

Homework 3

Stat 597a: Spatial Models

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Problem 1:

Fit a linear model relating rent per square meter to the covariates using least squares, and extract the coefficient estimates. You can ignore the Location variable for now since we will later treat this as a random effect. Also note that the room indicator variables include one that is redundant, so treat a single room as the baseline (i.e., leave Room1 out of the model, so the intercept corresponds to a single room and coefficients for the others represent adjustments to the intercept for a different number of rooms).

For this problem, I fit a linear model, but I do not include the intercept, because there is no intercept included in part 4 for the spatial model. As instructed, I also do not include Location or Room1 in this model. I include the estimates for the coefficients below.

##	(Intercept)	Year	NoHotWater	NoCentralHeat
##	-16.01232971	0.01333062	-1.87050503	-1.22576091
##	NoBathTiles	SpecialBathroom	SpecialKitchen	Room2
##	-0.72571150	0.66037195	1.46288743	-1.31937203
##	Room3	Room4	Room5	Room6
##	-1.88640798	-2.46463131	-2.37815761	-2.48262146

Problem 2:

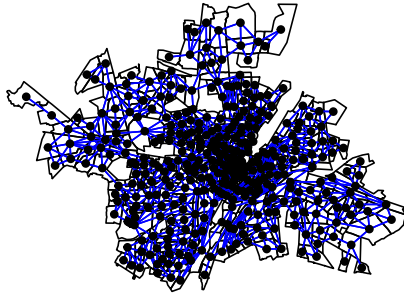
There are two SpatialPolygons objects associated with this dataset, districts.sp and parks.sp. The first corresponds to city districts in which apartments may be located. The second corresponds to districts with no possible apartments, such as parks or fields. Create an nb object with neighbors for the districts, defining neighbors as districts that share a common boundary. Make a plot showing the districts, then add the parks shaded a different color. Think about the way parks are treated; what else could we do?

First, I created an nb object with neighbors for the districts. I use the rook pattern because the problem states to define neighbors as districts that share a common boundary. The queen pattern would define neighbors as those districts which share a boundary and/or a corner.

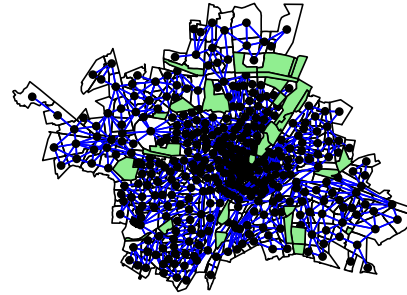
In my first plot below, I include a plot showing the districts, with the parks added in a different color (green). In my second plot below, I also add the neighborhood structure to this first map. In the context of this assignment, we do not treat the parks as districts. In other words, these parks are ignored and treated as empty.

Alternatively, perhaps we should consider districts that both share a border with the park to also be neighbors, as they are probably close to each other.

