

STA610 Case Study 1 Team 4 Report

Cole Juracek (Programmer) Lauren Palazzo (Writer) Lingyu Zhou (Checker + Coordinator)
Fan Zhu (Presenter)

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Introduction

This study is an exploration of the relationship between the sale price per milligram of diverted pharmaceutical substances and other factors related to the sale, as reflected in self-reported data from the website StreetRx. StreetRx maintains a partnership with the Researched Abuse, Diversion, and Addiction-Related Surveillance System (RADARS), which aims for these data to inform public health measures such as health policy related to illicit drugs.

Our drug of interest in this study is oxymorphone. Oxymorphone is used to relieve moderate to severe pain in people whose pain is not controlled with other medications, and may be habit-forming with prolonged use. Our aim is to investigate factors related to the price per mg of oxymorphone, accounting for potential clustering by location and exploring heterogeneity in pricing by location.

Among the data available to us are location data pertaining to the city, state, and United States region of the sale. These variables present a natural grouping structure for the price data, which suggest that we employ a hierarchical model in our analysis. The benefits of a hierarchical model include the ability to “borrow information” from the data set as a whole to stabilize estimates of price for groups (e.g., states) with relatively small sample size, and to establish a estimated distribution of how price varies with respect to groups, which would in theory allow generalization to new groups not included in the original analysis.

Exploratory data analysis and data cleaning

Data cleaning

Group variables

We performed an initial examination of two grouping variables, `state` and `USA_region`. The state North Dakota had only one observation, and as such was not considered a well-defined “group” for which a variance could be estimated; thus this state’s observation was removed. States labeled simply as “USA” were recoded to “Unknown”. Region was also sometimes labeled as unknown or other; these observations coincided with the unknown states. (The variable `city` as a grouping variable is addressed later in the report.)

Date variable

We considered two different ways to engineer the date-of-sale feature. One intuitive and commonly-used approach would be to recode this as a continuous variables counting days from some start date, the intent of which would be to capture the influence of inflation for this price data that spans years. Another approach would be to consider whether there might be an effect of seasonality on drug prices.

For recoding date into a continuous variable, `date_num`, we filter out observations dated prior to the year 2010, as that is the year that the StreetRx web site was founded. For the seasonality variable, `date_quarter`, we extract the “quarter” of the year (numbered 1 through 4) as its own variable. Please see appendix for distributions.

Sources variable

As the variable `source` that describes the source of the price information is a user-input field, it is an inherently “messy” variable that calls for some recoding in order to make the levels of the variable more meaningful and interpretable, and to avoid overfitting. In the case of `source`, some user input is simply blank, while a small number of other entries indicate in some other way that the source is unknown or unclear. Since it is quite uncommon in the data for a source to be entered yet indicated as unknown, we decided to drop these unknown-source observations as being indicative of poor data quality. (If the user cannot identify their source of price information, why should we trust their price data?) We decided to distinguish between “Internet pharmacy” sources and other internet sources, since there may be a distinction between a source that is selling a drug versus perhaps just discussing a drug price in a forum, which may yield less legitimate information. Ultimately, we recoded the categorical variable “sources” into the following levels.

Table 1: Frequency Table of Variable `source`

Level	Freq
Blank	1510
Internet	162
Internet Pharmacy	51
Personal	1611
Heard it	875

Bulk purchase variable

Our visualization of the bulk purchase variable indicates nothing unexpected that calls for data cleaning.

Purchase reason variable

`Primary_Reason` for purchase is another user-input variable that calls for recoding. Similarly to the `source` variable, we retain the information about which entries were blank. We also try to distinguish this from a different response level that is not blank, yet indicates the user did not provide a reason. We further still distinguished between these responses and a non-blank response of refusal to answer. We strove to avoid making presumptions of no difference between these similar-seeming response types, as they may indicate variations in data processing that hold some predictive power that we are not yet aware of. We ultimately recoded the categorical variable `Primary_Reason` into the following levels.

Table 2: Frequency Table of Variable `Primary_Reason`

Level	Freq
Blank	2473
Left unanswered	780
Other or unknown	97
Self-treat	427
Enjoyment	104
Resell	18
Refuse to answer	310

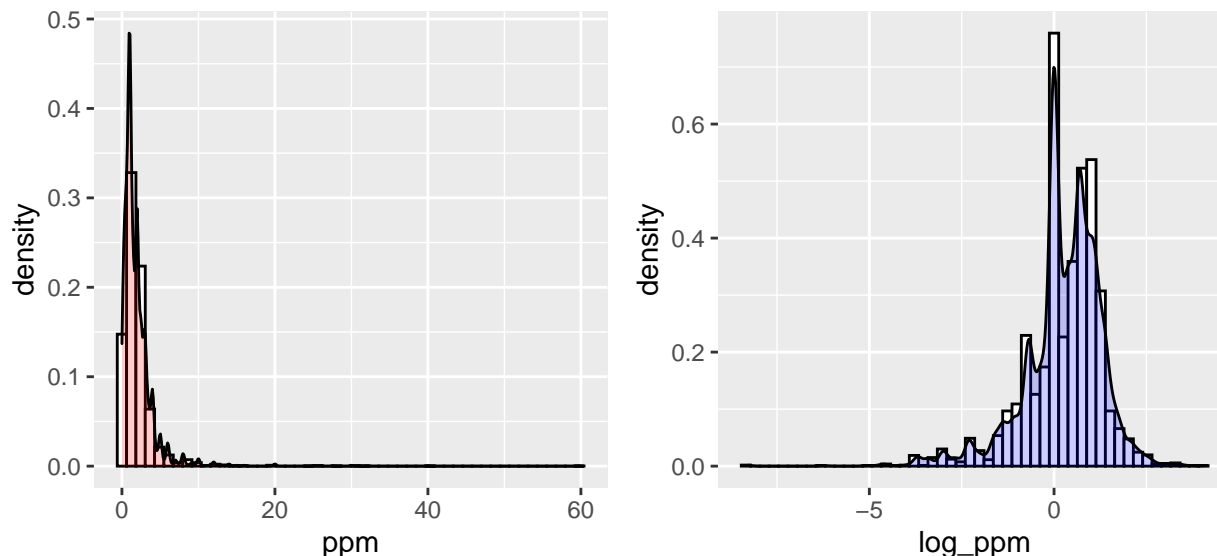
Drug strength variable

We renamed this variable `dose_per_mg` to make its meaning clearer to the reader. This is a continuous variable that nevertheless tends to take on only a few values (please see appendix), as depicted in the histogram. We considered a recoding of this variable into levels, but decided against it as the dose per milligram technically

may take on any value and the generalizability of our model may be improved if we are able to consider dose levels that are in a realistic range yet not present in our current data.

Response variable: Price per milligram

We are considering fitting a hierarchical model with a response variable that has normally distributed errors, and so we would like our response variable, here `ppm`, or price per milligram, to be close to normally-distributed. The histogram of the `ppm` is extremely right skewed (not unexpected for data related to financial measures such as price, salary, earnings, etc.), while the distribution of the `log(ppm)` looks roughly normal. Thus, we will use `log_ppm` as the dependent variable in our model.



