### 统计信号处理

第六章

# 最小二乘估计

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## 内容概要

- 一、最小二乘估计概念
- •二、线性最小二乘估计
- 三、序贯最小二乘估计
- 四、约束最小二乘估计
- 五、非线性最小二乘估计
- 六、小结

# 引言

- 最小方差无偏估计(MVU)
  - Cramer-Rao Lower Bound (CRLB)
  - (一般) 线性模型
  - Sufficient statistics (s.s.)
- 最大似然估计(MLE)
- 应用前提: 似然函数
- 实际应用中条件是否满足?
- 若无似然函数该当如何?
- 最小二乘估计(Least Squares)

### 一、最小二乘估计概念

已知信号模型 $s[n;\theta]$ , 观测数据为 $x[n]=s[n;\theta]+w[n]$ ,  $\theta$ 的最小二乘估计为

$$\hat{\theta} = \arg\min_{\theta} \left\{ \sum_{n=0}^{N-1} (x[n] - s[n; \theta])^2 \right\}$$

其中 $J(\theta) = \sum_{n=0}^{N-1} (x[n] - s[n;\theta])^2$  称为最小二乘误差

最小二乘 估计(LSE)

### **MVU**

$$\left\{ \min \left\{ \operatorname{var}(\hat{\theta}) \right\} \right.$$
s.t.  $E(\hat{\theta}) = \theta$ 

#### **MLE**

$$\hat{\theta} = \arg\max_{\theta} \left\{ p(\mathbf{x}; \theta) \right\}$$

例: 噪声中电平估计问题:

$$x[n] = A + w[n], n = 0,1,...,N-1$$

待估计参数为信号幅度A,w[n]为噪声。

$$J(A) = \sum_{n=0}^{N-1} (x[n] - A)^{2}$$

$$\frac{\partial J(A)}{\partial A} = -2\sum_{n=0}^{N-1} (x[n] - A)$$

$$\frac{\partial J(A)}{\partial A} = 0$$

$$\hat{A} = \frac{1}{N} \sum_{n=0}^{N-1} x[n]$$

### 二、线性最小二乘估计

若信号  $s = [s[0], s[1], s[2], ..., s[N-1]]^T$  与未知参数  $\theta = [\theta_1, \theta_2, \theta_3, ..., \theta_p]^T$  呈线性关系

$$s = H\theta$$

 $\mathbf{H}$ 是  $N \times p(N > p)$  的观测矩阵,秩为 p。观测数据模型为  $\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$ 

求 $\theta$ 的最小二乘估计?

$$J(\theta) = \sum_{n=0}^{N-1} (x[n] - s[n; \theta])^{2}$$

$$= (x - s)^{T} (x - s)$$

$$= (x - H\theta)^{T} (x - H\theta)$$

$$= x^{T} x - 2x^{T} H\theta + \theta^{T} H^{T} H\theta$$

$$\frac{\partial J(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = -2\mathbf{H}^T \boldsymbol{x} + 2\mathbf{H}^T \mathbf{H} \boldsymbol{\theta}$$
$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \boldsymbol{x}$$

相应的最小LS误差为:

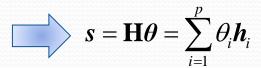
$$J_{\min} = \boldsymbol{x}^T \left( \boldsymbol{x} - \mathbf{H} \hat{\boldsymbol{\theta}} \right)$$

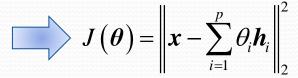
### • 几何解释

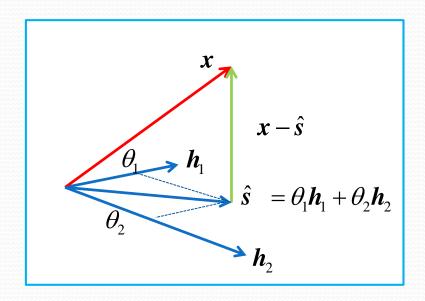
$$J(\boldsymbol{\theta}) = (\boldsymbol{x} - \mathbf{H}\boldsymbol{\theta})^{T} (\boldsymbol{x} - \mathbf{H}\boldsymbol{\theta})$$
$$= \|\boldsymbol{x} - \mathbf{H}\boldsymbol{\theta}\|_{2}^{2}$$

观测矩阵按列向量可记为:

$$\mathbf{H} = \begin{bmatrix} \mathbf{h}_1, \mathbf{h}_2, ..., \mathbf{h}_p \end{bmatrix}$$
 —信号矢量







