

统计信号处理

第九章

线性贝叶斯估计

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一、引言

- MMSE估计: $\hat{\theta} = E(\theta | x) = \begin{bmatrix} \int \theta_1 p(\theta | x) d\theta \\ \int \theta_2 p(\theta | x) d\theta \\ \vdots \\ \int \theta_p p(\theta | x) d\theta \end{bmatrix}$

多重积分

- MAP估计: $\hat{\theta} = \arg \max_{\theta} \{p(\theta | x)\}$

多维最大值

1. 不一定好求解!

2. 无pdf?

- 线性MMSE (LMMSE)估计

二、LMMSE估计

- 观察数据 $\{x[0], x[1], x[2], \dots, x[N-1]\}$
- LMMSE意味着：

$$\hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N$$

即限定估计量与观察数据间呈线性关系，然后使

$$\text{Bmse}(\hat{\theta}) = E\left[(\theta - \hat{\theta})^2\right] \quad \text{最小化}$$

即，LMMSE:

$$\begin{cases} \min \left\{ E\left[(\theta - \hat{\theta})^2\right] \right\} \\ s.t. \quad \hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N \end{cases}$$

LMMSE:

$$\begin{cases} \min \left\{ E \left[\left(\theta - \hat{\theta} \right)^2 \right] \right\} \\ s.t. \quad \hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N \end{cases}$$

$$\left. \begin{array}{l} \text{Bmse}(\hat{\theta}) = E \left[\left(\theta - \hat{\theta} \right)^2 \right] \\ \hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N \end{array} \right\} \Rightarrow \left. \begin{array}{l} \text{Bmse}(\hat{\theta}) = E \left[\left(\theta - \sum_{n=0}^{N-1} a_n x[n] - a_N \right)^2 \right] \\ \frac{\partial \text{Bmse}(\hat{\theta})}{\partial a_N} = 0 \end{array} \right\} \Rightarrow$$

$$E \left[\left(\theta - \sum_{n=0}^{N-1} a_n x[n] - a_N \right) \right] = 0 \Rightarrow a_N = E(\theta) - \sum_{n=0}^{N-1} a_n E(x[n])$$

$$\left. \begin{aligned} a_N &= E(\theta) - \sum_{n=0}^{N-1} a_n E(x[n]) \\ \hat{\theta} &= \sum_{n=0}^{N-1} a_n x[n] + a_N \end{aligned} \right\} \Rightarrow \left. \begin{aligned} \hat{\theta} &= E(\theta) + \sum_{n=0}^{N-1} a_n (x[n] - E(x[n])) \\ \text{Bmse}(\hat{\theta}) &= E[(\theta - \hat{\theta})^2] \end{aligned} \right\} \Rightarrow$$

$$\begin{aligned} \text{Bmse}(\hat{\theta}) &= E \left[\left(\theta - E(\theta) - \sum_{n=0}^{N-1} a_n (x[n] - E(x[n])) \right)^2 \right] \\ &= E \left[\left(\underbrace{\sum_{n=0}^{N-1} a_n (x[n] - E(x[n]))}_{\mathbf{a}} - (\theta - E(\theta)) \right)^2 \right] \\ &\quad \mathbf{a} = [a_0, a_1, a_2, \dots, a_{N-1}]^T
 \end{aligned}$$

$$\begin{aligned} &= E \left[\left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) - (\theta - E(\theta)) \right)^2 \right] \\ &= E \left[\left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) - (\theta - E(\theta)) \right) \left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) - (\theta - E(\theta)) \right)^T \right]
 \end{aligned}$$

$$\begin{aligned}
 \text{Bmse}(\hat{\theta}) &= E \left[\left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) - (\theta - E(\theta)) \right) \left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) - (\theta - E(\theta)) \right)^T \right] \\
 &= E \left[\begin{aligned} &\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) (\mathbf{x} - E(\mathbf{x}))^T \mathbf{a} - \mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) (\theta - E(\theta)) \\ &- (\theta - E(\theta)) (\mathbf{x} - E(\mathbf{x}))^T \mathbf{a} + (\theta - E(\theta))^2 \end{aligned} \right] \quad \boxed{\mathbf{C}_{\theta x}^T = \mathbf{C}_{x\theta}} \\
 &= \mathbf{a}^T \mathbf{C}_{xx} \mathbf{a} - \mathbf{a}^T \mathbf{C}_{x\theta} - \mathbf{C}_{\theta x} \mathbf{a} + \mathbf{C}_{\theta\theta} \\
 \left. \begin{aligned} &\frac{\partial \text{Bmse}(\hat{\theta})}{\partial \mathbf{a}} = 0 \end{aligned} \right\} \Rightarrow 2\mathbf{C}_{xx} \mathbf{a} - 2\mathbf{C}_{x\theta} = 0 \Rightarrow \mathbf{a} = \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \\
 \left. \begin{aligned} &\hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N = \mathbf{a}^T \mathbf{x} + a_N \\ &a_N = E(\theta) - \sum_{n=0}^{N-1} a_n E(x[n]) = E(\theta) - \mathbf{a}^T E(\mathbf{x}) \end{aligned} \right\} \Rightarrow \hat{\theta} = E(\theta) + \mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) \\
 \left. \begin{aligned} &\hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ &\text{Bmse}(\hat{\theta}) = E \left[(\theta - \hat{\theta})^2 \right] \end{aligned} \right\} \Rightarrow \underline{\text{Bmse}(\hat{\theta}) = \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}}
 \end{aligned}$$

