# Ahern-Clark-SP-2017-Appendix

### February 26, 2017

### 1 Introduction

This Jupyter notebook contains the python and R code for running the analysis in:

```
Ahern, Christopher and Robin Clark. Conflict, Cheap Talk, and Jespersen's Cycle. 2017.
```

Here is system and version information. Using the Anaconda python distribution is highly recommended for installing and configuring the SciPy stack and other dependencies.

Below we add add additional commentary and code about different aspects of the paper and the model proposed for the functional Jespersen cycle. The sections below can be grouped into three general categories.

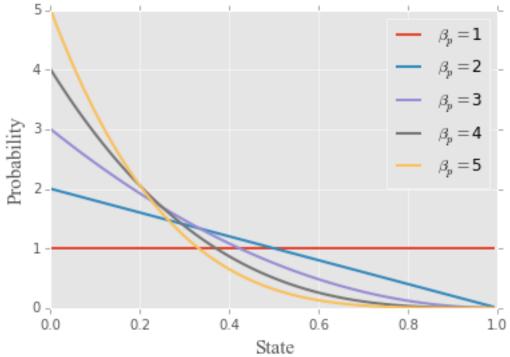
- **Model components**: here we discuss some of the model components in more detail and provide code for visualizations
- **Equilibria calculation**: here calculate the evolutionarily stable strategies of the signaling game defined in the document
- **Dynamics**: here we outline choices made in using the data, define the replicator dynamics in more detail, and fit the dynamic model to data

# 2 Model components

In this section we focus on the prior probability distribution over states and visualize the utility functions used in the game.

## 2.1 Prior probability

We use *beta distributions* as a prior probability over states T:[0,1], in the signaling game. The distribution is often written  $\mathcal{B}(\alpha,\beta)$  and is controlled by two shape parameters  $\alpha$  and  $\beta$ . Below we plot the prior where  $\alpha=1$  and  $\beta$  is free to vary.



For the dynamic model we consider a discretized version of the state space in order to keep fitting the model tractable. We use a *beta-binomial* over the states  $T:\{t_0,...,t_{n-1}\}$  and actions  $A:\{a_0,...,a_{n-1}\}$ , where  $t_i=a_i=\frac{i}{n}$ . In this case, we use one hundred states and actions, n=100 to approximate the beta distribution.

```
In [4]: from scipy.special import beta as beta_func
        from scipy.misc import comb
        def beta_binomial(n, alpha, beta):
            return np.array([comb(n - 1, k) * beta_func(k + alpha, n - 1 - k + beta) / \
                               beta_func(alpha, beta) for k in range(n)])
        plt.style.use('ggplot')
        for beta_var in range(1,6):
                 y = beta_binomial(len(x), 1, beta_var)
                 plt.plot(x, y, label=r'$\beta_p = $' + str(beta_var), linewidth=2)
        plt.legend(loc='upper right')
        plt.ylabel("Probability", fontsize=15, **hfont)
        plt.xlabel("State", fontsize=15, **hfont)
        plt.show()
           0.05 -
           0.04
       Probability
           0.03
                                                                       \beta_{\rm p} = 5
           0.02
           0.01
           0.00 -
                                        0.4
               0.0
                            0.2
                                                     0.6
                                                                  0.8
                                                                               1.0
                                             State
```

The only thing that changes when using the beta-binomial distribution is the scale of the y-axis because the beta-binomial is a probability mass function whereas the beta distribution is a

probability density function. Note that the probability for any discrete state  $t_i$  is approximately the value of the beta distribution at that point scaled by the number of discrete states.

$$p(t_i) \approx \frac{\mathcal{B}(\alpha, \beta)(t_i)}{n} \tag{1}$$

### 2.2 Utility functions

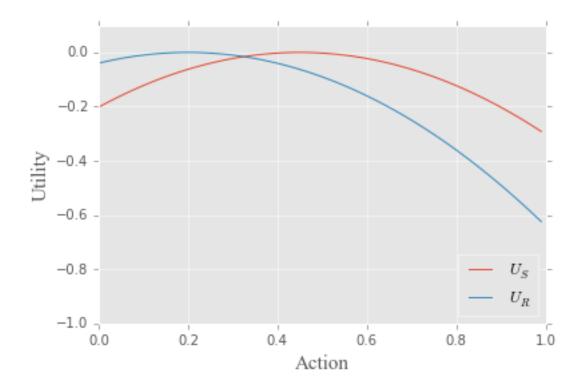
In the signaling game model we use a modified version of the payoffs defined in Crawford and Sobel (1982). The main motivation for taking these as a starting point is that it offers a means of modeling signaling where the state and action spaces are continuous. This is exactly what we want if we assume that standards of evidence and actions taken in response to them are continuous. More importantly, these utilities allow us to represent a preference for proximity between the standard of precision and the action taken.

Now, as a point of reference, the utilities used in Crawford and Sobel (1982) are the following.

$$U_S(s,r) = -(a-t-b)^2 U_R(s,r) = -(a-t)^2$$
 (2)

We can visualize these using the following code.

```
In [5]: x = np.arange (0, 1, 0.01)
# The actual standard of evidence
t = .2
# The degree of speaker bias
b = .25
U_S = [-(a - t - b)**2 for a in x]
U_R = [-(a - t)**2 for a in x]
plt.plot(x, U_S, label=r'$U_S$')
plt.plot(x, U_R, label=r'$U_R$')
plt.ylim(-1.0, .1)
plt.legend(loc='lower right')
plt.ylabel('Utility', fontsize=15, **hfont)
plt.xlabel("Action", fontsize=15, **hfont)
plt.show()
```



