CAS CS 365 InClass Note 2

Probability

1. Fair Coin

- a. $Pr(H) = Pr(T) = \frac{1}{2}$
- b. Coin flip can be "rigged" → the process is deterministic (person tossing the coin can use the knowledge of coin to predict the outcome with better probability)
- c. Tossing of coin is fair only and only if $Pr(H) = \frac{1}{2}$
- d. Suppose we throw the fair coin ten times
 - What is the expected number of heads?

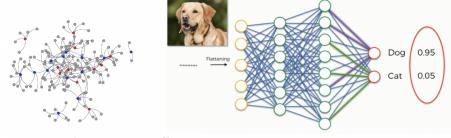
$$X_i = 1$$
 if the ith toss gives Head
= 0 if it gives Tail

 S_{10} = number of heads in 10 tosses = $X_1 + X_2 + ... + X_{10}$

E[S₁₀]
= E[
$$\Sigma$$
 i=1 to i = 10 of X]
= Σ i=1 to i = 10 E[X_i]
E[X_i] = 1 * ½ + 0* ½
 Σ i=1 to i = 10 E[X_i] = 10 * ½ = 5
X_i \rightarrow Ber(½)

$$S_{10} \rightarrow Binomial (10, \frac{1}{2})$$

2. Modeling uncertainty



Disease spreading

b. Information theory, modeling the reliability of numerous complex systems, insurance companies, investments, etc.

3. Reminder

a.

- a. Probability space (Ω, F, P) consists of three elements
- b. Sample space (Ω) : set of all possible outcomes

- c. Event Space (F): a set of events, an event being a set of outcomes in the sample space
- d. Probability (P): it is a function that assigns each event in the vent space a probability, which is a number between 0 and 1
- e. Example:

Example 2.29. Consider a sample space containing three elements $\Omega = \{\clubsuit, \heartsuit, \maltese\}$. The event space is then $\mathcal{F} = \{\emptyset, \{\clubsuit\}, \{\heartsuit\}, \{\clubsuit\}, \{\diamondsuit, \heartsuit\}, \{\heartsuit, \maltese\}, \{\clubsuit, \Psi\}, \{\clubsuit, \heartsuit, \maltese\}\}\}$. One possible $\mathbb P$ we could define would be

$$\begin{split} \mathbb{P}[\emptyset] &= 0, \quad \mathbb{P}[\{\clubsuit\}] = \mathbb{P}[\{\heartsuit\}] = \mathbb{P}[\{\maltese\}] = \frac{1}{3}, \\ \mathbb{P}[\{\clubsuit,\heartsuit\}] &= \mathbb{P}[\{\clubsuit,\maltese\}] = \mathbb{P}[\{\heartsuit,\maltese\}] = \frac{2}{3}, \quad \mathbb{P}[\{\clubsuit,\heartsuit,\maltese\}] = 1. \end{split}$$

space a probability, which is a number between 0 and 1.

4. Monty-hall problem

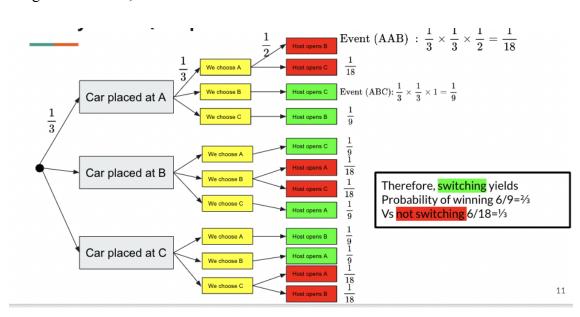
f.