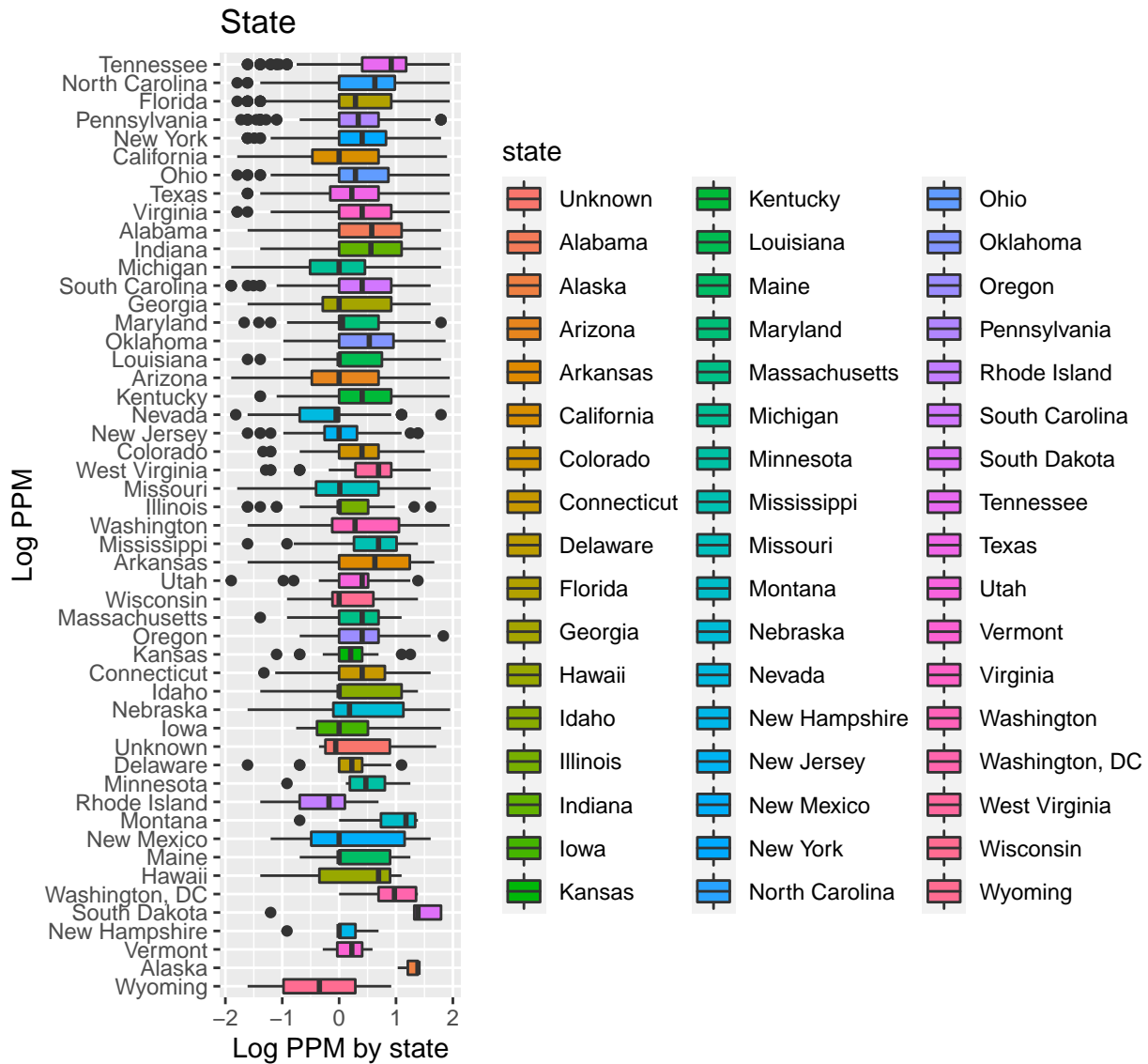
In order to cut off extreme price values that may skew the results for typical prices, we restrict the `log_ppm` distribution to more realistic and representative values:

Exploratory data analysis for random intercepts

We now take a closer look at our grouping variables with a focus on the question of whether random intercepts might be appropriate.

State

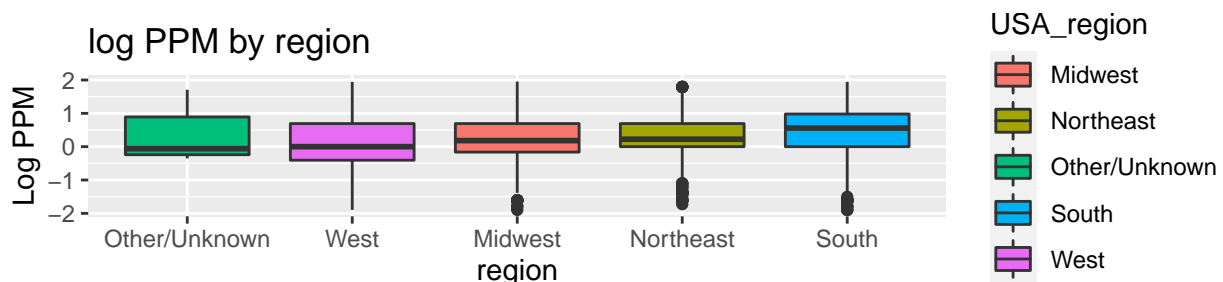
We show boxplots of `log_ppm` by `state`, arranged in the order of increasing sample size from the bottom to the top.



We see that some states with small sample sizes have distributions of `log_price` that seem unstable, with respect to how much they differ from other boxplots; this provides support for the idea that the stabilizing effect of a hierarchical model would benefit our analysis. The fact that some distributions differ from each other in a noteworthy way (e.g., boxplots hardly overlapping) suggests group-based differences in price that should be taken into account in modeling, as opposed to lumping all states' data together without distinction.

Region

Viewing a similar plot for the regions of the USA, we do see some variability in price among regions, but it is not as pronounced as for the states.



City

We find that there are many cities with a sample size of one, which makes `city` perhaps not the most natural choice for a grouping variable (please see appendix). We decided to focus on `state` as our primary candidate for a grouping variable moving forward.

Exploratory data analysis for random slopes

To consider whether we may want to include random slopes by state for our predictor variables, we examined plots of the predictor versus response among each state individually to determine whether it seemed that slopes fit to the group-level data would differ. (We present a sample of visualizations rather than visualizations for every state, due to space constraints. Please see appendix)

For our continuous variables (continuous date and dosage strength), in by-state plots we often saw strikingly similar patterns in predictor vs. response, suggesting that random slopes may not be necessary for these variables. For categorical variables, we generally observed that there was indeed variation among the states, but it was not clear whether this variation had any well-defined structure. It was not clear whether our models would benefit from random slopes for these variables, so we decided to further investigate this in our modeling process.

Model

Per the research question of interest, we are interested in exploring factors related to price per mg of oxymorphone. We are also specifically asked to control for location and how drug price changes with it. Cities seemed to be too noisy to use as a location predictor. On the flip side, region did not seem to be associated with a significant change in drug prices. Therefore, state provides the best granularity of accounting for location in this instance.

Linear regression models are usually a good standard approach to this, providing clear interpretations of how various predictors are associated with drug price. As seen in the EDA, however, many states have excessively small sample sizes. This produces corresponding $\hat{\beta}_{OLS}$ estimates with large variance. We would like to be able to *partially pool* information across states, letting us borrow information to help stabilize state estimates with small sample sizes.

This prompts a **mixed effects model**, where we instead model the state coefficients as having been drawn $N(0, \tau^2)$.

Model Selection

Based off of the results of the EDA, we want to initially fit a model with the following variables: `source`, `dose_per_mg`, `bulk_purchase`, `Primary_Reason`, `date_num`, and `date_quarter`. In terms of the research question, we would also like to group based on location. There are 3 options for this: `region`, `state`, or `city`.

In order to test whether keeping a variable in the model is “worth it”, we may perform a likelihood ratio test for nested models. For adding k fixed effects, the LRT test statistic follows a $\chi^2(k)$ distribution, and for adding k random effects, the LRT test statistic follows a 50-50 mixture of $\chi^2(k)$ and $\chi^2(k+1)$ distribution. For non-nested models, we can use BIC as a rough indicator of which model is “better”. BIC balances likelihood of model with how parsimonious it is.

First, let’s decide what to group on for region:

Table 3: BIC for Models with Different Random Intercepts

Grouping	BIC
State	8643.014
City	8769.939
Region	8788.861

State provides the best level of BIC, so we will choose this for our grouping variable.

