

统计信号处理

第四章

最小方差无偏估计 III

——线性模型与最佳线性无偏估计

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内容概要

- 一、引言
- 二、线性模型
- 三、最佳线性无偏估计
- 四、应用案例
- 五、小结

一、引言

- MVU估计量求解方法

- 借助于CRLB:

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$$

- ✓ 正则条件不一定满足
- ✓ 无法达到CRLB
- ✓ 不一定好求解
- ✓ 无似然函数

- 借助于充分统计量:

Neyman-Fisher因子分解: $p(\mathbf{x}; \boldsymbol{\theta}) = g(T(\mathbf{x}), \boldsymbol{\theta})h(\mathbf{x})$

若充分统计量完备, 构造无偏估计量

$$\hat{\boldsymbol{\theta}} = g(T(\mathbf{x}))$$

或求条件期望

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta} | T(\mathbf{x}))$$

- ✓ 不一定完备
- ✓ 不一定好求解
- ✓ 无似然函数

一、引言

- 特例——线性模型与一般线性模型
 - 观测数据是待估计参数的线性函数
 - 噪声为高斯的
- 妥协——最佳线性无偏估计(BLUE)
 - BLUE: best linear unbiased estimator
 - 线性 (linear)
 - 无偏 (unbiased)
 - 方差最小 (best)

二、线性模型

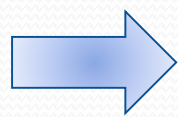
1. 线性模型

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

其中, \mathbf{x} 是 $N \times 1$ 维的观测数据, \mathbf{H} 是 $N \times p$ ($N > p$) 维、秩为 p 的观测矩阵, $\boldsymbol{\theta}$ 是 $p \times 1$ 维的待估计参数矢量, \mathbf{w} 是 $N \times 1$ 维的独立噪声矢量且服从 $N(\mathbf{0}, \sigma^2 \mathbf{I})$ 。

观测数据的似然函数为:

$$p(\mathbf{x}; \boldsymbol{\theta}) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} (\mathbf{x} - \mathbf{H}\boldsymbol{\theta})^T (\mathbf{x} - \mathbf{H}\boldsymbol{\theta}) \right\}$$


$$\ln p(\mathbf{x}; \boldsymbol{\theta}) = -\frac{N}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} (\mathbf{x} - \mathbf{H}\boldsymbol{\theta})^T (\mathbf{x} - \mathbf{H}\boldsymbol{\theta})$$

$$\rightarrow \frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = -\frac{1}{2\sigma^2} \frac{\partial \{ \mathbf{x}^T \mathbf{x} - 2\mathbf{x}^T \mathbf{H} \boldsymbol{\theta} + \boldsymbol{\theta}^T \mathbf{H}^T \mathbf{H} \boldsymbol{\theta} \}}{\partial \boldsymbol{\theta}} = \frac{1}{\sigma^2} \{ \mathbf{H}^T \mathbf{x} - \mathbf{H}^T \mathbf{H} \boldsymbol{\theta} \} \rightarrow$$

$$\frac{\partial \mathbf{b}^T \boldsymbol{\theta}}{\partial \boldsymbol{\theta}} = \mathbf{b}$$

$$\frac{\partial \boldsymbol{\theta}^T \mathbf{A} \boldsymbol{\theta}}{\partial \boldsymbol{\theta}} = 2\mathbf{A} \boldsymbol{\theta}$$

(\mathbf{A} 为对称矩阵)

CRLB定理: $\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \frac{\mathbf{H}^T \mathbf{H}}{\sigma^2} \left\{ \underbrace{(\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x}}_{\mathbf{I}(\boldsymbol{\theta})} - \underbrace{\boldsymbol{\theta}}_{\hat{\boldsymbol{\theta}}} \right\}$$

即，线性模型对应的MVU估计量是：

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x} \quad \text{——且是有效估计量}$$

相应的协方差阵为

$$\mathbf{C}_{\hat{\boldsymbol{\theta}}} = \sigma^2 (\mathbf{H}^T \mathbf{H})^{-1}$$

进一步地，该MVU估计量服从

$$\hat{\boldsymbol{\theta}} \sim N\left(\boldsymbol{\theta}, \sigma^2 (\mathbf{H}^T \mathbf{H})^{-1}\right) \quad \text{——这是普通MVU估计量不具备的}$$

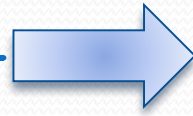
例：直线拟合

$$x[n] = A + Bn + w[n], \quad n = 0, 1, \dots, N-1$$

噪声为高斯白噪声，且 $w[n] \sim N(0, \sigma^2)$ ，待估计参数为 $\theta = [A, B]^T$ ，其 MVU 估计量是？

方法一：直接利用矢量参数 CRLB 定理

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$$

$$p(x[n]; \boldsymbol{\theta}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x[n] - A - Bn)^2\right\}$$


$$p(\mathbf{x}; \boldsymbol{\theta}) = \prod_{n=0}^{N-1} p(x[n]; \boldsymbol{\theta})$$

$$p(\mathbf{x}; \boldsymbol{\theta}) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A - Bn)^2\right\}$$

$$p(\mathbf{x};\boldsymbol{\theta}) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A - Bn)^2 \right\}$$

➡ $\ln p(\mathbf{x};\boldsymbol{\theta}) = -\frac{N}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A - Bn)^2$

➡
$$\left\{ \begin{array}{l} \frac{\partial \ln p(\mathbf{x};\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \left\{ \begin{array}{l} \frac{\partial \ln p(\mathbf{x};\boldsymbol{\theta})}{\partial A} \\ \frac{\partial \ln p(\mathbf{x};\boldsymbol{\theta})}{\partial B} \end{array} \right\} = \left\{ \begin{array}{l} \frac{1}{\sigma^2} \sum_{n=0}^{N-1} (x[n] - A - Bn) \\ \frac{1}{\sigma^2} \sum_{n=0}^{N-1} (x[n] - A - Bn)n \end{array} \right\} \\ \\ \mathbf{I}(\boldsymbol{\theta}) = \frac{1}{\sigma^2} \begin{bmatrix} N & \frac{N(N-1)}{2} \\ \frac{N(N-1)}{2} & \frac{N(N-1)(2N-1)}{6} \end{bmatrix} \end{array} \right.$$

$$\frac{\partial \ln p(\mathbf{x};\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta}) \quad \text{如何构建?}$$