-观测数据到信号的距离的平方

欧式距离: 
$$\|\xi\| = \sqrt{\sum_{i=1}^{N} \xi_i^2}$$

$$\min \{J(\boldsymbol{\theta})\} \qquad (\boldsymbol{x} - \hat{\boldsymbol{s}}) \perp \boldsymbol{h}_i, i = 1, 2, ..., p \qquad (\boldsymbol{x} - \hat{\boldsymbol{s}})^T \boldsymbol{h}_i = 0 \qquad (\boldsymbol{x} - \hat{\boldsymbol{s}})^T \boldsymbol{H} = \boldsymbol{0}^T$$

$$(\mathbf{x} - \hat{\mathbf{s}})^T \mathbf{H} = \mathbf{0}^T$$

$$\hat{\mathbf{s}} = \mathbf{H}\hat{\boldsymbol{\theta}}$$

$$(\mathbf{x} - \mathbf{H}\hat{\boldsymbol{\theta}})^T \mathbf{H} = \mathbf{0}^T$$

$$\mathbf{x}^T \mathbf{H} - \hat{\boldsymbol{\theta}}^T \mathbf{H}^T \mathbf{H} = \mathbf{0}^T$$

$$\mathbf{H}^{T}\mathbf{x} - \mathbf{H}^{T}\mathbf{H}\hat{\boldsymbol{\theta}} = \mathbf{0} \qquad \qquad \hat{\boldsymbol{\theta}} = \left(\mathbf{H}^{T}\mathbf{H}\right)^{-1}\mathbf{H}^{T}\mathbf{x} \qquad \text{LSE估计量}$$

其中

$$\varepsilon = x - H\hat{\theta}$$
 — 称为误差矢量

$$\varepsilon^T \mathbf{H} = \mathbf{0}^T$$
 — 表示误差矢量与信号矢量是正交的!

——称为正交原理

若列矢量相互正交:

$$\mathbf{h}_{i}^{T}\mathbf{h}_{j} = \delta_{ij}$$
  $\mathbf{H}^{T}\mathbf{H} = \mathbf{I}$   $\hat{\boldsymbol{\theta}} = (\mathbf{H}^{T}\mathbf{H})^{-1}\mathbf{H}^{T}x$   $\hat{\boldsymbol{\theta}} = \mathbf{H}^{T}x$  ——合适的信号向量将 大大简化LSE求解

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### • 加权最小二乘

$$J(\theta) = (x - H\theta)^{T} \mathbf{W}(x - H\theta) \quad (权系数 \mathbf{W} - \mathbf{H}\theta) \text{为对称矩阵})$$
$$= x^{T} \mathbf{W}x - x^{T} \mathbf{W} \mathbf{H}\theta - \theta^{T} \mathbf{H}^{T} \mathbf{W}x + \theta^{T} \mathbf{H}^{T} \mathbf{W} \mathbf{H}\theta$$

$$\frac{\partial J(\theta)}{\partial \theta} = -\mathbf{H}^T \mathbf{W}^T x - \mathbf{H}^T \mathbf{W} x + 2\mathbf{H}^T \mathbf{W} \mathbf{H} \theta$$

$$\frac{\partial J(\theta)}{\partial \theta} = \mathbf{0}$$

$$\hat{\theta} = (\mathbf{H}^T \mathbf{W} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{W} x$$

相应的最小LS误差为:  $J_{\min} = x^T \Big( \mathbf{W} - \mathbf{W} \mathbf{H} \Big( \mathbf{H}^T \mathbf{W} \mathbf{H} \Big)^{-1} \mathbf{H}^T \mathbf{W} \Big) x$ 

若 
$$\mathbf{W} = \mathbf{C}^{-1}$$
?
$$\hat{\boldsymbol{\theta}} = \left(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H}\right)^{-1} \mathbf{H}^T \mathbf{C}^{-1} \boldsymbol{x}$$

例: 不相关噪声中电平估计问题:

$$x[n] = A + w[n], n = 0,1,...,N-1$$

待估计参数为A, w[n]为不相关噪声且  $var(w[n]) = \sigma_n^2$ , 其加权LSE?

目标函数: 
$$J(A) = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} (x[n] - A)^2$$



$$\hat{\boldsymbol{\theta}} = \left(\mathbf{H}^T \mathbf{W} \mathbf{H}\right)^{-1} \mathbf{H}^T \mathbf{W} \mathbf{x}$$

参数: 
$$\theta = ? A$$

$$\mathbf{H} = ? [1,1,...,1]^T$$

$$\mathbf{W} = ? \begin{bmatrix} \frac{1}{\sigma_0^2} & 0 & \cdots & 0 \\ 0 & \frac{1}{\sigma_1^2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \frac{1}{\sigma_{N-1}^2} \end{bmatrix}$$

$$\hat{A} = \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} x[n]}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}}$$

### 三、序贯最小二乘估计

例: 不相关噪声中电平估计问题:

$$x[n] = A + w[n], n = 0,1,...,N-1$$

待估计参数为A, w[n]为不相关噪声且  $var(w[n]) = \sigma_n^2$ 。

加权LSE: 
$$J(A) = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} (x[n] - A)^2$$

$$\hat{A} = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} x[n] / \left( \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right)$$

$$\int \hat{A}[N-1] = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} x[n] / \left(\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}\right)$$

$$\hat{A}[N] = \sum_{n=0}^{N} \frac{1}{\sigma_n^2} x[n] / \left(\sum_{n=0}^{N} \frac{1}{\sigma_n^2}\right)$$
 贯计算?

