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- 1 To obtain the marginal joint distribution from the multivariate normal random variables (T_i, ϵ_i) the irrelevant variables are dropped from the mean vector and the covariance matrix. Thus the joint distribution of (T_i, ϵ_i) is multivariate normal and can be expressed as

$$\begin{pmatrix} T_i \\ \epsilon_i \end{pmatrix} \sim N \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \right)$$

From this we can see that the two variables T_i and ϵ_i are uncorrelated as $\rho = 0$ in this case. As a result it is true that $T_i \perp \epsilon_i$. However, in the case of X_i it is unknown if that variable is correlated with either T_i or ϵ_i and so it is unknown if $T_i \perp X_i$. This does not change the relationship between T_i and ϵ_i .

2

$$2 \quad \begin{bmatrix} X_i \\ T_i \end{bmatrix} \sim N(0, \Omega) \quad \Omega = \begin{bmatrix} 1 & \rho_2 \\ \rho_2 & 1 \end{bmatrix}$$

$$X_i = a T_i + u_i$$

$$= \begin{bmatrix} E[X_i] & E[X_i \cdot T_i] \\ E[X_i \cdot T_i] & E[T_i^2] \end{bmatrix}$$

$$\rho_2 = E[X_i \cdot T_i]$$

$$E[T_i (a T_i + u_i)]$$

$$E[a T_i^2 + T_i^2 u_i]$$

$$E[a T_i^2] + E[T_i^2 u_i]$$

$$a E[T_i^2] + E[T_i^2] E[u_i]$$

$$a(1) + (1)(0)$$

$$\boxed{\rho_2 = a}$$

$$3 \quad Y_i = \gamma T_i + \beta(aT_i + u_i) + \epsilon_i$$

$$= \gamma T_i + a\beta T_i + \beta u_i + \epsilon_i$$

$$Y_i = T_i(\gamma + a\beta) + \beta u_i + \epsilon_i$$

$$E[Y_i] = E[T_i(\gamma + a\beta) + \beta u_i + \epsilon_i]$$

$$= E[(\gamma + a\beta)T_i] + E[\beta u_i] + E[\epsilon_i]$$

$$E[Y_i] = (\gamma + a\beta)E[T_i]$$

$$\boxed{E[Y_i | T_i] = (\gamma + \rho_2 \beta) T_i}$$

4 From above $E[Y_i | T_i = t_i] = (\alpha + \rho_2 \beta) t_i$

The regression coefficient $\beta = (\alpha + \rho_2 \beta)$ and only provides an unbiased estimate of α if the term $\rho_2 \beta = 0$ which can only be the case if $\rho_2 = 0$. We can see this condition as further evidence that the relationship between X_i and T_i is expressed through the coefficient ρ_2 .

5

$$\epsilon_i = bT_i + cX_i + v_i$$

$$5 \quad \text{Cov}(X_i, \epsilon_i) = \begin{pmatrix} 1 & \rho_1 \\ \rho_1 & 1 \end{pmatrix} = \begin{pmatrix} E[X_i^2] & E[X_i \cdot \epsilon_i] \\ E[\epsilon_i \cdot X_i] & E[\epsilon_i^2] \end{pmatrix}$$

$$\text{Cov}(T_i, \epsilon_i) = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} = \begin{pmatrix} E[T_i^2] & E[T_i \epsilon_i] \\ E[\epsilon_i T_i] & E[\epsilon_i^2] \end{pmatrix}$$

$$0 = E[T_i \cdot \epsilon_i] = E[T_i(bT_i + cX_i + v_i)]$$

$$= bE[T_i^2] + cE[T_i X_i] + E[v_i]E[T_i]$$

$$\boxed{0 = b + c(\rho_2)} \quad \leftarrow \text{From 2}$$

$$\rho_1 = E[X_i \cdot \epsilon_i] = E[X_i(bT_i + cX_i + v_i)]$$

$$= bE[T_i \cdot X_i] + cE[X_i^2] + E[X_i]E[v_i]$$

$$\boxed{\rho_1 = b\rho_2 + c}$$

$$6 \quad Y_i = \gamma T_i + \beta X_i + \epsilon_i = \gamma T_i + \beta X_i + bT_i + cX_i + v_i$$

$$E[Y_i] = E[\gamma T_i + \beta X_i + bT_i + cX_i + v_i]$$

$$= E[(\gamma + b)T_i] + E[(\beta + c)X_i] + E[v_i]$$

$$E[Y_i] = (\gamma + b)E[T_i] + (\beta + c)E[X_i]$$

$$E[Y_i | X_i = x, T_i = t] = (\gamma + b)t + (\beta + c)x$$

$$\blacksquare = (\gamma + b) \quad \blacklozenge = (\beta + c)$$

7 The regression coefficient $\hat{\alpha} = r + b$ identifies the true value of r only if the term b is equal to zero which from the previous equation in problem 5 we can see is only true when the two variables, X_i and T_i , are uncorrelated or the term ρ_2 is equal to zero.

8 This is not always true.

$$\hat{\alpha} = (r + \rho_2 \beta) \quad \hat{\alpha} = (r + b)$$

If $|\rho_2 \beta| < |b|$ then it is better to exclude the X_i as a condition