例：直线拟合

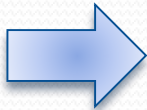
$$x[n] = A + Bn + w[n], \quad n = 0, 1, \dots, N-1$$

噪声为高斯白噪声，且 $w[n] \sim N(0, \sigma^2)$ ，待估计参数为 $\theta = [A, B]^T$ ，其 MVU 估计量是？

方法二：采用线性模型的方法

$$\hat{\theta} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x}$$

$$\mathbf{H} = \begin{Bmatrix} 1 & 0 \\ 1 & 1 \\ 1 & 2 \\ \vdots & \vdots \\ 1 & N-1 \end{Bmatrix} \Rightarrow \mathbf{H}^T \mathbf{H} = \begin{pmatrix} N & \frac{N(N-1)}{2} \\ \frac{N(N-1)}{2} & \frac{N(N-1)(2N-1)}{6} \end{pmatrix}$$



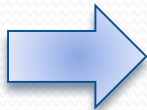
$$(\mathbf{H}^T \mathbf{H})^{-1} = \frac{12}{N^2(N-1)(N+1)} \begin{pmatrix} \frac{N(N-1)(2N-1)}{6} & -\frac{N(N-1)}{2} \\ -\frac{N(N-1)}{2} & N \end{pmatrix}$$

线性模型

- 本质上源自于CRLB定理
- 但用起来更直观、方便

$$\mathbf{H} = \begin{Bmatrix} 1 & 1 & 1 & \dots & 1 \\ 0 & 1 & 2 & \dots & N-1 \end{Bmatrix}^T$$


$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x}$$



$$\hat{\boldsymbol{\theta}} = \begin{cases} \frac{2(2N-1)}{N(N+1)} \sum_{n=0}^{N-1} x(n) - \frac{6}{N(N+1)} \sum_{n=0}^{N-1} x(n)n \\ -\frac{6}{N(N+1)} \sum_{n=0}^{N-1} x(n) + \frac{12}{N(N-1)(N+1)} \sum_{n=0}^{N-1} x(n)n \end{cases}$$

$= \hat{A}$
 $= \hat{B}$

$$C_{\hat{\boldsymbol{\theta}}} = \sigma^2 (\mathbf{H}^T \mathbf{H})^{-1} = \frac{12\sigma^2}{N^2(N-1)(N+1)} \begin{pmatrix} \frac{N(N-1)(2N-1)}{6} & -\frac{N(N-1)}{2} \\ -\frac{N(N-1)}{2} & N \end{pmatrix}$$


 $\begin{cases} \text{var}(\hat{A}) = \frac{2(2N-1)}{N(N+1)} \sigma^2 \\ \text{var}(\hat{B}) = \frac{12}{N(N^2-1)} \sigma^2 \end{cases}$

例：频率分量幅度分析

$$x[n] = \sum_{k=1}^M a_k \cos\left(\frac{2\pi kn}{N}\right) + \sum_{k=1}^M b_k \sin\left(\frac{2\pi kn}{N}\right) + w[n], \quad n = 0, 1, \dots, N-1$$

假定频率是基频 $1/N$ 的谐波，即 $f_k = k/N$ 。噪声为高斯白噪声。待估计参数为各频率分量的幅度 a_k, b_k ($1 \leq k \leq M$)。

观测数据： $\mathbf{x} = [x[0], x[1], \dots, x[N-1]]^T$

待估计参数矢量： $\boldsymbol{\theta} = [a_1, a_2, \dots, a_M, b_1, b_2, \dots, b_M]^T$

观测矩阵：

$$\mathbf{H} = \begin{bmatrix} \begin{matrix} \downarrow 1 \\ \cos\left(\frac{2\pi}{N}\right) \\ \vdots \\ \cos\left(\frac{2(N-1)\pi}{N}\right) \end{matrix} & \begin{matrix} \downarrow 1 \\ \cos\left(\frac{4\pi}{N}\right) \\ \vdots \\ \cos\left(\frac{4(N-1)\pi}{N}\right) \end{matrix} & \dots & \begin{matrix} \downarrow 1 \\ \cos\left(\frac{2M\pi}{N}\right) \\ \vdots \\ \cos\left(\frac{2M(N-1)\pi}{N}\right) \end{matrix} & \begin{matrix} \downarrow 0 \\ \sin\left(\frac{2\pi}{N}\right) \\ \vdots \\ \sin\left(\frac{2(N-1)\pi}{N}\right) \end{matrix} & \begin{matrix} \downarrow 0 \\ \sin\left(\frac{4\pi}{N}\right) \\ \vdots \\ \sin\left(\frac{4(N-1)\pi}{N}\right) \end{matrix} & \dots & \begin{matrix} \downarrow 0 \\ \sin\left(\frac{2M\pi}{N}\right) \\ \vdots \\ \sin\left(\frac{2M(N-1)\pi}{N}\right) \end{matrix} \\ \mathbf{h}_1 & \mathbf{h}_2 & & \mathbf{h}_M & \mathbf{h}_{M+1} & \mathbf{h}_{M+2} & & \mathbf{h}_{2M} \end{bmatrix}$$

该模型为线性模型，因此待估计参数 θ 的MVU为

$$\begin{aligned}\hat{\theta} &= (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x} \\ &= \frac{2}{N} \mathbf{H}^T \mathbf{x} \\ &= \left[\frac{2}{N} \mathbf{h}_1^T \mathbf{x}, \frac{2}{N} \mathbf{h}_2^T \mathbf{x}, \dots, \frac{2}{N} \mathbf{h}_{2M}^T \mathbf{x} \right]^T\end{aligned}$$