- a. Suppose you're on a game show, and you're given the choice of three doors
 - Behind one door is a car; behind the other, goats
 - You pick a door, Say 1, and the host, who knows what's behind the doors, opens another door, say 3, which has a goat
 - He then says to you, "Do you want to pick door 2?" Is it to your advantage to switch your choice?

b. Assumptions:

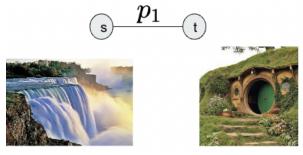
- Car is placed uniformly at random (uar) behind a door
- Out initial guess is also uar
- The host opens a door with a goat. When there exist two such doors, i.e., our guess is the car, he chooses uar

c. Steps



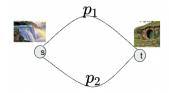
5. Transfer water

- a. Consider a water source s and a destination village t
- b. Each pipe i has probability of failure Pi, Pipes fail independently
- c. Question: What is the probability we cannot get water from s to t? In other words, when is the village t not reachable from the water source s?



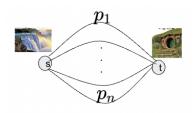
e. Exercise 1:

d.



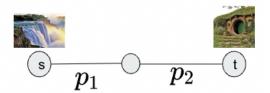
- Clearly, there is no path if both pipes fail
- Since they are independent, the probability of this event is the product of the probabilities of the individual events
- Thus, failure probability is $p_1 * p_2$

f. Exercise 2:



- Clearly, there is no path if all pipes fail
- Since they are independent, the probability of this event is the product of the probabilities of the individual events
- Thus, failure probability is $p_1 * p_2 * ... * p_n$
- E_i = pipe i fails
- $Pr(E_1 \ n \ E_2 \ n \ ... \ n \ E_n) = \prod Pr(E_n)$
- ** De Morgan's Law **
- Pr(not $E_1 U$ not $E_2 U ... U$ not E_n) = 1 $p_1 * p_2 *... * p_n$

g. Exercise 3:



- Two events are independent: $Pr(\text{not } E_1 \text{ n not } E_2) = Pr(\text{not } E_1) * Pr(\text{not } E_2)$
- $Pr(failure) = Pr(E_1 \cup E_2)$
 - $= 1 Pr(not (E_1 U E_2))$
 - = 1- Pr(not E_1 n not E_2) = 1 (1- p_1) * (1- p_2)
 - $= 1 (1 p_1 p_2 + p_1 * p_2)$
 - $= p_1 + p_2 p_1 * p_2$

OR

- Clearly, there is no path if at least one of the pipes fail
- We condition on whether the one of the two pipes (say the first) is broken or not
- Let A_i be the event that pipe i fails
- Then

$$egin{aligned} \Pr(A_1 \cup A_2) &= \Pr(A_1) + \Pr(A_2) - \Pr(A_1 \cap A_2) \ &= p_1 + p_2 - \Pr(A_1) \Pr(A_2) \ &= p_1 + p_2 - p_1 p_2 \end{aligned}$$

h. Exercise 4



- This is when all pipes are successful (if one pipe is unsuccessful, water will not go through)
- Instead of thinking of the probability that t will not be reachable from s, we think of the probability that it is. Reminder: Pr(not A) = 1 Pr(A)
- The probability of not failing is

$$\Pr\Bigl(\cap_{i=1}^nar{A_i}\Bigr)=\prod_{i=1}^n\Pr\Bigl(ar{A}_i\Bigr)=\prod_{i=1}^n(1-p_i)$$

- Therefore, the right answer is

$$1-\prod_{i=1}^n(1-p_i)$$

- Pr(sending water from s to t) = Pr(No Pipe fails)
 = Pr(not E₁ n not E₂ n ... n not E_n)
- $\Pi \operatorname{Pr}(\operatorname{not} E_i) = \Pi (1-\operatorname{Pr}(E_i))$
- 6. Reminders: independent events, conditional probability
 - a. Intuitive two events A, B are dependent if A's occurrence or non-occurrence provides us with some information about event B
 - b. Formally, A and B are independent events if and only iff

