

G-computation and propensity score weighting to estimate a marginal effect

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Plan

Introduction

Inverse probability weighting (IPW)

G-computation (GC)

Assumptions of causal inference

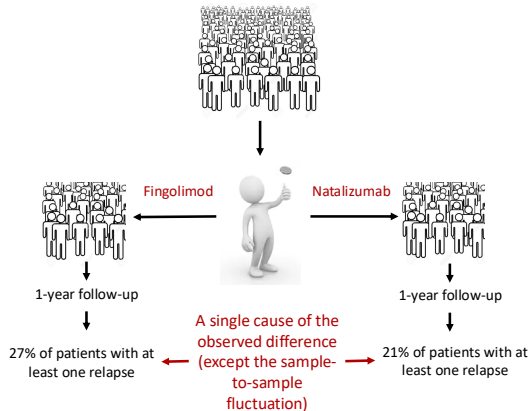
References



A use case in multiple sclerosis (MS)

- ▶ A MS relapse is the occurrence of new MS symptoms or the worsening of old ones.
- ▶ Fingolimod and natalizumab share the same indication for second-line treatment.
- ▶ Which treatment is most effective in preventing MS relapse?

Randomization allows us to demonstrate a causal relationship



Confounders in observational studies: data from Barbin et al.¹

<div>Comparative efficacy of fingolimod vs natalizumab</div> <div>A French multicenter observational study</div> <div>OPEN ▲</div>			
<div>Laetitia Barbin, PhD* Chloe Rousseau, MSc* Natasha Jousset, BSc Romain Casey, PhD Marc Debouverie, MD, PhD</div> <div>ABSTRACT</div> <div>Objective: To compare natalizumab and fingolimod on both clinical and MRI outcomes in patients with relapsing-remitting multiple sclerosis (RRMS) from 27 multiple sclerosis centers participating in the French follow-up cohort Observatoire de Multiple Sclerose.</div>			
Baseline characteristics	All patients	Fingolimod	Natalizumab
No. of patients considered for analysis	629	303	326
Female, n (%)	479 (76.2)	225 (74.3)	254 (77.9)
Age at treatment initiation, y, mean (SD)	37.0 (9.6)	37.2 (9.2)	36.8 (9.9)
Disease duration at treatment initiation, y, mean (SD)	8.5 (6.4)	9.0 (6.8)	8.0 (6.1)
EDSS score, mean (SD)	2.6 (1.3)	2.4 (1.3)	2.8 (1.3)
EDSS score between 3.0 and 5.5, n (%)	288 (45.8)	122 (40.3)	166 (50.9)
Previous immunomodulatory treatment, ^a n (%)	556 (88.4)	263 (86.8)	293 (89.2)
Relapse in the preceding year, n (%)			
≥1 relapse	526 (83.6)	233 (76.9)	293 (89.9)
≥2 relapses	264 (42.0)	92 (30.4)	172 (52.8)

¹Barbin et al. Neurology (2016)

Confounders in observational studies

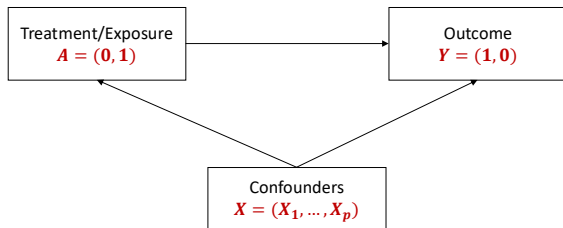
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≥2 relapses	264 (42.0)	92 (30.4)	172 (52.8)

Endpoints	All patients	Fingolimod	Natalizumab
Clinical			
No. of patients considered for relapse analysis	629	303	326
At least one relapse, n (%)	151 (24.0)	82 (27.1)	69 (21.2)

The greater efficacy of the Natalizumab may be underestimated

Context of this session

- Binary exposure and binary outcome.