Now let's investigate dropping various fixed effects and see results of our hypothesis tests:

Model selection is omitted for brevity (see code appendix for full demonstration). Essentially we compared nested models with fewer and fewer variables, provided the results of the Anova test suggested the simpler model was preferred. Ultimately, this led to us dropping `date_num` and `date_quarter`, while keeping: `dose_per_mg`, `bulk_purchase`, `primary_reason`, `source`.

Model Specification

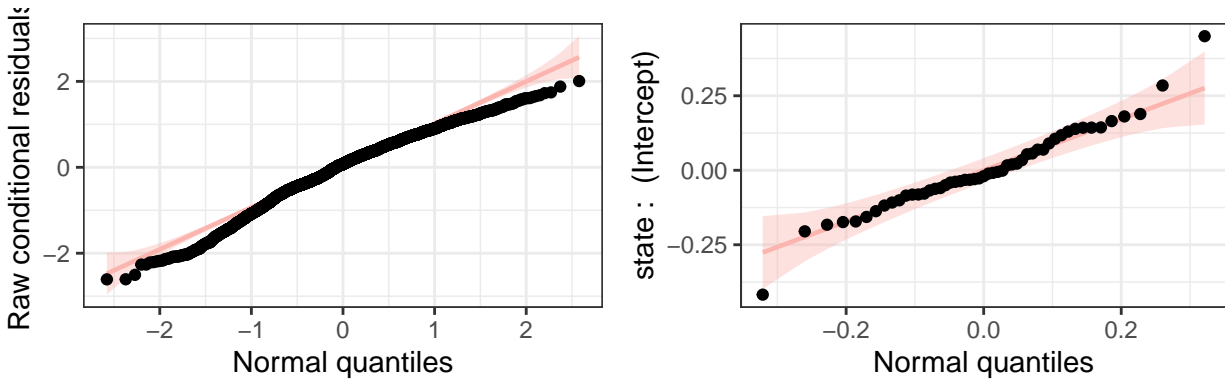
Our final model is as specified:

$$\begin{aligned} \log(y_{ij}) &= \beta_{0,j} + \beta_1 x_{ij1} + \beta_2 x_{ij2} + \beta_3 x_{ij3} + \beta_4 x_{ij4} + \epsilon_{ij} \\ \beta_{0,j} &= \beta_0 + b_{0,j}, b_{0,j} \sim N(0, \tau^2) \\ \epsilon_{ij} \perp b_{0,j} &\stackrel{iid}{\sim} N(0, \sigma^2) \end{aligned}$$

This linear model estimates the log price per milligram $\log(y)$ of oxymorphone for an individual i in state j (x_{ij}). The following predictors are defined as:

- x_1 : Dosage strength in mg of the units purchased
- x_2 : Bulk purchase, an indicator for whether 10+ units were purchased at once
- x_3 : The primary reason for purchase. Factors include “blank”, “left unanswered”, “other or unknown”, “self-treatment”, “enjoyment”, “resell”, and “refuse to answer”
- x_4 : Source of information (allows for reporting of transactions). Factors include “unkown”, “internet”, “internet pharmacy”, “heard it”, and “personal”

Model Diagnostics



From the normal quantile plot of the raw conditional residuals, although there are many residuals out of bounds, these residuals are not too far away from the bounds. Thus, we can still accept the raw conditional residuals are normally distributed. From the normal quantile plot on random effect state, there are only two outliers. Thus, we can accept random effect of state is normally distributed.

Results

Fixed Effects

	Estimate	Std.Error	t-value	P-value
(Intercept)	0.5204313	0.0385012	13.5172754	0.0000000
dose_per_mg	-0.0152224	0.0009142	-16.6507059	0.0000000
bulk_purchase1 Bulk purchase	-0.1716889	0.0307536	-5.5827296	0.0000000

	Estimate	Std.Error	t-value	P-value
Primary_ReasonLeft unanswered	0.0166237	0.0302059	0.5503465	0.5820818
Primary_ReasonOther or unknown	0.0289116	0.0756838	0.3820052	0.7024575
Primary_ReasonSelf-treat	0.1042256	0.0389198	2.6779570	0.0074073
Primary_ReasonEnjoyment	0.1691436	0.0743417	2.2752197	0.0228928
Primary_ReasonResell	-0.3148221	0.1686888	-1.8662895	0.0620009
Primary_ReasonRefuse to answer	0.0523060	0.0446617	1.1711600	0.2415345
sourceInternet	0.1361942	0.0627474	2.1705151	0.0299678
sourceInternet Pharmacy	-0.4135162	0.1078794	-3.8331344	0.0001265
sourcePersonal	0.0591284	0.0267121	2.2135460	0.0268600
sourceHeard it	0.1788115	0.0314385	5.6876587	0.0000000

- *dose_per_mg*: Controlling for other variables, for every one milligram increase of dosage, the price per mg of oxymorphone reduces by a multiplicative effect of $\exp(-0.0152224) = 0.9848929$, about a 1.51% reduction.
- *bulk_purchase* (“not bulk purchase” is the reference): Controlling for other variables, bulk purchase reduces the price per mg of oxymorphone by a multiplicative effect of $\exp(-0.1716889) = 0.8422412$, about a 15.76% reduction than the price per mg of oxymorphone for people without bulk purchase.
- *source* (“Source: Unknown” is the reference, other levels omitted for brevity):
 - Controlling for other variables, purchasing oxymorphone with Internet source tends to increase the price per mg by a multiplicative effect of $\exp(0.1361942) = 1.145904$, about a 14.59% increase than the price per mg of oxymorphone for people with unknown drug source.
- *Primary Reason* (“Blank” is the reference, other levels omitted for brevity):
 - Controlling for other variables, purchasing oxymorphone with primary reason Self-treat tends to increase the price per mg by a multiplicative effect of $\exp(0.1042256) = 1.109851$, about a 10.99% increase than the price per mg of oxymorphone for people with Blank primary reason.

Random Effects

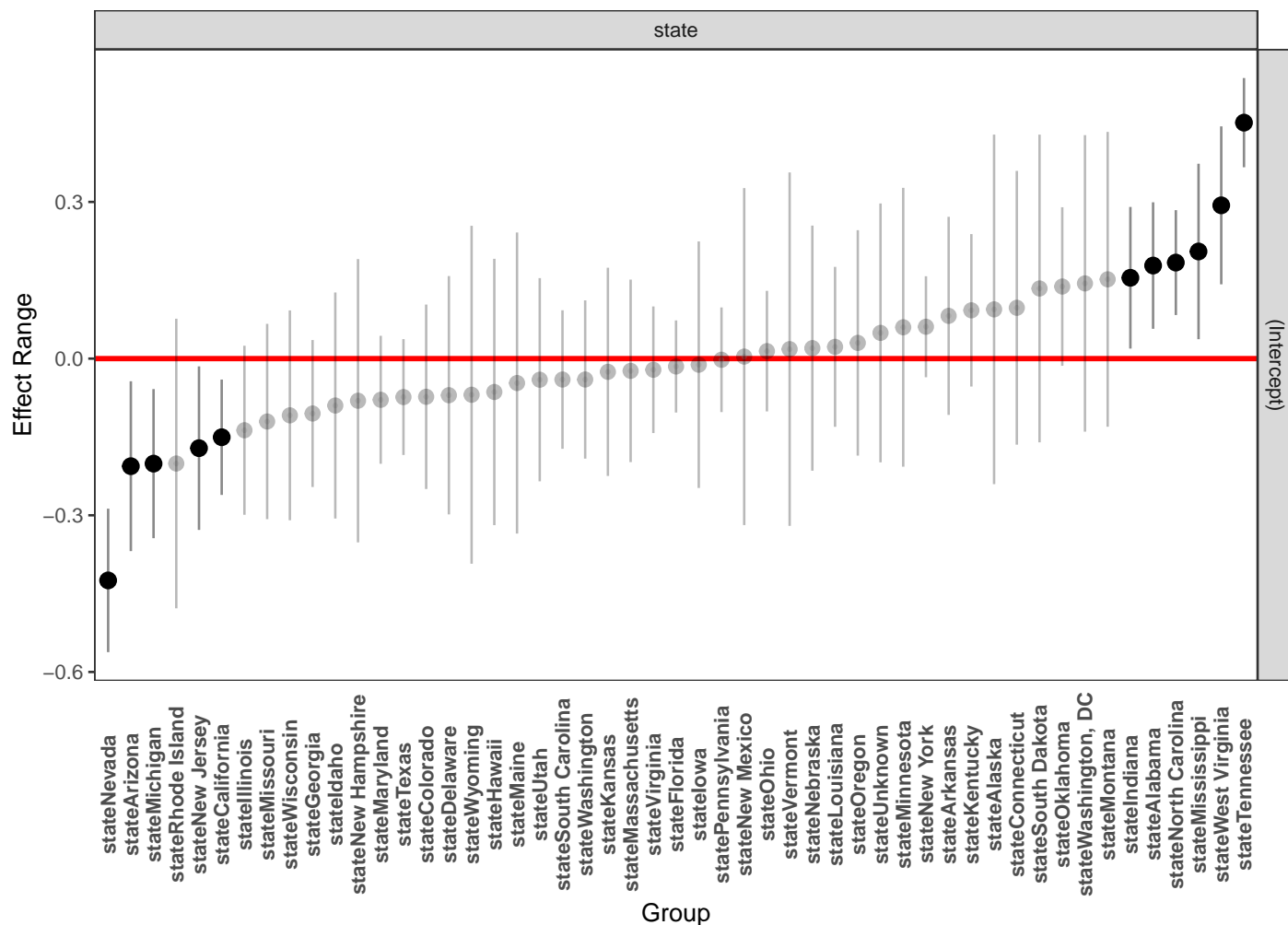
Group	Variance	Std.Dev
state (Intercept)	0.02292	0.1514
Residual	0.5006	0.70753