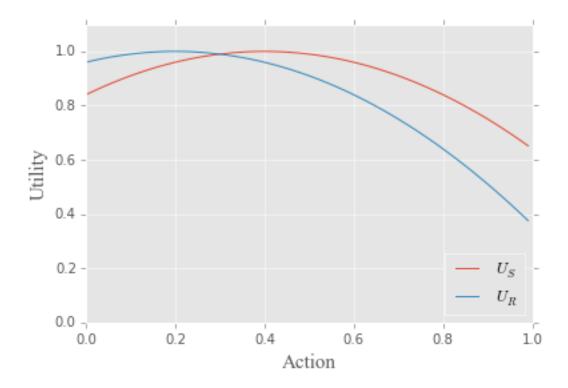
We modify these utility functions in two ways. First, we add a constant value so that the utilities are always positive and always fall in the unit interval. Second, we constrain speakers to only prefer actions that hearers can actually take. That is, speakers cannot prefer an action a > 1 for any value of t and t.

$$U_S(s,r) = 1 - (a - t - (1 - t)b)^2$$

$$U_R(s,r) = 1 - (a - t)^2$$
(3)

We can visualize these using the code below. Note that for these values, the sets of utility functions are slightly but not dramatically different.

```
In [6]: x = np.arange (0, 1, 0.01)
    # The actual standard of evidence
    t = .2
    # The degree of speaker bias
    b = .25
    U_S = [1 - (a - t - (1-t)*b)**2 for a in x]
    U_R = [1 - (a - t)**2 for a in x]
    plt.plot(x, U_S, label=r'$U_S$')
    plt.plot(x, U_R, label=r'$U_R$')
    plt.ylim(0, 1.1)
    plt.legend(loc='lower right')
    plt.ylabel('Utility', fontsize=15, **hfont)
    plt.xlabel("Action", fontsize=15, **hfont)
    plt.show()
```

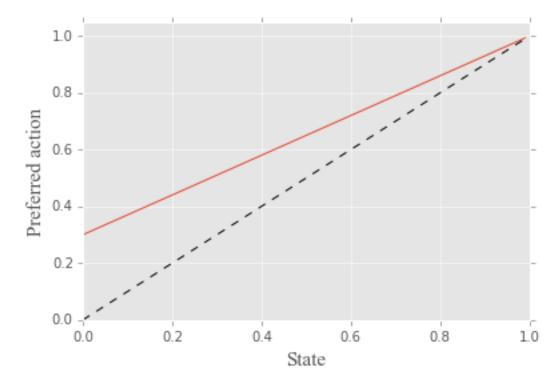


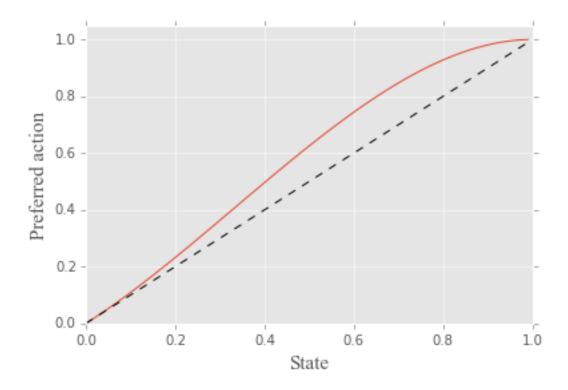
Now, there is nothing crucial about the linear relationship between speaker state and the speaker's preferred action on the part of the hearer the linear relationship between t and b. In fact, the relationship is also linear in the original formulation provided by Crawford and Sobel (1982): it has the same intercept, but a different slope. All that we have done is to constrain the values for those preferences so that speakers can't prefer non-available actions on the part of hearers. Note that we could just as well pick any arbitrary function such that t < f(t) < 1. Intuitively, in the unit square, draw a line continuously right that never goes below the diagonal and ends in the upper right. Any such function will do. Now, the function we have chosen stems from constraining speaker preferences to possible actions, but it also offers a convenient functional form for working with.

We could just as well define any function that models speakers preferring higher actions than the given state, while also being constrained to actions that exist. For example, consider the functions below.

```
In [7]: x = np.arange (0, 1, 0.01)
    b = .3
    plt.plot(x, x + (1-x)*b)
    plt.ylim(0, 1.05)
    plt.plot(x, x, 'k--')
    plt.ylabel('Preferred action', fontsize=15, **hfont)
    plt.xlabel("State", fontsize=15, **hfont)
    plt.show()
    x = np.arange (0, 1, 0.01)
    b = .3
    y = -x**3 + x**2 + x
```

```
plt.plot(x, y)
plt.ylim(0, 1.05)
plt.plot(x, x, 'k--')
plt.ylabel('Preferred action', fontsize=15, **hfont)
plt.xlabel("State", fontsize=15, **hfont)
plt.show()
```





# 3 Evolutionarily stable strategies of the signaling game

Calculating the evolutionarily stables strategies of the signaling game requires defining the expected utilities of speakers and hearers, taking partial derivatives, and finding the values that maximize the expected utilities. Note that this analysis is done using the continuous state space, whereas for the dynamic model we assume a discretized version of the state space.

Before moving on to the analysis, we note that the strategies we have defined are the only ones that can constitute evolutionarily stable strategies. For example, one might wonder whether mixed strategies need to be considered. Jäger et al. (2011, Lemma 2) suffices to show that mixed hearer strategies cannot ever be part of a strict Nash equilibria. It remains to show that mixed strategies for speakers cannot be either. To do so, we show that in all states it is always strictly worse to mix signals.

First, we define a mixed speaker strategy  $\sigma$  as a probability measure over speaker strategies where  $p_i$  is the probability of strategy  $s_i$ , such that  $\int p_i dp_i = 1$ . Now, we fix a state t and note that for a pair of pure sender and receiver strategies the function  $f: a \to U_s(t, a) = 1 - (r(s(t)) - t - (1-t)b)^2$ , is strictly concave.

