

# CSC0056 Data Communication

## Week 6, Part I - Poisson Process

Instructor: Chao Wang

Networked Cyber-Physical Systems Laboratory  
Department of Computer Science and Information Engineering  
National Taiwan Normal University

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# Outline

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# References

- ① Harchol-Balter, Mor. Performance modeling and design of computer systems: queueing theory in action. Cambridge University Press, 2013. ISBN 9781107027503. (Chapters 3 and 11)
- ② Bertsekas, Dimitri and Gallager, Robert. Data networks (2nd edition). Prentice Hall, 1992. ISBN 0132009161. (Section 3.3 to 3.3.1, before Markov chain formulation)
- ③ William Feller. An Introduction to Probability Theory and Its Applications, Volume I (3rd edition). Wiley, 1968. ISBN 0471257087.
- ④ 曹亮吉, Poisson 分佈.  
[http://episte.math.ntu.edu.tw/articles/sm/sm\\_16\\_07\\_1/](http://episte.math.ntu.edu.tw/articles/sm/sm_16_07_1/)

# Motivation for the following math: the M/M/1 system

- The Kendall notation: a X/Y/Z queueing system
  - X: the distribution of interarrival times for the arrival process
  - Y: the distribution of the service time
  - Z: the number of servers
- The M/M/1 queueing system:
  - Arrival statistics: data arrives according to the **Poisson process** with rate  $\lambda$ . The interarrival times follow an **exponential distribution** (the first M in M/M/1; M for *memoryless*)
  - Service statistics: the data service times follow an **exponential distribution** with rate  $\mu$  (the second M in M/M/1; M for *memoryless*)
  - With the above two statistical assumptions, the number of data items in the next moment will only depend on the current number of data items in the system. Thus, **Markov chain theory** can be applied here to analyze the probability for a system to have N data items, and subsequently we may get a prediction on the response time by applying Little's Law :)

## Background: Bernoulli trials and binomial distribution [3]

- Definition: Bernoulli trials are repeated independent trials that have only two possible outcomes for each trial and their probabilities remain the same throughout trials.
  - Example: tosses of a coin
- Definition: Let  $b(k; n, p)$  be the probability that  $n$  Bernoulli trials (with probability  $p$  for success and  $q = 1 - p$  for failure) result in  $k$  successes and  $n - k$  failures. Then

$$P\{S_n = k\} = b(k; n, p) = \binom{n}{k} p^k q^{n-k}$$

is called the binominal distribution of  $S_n$ .

## Poisson distribution [3]

- Poisson distribution is an approximation of the binomial distribution. For large  $n$  and small  $p$ , we have

$$b(k; n, p) \approx \frac{\lambda'^k}{k!} e^{-\lambda'}$$

where we define parameter  $\lambda' = np$ . Now, let

$$p(k; \lambda') = \frac{\lambda'^k}{k!} e^{-\lambda'}$$

and we call  $p(k; \lambda')$  the Poisson distribution.

# Proof of $b(k; n, p) \approx \frac{\lambda'^k}{k!} e^{-\lambda'}$

From the definition of  $b(k; n, p) = \binom{n}{k} p^k q^{n-k}$ , for small  $p$  we have

$$\frac{b(k; n, p)}{b(k-1; n, p)} = \frac{\lambda' - (k-1)p}{kq} = \frac{\lambda' - (k-1)p}{k(1-p)} \approx \frac{\lambda'}{k}$$

which implies

$$b(k; n, p) \approx \frac{\lambda'^k}{k!} b(0; n, p).$$

Now, for large  $n$  we have  $b(0; n, p) \approx e^{-\lambda'}$  using Taylor expansion (i.e.,  $\ln(1+t) = t - \frac{1}{2}t^2 + \frac{1}{3}t^3 - \dots$ ). And therefore, we write

$$b(k; n, p) \approx \frac{\lambda'^k}{k!} e^{-\lambda'}.$$

# Exponential distribution

- A random variable  $X$  is said to be distributed exponentially with rate  $\lambda$  if  $X$  has the following probability density function:

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0; \\ 0 & \text{otherwise.} \end{cases}$$

- The cumulative distribution function is given by

$$F_x(s) = P\{X \leq s\} = \int_0^s f(x) dx = 1 - e^{-\lambda s}.$$

- The mean is given by  $E[X] = \int_{-\infty}^{\infty} x f(x) dx = \frac{1}{\lambda}$ .
- Exponential distribution is *memoryless* in the sense that, for  $s, t > 0$ ,

$$P\{X > s+t | X > s\} = \frac{P\{X > s+t\}}{P\{X > s\}} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} = e^{-\lambda t} = P\{X > t\}.$$



# Poisson process, definition [2]

- A *Poisson process with rate  $\lambda$*  is a stochastic process  $\{A(t)|t \geq 0\}$  where
  - 1  $A(t)$  is a counting process that represents the total # of arrivals in  $[0, t]$ , and  $A(0) = 0$ ;
  - 2 Number of arrivals in disjoint time intervals are independent;
  - 3 Number of arrivals in any intervals of length  $\tau$  is Poisson distributed with parameter  $\lambda\tau$ . That is, for all  $t, \tau > 0$ ,

$$P\{A(t + \tau) - A(t) = k\} = e^{-\lambda\tau} \frac{(\lambda\tau)^k}{k!}$$

where  $k$  is the number of arrivals in the interval.

(Reminder: a stochastic process is a sequence of random variables)

# Poisson process, properties [1][2]

- Property 1: the interarrival times of a Poisson process are independent and exponentially distributed with rate  $\lambda$ .
- Property 2: if two or more independent Poisson processes  $A_1, \dots, A_n$  are merged into a single process  $A = A_1 + A_2 + \dots + A_n$ , then  $A$  is a Poisson process with a rate equal to the sum of the rates of  $A_i$  for  $i = 1 \dots n$ .

## Proof of Property 1 [1]

Let  $T_1, T_2, \dots, T_n$  be the interarrival times of a sequence of events. By the definition of Poisson process, we have

$$P\{T_1 > \tau\} = P\{A(\tau) = 0\} = \frac{e^{-\lambda\tau}(\lambda\tau)^0}{0!} = e^{-\lambda\tau}.$$

Now, consider

$$\begin{aligned} P\{T_{n+1} > \tau \mid \sum_{i=1}^n T_i = s\} &= P\{\text{zero arrival in } (s, s + \tau] \mid \sum_{i=1}^n T_i = s\} \\ &= P\{\text{zero arrival in } (s, s + \tau]\} \\ &\quad (\text{by independent increments}) \\ &= e^{-\lambda\tau} \quad (\text{by stationary increments}), \end{aligned}$$

and this shows that

$$P\{T_{n+1} > \tau \mid \sum_{i=1}^n T_i = s\} = P\{T_1 > \tau\}.$$