The within-state variance σ^2 is 0.5006, which is much larger than the across-state variance τ^2 , 0.02292.

Intraclass Correlation
0.044

Intraclass correlation is 0.044, which means values are not very similar from the same state.

Effect Ranges



The value of the random intecepts range from $\exp(0.450007131) = 1.568323$ for Tennessee to $\exp(-0.417349014) = 0.658791$ for Nevada. We interpret this as follows: the estimated average value for ppm of oxymorphone is highest for Tennessee. The price is estimated assuming a non-bulk purchase, unknown source, blank reason, and zero dose per mg (which is obviously not a realistic value, but is noted for the sake of interpretation).

Limitation

Our analysis has several limitations. Our filtering of the `log_ppm` variable to non-extreme values must be kept in mind, as we are essentially modeling typical values only, so some unusual price values may not be predicted correctly by our model. Small sample sizes for some states are not ideal, although our hierarchical model structure helps to mitigate this issue. User-input data is naturally messy and should be taken with a grain of salt. Furthermore, our sample is essentially self-selected (people who choose to enter their data into the StreetRx web site) rather than randomly sampled, and this may affect the generalizability of our results.

Appendix

```
knitr::opts_chunk$set(echo = FALSE, cache=TRUE, fig.height = 3)
library(pacman)
devtools::install_github("goodekat/redres")
pacman::p_load(tidyverse, lme4, gridExtra, grid, ggplot2, lattice, redres, stringr,
               influence.ME, knitr, merTools)
load("streetrx.rdata")
df <- streetrx %>% filter(api_temp == "oxymorphone", !is.na(ppm))
# cleaning state:
# table(df$state)
# remove levels corresponding to states with 0 observations in df
df$state <- droplevels(df$state)
# table(df$state)
# drop North Dakota - one observation
df <- df %>% filter(state != "North Dakota")
df$state <- droplevels(df$state)
# table(df$state)
# replace state "USA" with "Unknown"
df$state <- recode_factor(df$state, "USA" = "Unknown")
# table(df$state)
### PUT IN APPENDIX: the state-by-region table of counts below
# table(df$state, df$USA_region)
# Change date to numeric (easier for interpretation)
# this sets
df$date_num <- as.numeric(as.Date(df$price_date, "%m/%d/%y")) - 14621 # sets 1/12/2010 (first post-
tempdf <- df %>% dplyr::select(date_num, price_date)
# this filters out all data before the 1/1/2000, which probably shouldn't be included
df <- df %>% filter(date_num >= 0)
df$date_quarter <- as.numeric(substr(df$yq_pdate, 5, 5))
# recode sources
# remove unused levels
df$source <- droplevels(df$source)
source_df <- data.frame(table(df$source))
# levels(df$source)
levels(df$source)[1] <- "Blank"
source_df <- data.frame(table(df$source))
df$source <- recode_factor(df$source, "Blank" = "Blank",
                          "Internet" = "Internet",
                          "Internet Pharmacy" = "Internet Pharmacy",
                          "Personal" = "Personal",
                          "Heard it" = "Heard it",
                          "Idk" = "Unknown",
                          "w" = "Unknown",
                          "STREET PRICE" = "Heard it",
                          .default = "Internet") # everything else is a URL
source_df_2 <- data.frame(table(df$source))
# drop unknowns
df <- df %>% filter(df$source != "Unknown")
df$source <- droplevels(df$source)
knitr::kable(table(df$source), col.names = c("Level", "Freq"), caption = "Frequency Table of Variabl
# remove unused levels
```

```

df$Primary_Reason <- droplevels(df$Primary_Reason)
reason_df <- data.frame(table(df$Primary_Reason))
# levels(df$Primary_Reason)
levels(df$Primary_Reason)[1] <- "Blank"
reason_df <- data.frame(table(df$Primary_Reason))
df$Primary_Reason <- recode_factor(df$Primary_Reason, "Blank" = "Blank",
                                "0 Reporter did not answer the question" = "Left unanswered",
                                "Other or unknown" = "Other or unknown",
                                "9 To self-treat my pain" = "Self-treat",
                                "3 To prevent or treat withdrawal" = "Self-treat",
                                "10 To treat a medical condition other than pain" = "Self-treat",
                                "4 For enjoyment/to get high" = "Enjoyment",
                                "5 To resell" = "Resell",
                                "8 Prefer not to answer" = "Refuse to answer",
                                .default = "Other or unknown")
reason_df_2 <- data.frame(table(df$Primary_Reason))
#print(reason_df)
knitr::kable(reason_df_2, col.names = c("Level", "Freq"), caption = "Frequency Table of Variable Primary Reason")
# rename variable for ease of interpretability
df <- df %>% rename(dose_per_mg = mgstr)
# log transform ppm
df$log_ppm <- log(df$ppm)
hist_ppm <- ggplot(df, aes(x=ppm)) +
  geom_histogram(aes(y=..density..), colour="black", fill="white", bins = 50)+
  geom_density(alpha=.2, fill="red")
hist_log_ppm <- ggplot(df, aes(x=log_ppm)) +
  geom_histogram(aes(y=..density..), colour="black", fill="white", bins = 50)+
  geom_density(alpha=.2, fill="blue")
grid.arrange(hist_ppm, hist_log_ppm, ncol=2)
df <- df %>% filter(log_ppm > -2 & log_ppm < 2)
ggplot(df, aes(x=log_ppm, y=reorder(state, state, length), fill=state)) +
  geom_boxplot() +
  labs(title="State", x="Log PPM by state",y="Log PPM")
ggplot(df, aes(x=reorder(USA_region, USA_region, length), y=log_ppm, fill=USA_region)) +
  geom_boxplot() +
  labs(title="log PPM by region", x="region",y="Log PPM")
# Fit 3 models grouping on different location specificity
mod_s <- lmer(data=df, log_ppm ~ (1 |state) + date_num + date_quarter + dose_per_mg +
  bulk_purchase + Primary_Reason + source, REML=F)
mod_c <- lmer(data=df, log_ppm ~ (1 |city) + date_num + date_quarter + dose_per_mg +
  bulk_purchase + Primary_Reason + source, REML=F)
mod_r <- lmer(data=df, log_ppm ~ (1 |USA_region) + date_num + date_quarter + dose_per_mg +
  bulk_purchase + Primary_Reason + source, REML=F)
BIC <- sapply(c(mod_s, mod_c, mod_r), BIC)
re_results <- data.frame('Grouping' = c('State', 'City', 'Region'), 'BIC' = BIC)
kable(re_results, caption = 'BIC for Models with Different Random Intercepts')
# Test date_num
drop_date_num <- anova(mod_s,
  lmer(data=df, log_ppm ~ (1 |state) + date_quarter + dose_per_mg + bulk_purchase +
    Primary_Reason + source, REML=F))
date_num_pval <- drop_date_num$`Pr(>Chisq)`[2]