即

$$\begin{aligned}\hat{a}_k &= \frac{2}{N} \sum_{n=0}^{N-1} x[n] \cos\left(\frac{2\pi kn}{N}\right) \\ \hat{b}_k &= \frac{2}{N} \sum_{n=0}^{N-1} x[n] \sin\left(\frac{2\pi kn}{N}\right)\end{aligned}$$

离散傅里叶变换的系数！

2.一般线性模型

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

其中, \mathbf{x} 是 $N \times 1$ 维的观测数据, \mathbf{H} 是 $N \times p$ ($N > p$) 维、秩为 p 的观测矩阵, $\boldsymbol{\theta}$ 是 $p \times 1$ 维的待估计参数矢量, \mathbf{w} 是 $N \times 1$ 维的独立噪声矢量且服从 $N(\mathbf{0}, \sigma^2 \mathbf{I})$ 。

拓展

拓展一: 允许存在已知信号

拓展二: 可以是高斯有色噪声

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{s} + \mathbf{w}$$

其中, \mathbf{x} 是 $N \times 1$ 维的观测数据, \mathbf{H} 是 $N \times p$ ($N > p$) 维、秩为 p 的观测矩阵, $\boldsymbol{\theta}$ 是 $p \times 1$ 维的待估计参数矢量, \mathbf{s} 是 $N \times 1$ 维的已知信号矢量, \mathbf{w} 是 $N \times 1$ 维的噪声矢量且服从 $N(\mathbf{0}, \mathbf{C})$ 。

观测数据似然函数为:

$$p(\mathbf{x}; \boldsymbol{\theta}) = \frac{1}{(2\pi)^{N/2} (\det(\mathbf{C}))^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta})^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta}) \right\} \rightarrow$$

$$\ln p(\mathbf{x}; \boldsymbol{\theta}) = -\ln \left\{ (2\pi)^{N/2} (\det(\mathbf{C}))^{1/2} \right\} - \frac{1}{2} (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta})^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta}) \rightarrow$$

$$\begin{aligned} \frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} &= -\frac{1}{2} \frac{\partial \left\{ (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta})^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s} - \mathbf{H}\boldsymbol{\theta}) \right\}}{\partial \boldsymbol{\theta}} = -\frac{1}{2} \frac{\partial \left\{ -2(\mathbf{x} - \mathbf{s})^T \mathbf{C}^{-1} \mathbf{H}\boldsymbol{\theta} + \boldsymbol{\theta}^T \mathbf{H}^T \mathbf{C}^{-1} \mathbf{H}\boldsymbol{\theta} \right\}}{\partial \boldsymbol{\theta}} \\ &= \mathbf{H}^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s}) - \mathbf{H}^T \mathbf{C}^{-1} \mathbf{H}\boldsymbol{\theta} \end{aligned}$$

CRLB定理: $\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{H}^T \mathbf{C}^{-1} \mathbf{H} \left((\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s}) - \boldsymbol{\theta} \right)$$

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s}) \quad \text{——是有效估计量}$$

$$\mathbf{C}_{\hat{\boldsymbol{\theta}}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \quad \text{且 } \hat{\boldsymbol{\theta}} \sim N(\boldsymbol{\theta}, (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1})$$

例：有色噪声电平估计

$$x[n] = A + w[n], \quad n = 0, 1, \dots, N-1$$

噪声为有色噪声，即 $\mathbf{w} \sim N(\mathbf{0}, \mathbf{C})$ ，待估计参数为电平 A ，其MVU为？

其MVU估计为

$$\hat{A} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s})$$

观测矩阵为 $\mathbf{H} = [1, 1, \dots, 1]^T = \mathbf{1}$ ，

$$\text{故 } \hat{A} = \frac{\mathbf{1}^T \mathbf{C}^{-1} \mathbf{x}}{\mathbf{1}^T \mathbf{C}^{-1} \mathbf{1}}$$

令 $\mathbf{C}^{-1} = \mathbf{D}^T \mathbf{D}$ ，则有

$$\hat{A} = \frac{(\mathbf{D}\mathbf{1})^T (\mathbf{D}\mathbf{x})}{\mathbf{1}^T \mathbf{D}^T \mathbf{D} \mathbf{1}} = \frac{(\mathbf{D}\mathbf{1})^T \mathbf{x}'}{\mathbf{1}^T \mathbf{D}^T \mathbf{D} \mathbf{1}} = \sum_{n=0}^{N-1} d_n x'[n], \quad \text{其中 } d_n = \frac{[\mathbf{D}\mathbf{1}]_n}{\mathbf{1}^T \mathbf{D}^T \mathbf{D} \mathbf{1}}$$

若为高斯白噪声，有

$$\hat{A} = \frac{\mathbf{1}^T \mathbf{x}}{N} = \frac{1}{N} \sum_{n=0}^{N-1} x[n]$$

三、最佳线性无偏估计 (BLUE)

1. 标量参数的BLUE

- 观察数据 $\mathbf{x} = \{x[0], x[1], x[2], \dots, x[N-1]\}^T$
- BLUE意味着:

1 限定估计量与观察数据间呈线性关系

$$\hat{\theta} = \sum_{n=0}^{N-1} a_n x[n] = \mathbf{a}^T \mathbf{x} \quad , \quad \text{其中 } \mathbf{a} = [a_0, a_1, \dots, a_{N-1}]^T$$

