

## IS 604 – Homework #5

**Include the programming code (R, MATLAB, or Python) you used for each problem.**

1. At the start of each week, the condition of a machine is determined by measuring the amount of electrical current it uses. According to its amperage reading, the machine is categorized as being in one of the following four states: low, medium, high and failed. A machine in the low state has a probability of 0.05, 0.03, and 0.02 of being in the medium, high, or failed state, respectively, at the start of the next week. A machine in the medium state has a probability of 0.09 and 0.06 of being in the high or failed state, respectively, at the start of the next week (it cannot, by itself, go to the low state). And, a machine in the high state has a probability of 0.1 of being in the failed state at the start of the next week (it cannot, by itself, go to the low or medium state). If a machine is in the failed state at the start of a week, repair is immediately begun on the machine so that it will (with probability 1) be in the low state at the start of the following week. Let  $X$  be a Markov chain where  $X_n$  is the state of the machine at the start of week  $n$ .

- a) Give the Markov transition matrix for  $X$ .
- b) A new machine always starts in the low state. What is the probability that the machine is in the failed state three weeks after it is new?
- c) What is the probability that a machine has at least one failure three weeks after it is new?
- d) What is the expected number of weeks after a new machine is installed until the first failure occurs?
- e) On average, how many weeks per year is the machine working?
- f) Each week that the machine is in the low state, a profit of \$1000 is realized; each week that the machine is in the medium state, a profit of \$500 is realized; each week that the machine is in the high state, a profit of \$400 is realized; and the week in which a failure is fixed, a cost of \$700 is incurred. What is the long-run average profit per week realized by the machine?
- g) A suggestion has been made to change the maintenance policy for the machine. If at the start of a week the machine is in the high state, the machine will be taken out of service and repaired so that at the start of the next week it will again be in the low state. When a repair is made due to the machine being in the high state instead of a failed state, a cost of \$600 is incurred. Is this new policy worthwhile?

2. Rao (1973) presented an example on genetic linkage of 197 animals in four categories. The group sizes are (125, 18, 20, 34). Assume that the probabilities of the corresponding multinomial distribution are  $\left(\frac{1}{2} + \frac{\theta}{4}, \frac{1-\theta}{4}, \frac{1-\theta}{4}, \frac{\theta}{4}\right)$ . Estimate the posterior distribution of  $\theta$  given the observed sample, using the Metropolis-Hasting random walk sampler with a uniform proposal distribution. Note that the posterior distribution of  $\theta$  given the observed frequencies  $k = (k_1, k_2, k_3, k_4)$  is

$$f_{\theta|K}(\theta) \propto (2 + \theta)^{k_1} (1 - \theta)^{k_2 + k_3} \theta^{k_4}$$

3. Let  $y_1, \dots, y_n$  be a sample from a Poisson distribution for which there is a suspicion of a change point  $m$  along the observation process where the means change,  $m = 1, \dots, n$ .

Given  $m$ , the observation distributions are:

$$y_i | \lambda \sim \text{Poi}(\lambda), i = 1, \dots, m$$

$$y_i | \phi \sim \text{Poi}(\phi), i = m+1, \dots, n.$$

The model is completed with independent prior distributions

$$\lambda \sim \text{Gamma}(\alpha, \beta)$$

$$\phi \sim \text{Gamma}(\gamma, \delta),$$

$$m \sim \text{Uniform Discrete}(1, n).$$

Then the posterior joint density is:

$$\begin{aligned} \pi(\lambda, \phi, m | y_1, \dots, y_n) &\propto f(y_1, \dots, y_n | \lambda, \phi, m) p(\lambda, \phi, m) \\ &= \left[ \prod_{i=1}^m f(y_i | \lambda) \prod_{i=m+1}^n f(y_i | \phi) \right] f(\lambda | \alpha, \beta) f(\phi | \gamma, \delta) \frac{1}{n} \\ &\propto \left[ \prod_{i=1}^m e^{-\lambda} \lambda^{y_i} \prod_{i=m+1}^n e^{-\phi} \phi^{y_i} \right] (\lambda^{\alpha-1} e^{-\beta\lambda}) (\phi^{\gamma-1} e^{-\delta\phi}) \\ &\propto \lambda^{\alpha+s_m-1} e^{-(\beta+m)\lambda} \phi^{\gamma+s_n-s_m-1} e^{-(\delta+n-m)\phi} \\ s_m &= \sum_{i=1}^m y_i \\ s_n &= \sum_{i=1}^n y_i \end{aligned}$$

Use the vector  $Y$  below, containing count data of coal mining accidents in the U.K. from 1851-1962, and perform Gibbs sampling on this data to perform inference on the joint posterior distribution of  $(\lambda, \phi, m)$ .