$$\Pr(A \cap B) = \Pr(A) \Pr(B)$$

c. By rearranging, we get

$$\Pr(A) = rac{\Pr(A \cap B)}{\Pr(B)}$$

d. Recall that by the law of conditional probability

$$\Pr(A|B) = rac{\Pr(A \cap B)}{\Pr(B)}$$

- e. Therefore, when A and B are independent events Pr(A) = Pr(A|B) and Pr(B) = Pr(B|A)
- 7. Reminders: Law of total probability
 - a. Let Ω be a probability space. Let $B_1,...,B_m$ be a partition of Ω . Then,

$$\Pr(A) = \sum_{i=1}^m \Pr(A \cap B_i) = \sum_{i=1}^m \Pr(B_i) \Pr(A|B_i)$$

- 8. Reminder: chain rule
 - a. Chain rule:

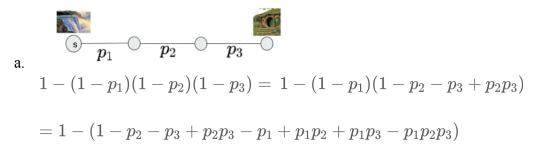
$$\Pr(A_1 \cap \ldots \cap A_n) = \Pr(A_1) \Pr(A_2 | A_1) \Pr(A_3 | A_2 A_1) \ldots \Pr(A_n | A_{n-1} \ldots A_1)$$

- b. In our case the events are mutually independent, so this simplifies to the product of the individual probabilities of the events A_i
- c. Pairwise Independent
 - Events A, B, C
 - P(A n B) = P(A n C) = P(B n C)
- d. Mutually Independent
 - Events A, B, C
 - $P(A \cap B \cap C) = P(A) * P(B) * P(C)$
- 9. Reminder: conditional probability + Law of total probability → Bayes rule

$$\Pr(B_i|A) = rac{\Pr(B_i \cap A)}{\Pr(A)} = rac{\Pr(B_i)\Pr(A|B_i)}{\sum_{j=1}^n\Pr(B_j)\Pr(B_j|A)}$$

a.

10. Example n = 3



b.
$$= p_1 + p_2 + p_3 - p_1p_2 - p_1p_3 - p_2p_3 + p_1p_2p_3$$

11. Example n = 4



a

b.
$$1 - (1-p_1)(1-p_2)(1-p_3)(1-p_4)$$

 $= 1 - (1-p_1-p_2+p_1p_2)(1-p_3-p_4+p_3p_4)$
 $= 1 - \dots$
 $= p_1 + p_2 + p_3 + p_4$
 $- p_1p_2 - p_1p_3 - p_1p_4 - p_2p_3 - p_2p_4 - p_3p_4$
 $+ p_1p_2p_3 + p_1p_2p_4 + p_2p_3p_4$
 $- p_1p_2p_3p_4$

12. Reminder: Inclusion exclusion

a. Another convenient way to write the IE formula is the following

$$\mathbf{P}igg(igcup_{i=1}^{n}A_{i}igg)=S_{1}-S_{2}+S_{3}-\ldots+(-1)^{n-1}S_{n}$$

wher

$$S_k = \sum_{1 \leq i_1 < i_2 < \ldots < i_k \leq n} \mathbf{P}(A_{i_1} \cap A_{i_2} \cap \ldots \cap A_{i_k}).$$

c. In our setting, due to the independence of the events A_i, we can write the following expression:

$$\Pr(\cup A_i) = \sum_{k=1}^n \left(-1
ight)^{k+1} \sum_{I\subseteq [n], |I|=k} \prod_{i\in I} \Pr(A_i)$$

let's write down some terms

$$egin{aligned} \Pr(\cup A_i) &= p_1 + \ldots + p_n \ &- (p_1 p_2 + \ldots + p_{n-1} p_n) \ &+ (p_1 p_2 p_3 + \ldots + p_{n-2} p_{n-1} p_n) \ &- \ldots \end{aligned}$$