Second, note that for a fixed state, a mixed speaker strategy  $\sigma$ , and a pure hearer strategy, the probability of an action being taken is determined by the probability a message being sent. Let  $g_k(t,s)$  be an indicator function that is equal to one when s sends message  $m_k$  in state t, and zero otherwise. The probability that an action  $a_k$  will be taken in a fixed state t, given a mixed speaker strategy  $\sigma$  and a pure hearer strategy r, is the following.

$$\gamma_k = \int g_k(t, s_i) p_i dp_i \tag{4}$$

Thus, the utility of a mixed speaker strategy for a fixed state and pure hearer strategy is the following.

$$E[f(a)] = \sum_{i} \gamma_i \left( 1 - (a_i - t - (1 - t)b)^2 \right)$$
 (5)

Finally, by Jensen's inequality, for any concave function E[f(a)] < f(E[a]). Here  $E[a] = \sum_i \gamma_i a_i$ , thus we have the following.

$$\sum_{i} \gamma_{i} \left( 1 - (a_{i} - t - (1 - t)b)^{2} \right) \le 1 - \left( \left( \sum_{i} \gamma_{i} a_{i} \right) - t - (1 - t)b \right)^{2} \tag{6}$$

This inequality is strict unless there is a single action used. That is, a mixed speaker strategy always does worse than a strategy where there is some action such that  $\alpha_i = 1$ . This means that in each state it is always strictly worse to mix messages.

So, the only speaker strategies that can strictly maximize expected utilities send a single message per state for all states. Since mixing signals is strictly worse, only pure strategies can be components of strict Nash equilibria and thus evolutionarily stable strategies. Thus, only speaker strategies that partition the state space in the manner described in the text can be components of evolutionarily stable strategies.

Moving on to our analysis, first, we import sympy, which is a library for symbolic math in python.

Next, we define the symbols that we'll use in constructing the utility functions and the prior probability.

Then, we build the actual utility functions and the respective expected utilities. Note that we can save a bit of effort by noting that the speaker's utility is defined by the hearer's response and the hearer's utility is the speaker's where b=0.

```
In [10]: Utility_S_0 = 1 - (a_0 - t - (1-t)*b)**2
Utility_S_1 = 1 - (a_1 - t - (1-t)*b)**2
```

From the utility functions we can calculate the expected utilities.

Now, we can differentiate the expected utility functions by the actions available to speakers and hearers, and solve for the values of  $t_0$ ,  $a_0$ ,  $a_1$  that maximize the expected utilities.

This can take a bit of time, so by default we've supplied the set of solutions to this system of equations. Note that there are two solutions because we haven't supplied the information to the solver that the values are constrained to be positive. In fact, we are only interested in the second solution that yields an ESS.

```
ESS = solve([t0_sol, a0_sol, a1_sol], [t_0, a_0, a_1])
```

We load this solution from disk that we've run and saved before.

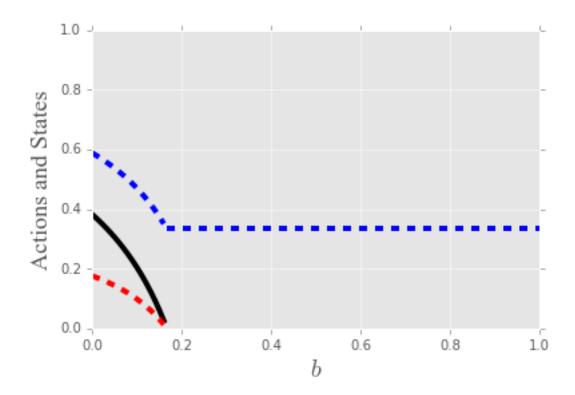
```
In [13]: import pickle
     with open('ESS.pkl', 'rb') as f:
          ESS = pickle.load(f)
```

We can visualize this result in the following manner. First, we note that speakers only use a single message if bias is sufficiently large.

```
In [14]: b_critical = solve(ESS[1], b)[0]
    b_critical

Out[14]: 0.166666666666667

In [15]: x = np.arange (0, b_critical, 0.01)
    plt.style.use('ggplot')
    plt.plot(x, [ESS[0].subs(b, value) for value in x], 'k', linewidth=4)
    plt.plot(x, [ESS[1].subs(b, value) for value in x], 'r', linewidth=4, linestyle='--')
    plt.plot(x, [ESS[2].subs(b, value) for value in x], 'b', linewidth=4, linestyle='--')
    plt.axhline(1/3.0, b_critical, 1, color='b', linewidth=4, ls='--')
    plt.ylim(0,1)
    plt.xlim(0,1)
    plt.xlabel(r"$b$", fontsize=18, **hfont)
    plt.ylabel(r"Actions and States", fontsize=18, **hfont)
    plt.savefig("../local/out/ESS-beta.pdf", format='pdf', dpi=1000, fontsize=18)
    plt.show()
```



We can find the values of the variables for a given value of b. For example, for b = 0, we can do the following.

```
In [16]: [item.subs(b, 0).evalf() for item in ESS]
Out[16]: [0.381966011250105, 0.175954681666807, 0.587977340833403]
```

# 4 Dynamics

# 4.1 Visualizing and formatting the data

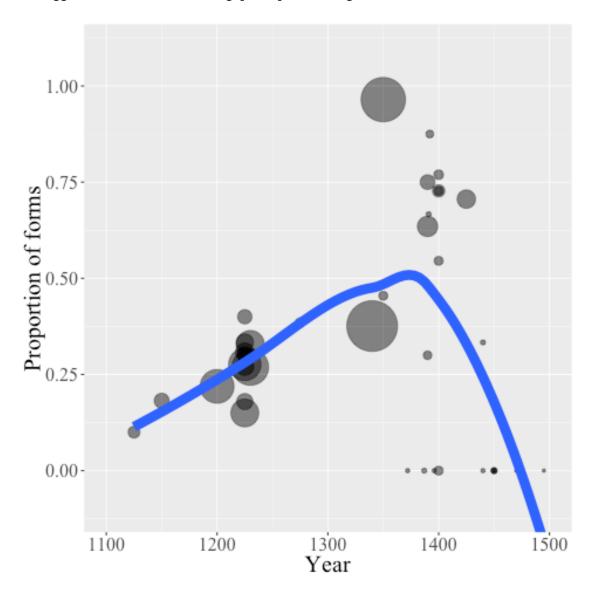
We load the data into R, plot the data and output in a format that can be used to fit data to the functional cycle.