2 无偏: $E(\hat{\theta}) = \sum_{n=0}^{N-1} a_n E(x[n]) = \theta$

\downarrow

$E(x[n]) = s[n]\theta$

\downarrow

$\sum_{n=0}^{N-1} a_n s[n] = 1 \quad \text{即} \quad \mathbf{a}^T \mathbf{s} = 1$

$$E(\mathbf{x}) = \mathbf{s}\theta \quad \text{其中 } \mathbf{s} = [s[0], s[1], \dots, s[N-1]]^T$$

$$\left. \begin{aligned} \text{var}(\hat{\theta}) &= E\left\{\left(\hat{\theta} - E(\hat{\theta})\right)^2\right\} \\ \hat{\theta} &= \sum_{n=0}^{N-1} a_n x[n] = \mathbf{a}^T \mathbf{x} \end{aligned} \right\} \Rightarrow \begin{aligned} \text{var}(\hat{\theta}) &= E\left\{\left(\mathbf{a}^T \mathbf{x} - \mathbf{a}^T E(\mathbf{x})\right)^2\right\} \\ &= E\left\{\left(\mathbf{a}^T (\mathbf{x} - E(\mathbf{x}))\right)^2\right\} \\ &= E\left\{\mathbf{a}^T (\mathbf{x} - E(\mathbf{x})) (\mathbf{x} - E(\mathbf{x}))^T \mathbf{a}\right\} \\ &= \mathbf{a}^T \mathbf{C} \mathbf{a} \end{aligned}$$

3 $\min \{\mathbf{a}^T \mathbf{C} \mathbf{a}\}$

标量参数BLUE:

$$\begin{cases} \min \{\mathbf{a}^T \mathbf{C} \mathbf{a}\} & \text{方差最小(Best)} \\ s.t. \quad \mathbf{a}^T \mathbf{s} = 1 & \text{线性(Linear)、无偏(Unbiased)} \end{cases}$$

采用拉格朗日乘子法

$$\text{令 } J = \mathbf{a}^T \mathbf{C} \mathbf{a} + \lambda (\mathbf{a}^T \mathbf{s} - 1)$$

$$J = \mathbf{a}^T \mathbf{C} \mathbf{a} + \lambda (\mathbf{a}^T \mathbf{s} - 1)$$

$$\begin{aligned} \Rightarrow \left. \begin{aligned} \frac{\partial J}{\partial \mathbf{a}} &= 2\mathbf{C}\mathbf{a} + \lambda \mathbf{s} \\ \frac{\partial J}{\partial \lambda} &= 0 \end{aligned} \right\}$$

$$\begin{aligned} \Rightarrow \left. \begin{aligned} \mathbf{a} &= -\frac{\lambda}{2} \mathbf{C}^{-1} \mathbf{s} \\ \mathbf{a}^T \mathbf{s} &= 1 \end{aligned} \right\}$$

$$\Rightarrow -\frac{\lambda}{2} = \frac{1}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}}$$

$$\Rightarrow \mathbf{a}_{opt} = \frac{\mathbf{C}^{-1} \mathbf{s}}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}}$$

$$\begin{aligned} \text{BLUE: } \hat{\theta} &= \mathbf{a}_{opt}^T \mathbf{x} = \frac{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{x}}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}} \Rightarrow E(\hat{\theta}) = \frac{\mathbf{s}^T \mathbf{C}^{-1} E(\mathbf{x})}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}} \\ &= \theta \quad \text{无偏!} \\ \text{相应的方差: } \text{var}(\hat{\theta}) &= \mathbf{a}_{opt}^T \mathbf{C} \mathbf{a}_{opt} = \frac{1}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}} \end{aligned}$$

BLUE估计所需条件:

- 成比例的均值: \mathbf{s}
- 协方差矩阵: \mathbf{C}

优缺点:

- 无需PDF, 只需一、二阶矩
- 假定线性, 是否合理?

成比例的均值

协方差矩阵

例：白噪声中电平估计

$$x[n] = A + w[n], \quad n = 0, 1, \dots, N-1$$

待估计参数为信号幅度 A ， $w[n]$ 为高斯白噪声，且 $w[n] \sim N(0, \sigma^2)$
求 A 的BLUE?

$$\begin{aligned} \hat{A} &= \frac{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{x}}{\mathbf{s}^T \mathbf{C}^{-1} \mathbf{s}} \\ \mathbf{s} &= [1, 1, 1, \dots, 1]^T = \mathbf{1} \\ \mathbf{C} &= \sigma^2 \mathbf{I} \Rightarrow \mathbf{C}^{-1} = \frac{1}{\sigma^2} \mathbf{I} \end{aligned} \quad \left\{ \begin{array}{l} \Rightarrow \hat{A} = \frac{\mathbf{1}^T \frac{1}{\sigma^2} \mathbf{I} \mathbf{x}}{\mathbf{1}^T \frac{1}{\sigma^2} \mathbf{I} \mathbf{1}} \\ = \frac{1}{N} \sum_{n=0}^{N-1} x[n] \quad \text{——BLUE估计量!} \end{array} \right.$$

- 同时也是MVU估计量!
- BLUE = MVU ?

2. 矢量参数的BLUE

- 观察数据 $\mathbf{x} = \{x[0], x[1], x[2], \dots, x[N-1]\}^T$

- BLUE意味着:

1 限定每个估计量与观察数据间呈线性关系

$$\begin{aligned}\hat{\theta}_i &= \sum_{n=0}^{N-1} a_{in} x[n], \quad i = 1, 2, \dots, p \\ &= \mathbf{a}_i^T \mathbf{x}, \quad \mathbf{a}_i = [a_{i0}, a_{i1}, \dots, a_{i,N-1}]^T\end{aligned} \quad \Rightarrow \quad \hat{\boldsymbol{\theta}}_{p \times 1} = \mathbf{A}_{p \times N} \mathbf{x}_{N \times 1}, \quad \mathbf{A}_{p \times N} = \begin{bmatrix} \mathbf{a}_1^T \\ \mathbf{a}_2^T \\ \vdots \\ \mathbf{a}_p^T \end{bmatrix}$$

