The negative binomial Indian buffet process

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Joint work with **Daniel M. Roy** (Cambridge)

BNP 9, Amsterdam June 2013

Outline

- 1. The Indian buffet process (IBP; [GG06, GGS07]) induces a distribution on allocations of features to individuals. We are interested in **count extensions of the IBP**, i.e., each individual may have multiple copies of each feature.
- 2. ([TJ07]). The IBP is the combinatorial structure of an exchangeable sequence of Bernoulli processes directed by a beta process base measure.
- 3. We are interested in the combinatorial structure of an exchangeable sequence of negative binomial processes ([BMPJ11, ZHDC12]), directed by a beta process, which we describe as the negative binomial Indian buffet process, a count extension of the IBP.

Plan

- 1. Review the IBP and connect it to the theory of completely random measures.
- 2. Develop a different continuum of Pólya urn schemes perspective for sampling the IBP.
- 3. Develop a method to **sample** a beta negative binomial process using the continuum of Pólya urn schemes intuition.
- 4. Present the corresponding negative binomial IBP.

Let $\alpha, c > 0$. A sequence of customers go through an Indian buffet with an infinite number of dishes.

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 - ▶ takes $\frac{\text{Bernoulli}\left(\frac{m_k}{c+n}\right)}{c+n}$ servings of every previously sampled dish k, where m_k is the total number of servings taken of dish k by the previous n customers;

- ▶ The first customer n = 1 tastes $Poisson(\alpha)$ dishes;
- ▶ The n + 1-st customer:
 - ▶ takes Bernoulli $\left(\frac{m_k}{c+n}\right)$ servings of every previously sampled dish k, where m_k is the total number of servings taken of dish k by the previous n customers;
 - ▶ tastes $Poisson(\alpha \frac{c}{c+n})$ new dishes.

rows = individuals/customers columns = features/dishes

| n = 1 | 1 | 1 | | | | |
|-------|---|---|---|---|---|---|
| n=2 | | 1 | 1 | 1 | | |
| : | 1 | | 1 | | 1 | |
| : | | 1 | | 1 | 1 | 1 |

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| : | | 1 | | 1 | 1 | 1 |

rows = individuals/customers Bernoulli processes columns = features/dishes locations in Ω

| | ω_1 | ω_2 | ω_3 | ω_4 | • | • • |
|-------|------------|------------|------------|------------|---|-----|
| X_1 | 1 | 1 | | | | |
| X_2 | | 1 | 1 | 1 | | |
| X_3 | 1 | | 1 | | 1 | |
| X_4 | | 1 | | 1 | 1 | 1 |

rows = $\frac{\text{individuals/customers}}{\text{columns}}$ = $\frac{\text{features/dishes}}{\text{locations}}$ in Ω

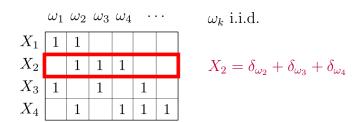
| ω_1 | ω_2 | ω_3 | ω_4 | • | • |
|------------|----------------------|---|---|---|---|
| 1 | 1 | | | | |
| | 1 | 1 | 1 | | |
| 1 | | 1 | | 1 | |
| | 1 | | 1 | 1 | 1 |
| | $\frac{\omega_1}{1}$ | $ \begin{array}{c cccc} \omega_1 & \omega_2 \\ \hline 1 & 1 \\ \hline 1 & 1 \\ \hline 1 & 1 \end{array} $ | $\begin{array}{c cccc} \omega_1 & \omega_2 & \omega_3 \\ \hline 1 & 1 & & \\ & 1 & 1 \\ \hline 1 & & 1 \\ & & 1 \\ \end{array}$ | $\begin{array}{c ccccccccccccccccccccccccccccccccccc$ | $\begin{array}{c ccccccccccccccccccccccccccccccccccc$ |

 ω_k i.i.d.