Districts with Neighbors



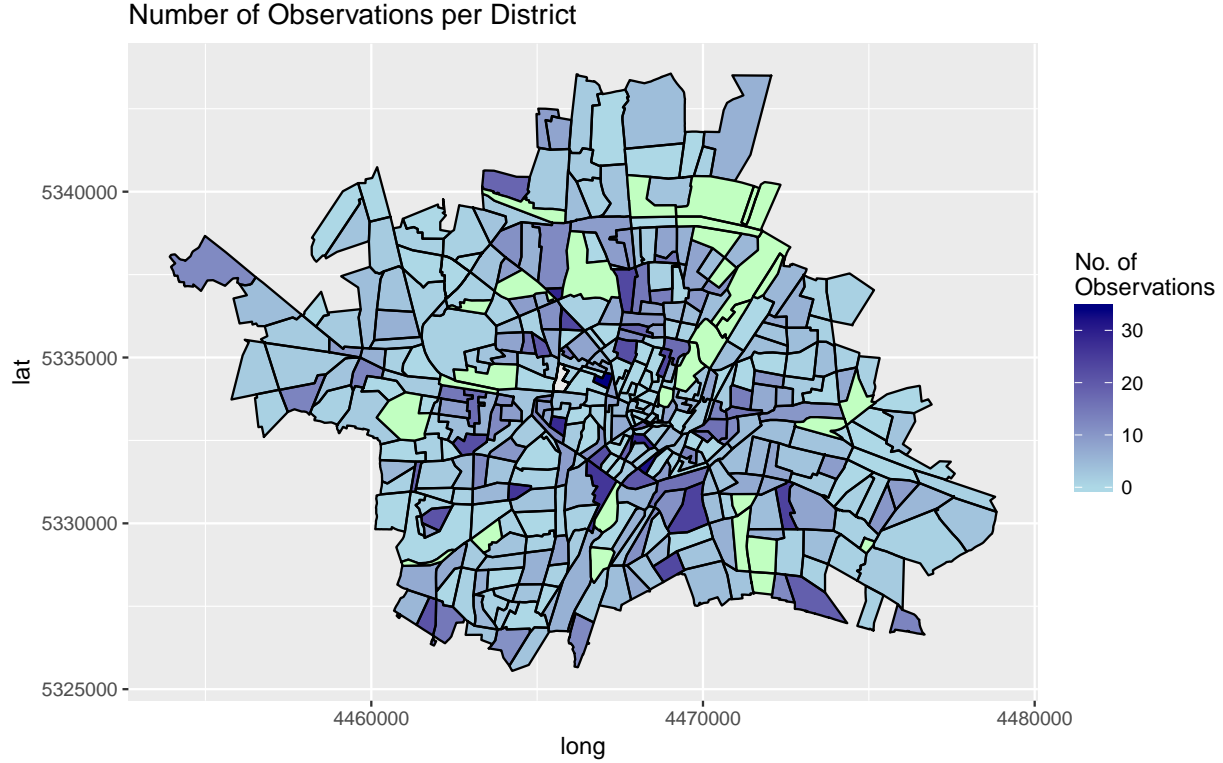
Map of the Districts, Parks in Green



Problem 3:

There are 380 districts in `districts.sp`, and the corresponding district numbers are indicated by the `Location` variable in `rents`. How many of the 380 districts appear in the rent dataset? I've included a matrix `H` that provides a mapping between the districts as they're ordered in `districts.sp` and as they appear in the `rents` dataframe. Use `H` to create a new vector containing the number of observations in each district, and make a color or grayscale plot to illustrate this. Note that inference for unobserved districts will still be possible under a hierarchical mixed effects model, since we can "borrow strength" from nearby districts that do have observations.

312 of the 380 districts appear in the rent dataset. Below, I include a map with the number of apartments in the rent dataset for each district. Once again, I include the empty districts with a red etching.



Problem 4

We will now create a Gibbs sampler to sample from the posterior distribution under the following Bayesian model. Let X be the matrix of covariates, including the intercept term. Let n be the number of data points in Y and m be the number of spatial locations in η . Data model:

$$Y|\beta, \eta, \sigma^2 \sim N(X\beta + H\eta, \sigma^2 I)$$

Process model:

$$p(\eta|\tau^2) \propto (\tau^2)^{-(m-1)/2} \exp\left\{\frac{-1}{2\tau^2} \eta^T (D_w - W)\eta\right\}$$

where W is the matrix of 0's and 1's indicating the neighborhood structure from problem 2, and D_w is a diagonal matrix with diagonal entries $\sum_j W_{1j}, \dots, \sum_j W_{nj}$. That is, η follows an (improper) intrinsic autoregressive model. The $-(m-1)/2$ exponent on τ^2 is due to the fact that the matrix $D_w - W$ has rank $m-1$ rather than m .