$$\hat{A}[N] = \left\{ \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} x[n] + \frac{1}{\sigma_N^2} x[N] \right\} / \sum_{n=0}^{N} \frac{1}{\sigma_n^2}$$

$$\hat{A}[N] = \left\{ \left( \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right) \hat{A}[N-1] + \frac{1}{\sigma_N^2} x[N] \right\} / \left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)$$

$$\hat{A}[N] = \left\{ \left( \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} + \frac{1}{\sigma_N^2} - \frac{1}{\sigma_N^2} \right) \hat{A}[N-1] + \frac{1}{\sigma_N^2} x[N] \right\} / \left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)$$

$$\hat{A}[N] = \hat{A}[N-1] + \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^N \frac{1}{\sigma_n^2}} \left( x[N] - \hat{A}[N-1] \right)$$

$$K[N] : 增益因子$$

$$\hat{A}[N] = \hat{A}[N-1] + K[N](x[N] - \hat{A}[N-1])$$

• 估计量可以序贯计算

### LSE的方差

$$\hat{A}[N-1] = \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} x[n]}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}}$$

$$\operatorname{var}(\hat{A}[N-1]) = \frac{\sum_{n=0}^{N-1} \left(\frac{1}{\sigma_n^2}\right)^2 \operatorname{var}(x[n])}{\left(\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}\right)^2} = \frac{1}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}}$$



新的估计量: 
$$\hat{A}[N] = \frac{\sum_{n=0}^{N} \frac{1}{\sigma_n^2} x[n]}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}}$$

对应的方差为:

$$\operatorname{var}(\hat{A}[N]) = \frac{1}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}}$$
 能否序贯  
计算?

$$\operatorname{var}(\hat{A}[N]) = \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \times \frac{1}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}}}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}} = \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \times \operatorname{var}(\hat{A}[N-1])}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}}$$

$$= \frac{\left(\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} + \frac{1}{\sigma_N^2} - \frac{1}{\sigma_N^2}\right) \times \operatorname{var}(\hat{A}[N-1])}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}}$$

$$= \frac{1 - \frac{1}{\sigma_N^2}}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}} \operatorname{var}(\hat{A}[N-1])$$

$$\operatorname{var}(\hat{A}[N]) = (1 - K[N])\operatorname{var}(\hat{A}[N-1])$$

- 估计量的方差可以序贯计算
- 估计量的方差在不断减小

### 最小LS误差

$$J_{\min} = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( x[n] - \hat{A} \right)^2$$

$$= \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( x[n] - \hat{A}[N-1] \right)^2$$

$$J_{\min} \left[ N \right] = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( x[n] - \hat{A}[N] \right)^2$$

$$J_{\min} \left[ N \right] = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( x[n] - \hat{A}[N] \right)^2 + \frac{1}{\sigma_N^2} \left( x[N] - \hat{A}[N] \right)^2$$

$$= \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( x[n] - \hat{A}[N-1] + \hat{A}[N-1] - \hat{A}[N] \right)^2 + \frac{1}{\sigma_N^2} \left( x[N] - \hat{A}[N] \right)^2$$

$$= \sum_{n=0}^{N-1} \left\{ \frac{1}{\sigma_n^2} \left( x[n] - \hat{A}[N-1] \right)^2 + \frac{1}{\sigma_n^2} \left( \hat{A}[N-1] - \hat{A}[N] \right)^2 + \frac{1}{\sigma_N^2} \left( x[N] - \hat{A}[N] \right)^2 + \frac{1}{\sigma_N^2} \left( x[N] - \hat{A}[N] \right)^2 \right\}$$

$$\hat{A}[N] = \hat{A}[N-1] + K[N](x[N] - \hat{A}[N-1])$$

$$= J_{\min}[N-1] + \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} K^2[N] (x[N] - \hat{A}[N-1])^2 + \frac{1}{\sigma_N^2} (1 - K[N])^2 (x[N] - \hat{A}[N-1])^2$$
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$$J_{\min}[N] = J_{\min}[N-1] + \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} K^2[N] \left(x[N] - \hat{A}[N-1]\right)^2 + \frac{1}{\sigma_N^2} \left(1 - K[N]\right)^2 \left(x[N] - \hat{A}[N-1]\right)^2$$

$$\left\{ \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} K^2 [N] + \frac{1}{\sigma_N^2} (1 - K[N])^2 \right\} (x[N] - \hat{A}[N-1])^2 \\
= \left\{ \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}} \right)^2 + \frac{1}{\sigma_N^2} \left( 1 - \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}} \right)^2 \right\} (x[N] - \hat{A}[N-1])^2 \\
= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( \frac{1}{\sigma_N^2} \right)^2 + \frac{1}{\sigma_N^2} \left( \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right)^2}{\left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)^2} (x[N] - \hat{A}[N-1])^2 \\
= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( \frac{1}{\sigma_N^2} \right) \left\{ \frac{1}{\sigma_N^2} + \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right\}}{\left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)^2} (x[N] - \hat{A}[N-1])^2 \\
= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( \frac{1}{\sigma_N^2} \right) \left\{ \frac{1}{\sigma_N^2} + \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right\}}{\left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)^2} (x[N] - \hat{A}[N-1])^2 \\
= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left( \frac{1}{\sigma_N^2} \right) \left\{ \frac{1}{\sigma_N^2} + \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \right\}}{\left( \sum_{n=0}^{N} \frac{1}{\sigma_n^2} \right)^2} (x[N] - \hat{A}[N-1])^2$$