Here we filter out texts that are known to be outliers and exclude everything but sentential negation in declaratives.

```
In [22]: %%R
         # Exclude texts that are known to be outliers
         excluded.texts = c("CMORM","CMBOETH","CMNTEST","CMOTEST")
         # Filter out tokens without do-support label, year, or type
         neg.data = neg.data.full %>%
                     filter(finite != "-") %>% # Exclude non-finite clauses
                     filter(clausetype != "imperative") %>% # Exclude imperatives
                     filter(exclude != "only") %>% # Exclude focus constructions
                     filter(exclude != "constituent") %% # Exclude constituent negation
                     filter(exclude != "contraction") %>% # Exclude contraction
                     filter(exclude != "coordination") %>% # Exclude coordinated clauses
                     filter(exclude != "concord") %>% # Exclude cases of negative concord
                     filter(exclude != "X") %>% # Exclude corpus errors
                     filter(! author %in% excluded.texts) %>% # Exclude texts
                     mutate(stages = ifelse(has.both, 2, ifelse(has.ne, 1, 3))) %>%
                     select(year, author, stages)
```

We plot the data for all three variants over the course of Middle English.

```
print(p)
ggsave('../local/out/neg-plot.pdf', height=6, width=8)
```



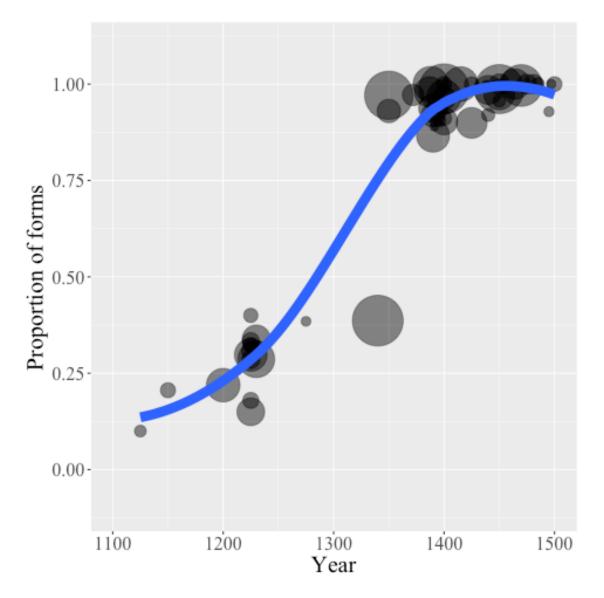
We also plot the data treating post-verbal tokens *as if* they were bipartite tokens in order to model the functional cycle.

```
In [24]: %%R
    # Compare ne to ne...not and not
    first.data = neg.data %>% group_by(year) %>%
        mutate(value = as.integer(! stages==1)) %>%
        select(year, author, value)

first.plot.data = first.data %>%
        group_by(year, author) %>%
```

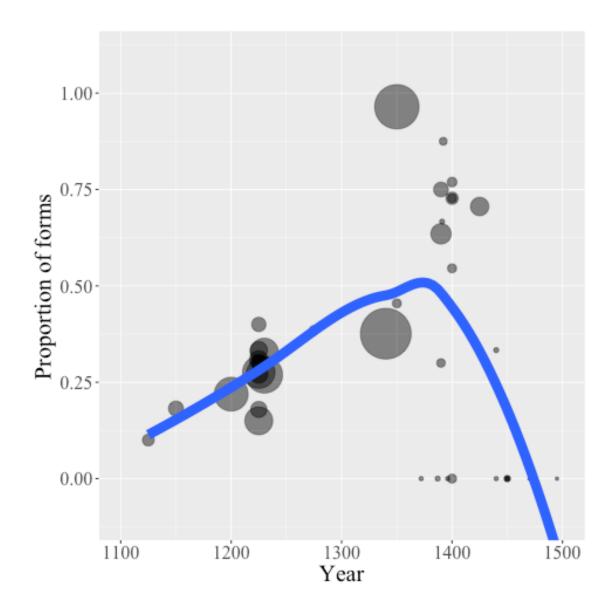
```
summarize(p = sum(value)/n(), total=n())
```

```
p = ggplot(aes(x = year, y = p), data = first.plot.data) +
  geom_point(aes(size = total), alpha = 0.5, position = "identity") +
  geom_smooth(method="loess", se = F, size=4) +
  scale_x_continuous(name="Year", limits=c(1100, 1500)) +
  scale_y_continuous(name="Proportion of forms", breaks=seq(0,1,.25)) +
  theme(text = element_text(size=20, family="Times New Roman")) +
  theme(legend.position="none") +
  scale_size_area("N", max_size = 20) +
  coord_cartesian(xlim = c(1090,1540)) +
  coord_cartesian(ylim = c(-.1,1.1))
  print(p)
  ggsave('../local/out/func-plot.pdf', height=6, width=8)
```



Note that if we were to compare only *ne* and *ne...not*, excluding *not* entirely from our analysis, we would run the risk of attributing too much to noisy fluctuations after 1350 when *ne* and *ne...not* combined cease to be the majority of forms. Indeed, after 1350 *ne* becomes more frequent than *ne...not* again.

```
In [25]: %%R
         # Compare ne to ne...not and not
         exclude.data = neg.data %>%
           filter(! stages == 3) %>%
           group_by(year) %>%
           mutate(value = as.integer(! stages==1)) %>%
           select(year, author, value)
         exclude.plot.data = exclude.data %>%
           group_by(year, author) %>%
           summarize(p = sum(value)/n(), total=n())
         p = ggplot(aes(x = year, y = p), data = exclude.plot.data) +
           geom_point(aes(size = total), alpha = 0.5, position = "identity") +
           geom_smooth(method="loess", se = F, size=4) +
           scale_x_continuous(name="Year", limits=c(1100, 1500)) +
           scale_y_continuous(name="Proportion of forms", breaks=seq(0,1,.25)) +
           theme(text = element_text(size=20, family="Times New Roman")) +
           theme(legend.position="none") +
           scale_size_area("N", max_size = 20) +
           coord_cartesian(xlim = c(1090, 1540)) +
           coord_cartesian(ylim = c(-.1,1.1))
         print(p)
```



Finally, we output the data in a format that will make it easy to calculate a loss function in a vectorized format.

