Classification Rule Derivation

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0.0.1 a) Classification rule devivation

Starting with some definitions:

- π_i is the prior probability that an individual selected at random belongs to population i.
- C(i|j) is the cost of incorrectly allocating an individual to population i, when they really belong to population j.

Bayes rule:

Allocate population 1 if:

$$\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} \ge \frac{\pi_2 C(1|2)}{\pi_1 C(2|1)},$$

otherwise allocate population 2.

In the multivariate normal case (with p variables), where the observation vectors $\mathbf{X}_i \sim \text{MVN}_p(\mu_i, \Sigma)$, then our two probability density functions for population 1 and 2 are:

$$f_1(\mathbf{x}) = \frac{1}{|2\pi\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \mu_1)^T \Sigma^{-1}(\mathbf{x} - \mu_1)\right\},$$

and

$$f_2(\mathbf{x}) = \frac{1}{|2\pi\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \mu_2)^T \Sigma^{-1}(\mathbf{x} - \mu_2)\right\},$$

respectively.

We can therefore calculate the likelihood ratio as:

$$\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} = \exp\left\{-\frac{1}{2}(\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1) + \frac{1}{2}(\mathbf{x} - \mu_2)^T \Sigma^{-1} (\mathbf{x} - \mu_2)\right\},\,$$

where the $\frac{1}{|2\pi\Sigma|^{1/2}}$ terms in each pdf cancel out.

We can then calculate the log-likelihood by taking natural logs of both sides:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) = -\frac{1}{2}(\mathbf{x} - \mu_1)^T \Sigma^{-1}(\mathbf{x} - \mu_1) + \frac{1}{2}(\mathbf{x} - \mu_2)^T \Sigma^{-1}(\mathbf{x} - \mu_2)$$

and factorising the $\frac{1}{2}$:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) = -\frac{1}{2}\left((\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1) - (\mathbf{x} - \mu_2)^T \Sigma^{-1} (\mathbf{x} - \mu_2)\right).$$

We can expand out the $(\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1) - (\mathbf{x} - \mu_2)^T \Sigma^{-1} (\mathbf{x} - \mu_2)$ terms:

Firstly
$$(\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1)$$

$$= (\mathbf{x}^T \Sigma^{-1} - \mu_1^T \Sigma^{-1}) (\mathbf{x} - \mu_1)$$

$$= \mathbf{x}^T \Sigma^{-1} \mathbf{x} - \mathbf{x}^T \Sigma^{-1} \mu_1 - \mu_1^T \Sigma^{-1} \mathbf{x} + \mu_1^T \Sigma^{-1} \mu_1$$

Each of these four components is a scalar, and the transpose of a scalar is the scalar itself, i.e. $a^T = a$, therefore:

$$\mu_{\mathbf{1}}^{T} \Sigma^{-1} \mathbf{x} = (\mu_{\mathbf{1}}^{T} \Sigma^{-1} \mathbf{x})^{T} = \mathbf{x}^{T} (\Sigma^{-1})^{T} \mu_{\mathbf{1}} = \mathbf{x}^{T} (\Sigma^{T})^{-1} \mu_{\mathbf{1}} = \mathbf{x}^{T} \Sigma^{-1} \mu_{\mathbf{1}},$$

since Σ is a symmetric matrix, $\Sigma = \Sigma^T$.

We can therefore write:

$$(\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1) = \mathbf{x}^T \Sigma^{-1} \mathbf{x} - 2 \mathbf{x}^T \Sigma^{-1} \mu_1 + \mu_1^T \Sigma^{-1} \mu_1$$

and

$$(\mathbf{x} - \mu_2)^T \Sigma^{-1} (\mathbf{x} - \mu_2) = \mathbf{x}^T \Sigma^{-1} \mathbf{x} - 2 \mathbf{x}^T \Sigma^{-1} \mu_2 + \mu_2^T \Sigma^{-1} \mu_2.$$

Thus we can expand out:

$$(\mathbf{x} - \boldsymbol{\mu_1})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu_1}) - (\mathbf{x} - \boldsymbol{\mu_2})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu_2})$$

to

$$-2\mathbf{x}^{T} \Sigma^{-1} \mu_{1} + \mu_{1}^{T} \Sigma^{-1} \mu_{1} + 2\mathbf{x}^{T} \Sigma^{-1} \mu_{2} - \mu_{2}^{T} \Sigma^{-1} \mu_{2},$$

where we can factor out the $-2\mathbf{x}^T\Sigma^{-1}$ terms:

$$-2\mathbf{x}^{T}\Sigma^{-1}(\mu_{1}-\mu_{2})+\mu_{1}^{T}\Sigma^{-1}\mu_{1}-\mu_{2}^{T}\Sigma^{-1}\mu_{2}.$$

The $\mu_1^T \Sigma^{-1} \mu_1 - \mu_2^T \Sigma^{-1} \mu_2$ terms follow the difference of two squares, and can be written as $(\mu_1 + \mu_2)^T \Sigma^{-1} (\mu_1 - \mu_2)$.

Our log-likelihood can therefore be written:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) = -\frac{1}{2}\left(-2\mathbf{x}^T \Sigma^{-1}(\mu_1 - \mu_2) + (\mu_1 + \mu_2)^T \Sigma^{-1}(\mu_1 - \mu_2)\right)$$
$$= \mathbf{x}^T \Sigma^{-1}(\mu_1 - \mu_2) - \frac{1}{2}(\mu_1 + \mu_2)^T \Sigma^{-1}(\mu_1 - \mu_2)$$

The common factor in both terms is $\Sigma^{-1}(\mu_1 - \mu_2)$. Let $\mathbf{L} = \Sigma^{-1}(\mu_1 - \mu_2)$ such that our log-likelihood takes the form:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) = \mathbf{x}^T \mathbf{L} - \frac{1}{2}(\mu_1 + \mu_2)^T \mathbf{L}$$

Note again that each of the two above terms is a scalar (**L** is a $p \times 1$ vector), so we can shuffle the order of the terms by taking the transpose of each:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) = \mathbf{L}^T \mathbf{x} - \frac{1}{2} \mathbf{L}^T (\mu_1 + \mu_2).$$

Recall that as:

$$\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})} \ge \frac{\pi_2 C(1|2)}{\pi_1 C(2|1)},$$

the log-likelihood relationship must be:

$$\ln\left(\frac{f_1(\mathbf{x})}{f_2(\mathbf{x})}\right) \ge \ln\left(\frac{\pi_2 C(1|2)}{\pi_1 C(2|1)}\right),\,$$

therefore:

$$\mathbf{L}^T \mathbf{x} - \frac{1}{2} \mathbf{L}^T (\mu_1 + \mu_2) \ge \ln \left(\frac{\pi_2 C(1|2)}{\pi_1 C(2|1)} \right).$$

Our allocation rule is therefore:

Allocate to population 1 if:

$$\mathbf{L}^T \mathbf{x} - \frac{1}{2} \mathbf{L}^T (\mu_1 + \mu_2) \ge \ln \left(\frac{\pi_2 C(1|2)}{\pi_1 C(2|1)} \right),$$

otherwise allocate to population 2.