

$$\underline{y}^T \underline{C}^{-1} \underline{x} = \underline{x}^T \underline{C}^{-1} \underline{y} > 0$$

$$4.12) \quad \underline{x}^T \underline{C}^{-1} \underline{y} = \underline{x}^T \sum_{i=0}^{N-1} \frac{1}{\lambda_i} \underline{v}_i \underline{v}_i^H \underline{y}$$

$$= \sum_{i=0}^{N-1} \frac{(\underline{v}_i^T \underline{x})(\underline{v}_i^H \underline{y})}{\lambda_i}$$

$$\approx \sum_{i=0}^{N-1} \frac{\frac{1}{\sqrt{N}} \sum_{n=0}^{N-1} x(n) e^{j2\pi f_i n}}{\frac{1}{\sqrt{N}} \sum_{n=0}^{N-1} y(n) e^{-j2\pi f_i n}}$$

$P_{WW}(f_i/N)$

$$= \sum_{i=0}^{N-1} \frac{X^*(f_i) Y(f_i)}{P_{WW}(f_i)} \frac{1}{N}$$

where $f_i = 1/N$

$$= \sum_{i=0}^{N-1} \frac{X(f_i) Y^*(f_i)}{P_{WW}(f_i)} \frac{1}{N} \quad \text{since } \underline{x}^T \underline{C}^{-1} \underline{y} \text{ is real}$$

$$\rightarrow \int_0^1 \frac{X(f) Y^*(f)}{P_{WW}(f)} df = \int_0^{\frac{1}{2}} \frac{X(f) Y^*(f)}{P_{WW}(f)} df$$

$$4.13) \quad d^2 = \underline{y}^T \underline{C}^{-1} \underline{y}$$

$$= \underline{y}^T \sum_i \frac{1}{\lambda_i} \underline{v}_i \underline{v}_i^H \underline{y}$$

$$= \sum_i \frac{|\underline{v}_i^H \underline{y}|^2}{\lambda_i}$$

$$\approx \sum_{i=0}^{N-1} \frac{\frac{1}{N} \left| \sum_{n=0}^{N-1} y(n) e^{-j2\pi f_i n} \right|^2}{P_{WW}(f_i)}$$

$$= \sum_{i=0}^{N-1} \frac{|S(f_i)|^2}{P_{WW}(f_i)} \frac{L}{N}$$

$$\rightarrow \int_0^1 \frac{|S(f)|^2}{P_{WW}(f)} df$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{|S(f)|^2}{P_{WW}(f)} df$$

$$4.14) y(n) = \sum_{k=-\infty}^{\infty} h(k) s(n-k)$$

$$\Rightarrow Y(f) = H(f) S(f)$$

$$y(n) = \int_{-\frac{1}{2}}^{\frac{1}{2}} H(f) S(f) e^{j2\pi f n} df$$

$$y(N-1) = \int_{-\frac{1}{2}}^{\frac{1}{2}} H(f) S(f) e^{j2\pi f (N-1)} df$$

$$w(n) \rightarrow \boxed{H(f)} \rightarrow y(n) = \sum_{k=-\infty}^{\infty} h(k) w(n-k)$$

A standard result is that if $w(n)$ is WSS then so is $y(n)$ and

$$P_{yy}(f) = |H(f)|^2 P_{ww}(f)$$

$$\Rightarrow E(y^2[N-1]) = r_{yy}(0)$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} P_{yy}(f) df$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} |H(f)|^2 P_{ww}(f) df$$

To maximize η let

$$g(f) = H(f) \sqrt{P_{WW}(f)}$$

$$h(f) = \frac{S(f) e^{j2\pi f(N-1)}}{\sqrt{P_{WW}(f)}}$$

$$\eta \leq \int_{-\frac{1}{2}}^{\frac{1}{2}} |h(f)|^2 df$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{|S(f)|^2}{P_{WW}(f)} df = \eta_{\text{MAX}}$$

and the maximum is attained for

$$g(f) = c h^*(f)$$

$$\frac{H(f) \sqrt{P_{WW}(f)}}{\sqrt{P_{WW}(f)}} = \frac{c S^*(f) e^{-j2\pi f(N-1)}}{\sqrt{P_{WW}(f)}}$$

or letting $c=1$

$$H(f) = \frac{S^*(f) e^{-j2\pi f(N-1)}}{P_{WW}(f)}$$

$$4.15) \quad T(\underline{x}) = \underline{x}^T \underline{C}^{-1} \underline{S}$$

$$= \sum_{n=0}^{N-1} \frac{x(n) S(n)}{\sigma_n^2}$$

$$= A \sum_{n=0}^{N-1} \frac{x(n)}{\sigma^2 r^n}$$

or we decide H_1 if

$$\sum_{n=0}^{N-1} \frac{x(n)}{r^n} > \tau$$

$$P_D = Q\left(Q^{-1}(P_{FA}) - \sqrt{\underline{s}^T \underline{C}^{-1} \underline{s}}\right)$$

$$\underline{s}^T \underline{C}^{-1} \underline{s} = A^2 \sum_{n=0}^{N-1} \frac{1}{\sigma^2 r^n}$$

$$= \frac{A^2}{\sigma^2} \sum_{n=0}^{N-1} 1/r^n$$

As $N \rightarrow \infty$ if $0 < r \leq 1$, $\underline{s}^T \underline{C}^{-1} \underline{s} \rightarrow \infty$
 $\Rightarrow P_D \rightarrow 1$

if $r > 1$, $\underline{s}^T \underline{C}^{-1} \underline{s} \rightarrow \frac{A^2}{\sigma^2} \frac{1}{1-1/r}$

$$= \frac{A^2}{\sigma^2} \frac{r}{r-1}$$

In first case the noise "dies out"

if $0 < r < 1$ and if $r = 1$ we have a DC level in WGN so that averaging causes $P_D \rightarrow 1$.

4.16) $d^2 = \underline{s}^T \underline{C}^{-1} \underline{s}$

Using Woodbury's identity

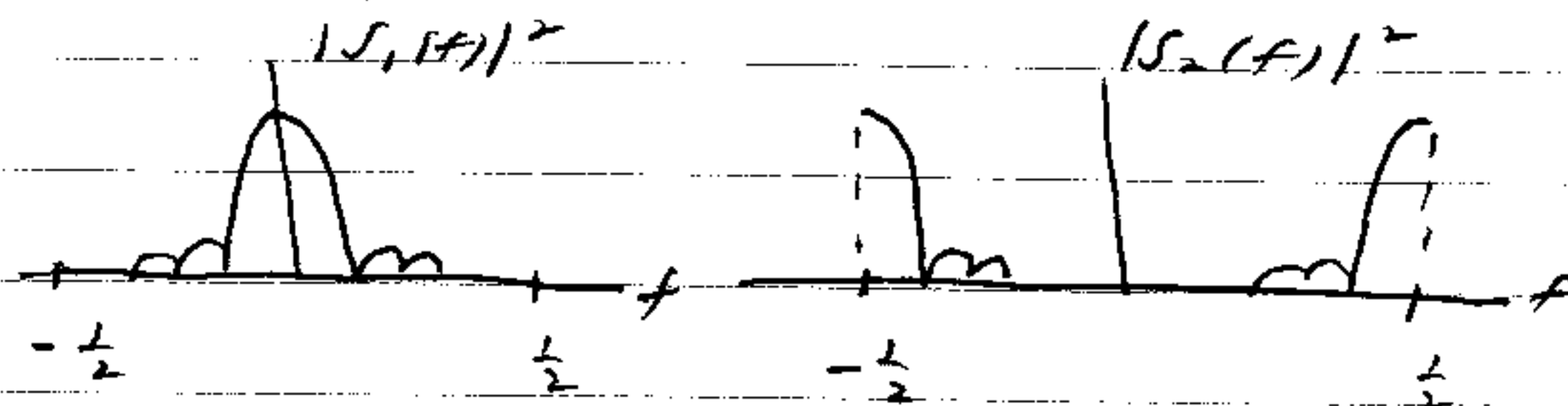
$$\underline{C}^{-1} = \frac{1}{\sigma^2} \underline{I} - \frac{1}{\sigma^4} \frac{\underline{P} \underline{1} \underline{1}^T}{1 + NP/\sigma^2}$$

$$d^2 = \frac{\underline{s}^T \underline{s}}{\sigma^2} - \frac{P/\sigma^4 (\underline{s}^T \underline{1})^2}{1 + NP/\sigma^2}$$

$$= \frac{\underline{s}^T \underline{s}}{\sigma^2} - \frac{1}{\sigma^2} \frac{P}{NP + \sigma^2} (\underline{s}^T \underline{1})^2$$

The better signal is the one that minimizes $\underline{s}^T \underline{1} = \sum_{n=0}^{N-1} s(n)$ which is $s_2(n)$.

This is because the noise PSD is $P_{NW}(f) = \sigma^2 + P \delta(f)$ or there is an impulse at DC. The FT magnitudes of the signals are



$s_2(n)$ has better overall SNR.

$$4.17) \quad T(x) = \int \frac{X(f) S^*(f)}{P_{NW}(f)} df$$

$$= \int \frac{1}{\sigma^2} X(f) S^*(f) |1 + a e^{-j2\pi f}|^2 df$$

$$= \frac{1}{\sigma^2} \int [A(f) X(f)] [A(f) S(f)]^* df$$

$$= \frac{1}{\sigma^2} \sum_{n=-\infty}^{\infty} (x(n) + a x(n-1)) (s(n) + a s(n-1))$$

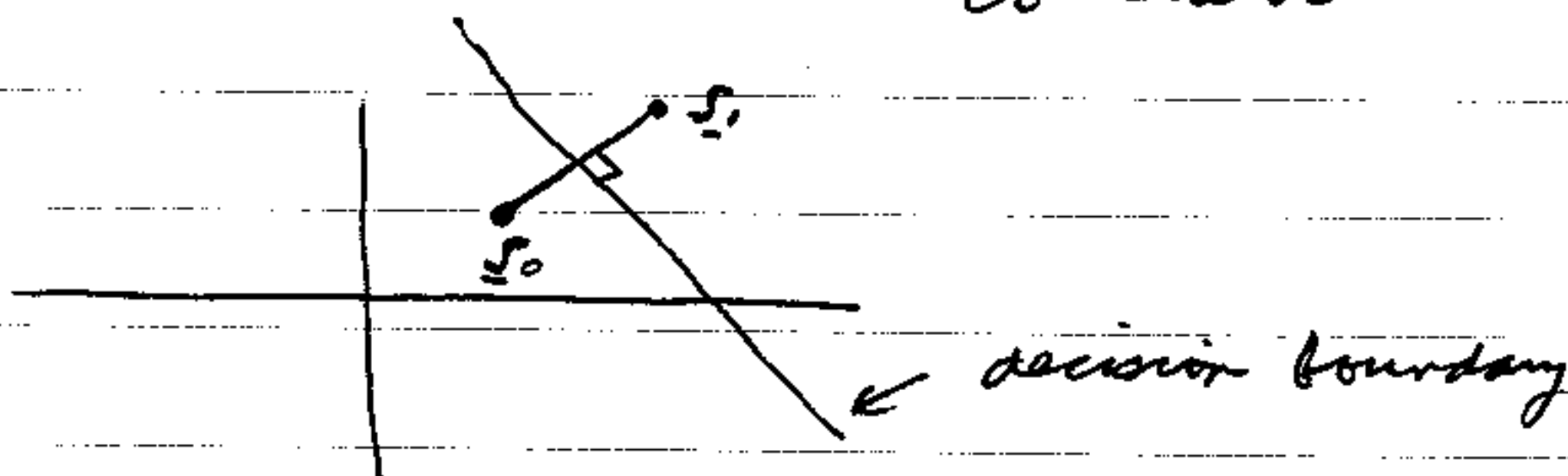
by Parseval's theorem

$$= \frac{1}{\sigma^2} \sum_{n=0}^N (x(n) + a x(n-1)) (s(n) + a s(n-1))$$

$$\approx \frac{1}{\sigma^2} \sum_{n=1}^{N-1} \underbrace{(x(n) + a x(n-1))}_{\text{prewhitener}} \underbrace{(s(n) + a s(n-1))}_{\text{modified signal}}$$

Correlator

4.18)



A decision boundary is a hyperplane whose points satisfy $\|x - s_0\| = \|x - s_1\|$

The midpoint of the line segment is

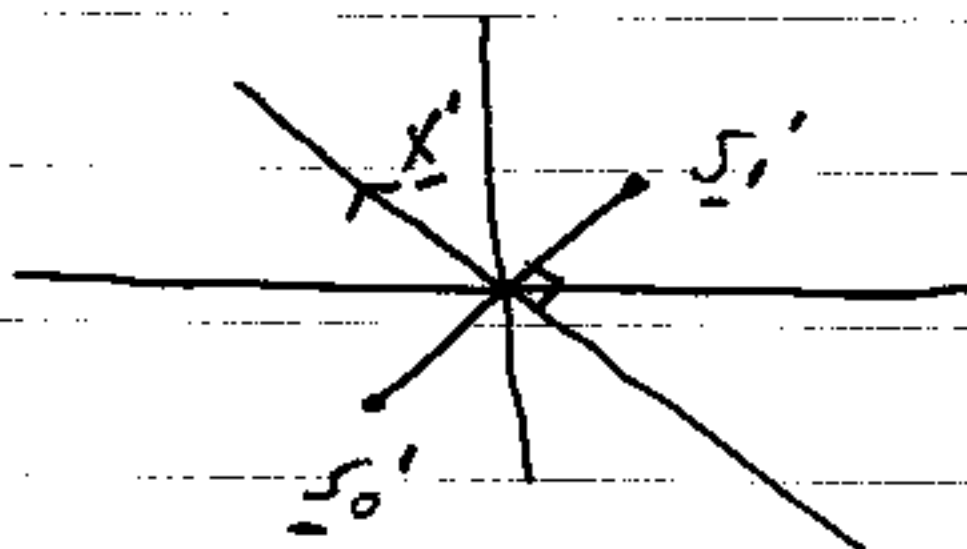
$$\frac{s_0 + s_1}{2} \text{ so that } \left\| \frac{s_0 + s_1}{2} - s_0 \right\| =$$

$$\left\| \frac{s_1}{2} - \frac{s_0}{2} \right\| = \left\| \frac{s_0 + s_1}{2} - s_1 \right\| \text{ and}$$

it lies on hyperplane. To show that

line segment is perpendicular let

$$\underline{x}' = \underline{x} - \frac{\underline{s}_0 + \underline{s}_1}{2} \text{ so that}$$



$$\underline{x}'^T \left(\frac{\underline{s}_1' - \underline{s}_0'}{2} \right) = \left(\underline{x} + \frac{\underline{s}_0 + \underline{s}_1}{2} \right)^T \left(\frac{\underline{s}_1 - \underline{s}_0}{2} \right)$$

$$= \frac{1}{2} \underline{x}^T (\underline{s}_1 - \underline{s}_0) + \frac{1}{4} \underline{s}_0^T \underline{s}_0 - \frac{1}{4} \underline{s}_1^T \underline{s}_1$$

$$= \frac{1}{4} \left[(\underline{x} - \underline{s}_0)^T (\underline{x} - \underline{s}_0) - (\underline{x} - \underline{s}_1)^T (\underline{x} - \underline{s}_1) \right]$$

$$= \frac{1}{4} (||\underline{x} - \underline{s}_0||^2 - ||\underline{x} - \underline{s}_1||^2) = 0$$