```

```

# Test date_quarter
drop_quarter <- anova(lmer(data=df, log_ppm ~ (1 |state) + date_quarter + dose_per_mg +
                        bulk_purchase + Primary_Reason + source, REML=F),
                    lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase +
                        Primary_Reason + source, REML=F))
quarter_pval <- drop_quarter$`Pr(>Chisq)`[2]
# Test dose_per_mg
drop_dose <- anova(lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + Primary_Reason + source, REML=F),
                    lmer(data=df, log_ppm ~ (1 |state) + bulk_purchase + Primary_Reason
                        + source, REML=F))
dose_pval <- drop_dose$`Pr(>Chisq)`[2]
# Test bulk_purchase
drop_bulk <- anova(lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + Primary_Reason + source, REML=F),
                    lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + Primary_Reason
                        + source, REML=F))
bulk_pval <- drop_bulk$`Pr(>Chisq)`[2]
# Test primary reason
drop_reason <- anova(lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + Primary_Reason + source, REML=F),
                    lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + source, REML=F))
reason_pval <- drop_reason$`Pr(>Chisq)`[2]
# Test source
drop_source <- anova(lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + Primary_Reason + source, REML=F),
                    lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase
                        + Primary_Reason, REML=F))
source_pval <- drop_source$`Pr(>Chisq)`[2]
mod_final <- lmer(data=df, log_ppm ~ (1 |state) + dose_per_mg + bulk_purchase + Primary_Reason + source)
resqq <- plot_resqq(mod_final)
ranefci <- plot_ranef(mod_final)
grid.arrange(resqq, ranefci, ncol = 2)
coefs <- data.frame(coef(summary(mod_final)))
# use normal distribution to approximate p-value
coefs$p.z <- 2 * (1 - pnorm(abs(coefs$t.value)))
knitr::kable(coefs, col.names = c("Estimate", "Std.Error", "t-value", "P-value"))
VarCorr(mod_final)
Variance <- c(0.02292, 0.50060)
Std.Dev <- c(0.15140, 0.70753)
Groups <- c("state (Intercept)", "Residual")
REvar <- cbind(Groups, Variance, Std.Dev)
knitr::kable(REvar, col.names = c("Group", "Variance", "Std.Dev"))
library(sjstats)
icc(mod_final)
ICC <- 0.044
knitr::kable(ICC, col.names = "Intraclass Correlation")
plotREsim(RESim(mod_final, n.sims=100), stat='median', sd=TRUE, labs = TRUE)
hist(df$date_num, xlab = "date_num", main = "Histogram of date_num")
knitr::kable(table(df$date_quarter), caption = "Frequency Table of Quarter in date_num")

```

```

hist(df$dose_per_mg, xlab = "does_per_mg", main = "Histogram of dose_per_mg")
ggplot(df, aes(x=reorder(USA_region, USA_region, length), y=log_ppm, fill=USA_region)) +
  geom_boxplot() +
  labs(title="log PPM by region", x="region",y="Log PPM")
ggplot(data = df[df$state=="Florida" | df$state=="Texas" | df$state=="North Carolina",],
  aes(x = dose_per_mg, y = log_ppm)) +
  geom_point() +
  facet_grid(~ state)
ggplot(data = df[df$state=="Florida" | df$state=="Texas" | df$state=="North Carolina",],
  aes(x = date_num, y = log_ppm)) +
  geom_point() +
  facet_grid(~ state)
# categorical variables -- looking at South only
set.seed(42)
states <- sample(df$state, 5)
ggplot(data = df %>% filter(state %in% states), aes(x = bulk_purchase, y = log_ppm)) +
  geom_boxplot() +
  facet_grid(~ state) +
  theme(axis.text.x = element_text(angle = 90))
ggplot(data = df[df$state=="Florida" | df$state=="Texas" | df$state=="North Carolina",],
  aes(x = date_quarter, y = log_ppm, group = date_quarter)) +
  geom_boxplot() +
  facet_grid(~ state)
ggplot(data = df[df$state=="Florida" | df$state=="Texas" | df$state=="North Carolina",],
  aes(x = Primary_Reason, y = log_ppm, group = Primary_Reason)) +
  geom_boxplot() +
  facet_grid(~ state) +
  theme(axis.text.x = element_text(angle = 90))
ggplot(data = df[df$state=="Florida" | df$state=="Texas" | df$state=="North Carolina",],
  aes(x = source, y = log_ppm, group = source)) +
  geom_boxplot() +
  facet_grid(~ state) +
  theme(axis.text.x = element_text(angle = 90))
ggplot(df[df$state=="Florida",], aes(x=log_ppm, y=reorder(city, city, length))) +
  geom_boxplot() +
  labs(title="log PPM by city (FL)", x="Log PPM",y="city")
# To get a better sense of how state-level groups and slopes are distributed, we created the follow

### Dose strength variable

group_int_list <- NULL
group_slope_list <- NULL
for(statej in unique(df$state)) {
  m <- lm(data = df[df$state==statej, ], log_ppm ~ dose_per_mg)
  group_int_list <- c(group_int_list, coef(m)[1])
  group_slope_list <- c(group_slope_list, coef(m)[2])
}

par(mfrows=c(2, 1))
hist(group_int_list)
hist(group_slope_list)

```

```

### Continuous date variable

group_int_list <- NULL
group_slope_list <- NULL
for(statej in unique(df$state)) {
  m <- lm(data = df[df$state==statej, ], log_ppm ~ date_num)
  group_int_list <- c(group_int_list, coef(m)[1])
  group_slope_list <- c(group_slope_list, coef(m)[2])
}

par(mfrows=c(2, 1))
hist(group_int_list)
hist(group_slope_list)