2 无偏: $E(\hat{\theta}_i) = \sum_{n=0}^{N-1} a_{in} E(x[n]) = \theta_i \quad \Rightarrow \quad E(\hat{\boldsymbol{\theta}}) = \mathbf{A}E(\mathbf{x}) = \boldsymbol{\theta}$

\Downarrow
 $E(\mathbf{x}) = \mathbf{H}\boldsymbol{\theta}$

$\Rightarrow \mathbf{A}\mathbf{H} = \mathbf{I}$

由于 $\mathbf{A} = \begin{bmatrix} \mathbf{a}_1^T \\ \mathbf{a}_2^T \\ \vdots \\ \mathbf{a}_p^T \end{bmatrix}, \quad \mathbf{H} = [\mathbf{h}_1, \mathbf{h}_2, \dots, \mathbf{h}_p]$

$\Rightarrow \mathbf{a}_i^T \mathbf{h}_j = \delta_{ij}$

$$\begin{aligned}
 \left. \begin{aligned} \text{var}(\hat{\theta}_i) &= E\left\{\left(\hat{\theta}_i - E(\hat{\theta}_i)\right)^2\right\} \\ \hat{\theta}_i &= \sum_{n=0}^{N-1} a_{in} x[n] = \mathbf{a}_i^T \mathbf{x} \end{aligned} \right\} \Rightarrow \begin{aligned} \text{var}(\hat{\theta}_i) &= E\left\{\left(\mathbf{a}_i^T \mathbf{x} - \mathbf{a}_i^T E(\mathbf{x})\right)^2\right\} \\ &= E\left\{\left(\mathbf{a}_i^T (\mathbf{x} - E(\mathbf{x}))\right)^2\right\} \\ &= E\left\{\mathbf{a}_i^T (\mathbf{x} - E(\mathbf{x})) (\mathbf{x} - E(\mathbf{x}))^T \mathbf{a}_i\right\} \\ &= \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i \end{aligned}
 \end{aligned}$$

$$3 \quad \min \left\{ \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i \right\}$$

矢量参数BLUE:

$$\begin{cases} \min \left\{ \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i \right\} \\ \text{s.t. } \mathbf{a}_i^T \mathbf{h}_j = \delta_{ij}, \quad 1 \leq i, j \leq p \end{cases}$$

采用拉格朗日乘子法

$$\text{令 } J_i = \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i + \sum_{j=1}^p \lambda_j^{(i)} (\mathbf{a}_i^T \mathbf{h}_j - \delta_{ij})$$

$$J_i = \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i + \sum_{j=1}^p \lambda_j^{(i)} (\mathbf{a}_i^T \mathbf{h}_j - \delta_{ij})$$

$$\Rightarrow \frac{\partial J_i}{\partial \mathbf{a}_i} = 2\mathbf{C}\mathbf{a}_i + \sum_{j=1}^p \lambda_j^{(i)} \mathbf{h}_j$$

$$\text{令 } \boldsymbol{\lambda}_i = [\lambda_1^{(i)}, \lambda_2^{(i)}, \dots, \lambda_p^{(i)}]^T$$

$$\text{已知 } \mathbf{H} = [\mathbf{h}_1, \mathbf{h}_2, \dots, \mathbf{h}_p]$$

$$\Rightarrow \frac{\partial J_i}{\partial \mathbf{a}_i} = 2\mathbf{C}\mathbf{a}_i + \mathbf{H}\boldsymbol{\lambda}_i$$

$$\frac{\partial J_i}{\partial \mathbf{a}_i} = 0$$

$$\Rightarrow \mathbf{a}_i = -\frac{1}{2} \mathbf{C}^{-1} \mathbf{H} \boldsymbol{\lambda}_i$$

另一方面，对约束条件：

$$\mathbf{a}_i^T \mathbf{h}_j = \delta_{ij}, \quad 1 \leq i, j \leq p$$

$$\Rightarrow \mathbf{h}_j^T \mathbf{a}_i = \delta_{ij}, \quad 1 \leq i, j \leq p$$

$$\Rightarrow \begin{bmatrix} \mathbf{h}_1^T \\ \vdots \\ \mathbf{h}_{i-1}^T \\ \mathbf{h}_i^T \\ \mathbf{h}_{i+1}^T \\ \vdots \\ \mathbf{h}_p^T \end{bmatrix} \mathbf{a}_i = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \mathbf{e}_i$$

$$\Rightarrow \mathbf{H}^T \mathbf{a}_i = \mathbf{e}_i$$

$$\Rightarrow \mathbf{H}^T \mathbf{a}_i = -\frac{1}{2} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{H} \lambda_i = \mathbf{e}_i$$

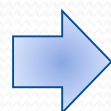
$$\text{var}(\hat{\theta}_i) = \mathbf{e}_i^T (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{e}_i$$

$$\Rightarrow \left. \begin{aligned} -\frac{1}{2} \lambda_i &= (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{e}_i \\ \mathbf{a}_i &= -\frac{1}{2} \mathbf{C}^{-1} \mathbf{H} \lambda_i \end{aligned} \right\}$$

$$\Rightarrow \mathbf{a}_{i,opt} = \mathbf{C}^{-1} \mathbf{H} (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{e}_i$$

$$\text{var}(\hat{\theta}_i) = \mathbf{a}_i^T \mathbf{C} \mathbf{a}_i$$

$$\text{已知 } \hat{\theta}_i = \sum_{n=0}^{N-1} a_{in} x[n] = \mathbf{a}_i^T \mathbf{x}$$



$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} \mathbf{a}_{1,opt}^T \mathbf{x} \\ \mathbf{a}_{2,opt}^T \mathbf{x} \\ \vdots \\ \mathbf{a}_{p,opt}^T \mathbf{x} \end{bmatrix}$$

$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} \mathbf{e}_1^T (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x} \\ \mathbf{e}_2^T (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x} \\ \vdots \\ \mathbf{e}_p^T (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x} \end{bmatrix} = \begin{bmatrix} \mathbf{e}_1^T \\ \mathbf{e}_2^T \\ \vdots \\ \mathbf{e}_p^T \end{bmatrix} (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x}$$