4.19) Using the MAP rule we decide \mathcal{H}_1 if

$$Z_1 = \ln p(\underline{x} | \mathcal{H}_1) + \ln P(\mathcal{H}_1)$$

$$> \ln p(\underline{x} | \mathcal{H}_0) + \ln P(\mathcal{H}_0) = Z_0$$

$$\text{But } p(\underline{x} | \mathcal{H}_i) = \frac{1}{2\pi} e^{-\frac{1}{2} (\underline{x} - \underline{s}_i)^T (\underline{x} - \underline{s}_i)}$$

for $N=2$ and $\sigma^2=1$

$$Z_i = -\ln 2\pi - \frac{1}{2} (\underline{x} - \underline{s}_i)^T (\underline{x} - \underline{s}_i) + \ln P(\mathcal{H}_i)$$

The decision boundary is obtained by letting $\underline{z}_1 = \underline{z}_0$ or

$$-\frac{1}{2}(\underline{x} - \underline{z}_1)^T(\underline{x} - \underline{z}_1) + \ln P(\mathcal{H}_1) =$$

$$-\frac{1}{2}(\underline{x} - \underline{z}_0)^T(\underline{x} - \underline{z}_0) + \ln P(\mathcal{H}_0)$$

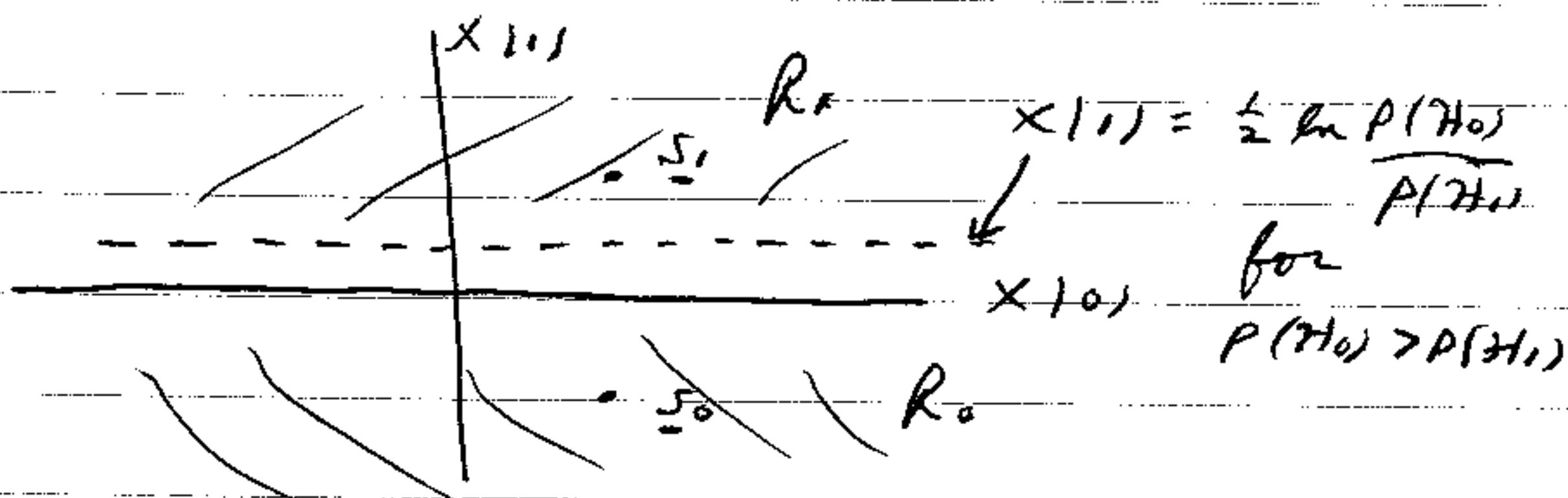
$$\underline{x}^T \underline{z}_1 - \frac{1}{2} \varepsilon_1 + \ln P(\mathcal{H}_1) = \underline{x}^T \underline{z}_0 - \frac{1}{2} \varepsilon_0 + \ln P(\mathcal{H}_0)$$

$$\underline{x}^T (\underline{z}_1 - \underline{z}_0) = \frac{1}{2} (\varepsilon_1 - \varepsilon_0) + \ln \frac{P(\mathcal{H}_0)}{P(\mathcal{H}_1)}$$

For given $\underline{z}_0, \underline{z}_1$ we have

$$\underline{x}^T \begin{bmatrix} 0 \\ 2 \end{bmatrix} = \ln P(\mathcal{H}_0)/P(\mathcal{H}_1)$$

$$x(1) = \frac{1}{2} \ln P(\mathcal{H}_0)/P(\mathcal{H}_1)$$



If $P(\mathcal{H}_0) = P(\mathcal{H}_1)$, we decide \mathcal{H}_1 if $x(1) > 0$ (minimum distance) and if $P(\mathcal{H}_0) > P(\mathcal{H}_1)$ we move decision boundary to include more of \underline{z} in R_0 .

$$4.20) \quad (\underline{s}_1 + \underline{s}_0)^T (\underline{s}_1 + \underline{s}_0) \geq 0$$

$$\underline{s}_1^T \underline{s}_1 + 2 \underline{s}_1^T \underline{s}_0 + \underline{s}_0^T \underline{s}_0 \geq 0$$

$$\frac{1}{2} (\underline{s}_1^T \underline{s}_1 + \underline{s}_0^T \underline{s}_0) \geq -\underline{s}_1^T \underline{s}_0$$

$$1 \geq \mp \rho_r \Rightarrow \rho_r \leq 1, -\rho_r \leq 1$$

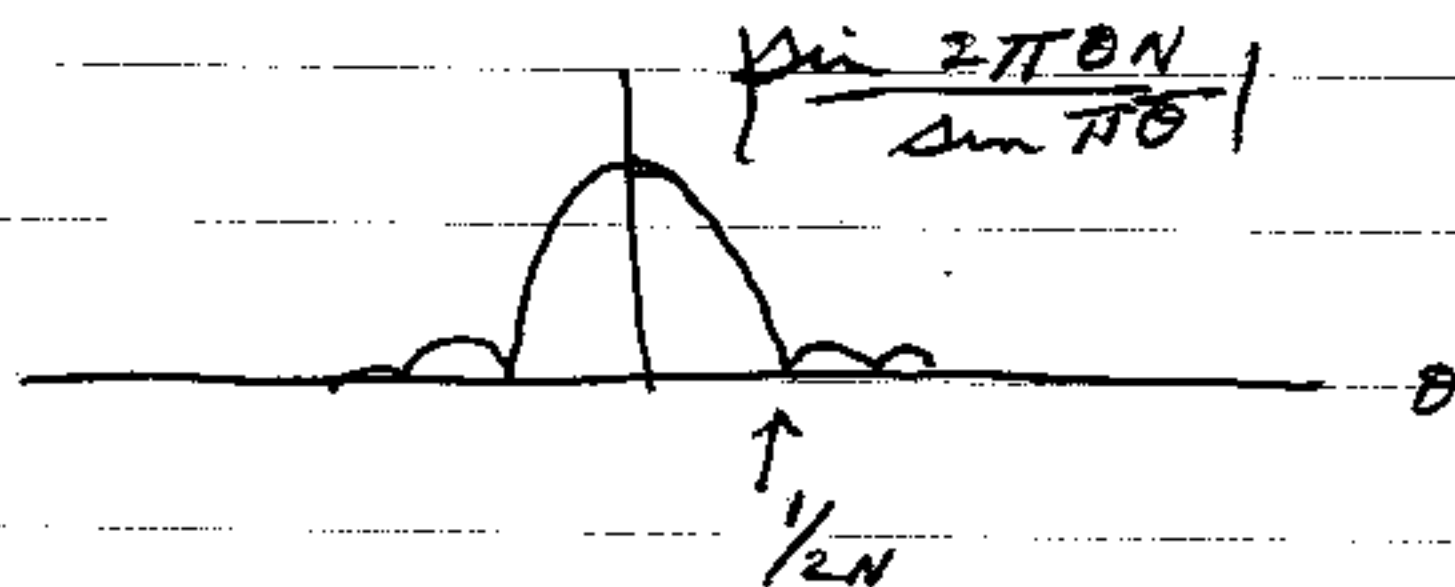
$$\text{or } \rho_r \geq -1 \Rightarrow |\rho_r| \leq 1$$

$$4.21) \quad \sum_{n=0}^{N-1} s_0[n] s_1[n] = A^2 \sum_n \cos 2\pi f_0 n \cdot \cos 2\pi f_1 n$$

$$= \frac{A^2}{2} \sum_n [\cos 2\pi (f_0 + f_1) n + \cos 2\pi (f_1 - f_0) n]$$

$$\approx \frac{A^2}{2} \left[\frac{1}{2} \frac{\sin 2\pi (f_0 + f_1) N}{\sin \pi (f_0 + f_1)} + \frac{1}{2} \frac{\sin 2\pi (f_1 - f_0) N}{\sin \pi (f_1 - f_0)} \right]$$

For $(f_1 - f_0) \gg 1/2N$ and f_0, f_1 not near 0 or $\frac{1}{2}$, this is ≈ 0



4.22) Must minimize ρ_r .
But from Problem 4.21

$$p_s = \underline{J_1}^T \underline{J_0}$$

$$\underline{\frac{1}{2}(\underline{J_1}^T \underline{J_1} + \underline{J_0}^T \underline{J_0})}$$

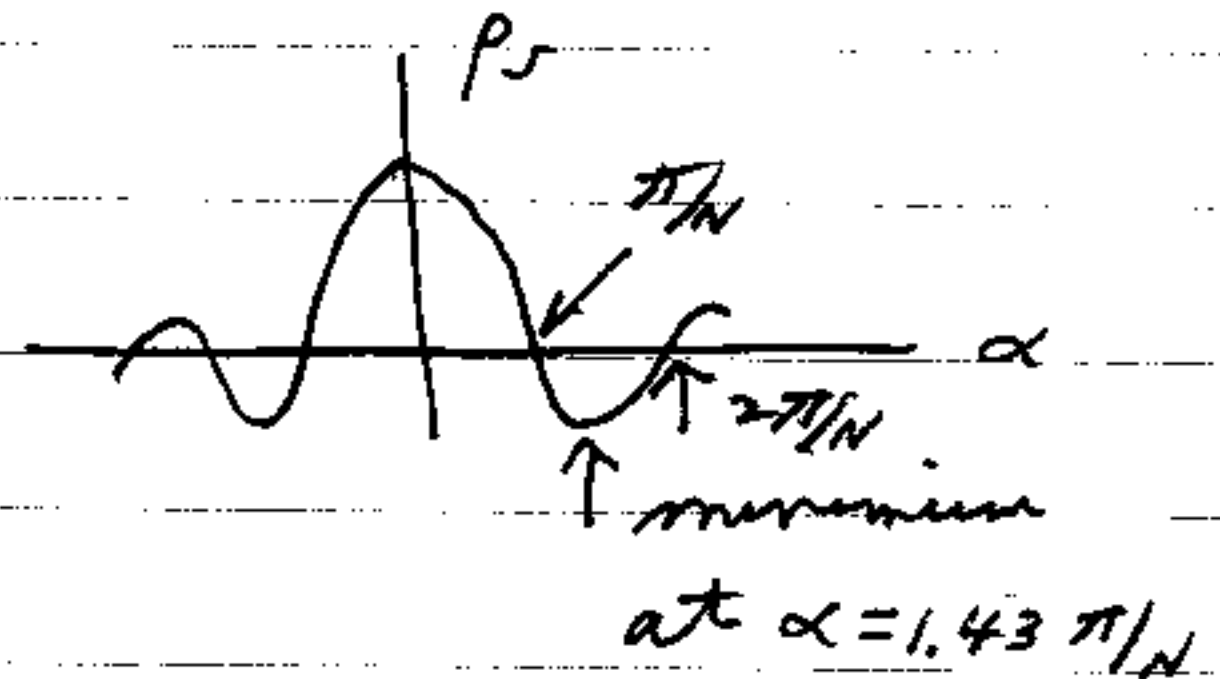
$$\approx \frac{\sin 2\pi(f_1 - f_0)N}{2N \sin \pi(f_1 - f_0)}$$

$$\text{since } \underline{J_1}^T \underline{J_1} \approx \underline{J_0}^T \underline{J_0} \approx NA^2/2.$$

$$\text{Let } \alpha = 2\pi(f_1 - f_0)$$

$$p_s \approx \frac{\sin N\alpha}{2N \sin \alpha/2}$$

$$\approx \frac{\sin N\alpha}{N\alpha}$$



$$\text{or } 2\pi|f_1 - f_0| = 1.4\pi/N$$

$$\Rightarrow |f_1 - f_0| = 0.7/N$$

$$4.23) \quad p_e = Q\left(\sqrt{\frac{\bar{\epsilon}(1-p_s)}{2\sigma^2}}\right)$$

$$\bar{\epsilon} \approx \frac{1}{2}(NA^2/2 + 0) = NA^2/4$$

$$p_s = 0$$

$$\Rightarrow p_e = Q\left(\sqrt{\frac{NA^2}{8\sigma^2}}\right)$$

$$\text{For PSK } P_e = Q\left(\sqrt{\frac{E}{\sigma^2}}\right) \\ = Q\left(\sqrt{\frac{NA^2}{2\sigma^2}}\right)$$

$$\text{For FSK } P_e = Q\left(\sqrt{\frac{E}{2\sigma^2}}\right) \\ = Q\left(\sqrt{\frac{NA^2}{4\sigma^2}}\right)$$

OOK is 6 dB poorer than PSK and 3 dB poorer than FSK, assuming a peak power constraint or the sinusoidal amplitude A is the same for all three systems.