Prior model: Specify independent priors for β , σ^2 , and τ^2 , with

$$p(\beta) \propto 1, \text{ and } \sigma^2, \tau^2 \sim \text{InverseGamma}(0 : 001, 0 : 001)$$

The full conditional distributions for β , η , σ^2 , and τ^2 are given at the end of this assignment. Construct a Gibbs sampler that cycles through each of the full conditionals and stores the results for $B = 10,000$ iterations. The full conditionals are given below.

A few notes to keep in mind when constructing the sampler:

- The matrix W can be computed from your `nb` object in problem 2; see `help(nb2mat)`. I also included objects `X` and `y` with the data file.
- The function `rinvgamma` is in the library `MCMCpack`.

- **IMPORTANT:** The intrinsic autoregressive model is an example of a pair-wise difference prior. It defines proper distributions for the differences $n_i - n_j$, but it also implicitly contains a distribution for $\frac{1}{m} \sum_{i=1}^m \eta_i$ that has infinite variance. In practice, since there is also an intercept term in $X\beta$, we impose the constraint $\sum_{i=1}^m \eta_i = 0$ when we sample from the full conditional for η . **Do this numerically by subtracting the mean $\frac{1}{m} \sum_{i=1}^m \eta_i^{(j)}$ from $\eta^{(j)}$ in each iteration j.**

Full Conditionals:

$$\beta|Rest \sim N((X^T X)^{-1} X^T (Y - H\eta), \sigma^2 (X^T X)^{-1})$$

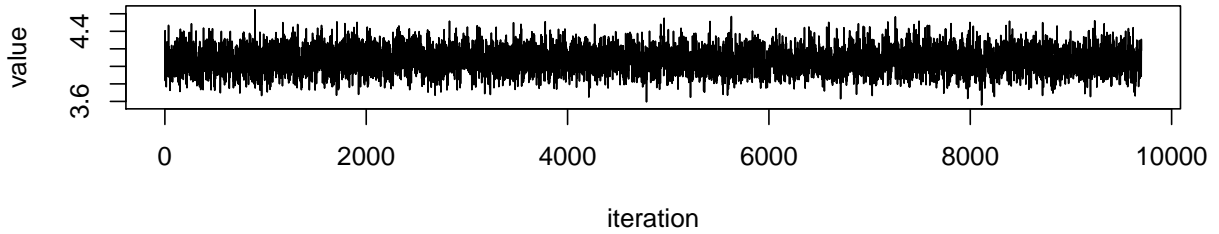
$$\eta|Rest \sim N([H^T H / \sigma^2 + (D_w - W) / \tau^2]^{-1} H^T (Y_X \beta) / \sigma^2, [H^T H / \sigma^2 + (D_w - W) / \tau^2]^{-1})$$

$$\sigma^2|Rest \sim InverseGamma(0.001 + n/2, 0.001 + (Y - X\beta - H\eta)^T (Y - X\beta - H\eta) / 2)$$

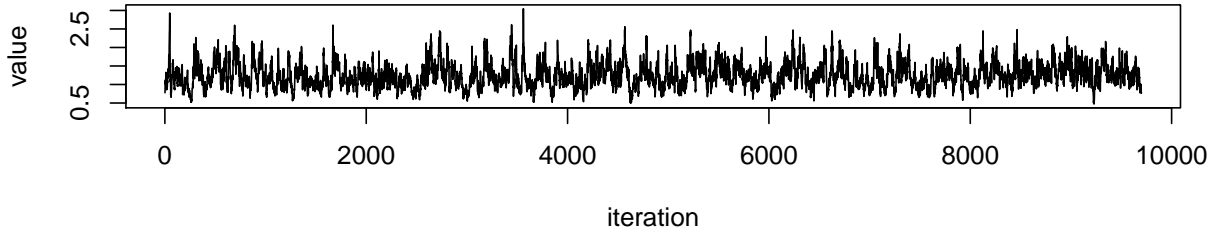
$$\tau^2|Rest \sim InverseGamma(0.001 + (m - 1) / 2, 0.001 + \eta^T (D_w - W) \eta / 2)$$

I ran the Gibbs Sampler for the B=10,000 iterations and chose a burnin value of 300. After discarding the burnin, I include the trace and ACF plots for σ^2 and τ^2 below.