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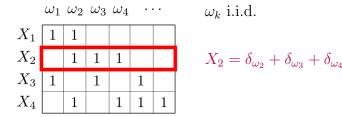
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$$X_{n+1} \mid X_1, \dots, X_n \sim \operatorname{BeP}\left(\frac{c}{c+n}\widetilde{B}_0 + \frac{1}{c+n}\sum_{i=1}^n X_i\right)$$



IBP: completely random measures perspective

Let $(X_n)_{n\in\mathbb{N}}$ satisfy

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Thm ([TJ07]): Then there exists a beta process

$$B \sim \mathrm{BP}(c, \widetilde{B}_0)$$

such that

$$(X_n)_{n\in\mathbb{N}}\mid B\stackrel{iid}{\sim} \operatorname{BeP}(B).$$

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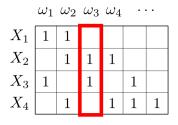
$$\delta_{\omega_1} + \dots$$

$$(X_n)_{n\in\mathbb{N}}\mid B\stackrel{iid}{\sim} \operatorname{BeP}(B)$$

| $\omega_1 \ \omega_2 \ \omega_3 \ \omega_4 \ \cdots$ | | | | | | | | |
|--|---|---|---|---|---|---|--|--|
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$$\delta_{\omega_1} + 2 \delta_{\omega_2} + \dots$$

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$$\delta_{\omega_1} + 2 \, \delta_{\omega_2}$$

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$$\delta_{\omega_1} + 2 \, \delta_{\omega_2} \sim \mathrm{BGP}(\widetilde{B}_0)$$

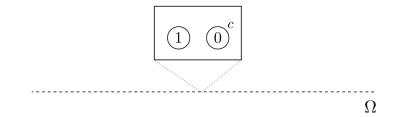
$$(X_n)_{n\in\mathbb{N}}\mid B\stackrel{iid}{\sim} \operatorname{BeP}(B)$$

 $Bernoulli(p) \to geometric(p)$: # successes before the first failure

$$\delta_{\omega_1} + 2 \, \delta_{\omega_2} \sim \mathrm{BGP}(\widetilde{B}_0)$$

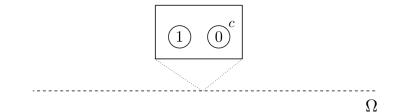
NB(r, p): # successes before the r-th failure

Imagine an infinite number of independent Pólya urn schemes, each with two tables labelled 1 and 0 $\,$



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i.e., a continuum of Pólya urn schemes



How to sample $X_1 \sim \text{BeP}(\widetilde{B}_0)$



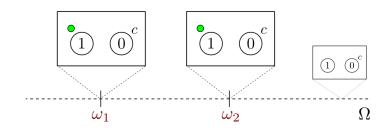
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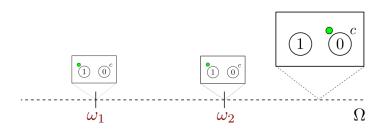
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$$\widetilde{B}_0(\Omega) = \alpha$$

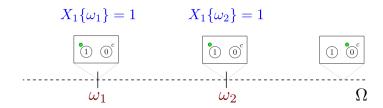
How to sample $X_1 \sim \text{BeP}(\widetilde{B}_0)$

- A $Poisson(\alpha)$ number of urn schemes succeed.
- The remaining urn schemes on the continuum fail.

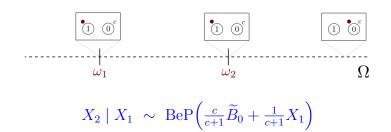


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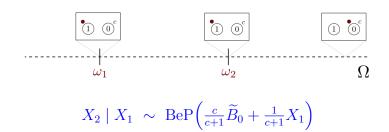
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 X_1 done.

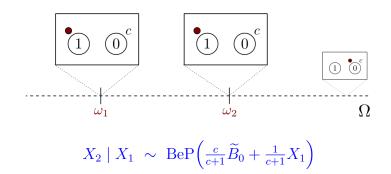


Run the urn schemes forward one more step



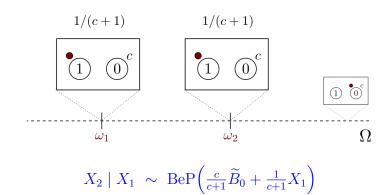
Run the urn schemes forward one more step