}
write.csv(functional.cycle.data, "../data/functional-cycle-data.csv", row.names=F)

### 4.2 Defining the evolutionary game dynamics

Next, we define the discrete-time version of the behavioral replicator dynamics outlined in Huttegger and Hofbauer (2015). The basic intuition is that we treat each state and each message *as if* it were its own population. This reduces calculating the dynamics to matrix multiplication.

First, we define the payoff matrices for senders and receivers, which depend solely on the utility functions of senders and receivers respectively.

**A** is an  $n \times n$  matrix such that  $\mathbf{A}_{ij} = U_S(t_i, a_j)$ :

$$\mathbf{A} = \begin{pmatrix} U_{S}(t_{1}, a_{1}) & \cdots & U_{S}(t_{1}, a_{j}) & \cdots & U_{S}(t_{1}, a_{n}) \\ \vdots & \ddots & \vdots & & \vdots \\ U_{S}(t_{i}, a_{1}) & \cdots & U_{S}(t_{i}, a_{j}) & \cdots & U_{S}(t_{i}, a_{n}) \\ \vdots & & \vdots & \ddots & \vdots \\ U_{S}(t_{n}, a_{1}) & \cdots & U_{S}(t_{n}, a_{j}) & \cdots & U_{S}(t_{n}, a_{n}) \end{pmatrix}$$
(7)

**B** is an  $n \times n$  matrix such that  $\mathbf{B}_{ij} = U_R(t_i, a_j)$ :

$$\mathbf{B} = \begin{pmatrix} U_{R}(t_{1}, a_{1}) & \cdots & U_{R}(t_{1}, a_{j}) & \cdots & U_{R}(t_{1}, a_{n}) \\ \vdots & \ddots & \vdots & & \vdots \\ U_{R}(t_{i}, a_{1}) & \cdots & U_{R}(t_{i}, a_{j}) & \cdots & U_{R}(t_{i}, a_{n}) \\ \vdots & & \vdots & \ddots & \vdots \\ U_{R}(t_{n}, a_{1}) & \cdots & U_{R}(t_{n}, a_{j}) & \cdots & U_{R}(t_{n}, a_{n}) \end{pmatrix}$$
(8)

Second, we define the speaker and hearer populations.

**X** is a stochastic population matrix such that the proportion of the population in  $x_i$  using  $m_j$  is  $x_{ij}$ , with  $\sum_i x_{ij} = 1$ .

$$\mathbf{X} = \begin{pmatrix} x_{11} & x_{12} & x_{13} \\ \vdots & \vdots & \vdots \\ x_{i1} & x_{i2} & x_{i3} \\ \vdots & \vdots & \vdots \\ x_{n1} & x_{n2} & x_{n3} \end{pmatrix}$$
(9)

Intuitively, each row corresponds to a given state. Each element in the row corresponds to the proportion of use in that population. Each row sums to one because the proportion using the various signals must sum to one.

**Y** is a population matrix such that the proportion of the population in  $y_i$  responding with action  $a_j$  is  $y_{ij}$ , with  $\sum_i y_{ij} = 1$ .

$$\mathbf{Y} = \begin{pmatrix} y_{11} & \cdots & y_{1j} & \cdots & y_{1n} \\ y_{21} & \cdots & y_{2j} & \cdots & y_{2n} \\ y_{31} & \cdots & y_{3i} & \cdots & y_{3n} \end{pmatrix}$$
(10)

Again, intuitively, each row corresponds to a given message. Each element in the row corresponds to the proportion of different responses to the message. Each row sums to one because the proportion using the various responses must sum to one.

**P** is a stochastic matrix such that  $\forall i$ **P** $_i = P(t_1), ..., P(t_n)$ . That is, **P** is just n rows of the prior probability distribution over states.

$$\mathbf{P} = \begin{pmatrix} P(t_1) & \cdots & P(t_i) & \cdots & P(t_n) \\ \vdots & & \vdots & & \vdots \\ P(t_1) & \cdots & P(t_i) & \cdots & P(t_n) \end{pmatrix}$$
(11)

With slight abuse of notation, the expected utility of sending message  $m_j$  in state  $t_i$ , where  $\mathbf{Y}^T$  is the transpose of  $\mathbf{Y}$ :

$$E[U_S(\mathbf{X}_{ij}, \mathbf{Y})] = (\mathbf{A}\mathbf{Y}^T)_{ij}$$
(12)

The average expected utility in a speaker population  $x_i$ :

$$E[U_S(\mathbf{X}_i, \mathbf{Y})] = \sum_{i} \mathbf{X}_{ij} E[U_S(\mathbf{X}_{ij}, \mathbf{Y})] = (\mathbf{X}(\mathbf{A}\mathbf{Y}^T)^T)_{ii}$$
(13)

Let  $\hat{\mathbf{X}}$  be the speaker expected utility matrix normalized by the average expected utilities such that:

$$\hat{\mathbf{X}}_{ij} = \frac{(\mathbf{A}\mathbf{Y}^T)_{ij}}{(\mathbf{X}(\mathbf{A}\mathbf{Y}^T)^T)_{ij}}$$
(14)

For both populations, under the discrete-time replicator dynamics strategies grow in proportion to the amount by which they exceed the average payoff in the population. The discrete-time replicator dynamic for message  $m_i$  in state  $t_i$ :

$$\mathbf{X}'_{ij} = \mathbf{X}_{ij} \frac{E[U_S(\mathbf{X}_{ij}, Y)]}{E[U_S(\mathbf{X}_{i\prime}, Y)]}$$
(15)

The speaker population at the next point in time is then given by the following, where  $\otimes$  indicates the element-wise Hadamard product:

$$\mathbf{X}' = \mathbf{X} \otimes \mathbf{\hat{X}} \tag{16}$$

Now that we have defined the discrete-time replicator dynamics for the speaker population, we can do the same for the hearer population with a few additions. Let  $\mathbf{C}$  be the conditional probability of a state given a message. That is,  $\mathbf{C}_{ij} = P(t_i|m_j)$ , where  $\otimes$  indicates element-wise Hadamard multiplication and  $\otimes$  indicates the element-wise Hadamard division.

$$\mathbf{C} = (P^T \otimes X) \oslash (PX) \tag{17}$$

The expected utility of hearer responding to message  $m_i$  with action  $a_i$ :

$$E[U_R(\mathbf{X}, \mathbf{Y}_{ij})] = (\mathbf{B}^T \mathbf{C})_{ji}$$
(18)

Since the resulting matrix is  $n \times m$ , we swap the indices to get the appropriate value. Each column corresponds to a hearer population, and each row corresponds to a response action.