4.24) Use ML rule \Rightarrow use (4.26)

$$T_i(\underline{x}) = \sum_n x(n) s_i[n] - \frac{1}{2} E_i$$

$$\text{with } s_i[n] = A_i$$

$$T_i(\underline{x}) = NA_i \bar{x} - \frac{1}{2} NA_i^2$$

Decide H_k for which $T_i(\underline{x})$ is maximum,

For $M=2$ we have

$$P_e = Q\left(\sqrt{\frac{E(1-\rho_s)}{2\sigma^2}}\right)$$

To minimize P_e must minimize ρ_s

But

$$\rho_s = \frac{\sum_n s_0(n) s_1(n)}{\frac{1}{2} \left(\sum_n s_0^2(n) + \sum_n s_1^2(n) \right)}$$

$$= \frac{N A_0 A_1}{\frac{1}{2} (N A_0^2 + N A_1^2)}$$

$$= \frac{A_0 A_1}{\frac{1}{2} (A_0^2 + A_1^2)}$$

Choose $A_1 = -A_0 \Rightarrow$ antipodal signals

Then, $\rho_s = -1$.

$$4.25) P_e = 1 - \int_{-\infty}^{\infty} \Phi(u) \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(u - \sqrt{E/\sigma^2})^2} du$$

$$= 1 - \int_{-\infty}^{\infty} \int_{-\infty}^u \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(u - \sqrt{E/\sigma^2})^2} du$$

$$= 1 - \int_{-\infty}^{\infty} \int_{-\infty}^u \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(t^2 + (u - \sqrt{E/\sigma^2})^2)} dt du$$

$$\text{Let } v = u - \sqrt{E/2}t$$

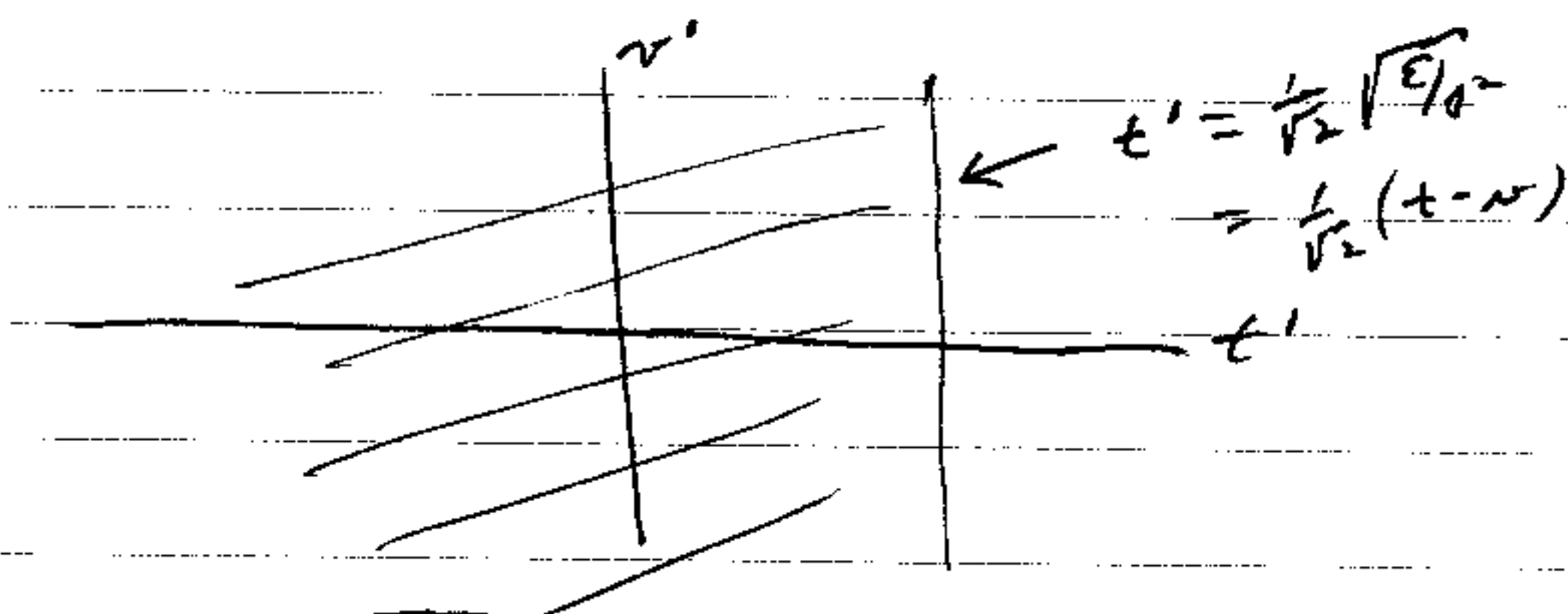
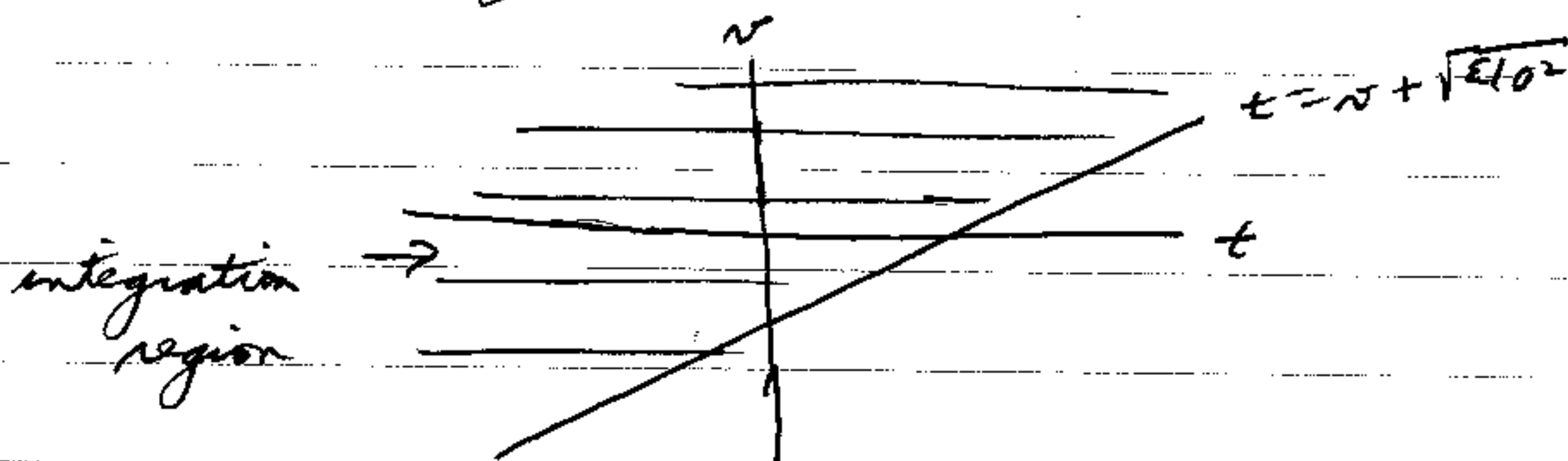
$$P_e = 1 - \int_{-\infty}^{\infty} \int_{-\infty}^{v + \sqrt{E/2}t} \frac{1}{2\pi} e^{-\frac{1}{2}(t^2 + v^2)} dt dv$$

$$\text{Let } \begin{bmatrix} t' \\ v' \end{bmatrix} = \underbrace{\begin{pmatrix} 1/\sqrt{2} & -1/\sqrt{2} \\ 1/\sqrt{2} & 1/\sqrt{2} \end{pmatrix}}_{\underline{A}} \begin{bmatrix} t \\ v \end{bmatrix}$$

$$\Rightarrow t^2 + v^2 = t'^2 + v'^2 \text{ since } \underline{A}^T = \underline{A}^{-1}$$

and the Jacobian has determinant = 1

since the Jacobian is \underline{A}



$$P_e = 1 - \int_{-\infty}^{\infty} \int_{-\infty}^{\sqrt{E/2}t'} \frac{1}{2\pi} e^{-\frac{1}{2}(t'^2 + v'^2)} dt' dv'$$

$$= 1 - \int_{-\infty}^{\sqrt{E/2}t'} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t'^2} dt'$$

$$P_e = Q\left(\sqrt{\frac{\mathcal{E}}{2\sigma^2}}\right)$$

$$4.26) \quad H = \begin{bmatrix} 1 & 0 \\ \cos 2\pi f_0 & \sin 2\pi f_0 \\ \vdots & \vdots \\ \cos 2\pi f_0(N-1) & \sin 2\pi f_0(N-1) \end{bmatrix}$$

To show it has rank 2 we need only show that there is a 2×2 submatrix that has a nonzero determinant.

$$\begin{vmatrix} 1 & 0 \\ \cos 2\pi f_0 & \sin 2\pi f_0 \end{vmatrix} = \sin 2\pi f_0 \neq 0$$

for $0 < f_0 < \frac{1}{2}$

$$4.27) \quad \begin{pmatrix} \hat{a} \\ \hat{b} \end{pmatrix} = \begin{bmatrix} \frac{2}{N} \sum_n x[n] \cos 2\pi f_0 n \\ \frac{2}{N} \sum_n x[n] \sin 2\pi f_0 n \end{bmatrix}$$

$$\begin{aligned} E(\hat{a}) &= \frac{2}{N} \sum_n (a \cos 2\pi f_0 n + b \sin 2\pi f_0 n \cos 2\pi f_0 n) \\ &= \frac{2}{N} \sum_n \left(\frac{a}{2} + \frac{a}{2} \cos 4\pi f_0 n + \frac{b}{2} \sin 4\pi f_0 n \right) \\ &\approx \frac{2}{N} a \frac{N}{2} = a \end{aligned}$$

$$E(\hat{b}) = \frac{2}{N} \sum_n (a \cos 2\pi f_0 n \sin 2\pi f_0 n + b \sin^2 2\pi f_0 n)$$

$$= \frac{2}{N} \sum_n a_2 \sin 4\pi f_0 n + b/2 - b/2 \cos 4\pi f_0 n$$

$$\approx \frac{2}{N} N \frac{b}{2} = b$$

4.28) $\underline{x} = \underline{H}\underline{\theta} + \underline{w}$

where $\underline{H} = \begin{bmatrix} 1 & 0 \\ 1 & 1 \\ \vdots & \vdots \\ 1 & N-1 \end{bmatrix}$ $\underline{\theta} = \begin{bmatrix} A \\ B \end{bmatrix}$

NP detector is to decide H_1 if

$$T(\underline{x}) = \underline{x}^T \underline{C}^{-1} \underline{S} > \gamma'$$

$$T(\underline{x}) = \frac{1}{\sigma^2} \underline{x}^T \underline{H} \underline{\theta}_1$$

$$= \frac{1}{\sigma^2} \sum_{n=0}^{N-1} x(n)(A+Bn)$$

$$P_D = Q(Q^{-1}(P_{FA}) - \sqrt{\underline{S}^T \underline{C}^{-1} \underline{S}})$$

$$\underline{S}^T \underline{C}^{-1} \underline{S} = \frac{1}{\sigma^2} \underline{S}^T \underline{S} = \frac{\underline{S}^T \underline{S}}{\sigma^2} = \frac{\underline{\theta}_1^T \underline{H}^T \underline{H} \underline{\theta}_1}{\sigma^2}$$

or

$$\underline{S}^T \underline{C}^{-1} \underline{S} = \frac{\sum_{n=0}^{N-1} (A+Bn)^2}{\sigma^2}$$

4.29) Since $L = \log_2 M$ we have

$L = \ln M / \ln 2$ and the first step follows by letting $v = u - \sqrt{\frac{LFT}{\Delta 0^2}}$

Now let $\alpha = \frac{PT}{\Delta 0^2}$ and define

$$g(x) = \ln \left[\Phi \left(v + \sqrt{\frac{\alpha \ln x}{\ln 2}} \right)^{x-1} \right]$$

want $\lim_{x \rightarrow 0} g(x)$, or $\lim_{y \rightarrow 0} g(1/y)$

$$\lim_{y \rightarrow 0} g(1/y) = \lim_{y \rightarrow 0} \ln \Phi \left(v + \sqrt{\frac{\alpha \ln(1/y)}{\ln 2}} \right)$$

$$= \lim_{y \rightarrow 0} \ln \Phi \left(\sqrt{-\frac{\alpha}{\ln 2} \ln y} \right)$$

Using L'Hospital's rule

$$= \lim_{y \rightarrow 0} \frac{\Phi' \left(\sqrt{-\frac{\alpha}{\ln 2} \ln y} \right) \cdot \frac{1}{2} \left(-\frac{\alpha}{\ln 2} \ln y \right)^{-1/2} \left(-\frac{\alpha}{y \ln 2} \right)}{\Phi \left(\sqrt{-\frac{\alpha}{\ln 2} \ln y} \right)}$$

$$= \lim_{y \rightarrow 0} \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \left(-\frac{\alpha}{\ln 2} \ln y \right)} \cdot \frac{-\alpha}{2y \ln 2}}{\sqrt{-\frac{\alpha}{\ln 2} \ln y}}$$

$$\text{since } \Phi'(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}$$

$$\text{and } \Phi(\infty) = 1$$

$$= \lim_{y \rightarrow 0} \frac{1}{\sqrt{2\pi}} e^{\ln y \frac{\alpha}{2 \ln 2}} = \frac{e^{\frac{\alpha}{2 \ln 2}}}{\sqrt{\frac{\alpha}{\ln 2}}} \frac{1}{y \sqrt{\ln 1/y}}$$

$$= \lim_{y \rightarrow 0} C \frac{y^{\frac{\alpha}{2 \ln 2}}}{y \sqrt{\ln 1/y}} \quad \text{where } C < 0$$

$$= C \lim_{y \rightarrow 0} \frac{y^{\frac{\alpha}{2 \ln 2} - 1}}{\sqrt{\ln 1/y}}$$