✓ 仅需一阶矩、二阶矩

✓ 无需PDF

● LMMSE Vs MMSE

LMMSE

LMMSE优化目标:

$$\begin{cases} \min E\left[(\theta - \hat{\theta})^2\right] \\ s.t. \quad \hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] + a_N \end{cases}$$

相应的LMMSE估计量及Bmse:

$$\begin{cases} \hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ \text{Bmse}(\hat{\theta}) = \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \end{cases}$$

- ✓ 附加了线性约束
- ✓ 仅需一阶矩和二阶矩
- ✓ 可得显示解——好求
- ✓ 仅在“线性”中最优

MMSE

MMSE优化目标: $\min \left\{ E\left((\theta - \hat{\theta})^2\right) \right\}$

相应的MMSE估计量及Bmse:

$$\begin{cases} \hat{\theta} = E(\theta | \mathbf{x}) \\ \text{Bmse}(\hat{\theta}) = E\left((\theta - \hat{\theta})^2\right) \end{cases}$$

- ✓ 无附加约束
- ✓ 需PDF
- ✓ 可能难以求得显示解
- ✓ 全局最优

若观测数据 \mathbf{x} 满足如下数学模型：

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

其中， \mathbf{H} 为观测矩阵， $\boldsymbol{\theta}$ 为待估计参数，具有先验概率密度函数 $N(\boldsymbol{\mu}_\theta, \mathbf{C}_\theta)$ ，噪声矢量 \mathbf{w} 服从 $N(\mathbf{0}, \mathbf{C}_w)$ ，且与 $\boldsymbol{\theta}$ 无关，则MMSE为：

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta} | \mathbf{x}) = E(\boldsymbol{\theta}) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$\mathbf{C}_{\theta|x} = \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

Vs

LMMSE

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$\text{Bmse}(\hat{\boldsymbol{\theta}}) = \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

对贝叶斯一般线性模型，LMMSE与MMSE具有相同形式

例：白噪声中电平估计问题：

$$x[n] = A + w[n], \quad n = 0, 1, \dots, N-1$$

待估计参数为信号幅度 A ，且 $A \sim U[-A_0, A_0]$ ， $w[n]$ 为高斯白噪声 $w[n] \sim N(0, \sigma^2)$ ，且和 A 相互独立。试比较其MMSE与LMMSE。

1. MMSE估计量

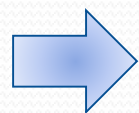
$$\hat{A} = \int A \underline{p(A | \mathbf{x})} dA$$

$$p(A | \mathbf{x}) = \frac{p(\mathbf{x} | A) p(A)}{\int p(\mathbf{x} | A) p(A) dA}$$

$$\begin{aligned} p_x(x[n] | A) &= p_w(x[n] - A | A) \\ &= p_w(x[n] - A) \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2} (x[n] - A)^2\right\} \end{aligned}$$

$$\begin{aligned}
 & \Rightarrow p(\mathbf{x} | A) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A)^2 \right\} \\
 & p(A) = \frac{1}{2A_0}, \quad |A| \leq A_0 \\
 & p(A | \mathbf{x}) = \frac{p(\mathbf{x} | A) p(A)}{\int p(\mathbf{x} | A) p(A) dA}
 \end{aligned}$$

$$\begin{aligned}
 p(A | \mathbf{x}) &= \begin{cases} \frac{\frac{1}{2A_0 (2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A)^2 \right\}}{\int_{-A_0}^{A_0} \frac{1}{2A_0 (2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A)^2 \right\} dA}, & |A| \leq A_0 \\ 0, & |A| > A_0 \end{cases} \\
 &= \begin{cases} \frac{\exp \left\{ -\frac{1}{2\sigma^2} N (A - \bar{x})^2 \right\}}{\int_{-A_0}^{A_0} \exp \left\{ -\frac{1}{2\sigma^2} N (A - \bar{x})^2 \right\} dA}, & |A| \leq A_0 \\ 0, & |A| > A_0 \end{cases}
 \end{aligned}$$



$$p(A|\mathbf{x}) = \begin{cases} \frac{\exp\left\{-\frac{1}{2\frac{\sigma^2}{N}}(A-\bar{x})^2\right\}}{\int_{-A_0}^{A_0} \exp\left\{-\frac{1}{2\frac{\sigma^2}{N}}(A-\bar{x})^2\right\} dA}, & |A| \leq A_0 \\ 0, & |A| > A_0 \end{cases} = \begin{cases} \frac{\frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp\left\{-\frac{1}{2\frac{\sigma^2}{N}}(A-\bar{x})^2\right\}}{\int_{-A_0}^{A_0} \frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp\left\{-\frac{1}{2\frac{\sigma^2}{N}}(A-\bar{x})^2\right\} dA}, & |A| \leq A_0 \\ 0, & |A| > A_0 \end{cases}$$

$$\text{MMSE: } \hat{A} = E(A|\mathbf{x}) = \int A p(A|\mathbf{x}) dA$$

$$= \frac{\int_{-A_0}^{A_0} A \frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp\left\{-\frac{1}{2\frac{\sigma^2}{N}} \sum_{n=0}^{N-1} (A-\bar{x})^2\right\} dA}{\int_{-A_0}^{A_0} \frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp\left\{-\frac{1}{2\frac{\sigma^2}{N}} \sum_{n=0}^{N-1} (A-\bar{x})^2\right\} dA}$$

——化简困难!


