

1 Experimental Design

Experimental design

Experimental design can be thought of as the study of careful perturbation. Many hypotheses can be tested through observation of natural phenomena, but all non-trivial inference from observation is correlative. Through manipulation of natural phenomena, it is possible to tease causality from correlation

Objectives

The first step in experimental design is to clearly state the hypothesis and in the specific predictions from that hypothesis. It is often tempting to try a complex analysis to answer a simple question.

Example: *Completely randomized design with one factor* . This is a design where you are interested in the effect of a single factor and you can perform all the replicates of all the treatment levels at one time.

- Can test for differences among means of the treatment or factor levels.

Factors

Factors are discrete categorical variables. They can be ordered or unordered. Factors are typically used as the independent variable in Anova

```
> vec <- c(1, 2, 3, 2, 5, 5, 1, 3)
```

```
> vec
```

```
[1] 1 2 3 2 5 5 1 3
```

```
> vec2 <- as.factor(vec)
```

```
> vec2
```

```
[1] 1 2 3 2 5 5 1 3
```

```
Levels: 1 2 3 5
```

```
> is.factor(vec)
```

```
[1] FALSE
```

```
> is.factor(vec2)
```

```
[1] TRUE
```

Ordered factors

```
> vec3 <- ordered(vec)
> vec3
```

```
[1] 1 2 3 2 5 5 1 3
Levels: 1 < 2 < 3 < 5
```

```
> is.factor(vec3)
```

```
[1] TRUE
```

Basic process

- assign individuals to experimental units
- assign units to treatments (manipulations)
- arrange treatments through time and/or space
- arrange data collection through time and/or space

Assigning individuals

As mentioned before, individuals must be assigned to the experimental units at random.

Replication

For each level of a factor, usually multiple *replicates* are performed.

Multiple Replicates:

- Give better estimates of τ_i
- Allow estimation of ϵ in linear models (and comparison to expected distribution)
- Allow statistical tests

Formal expression of randomized design with one factor

$$Y_{ij} = \mu + \tau_i + \epsilon_{ij} \tag{1}$$

This should look familiar. The important parameters for a completely randomized, one factor design are:

- f , the number of factors (here f is defined as 1)
- k , the number of levels for a factor (i ranges from 1 – k)
- n , number of replications per level of the factor (j ranges from 1 – n_i)

The total sample size for the experiment is: $f \times k \times n$ This number is relevant in terms of experimental cost but not necessarily in terms of statistical power. If k approaches n in this design, it is possible to lose power.

Assignment of units to treatments

In a completely randomized design, both the order of treatment level applied and the assignment of experimental units to treatments are randomized.

Why randomize

The effect of randomization is to eliminate correlation that may exist among replicates or experimental units.

- changes in functionality of instrument (DNA flurometer)
- genetic relatedness of experimental subjects

The variation in the total experiment that might be due to that correlation gets subsumed into the ϵ term.

One implication of this is that both random *and* systematic errors tend to increase ϵ and reduce F .

2 Blocking designs

Complete randomized w/ blocking

Imagine a situation where you cannot perform all of your treatments at one time

- limited incubator space
- limited tank space
- limited time

If it is possible to break up the experiment into subgroups that can accomodate every treatment, the subgroups can be used as *blocks* .

Blocking allows the effects of random variation over experimental runs to be explicitly modeled. This takes some of the variation to be explained by the model out of the within-treatment error term.

Complete randomized block design

$$Y_{ijl} = \mu + \text{block}_l + \tau_i + \epsilon_{ijk}$$

Snake data

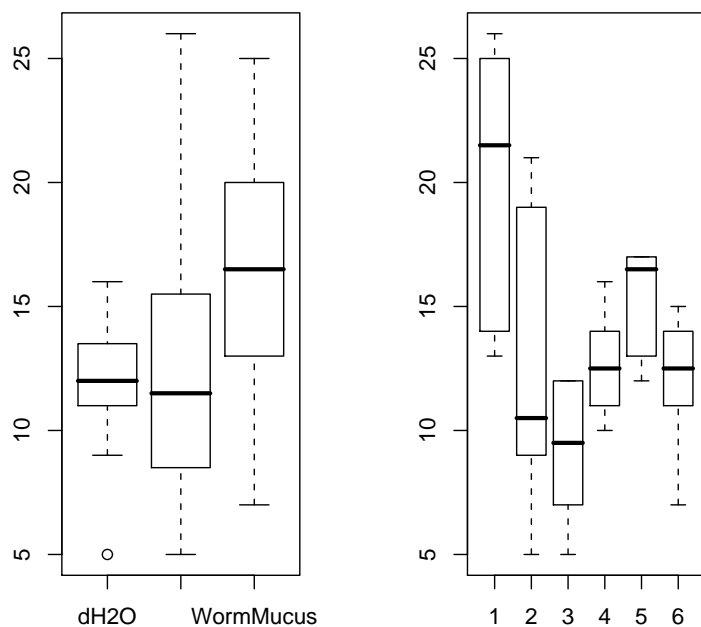
```
> snake <- as.factor(rep(rep(1:6, 3), 2))
> flicks <- c(13, 9, 17, 10, 13, 11, 18, 19, 12, 16, 17, 14, 8,
+ 12, 10, 11, 12, 12)
> flicks <- c(flicks, flicks + round(rnorm(18, 0, 4), 0))
> stimulus <- as.factor(c(rep("FishMucus", 6), rep("WormMucus",
+ 6), rep("dH2O", 6), rep("FishMucus", 6), rep("WormMucus",
+ 6), rep("dH2O", 6)))
> snakedat1 <- data.frame(snake, stimulus, flicks)
> snakedat2 <- snakedat1
> snakedat2[snakedat2$snake == "1", 3] <- snakedat2[snakedat2$snake ==
+ "1", 3] + 6
> snakedat2[snakedat2$snake == "3", 3] <- snakedat2[snakedat2$snake ==
+ "3", 3] - 5
> snakedat2
```

	snake	stimulus	flicks
1	1	FishMucus	19
2	2	FishMucus	9
3	3	FishMucus	12
4	4	FishMucus	10
5	5	FishMucus	13
6	6	FishMucus	11
7	1	WormMucus	24
8	2	WormMucus	19
9	3	WormMucus	7
10	4	WormMucus	16
11	5	WormMucus	17
12	6	WormMucus	14
13	1	dH2O	14
14	2	dH2O	12