```

Visualizations

Date variable

Histogram of date_num

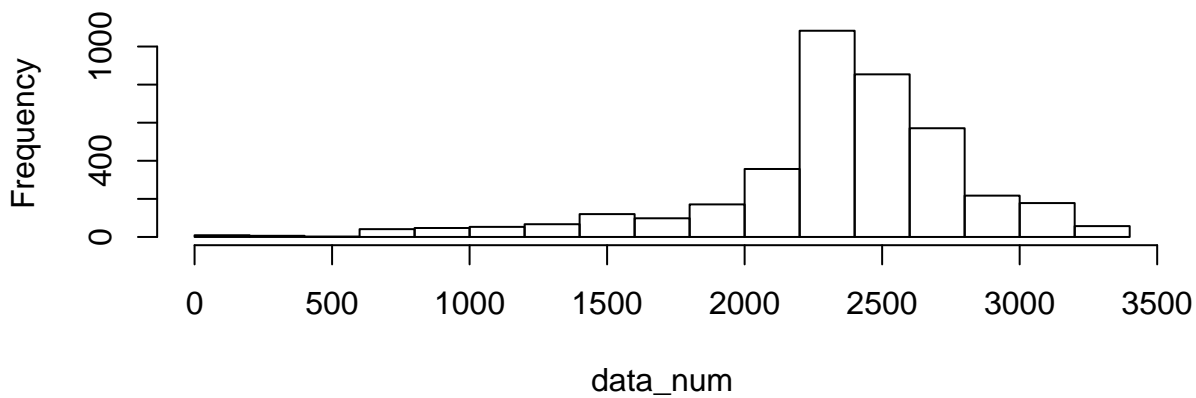
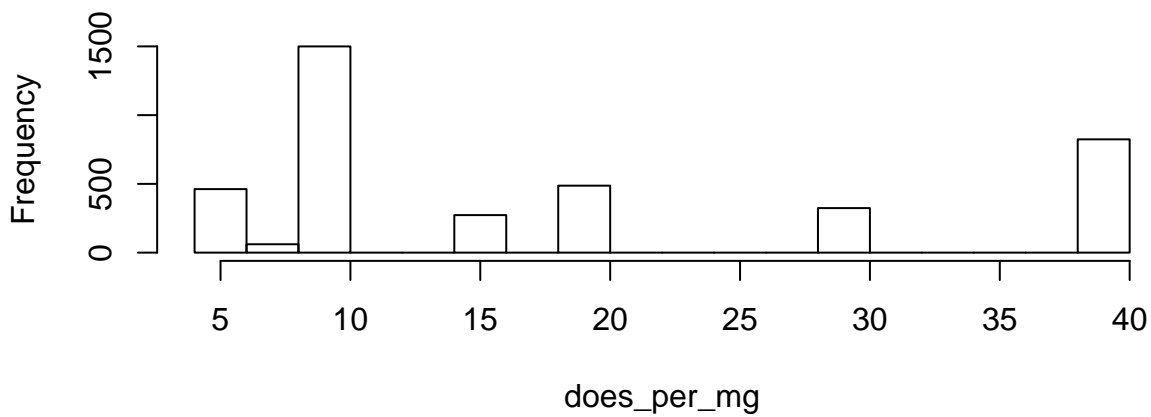