2. LMMSE估计量

$$\hat{A} = E(A) + \mathbf{C}_{Ax} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

其中, $E(A) = 0$

$$\mathbf{C}_{Ax} = E\left((A - E(A))(\mathbf{x} - E(\mathbf{x}))^T\right) = E\left(A(A\mathbf{1} + \mathbf{w})^T\right) = E(A^2)\mathbf{1}^T = \sigma_A^2 \mathbf{1}^T$$

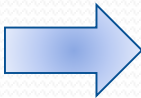
$$\mathbf{C}_{xx} = E\left((\mathbf{x} - E(\mathbf{x}))(\mathbf{x} - E(\mathbf{x}))^T\right) = E\left((A\mathbf{1} + \mathbf{w})(A\mathbf{1} + \mathbf{w})^T\right) = \sigma_A^2 \mathbf{1}\mathbf{1}^T + \sigma^2 \mathbf{I}$$


$$\hat{A} = \sigma_A^2 \mathbf{1}^T (\sigma_A^2 \mathbf{1}\mathbf{1}^T + \sigma^2 \mathbf{I})^{-1} \mathbf{x}$$

$$= \frac{\sigma_A^2}{\sigma^2} \mathbf{1}^T \left(\frac{\sigma_A^2}{\sigma^2} \mathbf{1}\mathbf{1}^T + \mathbf{I} \right)^{-1} \mathbf{x}$$

Woodbury恒等式:

$$(\mathbf{B} + \mathbf{u}\mathbf{u}^T)^{-1} = \mathbf{B}^{-1} - \frac{\mathbf{B}^{-1}\mathbf{u}\mathbf{u}^T\mathbf{B}^{-1}}{1 + \mathbf{u}^T\mathbf{B}^{-1}\mathbf{u}}$$


$$\begin{aligned} \hat{A} &= \frac{\sigma_A^2}{\sigma^2} \mathbf{1}^T \left(\mathbf{I} - \frac{\frac{\sigma_A^2}{\sigma^2} \mathbf{1}\mathbf{1}^T}{1 + \frac{\sigma_A^2}{\sigma^2} N} \right) \mathbf{x} \\ &= \frac{\sigma_A^2}{\sigma_A^2 + \frac{\sigma^2}{N}} \bar{x} \quad \text{——显示解} \end{aligned}$$

✓ 比MMSE更好求解!

✓ 但, 是**准最佳的!**

● 几何解释

假定 θ 和 $x[n]$ 是零均值的

矢量“长度”： $\|x\| = \sqrt{E(x^2)}$

线性约束： $\hat{\theta} = \sum_{n=0}^{N-1} a_n x[n]$

$$\text{Bmse}(\hat{\theta}) = E\left((\theta - \hat{\theta})^2\right)$$

$$\text{Bmse}(\hat{\theta}) = E\left(\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)^2\right)$$

$$\text{Bmse}(\hat{\theta}) = \left\| \theta - \sum_{n=0}^{N-1} a_n x[n] \right\|^2$$

——误差矢量长度的平方

误差： $\varepsilon = \theta - \hat{\theta}$

优化目标： $\min \{ \text{Bmse}(\hat{\theta}) \}$ \longleftrightarrow 误差矢量长度的平方最小化

$\varepsilon \perp \{x[0], x[1], x[2], \dots, x[N-1]\} \longrightarrow E((\theta - \hat{\theta})x[m]) = 0$ 正交原理

$$\left. \begin{aligned} E((\theta - \hat{\theta})x[m]) &= 0 \\ \hat{\theta} &= \sum_{n=0}^{N-1} a_n x[n] \end{aligned} \right\} \Rightarrow E\left(\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)x[m]\right) = 0 \Rightarrow$$

$$E(\theta x[m]) = \sum_{n=0}^{N-1} a_n E(x[m]x[n]) \Rightarrow$$

$$\begin{bmatrix} E(\theta x[0]) \\ E(\theta x[1]) \\ \vdots \\ E(\theta x[N-1]) \end{bmatrix} = \begin{bmatrix} E(x[0]x[0]) & E(x[0]x[1]) & \cdots & E(x[0]x[N-1]) \\ E(x[1]x[0]) & E(x[1]x[1]) & \cdots & E(x[1]x[N-1]) \\ \vdots & \vdots & \ddots & \vdots \\ E(x[N-1]x[0]) & E(x[N-1]x[1]) & \cdots & E(x[N-1]x[N-1]) \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_{N-1} \end{bmatrix}$$

$\mathbf{C}_{x\theta}$

\mathbf{C}_{xx}

\mathbf{a}

$$\Rightarrow \mathbf{C}_{x\theta} = \mathbf{C}_{xx} \mathbf{a} \Rightarrow \mathbf{a} = \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$

$$\left. \begin{aligned} \mathbf{a} &= \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \\ \hat{\theta} &= \sum_{n=0}^{N-1} a_n x[n] = \mathbf{a}^T \mathbf{x} \end{aligned} \right\} \Rightarrow \hat{\theta} = \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{x}$$

➤ 相应的贝叶斯均方误差:

$$\begin{aligned} \text{Bmse}(\hat{\theta}) &= E\left(\left(\theta - \hat{\theta}\right)^2\right) = E\left(\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)^2\right) \\ &= E\left(\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)\right) \\ &= E\left(\underbrace{\theta^2}_{\mathbf{C}_{\theta\theta}} - \underbrace{\sum_{n=0}^{N-1} a_n x[n]\theta}_{\mathbf{C}_{x\theta}} - \underbrace{\sum_{n=0}^{N-1} a_n x[n]\left(\theta - \sum_{n=0}^{N-1} a_n x[n]\right)}_0\right) \\ &= \mathbf{C}_{\theta\theta} - \mathbf{a}^T \mathbf{C}_{x\theta} \\ &= \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \end{aligned}$$

