

## Graphs

1. Goals of this week
  - a.  $G(n,p)$  model
  - b. First and second moment for phase transitions
  - c. Exercises on  $E[x]$ ,  $\text{Var}[x]$  where  $X$  is a RV in  $G$
2. Isolated Nodes
  - a. There exist a  $p^*$  value that if  $p > p^*$  in  $G(n,p)$ , there are no isolated nodes
  - b.  $E[x] = n(1-p)^{(n-1)}$ 
    - i.  $X_i = 1$  if node  $i$  is isolated, 0 otherwise
    - ii.  $X = X_1 + \dots + X_n$
    - iii. 1st moment method by using Markov's inequality
      1.  $\Pr(X > 1) < E[x]$  and if  $E[x]$  goes to 0 as  $n \rightarrow \text{infinity}$ ,  $\Pr(X > 0)$  is 0
      2. Therefore,  $\Pr(X = 0) = 1 - O(1)$
  - c. Asymptotic
    - i. For small  $p$ ,  $1-p = e^{-p}$   
 $(1-p) = e^{\ln(1-p)}$   
 $= e^{(-p + p^2/2 + \dots)}$  by using Taylor Series, which is  $e^{(-p)}$
    - ii.  $\lim_{n \rightarrow \text{infinity}} (1+1/n)^n = e^{(n \ln(1+1/n))} = e^1 = e$
  - d.  $E[x] = n(1-p)^{(n-1)} = n \cdot e^{(-p(n-1))}$   
 $= n \cdot e^{(-c \ln(n)/n \cdot (n-1))}$   
 $= n \cdot e^{(-c \ln(n))}$   
 $= n \cdot n^{(-c)}$   
 $= n^{(1-c)}$
  - e. So if  $c > 1$ , then  $E[x] = o(1)$ , so by Markov's inequality,  
 $\Pr(X > 1) \leq E(x) = O(1)$   
 Therefore,  $X = 0$  with probability  $1 - O(1)$
  - f.  $Z = 0$  if  $1-1/n = 1 - O(1)$   
 $= n^2$  if  $1/n$
  - g.  $E[z] = n$
  - h. If  $p < p^* = \ln(n)/n$ , then  $X > 0$  with high probability  
 $\Pr(X=0) \leq \Pr(|X-E[X]| \geq E[X]) \leq \text{Var}(X)/E[X]^2$  as  $n \rightarrow \text{infinity}$  by using Chebyshev inequality
  - i.  $\text{Var}(X) = \text{Var}(X_1 + \dots + X_n)$   
 $= \sum_{i=1}^n \text{Var}(X_i) + \sum_{(i \text{ does not equal } j)} \text{Cov}(X_i, X_j)$

$$\leq \sum_{i=1}^n E(X_i) + \sum_{(i \text{ does not equal } j)} \text{Cov}(X_i, X_j) \\ = E[X] + \sum_{(i \text{ does not equal } j)} \text{Cov}(X_i, X_j)$$

**\*\* proof \*\***

$$X_i = 1 \text{ with probability } (1-p)^{(n-1)} \\ = 0 \text{ with probability } 1-(1-p)^{(n-1)}$$

$$E[X_i] = (1-p)^{(n-1)}$$

$$\text{Var}[X_i] = (1-p)^{(n-1)} * (1-(1-p)^{(n-1)}) \leq (1-p)^{(n-1)} = E[X_i] \text{ since} \\ (1-(1-p)^{(n-1)}) \text{ is at most } 1$$


**\*\*End proof \*\***

$$\frac{\text{Var}(X)}{E[x]^2} \leq \frac{1}{E[x]} + \frac{\sum \text{Cov}()}{E[x]^2}$$

- j.  $\text{Cov}(X_i, X_j)$   
 $= E(X_i * X_j) - E[X_i] * E[X_j]$   
 $= E[X_i * X_j] - (1-p)^{(n-1)} * (1-p)^{(n-1)}$   
 $= E[X_i * X_j] - (1-p)^{2(n-1)}$   
 $= \text{Pr}[X_i = X_j = 1] - (1-p)^{2(n-1)}$   
 $= (1-p)^{(2(n-2)+1)} - (1-p)^{2(n-1)}$   
 $\leq (1-p)^{(2(n-2)+1)}$
- k.  $\sum_{(i \text{ does not equal } j)} \text{Cov}[X_i, X_j]$   
 $= n(n-1) * ((1-p)^{(2(n-2)+1)} - (1-p)^{2(n-1)})$   
 $\leq n * (n-1) * (1-p)^{(2(n-2)+1)}$
- l. Now, we have to prove  $\sum \text{Cov}() / E[x]^2$  goes to 0  

$$\frac{(n * (n-1) * (1-p)^{(2(n-2)+1)})^2}{n^2(1-p)^{(2(n-2))}} \text{ goes to } 0$$

3.