Table 7: Frequency Table of Quarter in date_num

Var1	Freq
1	1106
2	1161
3	867
4	797

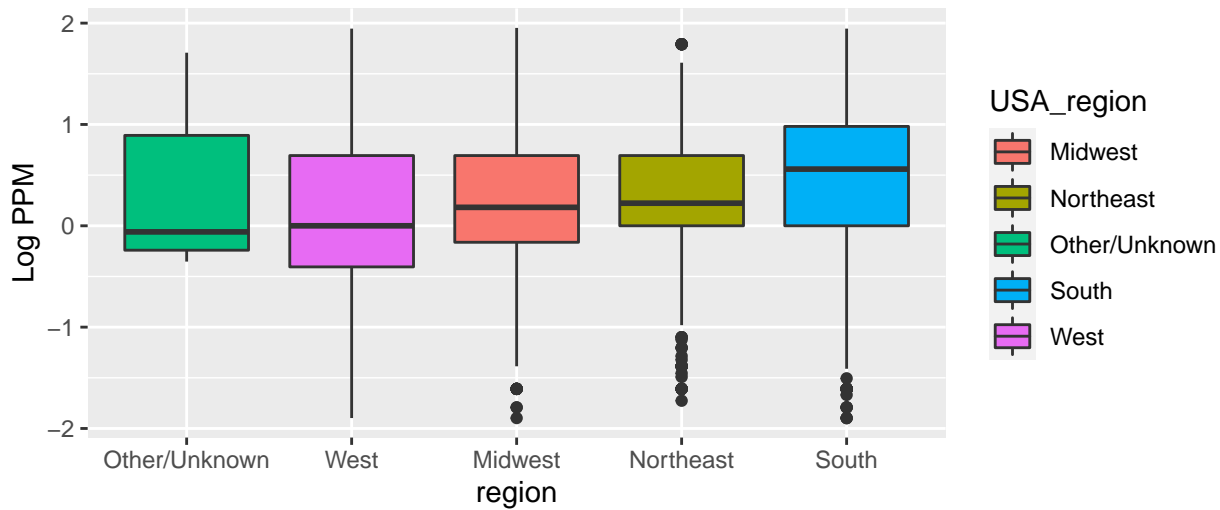
Dose strength variable

Histogram of dose_per_mg

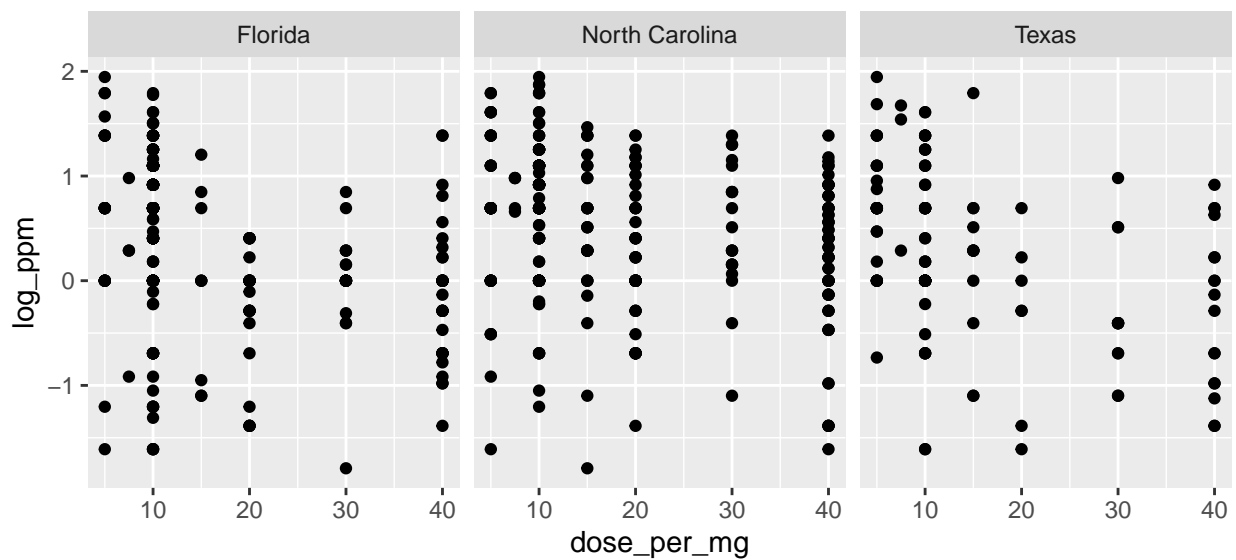


Region box plots

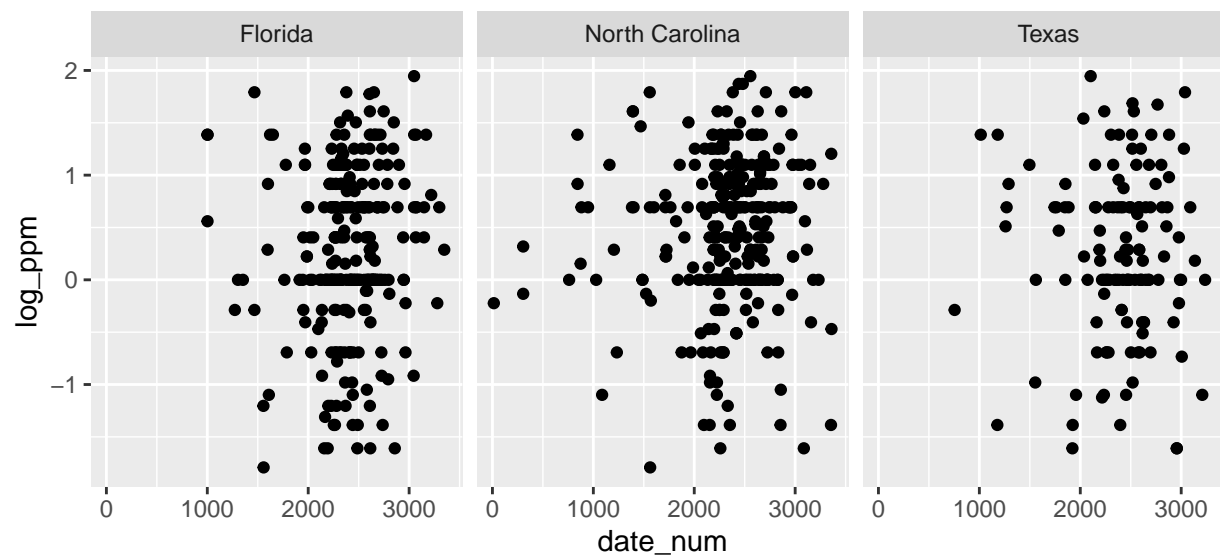
log PPM by region



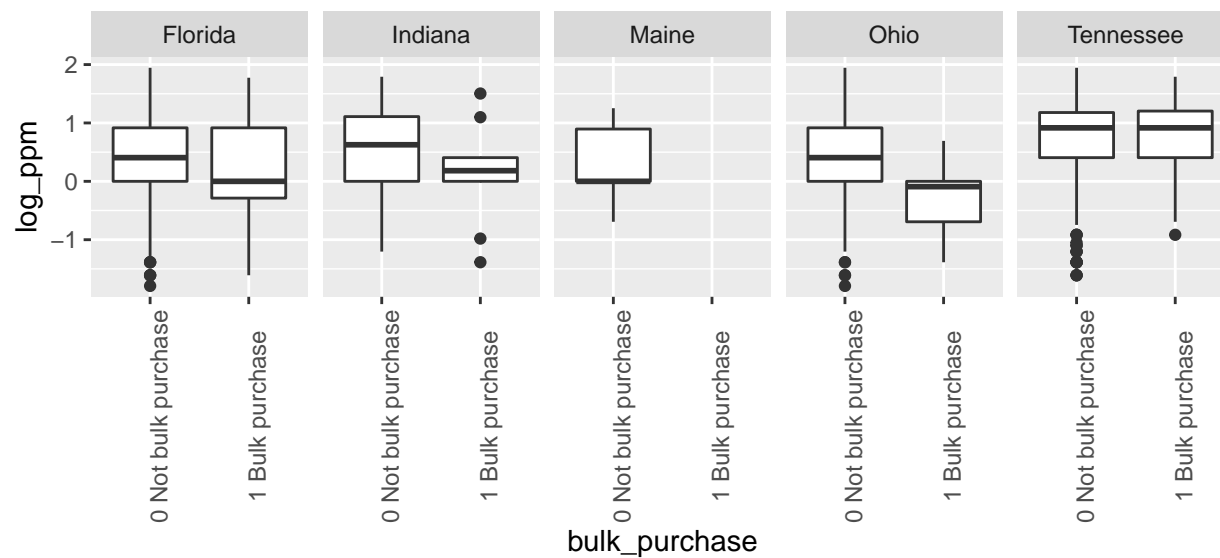
dose_per_mg



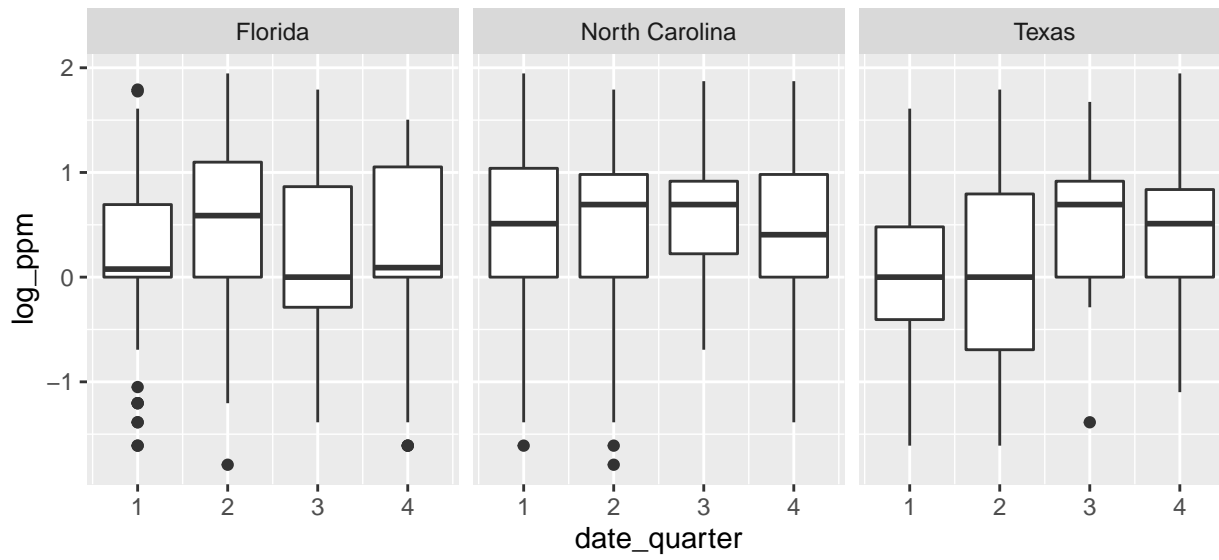
date_num



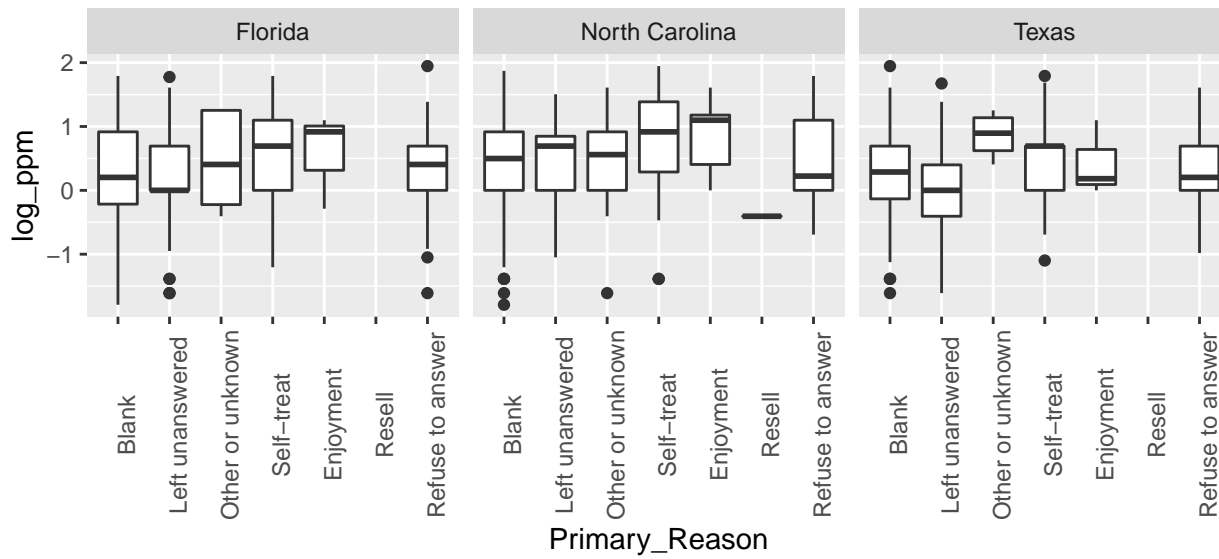
bulk_purchase



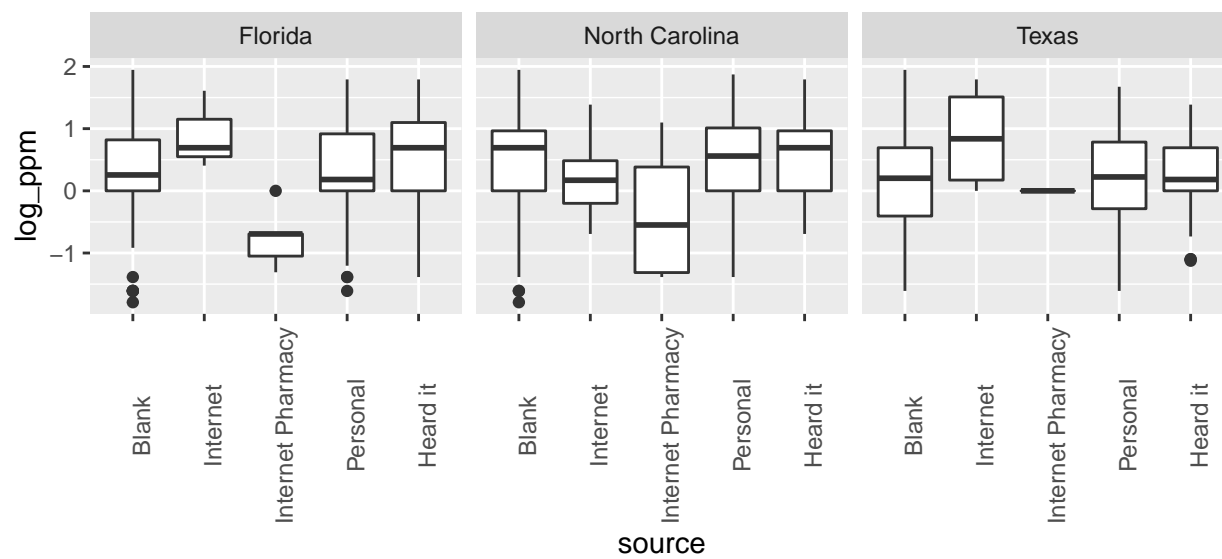
date_quarter



Primary_Reason

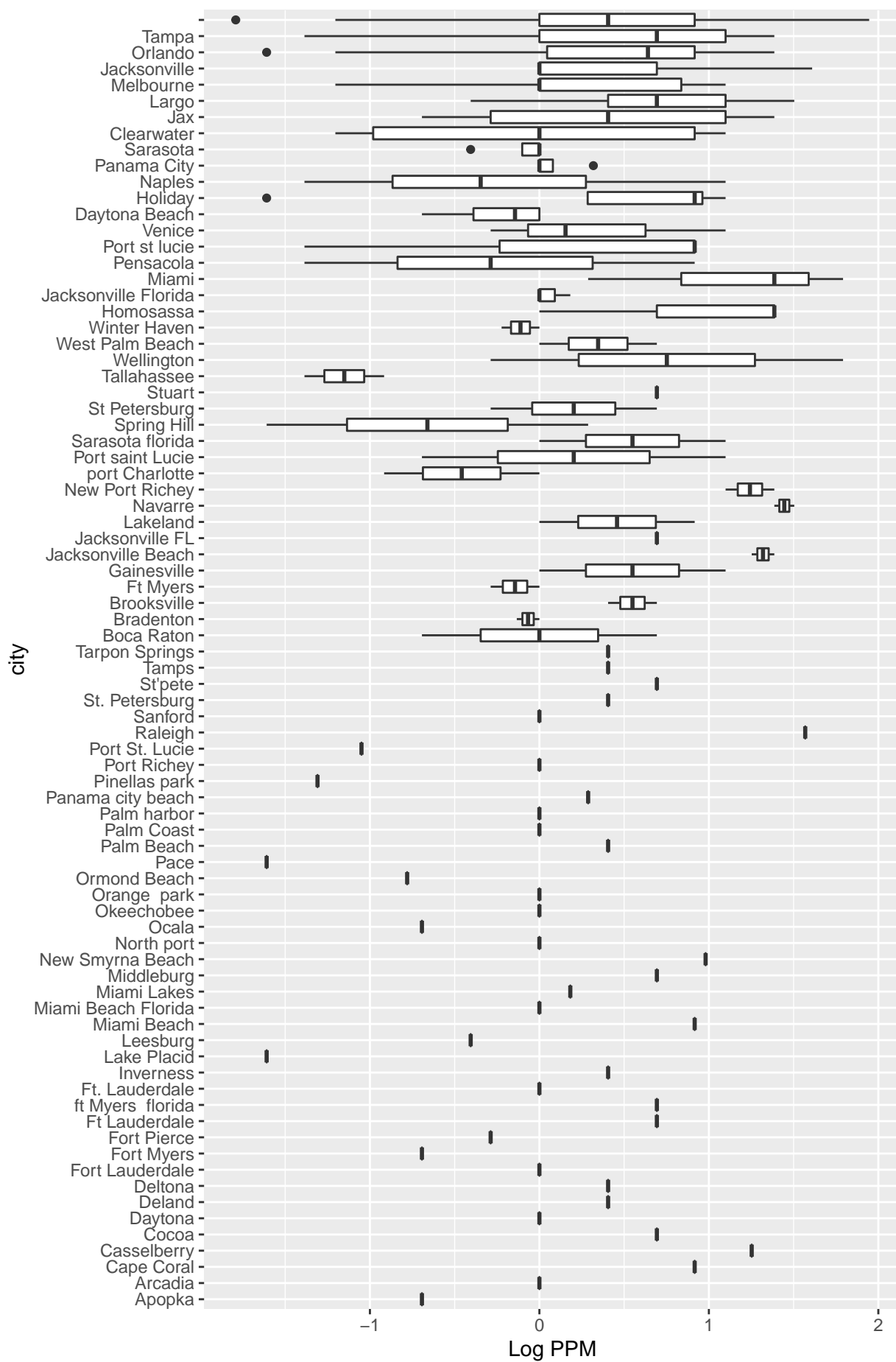


