# Part IB — Statistics Example Sheet 3 $\,$

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Let  $X \sim N_n(\mu, \Sigma)$ , and let  $A \in \mathbb{R}^{m \times n}$ . X has a *n*-variable normal distribution, so for every  $t \in \mathbb{R}^n$ , the rv.  $t^T X$  has a univariate normal distribution. Have

$$AX = \begin{pmatrix} a_{11}X_1 + \dots + a_{1n}X_n \\ \vdots \\ \vdots \\ a_{m1}X_1 + \dots + a_{mn}X_n \end{pmatrix}$$

Want to show that for every  $s \in \mathbb{R}^m$ , the variable  $s^TAX$  has a univariate normal distribution. Not sure. Assuming it does, we calculate the mean and variance.

Have

$$\mathbb{E}[AX] = \begin{pmatrix} a_{11}\mathbb{E}[X_1] + \dots + a_{1n}\mathbb{E}[X_n] \\ \vdots \\ \vdots \\ a_{m1}\mathbb{E}[X_1] + \dots + a_{mn}\mathbb{E}[X_n] \end{pmatrix} = \begin{pmatrix} a_{11}\mu_1 + \dots + a_{1n}\mu_n \\ \vdots \\ \vdots \\ a_{m1}\mu_1 + \dots + a_{mn}\mu_n \end{pmatrix} = A\mu$$

and

$$\operatorname{Var}\left[AX\right] = \begin{pmatrix} a_{11}^{2}\operatorname{Var}\left[X_{1}\right] + \dots + a_{1n}^{2}\operatorname{Var}\left[X_{n}\right] \\ \vdots \\ a_{m1}^{2}\operatorname{Var}\left[X_{1}\right] + \dots + a_{mn}^{2}\operatorname{Var}\left[X_{n}\right] \end{pmatrix} = \begin{pmatrix} a_{11}^{2}\sigma_{1}^{2} + \dots + \sigma_{n}^{2}\sigma_{n}^{2} \\ \vdots \\ a_{m1}^{2}\sigma_{1}^{2} + \dots + a_{mn}^{2}\sigma_{n}^{2} \end{pmatrix} = A\Sigma A^{T}$$

Choosing A be the  $n_1 \times n$  matrix

$$A = \begin{pmatrix} 1 & & 0 \\ & \ddots & & 0 \\ 0 & & 1 \end{pmatrix}$$

Have  $AX = X_1$ , and from Question 1 we have

$$X_1 \sim N_{n_1}(A\mu, A\Sigma A^T)$$

And we can see that  $AX_1 = \mu_1$  and  $A\Sigma A^T = \Sigma_{11}$  as required.

$$Y_i = a + bx_i + \varepsilon_i$$

This is a linear model; we have  $Y_i \sim N(a + bx_i, \sigma^2)$ , with likelihood

$$f_{Y_i} = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - a - bx_i)^2\right)$$

Then

$$\frac{\partial f_{Y_i}}{\partial a} \propto \sum_{i=1}^{n} (y_i - a - bx_i)$$

and since  $\sum_{i=1}^{n} x_i = 0$  we have

$$\hat{a} = \frac{1}{n} \sum_{i=1}^{n} y_i = \bar{Y}$$

Next, we have

$$\frac{\partial f_{Y_i}}{\partial b} \propto \sum_{i=1}^n (y_i - a - bx_i) x_i$$

then

$$\hat{b} = \frac{\sum_{i=1}^{n} x_i y_i}{\sum_{i=1}^{n} x_i^2}$$

Now the log-likelihood is given by

$$l(\alpha, \beta, \sigma) = -\frac{n}{2}\log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - a - bx_i)^2$$

Here,

$$\frac{\partial l}{\partial \sigma^2} = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - a - bx_i)^2$$

Thus setting this to zero at  $a = \hat{a}$ ,  $b = \hat{b}$  yields

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{a} - \hat{b}x_i)^2$$

This is a linear model with

$$y_i = \underbrace{\frac{v^2}{g}}_{\hat{\beta}} \underbrace{\sin 2\alpha}_{x_i} + \varepsilon_i$$

Here,

$$\hat{b} = \frac{\sum_{i=1}^{n} x_i y_i}{\sum_{i=1}^{n} x_i^2}$$

$$= 28156$$

Hence  $v = \sqrt{g\hat{\beta}}$  which is approximately 526 ms<sup>-1</sup>.

$$Y_{ij} = \mu_i + \varepsilon_{ij}$$

This is a linear model; we have  $Y_{ij} \sim N(\mu_i, \sigma^2)$ , with likelihood

$$f_{Y_{ij}} = \frac{1}{(2\pi\sigma^2)^{n_i/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^{I} \sum_{i=j}^{n_i} (y_{ij} - \mu_i)^2\right)$$

Then

$$\frac{\partial f_{Y_{ij}}}{\partial \mu_i} \propto \left(2n_i\hat{\mu}_i - 2\sum_{j=1}^{n_i} y_{ij}\right)$$

thus

$$\hat{\mu}_i = \frac{1}{n} \sum_{i=1}^{n_i} y_{ij} = \hat{Y}_i$$

Now the log-likelihood is given by

$$l(\mu_i, \sigma) = -\frac{n_i}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - a - bx_i)^2$$

Here,

$$\frac{\partial l}{\partial \sigma^2} = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^2} \sum_{i=1}^{I} \sum_{i=j}^{n_i} (y_{ij} - \mu_i)^2$$

Thus setting this to zero yields

$$0 = -\frac{n_i}{2\hat{\sigma}^2} + \frac{1}{2\hat{\sigma}^4} \sum_{i=1}^{I} \sum_{i=j}^{n_i} (y_{ij} - \mu_i)^2$$

Thus

$$\hat{\sigma}^2 = \frac{1}{n_i} \sum_{i=1}^{I} \sum_{i=j}^{n_i} (y_{ij} - \mu_i)^2$$

We can write the distribution of the linearly transformed variable

$$(\bar{X}, X_1 - \bar{X}, X_2 - \bar{X}, \cdots, X_n - \bar{X})$$

as

$$\frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=j}^{n} (x_i - \bar{x})^2\right) \exp\left(\frac{n(\bar{x} - \mu)^2}{2\sigma^2}\right)$$

Not sure how to show  $\bar{X}$  and  $X_i - \bar{X}$  are independent. But if we do that, then it follows  $S_{XX}$  and  $\bar{X}$  are independent, as  $S_{XX}$  is a function of the  $X_i - \bar{X}$ .