{-\infty}^{\infty} f_{X}(t) f_{Y|X}(y|t) dt}$$

Inference about discrete mi r.v.

- unobsenced phenom is discrete (disease, no disease) P(A), $f_{Y|A}(y)$, $f_{Y|A}(y)$ event

want P(A|Y=y)

- suppose A has from
$$\{N=n\}$$
, $N \sim PN(PMF)$

$$P(N=n|Y=y) = \frac{PN(n) f_{Y}|N(y|n)}{\sum_{i} P_{N}(i) f_{Y}|N(y|i)}$$