source



City visualization for FL

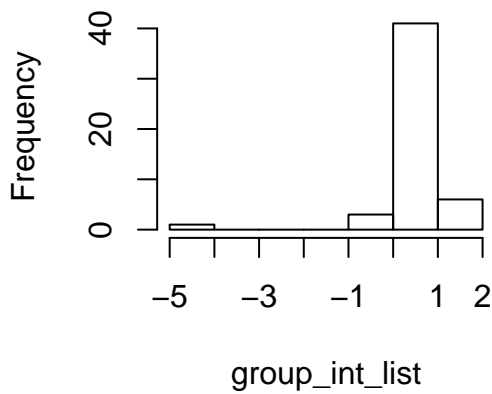
log PPM by city (FL)



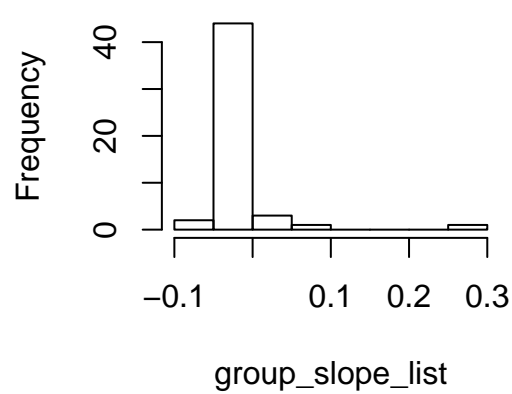
More EDA: slope and intercept distributions for predictors by state

```
## Warning in par(mfrows = c(2, 1)): "mfrows" is not a graphical parameter
```

Histogram of group_int_list

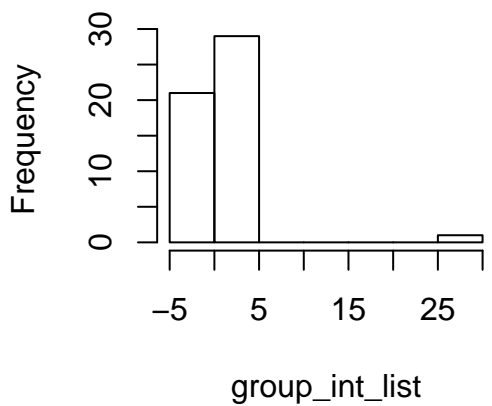


Histogram of group_slope_list



```
## Warning in par(mfrows = c(2, 1)): "mfrows" is not a graphical parameter
```

Histogram of group_int_list



Histogram of group_slope_list