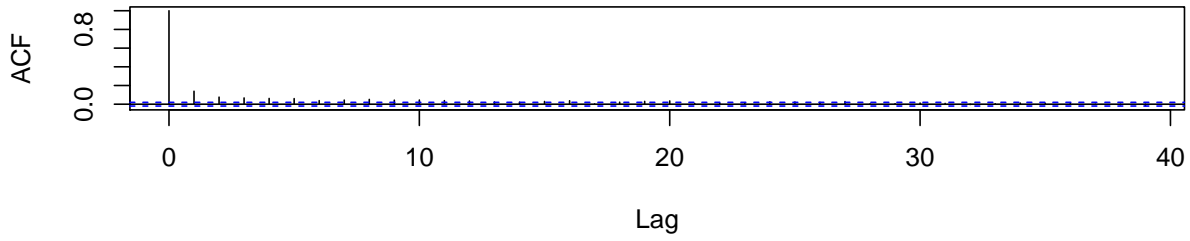
Trace Plot, sigma^2



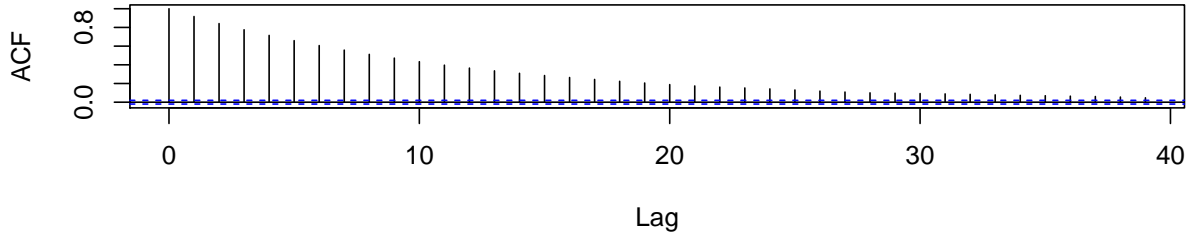
Trace Plot, tau^2



Series s2.final



Series t2.final

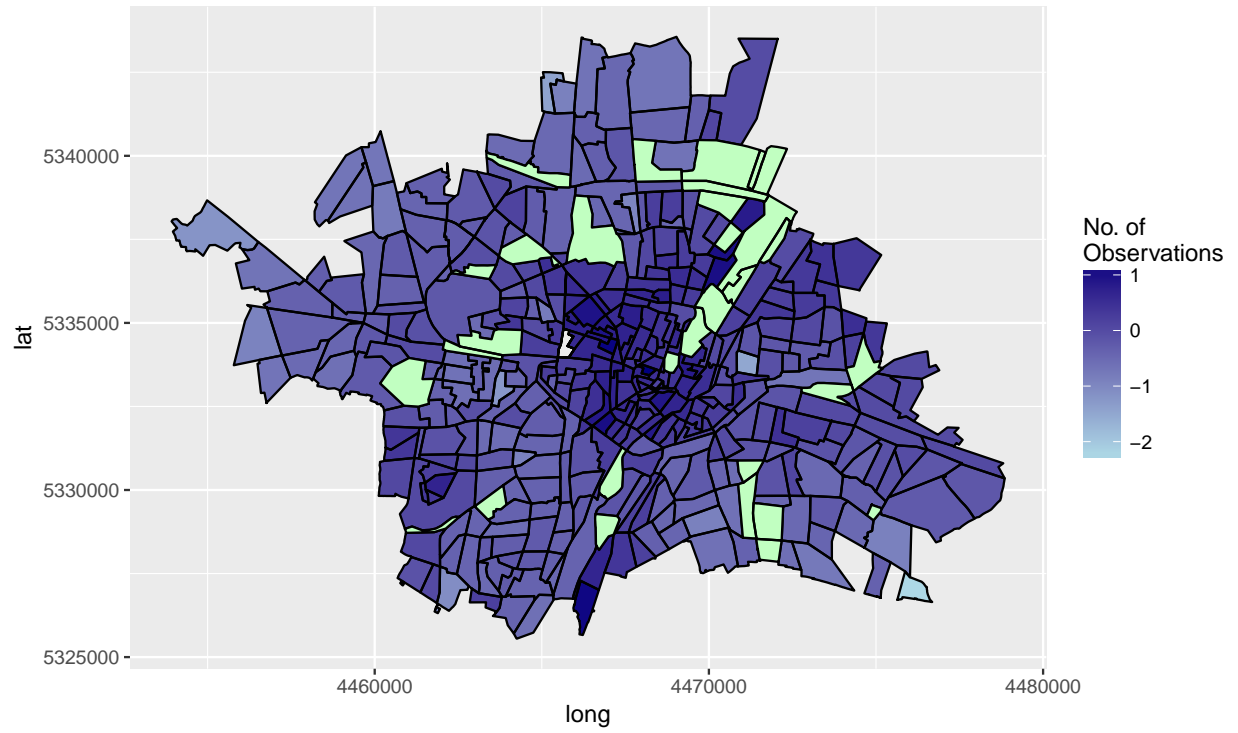


Below, I also include a table with posterior means of the β 's and 95% credible intervals constructed using the 0.025 and 0.975 quantiles of the posterior samples. We also see that the initial estimates are somewhat close to the posterior mean.