Now if $y \rightarrow 0$ and $\beta = \frac{\alpha}{2 \ln 2} - 1 > 0$, we have

$$C \lim_{y \rightarrow 0} \frac{y^\beta}{\sqrt{\ln 1/y}} = 0$$

or if $\frac{\alpha}{2 \ln 2} - 1 < 0$, we have

$$C \lim_{y \rightarrow 0} \frac{1}{y^\beta \sqrt{\ln 1/y}} = -\infty \quad \text{since } C < 0.$$

$$\Rightarrow \lim = -\infty \quad \text{if } \alpha < 2 \ln 2$$

$$0 \quad \text{if } \alpha > 2 \ln 2$$

$$4.30) \det(\underline{I}_{N-1} - \frac{1}{N} \underline{1}_{N-1} \underline{1}_{N-1}^T) =$$

$$\det(1 - \frac{1}{N} \underline{1}_{N-1} \underline{1}_{N-1}^T)$$

Since $M=1$

$$= 1 - \frac{1}{N} (N-1) = \frac{1}{N} \neq 0$$

$$\begin{aligned}
 4.31) \quad (\underline{x} - \underline{H}\underline{\theta})^T \underline{C}^{-1} (\underline{x} - \underline{H}\underline{\theta}) &= \\
 &= [\underline{x} - \underline{H}\hat{\underline{\theta}} - (\underline{H}\underline{\theta} - \underline{H}\hat{\underline{\theta}})]^T \underline{C}^{-1} [\underline{x} - \underline{H}\hat{\underline{\theta}} - (\underline{H}\underline{\theta} - \underline{H}\hat{\underline{\theta}})] \\
 &= [(\underline{x} - \underline{H}\hat{\underline{\theta}}) - \underline{H}(\underline{\theta} - \hat{\underline{\theta}})]^T \underline{C}^{-1} [(\underline{x} - \underline{H}\hat{\underline{\theta}}) - \underline{H}(\underline{\theta} - \hat{\underline{\theta}})] \\
 &= (\underline{x} - \underline{H}\hat{\underline{\theta}})^T \underline{C}^{-1} (\underline{x} - \underline{H}\hat{\underline{\theta}}) + (\underline{\theta} - \hat{\underline{\theta}})^T \underline{H}^T \underline{C}^{-1} \underline{H} (\underline{\theta} - \hat{\underline{\theta}}) \\
 &\quad - 2 (\underline{\theta} - \hat{\underline{\theta}})^T \underline{H}^T \underline{C}^{-1} (\underline{x} - \underline{H}\hat{\underline{\theta}})
 \end{aligned}$$

$$\text{But } \underline{H}^T \underline{C}^{-1} (\underline{x} - \underline{H}\hat{\underline{\theta}}) = \underline{H}^T \underline{C}^{-1} \underline{x} - (\underline{H}^T \underline{C}^{-1} \underline{H}) \hat{\underline{\theta}} = \underline{0}$$

Now

$$p(\underline{x}; \underline{\theta}) = \frac{1}{(2\pi)^{N/2} \det^{1/2}(\underline{C})}$$

$$\underbrace{e^{-\frac{1}{2} (\underline{x} - \underline{H}\hat{\underline{\theta}})^T \underline{C}^{-1} (\underline{x} - \underline{H}\hat{\underline{\theta}})}}_{h(\underline{x})} \underbrace{e^{-\frac{1}{2} (\underline{\theta} - \hat{\underline{\theta}})^T \underline{H}^T \underline{C}^{-1} \underline{H} (\underline{\theta} - \hat{\underline{\theta}})}}_{g(\underline{\tau}(\underline{x}), \underline{\theta})}$$

Chapter 5

5.1) For $N=6$ and $P_{FA} = 10^{-3}$

$$\gamma_{k+1}'' = -\ln 10^{-3} + \ln \left[1 + \gamma_k'' + \frac{1}{2} \gamma_k''^2 \right]$$

k	γ_k''
0	1
1	7.8240
2	10.5823
3	11.1210
4	11.2113
5	11.2260
6	11.2284
7	11.2289
8	11.2289

$$\Rightarrow \gamma' = 2\sigma^2 (11.2289) = 22.4578 \sigma^2$$

5.2) Let $X = \sum_{i=1}^N X_i^2$

$$X \stackrel{a}{\sim} N(N E(X_i^2), N \text{var}(X_i^2))$$

$$\text{But } E(X_i^2) = E(X_i^2) = 1$$

$$\text{var}(X_i^2) = \text{var}(X_i^2) = 2$$

(see Chapter 2) $\Rightarrow X \stackrel{a}{\sim} N(N, 2N)$

or

$$\frac{X-N}{\sqrt{2N}} \stackrel{a}{\sim} N(0, 1)$$

$$\begin{aligned} Q_{X_N^2}(x) &= P_r \{ X_N^2 > x \} \\ &= P_r \left\{ \frac{X_N^2 - N}{\sqrt{2N}} > \frac{x-N}{\sqrt{2N}} \right\} \\ &\approx P_r \left\{ N(0, 1) > \frac{x-N}{\sqrt{2N}} \right\} \\ &= Q \left(\frac{x-N}{\sqrt{2N}} \right) \end{aligned}$$

Now

$$P_D = Q_{X_N^2} \left(\frac{\delta'/\sigma^2}{\sigma_s^2/\sigma^2 + 1} \right)$$

$$\approx Q \left[\frac{\frac{\delta'/\sigma^2}{\sigma_s^2/\sigma^2 + 1} - N}{\sqrt{2N}} \right]$$

$$P_{FA} = Q_{X_N^2} \left(\frac{\delta'}{\sigma^2} \right)$$

$$\approx Q \left(\frac{\delta'/\sigma^2 - N}{\sqrt{2N}} \right)$$

$$\Rightarrow \delta'/\sigma^2 = N + \sqrt{2N} Q^{-1}(P_{FA})$$

$$P_D = Q \left[\frac{\delta'/\sigma^2 - N\sigma_s^2/\sigma^2 - N}{(\sigma_s^2/\sigma^2 + 1)\sqrt{2N}} \right]$$

$$= Q \left(\frac{N + \sqrt{2N} Q^{-1}(PFA) - N\sigma_s^2/\sigma^2 - N}{(\sigma_s^2/\sigma^2 + 1) \sqrt{2N}} \right)$$

$$= Q \left(\frac{Q^{-1}(PFA) - \sqrt{N/2} \sigma_s^2/\sigma^2}{\sigma_s^2/\sigma^2 + 1} \right)$$

5.3) From (5.5) decide H_1 if

$$T(\underline{x}) = \underline{x}^T \underline{C}_J (\underline{C}_J + \sigma^2 \underline{I})^{-1} \underline{x} > \gamma''$$

$$\hat{\underline{J}} = \underline{C}_J (\underline{C}_J + \sigma^2 \underline{I})^{-1} \underline{x}$$

$$= \begin{bmatrix} \frac{\sigma_{s_0}^2}{\sigma_{s_0}^2 + \sigma^2} & & 0 \\ & \ddots & \\ 0 & & \frac{\sigma_{s_{N-1}}^2}{\sigma_{s_{N-1}}^2 + \sigma^2} \end{bmatrix} \underline{x}$$

$$T(\underline{x}) = \sum_{n=0}^{N-1} \frac{\sigma_{s_n}^2}{\sigma_{s_n}^2 + \sigma^2} x^2[n]$$

5.4) $\hat{\underline{J}} = \underline{C}_J (\underline{C}_J + \sigma^2 \underline{I})^{-1} \underline{x}$

But $\underline{J} = A \underline{1}$

$$\underline{C}_J = E(\underline{J} \underline{J}^T) = E(A^2 \underline{1} \underline{1}^T) \\ = \sigma_A^2 \underline{1} \underline{1}^T$$

By Woodbury's identity

$$(\underline{C}_J + \sigma^2 \underline{I})^{-1} = (\sigma^2 \underline{I} + \sigma_A^2 \underline{1} \underline{1}^T)^{-1}$$

$$= \frac{1}{\sigma^2} \underline{I} - \frac{1}{\sigma^4} \frac{\sigma_A^2 \underline{1} \underline{1}^T}{1 + \frac{\sigma_A^2}{\sigma^2} \underline{1} \underline{1}^T}$$

$$\begin{aligned}
 \hat{\underline{J}} &= \sigma_A^2 \underline{11}^T \left(\frac{1}{\sigma^2} \underline{I} - \frac{\sigma_A^2}{\sigma^2} \frac{\underline{11}^T}{\sigma^2 + N\sigma_A^2} \right) \underline{x} \\
 &= \frac{\sigma_A^2}{\sigma^2} N \bar{x} \underline{1} - \frac{N\sigma_A^4 / \sigma^2}{\sigma^2 + N\sigma_A^2} N \bar{x} \underline{1} \\
 &= \frac{N\sigma_A^2(\sigma^2 + N\sigma_A^2) - N\sigma_A^4}{\sigma^2(\sigma^2 + N\sigma_A^2)} \bar{x} \underline{1} \\
 &= \frac{\sigma_A^2}{\sigma_A^2 + \sigma^2/N} \bar{x} \underline{1}
 \end{aligned}$$

$$\begin{aligned}
 J(\underline{x}) &= \underline{x}^T \hat{\underline{J}} = \frac{\sigma_A^2}{\sigma_A^2 + \sigma^2/N} \bar{x} N \bar{x} \\
 &= \frac{N\sigma_A^2}{\sigma_A^2 + \sigma^2/N} (\bar{x})^2
 \end{aligned}$$

$$\begin{aligned}
 5.5) \quad J &= E[(\underline{\alpha}^T(\underline{J} - \hat{\underline{J}}))^2] \\
 &= E[(\underline{\alpha}^T(\underline{J} - \underline{W}\underline{x}))^2] \\
 &= E[(\underline{\alpha}^T \underline{J} - \underline{\alpha}^T(\underline{W}_{opt} + \epsilon \delta \underline{W}) \underline{x})^2] \\
 \frac{\partial J}{\partial \epsilon} &= E[2(\underline{\alpha}^T \underline{J} - \underline{\alpha}^T(\underline{W}_{opt} + \epsilon \delta \underline{W}) \underline{x}) \cdot (-\underline{\alpha}^T \delta \underline{W} \underline{x})]
 \end{aligned}$$

$$\left. \frac{\partial J}{\partial \epsilon} \right|_{\epsilon=0} = 0 \Rightarrow$$

$$E[(\underline{\alpha}^T \underline{J} - \underline{\alpha}^T \underline{W}_{opt} \underline{x})(-\underline{\alpha}^T \delta \underline{W} \underline{x})] = 0$$

$$E[\underline{\alpha}^T(\underline{J} - \underline{W}_{opt} \underline{x})(\underline{x}^T \delta \underline{W}^T \underline{\alpha})] = 0$$

$$\underline{\alpha}^T E[(\underline{J} - \underline{W}_{opt} \underline{x}) \underline{x}^T] \delta \underline{W}^T \underline{\alpha} = 0$$

Now since we must have

$$E[(\underline{y} - \underline{W}_{opt} \underline{x}) \underline{x}^T] = \underline{0}$$

$$E(\underline{y} \underline{x}^T) - \underline{W}_{opt} E(\underline{x} \underline{x}^T) = \underline{0}$$

or

$$\begin{aligned} \underline{W}_{opt} &= E(\underline{y} \underline{x}^T) E(\underline{x} \underline{x}^T)^{-1} \\ &= E(\underline{y} (\underline{y} + \underline{w})^T) E((\underline{y} + \underline{w})(\underline{y} + \underline{w})^T)^{-1} \\ &= \underline{C}_y (\underline{C}_y + \sigma^2 \underline{I})^{-1} \end{aligned}$$

$$5.6) \quad \underline{V}_0 = \frac{1}{\sqrt{N}} \underline{1}, \quad \underline{V}_1 = \frac{1}{\sqrt{N}} \underline{1}$$

$$\lambda_{s_0} = N\sigma_s^2/2 \quad \lambda_1 = N\sigma_s^2/2$$

From (5.10), (5.11) with

$$\alpha_0 = \frac{N\sigma_s^2/2 \cdot \sigma^2}{N\sigma_s^2/2 + \sigma^2} = \alpha_1 = \alpha$$

$$\begin{aligned} P_{FA} &= \int_{\gamma''}^{\infty} \int_{-\infty}^{\infty} \underbrace{\left(\frac{1}{\sqrt{1-2\gamma\alpha u}} \right)^2}_{\frac{1}{2\alpha} e^{-t/2\alpha} \quad t > 0} e^{-j\omega t} \frac{d\omega}{2\pi} dt \\ &\quad \text{from (5.13)} \end{aligned}$$

$$= \int_{\gamma''}^{\infty} \frac{1}{2\alpha} e^{-t/2\alpha} dt = e^{-\gamma''/2\alpha}$$

also,

$$P_D = e^{-\gamma''/2\lambda_s}$$

$$\text{Now } \gamma'' = 2\alpha \ln(1/P_{FA})$$

$$\ln P_D = -\frac{\sigma^2}{2\lambda s} = -\frac{2\alpha \ln 1/P_{FA}}{2\lambda s}$$