● 对具有一般线性模型形式的观测数据

观测数据模型：

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

——噪声零均值、协方差为 \mathbf{C} ，具体分布未知

BLUE: $\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x}$

→ $\hat{\boldsymbol{\theta}} = \boldsymbol{\theta} + (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{w}$

$\mathbf{C}_{\hat{\boldsymbol{\theta}}} = E \left[(\hat{\boldsymbol{\theta}} - E(\hat{\boldsymbol{\theta}})) (\hat{\boldsymbol{\theta}} - E(\hat{\boldsymbol{\theta}}))^T \right]$

$\mathbf{C}_{\hat{\boldsymbol{\theta}}} = E \left[\left((\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{w} \right) \left((\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{w} \right)^T \right]$

$= E \left[(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{w} \mathbf{w}^T \mathbf{C}^{-1} \mathbf{H} (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \right]$

$= (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{C} \mathbf{C}^{-1} \mathbf{H} (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}$

$= (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}$

由上一页推导知: $\text{var}(\hat{\theta}_i) = \mathbf{e}_i^T (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{e}_i$ → $\text{var}(\hat{\theta}_i) = [\mathbf{C}_{\hat{\boldsymbol{\theta}}}]_{ii} = [(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}]_{ii}$

如果观测数据具有一般线性模型的形式，即

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

其中 \mathbf{H} 是已知的 $N \times p$ 矩阵， $\boldsymbol{\theta}$ 是 $p \times 1$ 的待估计参数， \mathbf{w} 是
的均值为零、协方差为 \mathbf{C} 的噪声矢量（**不一定为高斯**），
那么 $\boldsymbol{\theta}$ 的BLUE估计量是

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x}$$

$\hat{\boldsymbol{\theta}}$ 的协方差矩阵为

$$\mathbf{C}_{\hat{\boldsymbol{\theta}}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}$$

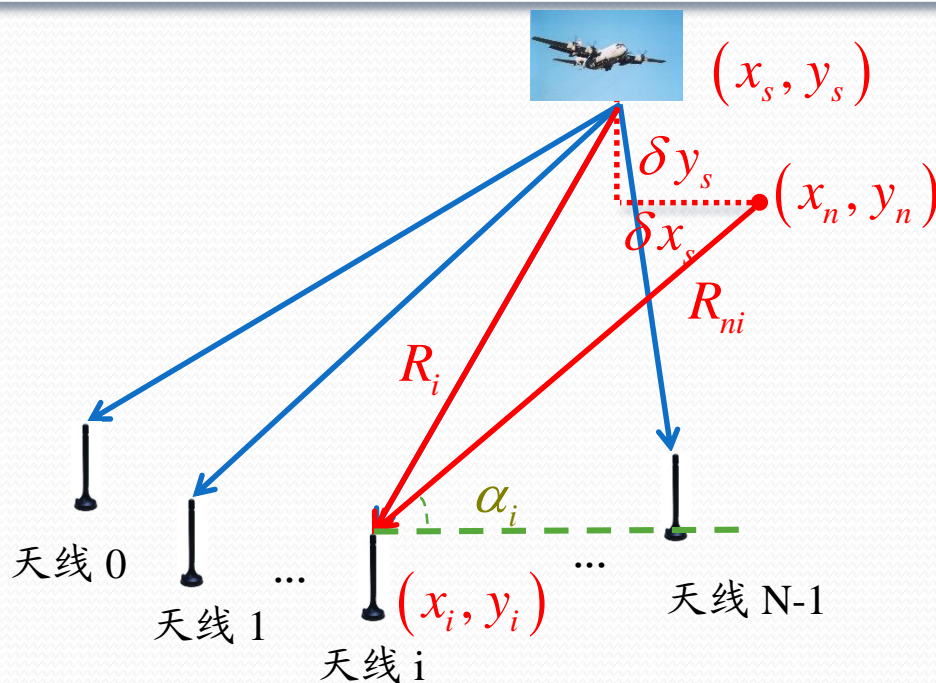
每个估计量 $\hat{\theta}_i$ 的方差为

$$\text{var}(\hat{\theta}_i) = \left[(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \right]_{ii}$$

高斯-马尔
可夫定理

- 若为高斯噪声，则BLUE为MVU，且为有效估计量

四、应用案例



到第 i 颗天线的时间: $t_i = T_0 + \frac{R_i}{c} + \varepsilon_i$

$$R_i = \sqrt{(x_s - x_i)^2 + (y_s - y_i)^2}$$

$$= \underbrace{\sqrt{(x_n - x_i)^2 + (y_n - y_i)^2}}_{R_{ni}} + \frac{x_n - x_i}{R_{ni}} \delta x_s + \frac{y_n - y_i}{R_{ni}} \delta y_s = R_{ni} + \underbrace{\frac{x_n - x_i}{R_{ni}}}_{\cos \alpha_i} \delta x_s + \underbrace{\frac{y_n - y_i}{R_{ni}}}_{\sin \alpha_i} \delta y_s$$

$$t_i = T_0 + \frac{R_{ni}}{c} + \frac{\cos \alpha_i}{c} \delta x_s + \frac{\sin \alpha_i}{c} \delta y_s + \varepsilon_i, i = 0, 1, 2, \dots, N-1$$

$$\underline{t_i - \frac{R_{ni}}{c}} = T_0 + \frac{\cos \alpha_i}{c} \delta x_s + \frac{\sin \alpha_i}{c} \delta y_s + \varepsilon_i$$