The average expected utility in a hearer population  $y_i$ :

$$E[U_R(\mathbf{X}, \mathbf{Y}_i)] = \sum_j \mathbf{Y}_{ij} E[U_R(\mathbf{X}, \mathbf{Y}_{ij})] = (\mathbf{Y}(\mathbf{B}^T \mathbf{C})_{ji})_{ii}$$
(19)

Let  $\hat{\mathbf{Y}}$  be the population normalized matrix for hearer expected utilities such that:

$$\hat{\mathbf{Y}}_{ji} = \frac{(\mathbf{B}^T \mathbf{C})_{ji}}{(\mathbf{Y}(\mathbf{B}^T \mathbf{C})_{ii})_{ii}}$$
(20)

The discrete-time replicator dynamic for action  $a_i$  in response to message  $m_i$ :

$$\mathbf{Y}'_{ij} = \mathbf{Y}_{ij} \frac{E[U_R(\mathbf{X}, \mathbf{Y}_{ij})]}{E[U_R(\mathbf{X}, \mathbf{Y}_i)]}$$
(21)

The hearer population at the next point in time is then given by the following:

$$\mathbf{Y}' = \mathbf{Y} \otimes \mathbf{\hat{Y}}^T \tag{22}$$

In [27]: def discrete\_time\_replicator\_dynamics(n\_steps, X, Y, A, B, P):

"""Simulate the discrete-time replicator dynamics.

```
Parameters
_____
n_steps : int, the number of discrete time steps to simulate
X : stochastic sender matrix
Y : stochastic receiver matrix
A : sender utility matrix
B : receiver utility matrix
P : prior probability over states matrix
Returns
_____
X_t : array-like, the state of the sender population at each year
Y_{-}t: array-like, the state of the receiver population at each year
# Get the number of states
X_nrow = X.shape[0]
# Get the number of messages
X_{ncol} = X.shape[1]
# Get the number of actions
Y_nrow = Y.shape[0]
Y_ncol = Y.shape[1]
# Create empty arrays to hold flattened matrices for the population over time
X_t = np.empty(shape=(n_steps, X_nrow*X_ncol), dtype=float)
Y_t = np.empty(shape=(n_steps, X_nrow*X_ncol), dtype=float)
# Set the initial state
X_t[0,:] = X.ravel()
Y_t[0,:] = Y.ravel()
# Iterate forward over (n-1) steps
for i in range(1,n_steps):
    # Get the previous state
    X_prev = X_t[i-1,:].reshape(X_nrow, X_ncol)
    Y_prev = Y_t[i-1,:].reshape(Y_nrow, Y_ncol)
```

```
# Calculate the scaling factors
E_X = A * Y_prev.T
X_bar = (((A * Y_prev.T) * X_prev.T).diagonal()).T
X_hat = E_X / X_bar
# Calculate probability of states given messages
C = np.divide(np.multiply(P.T, X_prev), (P * X_prev)[0])
E_Y = (B.T * C).T
Y_bar = ((E_Y*Y_prev.T).diagonal()).T
Y_hat = np.divide(E_Y, Y_bar)
# Calculate current states
X_t[i,:] = np.multiply(X_prev, X_hat).ravel()
Y_t[i,:] = np.multiply(Y_prev, Y_hat).ravel()
return X_t, Y_t
```

Now that we have defined the dynamics, we define several functions to construct the model, including one that takes the parameters of the model and constructs the initial state.

```
In [28]: # Define various components used to construct the model
         import numpy as np
         import pandas as pd
         from scipy.optimize import minimize
         from scipy.optimize import brute
         from scipy.optimize import fmin
         from scipy.special import beta as beta_func
         from scipy.special import binom
         from scipy.misc import comb
         from scipy.stats import chi2
         from functools import partial
         import matplotlib.pyplot as plt
         plt.style.use('ggplot')
         hfont = {'fontname':'Times New Roman'}
         def beta_binomial(alpha, beta, n=100):
             return np.matrix([comb(n-1,k) * beta_func(k+alpha, n-1-k+beta) / beta_func(alpha,beta_func)
                                  for k in range(n)])
         def U_S(state, action, b):
             return 1 - (action - state - (1-state)*b)**2
         def U_R(state, action):
             return 1 - (action - state) **2
```

def t(i, n):

def a(i, n):

return i/float(n)

return i/float(n)

```
def sender_matrix(b, number=100):
    return np.matrix([[U_S(t(i, number-1), a(j,number-1), b)
                       for j in range(number)] for i in range(number)])
def receiver_matrix(number=100):
    return np.matrix([[U_R(t(i, number-1), a(j,number-1))
                       for j in range(number)] for i in range(number)])
def construct_initial_state(a_s, b_p, b=0):
    """Construct the initial state of the model.
    Parameters
    _____
    a_s, b_p, b : parameters defined in document
    Returns
    XO: array-like, the initial state of the speaker population
    YO: array-like, the initial state of the hearer population
    prior: prior probability over states
    # Define prior probability
    a_p = 1
    prior = beta_binomial(a_p, b_p)
   P = np.repeat(prior, 2, axis=0)
    # Define payoff matrices
    A = sender_matrix(b)
    B = receiver_matrix(
    # Define speaker population
    X0_m2 = beta_binomial(a_s, 1)
    XO_m1 = 1 - XO_m2
    X0 = np.vstack((X0_m1, X0_m2)).T
    # Calculate probability of state given m2
    p_ti_m2 = np.multiply(X0[:,1], prior.T)
   p_m2 = prior * X0[:,1]
   p_t_m2 = p_t_m2 / p_m2
    # Calculate probability of state given m1
   p_ti_m1 = np.multiply(X0[:,0], prior.T)
   p_m1 = prior * X0[:,0]
   p_t_m1 = p_t_m1 / p_m1
    # Calculate expected utility for receiver of action given m1
    E_ai_m1 = p_t_m1.T * B
    E_a_m1 = E_ai_m1 / E_ai_m1.sum()
    # Calculate expected utility for receiver of action given m2
    E_ai_m2 = p_t_m2.T * B
    E_a_m2 = E_ai_m2 / E_ai_m2.sum()
    # Define hearer population
    Y0 = np.vstack([E_a_m1, E_a_m2])
```

```
return XO, YO, A, B, prior
```