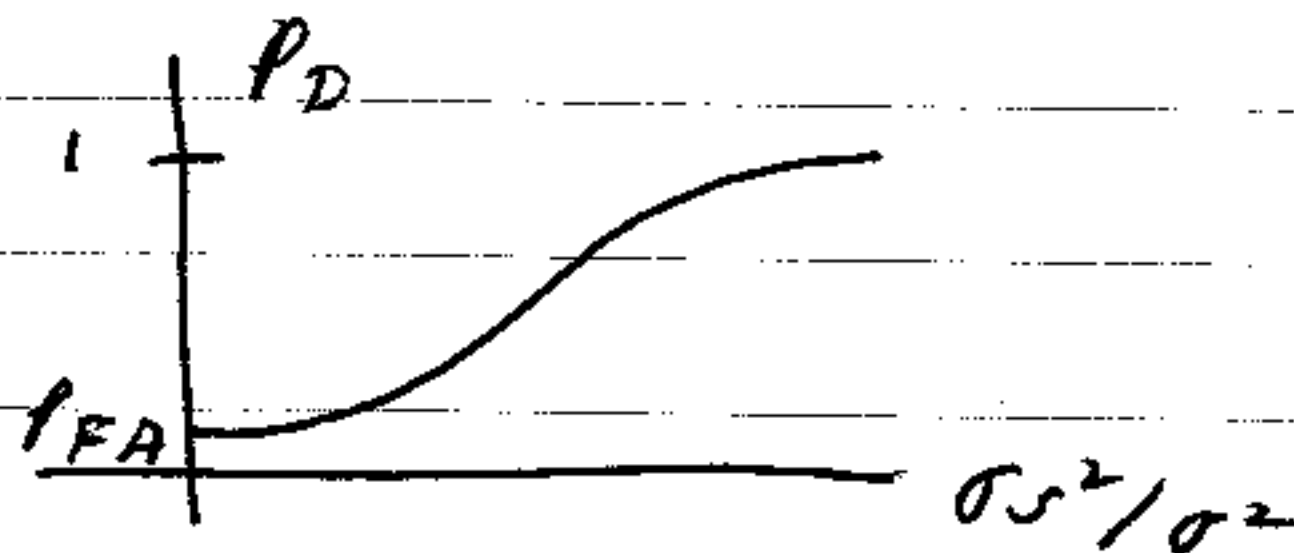
$$= \frac{\alpha}{\lambda s} \ln P_{FA}$$

$$P_D = P_{FA}^{\alpha/\lambda s}$$

$$\frac{\alpha}{\lambda s} = \frac{\sigma^2}{N\sigma_s^2/2 + \sigma^2} = \frac{1}{1 + \frac{N}{2} \sigma_s^2/\sigma^2}$$

$$P_D = P_{FA}^{\frac{1}{1 + \frac{N}{2} \sigma_s^2/\sigma^2}}$$

$$= (0.01)^{\frac{1}{1 + 5\sigma_s^2/\sigma^2}}$$



5.7) as $\rho \rightarrow 1$ $T(\underline{x}) \rightarrow \frac{2\sigma_s^2}{2\sigma_s^2 + \sigma^2} y^2(0)$

$\Rightarrow y(1)$ is discarded

But $y(1) = (\underline{V}^T \underline{x})_1 = \begin{bmatrix} 1/\sqrt{2} & 1/\sqrt{2} \\ 1/\sqrt{2} & -1/\sqrt{2} \end{bmatrix} \begin{bmatrix} x(0) \\ x(1) \end{bmatrix}$

$$= \frac{1}{\sqrt{2}} (x(0) - x(1))$$

If $p \rightarrow 1$, then $S(0) \rightarrow S(1)$ since the PDF of $(S(0) S(1))^T$ is concentrated along $S(1) = S(0)$. Thus $y(1) = \frac{1}{\sqrt{2}} (S(0) + W(0) - S(1) - W(1)) \rightarrow \frac{1}{\sqrt{2}} (W(0) - W(1))$ under H_1 .

Thus $y(1)$ provides no discrimination.

5.8) $\underline{x}' = \underline{W} \underline{x} = \underline{C}_S (\underline{C}_S + \sigma^2 \underline{I})^{-1} \underline{x}$
 But $\underline{V}^T \underline{C}_S \underline{V} = \underline{\Lambda}_S \Rightarrow \underline{C}_S = \underline{V} \underline{\Lambda}_S \underline{V}^T$

$$\begin{aligned} \underline{x}' &= \underline{V} \underline{\Lambda}_S \underline{V}^T (\underline{V} \underline{\Lambda}_S \underline{V}^T + \sigma^2 \underline{I})^{-1} \underline{x} \\ &= \underline{V} \underline{\Lambda}_S \underline{V}^T (\underline{V} (\underline{\Lambda}_S + \sigma^2 \underline{I}) \underline{V}^T)^{-1} \underline{V} \underline{V}^T \underline{x} \\ &= \underline{V} \underline{\Lambda}_S \underline{V}^T \underline{V}^T \underline{V}^T (\underline{\Lambda}_S + \sigma^2 \underline{I})^{-1} \underline{V}^T \underline{V} \underline{V}^T \underline{x} \\ &= \underline{V} \underline{\Lambda}_S (\underline{\Lambda}_S + \sigma^2 \underline{I})^{-1} \underbrace{\underline{V}^T \underline{x}}_{\underline{y}} \end{aligned}$$

$$= \underline{V} \begin{bmatrix} \frac{\lambda_{S0}}{\lambda_{S0} + \sigma^2} & & \\ & \ddots & \\ & & \frac{\lambda_{S(N-1)}}{\lambda_{S(N-1)} + \sigma^2} \end{bmatrix} \underline{y}$$

$$= \underline{V} \underbrace{\begin{bmatrix} \frac{\lambda_{S0}}{\lambda_{S0} + \sigma^2} y(0) \\ \vdots \\ \frac{\lambda_{S(N-1)}}{\lambda_{S(N-1)} + \sigma^2} y(N-1) \end{bmatrix}}_{\underline{y}'}$$

5.9) $N=4$ Let $M=2$ in Ex 5.4

$$\Rightarrow P_{FA} = A_0 e^{-\gamma''/2\alpha_0} + A_1 e^{-\gamma''/2\alpha_1}$$

$$A_0 = \frac{1}{1 - \alpha_1/\alpha_0} \quad A_1 = \frac{1}{1 - \alpha_0/\alpha_1} = 1 - A_0$$

$$\alpha_0 = \frac{\lambda_{s0} \sigma^2}{\lambda_{s0} + \sigma^2} \quad \alpha_1 = \frac{\lambda_{s1} \sigma^2}{\lambda_{s1} + \sigma^2}$$

$$\lambda_{s0} = 2, \quad \lambda_{s1} = 1, \quad \sigma^2 = 1$$

$$\Rightarrow \alpha_0 = \frac{2}{3} \quad \alpha_1 = \frac{1}{2}$$

$$A_0 = \frac{1}{1 - \frac{1/2}{2/3}} = 4$$

$$A_1 = -3$$

$$P_{FA} = 4 e^{-\frac{3}{4}\gamma''} - 3 e^{-\gamma''}$$

$$P_D = B_0 e^{-\gamma''/2\lambda_{s0}} + B_1 e^{-\gamma''/2\lambda_{s1}}$$

$$B_0 = \frac{1}{1 - \lambda_{s1}/\lambda_{s0}} \quad B_1 = \frac{1}{1 - \lambda_{s0}/\lambda_{s1}} = 1 - B_0$$

$$B_0 = \frac{1}{1 - 1/2} = 2 \quad B_1 = -1$$

$$P_D = 2 e^{-\frac{1}{4}\gamma''} - e^{-\frac{1}{2}\gamma''}$$

5.10) Decide H_1 if

$$L(\underline{x}) = \frac{p(\underline{x}; H_1)}{p(\underline{x}; H_0)} > \gamma$$

$$L(\underline{x}) = \frac{\frac{1}{(2\pi)^{N/2} \det^{1/2}(\underline{C}_S + \underline{C}_W)} e^{-\frac{1}{2} \underline{x}^T (\underline{C}_S + \underline{C}_W)^{-1} \underline{x}}}{\frac{1}{(2\pi)^{N/2} \det^{1/2}(\underline{C}_W)} e^{-\frac{1}{2} \underline{x}^T \underline{C}_W^{-1} \underline{x}}}$$

or decide H_1 if

$$-\frac{1}{2} \underline{x}^T (\underline{C}_S + \underline{C}_W)^{-1} \underline{x} + \frac{1}{2} \underline{x}^T \underline{C}_W^{-1} \underline{x} > \delta'$$

$$T(\underline{x}) = \underline{x}^T [\underline{C}_W^{-1} - (\underline{C}_S + \underline{C}_W)^{-1}] \underline{x} > 2\delta'$$

From matrix inversion lemma

$$(\underline{C}_S + \underline{C}_W)^{-1} = \underline{C}_W^{-1} - \underline{C}_W^{-1} \underline{C}_S (\underline{C}_W^{-1} \underline{C}_S + \underline{I})^{-1} \underline{C}_W^{-1}$$

$$T(\underline{x}) = \underline{x}^T \underline{C}_W^{-1} \underbrace{[\underline{C}_S (\underline{C}_W^{-1} \underline{C}_S + \underline{I})^{-1} \underline{C}_W^{-1}]}_{\hat{\underline{J}}_S} \underline{x}$$

$$\hat{\underline{J}}_S = [\underline{C}_W (\underline{C}_W^{-1} \underline{C}_S + \underline{I}) \underline{C}_S^{-1}]^{-1} \underline{x}$$

$$= [(\underline{C}_S + \underline{C}_W) \underline{C}_S^{-1}]^{-1} \underline{x}$$

$$= \underline{C}_S (\underline{C}_S + \underline{C}_W)^{-1} \underline{x}$$

$$5.11) \quad \underline{V}_N^T \underline{C}_N \underline{V}_N = \underline{\Lambda}_N$$

$$= \sqrt{\underline{\Lambda}_N} \sqrt{\underline{\Lambda}_N}$$

$$\underbrace{\sqrt{\underline{\Lambda}_N}^{-1}}_{\underline{A}^T} \underline{V}_N^T \underline{C}_N \underbrace{\underline{V}_N \sqrt{\underline{\Lambda}_N}^{-1}}_{\underline{A}} = \underline{I}$$

$$\Rightarrow \underline{C}_N = \underline{A}^T^{-1} \underline{A}^{-1}$$

$$T(\underline{x}) = \underline{x}^T \underline{C}_W^{-1} \underline{C}_J (\underline{C}_J + \underline{C}_N)^{-1} \underline{x}$$

$$= \underline{x}^T \underline{A} \underline{A}^T \underline{C}_J (\underline{C}_J + \underline{A}^T^{-1} \underline{A}^{-1})^{-1} \underline{x}$$

$$= \underline{x}^T \underline{A} \underline{A}^T \underline{C}_J \underline{A} \underline{A}^{-1} (\underline{A}^T (\underline{A}^T \underline{C}_J \underline{A} + \underline{I}) \underline{A}^{-1})^{-1} \underline{x}$$

$$= \underline{x}^T \underline{A} \underline{B} \underline{A}^{-1} \underline{A} (\underline{B} + \underline{I})^{-1} \underline{A}^T \underline{x}$$

$$= \underline{x}^T \underline{A} \underline{B} (\underline{B} + \underline{I})^{-1} \underline{A}^T \underline{x}$$

$$\text{Let } \underline{y} = \underline{V}_B^T \underline{A}^T \underline{x} \Rightarrow \underline{A}^T \underline{x} = \underline{V}_B \underline{y}$$

$$T(\underline{x}) = \underline{y}^T \underline{V}_B^T \underline{B} \underline{V}_B \underline{V}_B^{-1} (\underline{B} + \underline{I})^{-1} \underline{V}_B \underline{y}$$

$$= \underline{y}^T \underline{\Lambda}_B (\underline{V}_B^{-1} (\underline{B} + \underline{I}) \underline{V}_B)^{-1} \underline{y}$$

$$= \underline{y}^T \underline{\Lambda}_B (\underline{\Lambda}_B + \underline{I})^{-1} \underline{y}$$

$$5.12) \quad T(\underline{x}) = \underline{x}^T (\underline{C}_J (\underline{C}_J + \sigma^2 \underline{I})^{-1}) \underline{x}$$

$$\underline{J} = A \underline{h} \quad \underline{h} = \begin{bmatrix} 1 \\ r \\ \vdots \\ r^{N-1} \end{bmatrix}$$

$$\underline{C}_J = E(A \underline{h} A \underline{h}^T) = \sigma_A^2 \underline{h} \underline{h}^T$$

$$(\underline{C}_J + \sigma^2 \underline{I})^{-1} = (\sigma^2 \underline{I} + \sigma_A^2 \underline{h} \underline{h}^T)^{-1}$$

Using Woodbury's identity

$$= \frac{1}{\sigma^2} \underline{I} - \frac{1}{\sigma^4} \frac{\sigma_A^2 / \sigma^2 \underline{h} \underline{h}^T}{1 + \underline{h}^T \underline{h} \sigma_A^2 / \sigma^2}$$

$$T(\underline{x}) = \underline{x}^T \sigma_A^2 \underline{h} \underline{h}^T \left[\frac{1}{\sigma^2} \underline{I} - \frac{\sigma_A^2 / \sigma^4 \underline{h} \underline{h}^T}{1 + \underline{h}^T \underline{h} \frac{\sigma_A^2}{\sigma^2}} \right] \underline{x}$$

$$= (\underline{h}^T \underline{x})^2 \left[\frac{\sigma_A^2}{\sigma^2} - \frac{\sigma_A^4 / \sigma^4 \underline{h}^T \underline{h}}{1 + \underline{h}^T \underline{h} \sigma_A^2 / \sigma^2} \right]$$

$$\frac{\sigma_A^2}{\sigma^2} - \frac{\sigma_A^4 \underline{h}^T \underline{h}}{\sigma^2 (\sigma^2 + \sigma_A^2 \underline{h}^T \underline{h})}$$

$$= \frac{\sigma_A^2 (\sigma^2 + \sigma_A^2 \underline{h}^T \underline{h}) - \sigma_A^4 \underline{h}^T \underline{h}}{\sigma^2 (\sigma^2 + \sigma_A^2 \underline{h}^T \underline{h})}$$

$$= \frac{\sigma_A^2}{\sigma_A^2 \underline{h}^T \underline{h} + \sigma^2}$$

or we decide H_1 if

$$T'(x) = (\underline{h}^T \underline{x})^2 \\ = \left(\sum_{n=0}^{N-1} x(n) r^n \right)^2 > \gamma''$$

5.13) $\underline{y}(n) = A \underline{h}(n)$

$$\Rightarrow \underline{y} = A \underline{h}$$

$$C_y = E(\underline{y} \underline{y}^T) = E(A^T \underline{h} \underline{h}^T A) \\ = \sigma A^T \underline{h} \underline{h}^T A$$

$$\Rightarrow \underline{u} = \sigma A \underline{h}$$

5.14) See Problem 5.12. To find P_{FA} , P_D

$$T'(x) = (\underline{h}^T \underline{x})^2$$

Under H_0 $\underline{x} \sim N(\underline{0}, \sigma^2 \underline{I})$

$$\Rightarrow \underline{h}^T \underline{x} \sim N(0, \sigma^2 \underline{h}^T \underline{h})$$

$$\frac{(\underline{h}^T \underline{x})^2}{\sigma^2 \underline{h}^T \underline{h}} \sim \chi_1^2$$

$$P_{FA} = P_r \{ T'(x) > \gamma'' ; H_0 \} \\ = P_r \left\{ \underbrace{\frac{T'(x)}{\sigma^2 \underline{h}^T \underline{h}}}_{\chi_1^2} > \frac{\gamma''}{\sigma^2 \underline{h}^T \underline{h}} ; H_0 \right\}$$