  $K_4 \rightarrow 4$  nodes and  $\binom{4}{2}$  edges




Let  $X$  to the # of  $K_4$ 's in  $G(n, p)$

$X = \sum_{i=1}^{\binom{n}{4}} x_i$  where  $x_i = \begin{cases} 1 & \text{if the } i\text{-th } 4\text{-clique is } \boxed{\times} \\ 0 & \text{otherwise} \end{cases}$

$$E[X] = \sum_{i=1}^{\binom{n}{4}} E(x_i) = \binom{n}{4} p^6$$

$\hookrightarrow$  since  $\Pr(x_i = 1) = p^6$  for all  $i$

  $\rightarrow 6$  edges connected

- what value of  $p$  makes  $\binom{n}{4} p^6$  go to 0?

$\hookrightarrow$  threshold value for  $p^*$

$$\binom{n}{4} = \frac{n(n-1)(n-2)(n-3)}{4!} \leq n^4$$

$$\binom{n}{4} p^6 \leq n^4 p^6 \rightarrow 0$$

$$n^4 p^6 \ll 1$$

$$p \ll \frac{n^{-2/3}}{n}$$

$\hookrightarrow$  threshold value to guess

Prove

If  $p \ll n^{-2/3}$ , then  $X=0$  with high probability

$$E(X) = \binom{n}{4} p^6 \leq n^4 p^6 \ll 1$$

By Markov's inequality,  $\Pr(X \geq 1) \leq E(X)$   
 $\hookrightarrow$  goes to 0

✶

$$p = o(n^{-2/3})$$

$\hookrightarrow p = \frac{1}{n^{2/3} q(n)}$  such that  $q(n) \rightarrow \infty$  as  $n \rightarrow \infty$

$$p = o(n^{-2/3})$$

$$\hookrightarrow p = \frac{q(n)}{n^{2/3}}$$

$$n^4 p^6 = n^4 \left( \frac{1}{n^{2/3} q(n)} \right)^6$$

$$= n^4 \left( \frac{1}{n^4 q(n)^6} \right) = \frac{1}{q(n)^6} \rightarrow 0$$


✶

$$\text{Var}(X) \leq E(X) + \sum_{i \neq j} \text{cov}(X_i, X_j)$$

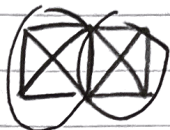
$$\frac{\text{Var}(X)}{E(X)^2} \leq \frac{E(X)}{E(X)^2} + \frac{\sum_{i \neq j} \text{cov}(X_i, X_j)}{E(X)^2} \rightarrow 0$$

$$\leq \frac{1}{E(X)} + \frac{\sum \text{cov}(X_i, X_j)}{E(X)^2} \rightarrow 0$$



$\hookrightarrow$  cov is not 0 when two  share two or more nodes

① share two nodes



$$\begin{aligned} x_i, x_j \quad \text{Cov}(x_i, x_j) &= E(x_i, x_j) - E(x_i) E(x_j) \\ &= p^{11} - p^6 \cdot p^6 \\ &= p^{11} - p^{12} \end{aligned}$$

$$\binom{7}{6} \binom{6}{2}$$

↳ possible



$$\leq n^6$$

② share three nodes



$$\begin{aligned} \text{Cov}(x_i, x_j) &= E(x_i, x_j) - E(x_i) E(x_j) \\ &= p^9 - p^{12} \end{aligned}$$

$$\begin{aligned} \sum_{i,j} \text{Cov}(x_i, x_j) &\leq n^6 (p^{11} - p^{12}) + n^5 (p^9 - p^{12}) \\ &\leq n^6 p^{11} + n^5 p^9 \end{aligned}$$

$$\frac{\sum_{i,j} \text{Cov}(x_i, x_j)}{E(x)^2} \leq \frac{n^6 p^{11} + n^5 p^9}{n^8 p^{12}}$$

- The Covariance of  $x_i$  is variance

$$\frac{\text{Var}(x)}{E(x)^2} \leq \frac{1}{E(x)} + \frac{n^6 p^{11} + n^5 p^9}{n^8 p^{12}} \quad \text{if } p \gg n^{-2/3}, \text{ then } \frac{\text{Var}(x)}{E(x)^2} \rightarrow 0$$

$$p = \frac{q(n)}{n^{2/3}} \text{ such that } q(n) \rightarrow \infty$$

$$\leq \frac{1}{q(n)^6} + \frac{1}{n^2 p} + \frac{1}{n^2 p^3}$$

↳ goes to 0