 \Rightarrow Consider the previous atoms in X_1



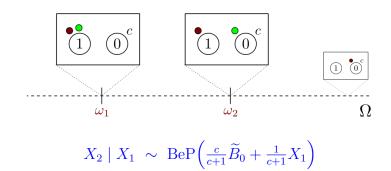
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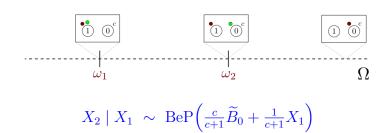
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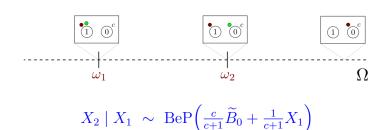


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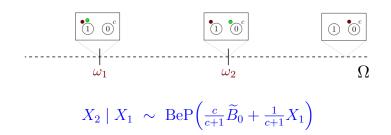
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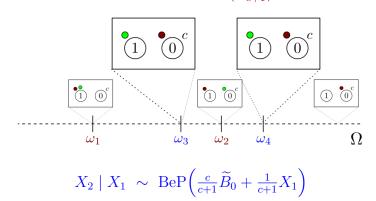
- \Rightarrow Consider the previous atoms in X_1
- \Rightarrow Consider urns **not in** X_1



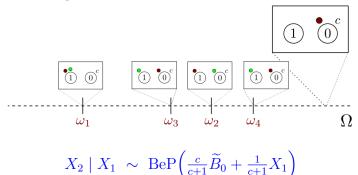
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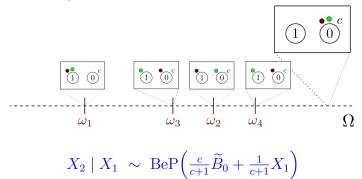
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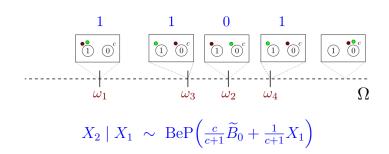
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 X_2 done.



A count extension

Bernoulli(p): distribution of successful trial with success probability p;

NB(r, p): distribution of # successes until r failures.

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NB(r, p): distribution of # successes until r failures.

Intuition: The Bernoulli process runs one trial at every urn scheme. So the *negative binomial process* continues running trials until r failures at each urn scheme.

Key point: Instead of binary indicators at each urn scheme, we get integer-valued counts.

Negative binomial processes are parameterized by a "base measure"

$$B_0 = \widetilde{B}_0 + \sum_{k=1}^{\kappa} b_k \delta_{\omega_k}. \tag{1}$$

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Defn. We call $X \sim \text{NBP}(r, B_0)$ a <u>negative binomial process</u>, when it is a completely random measure with

• fixed component (as defined by [BMPJ11, ZHDC12]):

$$\sum_{k=1}^{\kappa} \zeta_k \delta_{\omega_k}, \qquad \zeta_k \stackrel{ind}{\sim} NB(r, b_k), \quad \kappa \in \mathbb{N} \cup \{\infty\},$$
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• ordinary component (our definition): a Poisson process with intensity $r\widetilde{B}_0$.

Defn. We call $X \sim \text{BNBP}(r, c, B_0)$ a beta negative binomial process when there exists a beta process

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- ▶ The beta process has an infinite number of atoms w.p. one
- ▶ [BMPJ11, ZHDC12] use finite approximations to the beta process in order to produce samples of X

Our goal: give a method to directly sample X without representing the underlying beta process.

Start simple

Consider a NB(r, p) distribution when r = 1, i.e., a geometric distribution

$$NB(1, p) = geometric(p),$$
 (4)

which counts the number of successful trials (with success probability p) before the **first failure**.

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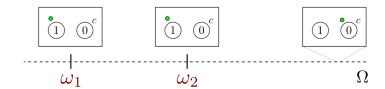
Intuition: Run the urn scheme until the first failure.