Now we define a function that takes a set of parameters that define a starting state and simulates the dynamics for a specified amount of time.

```
In [29]: def simulate_dynamics(params, n_years=401, time_scale=1, number=100):
             """Simulate the discrete-time behavioral replicator dynamics for the game.
             Parameters
             n_years : int, the number of discrete time steps to simulate
             time_scale : int (optional), the number of discrete time steps per year
             number: int, (optional), the number of discretized states and actions
             params: array-like, parameters that determine starting state of population
             Returns
             X_sol : array-like, the state of the speaker population at each year
             Y_{-}sol: array-like, the state of the hearer population at each year
             prior : prior probability over states
             # Unpack the parameters
             a_s, b_p, b = params
             # Construct the initial state
             X0, Y0, A, B, prior = construct_initial_state(a_s, b_p, b)
             # Create prior probability matrix
             P = np.repeat(prior, 2, axis=0)
             # Iterate using dynamics to get values for the number of years
             X_sol, Y_sol = discrete_time_replicator_dynamics(
                                     n_years*time_scale, XO, YO, A, B, P)
             X_sol = X_sol[0::time_scale,:]
             Y_sol = Y_sol[0::time_scale,:]
             return X_sol, Y_sol, prior
```

### 4.3 Fitting the dynamic model to historical corpus data

Now that we have defined the dynamics, we import the data from the functional cycle and define a loss function to minimize. In this case the loss function is the negative log-likelihood of the parameters.

```
In [30]: func_data = pd.read_csv('../data/functional-cycle-data.csv')

def loss_function(params, func=simulate_dynamics, time_scale=1, df=func_data):
    """Calculate the loss function.

Parameters
------
params : array-like, parameters that determine starting state of population
```

```
func : model begin fit to the data
time_scale : int (optional), the number of discrete time steps per year
number: int, (optional), the number of discretized states and actions
         default is one "generation" of the dynamics per year
df : data to use, default is from func_data
Returns
negLL: float, negative log likelihood to be minimized
# Simulate the dynamics from 1100 to 1500
X_sol, Y_sol, prior = func(params, n_years=401, time_scale=time_scale)
# Get p(m_2) over time
m2_sol = np.asarray([prior.dot(line)[0,0] for line in X_sol[:,1::2]])
# Append solution trajectory to data frame
# Use approximate value when trajectory reaches boundary
df['p'] = np.minimum(m2_sol.ravel(), np.repeat(.9999999, len(m2_sol)))
# Add binomial coefficient
df['binom'] = binom(df.ones + df.zeros, df.ones)
# Calculate log-likelihood for
df['LL'] = np.log(df.binom) + (df.ones * np.log(df.p)) + (df.zeros * np.log(1 - df.
# Only use years that have tokens
df = df[df['has.tokens'] == 1]
# Calculate log likelihood given m2_sol
LL = np.sum(df['LL'])
negLL = -1*LL
# Minimizing negative log-likelihood is equivalent to maximizing log-likelihood
return negLL
```

Here we fit the model to the data and print the maximum likelihood parameters.

We simulate proportion of the incoming form using these parameters

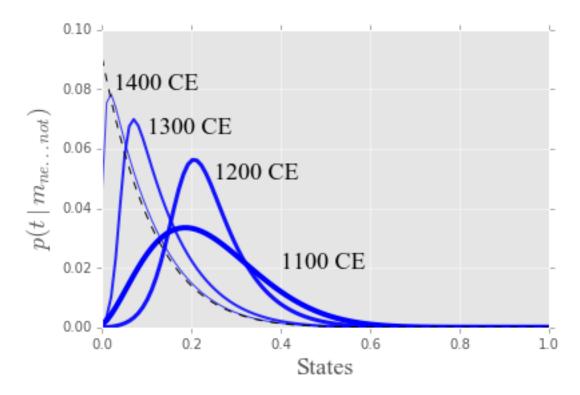
```
In [32]: X_sol, Y_sol, prior = simulate_dynamics(model_results)
```

We visualize the results by looking at the proportion of the incoming form over time, the conditional probability of states over time, and the KL-divergence of the form over time.

```
In [33]: states = np.linspace(0,1, num=100)
         timesteps=401
         m2_sol = [prior.dot(line)[0,0] for line in X_sol[:,1::2]]
         years = [1100 + item for item in range(401)]
         plt.plot(years, m2_sol, 'b', linewidth=4, zorder=3)
         plt.ylim(-.1,1.1)
         plt.xlim(1075, 1525)
         plt.xlabel('Year', fontsize=18, **hfont)
         plt.yticks(np.linspace(0, 1, num=5))
         plt.ylabel(r'$p(m_{ne...not})$', fontsize=18, **hfont)
         plt.savefig("../local/out/p-ne-not.pdf", format='pdf', dpi=1000)
         plt.show()
           1.00 -
            0.75 -
            0.50 -
            0.25
            0.00 -
                                1200
                                             1300
                                                           1400
                                                                         1500
                  1100
```

Year

```
plt.text(.1, .065, str(1300) + " CE", fontsize=18, **hfont)
plt.text(.025, .08, str(1400) + " CE", fontsize=18, **hfont)
plt.plot(states, prior.tolist()[0], 'k--')
plt.ylabel(r'$p(t \mid m_{ne...not})$', fontsize=18, **hfont)
plt.xlabel('States', fontsize=18, **hfont)
plt.savefig("../local/out/pt-ne-not.pdf", format='pdf', dpi=1000)
plt.show()
```



### 4.4 Comparison with simplified model

1100

0.5

0.0

We can assess whether positing a bias on the part of speakers is justified by comparing the full dynamic model to a simplified model where b = 0 and  $\alpha_s$  and  $\beta_p$  are free to vary.