● 推广至矢量参数情况

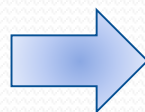
待估计参数 $\boldsymbol{\theta} = [\theta_1, \theta_2, \dots, \theta_p]^T$ ，其每个参数的LMMSE定义为：

$$\begin{cases} \min E\left[(\theta_i - \hat{\theta}_i)^2\right] \\ s.t. \quad \hat{\theta}_i = \sum_{n=0}^{N-1} a_{in} x[n] + a_{iN} \end{cases}$$

标量参数LMMSE:

$$\hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$\text{Bmse}(\hat{\theta}) = C_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta}$$



$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} E(\theta_1) \\ E(\theta_2) \\ \vdots \\ E(\theta_p) \end{bmatrix} + \begin{bmatrix} \mathbf{C}_{\theta_1 x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ \mathbf{C}_{\theta_2 x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ \vdots \\ \mathbf{C}_{\theta_p x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \end{bmatrix}$$

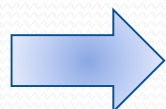
$$= E(\boldsymbol{\theta}) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

三、LMMSE性能评估及性质

1. LMMSE性能评估

$$\left. \begin{array}{l} \text{LMMSE估计量: } \hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ \text{估计量误差: } \varepsilon = \theta - \hat{\theta} \end{array} \right\}$$

——利用误差来进行性能评估


$$\varepsilon = \theta - E(\theta) - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

● 误差均值

$$\begin{aligned} E(\varepsilon) &= E\left\{\theta - E(\theta) - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))\right\} \\ &= E_{x,\theta}\left\{\theta - E(\theta) - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))\right\} \\ &= E_{x,\theta}(\theta - E(\theta)) - E_{x,\theta}(\mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))) \\ &= 0 \end{aligned}$$

即LMMSE估计量的误差是零均值的!

● 误差协方差矩阵:

其对角线元素为: $\left[E_{x,\theta}(\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^T) \right]_{ii} = E_{x,\theta} \left((\theta_i - \hat{\theta}_i)(\theta_i - \hat{\theta}_i)^T \right)$

$$\left. \begin{aligned} E_{x,\theta}(\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^T) &= E_{x,\theta} \left((\boldsymbol{\theta} - \hat{\boldsymbol{\theta}})(\boldsymbol{\theta} - \hat{\boldsymbol{\theta}})^T \right) \\ \hat{\boldsymbol{\theta}} &= E(\boldsymbol{\theta}) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \end{aligned} \right\} \Rightarrow$$

$$\begin{aligned} E_{x,\theta}(\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^T) &= E \left(\left(\boldsymbol{\theta} - E(\boldsymbol{\theta}) - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \right) \left(\boldsymbol{\theta} - E(\boldsymbol{\theta}) - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \right)^T \right) \\ &= E \left(\begin{aligned} &(\boldsymbol{\theta} - E(\boldsymbol{\theta}))(\boldsymbol{\theta} - E(\boldsymbol{\theta}))^T - (\boldsymbol{\theta} - E(\boldsymbol{\theta})) \left(\mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \right)^T \\ &- \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) (\boldsymbol{\theta} - E(\boldsymbol{\theta}))^T + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \left(\mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \right)^T \end{aligned} \right) \\ &= \mathbf{C}_{\theta\theta} - E \left\{ (\boldsymbol{\theta} - E(\boldsymbol{\theta})) (\mathbf{x} - E(\mathbf{x}))^T \right\} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \\ &\quad - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} E \left((\mathbf{x} - E(\mathbf{x})) (\boldsymbol{\theta} - E(\boldsymbol{\theta}))^T \right) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} E \left((\mathbf{x} - E(\mathbf{x})) (\mathbf{x} - E(\mathbf{x}))^T \right) \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \\ &= \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \end{aligned}$$

✓ 其对角线元素表示相应待估计参数的**最小贝叶斯均方误差**

✓ 也称为**贝叶斯均方误差矩阵**, 并常记: $\mathbf{M}_{\hat{\boldsymbol{\theta}}}$

2. LMMSE的性质

(1) 线性变换不变性

若 $\alpha = \mathbf{A}\theta + \mathbf{b}$ ， θ 的LMMSE估计是 $\hat{\theta}$ ，则 α 的LMMSE估计为：

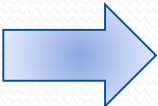
$$\hat{\alpha} = \mathbf{A}\hat{\theta} + \mathbf{b}$$

证明：

按定义，有 $\hat{\alpha} = E(\alpha) + \mathbf{C}_{\alpha x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$

$$E(\alpha) = E(\mathbf{A}\theta + \mathbf{b}) = \mathbf{A}E(\theta) + \mathbf{b}$$

$$\mathbf{C}_{\alpha x} = E\left((\alpha - E(\alpha))(\mathbf{x} - E(\mathbf{x}))^T\right) = E\left(\mathbf{A}(\theta - E(\theta))(\mathbf{x} - E(\mathbf{x}))^T\right) = \mathbf{A}\mathbf{C}_{\theta x}$$


$$\hat{\alpha} = \mathbf{A}E(\theta) + \mathbf{b} + \mathbf{A}\mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$= \mathbf{A}\left(E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))\right) + \mathbf{b}$$

$$= \mathbf{A}\hat{\theta} + \mathbf{b} \quad \text{得证!}$$

(2) 可加性

若 $\alpha = \theta_1 + \theta_2$, θ_1 和 θ_2 的LMMSE估计分别是 $\hat{\theta}_1$ 、 $\hat{\theta}_2$, 则 α 的LMMSE估计为:

$$\hat{\alpha} = \hat{\theta}_1 + \hat{\theta}_2$$

证明:

按定义, 有 $\hat{\alpha} = E(\alpha) + \mathbf{C}_{\alpha x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$

$$\alpha = \theta_1 + \theta_2 \quad \Rightarrow \quad E(\alpha) = E(\theta_1) + E(\theta_2)$$

$$\begin{aligned} \mathbf{C}_{\alpha x} &= E\left((\alpha - E(\alpha))(\mathbf{x} - E(\mathbf{x}))^T\right) \\ &= E\left((\theta_1 + \theta_2 - E(\theta_1 + \theta_2))(\mathbf{x} - E(\mathbf{x}))^T\right) \\ &= E\left(((\theta_1 - E(\theta_1)) + (\theta_2 - E(\theta_2))) (\mathbf{x} - E(\mathbf{x}))^T\right) \\ &= \mathbf{C}_{\theta_1 x} + \mathbf{C}_{\theta_2 x} \end{aligned}$$

$$\hat{\alpha} = E(\theta_1) + E(\theta_2) + (\mathbf{C}_{\theta_1 x} + \mathbf{C}_{\theta_2 x}) \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$= E(\theta_1) + \mathbf{C}_{\theta_1 x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) + E(\theta_2) + \mathbf{C}_{\theta_2 x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) = \hat{\theta}_1 + \hat{\theta}_2 \quad \text{得证!}$$

(3) 贝叶斯高斯-马尔可夫定理

如果数据具有贝叶斯一般线性模型的形式，即

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

其中 \mathbf{H} 是已知的 $N \times p$ 观测矩阵； $\boldsymbol{\theta}$ 是 $p \times 1$ 的随机矢量，其均值为 $E(\boldsymbol{\theta})$ ，协方差矩阵为 $\mathbf{C}_{\theta\theta}$ ，其现实待估计； \mathbf{w} 是 $N \times 1$ 的随机矢量，其均值为零、协方差为 \mathbf{C}_w ，且与 $\boldsymbol{\theta}$ 不相关（联合PDF $p(\mathbf{x};\boldsymbol{\theta})$ 是任意的）。那么， $\boldsymbol{\theta}$ 的LMMSE估计量是

$$\begin{aligned}\hat{\boldsymbol{\theta}} &= E(\boldsymbol{\theta}) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x})) \\ &= E(\boldsymbol{\theta}) + \mathbf{C}_{\theta\theta} \mathbf{H}^T (\mathbf{H} \mathbf{C}_{\theta\theta} \mathbf{H}^T + \mathbf{C}_w)^{-1} (\mathbf{x} - \mathbf{H} E(\boldsymbol{\theta})) \\ &= E(\boldsymbol{\theta}) + (\mathbf{C}_{\theta\theta}^{-1} + \mathbf{H}^T \mathbf{C}_w^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}_w^{-1} (\mathbf{x} - \mathbf{H} E(\boldsymbol{\theta}))\end{aligned}$$

误差 $\boldsymbol{\varepsilon} = \boldsymbol{\theta} - \hat{\boldsymbol{\theta}}$ 的均值为零、协方差矩阵为

$$\begin{aligned}\mathbf{C}_{\varepsilon} &= E(\boldsymbol{\varepsilon} \boldsymbol{\varepsilon}^T) \\ &= \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{C}_{x\theta} \\ &= \mathbf{C}_{\theta\theta} - \mathbf{C}_{\theta\theta} \mathbf{H}^T (\mathbf{H} \mathbf{C}_{\theta\theta} \mathbf{H}^T + \mathbf{C}_w)^{-1} \mathbf{H} \mathbf{C}_{\theta\theta} \\ &= (\mathbf{C}_{\theta\theta}^{-1} + \mathbf{H}^T \mathbf{C}_w^{-1} \mathbf{H})^{-1}\end{aligned}$$

若为高斯分布，
则LMMSE为
MMSE

$$\text{Bmse}(\hat{\theta}_i) = [\mathbf{C}_{\varepsilon}]_{ii} = [\mathbf{M}_{\hat{\theta}}]_{ii}$$

四、序贯LMMSE估计

例：白噪声中电平估计问题：

$$x[n] = A + w[n], \quad n = 0, 1, \dots, N-1$$

待估计参数为信号幅度 A ，其服从 $N(0, \sigma_A^2)$ ， $w[n]$ 为高斯白噪声，且其方差为 σ^2 ，即 $w[n] \sim N(0, \sigma^2)$ ，且与幅度不相关。 A 的LMMSE估计量为？

$$\hat{\theta} = E(\theta) + (\mathbf{C}_{\theta\theta}^{-1} + \mathbf{H}^T \mathbf{C}_w^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}_w^{-1} (\mathbf{x} - \mathbf{H} E(\theta))$$

$$\text{Bmse}(\hat{\theta}) = (\mathbf{C}_{\theta\theta}^{-1} + \mathbf{H}^T \mathbf{C}_w^{-1} \mathbf{H})^{-1}$$

$$E(\theta) = 0$$

$$\mathbf{C}_{\theta\theta} = \sigma_A^2$$

$$\mathbf{H} = [1, 1, 1, \dots, 1]^T$$

$$\mathbf{C}_w = \sigma^2 \mathbf{I}$$

$$\hat{A} = \frac{\sigma_A^2}{\sigma_A^2 + \frac{\sigma^2}{N}} \bar{x}$$

$$\text{Bmse}(\hat{A}) = \frac{\sigma_A^2 \sigma^2}{N \sigma_A^2 + \sigma^2}$$

LMMSE估计量:

$$\hat{A} = \frac{\sigma_A^2}{\sigma_A^2 + \frac{\sigma^2}{N}} \bar{x}$$

重记为:

$$\hat{A}[N-1] = \frac{\sigma_A^2}{\sigma_A^2 + \frac{\sigma^2}{N}} \frac{1}{N} \sum_{n=0}^{N-1} x[n]$$

$$\hat{A}[N] = \frac{\sigma_A^2}{\sigma_A^2 + \frac{\sigma^2}{N+1}} \frac{1}{N+1} \sum_{n=0}^N x[n]$$

$$= \frac{N\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} \frac{1}{N} \left(\sum_{n=0}^{N-1} x[n] + x[N] \right)$$

$$\begin{aligned} &= \frac{N\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} \frac{1}{N} \sum_{n=0}^{N-1} x[n] + \frac{N\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} \frac{1}{N} x[N] \\ &= \frac{N\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} \frac{\sigma_A^2 + \frac{\sigma^2}{N}}{\sigma_A^2} \hat{A}[N-1] + \frac{\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} x[N] \\ &= \frac{N\sigma_A^2 + \sigma^2}{(N+1)\sigma_A^2 + \sigma^2} \hat{A}[N-1] + \frac{\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} x[N] \\ &= \hat{A}[N-1] + \underbrace{\frac{\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2}}_{K[N]} (x[N] - \hat{A}[N-1]) \end{aligned}$$

$K[N]$ 增益因子

$$\begin{aligned} &\frac{\frac{1}{\sigma^2}}{\frac{1}{\text{Bmse}(\hat{A}[N-1])} + \frac{1}{\sigma^2}} = \frac{\text{Bmse}(\hat{A}[N-1])}{\text{Bmse}(\hat{A}[N-1]) + \sigma^2} \\ &= \frac{\frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2}}{\frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2} + \sigma^2} = \frac{\sigma_A^2}{(N\sigma_A^2 + \sigma^2) + \sigma_A^2} = K[N] \end{aligned}$$

最小贝叶斯MSE:

$$\text{Bmse}(\hat{A}) = \frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2}$$

重记为:

$$\text{Bmse}(\hat{A}[N-1]) = \frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2}$$

$$\text{Bmse}(\hat{A}[N]) = \frac{\sigma_A^2 \sigma^2}{(N+1)\sigma_A^2 + \sigma^2}$$

$$= \frac{N\sigma_A^2 + \sigma^2}{(N+1)\sigma_A^2 + \sigma^2} \frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2}$$

$$= \left(1 - \frac{\sigma_A^2}{(N+1)\sigma_A^2 + \sigma^2} \right) \frac{\sigma_A^2 \sigma^2}{N\sigma_A^2 + \sigma^2}$$

$$= (1 - K[N]) \text{Bmse}(\hat{A}[N-1])$$

序贯计算方法

初始化:

$$\hat{A}[-1] = E(A)$$

$$\text{Bmse}(\hat{A}[-1]) = \text{var}(A)$$

增益因子:

$$K[N] = \frac{\text{Bmse}(\hat{A}[N-1])}{\text{Bmse}(\hat{A}[N-1]) + \sigma^2}$$

估计量更新:

$$\hat{A}[N] = \hat{A}[N-1] + K[N](x[N] - \hat{A}[N-1])$$

最小贝叶斯MSE更新:

$$\text{Bmse}(\hat{A}[N]) = (1 - K[N]) \text{Bmse}(\hat{A}[N-1])$$

一般化，且可推广至矢量参数情况

信号模型：

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

待估计参数 $\boldsymbol{\theta}$ ，其均值为 $E(\boldsymbol{\theta})$ ，协方差矩阵为 $\mathbf{C}_{\theta\theta}$ 。 \mathbf{w} 为零均值**不相关噪声**，且 $\text{var}(w[n]) = \sigma_n^2$ ， \mathbf{w} 与 $\boldsymbol{\theta}$ 不相关。

对 $\boldsymbol{\theta}$ 的LMMSE估计可采用如下**序贯**方式进行：

初始化： $\hat{\boldsymbol{\theta}}[-1] = E(\boldsymbol{\theta}) \quad \mathbf{M}[-1] = \mathbf{C}_{\theta\theta}$

增益因子： $\mathbf{K}[n] = \frac{\mathbf{M}[n-1]\mathbf{h}[n]}{\sigma_n^2 + \mathbf{h}^T[n]\mathbf{M}[n-1]\mathbf{h}[n]}$

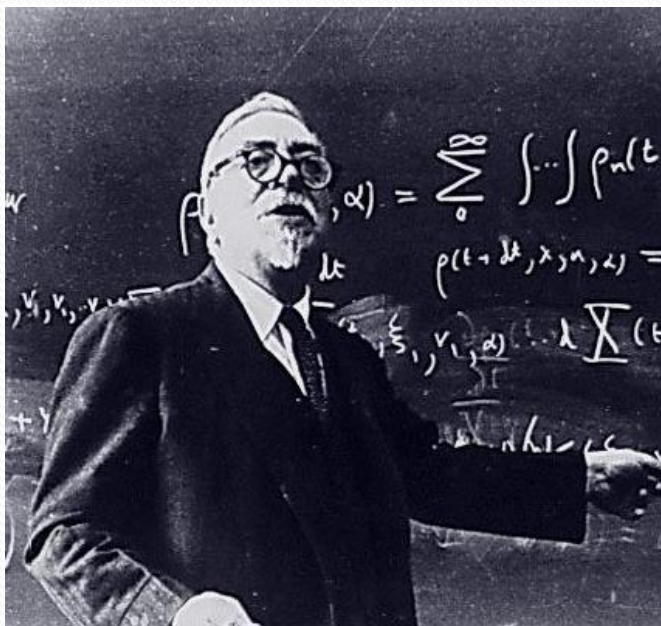
$$\mathbf{H}[n] = \begin{bmatrix} \mathbf{H}[n-1] \\ \mathbf{h}^T[n] \end{bmatrix}$$
$$x[n] = \mathbf{h}^T[n]\boldsymbol{\theta} + w[n]$$