- From the observations (y_i, a_i, x_i) , with $i = 1, \dots, N.$, we aim to estimate the average treatment effect (no mediation):

$$\text{ATE} = \mathbb{E}_X[\Pr(Y = 1 \mid A = 1)] - \mathbb{E}_X[\Pr(Y = 1 \mid A = 0)]$$

The ATE in our use case

Baseline characteristics	All patients	Fingolimod	Natalizumab
No. of patients considered for analysis	629	303	326
Female, n (%)	479 (76.2)	479 (76.2)	479 (76.2)
Age at treatment initiation, y, mean (SD)	37.0 (9.6)	37.0 (9.6)	37.0 (9.6)
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Relapse in the preceding year, n (%)			
≥1 relapse	526 (83.6)	526 (83.6)	526 (83.6)
≥2 relapses	264 (42.0)	264 (42.0)	264 (42.0)

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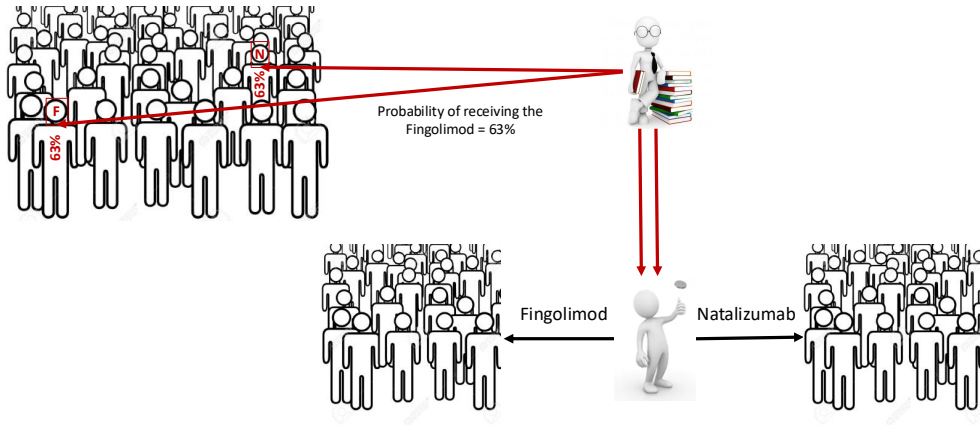
The definition of a propensity score (PS)

- ▶ The PS is the individual probability to be exposed:

$$\pi_i = \Pr(A = 1 \mid X_i)$$

- ▶ The related **exposure model** can be obtained from a logistic regression.
- ▶ Once estimated, two main methods are used:
 - ▶ Matching on the PS.
 - ▶ Weighting on the Inverse of the Treatment Probability (IPW).

The principle of matching on the PS



Three main limitations of matching on the PS^2

- ▶ The targeted population cannot be defined a priori.
- ▶ The sample size is reduced due to non-matched individuals.
- ▶ The matched samples vary due to the random matching process.



²Chatton et al. Scientific report (2020)

The principle of weighting on the PS

Fingolimod ($A = 1$)				Natalizumab ($A = 0$)			
relapse +1y (Y)	relapse -1y (X)	relapse +1y (Y)	relapse -1y (X)	relapse +1y (Y)	relapse -1y (X)	relapse +1y (Y)	relapse -1y (X)
0	1			0	0		
0	1			1	0		
1	1			0	1		
0	0			0	0		
0	1			1	1		
0	0			0	0		
1	1			0	0		
0	1			0	0		
1	0			0	0		
0	1			0	0		
30% 70%				20% 20%			
Real world		Conterfactual world		Real world		Conterfactual world	

The principle of weighting on the PS

The PS represents the individual probability of receiving Fingolimod.

$$P(A = 1|X)$$

Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1			0	0		
0	1			1	0		
1	1			0	1		
0	0			0	0		
0	1			1	1		
0	0			0	0		
1	1			0	0		
0	1			0	0		
1	0			0	0		
0	1			0	0		