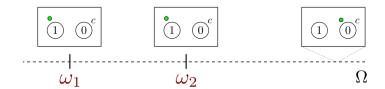


Run all urn schemes at once



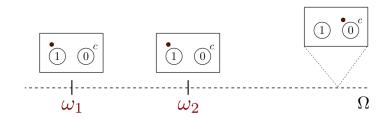
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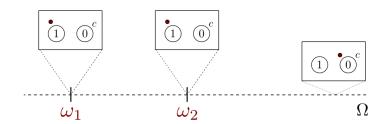


We continue until all urn schemes have failed once!

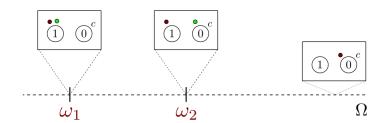
⇒ The remaining urns on the continuum have already failed



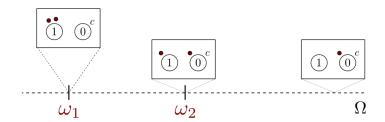
- ⇒ The remaining urns on the continuum have already failed
- \Rightarrow We **complete** the atoms



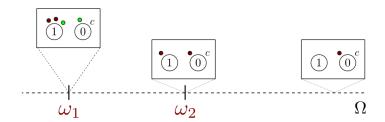
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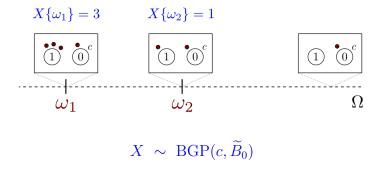


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Ordinary component: BGP

Thm ([HR13]). Let $Y \sim PP(\widetilde{B}_0)$ and let

$$\zeta_s \stackrel{ind}{\sim} \text{beta-geometric}(1, c), \quad s \in \Omega,$$

be independent from Y. Then

$$X = \sum_{c \in Y} (1 + \zeta_s) \delta_s \sim \mathrm{BGP}(c, \widetilde{B}_0).$$

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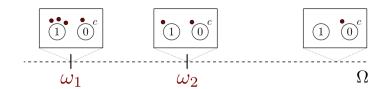
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How to sample $X \sim \text{BNBP}(r, c, \widetilde{B}_0)$



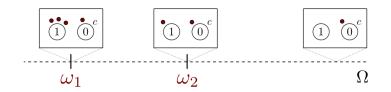
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Start where $X \sim \mathrm{BGP}(c, \widetilde{B}_0)$ left off



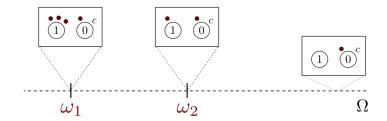
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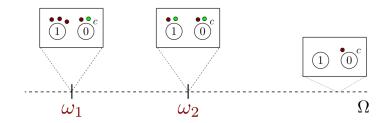
Continue running the urn schemes untill r failures!

Start where $X \sim \mathrm{BGP}(c, \widetilde{B}_0)$ left off \Rightarrow complete the atoms



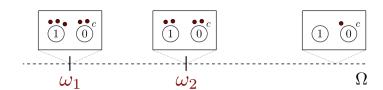
Continue running the urn schemes untill r failures!

Start where $X \sim \mathrm{BGP}(c, \widetilde{B}_0)$ left off \Rightarrow complete the atoms E.g., if r = 2, we stop.



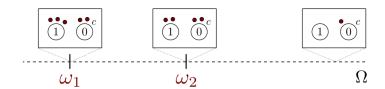
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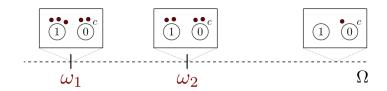
Start where $X \sim \mathrm{BGP}(c, \widetilde{B}_0)$ left off

- \Rightarrow complete the atoms
- \Rightarrow advance the remaining urns on the continuum



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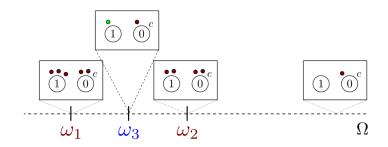
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- A Poisson $\left(\alpha \frac{c}{c+1}\right)$ number of new urns succeed