$$= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left(\frac{1}{\sigma_N^2}\right)}{\left(\sum_{n=0}^{N} \frac{1}{\sigma_n^2}\right)} \left(x[N] - \hat{A}[N-1]\right)^2$$

$$= \frac{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} \left(\frac{1}{\sigma_N^2}\right)}{\left(\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} + \frac{1}{\sigma_N^2}\right)} \left(x[N] - \hat{A}[N-1]\right)^2$$

$$= \frac{1}{\left(\sigma_N^2 + \frac{1}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2}}\right)} \left(x[N] - \hat{A}[N-1]\right)^2$$

$$= \frac{1}{\left(\sigma_N^2 + \text{var}(\hat{A}[N-1])\right)} \left(x[N] - \hat{A}[N-1]\right)^2$$

$$J_{\min}[N] = J_{\min}[N-1] + \frac{1}{\text{var}(\hat{A}[N-1]) + \sigma_N^2} (x[N] - \hat{A}[N-1])^2$$

$$= J_{\min} \left[ N - 1 \right] + \frac{\sigma_N^2}{\operatorname{var} \left( \hat{A} \left[ N - 1 \right] \right) + \sigma_N^2} \frac{\left( x \left[ N \right] - \hat{A} \left[ N - 1 \right] \right)^2}{\sigma_N^2}$$

$$K[N] = \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^{N} \frac{1}{\sigma_n^2}} = \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} + \frac{1}{\sigma_N^2}} = \frac{\frac{1}{\sigma_N^2}}{\frac{1}{\text{var}(\hat{A}[N-1])} + \frac{1}{\sigma_N^2}} = \frac{\text{var}(\hat{A}[N-1])}{\text{var}(\hat{A}[N-1]) + \sigma_N^2}$$



$$J_{\min}[N] = J_{\min}[N-1] + (1-K[N]) \frac{\left(x[N] - \hat{A}[N-1]\right)^{2}}{\sigma_{N}^{2}}$$

- 最小LS误差可以序贯计算
- 最小LS误差在不断增加

—因要拟合的数据在增加

例: 不相关噪声中电平估计问题:

$$x[n] = A + w[n], n = 0,1,...,N-1$$

待估计参数为A, w[n]为不相关噪声且  $var(w[n]) = \sigma_n^2$ 。

加权LSE: 
$$J(A) = \sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} (x[n] - A)^2$$

估计量更新:  $\hat{A}[N] = \hat{A}[N-1] + K[N](x[N] - \hat{A}[N-1])$ 

增益因子: 
$$K[N] = \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^N \frac{1}{\sigma_n^2}} = \frac{\frac{1}{\sigma_N^2}}{\sum_{n=0}^{N-1} \frac{1}{\sigma_n^2} + \frac{1}{\sigma_N^2}} = \frac{\frac{1}{\sigma_N^2}}{\frac{1}{\operatorname{var}(\hat{A}[N-1])} + \frac{1}{\sigma_N^2}}$$
 总的信息量

LSE方差更新:  $\operatorname{var}(\hat{A}[N]) = (1 - K[N]) \operatorname{var}(\hat{A}[N-1])$ 

最小LS误差更新: 
$$J_{\min}[N] = J_{\min}[N-1] + \left(1 - \underline{K[N]}\right) \frac{\left(x[N] - \hat{A}[N-1]\right)^2}{\sigma_N^2}$$

初始化: 
$$\hat{A}[0] = x[0]$$
  $\operatorname{var}(\hat{A}[0]) = \sigma_0^2$ 

### ● 推广至矢量参数情况

信号模型:  $x = \mathbf{H}\theta + \mathbf{w}$ 

w 为噪声, 其协方差矩阵为  $C_w$ 。

A. Giordano and Frank M. Hsu. Least square estimation with application to digital signal processing. Wiley, New York, 1985

加权最小LS误差: 
$$J(\theta) = (x - H\theta)^T C_w^{-1} (x - H\theta)$$

加权LSE:  $\hat{\boldsymbol{\theta}} = \left(\mathbf{H}^T \mathbf{C}_w^{-1} \mathbf{H}\right)^{-1} \mathbf{H}^T \mathbf{C}_w^{-1} \boldsymbol{x}$ 

LSE的协方差阵:  $\mathbf{C}_{\hat{\theta}} = \left(\mathbf{H}^T \mathbf{C}_{w}^{-1} \mathbf{H}\right)^{-1}$ 

若 $\mathbf{C}_{\mathbf{w}}$ 是对角矩阵,即噪声是不相关的,则可按如下方式序贯计算:

估计量更新: 
$$\hat{\theta}[n] = \hat{\theta}[n-1] + k[n](x[n] - h^T[n]\hat{\theta}[n-1])$$

增益因子: 
$$k[n] = \frac{\mathbf{C}_{\hat{\theta}}[n-1]h[n]}{\sigma_n^2 + h^T[n]\mathbf{C}_{\hat{\theta}}[n-1]h[n]}$$