估计量更新： $\hat{\boldsymbol{\theta}}[n] = \hat{\boldsymbol{\theta}}[n-1] + \mathbf{K}[n](x[n] - \mathbf{h}^T[n]\hat{\boldsymbol{\theta}}[n-1])$

最小贝叶斯MSE更新： $\mathbf{M}[n] = (\mathbf{I} - \mathbf{K}[n]\mathbf{h}^T[n])\mathbf{M}[n-1]$

五、LMMSE的应用

● 维纳滤波



- ❑ 诺伯特·维纳(Norbert Wiener, 1894-1964), 美国数学家, 控制论的创始人。1894年11月26日生于密苏里州的哥伦比亚, 1964年3月18日卒于斯德哥尔摩
- ❑ 维纳在其50年的科学生涯中, 先后涉足哲学、数学、物理学和工程学, 最后转向生物学, 在各个领域中都取得了丰硕成果, 是一位多才多艺、学识渊博的科学巨人
- ❑ 他一生发表论文240多篇, 著作14本。他的主要著作有《控制论》(1948)、《维纳选集》(1964)和《维纳数学论文集》(1980)等

$$x[n] = s[n] + w[n], \quad n = 0, 1, \dots, N-1$$

假定观测数据是零均值的WSS，信号也是零均值的，信号与噪声不相关

- 滤波

$\theta = s[n]$ 用 $x[0], x[1], x[2], \dots, x[n]$ 来估计

- 平滑

$\theta = s[n]$ 用 $x[0], x[1], x[2], \dots, x[N-1]$ 来估计

- 预测

$\theta = x[N-1+l]$ 用 $x[0], x[1], x[2], \dots, x[N-1]$ 来估计

● 滤波

$\theta = s[n]$ 用 $x[0], x[1], x[2], \dots, x[n]$ 来估计

$$\text{LMMSE: } \hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$\hat{s}[n] = \hat{\theta} = \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{x}$$

$$\begin{aligned} \mathbf{C}_{\theta x} &= E\left(s[n] \begin{bmatrix} x[0], x[1], \dots, x[n] \end{bmatrix}\right) \\ &= E\left(s[n] \begin{bmatrix} s[0], s[1], \dots, s[n] \end{bmatrix}\right) \\ &\quad + E\left(s[n] \begin{bmatrix} w[0], w[1], \dots, w[n] \end{bmatrix}\right) \\ &= \begin{bmatrix} r_{ss}[n], r_{ss}[n-1], \dots, r_{ss}[0] \end{bmatrix} \\ &\triangleq \mathbf{r}_{ss}^T \end{aligned}$$

$$\begin{aligned} \mathbf{C}_{xx} &= E(\mathbf{x} \mathbf{x}^T) \\ &= E\left((\mathbf{s} + \mathbf{w})(\mathbf{s} + \mathbf{w})^T\right) \\ &= E(\mathbf{s} \mathbf{s}^T) + E(\mathbf{w} \mathbf{w}^T) \\ &= \mathbf{C}_{ss} + \mathbf{C}_{ww} \\ &= \mathbf{R}_{ss} + \mathbf{R}_{ww} \end{aligned}$$

$$\Rightarrow \hat{s}[n] = \mathbf{r}_{ss}^T (\mathbf{R}_{ss} + \mathbf{R}_{ww})^{-1} \mathbf{x}$$

滤波器: $\hat{s}[n] = \sum_{k=0}^n h^{(n)}[n-k]x[k] = \sum_{k=0}^n h^{(n)}[k]x[n-k]$

$$\left. \begin{aligned} \hat{s}[n] &= \mathbf{r}_{ss}^T (\mathbf{R}_{ss} + \mathbf{R}_{ww})^{-1} \mathbf{x} \\ \hat{s}[n] &= \mathbf{a}^T \mathbf{x} = \sum_{k=0}^n a_k x[k] \end{aligned} \right\} \Rightarrow \mathbf{a} = (\mathbf{R}_{ss} + \mathbf{R}_{ww})^{-1} \mathbf{r}_{ss} \Rightarrow (\mathbf{R}_{ss} + \mathbf{R}_{ww}) \mathbf{a} = \mathbf{r}_{ss} \left. \begin{aligned} &\text{令 } h^{(n)}[k] = a_{n-k}, k = 0, 1, \dots, n \end{aligned} \right\} \Rightarrow$$

$(\mathbf{R}_{ss} + \mathbf{R}_{ww}) \mathbf{h} = \mathbf{r}_{ss}$, 其中 $\mathbf{r}_{ss} = [r_{ss}[0], r_{ss}[1], \dots, r_{ss}[n]]^T$

\mathbf{R}_{xx}

$$\Rightarrow \begin{bmatrix} r_{xx}[0] & r_{xx}[1] & \cdots & r_{xx}[n] \\ r_{xx}[1] & r_{xx}[0] & \cdots & r_{xx}[n-1] \\ \vdots & \vdots & \ddots & \vdots \\ r_{xx}[n] & r_{xx}[n-1] & \cdots & r_{xx}[0] \end{bmatrix} \begin{bmatrix} h^{(n)}[0] \\ h^{(n)}[1] \\ \vdots \\ h^{(n)}[n] \end{bmatrix} = \begin{bmatrix} r_{ss}[0] \\ r_{ss}[1] \\ \vdots \\ r_{ss}[n] \end{bmatrix}$$