$Y =$   
 $[4, 5, 4, 1, 0, 4, 3, 4, 0, 6, 3, 3, 4, 0, 2, 6, 3, 3, 5, 4, 5, 3, 1, 4, 4, 1, 5, 5, 3, 4, 2, 5, 2, 2, 3, 4, 2, 1, 3, 2, 2, 1, 1, 1, 1, 3, 0, 0, 1, 0, 1,$   
 $1, 0, 0, 3, 1, 0, 3, 2, 2, 0, 1, 1, 1, 0, 1, 0, 1, 0, 0, 0, 2, 1, 0, 0, 0, 1, 1, 0, 2, 3, 3, 1, 1, 2, 1, 1, 1, 1, 2, 4, 2, 0, 0, 0, 1, 4, 0, 0, 0, 1, 0, 0,$   
 $0, 0, 0, 1, 0, 0, 1, 0, 1];$

### Guide to Implementation:

It turns out to work even better if you add two more distributions to the above conditional. Assume that the scale parameters  $\beta$  and  $\delta$  for the distributions of  $\lambda$  and  $\phi$ , respectively are themselves distributed as Inverse Gamma

$$\beta \sim \text{InvGamma}(\alpha, \lambda + 1)$$

$$\delta \sim \text{InvGamma}(\gamma, \phi + 1)$$

Then implement the Gibbs sampling as a Markov chain of 5 values for each iteration:  $(\lambda, \phi, m, \beta, \delta)$ .

### **Pseudo-code for Implementation**

Initialize # of iterations, hyperparameters, initial guesses of the five parameters

Create a vector Mprob to contain the discrete probability distribution over all possible values of  $m$  (size  $N$ , the number of observations in  $Y$ ).

Create a matrix Chain with 5 columns and as many rows as iterations

For  $i = 1$  to number of iterations

    % Conditional on  $m$ ,  $b$ , and  $d$ , update samples of  $l$  and  $f$

    Calculate  $\text{sum1} = \text{sum}(Y(1) \text{ to } Y(m))$

    Calculate  $\text{aum2} = \text{sum}(Y(m+1) \text{ to } Y(N))$

    Sample  $\lambda$  from  $\text{Gamma}(\text{sum1} + \alpha, \beta / (\beta * m + 1))$

    Sample  $\phi$  from  $\text{Gamma}(\text{sum2} + \gamma, \delta / (\delta * (n - m) + 1))$

    % Update probability distribution of  $m$ , using function fullcondm() defined below

    For  $j = 1$  to  $n$

        Mprob( $j$ ) = fullcondm( $j$ ,  $\lambda$ ,  $\phi$ ,  $Y$ ,  $n$ ,  $\alpha$ ,  $\beta$ ,  $\gamma$ ,  $\delta$ )

    end

    sample new  $m$  from discrete distribution in mprob

    % update samples of  $\beta$  and  $\delta$  conditional on  $\lambda$ ,  $\phi$ , and  $m$

    Beta =  $1 / (\text{sample from } \text{Gamma}(\alpha, \lambda + 1))$

    Delta =  $1 / (\text{sample from } \text{Gamma}(\gamma, \phi + 1))$

HINT: I found that occasionally, this sampling would give crazy huge numbers for  $\beta$ , and  $\delta$  which would break the MCMC, so to be safe, embed the above two statements in loop that only exits when new  $\beta$  and  $\delta$  samples are  $\leq 99$ .

    Store new samples for ( $\lambda$ ,  $\phi$ ,  $m$ ,  $\beta$ ,  $\delta$ )

End (for loop)

Define function fullcondm(  $m, \lambda, \phi, y, n, \alpha_0, \beta_0, \gamma_0, \delta_0$  )

– the conditional probability that the true change point is the parameter value passed in  $m$ , given the other values.