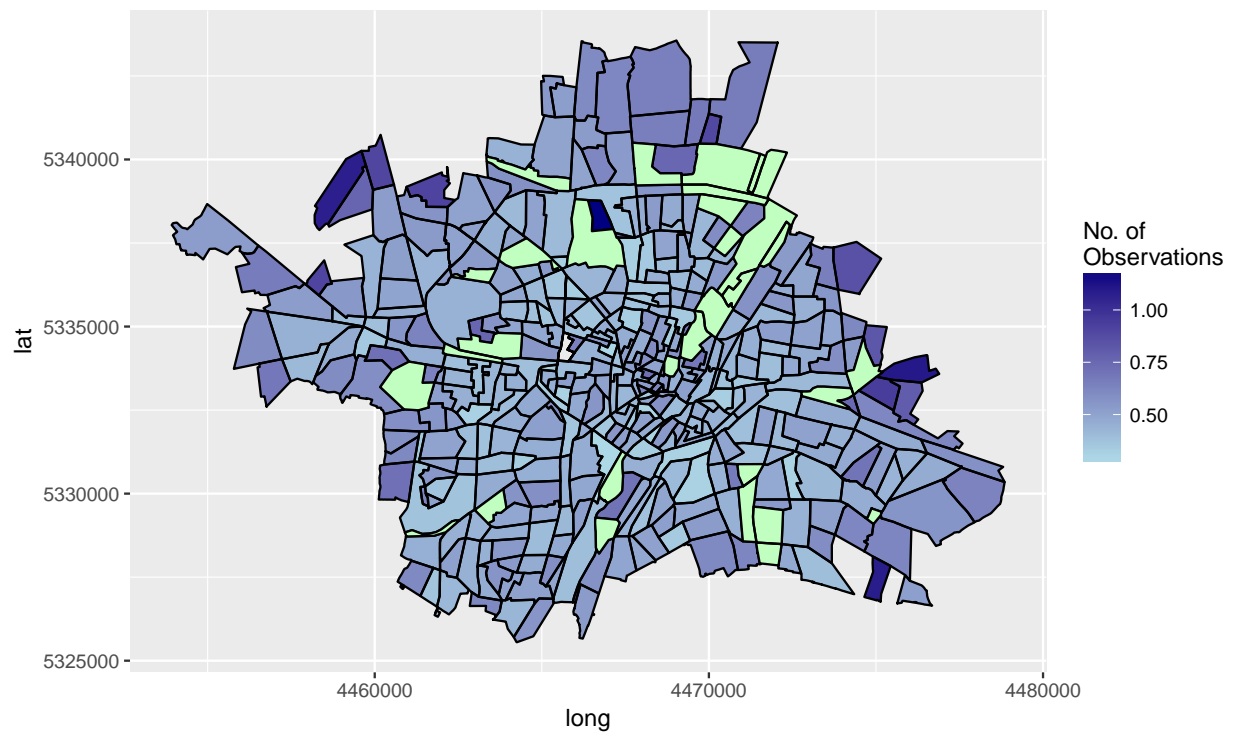
##	0.025 Quantile	Posterior Mean	0.975 Quantile	Initial Estimate
## b1	-34.18411593	-25.58844205	-17.05064262	-16.01232971
## b2	0.01383148	0.01816817	0.02253364	0.01333062
## b3	-2.47055631	-1.89206710	-1.32404427	-1.87050503
## b4	-1.70479634	-1.32081845	-0.92790128	-1.22576091
## b5	-0.87824510	-0.64699301	-0.40974553	-0.72571150
## b6	0.30212687	0.62302351	0.93997565	0.66037195
## b7	1.00411453	1.36144407	1.71809378	1.46288743
## b8	-1.53984412	-1.23904601	-0.93397340	-1.31937203
## b9	-2.05097045	-1.75365098	-1.45354411	-1.88640798
## b10	-2.64096266	-2.27465940	-1.90252622	-2.46463131
## b11	-2.89235823	-2.23359527	-1.56152735	-2.37815761
## b12	-3.63843443	-2.48125371	-1.36286128	-2.48262146