协方差更新: 
$$\mathbf{C}_{\hat{\theta}}[n] = (\mathbf{I} - \mathbf{k}[n]\mathbf{h}^T[n])\mathbf{C}_{\hat{\theta}}[n-1]$$

$$\mathbf{H}[n] = \begin{bmatrix} \mathbf{H}[n-1] \\ \boldsymbol{h}^{T}[n] \end{bmatrix}$$
$$x[n] = \boldsymbol{h}^{T}[n]\boldsymbol{\theta} + w[n]$$

最小LS误差更新: 
$$J_{\min}[n] = J_{\min}[n-1] + \frac{\left(x[n] - \boldsymbol{h}^T[n]\hat{\boldsymbol{\theta}}[n-1]\right)^2}{\sigma_n^2 + \boldsymbol{h}^T[n]\mathbf{C}_{\hat{\boldsymbol{\theta}}}[n-1]\boldsymbol{h}[n]}$$

# 四、约束最小二乘估计

信号模型:

$$x = H\theta + w$$

w 为噪声, 其协方差矩阵为 C。待估计参数需满足如下约束:

$$\mathbf{A}\boldsymbol{\theta} = \boldsymbol{b}$$

$$\begin{cases}
\min_{\theta} \left\{ (\mathbf{x} - \mathbf{H}\boldsymbol{\theta})^T (\mathbf{x} - \mathbf{H}\boldsymbol{\theta}) \right\} \\
s.t. \quad \mathbf{A}\boldsymbol{\theta} = \boldsymbol{b}
\end{cases}$$

$$J_{c} = (\mathbf{x} - \mathbf{H}\boldsymbol{\theta})^{T} (\mathbf{x} - \mathbf{H}\boldsymbol{\theta}) + \lambda^{T} (\mathbf{A}\boldsymbol{\theta} - \mathbf{b})$$

$$\frac{\partial J_{c}}{\partial \boldsymbol{\theta}} = \mathbf{0}$$

$$\hat{\boldsymbol{\theta}}_{c} = (\mathbf{H}^{T}\mathbf{H})^{-1} \mathbf{H}^{T} \mathbf{x} - \frac{1}{2} (\mathbf{H}^{T}\mathbf{H})^{-1} \mathbf{A}^{T} \lambda$$

$$\hat{\boldsymbol{\theta}}_{c} = \hat{\boldsymbol{\theta}} - (\mathbf{H}^{T}\mathbf{H})^{-1} \mathbf{A}^{T} \frac{\lambda}{2}$$

约束条件: 
$$\mathbf{A}\hat{\boldsymbol{\theta}}_{c} = \mathbf{b}$$

$$\frac{\lambda}{2} = \left[ \mathbf{A} \left( \mathbf{H}^T \mathbf{H} \right)^{-1} \mathbf{A}^T \right]^{-1} \left( \mathbf{A} \hat{\boldsymbol{\theta}} - \boldsymbol{b} \right)$$

$$\hat{\boldsymbol{\theta}}_c = \hat{\boldsymbol{\theta}} - \left(\mathbf{H}^T \mathbf{H}\right)^{-1} \mathbf{A}^T \left[ \mathbf{A} \left(\mathbf{H}^T \mathbf{H}\right)^{-1} \mathbf{A}^T \right]^{-1} \left( \mathbf{A} \hat{\boldsymbol{\theta}} - \boldsymbol{b} \right)$$

• 在无约束LSE的基础上加一"修正项"

例: 信号模型:

$$x[n] = s[n] + w[n]$$

其中,信号为 
$$s[n] = \begin{cases} \theta_1, & n = 0 \\ \theta_2, & n = 1 \end{cases}$$
 , 观测数据为  $\{x[0], x[1], x[2]\}$ 。  $\{0, n = 2\}$ 

若已知  $\theta_1 = \theta_2$ , 对参数的估计为?

待估计参数:  $\boldsymbol{\theta} = [\theta_1, \theta_2]^T$ 

### 第一种方法: 无约束LSE

$$\hat{\boldsymbol{\theta}} = \left(\mathbf{H}^T \mathbf{H}\right)^{-1} \mathbf{H}^T \boldsymbol{x}$$

$$\mathbf{H} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 0 & 0 \end{bmatrix}$$

$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} x[0] \\ x[1] \end{bmatrix}$$

例:信号模型:

$$x[n] = s[n] + w[n]$$

其中,信号为 
$$s[n] = \begin{cases} \theta_1, & n = 0 \\ \theta_2, & n = 1 \end{cases}$$
,观测数据为  $\{x[0], x[1], x[2]\}$ 。  $\{0, n = 2\}$ 

若已知  $\theta_1 = \theta_2$ , 对参数的估计为?