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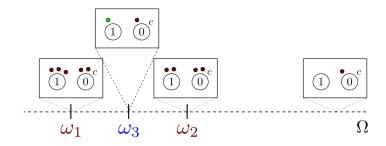
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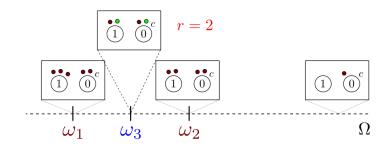
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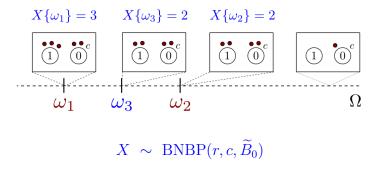
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- Complete the new atom

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- \Rightarrow advance the remaining urns on the continuum



- A Poisson $\left(\alpha \frac{c}{c+1}\right)$ number of new urns succeed
- Complete the new atom



Ordinary component: BNBP

Thm ([HR13]). Let

$$Y_{\ell} \stackrel{ind}{\sim} PP\left(\frac{c}{c+\ell-1}\widetilde{B}_{0}\right), \qquad \ell \in [r] = (1,\ldots,r),$$
 (5)

and let

$$\zeta_{\ell,s} \stackrel{ind}{\sim} \text{beta-NB}(r-\ell+1,1,c+\ell-1), \qquad \ell \in [r], \ s \in \Omega, \quad (6)$$

independent also from (Y_r) . Then

$$X = \sum_{\ell=1}^{r} \sum_{s \in Y_{\ell}} (1 + \zeta_{\ell,s}) \delta_s \sim \text{BNBP}(r, c, \widetilde{B}_0).$$
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Key point: atoms arise in a sequence of r rounds.

- ▶ n = 1: Independently for each of r rounds $\ell = 1, ..., r$,
 - ▶ tastes Poisson $\left(\alpha \frac{c}{c+\ell-1}\right)$ dishes;

- ightharpoonup n = 1: Independently for each of r rounds $\ell = 1, \ldots, r$,
 - ▶ tastes Poisson $\left(\alpha \frac{c}{c+\ell-1}\right)$ dishes;
 - returns to each dish for beta-NB $(r \ell + 1, 1, c + \ell 1)$ additional servings;

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- ▶ subsequent customers $n \ge 2$:
 - ▶ takes beta-NB $(r, m_k, c + (n-1)r)$ servings of each previously tasted dish k, where m_k is the total number of servings taken by first n customers;

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 - independently for each of r rounds $\ell = 1, \ldots, r$,
 - ▶ tries Poisson $\left(\alpha \frac{c}{c+(n-1)r+\ell-1}\right)$ new dishes;

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Probability function

Let $\mathcal{H}_n \equiv \mathbb{Z}_+^n \setminus \{0^n\}$, and for $h \in \mathcal{H}_n$, let M_h count the number of features k where every customer n has h(n) copies of feature k.

Probability function

Let $\mathcal{H}_n \equiv \mathbb{Z}_+^n \setminus \{0^n\}$, and for $h \in \mathcal{H}_n$, let M_h count the number of features k where every customer n has h(n) copies of feature k.

Claim. The probability function for the NB-IBP is given by

$$\frac{\alpha^{K_N}}{\prod_{h \in \mathcal{H}_N} M_h!} \exp\left(-\alpha \sum_{n=1}^N \sum_{\ell=1}^r \frac{c}{c + nr + \ell - 1}\right)$$

$$\prod_{h \in \mathcal{H}_N} \left[\frac{c}{c + Nr} \frac{\Gamma(s(h))\Gamma(c + Nr + 1)}{\Gamma(c + Nr + s(h))}\right]^{M_h}$$

$$\prod {r + h(n) - 1 \choose r - 1}^{M_h}$$

where $s(h) \equiv \sum_{n=1}^{N} h(n)$.

General r > 0

- ▶ The previous urn scheme was only valid for $r \in \mathbb{N}$.
- ▶ Urn schemes for general r > 0 reduces to the case of **fractional** $r \in (0,1)$, which is desirable for some applications.
 - \blacktriangleright We use Poisson process calculus to reduce the r rounds to one slightly more clever round.
 - ► Completions are no longer beta-NB variables, but what we call harmonic mixtures.