Lamexp = sum(Y(1:m)) if m > 1, 0 otherwise  
 Phiexp = sum(Y(m+1:n)) if m < n, 0 otherwise

Return prob =  $\lambda^{\alpha_0-1+\text{lamexp}} \exp(-(\beta_0+m)\lambda) * \phi^{\gamma_0-1+\text{phiexp}} \exp(-(\delta_0+n-m)\phi);$

## Questions

a) Run your code for 5000 iterations. Create and save the following figures:

Histogram of lambda  
 Histogram of phi  
 Histogram of m  
 Plot of lambda vs. phi  
 Plot of lambda vs. m  
 Plot of beta vs. delta

b) When do you think the change point occurred, based on your results? Can you put a 95% confidence bound on this result? What were the average rates of coal mining accidents before and after the change? Are these results consistent with the time series of the observations? Why or why not?

c) How is Gibbs sampling different from the Metropolis-Hastings approach?

4. Implement a Metropolis-Hastings simulated annealing optimization to find a least cost solution to the traveling salesman problem (TSP). A sample problem is provided below, which is a 17x17 matrix of cost (C) between every pair of cities in a 17-city TSP problem.

C =

0	633	257	91	412	150	80	134	259	505	353	324	70	211	268	246	121
633	0	390	661	227	488	572	530	555	289	282	638	567	466	420	745	518
257	390	0	228	169	112	196	154	372	262	110	437	191	74	53	472	142
91	661	228	0	383	120	77	105	175	476	324	240	27	182	239	237	84
412	227	169	383	0	267	351	309	338	196	61	421	346	243	199	528	297
150	488	112	120	267	0	63	34	264	360	208	329	83	105	123	364	35
80	572	196	77	351	63	0	29	232	444	292	297	47	150	207	332	29
134	530	154	105	309	34	29	0	249	402	250	314	68	108	165	349	36
259	555	372	175	338	264	232	249	0	495	352	95	189	326	383	202	236
505	289	262	476	196	360	444	402	495	0	154	578	439	336	240	685	390
353	282	110	324	61	208	292	250	352	154	0	435	287	184	140	542	238
324	638	437	240	421	329	297	314	95	578	435	0	254	391	448	157	301
70	567	191	27	346	83	47	68	189	439	287	254	0	145	202	289	55
211	466	74	182	243	105	150	108	326	336	184	391	145	0	57	426	96

268	420	53	239	199	123	207	165	383	240	140	448	202	57	0	483	153
246	745	472	237	528	364	332	349	202	685	542	157	289	426	483	0	336
121	518	142	84	297	35	29	36	236	390	238	301	55	96	153	336	0

To implement:

- Write a function that returns the total cost of a candidate solution for TSP, given the proposed solution and the cost matrix.
- Write a function to perform simulated annealing that takes three input parameters: the initial temperature  $T_0$ , beta (determines annealing schedule), and number of iterations, and returns two results: the solution vector  $x$  and the cost of the solution  $s$ .

The pseudo-code for this algorithm is:

```
Initialize: T=T0, C, n = length(C),
           initial guess x = random permutation of (1:17), Sx = cost(x)
           xbest = x, sbest = s
           set up vector to save the cost of current solution at each iteration

Loop from 1 to number of iterations
    Choose another random permutation of the integers 1-17
    Sort the first 2 in this list -> call this vector of two indices I
    Create a candidate solution by reordering current x between I(1) and I(2):
        Y = [x(1:I(1)-1), x(I(2):-1:I(1)), x(I(2)+1:end)];
        % y is now new proposed solution

    Set Sy = cost(y)
    If Sy < Sx then set alpha=1 otherwise set alpha = exp(-(Sy-Sx)/T);
    Draw u ~ U(0,1) and if u < alpha, set x=y, Sx=Sy
    Set T = beta * T
    Set xbest = x, sbest = Sx
    Save sbest in results(i)
End loop
```

Test this code with initial parameters  $T_0=1$ ,  $\beta=0.9999$ , iterations = 10000. How does the cost of the minimum solution compare with the cost of a random tour? Plot the cost of the current best solution as a function of iteration number. What do you notice about the performance of simulated annealing? Repeat the simulated annealing procedure 3 or 4 times and comment on the variability of the minimum cost and the optimal solution from repeated attempts.