The principle of weighting on the PS

The PS represents the individual probability of receiving Fingolimod:

$$P(A = 1|X)$$

For patients with a relapse before the prescription, 7/9 received Fingolimod:

$$PS = P(A = 1|X = 1) = 78\%$$

Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1			0	0		
0	1			1	0		
1	1			0	1		
0	0			0	0		
0	1			1	1		
0	0			0	0		
1	1			0	0		
0	1			0	0		
1	0			0	0		
0	1			0	0		
7				2			

The principle of weighting on the PS

The PS represents the individual probability of receiving Fingolimod:

$$P(A = 1|X)$$

For patients with a relapse before the prescription, 7/9 received Fingolimod:

$$PS = P(A = 1|X = 1) = 78\%$$

For patients without relapse before the prescription, 3/11 received Fingolimod:

$$PS = P(A = 1|X = 0) = 27\%$$

Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1			0	0		
0	1			1	0		
1	1			0	1		
0	0			0	0		
0	1			1	1		
0	0			0	0		
1	1			0	0		
0	1			0	0		
1	0			0	0		
0	1			0	0		
3				8			

The principle of weighting on the PS

For patients with a relapse before the prescription, 7/9 received Fingolimod:

$$PS = P(A = 1 | X = 1) = 78\%$$

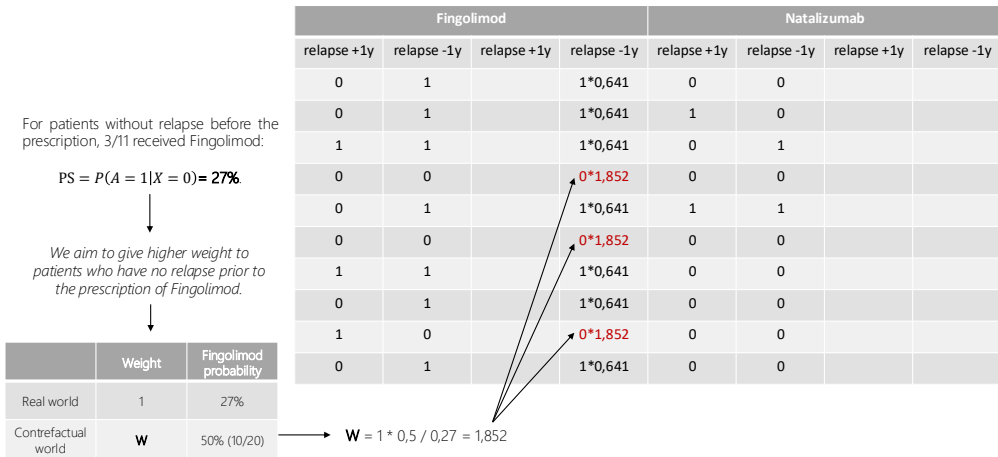
We aim to give less weight to patients who have had a relapse prior to the prescription of Fingolimod.

	Weight	Fingolimod probability
Real world	1	78%
Counterfactual world	W	50% (10/20)

Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1		$1 \cdot 0,641$	0	0		
0	1		$1 \cdot 0,641$	1	0		
1	1		$1 \cdot 0,641$	0	1		
0	0			0	0		
0	1		$1 \cdot 0,641$	1	1		
0	0			0	0		
1	1		$1 \cdot 0,641$	0	0		
0	1		$1 \cdot 0,641$	0	0		
1	0			0	0		
0	1		$1 \cdot 0,641$	0	0		

$$W = 1 \cdot 0,5 / 0,78 = 0,641$$

The principle of weighting on the PS



The principle of weighting on the PS

For patients with a relapse before the prescription, 2/9 received Natalizumab:

$$PS = P(A = 0|X = 1) = 22\%.$$

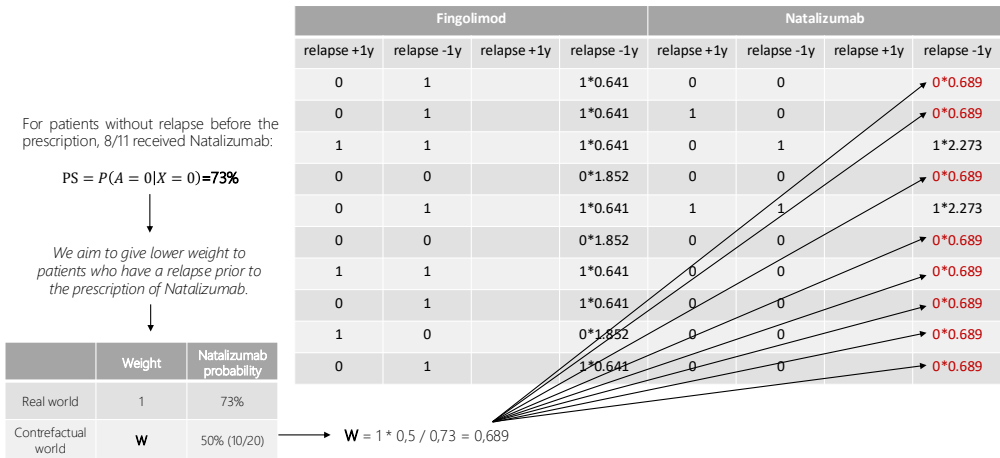
We aim to give higher weight to patients who have a relapse prior to the prescription of Natalizumab.

	Weight	Natalizumab probability
Real world	1	22%
Counterfactual world	W	50% (10/20)

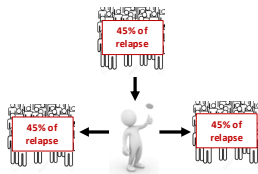
Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1		1*0,641	0	0		
0	1		1*0,641	1	0		
1	1		1*0,641	0	1		1*2,273
0	0		0*1,852	0	0		
0	1		1*0,641	1	1		1*2,273
0	0		0*1,852	0	0		
1	1		1*0,641	0	0		
0	1		1*0,641	0	0		
1	0		0*1,852	0	0		
0	1		1*0,641	0	0		

$$W = 1 * 0,5 / 0,22 = 2,273$$

The principle of weighting on the PS



The principle of weighting on the PS



Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1		1*0,641	0	0		0*0,689
0	1		1*0,641	1	0		0*0,689
1	1		1*0,641	0	1		1*2,273
0	0		0*1,852	0	0		0*0,689
0	1		1*0,641	1	1		1*2,273
0	0		0*1,852	0	0		0*0,689
1	1		1*0,641	0	0		0*0,689
0	1		1*0,641	0	0		0*0,689
1	0		0*1,852	0	0		0*0,689
0	1		1*0,641	0	0		0*0,689

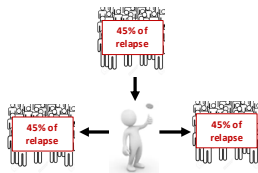
70%

44,7%

20%

45,3%

The principle of weighting on the PS



Fingolimod				Natalizumab			
relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y	relapse +1y	relapse -1y
0	1	0*0,641	1*0,641	0	0	0*0,689	0*0,689
0	1	0*0,641	1*0,641	1	0	1*0,689	0*0,689
1	1	1*0,641	1*0,641	0	1	0*2,273	1*2,273
0	0	0*1,852	0*1,852	0	0	0*0,689	0*0,689
0	1	0*0,641	1*0,641	1	1	1*2,273	1*2,273
0	0	0*1,852	0*1,852	0	0	0*0,689	0*0,689
1	1	1*0,641	1*0,641	0	0	0*0,689	0*0,689
0	1	0*0,641	1*0,641	0	0	0*0,689	0*0,689
1	0	1*1,852	0*1,852	0	0	0*0,689	0*0,689
0	1	0*0,641	1*0,641	0	0	0*0,689	0*0,689
30%	70%	31,2%	44,7%	20%	20%	29,5%	45,3%

$$ATE = 31.2 - 29.5 = 1.7\%$$

Main steps to estimate ATE by using IPW (1)