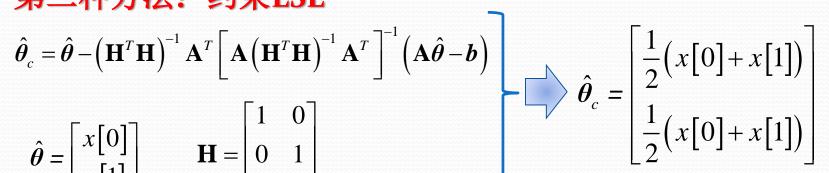
待估计参数:  $\theta = [\theta_1, \theta_2]^T$ 

### 第二种方法:约束LSE

$$\hat{\boldsymbol{\theta}}_{c} = \hat{\boldsymbol{\theta}} - \left(\mathbf{H}^{T}\mathbf{H}\right)^{-1}\mathbf{A}^{T} \left[\mathbf{A}\left(\mathbf{H}^{T}\mathbf{H}\right)^{-1}\mathbf{A}^{T}\right]^{-1} \left(\mathbf{A}\hat{\boldsymbol{\theta}} - \boldsymbol{b}\right)$$

$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} x[0] \\ x[1] \end{bmatrix} \qquad \mathbf{H} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 0 & 0 \end{bmatrix}$$

约束
$$\mathbf{A}\boldsymbol{\theta} = \boldsymbol{b}$$
 [1,-1] $\boldsymbol{\theta} = 0$ 



# 五、非线性最小二乘估计

已知信号模型 $s[n;\theta]$ ,观测数据为 $x[n]=s[n;\theta]+w[n]$ , $\theta$ 的最小二乘估计为

$$\hat{\theta} = \arg\min_{\theta} J(\theta)$$

$$J(\theta) = \sum_{n=0}^{N-1} (x[n] - s[n; \theta])^{2}$$

$$J(\theta) = (x - s(\theta))^T (x - s(\theta))$$
 非线性问题!

#### 1. 网格搜索法

一般适用于待搜索维数较小者

### 2. 参数变换法

$$J(\theta) = (x - s(\theta))^{T} (x - s(\theta))$$

#### 核心思想:将非线性参数转换为线性参数来求解

对该估计参数 $\theta$ , 进行某一对一变换:

$$\alpha = g(\theta)$$

使得非线性信号模型能转换为线性信号模型:

$$s(\theta) = \mathbf{H}\alpha$$

$$x = \mathbf{H}\alpha + \mathbf{w}$$

若α的LSE为:

$$\hat{\boldsymbol{\alpha}} = \left(\mathbf{H}^T \mathbf{H}\right)^{-1} \mathbf{H}^T \boldsymbol{x}$$

则待估计参数 $\theta$ 的LSE可由下式求出:

$$\hat{\boldsymbol{\theta}} = \boldsymbol{g}^{-1} \left( \hat{\boldsymbol{\alpha}} \right)$$

——LSE的不变性!

信号模型:

$$x[n] = A\cos(2\pi f_0 n + \phi) + w[n]$$

 $f_0$ 已知,待估计参数为信号幅度 (A>0) 和载波相位  $\phi$ 。

$$J = \sum_{n=0}^{N-1} (x[n] - A\cos(2\pi f_0 n + \phi))^2$$
 非线性问题!

$$A\cos(2\pi f_0 n + \phi) = \underline{A\cos(\phi)\cos(2\pi f_0 n)} - \underline{A\sin(\phi)\sin(2\pi f_0 n)} \alpha_1$$

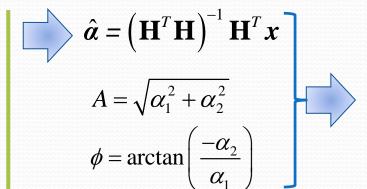
$$s[n] = \alpha_1 \cos(2\pi f_0 n) + \alpha_2 \sin(2\pi f_0 n)$$

$$\mathbf{s} = \begin{bmatrix} 1 & 0 \\ \cos(2\pi f_0) & \sin(2\pi f_0) \\ \vdots & \vdots \\ \cos(2\pi f_0(N-1)) & \sin(2\pi f_0(N-1)) \end{bmatrix} \underline{\begin{bmatrix} \alpha_1 \\ \alpha_2 \end{bmatrix}} \qquad \mathbf{\alpha}$$

$$A = \sqrt{\alpha_1^2 + \alpha_2^2} \qquad \phi = \arctan\left(\frac{-\alpha_2}{\alpha_1}\right)$$

H

$$x = \mathbf{H}\alpha + \mathbf{w}$$



$$\hat{\boldsymbol{\theta}} = \begin{bmatrix} \hat{A} \\ \hat{\phi} \end{bmatrix} = \begin{bmatrix} \sqrt{\hat{\alpha}_1^2 + \hat{\alpha}_2^2} \\ \arctan\left(\frac{-\hat{\alpha}_2}{\hat{\alpha}_1}\right) \end{bmatrix}$$

### 3. 参数分离法

$$J(\theta) = (x - s(\theta))^{T} (x - s(\theta))$$

### 核心思想:将非线性参数尽量转换为线性参数,以减小复杂度

尽量将信号变换为如下形式

$$s(\theta) = \mathbf{H}(\alpha)\beta$$

1. 参数分离

其中待估计参数为  $\theta = \begin{bmatrix} \alpha \\ \beta \end{bmatrix}$ 。此时,LS误差为

$$J(\alpha, \beta) = (x - \mathbf{H}(\alpha)\beta)^{T} (x - \mathbf{H}(\alpha)\beta)$$

对给定的  $\alpha$ , 使LS误差最小的  $\beta$  为

$$\hat{\boldsymbol{\beta}} = \left(\mathbf{H}^{T}(\boldsymbol{\alpha})\mathbf{H}(\boldsymbol{\alpha})\right)^{-1}\mathbf{H}^{T}(\boldsymbol{\alpha})\boldsymbol{x}$$