Lastly, I include a color map of the posterior means for the vector η as well as a map for the posterior standard deviations for the vector η . We see that the mean is highest in the center and that the standard deviation is highest near the border of the map, where there are the fewest neighbors.

Posterior mean of eta



Standard Dev of Post eta



Appendix: R Code

Below, I include all of the R code that has produced graphs and data.

```
#load the data
load("C:/Users/ckell/OneDrive/Penn State/2017-2018/597/spatial_statistics_597/Homework 3/data/munichrents.rda")

#head(rents)

#Fit a linear model relating rent per square meter to the covariates using least squares
# This model is fit without Location (used this later as a random effect)
# This model is also fit without Room1 because we treat a single room as the baseline
lin_mod <- lm(RentPerM2~ Year+ NoHotWater+NoCentralHeat+ NoBathTiles+
              SpecialBathroom + SpecialKitchen +Room2+Room3 +Room4 +Room5+ Room6, data=rents)
#summary(lin_mod)

#extract the coefficient estimates
beta <- lin_mod$coefficients
beta

plot(parks.sp)
plot(districts.sp)

#Create an nb object with neighbors for the districts, defining neighbors as districts that
#share a common boundary.
#queen is the option that has neighbor being at least one point shared, including corners
#rook does not include corners, only borders
district_neighb <- poly2nb(districts.sp, queen = TRUE)

par(mfrow = c(1,2))

#plot the neighborhoods
coords <- coordinates(districts.sp)
plot(districts.sp, main = "Districts with Neighbors")
plot(district_neighb, coords, col="blue", pch = 19, cex = 0.6, add = TRUE)

#Make a plot showing the districts, then add the parks shaded a different color.
plot(districts.sp, main = "Map of the Districts, Parks in Green")
plot(district_neighb, coords, col="blue", pch = 19, cex = 0.6, add = TRUE)
plot(parks.sp,col = "lightgreen", add=TRUE)

#How many of the 380 districts appear in the rent dataset?
length(unique(rents$Location))
# There are 312 districts represented in the rents dataset, of the 380 total districts.

# Use H to create a new vector containing the number of observations in each district
# There are 380 districts, so I want a 1x380 or 380x1 vector
dim(H) #H is of dimension 2035X380
# So, I need (1x2035)(2035x380) = 1x380
rents$indicator <- rep(1, nrow(rents))
num_per_dist <- rents$indicator%*% H
dim(num_per_dist) # this is my 1x380 vector
sum(num_per_dist) # This is equal to 2035, the total number of observations (rows) in the rent dataset
num_per_dist <- as.numeric(num_per_dist)
```

```

#Now I will create the data structure that I need to create a plot
sp_f <- fortify(districts.sp)
#head(sp_f)
districts.sp$id <- row.names(districts.sp)
#head(districts.sp@data, n=2)
districts.sp@data$num <- num_per_dist
sp_f <- left_join(sp_f, districts.sp@data)
parks.sp <- fortify(parks.sp)
num <- num_per_dist

#make a color or grayscale plot to illustrate this
obs_by_dist <- ggplot(data = sp_f, aes(long, lat, group = group, fill = num)) + geom_polygon(linetype=0,
  labs(fill = "No. of \nObservations") +
  ggtitle("Number of Observations per District")+ scale_fill_gradient(low = "lightblue", high = "darkblue")

obs_by_dist <- ggplot() + geom_polygon(data = sp_f, aes(long, lat, group = group, fill = num)) + coord_sf(
  labs(fill = "No. of \nObservations") +
  ggtitle("Number of Observations per District")+ scale_fill_gradient(low = "lightblue", high = "navyblue")

#final graph
obs_graph <- obs_by_dist + geom_polygon(data=sp_f,aes(long,lat, group = group), fill = NA, col = "black")

# Construct a Gibbs sampler that cycles through each of the full conditionals and stores
# the results for B = 10,000 iterations.

# I will start by creating the input parameters/vectors/matrices
n <- nrow(rents) # number of data points in Y
m <- nrow(districts.sp) # number of spatial locations in eta
W <- nb2mat(neighbours = district_neighb, style = "B")
#W is matrix of 0's and 1's indicating neighb structure from 2, need style = B to be binary (W is row s
#X and y are given in the data file
D <- diag(rowSums(W)) # diagonal matrix with row sums of W

#is there a typo in this question? row sum up to m, not n?

#how many iterations
B <- 10000

#Prior parameters
b <- 1
a.s2 <- 0.001; b.s2 <- 0.001
a.t2 <- 0.001; b.t2 <- 0.001

#setup and storage and starting values
beta.samps <- matrix(NA, nrow = 12, ncol = B)
beta.samps[,1] <- beta
#coefficients from problem 1

s2.samps <- t2.samps <- rep(NA, B)
#eta.obs.samps <- matrix(NA, nrow = n, ncol = B)
s2.samps[1] <- t2.samps[1] <- 1

eta.obs.samps <- matrix(NA, nrow = m, ncol = B)
eta.obs.samps[,1] <- rep(mean(y), m)