But $Q_{\chi^2_1}(x) = 2Q(\sqrt{x})$ (see Chapter 2)

$$\Rightarrow P_{FA} = 2Q\left(\sqrt{\frac{\gamma''}{\sigma^2 \underline{h}^T \underline{h}}}\right)$$

Under H_1 , $\underline{x} \sim N(\underline{0}, \sigma_A^2 \underline{h} \underline{h}^T + \sigma^2 \underline{I})$

$$\Rightarrow \frac{(\underline{h}^T \underline{x})^2}{\sigma_A^2 (\underline{h}^T \underline{h})^2 + \sigma^2 \underline{h}^T \underline{h}} \sim \chi^2_1$$

As before we have

$$P_D = 2Q\left(\sqrt{\frac{\gamma''}{\sigma_A^2 (\underline{h}^T \underline{h})^2 + \sigma^2 \underline{h}^T \underline{h}}}\right)$$

$$5.15) \quad \hat{\underline{\theta}} = \underline{C}_{\underline{\theta} \underline{x}} \underline{C}_{\underline{x} \underline{x}}^{-1} \underline{x}$$

$$\text{where } \underline{x} = \underline{H} \underline{\theta} + \underline{w}$$

$$\begin{aligned} \underline{C}_{\underline{x} \underline{x}} &= E(\underline{x} \underline{x}^T) = E((\underline{H} \underline{\theta} + \underline{w})(\underline{H} \underline{\theta} + \underline{w})^T) \\ &= E(\underline{H} \underline{\theta} \underline{\theta}^T \underline{H}^T) + E(\underline{w} \underline{w}^T) \end{aligned}$$

Since $\underline{\theta}$ and \underline{w} are independent

and zero mean

$$\underline{C}_{\underline{x} \underline{x}} = \underline{H} \underline{C}_{\underline{\theta} \underline{\theta}} \underline{H}^T + \sigma^2 \underline{I}$$

$$\underline{C}_{\underline{\theta} \underline{x}} = E(\underline{\theta} (\underline{H} \underline{\theta} + \underline{w})^T)$$

$$= E(\underline{\theta} \underline{\theta}^T \underline{H}^T) = \underline{C}_{\underline{\theta} \underline{\theta}} \underline{H}^T$$

$$\Rightarrow \hat{\underline{\theta}} = \underline{C}_{\underline{\theta} \underline{H}^T} (\underline{H} \underline{C}_{\underline{\theta} \underline{H}^T} + \sigma^2 \underline{I})^{-1} \underline{x}$$

$$5.16) \quad T(\underline{x}) = \frac{c}{N} \underline{x}^T \underline{H} \underline{H}^T \underline{x}$$

$$= \underline{x}^T \underline{\hat{S}} \quad \text{where } \underline{\hat{S}} = \frac{c}{N} \underline{H} \underline{H}^T \underline{x}$$

$$\underline{H}^T \underline{x} = \begin{bmatrix} \sum_n x(n) \cos 2\pi f_0 n \\ \sum_n x(n) \sin 2\pi f_0 n \end{bmatrix}$$

$$\underline{\hat{S}} = \frac{c}{N} \begin{bmatrix} 1 & 0 \\ \cos 2\pi f_0 & \sin 2\pi f_0 \\ \vdots & \vdots \\ \cos 2\pi f_0 (N-1) & \sin 2\pi f_0 (N-1) \end{bmatrix} \begin{bmatrix} \sum x(n) \cos 2\pi f_0 n \\ \sum x(n) \sin 2\pi f_0 n \end{bmatrix}$$

$$\Rightarrow \hat{S}(n) = \frac{c}{N} \left[\cos 2\pi f_0 n \left(\sum x(n) \cos 2\pi f_0 n \right) + \sin 2\pi f_0 n \left(\sum x(n) \sin 2\pi f_0 n \right) \right]$$

$$= \hat{a} \cos 2\pi f_0 n + \hat{b} \sin 2\pi f_0 n$$

$$\text{where } \hat{a} = \frac{\sigma_s^2}{\frac{N\sigma_s^2}{2} + \sigma^2} \sum_{n=0}^{N-1} x(n) \cos 2\pi f_0 n$$

$$\hat{b} = \frac{\sigma_s^2}{\frac{N\sigma_s^2}{2} + \sigma^2} \sum_{n=0}^{N-1} x(n) \sin 2\pi f_0 n$$

$$\text{or } \hat{a} = \frac{N\sigma_s^2/2}{N\sigma_s^2/2 + \sigma^2} \approx \frac{1}{N} \sum_{n=0}^{N-1} x(n) \cos 2\pi f_0 n$$

$$\hat{b} = \frac{N\sigma_s^2/2}{N\sigma_s^2/2 + \sigma^2} \approx \frac{1}{N} \sum_{n=0}^{N-1} x(n) \sin 2\pi f_0 n$$

$$5.17) \quad E(T; \mathcal{H}_0) = 0$$

$$E(T; \mathcal{H}_1) = \sum_n \sin(n) A \cos 2\pi f_0 n$$

$$= A^2 \sum_n \cos(2\pi f_0 n + \phi) \cos 2\pi f_0 n$$

$$= \frac{A^2}{2} \sum_n \cos \phi + \cos(4\pi f_0 n + \phi)$$

$$\approx \frac{NA^2}{2} \cos \phi$$

$$\text{var}(T; \mathcal{H}_0) = E\left(\left(\sum_n W(n) A \cos 2\pi f_0 n\right)^2\right)$$

$$= A^2 \sum_m \sum_n \underbrace{E(W(m)W(n))}_{\delta^2 \delta_{m-n}} \cos 2\pi f_0 m \cos 2\pi f_0 n$$

$$= A^2 \sum_n \delta^2 \cos^2 2\pi f_0 n$$

$$= \frac{A^2 \delta^2}{2} \sum_n (1 + \cos 4\pi f_0 n)$$

$$\approx \frac{NA^2 \delta^2}{2} = \text{var}(T; \mathcal{H}_1)$$

$$d^2 = \frac{(NA^2/2 \cos \phi)^2}{NA^2 \delta^2/2} = \frac{NA^2}{2\delta^2} \cos^2 \phi$$

$$\text{If } \phi = 0, \text{ then } d^2 = NA^2/2\delta^2$$

$$\text{If } \phi = 90^\circ, \text{ then } d^2 = 0 \Rightarrow \text{poor performance (} P_0 = P_{FA} \text{)}.$$

$$5.18) \quad T'(x) = \frac{1}{N} (I^2 + Q^2)$$

First note that I, Q are jointly Gaussian, being linear transformations of x . Also, since $S(n)$ is zero mean, all the means are zero.

To find the variances:

Under H_0

$$\begin{aligned} \text{var}(I) &= E(I^2) \\ &= E\left(\sum_m \sum_n W(m)W(n) \cos 2\pi f_0 m \cos 2\pi f_0 n\right) \\ &= \sum_m \sum_n E(W(m)W(n)) \cos 2\pi f_0 m \cos 2\pi f_0 n \\ &= \sigma^2 \sum_n \cos^2 2\pi f_0 n \\ &= \frac{\sigma^2}{2} \sum_n (1 + \cos 4\pi f_0 n) \\ &\sim N \sigma^2 / 2 \end{aligned}$$

and similarly for $\text{var}(Q)$. Under H_1 ,

$$\begin{aligned} I &= \sum_n (S(n) + W(n)) \cos 2\pi f_0 n \\ &= \sum_n (a \cos 2\pi f_0 n + b \sin 2\pi f_0 n + W(n)) \cdot \cos 2\pi f_0 n \\ &= a \sum_n \cos^2 2\pi f_0 n + b \sum_n \sin 2\pi f_0 n \cos 2\pi f_0 n \\ &\quad + \sum_n W(n) \cos 2\pi f_0 n \end{aligned}$$

$$\approx aN/2 + \sum_n W(n) \cos 2\pi f_0 n$$

since $\sum_n \sin 2\pi f_0 n \cos 2\pi f_0 n \approx 0$

$$\begin{aligned} \text{var}(I) &= (N/2)^2 \text{var}(a) + N\sigma^2/2 \\ &= \frac{N^2\sigma^2}{4} + N\sigma^2/2 \end{aligned}$$

And similarly for Q . To show that I and Q are uncorrelated under H_1 ,

$$I \approx aN/2 + \sum_n W(n) \cos 2\pi f_0 n$$

$$Q \approx bN/2 + \sum_n W(n) \sin 2\pi f_0 n$$

$$E(IQ) = (N/2)^2 \underbrace{E(ab)}_{=0} + E\left(\sum_n W(n) \cos 2\pi f_0 n \sum_n W(n) \sin 2\pi f_0 n\right)$$

$$= \sum_m \sum_n \underbrace{E(W(m)W(n))}_{\sigma^2 \delta(m-n)} \cos 2\pi f_0 m \sin 2\pi f_0 n$$

$$= \sigma^2 \sum_n \underbrace{\cos 2\pi f_0 n \sin 2\pi f_0 n}_{\frac{1}{2} \sin 4\pi f_0 n} \approx 0$$

And similarly under H_0 .

Now we decide H_1 if $T'(x) > \gamma'''$

where $T'(x) = 1/N(I^2 + Q^2)$

Under H_0

$$\frac{I^2 + Q^2}{N\sigma^2/2} \sim \chi^2_2$$

Under H_1

$$\frac{I^2 + Q^2}{\frac{N\sigma^2}{2} + \frac{N^2\sigma_s^2}{4}} \sim \chi^2_2$$

$$P_{FA} = P_r \left\{ \frac{I^2 + Q^2}{N} > \gamma'''; H_0 \right\}$$

$$= P_r \left\{ \chi^2_2 > \frac{\gamma'''}{\sigma^2/2} \right\} = e^{-1/2 \left(\frac{\gamma'''}{\sigma^2/2} \right)} = e^{-\gamma'''/\sigma^2}$$

$$\text{Similarly, } P_D = e^{-\frac{1}{2} \left(\frac{\gamma'''}{\sigma^2/2 + N\sigma_s^2/4} \right)}$$

$$= e^{-\frac{\gamma'''}{N\sigma_s^2/2 + \sigma^2}}$$

5.19) Let $\underline{c}_i = [1 \cos 2\pi f_i \dots \cos 2\pi f_i(N-1)]^T$

$$\underline{s}_i = [0 \sin 2\pi f_i \dots \sin 2\pi f_i(N-1)]^T$$

$i = 0, 1$

$$\underline{W}(f_i) = \begin{bmatrix} \underline{c}_i^T \underline{W} \\ \underline{s}_i^T \underline{W} \end{bmatrix}$$

$$E \left(\underline{W}(f_0) \underline{W}^T(f_1) \right) = E \begin{bmatrix} \underline{c}_0^T \underline{W} \underline{W}^T \underline{c}_1 & \underline{c}_0^T \underline{W} \underline{W}^T \underline{s}_1 \\ \underline{s}_0^T \underline{W} \underline{W}^T \underline{c}_1 & \underline{s}_0^T \underline{W} \underline{W}^T \underline{s}_1 \end{bmatrix}$$

$$= \sigma^2 \begin{bmatrix} \underline{C}_0^T \underline{C}_1 & \underline{C}_0^T \underline{J}_1 \\ \underline{J}_0^T \underline{C}_1 & \underline{J}_0^T \underline{J}_1 \end{bmatrix} \approx \underline{0}$$

for $|f_1 - f_0| \gg 1/N$ since as an example

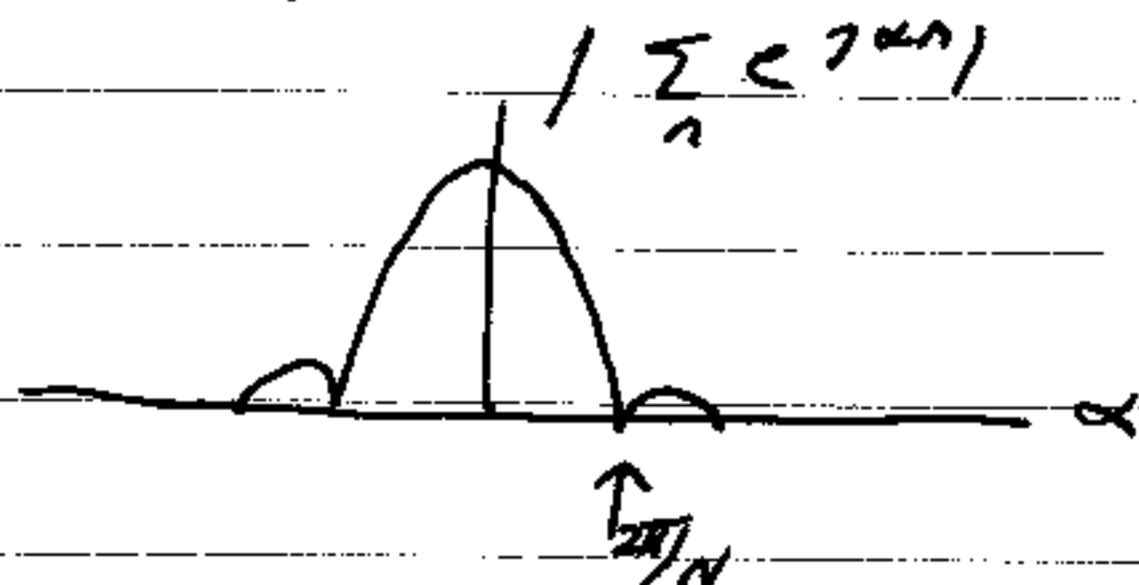
$$\underline{C}_0^T \underline{C}_1 = \sum_n \cos 2\pi f_0 n \cos 2\pi f_1 n$$

$$= \sum_n \left(\frac{1}{2} \cos 2\pi (f_0 + f_1) n \right.$$

$$\left. + \frac{1}{2} \cos 2\pi (f_0 - f_1) n \right)$$

$$= \frac{1}{2} \operatorname{Re} \left(\sum_n e^{j2\pi (f_0 + f_1) n} + e^{j2\pi (f_0 - f_1) n} \right)$$

And $\sum_{n=0}^{N-1} e^{j\alpha n} = e^{j(N-1)\alpha/2} \frac{\sin N\alpha/2}{\sin \alpha/2}$



For $|\alpha| \gg 2\pi/N$ this is ≈ 0 .