15	3	dH20	5
16	4	dH20	11
17	5	dH20	12
18	6	dH20	12
19	1	FishMucus	14
20	2	FishMucus	8
21	3	FishMucus	11
22	4	FishMucus	5
23	5	FishMucus	15
24	6	FishMucus	12
25	1	WormMucus	28
26	2	WormMucus	11
27	3	WormMucus	0
28	4	WormMucus	9
29	5	WormMucus	12
30	6	WormMucus	11
31	1	dH20	18
32	2	dH20	16
33	3	dH20	-3
34	4	dH20	9
35	5	dH20	11
36	6	dH20	12

Snake data-transmogrified

```
> par(mfrow = c(1, 2))
> boxplot(flicks ~ stimulus, data = snakedat2)
> boxplot(flicks ~ snake, data = snakedat2)
```



Snake data analyzed

```
> summary(aov(flicks ~ stimulus, data = snakedat2))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
stimulus	2	68.39	34.19	1.036	0.3661
Residuals	33	1089.17	33.01		

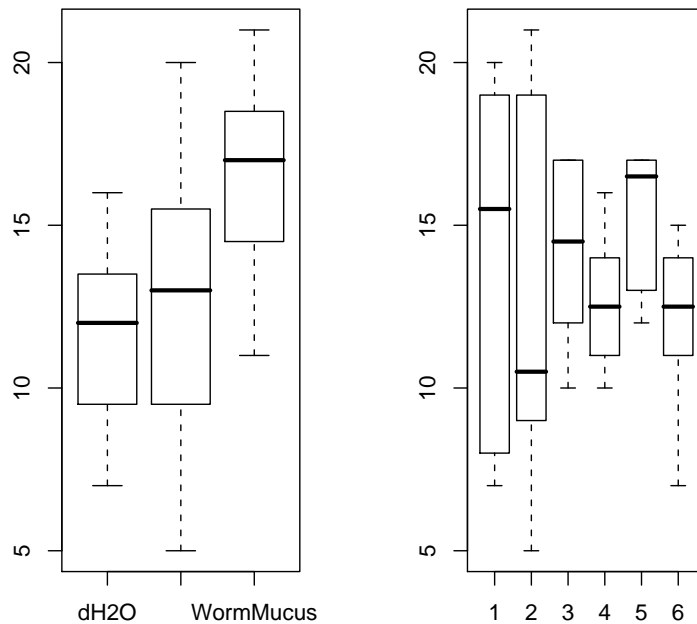
```
> summary(aov(flicks ~ stimulus + snake, data = snakedat2))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
stimulus	2	68.39	34.19	2.1311	0.1376
snake	5	639.89	127.98	7.9759	8.925e-05 ***
Residuals	28	449.28	16.05		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Snake data-actual

```
> par(mfrow = c(1, 2))
> boxplot(flicks ~ stimulus, data = snakedat1)
> boxplot(flicks ~ snake, data = snakedat1)
```



Snake data analyzed, cont

```
> summary(aov(flicks ~ stimulus + snake, data = snakedat1))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
stimulus	2	68.39	34.19	2.1311	0.1376
snake	5	66.22	13.24	0.8254	0.5423
Residuals	28	449.28	16.05		

If no effect of block, then subsequent experiments can ignore block and just randomly assign treatment replicates.

Latin square

There are times in an experimental design where there are clearly defined gradients built into the design. For example, there may be a field that has a gradient of moisture and

a gradient of nitrogen running perpendicular. The Latin Square design is a special type of complete blocking where 2-dimensional space is explicitly taken into account. It can also be used anytime there are two nuisance factors that should affect results.

```
A B C D
C D A B
D C B A
B A D C
```

Rabbit blisters

```
> rabbits <- read.csv("rabbits.csv", header = T)
> rmat <- matrix(rabbits$Order, ncol = 6)
> rownames(rmat) <- rabbits$Position[1:6]
> colnames(rmat) <- unique(rabbits$Rabbit)
> rmat
```

```
  1 2 3 4 5 6
1 3 5 4 1 6 2
2 4 2 6 5 3 1
3 1 3 5 6 2 4
4 6 1 3 2 4 5
5 2 4 1 3 5 6
6 5 6 2 4 1 3
```

Rabbit blisters cont

```
> summary(aov(Observation ~ as.factor(Rabbit) + as.factor(Position) +
+ as.factor(Order), data = rabbits))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(Rabbit)	5	12.8333	2.5667	3.9096	0.01235 *
as.factor(Position)	5	3.8333	0.7667	1.1678	0.35919
as.factor(Order)	5	0.5633	0.1127	0.1716	0.97013
Residuals	20	13.1300	0.6565		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Incomplete randomized blocked design

There are frequently times where there are more treatments than can be placed within a block. In this case, treatments are placed in only some blocks