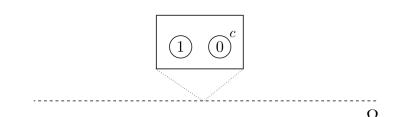
General r > 0

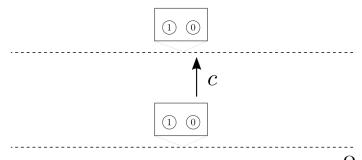
There is an analytical extension of the NB-IBP to general values of r > 0. The probability function looks similar:

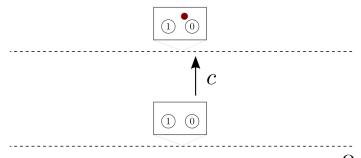
Claim ([HR13]). Let r > 0. The probability function for the NB-IBP is given by

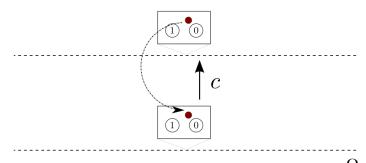
$$\frac{(c \alpha)^{K_N}}{\prod_{h \in \mathcal{H}_N} M_h!} \exp\left(-c \alpha \sum_{n=1}^N [\psi(c + (n+1)r) - \psi(c + nr)]\right) \times \prod_{h \in \mathcal{H}_N} \left[\frac{\Gamma(s(h))\Gamma(c + Nr)}{\Gamma(c + Nr + s(h))} \prod_{n \in \mathbb{N}} \frac{\Gamma(r + h(n))}{h(n)! \Gamma(r)}\right]^{M_h}.$$
(8)

Key difference: No longer a concept of *rounds*.



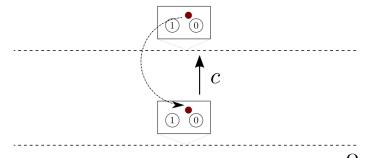






We can also characterize the combinatorial structure of an exchangeable sequence of negative binomial processes, directed by a generalized beta process [Roy13], parametrized by a measurable family of EPPFs.

A continuum of generalized Blackwell-MacQueen urn schemes



Special cases include:

- ▶ hierarchies of beta processes [TJ07]
- ▶ stable beta processes [TG09, BJP12]
 - corresponding power-law NB-IBP
- hierarchies of stable beta processes

Probability function

Claim. The probability function for the NB-IBP is given by

$$\frac{\alpha^{K_N}}{\prod_{h \in \mathcal{H}_N} M_h!} \exp\left(-\alpha \sum_{n=1}^N \sum_{\ell=1}^r \frac{c}{c+nr+\ell-1}\right)$$

$$\prod_{h \in \mathcal{H}_N} \left[\frac{c}{c+Nr} \frac{\Gamma(s(h))\Gamma(c+Nr+1)}{\Gamma(c+Nr+s(h))}\right]^{M_h}$$

$$\prod_{n \in \mathbb{N}} \binom{r+h(n)-1}{r-1}$$

Probability function from a generalized beta

Claim. The probability function for the generalized NB-IBP is given by

$$\frac{\alpha^{K_N}}{\prod_{h \in \mathcal{H}_N} M_h!} \exp\left(-\alpha \sum_{n=1}^N \sum_{\ell=1}^r \mathbb{P}(K_{nr+\ell} > K_{nr+\ell-1})\right)$$

$$\prod_{h \in \mathcal{H}_N} \left[\mathbb{P}(K_{Nr+1} > K_{Nr}, Z_{Nr+s(h)} = \dots = Z_{Nr+1}) \right] \prod_{n \in \mathbb{N}} {r+h(n)-1 \choose r-1}^{M_h}$$

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$$\prod_{h \in \mathcal{H}_N} \left[\mathbb{P}(K_{Nr+1} > K_{Nr}, Z_{Nr+s(h)} = \dots = Z_{Nr+1}) \right.$$

$$\left. \prod_{n \in \mathbb{N}} {r+h(n)-1 \choose r-1} \right]^{M_h}$$

 $\mathbb{P}(K_{n+1} > K_n)$: probability of a new table at step n+1 ... $\mathbb{P}(K_{n+1} > K_n, Z_{n+m} = \cdots = Z_{n+1})$: probability of also staying there for next m steps

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