- ▶ Construct the exposure model, for instance a logistic regression:

$$\text{logit}(\Pr(A = 1 \mid X)) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

- ▶ Predict the probability of being exposed according to the individual characteristics:

$$\hat{\pi}_i = \frac{\exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi})}{(1 + \exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi}))}$$

- ▶ Transform the predictions into individual (stabilized) weights:³

$$\hat{\omega}_i = \mathbb{1}\{A_i = 1\} \frac{\Pr(A_i = 1)}{\hat{\pi}_i} + \mathbb{1}\{A_i = 0\} \frac{1 - \Pr(A_i = 1)}{1 - \hat{\pi}_i}$$

³Austin and Stuart. Statistics in medicine (2015)

Main steps to estimate ATE by using IPW (2)

- ▶ Estimate the ATE in the weighted sample, for instance an univariate logistic regression obtained by maximising the weighted log-likelihood:

$$\text{logit}(\Pr(Y = 1 \mid A)) = \gamma_0 + \gamma_1 A$$

- ▶ Compute the ATE:

$$\text{ATE} = \mathbb{E}_X[\Pr(Y = 1 \mid A = 1)] - \mathbb{E}_X[\Pr(Y = 1 \mid A = 0)]$$

$$\widehat{\text{ATE}} = \frac{\exp(\hat{\gamma}_0 + \hat{\gamma}_1)}{(1 + \exp(\hat{\gamma}_0 + \hat{\gamma}_1))} - \frac{\exp(\hat{\gamma}_0)}{(1 + \exp(\hat{\gamma}_0))}$$

- ▶ Bootstrap all the previous steps to obtain the 95% confidence intervals.

ATE by using IPW and R software

```
1 # The exposure model
2 glm1 <- glm(A~X1+... +Xp, family=binomial(link="logit"))
3
4 # Individual predicted probabilities
5 pi <- glm1$fitted.values
6
7 # Weights
8 pA <- mean(A)
9 w <- (A==1)*pA/pi + (A==0)*(1-pA)/(1-pi)
10
11 # ATE estimation
12 glm2 <- glm(Y~A, weights = w, family=binomial(link = "logit"))
13 p0 <- exp(glm2[1,1])/(1+exp(glm2[1,1]))
14 p1 <- exp(glm2[1,1]+ glm2[2,1])/(1+exp(glm2[1,1]+glm2[2,1]))
15 ATE <- p1 - p0
```

The average treatment effect on the treated (ATT) as an alternative

Baseline characteristics	All patients	Fingolimod	Natalizumab
No. of patients considered for analysis	629	303	326
Female, n (%)	479 (76.2)	225 (74.3)	254 (77.9)
Age at treatment initiation, y, mean (SD)	37.0 (9.6)	37.2 (9.2)	36.8 (9.9)
Disease duration at treatment initiation, y, mean (SD)	8.5 (6.4)	9.0 (6.8)	8.0 (6.1)
EDSS score, mean (SD)	2.6 (1.3)	2.4 (1.3)	2.8 (1.3)
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Relapse in the preceding year, n (%)			
≥1 relapse	526 (83.6)	233 (76.9)	293 (89.9)
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ATE

Baseline characteristics	All patients	Fingolimod	Natalizumab
No. of patients considered for analysis	629	303	326
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EDSS score between 3.0 and 5.5, n (%)	288 (45.8)	288 (45.8)	288 (45.8)
Previous immunomodulatory treatment,* n (%)	556 (88.4)	556 (88.4)	556 (88.4)
Relapse in the preceding year, n (%)			
≥1 relapse	526 (83.6)	526 (83.6)	526 (83.6)
≥2 relapses	264 (42.0)	264 (42.0)	264 (42.0)