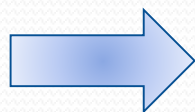
$$\tau_i = T_0 + \frac{\cos \alpha_i}{c} \delta x_s + \frac{\sin \alpha_i}{c} \delta y_s + \varepsilon_i$$

$$\xi_1 = \tau_1 - \tau_0 = \frac{1}{c} (\cos \alpha_1 - \cos \alpha_0) \delta x_s + \frac{1}{c} (\sin \alpha_1 - \sin \alpha_0) \delta y_s + (\varepsilon_1 - \varepsilon_0)$$

$$\xi_2 = \tau_2 - \tau_1 = \frac{1}{c} (\cos \alpha_2 - \cos \alpha_1) \delta x_s + \frac{1}{c} (\sin \alpha_2 - \sin \alpha_1) \delta y_s + (\varepsilon_2 - \varepsilon_1)$$

...

$$\xi_i = \tau_i - \tau_{i-1} = \frac{1}{c} (\cos \alpha_i - \cos \alpha_{i-1}) \delta x_s + \frac{1}{c} (\sin \alpha_i - \sin \alpha_{i-1}) \delta y_s + (\varepsilon_i - \varepsilon_{i-1})$$



$$\boldsymbol{\xi} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

数据模型： $\xi = \mathbf{H}\theta + \mathbf{w}$

待估计参数： $\theta = [\delta x_s, \delta y_s]^T$

观测矩阵： $\mathbf{H} = \frac{1}{c} \begin{bmatrix} \cos \alpha_1 - \cos \alpha_0 & \sin \alpha_1 - \sin \alpha_0 \\ \cos \alpha_2 - \cos \alpha_1 & \sin \alpha_2 - \sin \alpha_1 \\ \vdots & \vdots \\ \cos \alpha_{N-1} - \cos \alpha_{N-2} & \sin \alpha_{N-1} - \sin \alpha_{N-2} \end{bmatrix}$

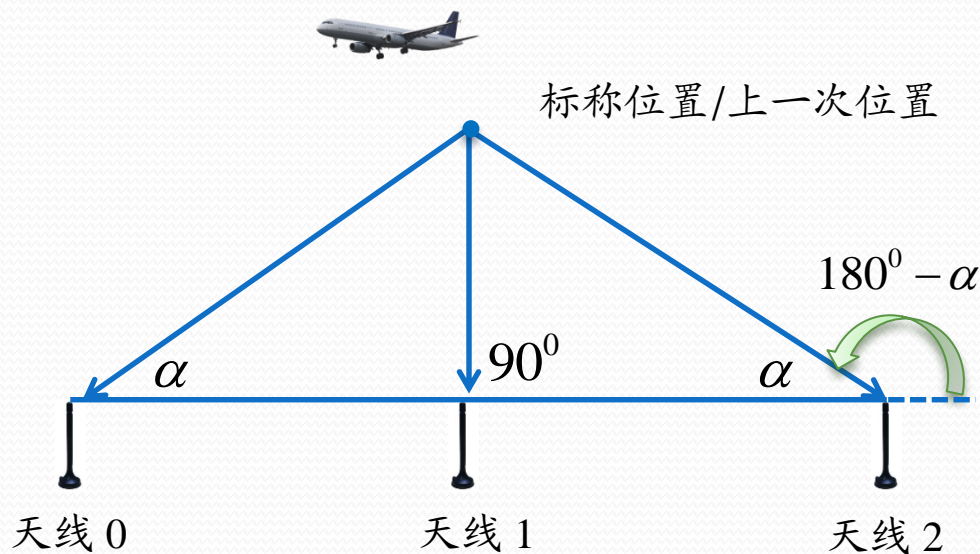
噪声矢量： $\mathbf{w} = \begin{bmatrix} \varepsilon_1 - \varepsilon_0 \\ \varepsilon_2 - \varepsilon_1 \\ \vdots \\ \varepsilon_{N-1} - \varepsilon_{N-2} \end{bmatrix} = \underbrace{\begin{bmatrix} -1 & 1 & 0 & 0 & \cdots & 0 \\ 0 & -1 & 1 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & 0 \\ 0 & 0 & 0 & 0 & -1 & 1 \end{bmatrix}}_{\mathbf{A}} \underbrace{\begin{bmatrix} \varepsilon_0 \\ \varepsilon_1 \\ \vdots \\ \varepsilon_{N-1} \end{bmatrix}}_{\boldsymbol{\varepsilon}}$

$$\mathbf{C} = E\left((\mathbf{w} - E(\mathbf{w}))(\mathbf{w} - E(\mathbf{w}))^T\right) = E(\mathbf{w}\mathbf{w}^T) = E(\mathbf{A}\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^T\mathbf{A}^T) = \sigma^2\mathbf{A}\mathbf{A}^T$$

BLUE估计量： $\hat{\theta} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x} = \left(\mathbf{H}^T (\mathbf{A}\mathbf{A}^T)^{-1} \mathbf{H} \right)^{-1} \mathbf{H}^T (\mathbf{A}\mathbf{A}^T)^{-1} \xi$

相应的协方差矩阵： $\mathbf{C}_{\hat{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} = \sigma^2 \left(\mathbf{H}^T (\mathbf{A}\mathbf{A}^T)^{-1} \mathbf{H} \right)^{-1}$

特例：三天线时



$$\alpha_0 = \alpha$$

$$\alpha_1 = 90^\circ$$

$$\alpha_2 = 180^\circ - \alpha$$

$$\mathbf{H} = \frac{1}{c} \begin{bmatrix} -\cos \alpha & 1 - \sin \alpha \\ -\cos \alpha & \sin \alpha - 1 \end{bmatrix}$$