Need $f_0 + f_1 \gg 1/N$ and $|f_0 - f_1| \gg 1/N$

$$\text{Now } E(W^*(f_0) W(f_1)) =$$

$$E((\underline{C}_0^T \underline{W} - j \underline{S}_0^T \underline{W})(\underline{C}_1^T \underline{W} + j \underline{S}_1^T \underline{W}))$$

$$= \underline{C}_0^T E(\underline{W} \underline{W}^T) \underline{C}_1 + j E(\underline{C}_0^T \underline{W} \underline{W}^T \underline{S}_1)$$

$$- j \underline{S}_0^T E(\underline{W} \underline{W}^T) \underline{C}_1 + \underline{S}_0^T E(\underline{W} \underline{W}^T) \underline{S}_1$$

$$= \sigma^2 \left(\underline{C}_0^T \underline{C}_1 + j \underline{C}_0^T \underline{S}_1 - j \underline{S}_0^T \underline{C}_1 + \underline{S}_0^T \underline{S}_1 \right) \approx 0$$

$$5.20) \quad T(\underline{x}) = \sum_{n=0}^{N-1} \frac{\Delta s_n}{\Delta s_n + \sigma^2} y^2(n)$$

$$\underline{y} = \underline{V}^T \underline{x} \Rightarrow y(n) = \underline{V}_n^T \underline{x}$$

$$y(n) = \frac{1}{\sqrt{N}} \sum_{k=0}^{N-1} x(k) e^{j2\pi f_n k} = \frac{1}{\sqrt{N}} x^*(f_n)$$

also, $\Delta s_n = P_{SS}(f_n)$. Using the hint, $y^2(n) + y^2(N-n) = 2|y(n)|^2 = \frac{2}{N} |x(f_n)|^2$

$$\begin{aligned} T(\underline{x}) &= \sum_{n=0}^{N/2-1} + \sum_{n=N/2}^{N-1} \\ &= \sum_{n=0}^{N/2-1} \frac{P_{SS}(f_n)}{P_{SS}(f_n) + \sigma^2} y^2(n) \\ &\quad + \sum_{n=N/2}^{N-1} \frac{P_{SS}(f_n)}{P_{SS}(f_n) + \sigma^2} y^2(n) \\ &= \sum_{n=0}^{N/2-1} \frac{P_{SS}(f_n)}{P_{SS}(f_n) + \sigma^2} y^2(n) + \sum_{n=1}^{N/2} \frac{P_{SS}(f_{N-n})}{P_{SS}(f_{N-n}) + \sigma^2} y^2(N-n) \\ &\approx \sum_{n=0}^{N/2} \frac{P_{SS}(f_n)}{P_{SS}(f_n) + \sigma^2} (y^2(n) + y^2(N-n)) \end{aligned}$$

$$\rightarrow 2 \int_0^{\frac{1}{2}} \frac{P_{SS}(f)}{P_{SS}(f) + \sigma^2} |x(f)|^2 df$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{P_{SS}(f)}{P_{SS}(f) + \sigma^2} |x(f)|^2 df$$

$$= N \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{P_{SS}(f)}{P_{SS}(f) + \sigma^2} I(f) df$$

$$5.21) \quad T(\underline{x}) \approx N \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{P_{SS}(f)}{\sigma^2} I(f) df$$

= correlation of PSD and
estimated PSD

By Parseval's theorem

$$T(\underline{x}) = N/\sigma^2 \sum_{k=-\infty}^{\infty} \mathcal{F}^{-1}\{P_{SS}(f)\} \mathcal{F}^{-1}\{I(f)\}$$

$$= \frac{N}{\sigma^2} \sum_{k=-\infty}^{\infty} r_{SS}[k] \mathcal{F}^{-1}\{I(f)\}$$

$$= \frac{N}{\sigma^2} \sum_{k=-(N-1)}^{N-1} r_{SS}[k] \hat{r}_{xx}[k]$$

To find $\mathcal{F}^{-1}\{I(f)\}$:

$$\hat{r}_{xx}[k] = \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{1}{N} |X(f)|^2 e^{j2\pi f k} df$$

$$= \frac{1}{N} \int \sum_m x(m) e^{-j2\pi f m} \sum_n x(n) e^{j2\pi f n} e^{j2\pi f k} df$$

$$= \frac{1}{N} \sum_m \sum_n x(m) x(n) \underbrace{\int_{-\frac{1}{2}}^{\frac{1}{2}} e^{j2\pi f (k+n-m)} df}_{1 \text{ for } m=n+k}$$

$$= \frac{1}{N} \sum_{m=-\infty}^{\infty} \sum_{n=-\infty}^{\infty} x(m) x(n) \delta(m-n-k)$$

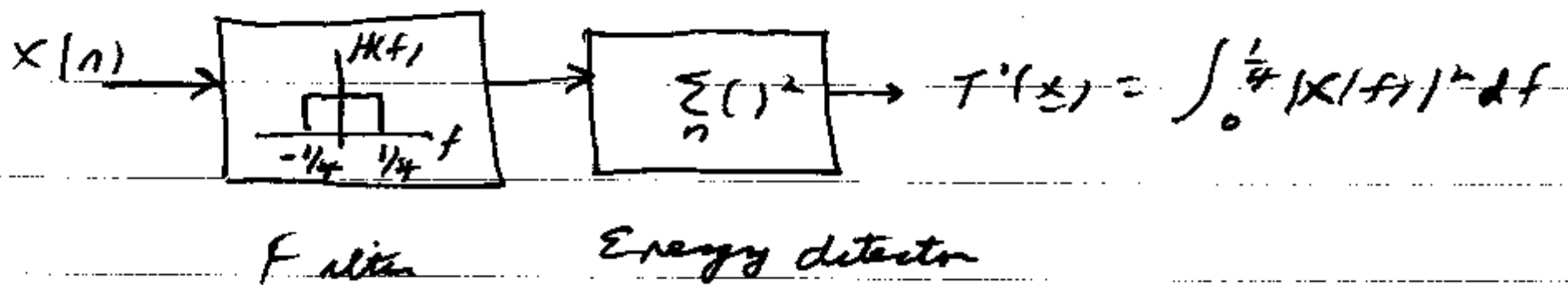
0 otherwise

where $x[n] = 0$ for $n < 0$, $n > N-1$

$$\begin{aligned}\hat{r}_{xx}(k) &= \frac{1}{N} \sum_{n=-\infty}^{\infty} x[n] x[n+k] \\ &= \frac{1}{N} \sum_{n=0}^{N-1-|k|} x[n] x[n+k]\end{aligned}$$

$$5.22) \quad T(\underline{x}) = 2N \int_0^{1/4} \frac{P_0}{P_0 + \sigma^2} I(f) df$$

$$= \frac{2P_0}{P_0 + \sigma^2} \int_0^{1/4} |X(f)|^2 df$$



5.23) From (5.30)

$$\begin{aligned}T'(\underline{x}) &= \frac{1}{2} \underline{x}^T \underline{C}_w^{-1} \underline{C}_s (\underline{C}_s + \underline{C}_w)^{-1} \underline{x} \\ &= \frac{1}{2} \underline{x}^T \underline{C}_w^{-1} \eta \underline{C}_w (\eta \underline{C}_w + \underline{C}_w)^{-1} \underline{x} \\ &= \frac{1}{2} \eta \underline{x}^T \underline{C}_w^{-1} \underline{x} \\ &= \frac{1}{2} \frac{\eta}{\eta+1} \underline{x}^T \underline{C}_w^{-1} \underline{x}\end{aligned}$$

$$\text{or } T''(\underline{x}) = \underline{x}^T \underline{C}_w^{-1} \underline{x}$$

This is just a prewhitener + energy detector. No Wiener filtering is possible.

But under H_0 $\underline{x} \sim N(\underline{0}, \underline{C}_w)$

$$\Rightarrow T''(\underline{x}) \sim \chi_N^2$$

$$P_{FA} = Q_{\chi_N^2}(\delta'')$$

Under H_1 $\underline{x} \sim N(\underline{0}, \underbrace{\underline{C}_s + \underline{C}_w}_{(\eta+1)\underline{C}_w})$

$$\Rightarrow \frac{T''(\underline{x})}{\eta+1} \sim \chi_N^2$$

$$P_D = P_r \{ T''(\underline{x}) > \delta''; H_1 \}$$

$$= P_r \left\{ \frac{T''(\underline{x})}{\eta+1} > \frac{\delta''}{\eta+1}; H_1 \right\}$$

$$= Q_{\chi_N^2}(\delta''/(\eta+1))$$

$$\text{For } N=2 \quad P_{FA} = e^{-\frac{1}{2}\delta''}$$

$$P_D = e^{-\frac{1}{2}\delta''/(\eta+1)}$$

$$P_D = (P_{FA})^{\frac{1}{\eta+1}}$$

$$S.24) \quad S(n) = \sum_{k=0}^{p-1} h(k) u(n-k)$$

$$= \sum_{k=0}^{p-1} h(k) \cos 2\pi f_0(n-k) \quad n \geq p-1$$

$$= \sum_k h(k) \cos 2\pi f_0 n \cos 2\pi f_0 k \\ + \sum_k h(k) \sin 2\pi f_0 n \sin 2\pi f_0 k$$

$$= a \cos 2\pi f_0 n + b \sin 2\pi f_0 n$$

$$\text{Let } \underline{c} = [1 \cos 2\pi f_0 \dots \cos 2\pi f_0 (p-1)]^T$$

$$\underline{s} = [0 \sin 2\pi f_0 \dots \sin 2\pi f_0 (p-1)]^T$$

$$\underline{z} = \begin{pmatrix} a \\ b \end{pmatrix} = \begin{pmatrix} \underline{c}^T \underline{h} \\ \underline{s}^T \underline{h} \end{pmatrix} \quad \text{where } \underline{h} \sim N(\underline{0}, \underline{C}_h)$$

\underline{z} is Gaussian since \underline{h} is Gaussian. Also,

$$E(\underline{z}) = \underline{0}$$

$$\underline{C}_z = E(\underline{z} \underline{z}^T) = E \left(\begin{bmatrix} \underline{c}^T \underline{h} \underline{h}^T \underline{c} & \underline{c}^T \underline{h} \underline{h}^T \underline{s} \\ \underline{s}^T \underline{h} \underline{h}^T \underline{c} & \underline{s}^T \underline{h} \underline{h}^T \underline{s} \end{bmatrix} \right)$$

$$= \begin{bmatrix} \underline{c}^T \underline{C}_h \underline{c} & \underline{c}^T \underline{C}_h \underline{s} \\ \underline{s}^T \underline{C}_h \underline{c} & \underline{s}^T \underline{C}_h \underline{s} \end{bmatrix}$$

$$= \begin{bmatrix} \sum_{k=0}^{p-1} \sigma_k^2 \cos^2 2\pi f_0 k & \sum_{k=0}^{p-1} \sigma_k^2 \cos 2\pi f_0 k \sin 2\pi f_0 k \\ \sum_{k=0}^{p-1} \sigma_k^2 \sin 2\pi f_0 k \cos 2\pi f_0 k & \sum_{k=0}^{p-1} \sigma_k^2 \sin^2 2\pi f_0 k \end{bmatrix}$$

$\frac{1}{2} \sin 4\pi f_0 k$

By limit the cross-diagonal terms ≈ 0

$$\sim \begin{bmatrix} \sum_k \frac{\sigma_k^2}{2} (1 + \cos 4\pi f_0 k) & 0 \\ 0 & \sum_k \frac{\sigma_k^2}{2} (1 - \cos 4\pi f_0 k) \end{bmatrix}$$

Again by the hint

$$\sim \begin{bmatrix} \frac{1}{2} \sum_k \sigma_k^2 & 0 \\ 0 & \frac{1}{2} \sum_k \sigma_k^2 \end{bmatrix}$$

Chapter 6

6.1) From Example 6.1 we decide H_1 if

$$A \sum_n x(n) > \sigma^2 \ln \gamma + NA^2/2$$

If $A < 0$, we decide H_1 if

$$\bar{x} = \frac{1}{N} \sum_{n=0}^{N-1} x(n) < A/2 + \frac{\sigma^2}{NA} \ln \gamma = \gamma'$$