ATT

Baseline characteristics	All patients	Fingolimod	Natalizumab
No. of patients considered for analysis	629	303	326
Female, n (%)	225 (74.3)	225 (74.3)	225 (74.3)
Age at treatment initiation, y, mean (SD)	37.2 (9.2)	37.2 (9.2)	37.2 (9.2)
Disease duration at treatment initiation, y, mean (SD)	9.0 (6.8)	9.0 (6.8)	9.0 (6.8)
EDSS score, mean (SD)	2.4 (1.3)	2.4 (1.3)	2.4 (1.3)
EDSS score between 3.0 and 5.5, n (%)	122 (40.3)	122 (40.3)	122 (40.3)
Previous immunomodulatory treatment,* n (%)	263 (86.8)	263 (86.8)	263 (86.8)
Relapse in the preceding year, n (%)			
≥1 relapse	233 (76.9)	233 (76.9)	233 (76.9)
≥2 relapses	92 (30.4)	92 (30.4)	92 (30.4)

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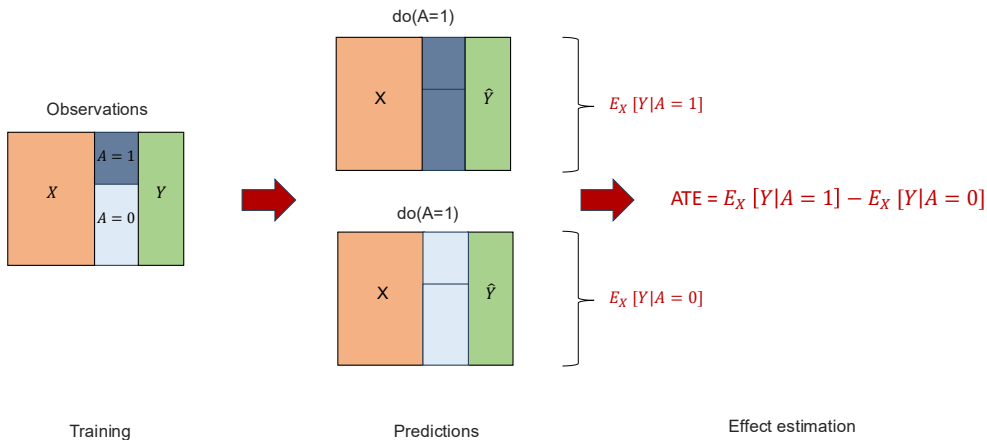
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The principle of the G-computation (GC) for ATE estimation



Main steps to estimate ATE by using GC (1)

- ▶ Construct the outcome model, for instance a logistic regression:

$$\text{logit}(Pr(Y = 1 | X, A)) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \beta_{p+1} A$$

- ▶ Predict the individual probabilities of event if all the subject are exposed:

$$\hat{\pi}_{1i} = \frac{\exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi} + \hat{\beta}_{p+1})}{(1 + \exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi} + \hat{\beta}_{p+1}))}$$

- ▶ Predict the individual probabilities of event if all the subject are unexposed:

$$\hat{\pi}_{0i} = \frac{\exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi})}{(1 + \exp(\hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_p X_{pi}))}$$

Main steps to estimate ATE by using GC (2)

- Compute the ATE:

$$\widehat{ATE} = n^{-1} \sum_i \hat{\pi}_{1i} - n^{-1} \sum_i \hat{\pi}_{0i}$$

- Bootstrap all the previous steps to obtain the 95% confidence intervals.