1200

```
In [37]: def simulate_simplified_dynamics(params, n_years=401, time_scale=1, number=100):
    """Simulate simplified dynamic model."""
    # Unpack the parameters
    a_s, b_p = params
    # Construct the initial state
    X0, Y0, A, B, prior = construct_initial_state(a_s, b_p)
    # Create prior probability matrix
    P = np.repeat(prior, 2, axis=0)
    # Iterate using dynamics to get values for the number of years
    X_sol, Y_sol = discrete_time_replicator_dynamics(n_years*time_scale, X0, Y0, A, B, X_sol = X_sol[0::time_scale,:]
    Y_sol = Y_sol[0::time_scale,:]
    return X_sol, Y_sol, prior
```

1300

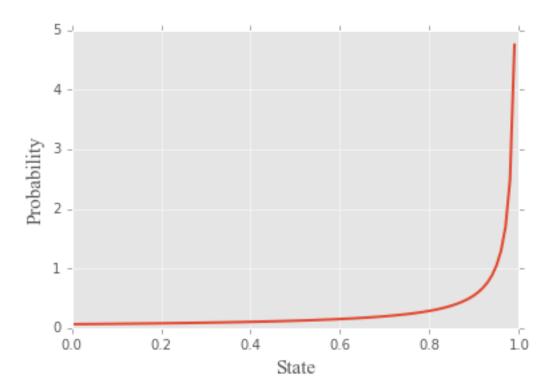
Year

1400

1500

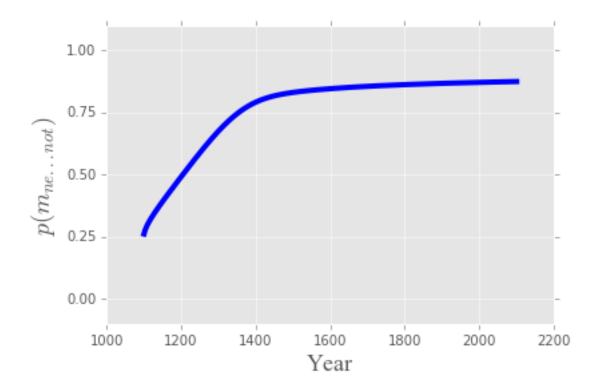
One problem with the simplified model is that it predicts that speakers are almost always absolutely sure of what they say. Below is the prior probability over states for the parameters inferred for the simplified model.

```
In [39]: from scipy.stats import beta
    x = np.arange (0, 1, 0.01)
    b_p = simplified_results[1]
    plt.style.use('ggplot')
    y = beta.pdf(x, 1, b_p)
    plt.plot(x, y, linewidth=2)
    plt.ylabel("Probability", fontsize=15, **hfont)
    plt.xlabel("State", fontsize=15, **hfont)
    plt.show()
```



As we would expect from the equilibrium analysis above, in the absence of a speaker bias, the functional cycle doesn't go to completion by 1500. Indeed, the use of the new form would still be approaching its equilibrium value in the present day.

```
In [40]: X_sol, Y_sol, prior = simulate_simplified_dynamics(simplified_results, n_years=1001)
         m2_sol = [prior.dot(line)[0,0] for line in X_sol[:,1::2]]
         years = [1100 + item for item in range(1001)]
         plt.plot(years[:401], m2_sol[:401], 'b', linewidth=4, zorder=3)
         plt.ylim(-.1,1.1)
         plt.xlim(1075, 1525)
         plt.xlabel('Year', fontsize=18, **hfont)
         plt.yticks(np.linspace(0, 1, num=5))
         plt.ylabel(r'$p(m_{ne...not})$', fontsize=18, **hfont)
         plt.savefig("../local/out/p-ne-not-simplified.pdf", format='pdf', dpi=1000)
         plt.show()
         plt.plot(years, m2_sol, 'b', linewidth=4, zorder=3)
         plt.ylim(-.1,1.1)
         plt.xlabel('Year', fontsize=18, **hfont)
         plt.yticks(np.linspace(0, 1, num=5))
         plt.ylabel(r'$p(m_{ne...not})$', fontsize=18, **hfont)
         plt.show()
           1.00
           0.75
           0.50 -
           0.25
           0.00 -
                                1200
                                             1300
                  1100
                                                          1400
                                                                        1500
                                            Year
```



We can compare the full and the simplified model using the difference in AIC.

The probability that the simplified model is better than the full model is given by its Akaike weight.

```
In [42]: 1 / (1 + np.exp(.5*delta_AIC))
Out[42]: 1.8491678206998729e-248
```

To put this in perspective, this probability is roughly the magnitude of flipping a fair coin eight hundred times and it only ever coming up heads.

```
In [43]: .5**800
Out[43]: 1.499696813895631e-241
```

1140.8527347

Since the full and simplified model are nested, we can also perform a Likelihood Ratio test, where the test statistic D is  $\chi^2$  distributed with one degree of freedom. We reject the null hypothesis that the simplified model is correct.

```
In [44]: from scipy.stats import chi2
    D = 2*(full_LL - simple_LL)
    chi2_result = chi2.pdf(D,1)
    print chi2_result
8.02780427052e-251
```

# 5 References

Jager, Gerhard, Lars P Metzger & Frank Riedel. 2011. Voronoi languages: Equilibria in cheaptalk games with high-dimensional types and few signals. Games and economic behavior 73(2). 517–537. http://dx.doi.org/10.1016/j.geb.2011.03.008.