- 2. 线性参数非线性表示
- 4. 线性参数求解

此时,LS误差为

$$J(\boldsymbol{\alpha}, \hat{\boldsymbol{\beta}}) = \boldsymbol{x}^{T} \left( \mathbf{I} - \mathbf{H}(\boldsymbol{\alpha}) (\mathbf{H}^{T}(\boldsymbol{\alpha}) \mathbf{H}(\boldsymbol{\alpha}))^{-1} \mathbf{H}^{T}(\boldsymbol{\alpha}) \right) \boldsymbol{x}$$

此时,  $\alpha$ 的LSE为

### 4. Gauss-Newton迭代法

$$J(\theta) = \sum_{n=0}^{N-1} (x[n] - s[n;\theta])^2$$

### 核心思想: 非线性问题线性化后迭代求解

某个标称的θ 附近对非线性函数进行线性化

$$s[n;\theta] \approx s[n;\theta_0] + \frac{\partial s[n;\theta]}{\partial \theta} \bigg|_{\theta=\theta_0} (\theta - \theta_0)$$

此时,LS误差为

$$J(\theta) = \sum_{n=0}^{N-1} (x[n] - s[n;\theta])^{2}$$

$$\approx \sum_{n=0}^{N-1} \left[ x[n] - s[n;\theta_{0}] + \frac{\partial s[n;\theta]}{\partial \theta} \Big|_{\theta=\theta_{0}} \theta_{0} - \frac{\partial s[n;\theta]}{\partial \theta} \Big|_{\theta=\theta_{0}} \theta \right]^{2}$$

$$= \left( x - s(\theta_{0}) + \mathbf{H}(\theta_{0}) \theta_{0} - \mathbf{H}(\theta_{0}) \theta \right)^{T} \left( x - s(\theta_{0}) + \mathbf{H}(\theta_{0}) \theta_{0} - \mathbf{H}(\theta_{0}) \theta \right)$$

其中, 
$$\left[\mathbf{H}(\theta)\right]_{i} = \frac{\partial s\left[i;\theta\right]}{\partial \theta}$$
 清华大学电子工程系 李洪 副教授

### LSE为

$$\hat{\theta} = (\mathbf{H}^{T}(\theta_{0})\mathbf{H}(\theta_{0}))^{-1}\mathbf{H}^{T}(\theta_{0})(\mathbf{x} - \mathbf{s}(\theta_{0}) + \mathbf{H}(\theta_{0})\theta_{0})$$

$$= \theta_{0} + (\mathbf{H}^{T}(\theta_{0})\mathbf{H}(\theta_{0}))^{-1}\mathbf{H}^{T}(\theta_{0})(\mathbf{x} - \mathbf{s}(\theta_{0}))$$

#### 其迭代解法:

$$\theta_{k+1} = \theta_k + \left(\mathbf{H}^T(\theta_k)\mathbf{H}(\theta_k)\right)^{-1}\mathbf{H}^T(\theta_k)(\mathbf{x} - \mathbf{s}(\theta_k))$$

进一步地,可推广至矢量参数时

$$\boldsymbol{\theta}_{k+1} = \boldsymbol{\theta}_k + \left(\mathbf{H}^T\left(\boldsymbol{\theta}_k\right)\mathbf{H}\left(\boldsymbol{\theta}_k\right)\right)^{-1}\mathbf{H}^T\left(\boldsymbol{\theta}_k\right)\left(\boldsymbol{x} - \boldsymbol{s}\left(\boldsymbol{\theta}_k\right)\right)$$

其中, 
$$\left[\mathbf{H}(\boldsymbol{\theta})\right]_{ij} = \frac{\partial s[i;\boldsymbol{\theta}]}{\partial \theta_j}$$

### • 存在收敛问题

Seber, G.A.F., Wild, C.J. Nonlinear Regression. J. Wiley, New York, 1989

其它类似方法: Newton-Raphson迭代法

清华大学电子工程系 李洪 副教授

### 5. 循环最小化方法

$$J(\theta) = (x - s(\theta))^{T} (x - s(\theta))$$

- 将待估计参数分为两部分:  $\theta = [\xi \ \varsigma]^T$  也可分为更多部分
- 给定 ς<sup>0</sup>
- 计算  $\xi^1 = \arg\min_{\xi} \left\{ J \left( \begin{bmatrix} \xi & \varsigma^0 \end{bmatrix}^T \right) \right\}$  和  $\varsigma^1 = \arg\min_{\varsigma} \left\{ J \left( \begin{bmatrix} \xi^1 & \varsigma \end{bmatrix}^T \right) \right\}$
- 计算  $\xi^2 = \arg\min_{\xi} \left\{ J(\begin{bmatrix} \xi & \varsigma^1 \end{bmatrix}^T) \right\}$  和  $\varsigma^2 = \arg\min_{\varsigma} \left\{ J(\begin{bmatrix} \xi^2 & \varsigma \end{bmatrix}^T) \right\}$
- ...
- 计算 $\xi^k = \arg\min_{\xi} \left\{ J([\xi \ \varsigma^{k-1}]^T) \right\}$  和  $\varsigma^k = \arg\min_{\varsigma} \left\{ J([\xi^k \ \varsigma]^T) \right\}$  直至收敛