ATE by using GC and R software

```
1 # The outcome model
2 glm1 <- glm(Y~X1+...+Xp+A, family=binomial(link="logit"), data=mydata)
3
4 # The conterfactual prediction
5 mydata0 <- mydata1 <- mydata
6 mydata0$A <- 0
7 mydata1$A <- 1
8 p0i <- predict(glm1, newdata=mydata0, type="response")
9 p1i <- predict(glm1, newdata=mydata1, type="response")
10
11 # The ATE estimation
12 p0 <- mean(p0i)
13 p1 <- mean(p1i)
14 ATE <- p1 - p0
```

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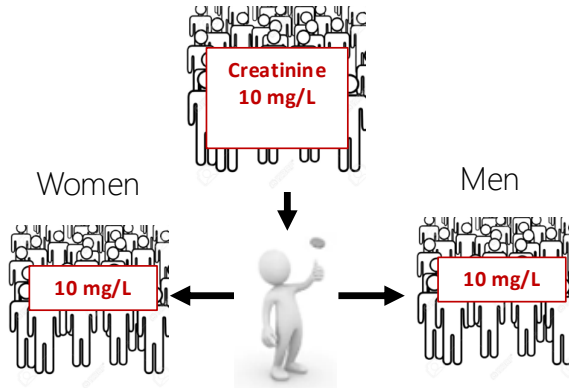
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The exposure should be modifiable



The correct specification of the working models

- ▶ Do not categorize continuous explanatory variables X .
- ▶ Respect the assumptions of the model (log-linearity in logistic regression).
- ▶ Consider the interaction when necessary.
- ▶ AUC maximization does not represent an objective.
- ▶ Standardized mean differences should not exceed 10%.⁴
 - ▶ For categorical covariate, SMD is defined by:

$$SMD = (\hat{p}_{T=1} - \hat{p}_{A=0}) / \left(\sqrt{(\hat{p}_{A=1}(1 - \hat{p}_{A=1}) + \hat{p}_{A=0}(1 - \hat{p}_{A=0}))/2} \right)$$

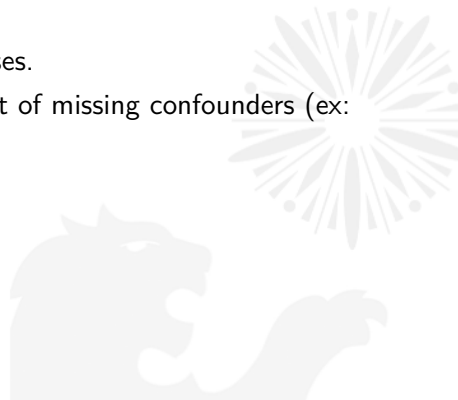
- ▶ For continuous covariates, SMD is defined by:

$$SMD = (\bar{X}_{A=1} - \bar{X}_{A=0}) / \left(\sqrt{(s_{A=1}(X)^2 + s_{A=0}(X)^2)/2} \right)$$

⁴Austin PC. Communications in Statistics - Simulation and Computation (2009)

No missing confounders

- ▶ Draw a directed acyclic graph (DAG) before analyses.
- ▶ Sensitivity analysis to evaluate the potential impact of missing confounders (ex: E-value).



Positivity assumption

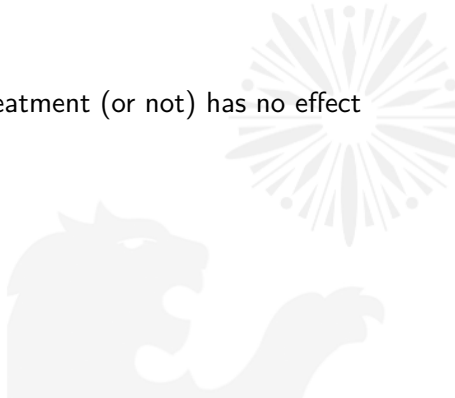
- ▶ Draw the histograms of the propensity scores in the two groups.
- ▶ Distinguish theoretical violation and near violation (i.e. by chance). ⁵
 - ▶ Inclusion criteria should be updated in case of theoretical violation.
 - ▶ GC is more robust to near-violation.



⁵Léger et al. Biometrical Journal (2022)

Consistency and no interference

- ▶ Consistency: well-defined expositions/treatment.
- ▶ No interference: whether one individual receives treatment (or not) has no effect on the potential outcomes of any other individual.



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