But $\bar{x} \sim N(0, \sigma^2/N)$ under H_0

$$\Rightarrow P_{FA} = Q(\gamma' / \sqrt{\sigma^2/N})$$

$$\text{or } \gamma' = \sqrt{\sigma^2/N} Q^{-1}(P_{FA})$$

which is (6.4)

$$6.2) \frac{p(x(0), x(1); H_1)}{p(x(0), x(1); H_0)} = \frac{\lambda^2 e^{-\lambda(x(0)+x(1))}}{\lambda_0^2 e^{-\lambda_0(x(0)+x(1))}} > \frac{\lambda_0^2 \gamma}{\lambda^2}$$

$$e^{-(\lambda-\lambda_0)(x(0)+x(1))}$$

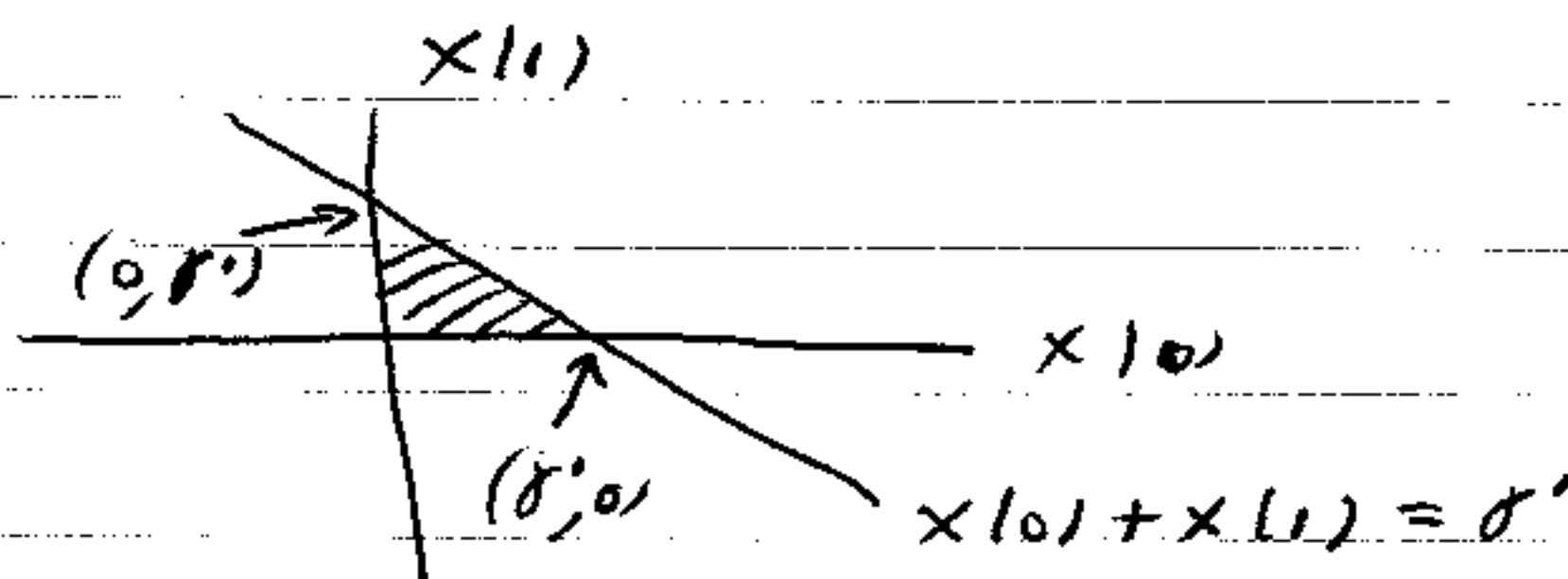
$$> \frac{\lambda_0^2 \gamma}{\lambda^2}$$

$$-(\lambda-\lambda_0)(x(0)+x(1)) > \ln \lambda_0^2 \gamma / \lambda^2$$

Since $\lambda > \lambda_0$ we decide H_1 if

$$T = x(0) + x(1) < \frac{-\ln \lambda_0^2 \gamma / \lambda^2}{(\lambda - \lambda_0)} = \gamma'$$

$$P_{FA} = P_{\lambda} \{x(0) + x(1) < \tau'; H_0\}$$



$$P_{FA} = \int_0^{\tau'} \int_0^{\tau' - x(1)} \lambda_0^2 e^{-\lambda_0(x(0) + x(1))} dx(0) dx(1)$$

$$= \int_0^{\tau'} \lambda_0 e^{-\lambda_0 x(1)} \left[-e^{-\lambda_0 x(0)} \right]_0^{\tau' - x(1)} dx(1)$$

$$= \int_0^{\tau'} (\lambda_0 e^{-\lambda_0 x(1)} - \lambda_0 e^{-\lambda_0 \tau'}) dx(1)$$

$$= -e^{-\lambda_0 x(1)} \Big|_0^{\tau'} - \tau' \lambda_0 e^{-\lambda_0 \tau'}$$

$$= 1 - e^{-\lambda_0 \tau'} - \tau' \lambda_0 e^{-\lambda_0 \tau'}$$

\Rightarrow For given P_{FA} can find threshold without knowledge of λ under H_1 .

6.3) No UMP test since hypothesis is two-sided.

For $\lambda > \lambda_0$ should use $x(0) + x(1) < \tau'$

For $\lambda < \lambda_0$ " " $x(0) + x(1) > \tau'$

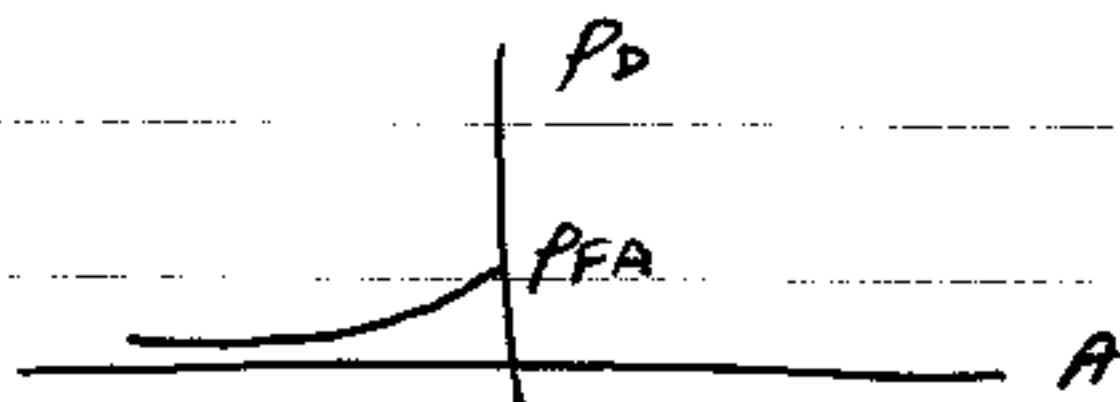
6.4) Decide H_1 if $\bar{x} > \tau' = \sqrt{\sigma^2/N} Q^{-1}(P_{FA})$

Under H_1 , $\bar{x} \sim N(A, \sigma^2/N)$ $A < 0$

$$P_D = Q\left(\frac{\tau' - A}{\sqrt{\sigma^2/N}}\right) = Q\left(Q^{-1}(P_{FA}) - \frac{A}{\sqrt{\sigma^2/N}}\right)$$

$$= Q\left(Q^{-1}(P_{FA}) + \sqrt{NA^2/\sigma^2}\right)$$

As $|A|$ increases P_D decreases



6.5) Decide H_1 if $|\bar{x}| > \tau''$

$$P_D = P_r\{|\bar{x}| > \tau''; H_1\}$$

$$= P_r\{\bar{x} > \tau''; H_1\} + P_r\{\bar{x} < -\tau''; H_1\}$$

$$= Q\left(\frac{\tau'' - A}{\sqrt{\sigma^2/N}}\right) + 1 - Q\left(\frac{-\tau'' - A}{\sqrt{\sigma^2/N}}\right)$$

$$= Q\left(\frac{\tau'' - A}{\sqrt{\sigma^2/N}}\right) + Q\left(\frac{\tau'' + A}{\sqrt{\sigma^2/N}}\right)$$

$$= Q\left(Q^{-1}\left(\frac{P_{FA}}{2}\right) - \frac{A}{\sqrt{\sigma^2/N}}\right)$$

$$+ Q\left(Q^{-1}\left(\frac{P_{FA}}{2}\right) + \frac{A}{\sqrt{\sigma^2/N}}\right)$$

Since $P_{FA} = P_r\{|\bar{x}| > \tau''; H_0\}$

$$= 2 P_r\{\bar{x} > \tau''; H_0\}$$

$$= 2 Q\left(\tau''/\sqrt{\sigma^2/N}\right)$$

$$P_D = Q \left(Q^{-1}(PFA/2) - \sqrt{NA^2/\sigma^2} \right) + Q \left(Q^{-1}(PFA/2) + \sqrt{NA^2/\sigma^2} \right)$$

$$6.6) p(\underline{x} | \sigma^2; \mathcal{H}_1) = \frac{1}{(2\pi\sigma^2)^{N/2}} e^{-\frac{1}{2\sigma^2} \sum_n (\underline{x}(n) - A)^2}$$

$$p(\underline{x}; \mathcal{H}_1) = \int_0^\infty p(\underline{x} | \sigma^2; \mathcal{H}_1) p(\sigma^2) d\sigma^2$$

$$= \int_0^\infty \frac{1}{(2\pi)^{N/2} \sigma^N} e^{-\frac{1}{2\sigma^2} Q} \lambda \frac{e^{-\lambda/\sigma^2}}{\sigma^4} d\sigma^2$$

$$= \frac{\lambda}{(2\pi)^{N/2}} \int_0^\infty \frac{1}{(\sigma^2)^{N/2+2}} e^{-(\lambda + 1/2 Q)/\sigma^2} d\sigma^2$$

$$\text{Let } u = 1/\sigma^2 \quad du = -1/\sigma^4 d\sigma^2 = -u^2 d\sigma^2$$

$$p(\underline{x}; \mathcal{H}_1) = \frac{\lambda}{(2\pi)^{N/2}} \int_0^\infty \frac{u^{N/2+2}}{u^2} e^{-(\lambda + \frac{1}{2} Q)u} du$$

$$= \frac{\lambda}{(2\pi)^{N/2}} (\lambda + 1/2 Q)^{\frac{N}{2}+1} \Gamma(\frac{N}{2}+1)$$

$$= \frac{\lambda}{(2\pi)^{N/2}} \left(\lambda + \frac{1}{2} \sum_n (\underline{x}(n) - A)^2 \right)^{\frac{N}{2}+1} \Gamma(\frac{N}{2}+1)$$

$$\Rightarrow p(\underline{x}; \mathcal{H}_0) = \frac{\lambda}{(2\pi)^{N/2}} \left(\lambda + \frac{1}{2} \sum_n \underline{x}^2(n) \right)^{\frac{N}{2}+1} \Gamma(\frac{N}{2}+1)$$

$$\text{Decide } \mathcal{H}_1 \text{ if } \frac{p(\underline{x}; \mathcal{H}_1)}{p(\underline{x}; \mathcal{H}_0)} > \tau$$

$$L(\underline{x}) = \left(\frac{\lambda + \frac{1}{2} \sum_n \underline{x}^2(n)}{\lambda + \frac{1}{2} \sum_n (\underline{x}(n) - A)^2} \right)^{\frac{N}{2}+1}$$

or we decide H_1 if

$$L(x)^{-\frac{N}{2}-1} = \frac{\lambda + \frac{1}{2} \sum x^2/n}{\lambda + \frac{1}{2} \sum (x/n - A)^2} > \delta'$$

If $\lambda \rightarrow 0$, we decide H_1 if

$$\frac{\frac{1}{N} \sum x^2/n}{\frac{1}{N} \sum (x/n - A)^2} > \delta'$$

$$\text{or } \frac{\hat{\sigma}_0^2}{\hat{\sigma}_1^2} > \delta'$$

As $\lambda \rightarrow 0$ $p(\sigma^2) \rightarrow \text{constant}$ and thus we have no prior knowledge. Hence, test statistic is ratio of fitting errors (similar to GLRT).

$$6.7) \quad |\bar{x}| > \sqrt{P_1}$$

$$\bar{x} \sim N(0, \sigma^2/N) \text{ under } H_0$$

$$N(0, \sigma_A^2 + \sigma^2/N) \text{ under } H_1$$

since under H_1 ,

$$\bar{x} = A + \bar{w} \leftarrow N(0, \sigma^2/N)$$

$$\uparrow \\ N(0, \sigma_A^2)$$

$$P_{FA} = P_r \{ |\bar{x}| > \sqrt{\gamma_1}; \mathcal{H}_0 \}$$

$$= 2Q(\sqrt{\gamma_1}/\sqrt{\sigma^2/N})$$

$$P_D = P_r \{ |\bar{x}| > \sqrt{\gamma_1}; \mathcal{H}_1 \}$$

$$= 2P_r \{ \bar{x} > \sqrt{\gamma_1}; \mathcal{H}_1 \}$$

$$= 2Q(\sqrt{\gamma_1}/\sqrt{\sigma_A^2 + \sigma^2/N})$$

$$\text{or } P_D = 2Q\left(\frac{\sqrt{\sigma^2/N} Q^{-1}(P_{FA}/2)}{\sqrt{\sigma_A^2 + \sigma^2/N}}\right)$$

$$= 2Q\left(\frac{Q^{-1}(P_{FA}/2)}{\sqrt{1 + \sigma_A^2/\sigma^2/N}}\right)$$

As $\sigma_A^2 \rightarrow \infty$, the argument $\rightarrow 0 \Rightarrow P_D \rightarrow 1$.

Signal power is much greater than noise power.

$$6.81 \quad L_G(\underline{x}) = \frac{p(\underline{x}; \hat{A}, \mathcal{H}_1)}{p(\underline{x}; \mathcal{H}_0)}$$

$$p(\underline{x}; A, \mathcal{H}_1) = \frac{1}{(2\pi\sigma^2)^{N/2}} e^{-\frac{1}{2\sigma^2} \sum_n (x(n) - Ar^n)^2}$$

To find \hat{A} we minimize $J(A) = \sum_n (x(n) - Ar^n)^2$

$$\frac{\partial J}{\partial A} = 0 \Rightarrow \sum_n (x(n) - \hat{A}r^n) r^n = 0$$

$$\Rightarrow \hat{A} = \frac{\sum_n x(n) r^n}{\sum_n r^{2n}}$$

$$L_G(\underline{x}) = \frac{e^{-\frac{1}{2\sigma^2} \sum_n (x[n] - \hat{A}r^n)^2}}{e^{-\frac{1}{2\sigma^2} \sum_n x^2[n]}}$$

$$\begin{aligned} \ln L_G(\underline{x}) &= -\frac{1}{2\sigma^2} \left[\sum_n (-2\hat{A}r^n x[n] + \hat{A}^2 r^{2n}) \right] \\ &= \frac{1}{\sigma^2} \sum_n (\hat{A}r^n x[n] - \frac{1}{2} \hat{A}^2 r^{2n}) \end{aligned}$$

$$\text{But } \sum_n r^n x[n] = \hat{A} \sum_n r^{2n}$$

$$\begin{aligned} \ln L_G(\underline{x}) &= \frac{\hat{A}}{\sigma^2} \sum_n r^n x[n] - \frac{1}{2} \frac{\hat{A}^2}{\sigma^2} \sum_n r^{2n} \\ &= \frac{\hat{A}^2}{\sigma^2} \sum_n r^{2n} - \frac{1}{2} \frac{\hat{A}^2}{\sigma^2} \sum_n r^{2n} \\ &= \frac{1}{2\sigma^2} \sum_n r^{2n} \hat{A}^2 \end{aligned}$$

Decide H_1 if $L_G(\underline{x}) > \gamma$ or

$$\frac{1}{2\sigma^2} \sum_n r^{2n} \hat{A}^2 > \ln \gamma$$

$$\hat{A}^2 > \frac{2\sigma^2 \ln \gamma}{\sum_n r^{2n}} = \gamma'$$

$$6.9) \quad L_G(\underline{x}) = \frac{p(\underline{x}; \hat{A}, \hat{\sigma}^2, H_1)}{p(\underline{x}; \hat{\sigma}^2, H_0)}$$