### 6. 放松估计方法

例: 正弦信号参数估计

$$x[n] = \sum_{p=1}^{P} \alpha_p e^{j2\pi f_p n} + w[n], \quad n = 0, 1, 2, ...N - 1$$

其中信号幅度 $\alpha_p$ 和频率 $f_p$ 未知。

$$\mathbf{x} = \begin{bmatrix} x[0], x[1], ..., x[N-1] \end{bmatrix}^{T}$$

$$\mathbf{\alpha} = \begin{bmatrix} \alpha_{1}, \alpha_{1}, ..., \alpha_{p} \end{bmatrix}^{T}$$

$$\mathbf{f} = \begin{bmatrix} f_{1}, f_{1}, ..., f_{p} \end{bmatrix}^{T}$$

$$\mathbf{h}(f_{p}) = \begin{bmatrix} 1, e^{j2\pi f_{p}}, e^{j2\pi f_{p}^{2}} ..., e^{j2\pi f_{p}(N-1)} \end{bmatrix}^{T}$$

$$\mathbf{w} = \begin{bmatrix} w[0], w[1], ..., w[N-1] \end{bmatrix}^{T}$$

$$\mathbf{H}(f) = \begin{bmatrix} \mathbf{h}(f_{1}), \mathbf{h}(f_{2}), \mathbf{h}(f_{1}), ..., \mathbf{h}(f_{p}) \end{bmatrix}$$

$$x = \mathbf{H}(f)\alpha + w$$

$$\hat{\boldsymbol{\alpha}} = \left(\mathbf{H}^{H}\left(f\right)\mathbf{H}\left(f\right)\right)^{-1}\mathbf{H}^{H}\left(f\right)\boldsymbol{x}$$

$$\hat{\boldsymbol{\alpha}} = (\mathbf{H}^{H}(f)\mathbf{H}(f))^{-1}\mathbf{H}^{H}(f)x$$

$$\max_{f} \left\{ \boldsymbol{x}^{H}\mathbf{H}(f)(\mathbf{H}^{H}(f)\mathbf{H}(f))^{-1}\mathbf{H}^{H}(f)x \right\}$$
问题: 若  $\hat{f}_{i} \approx \hat{f}_{i}$ ?

### 将导致幅度估计性能较差

假定 $\{\hat{\alpha}_i, \hat{f}_i\}_{i=1,i\neq n}^r$ 已知或通过其它方式估计得到,那么可构建新 的观测数据

$$\mathbf{x}_{p} = \mathbf{x} - \sum_{i=1, i \neq p}^{P} \hat{\alpha}_{i} \mathbf{h} \left( \hat{f}_{i} \right) \qquad \dots \dots (1)$$



$$J(\alpha_p, f_p) = (\mathbf{x}_p - \alpha_p \mathbf{h}(f_p))^H (\mathbf{x}_p - \alpha_p \mathbf{h}(f_p))$$



$$\hat{f}_p = \arg \max_{f_p} \left| \frac{\boldsymbol{h}^H (f_p) \boldsymbol{x}_p}{N} \right|$$

- Li J, Stoica P, Efficient mixed-spectrum estimation with applications to target feature extraction, IEEE Transactions on signal processing, 1996, 44(2): 281-295
- ▶ 假定信号数为1
- 吴仁彪等,通用鲁棒的放松估计方法,科学出版社,2017年
- 利用观测数据  $oldsymbol{x}$  和式 (2) 估计得到  $\left\{\hat{lpha}_{\!\scriptscriptstyle 1},\hat{f}_{\!\scriptscriptstyle 1}\right\}$
- ▶ 假定信号数为2
  - 利用式 (1) 和估计得到的  $\{\hat{\alpha}_1,\hat{f}_1\}$  构建数据  $x_2$  。利用  $x_2$  和式 (2) 估计得到  $\{\hat{\alpha}_2,\hat{f}_2\}$
  - 利用式 (1) 和估计得到的  $\{\hat{\alpha}_2, \hat{f}_2\}$  构建数据  $x_1$ 。利用  $x_1$  和式 (2) 估计得到  $\{\hat{\alpha}_1, \hat{f}_1\}$
  - 重复上述步骤直至收敛 洪化
- 收敛: 残差小于门限或 迭代次数达到门限
  - ► 假定信号数为3,按类似方法反复估计直至收敛 .....
  - ➤ 假定信号数为 P, 按类似方法反复估计直至收敛

### 六、小结

- 最小二乘估计
  - 无需似然函数,与MVU、MLE不同
  - 与同样也不需要似然函数的BLUE也不同
  - 核心思想:使由待估计参数构建的信号 与观测数据"尽量接近" ——信号模型至关重要

体会各类估计理论与方法核心思想的差异

- 线性最小二乘估计
- 加权最小二乘估计
- 序贯最小二乘估计
- 约束最小二乘估计
- 非线性最小二乘估计
  - 网格搜索法、参数变换、参数分离法、迭代方法,等