维纳-霍夫滤波方程

Levinson递推算法

- 平滑

$\theta = s[n]$ 用 $\dots, x[-1], x[0], x[1], x[2], \dots$ 来估计

$$\text{LMMSE: } \hat{s}[n] = \sum_{k=-\infty}^{\infty} a_k x[k]$$

令 $h[k] = a_{n-k}$

$$\hat{s}[n] = \sum_{k=-\infty}^{\infty} h[k] x[n-k]$$

正交原理

$$E\left((s[n] - \hat{s}[n])x[m]\right) = 0$$

$$E(s[n]x[m]) = E(\hat{s}[n]x[m])$$

$$E(s[n](s[m] + w[m])) = E\left(\sum_{k=-\infty}^{\infty} h[k] x[n-k] x[m]\right)$$

$$E\left(s[n](s[m] + w[m])\right) = E\left(\sum_{k=-\infty}^{\infty} h[k]x[n-k]x[m]\right)$$

$$r_{ss}[n-m] = \sum_{k=-\infty}^{\infty} h[k]E(x[n-k]x[m])$$

$$r_{ss}[n-m] = \sum_{k=-\infty}^{\infty} h[k]r_{xx}[n-m-k]$$

$$r_{ss}[l] = \sum_{k=-\infty}^{\infty} h[k]r_{xx}[l-k], \quad -\infty < l < \infty$$

$$r_{ss}[n] = h[n] * r_{xx}[n]$$

$$H(f) = \frac{P_{ss}(f)}{P_{xx}(f)} = \frac{P_{ss}(f)}{P_{ss}(f) + P_{ww}(f)} = \frac{\eta(f)}{\eta(f) + 1}, \text{ 其中 } \eta(f) = \frac{P_{ss}(f)}{P_{ww}(f)}$$

——无限维纳平滑器的频率响应

● 预测

$\theta = x[N-1+l]$ 用 $x[0], x[1], x[2], \dots, x[N-1]$ 来估计

$$\text{LMMSE: } \hat{\theta} = E(\theta) + \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

$$\hat{x}[N-1+l] = \mathbf{C}_{\theta x} \mathbf{C}_{xx}^{-1} \mathbf{x} \longrightarrow$$

$$\mathbf{C}_{\theta x} = E\left(x[N-1+l] \begin{bmatrix} x[0] & x[1] & \dots & x[N-1] \end{bmatrix}\right)$$

$$= \begin{bmatrix} r_{xx}[N-1+l] & r_{xx}[N-2+l] & \dots & r_{xx}[l] \end{bmatrix}$$

$$\triangleq \mathbf{r}_{xx}^T \longrightarrow$$

$$\mathbf{C}_{xx} = \mathbf{R}_{xx} \longrightarrow$$

$$\hat{x}[N-1+l] = \mathbf{r}_{xx}^T \mathbf{R}_{xx}^{-1} \mathbf{x}$$

$$\text{“滤波”器: } \hat{x}[N-1+l] = \sum_{k=0}^{N-1} h[N-k]x[k] = \sum_{k=1}^N h[k]x[N-k]$$

$$\left. \begin{aligned} \hat{x}[N-1+l] &= \mathbf{r}_{xx}^T \mathbf{R}_{xx}^{-1} \mathbf{x} \\ \hat{x}[N-1+l] &= \mathbf{a}^T \mathbf{x} = \sum_{k=0}^n a_k x[k] \end{aligned} \right\} \Rightarrow \mathbf{a} = \mathbf{R}_{xx}^{-1} \mathbf{r}_{xx} \Rightarrow \mathbf{R}_{xx} \mathbf{a} = \mathbf{r}_{xx} \left. \begin{aligned} \text{令 } h[k] &= a_{N-k} \end{aligned} \right\}$$

$$\Rightarrow \mathbf{R}_{xx} \mathbf{h} = \mathbf{r}_{xx}, \text{ 其中 } \mathbf{r}_{xx} = [r_{xx}[l], r_{xx}[l+1], \dots, r_{xx}[N-1+l]]^T$$

$$\Rightarrow \begin{bmatrix} r_{xx}[0] & r_{xx}[1] & \cdots & r_{xx}[N-1] \\ r_{xx}[1] & r_{xx}[0] & \cdots & r_{xx}[N-2] \\ \vdots & \vdots & \ddots & \vdots \\ r_{xx}[N-1] & r_{xx}[N-2] & \cdots & r_{xx}[0] \end{bmatrix} \begin{bmatrix} h[1] \\ h[2] \\ \vdots \\ h[N] \end{bmatrix} = \begin{bmatrix} r_{xx}[l] \\ r_{xx}[l+1] \\ \vdots \\ r_{xx}[N-1+l] \end{bmatrix}$$

线性预测维纳-霍夫方程

Levinson递推算法

六、小结

- LMMSE估计
 - 限定估计量是线性的，然后再找贝叶斯MSE最小者
 - 无需PDF，仅需前两阶矩即可
 - 一般情况下并非最佳（在贝叶斯MSE准则下），但相比MMSE较易得到显示解
- LMMSE的性质
 - 线性变换不变性
 - 可加性
- 贝叶斯高斯马尔科夫定理
- 对贝叶斯一般线性模型，LMMSE与MMSE具有相同形式
- 序贯LMMSE
- LMMSE的“应用”
 - 维纳滤波
 - 卡尔曼滤波