Probability and Distribution Theory:

Supplementary Exercises

Probability Theory

- 1. Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let A, B, C and D be events on Ω , i.e. A, B, C, D belong to the collection \mathcal{F} . Express the following statements in set theoretic terms.
 - (a) A occurs but neither B nor C nor D occurs.
 - (b) None of the events occur.
 - (c) All occur.
 - (d) At most, three occur.
 - (e) At least three of them occur.
 - (f) Exactly three occur.
 - (g) Exactly one of them occurs.
 - (h) Exactly one of A and B occurs, but D does not occur.
- 2. A survey of the students at a college who use email found that 50% have a Shotmail account, 30% have a Payserve account, and 25% have both Shotmail and Payserve accounts.

Which of the following statements are correct?

- (a) 0.55 that s/he has a Shotmail or a Payserve account.
- (b) 0.75 that s/he does not have both a Shotmail and a Payserve account.
- (c) 0.75 that s/he has neither a Shotmail nor a Payserve account.
- (d) 0.25 that s/he has either a Shotmail or a Payserve account, but not both.
- 3. A pair of fair dice are thrown once. What is the probability that
 - (a) the sum of the scores is divisible by 4?
 - (b) a three turns up exactly once?
 - (c) the sum of the scores is 5?
 - (d) both numbers are even?

- 4. A pair of *fair* five-sided spinners are spun, each one being numbered 1,2,3,4, or 5. What is the probability that the sum is 4, given that a sum of 4 or more is obtained?
- 5. Suppose that A and B are independent events. Show that:
 - (a) A^c and B are independent events;
 - (b) A^c and B^c are independent events.

Univariate Distribution Theory

6. Determine the constant c for the following distributions.

(a)
$$f_X(x) = \begin{cases} c(1-x^2) & -1 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

(b)
$$f_X(x) = \begin{cases} \frac{c}{x^2} & x > 5\\ 0 & \text{o.w.} \end{cases}$$

7. Calculate the mean and variance of the following distributions.

(a)
$$f_X(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x} \qquad x > 0$$

where $\alpha, \beta > 0$,

i.e. $X \sim Gamma(\alpha, \beta)$.

(b)
$$f_X(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha - 1} (1 - x)^{\beta - 1} \qquad 0 < x < 1$$

where $\alpha, \beta > 0$

i.e. $X \sim Beta(\alpha, \beta)$.

You may use the fact that $\Gamma(\alpha) = \int_0^\infty t^{\alpha-1} e^{-t} dt$ and $\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$ for $\alpha > 0$.

8. The Normal distribution

Assuming that

$$\int_{-\infty}^{\infty} e^{-\frac{y^2}{2}} dy = \sqrt{2\pi},$$

show that

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} - \infty < x < \infty$$

is a p.d.f. for any $\mu \in \mathbb{R}$, $\sigma > 0$.

[Hint: Consider using the substitution $y = \frac{x-\mu}{\sigma}$ for part of your analysis].

Joint and Conditional Distributions

9. Suppose X and Y are continuous random variables with p.d.f.

$$f_{(X,Y)}(x,y) = \begin{cases} cx^2 & 0 < x < 1, \ 0 < y < 1 \\ 0 & \text{o.w.} \end{cases}$$

- (a) Specify $R_{(X,Y)}$.
- (b) Find the constant c.
- (c) Find the marginal p.d.f.'s of X and Y, i.e. $f_X(\cdot)$ and $f_Y(\cdot)$ respectively.
- 10. Suppose (X,Y) is a continuous random vector which is 'uniformly' distributed over $R_{(X,Y)} = \{(x,y) : -1 \le x \le 1, -1 \le y \le 1\}$ (i.e. the joint p.d.f. is constant over the entire region).
 - (a) Specify the joint p.d.f. $f_{(X,Y)}$.
 - (b) Find:

i.
$$\mathbb{P}(2X > Y)$$

ii.
$$\mathbb{P}(X^2 + Y^2 < 1)$$

iii.
$$\mathbb{P}(|X + Y| < 1)$$
.

11. The discrete random vector (X,Y) has a p.m.f. given by the table below:

X

$$Y \begin{array}{|c|c|c|c|c|c|c|} \hline & 1 & 2 & 3 \\ \hline 1 & 1/6 & 1/12 & 1/12 \\ 2 & 1/6 & 0 & 1/6 \\ 3 & 0 & 2/9 & 1/9 \\ \hline \end{array}$$

(So, for e.g., $p_{(X,Y)}(2,1) = \frac{1}{12}$).