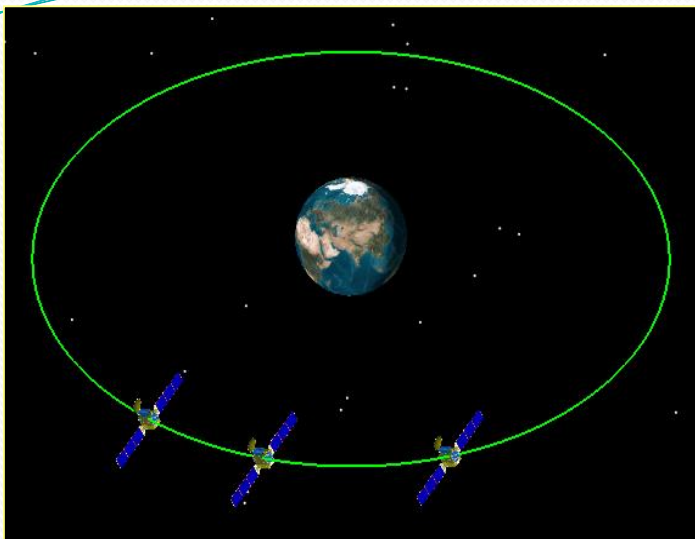
$$\mathbf{A} = \begin{bmatrix} -1 & 1 & 0 \\ 0 & -1 & 1 \end{bmatrix}$$

$$\mathbf{C}_{\hat{\theta}} = \sigma^2 \left(\mathbf{H}^T (\mathbf{A} \mathbf{A}^T)^{-1} \mathbf{H} \right)^{-1}$$

$$\mathbf{C}_{\hat{\theta}} = \sigma^2 c^2 \begin{bmatrix} \frac{1}{2 \cos^2 \alpha} & 0 \\ 0 & \frac{3/2}{(1 - \sin \alpha)^2} \end{bmatrix}$$

- α 越小，定位精度越好
- 增加天线阵距离

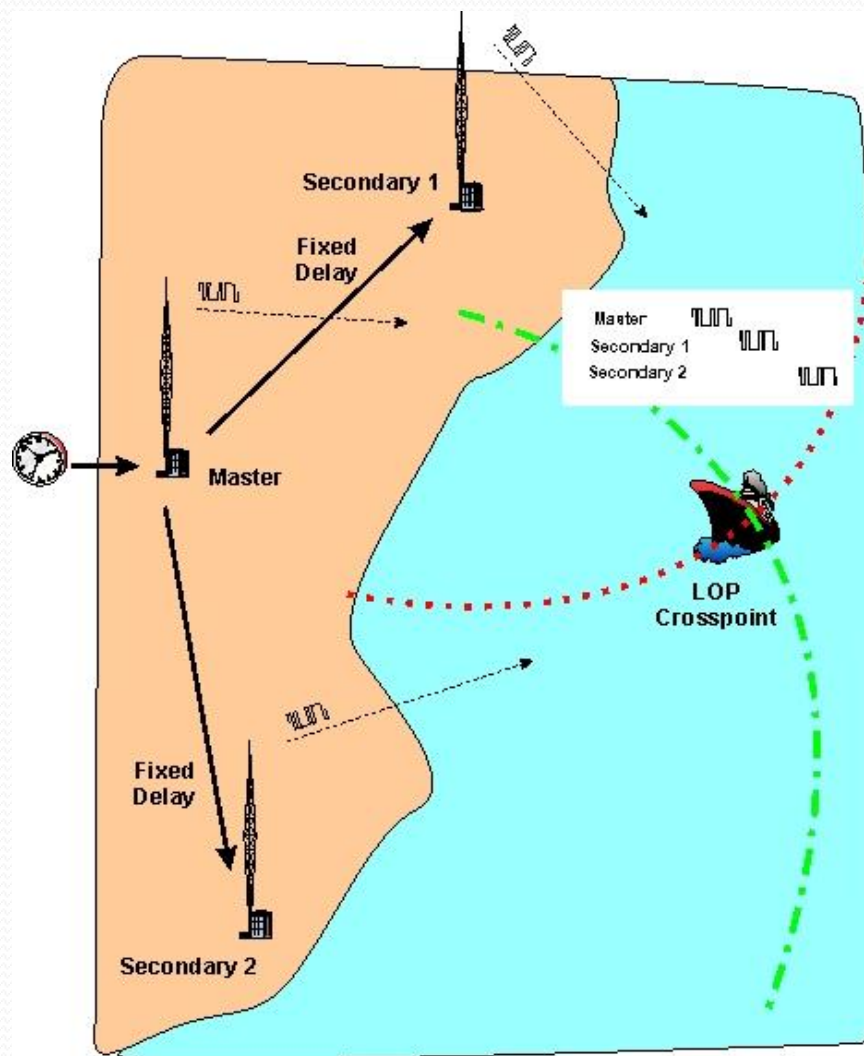
➤ 工程应用——北斗一号



北斗一号主要特点:

- ✓ 1994年开始建设, 2002年正式投入使用
- ✓ 由3颗GEO卫星构成, 其中一颗为备份卫星, 分别位于东经80度、140度和110.5度
- ✓ 覆盖范围北纬5~55度, 东经70~145度
- ✓ 定位精度: 约100米, 标较后20米
- ✓ 授时精度: 单向100纳秒, 双向20纳秒

罗兰系统 (Loran)



五、小结

- 求解MVU的难处
- 特例——线性模型与一般线性模型 ——易于求解

线性模型:

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$
$$N(\mathbf{0}, \sigma^2 \mathbf{I})$$

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x}$$

$$\hat{\boldsymbol{\theta}} \sim N\left(\boldsymbol{\theta}, \sigma^2 (\mathbf{H}^T \mathbf{H})^{-1}\right)$$

一般线性模型:

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{s} + \mathbf{w}$$
$$N(\mathbf{0}, \mathbf{C})$$

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} (\mathbf{x} - \mathbf{s})$$

$$\hat{\boldsymbol{\theta}} \sim N\left(\boldsymbol{\theta}, (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}\right)$$

- 妥协——最佳线性无偏估计 (BLUE) ——方便实用

- Best Linear Unbiased Estimator
- 假定估计量与观测数据间呈线性关系
- 全局来看 (包括线性、非线性估计), 并不一定最优
- 但仅需观测数据一、二阶矩, 无需PDF——很实用!
- 若数据高斯的, 则BLUE等效于MVU, 是有效的 (高斯-马尔可夫定理)