```



```

y <- as.matrix(y)

## MCMC sampler
for(i in 2:B){
  print(i)

  ## beta_obs | Rest
  mu_b <- solve(t(X)%*%X)%*%t(X)%*(y-H%*%as.matrix(eta.obs.samps[,i-1]))
  Sig_b <- s2.samps[i-1]*solve(t(X)%*%X)
  beta.samps[,i] <- rmvnorm(1, mean= mu_b, sigma = Sig_b, method = "svd")

  ## eta | Rest
  mu_e <- solve((t(H)%*%H)/s2.samps[i-1] + (D-W)/t2.samps[i-1])%*%t(H)%*(y-X%*%beta.samps[,i])/s2.samps[i-1]
  Sig_e <- solve(t(H)%*%H/s2.samps[i-1] + (D-W)/t2.samps[i-1])
  eta.pre <- rmvnorm(1, mean = mu_e, sigma = Sig_e, method = "svd")
  #need to adjust for the pairwise difference prior
  eta.obs.samps[,i] <- eta.pre - mean(eta.pre)

  ## s2 | Rest
  a <- 0.001 + (n/2)
  resid <- y-X%*%beta.samps[,i]
  second_part <- H%*%as.matrix(eta.obs.samps[,i])
  b <- 0.001 + 0.5*(t(resid-second_part)%*(resid-second_part))
  s2.samps[i] <- rinvgamma(1, a, b)

  ## t2 | Rest
  c <- 0.001 + ((m-1)/2)
  d <- 0.001 + 0.5*(t(as.matrix(eta.obs.samps[,i]))%*(D-W)%*%as.matrix(eta.obs.samps[,i]))
  t2.samps[i] <- rinvgamma(1, c, d)
}

save(t2.samps, s2.samps, beta.samps, eta.obs.samps, file = "C:/Users/ckell/OneDrive/Penn State/2017-2018/597/spatial_statistics_597/Homework 3/data/mcmc_data.Rsave")

load("C:/Users/ckell/OneDrive/Penn State/2017-2018/597/spatial_statistics_597/Homework 3/data/mcmc_data.Rsave")
## Diagnostics

plot(beta.samps[1,], type = "l")
plot(s2.samps, type = "l")
plot(t2.samps, type = "l")
plot(eta.obs.samps[1,], type = "l")

burnin <- 300
s2.final <- s2.samps[-(1:burnin)]
t2.final <- t2.samps[-(1:burnin)]
beta.final <- beta.samps[-(1:burnin)]
eta.obs.final <- eta.obs.samps[-(1:burnin)]

acf(s2.final)
acf(t2.final)
acf(beta.final[1,])
acf(eta.obs.final[1,])

```

```

effectiveSize(s2.final)
effectiveSize(t2.final)
effectiveSize(beta.final[1,])
effectiveSize(eta.obs.final[1,])

#table with posterior means and credible intervals
b_post_mean <- apply(beta.final, 1, mean)
b_CI <- apply(beta.final, 1, function(x) quantile(x, probs = c(0.025, 0.975)))
est_and_ci <- cbind(b_CI[1,], b_post_mean, b_CI[2,])

est_and_ci <- cbind(est_and_ci, beta)

colnames(est_and_ci) <- c("0.025 Quantile", "Posterior Mean", "0.975 Quantile", "Initial Estimate")
rownames(est_and_ci) <- c("b1", "b2", "b3", "b4", "b5", "b6", "b7", "b8", "b9", "b10", "b11", "b12")

## Find pointwise posterior means and sds
eta.m <- apply(eta.obs.final, 1, mean)
eta.sd <- apply(eta.obs.final, 1, sd)

districts.sp@data$eta.m <- eta.m
districts.sp@data$eta.sd <- eta.sd
#districts.sp@data <- districts.sp@data[,-c(2)]
sp_f <- left_join(sp_f, districts.sp@data)

m_by_dist <- ggplot() + geom_polygon(data = sp_f, aes(long, lat, group = group, fill = eta.m)) + coord_
  labs(fill = "No. of \nObservations") +
  ggtitle("Posterior mean of eta")+ scale_fill_gradient(low = "lightblue", high = "navyblue")

#final graph
m_map <- m_by_dist + geom_polygon(data=sp_f,aes(long,lat, group = group), fill = NA, col = "black")+geom

sd_by_dist <- ggplot() + geom_polygon(data = sp_f, aes(long, lat, group = group, fill = eta.sd)) + coord
  labs(fill = "No. of \nObservations") +
  ggtitle("Standard Dev of Post eta")+ scale_fill_gradient(low = "lightblue", high = "navyblue")

#final graph
sd_map <- sd_by_dist + geom_polygon(data=sp_f,aes(long,lat, group = group), fill = NA, col = "black")+g

```