- (a) Find the marginal p.m.f.'s of X and Y, i.e. $p_X(k)$ for each $k \in R_X$ and $p_Y(k)$ for each $k \in R_Y$.
- (b) Find the p.m.f. of XY and hence E[XY].
- (c) Calculate $\mathbb{P}(Y=2|X>1)$.
- 12. The conditional p.d.f. of Y, given X = x, is

$$f_{Y|X}(y|x) = \begin{cases} \frac{2y}{x^2} & 0 < x < 1, \ 0 < y < x \\ 0 & \text{o.w.} \end{cases}$$

and the marginal p.d.f. of X is

$$f_X(x) = \begin{cases} 7x^6 & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

- (a) Find the joint p.d.f. of (X, Y).
- (b) Find the marginal p.d.f. of Y.
- 13. Let the random variables X and Y have joint p.d.f.

$$f_{(X,Y)}(x,y) = \begin{cases} x+y & \text{if } 0 < x < 1, \ 0 < y < 1 \\ 0 & \text{o.w.} \end{cases}$$

Find

- (a) cov(X, Y)
- (b) $\rho(X,Y)$ and
- (c) var(2X Y + 4).
- 14. Suppose that the joint density function of X and Y is given by

$$f_{(X,Y)}(x,y) = \begin{cases} \frac{1}{2} & 0 < x < y < 2 \\ 0 & \text{o.w.} \end{cases}$$

- (a) Find the marginal probability density functions of X and Y.
- (b) Are X and Y independent? Justify your answer.

Transformations of Random Variables

15. (a) Show that if $U \sim Uniform(0,1)$, then

$$-\frac{1}{\lambda}\ln U \sim Exp(\lambda)$$

where $\lambda > 0$.

(b) Now suppose that $X \sim Exp(\lambda)$, where $\lambda > 0$. Show that

$$1 - e^{-\lambda X} \sim U(0, 1).$$

- 16. Suppose that $X \sim Exp(\lambda)$ and $Y \sim Exp(\mu)$ are independent r.v.'s. Show that $W = \min(X, Y) \sim Exp(\lambda + \mu)$.
- 17. Suppose that U and V are independently distributed non-negative random variables with joint p.d.f. given by $f_{(U,V)}(\cdot,\cdot)$. Define the random variables X and Y via the transformations

$$X = U,$$
 $Y = U + V.$

Let the joint p.d.f. of X and Y be denoted by $f_{(X,Y)}(\cdot,\cdot)$.

(a) State the numerical value of J(x,y) in the expression

$$f_{(X,Y)}(x,y) = f_{(U,V)}(x,y-x)|J(x,y)|.$$

(b) Suppose that

$$f_{(U,V)}(u,v) = \lambda^2 e^{-\lambda(u+v)}, \qquad u \ge 0, \ v \ge 0.$$

Find an expression in terms of x and/or y for $f_{(X,Y)}(x,y)$. What are the set of values (x,y) for which $f_{(X,Y)}(x,y)$ is positive?

Generating Functions for univariate distributions

18. Suppose that X is a discrete r.v. with p.g.f.

$$G_X(\theta) = 0.35 + 0.65\theta.$$

Find the following quantities:

(a)
$$\mathbb{P}(X=1)$$
,

- (b) $\mathbb{P}(X = 10)$,
- (c) $\mathbb{P}(X=0)$,
- (d) E[X].
- 19. Suppose that $X \sim Gamma(\alpha, \beta)$ i.e.

$$f_X(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x}$$
 $x > 0$

 $\alpha, \beta > 0$.

Compute the m.g.f. of X, and hence calculate E[X] and var(X).

20. Suppose that the r.v. W has m.g.f.

$$M_W(t) = (1 - 7t)^{-20}$$
.

Find the mean and variance of W.

21. Let X be the time it takes, starting from midnight, for a total of 3 buses to leave Trafalgar Square. It is assumed that X has the following probability density function

$$f_X(x) = \begin{cases} 108x^2e^{-6x} & x > 0\\ 0 & \text{o.w.} \end{cases}$$

- (a) Find the moment generating function, $M_X(t)$, of X (for t in a suitably defined range).
- (b) Compute $M_X^{(3)}(t)$ and hence find $E[X^3]$.
- 22. (a) Suppose that the random variable $X \sim Bernoulli(p)$ where 0 . Show that the characteristic function of <math>X is given by

$$\Psi_X(\theta) = (1 - p) + pe^{i\theta}.$$

(b) Suppose that $X_j \sim Bernoulli(p)$, i = 1, ..., n, are independent. Let $Y = \sum_{j=1}^n X_j$. Show that the characteristic function of Y is given by

$$\Psi_Y(\theta) = \left((1 - p) + pe^{i\theta} \right)^n.$$

Central Limit Theorem

23. Let X_1, X_2, \ldots, X_n be a random sample of size n = 40, with population mean $\mu = 6.53$, and population variance $\sigma^2 = 10$.

By using the *Central Limit Theorem*, find the approximate probability that $\sum_{i=1}^{40} X_i$ lies between 245 and 255.

24. The random variables X_1, \ldots, X_{625} are independent and identically distributed each with p.d.f. $f_X(x) = 3(1-x)^2$ for $0 \le x \le 1$. Approximate

$$\mathbb{P}(X_1 + X_2 + \ldots + X_{625} < 170)$$

in terms of the distribution function, $\Phi(\cdot)$, of the standard normal distribution.