

Time Series Homework 3

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Problem 11.

(a)

$$\begin{aligned} X_t &= 0.5X_{t+1} - 0.5Z_{t+1} \\ &= 0.5(X_{t+2} - 0.5Z_{t+2}) - 0.5Z_{t+1} \\ &\vdots \\ &= -\sum_{k=0}^{\infty} (0.5)^{k+1} Z_{t+1+k}, \end{aligned}$$

since stationary process X_t has a finite variance so it goes away as k goes to infinity.

(b)

mean

$$\begin{aligned} E[Z_t^*] &= 0.25EZ_t - E\frac{3}{4} \lim_{n \rightarrow \infty} \sum_{j=1}^n 2^{-j} Z_{t+j} \\ &= 0.25EZ_t - \lim_{n \rightarrow \infty} E\left[\sum_{j=1}^n 2^{-j} Z_{t+j}\right] = 0, \end{aligned}$$

where we use dominated convergence theorem(DCT) to interchange lim and expectation.

variance, $h=0$

$$\begin{aligned} E[Z_t^*]^2 &= 0.25^2 EZ_t^2 + E9/16 \lim_{n \rightarrow \infty} \sum_{j=1}^n Z_{t+j}^2 (1/4)^j \\ &= 0.25^2 \sigma^2 + 9/16 * \lim_{n \rightarrow \infty} \sigma^2 \sum_{j=1}^n (1/4)^j \\ &= \sigma^2 (0.25^2 + \frac{3}{16}). \end{aligned} \tag{DCT}$$

covariance, $h > 1$

$$\begin{aligned} EZ_t^* Z_{t+h}^* &= -3/42^{-h} E[Z_{t+h}^2] + 9/16 \sum_{j=1}^{\infty} 2^{-2j-h} EZ_{t+h+j}^2 \\ &= -3/42^{-h} \sigma^2 + 3/16 * 2^{-h} \sigma^2 = -9/16 * 2^{-h} \sigma^2. \end{aligned} \tag{DCT}$$

(c)

Using the formula in (a), we can write RHS $0.5X_{t-1} + Z_t^*$ as

$$\begin{aligned}
0.5X_{t-1} + Z_t^* &= -0.5 \sum_{k=0}^{\infty} (0.5)^{K+1} Z_{t+k} + 1/4 Z_t - 3/4 \sum_{k=0}^{\infty} (0.5)^{k+1} Z_{t+k+1} \\
&= -1/4 Z_t - 1/4 \sum_{k=0}^{\infty} (0.5)^{k+1} Z_{t+k+1} + 1/4 Z_t - 3/4 \sum_{k=0}^{\infty} (0.5)^{k+1} Z_{t+k+1} \\
&= - \sum_{k=0}^{\infty} (0.5)^{K+1} Z_{t+k+1},
\end{aligned}$$

which is the X_t .

Problem 12

(a)

$$\gamma_w(h) = E\phi(B)X_t\phi(B)X_{t+h} = \phi(B)^2\gamma_x(h) = \phi(B)^2\gamma_y(h),$$

where $\gamma_y(h) = 0$ for $q > 0$ because $\{Y_t\}$ is ARMA(p,q) process. To see this, let $M_t = \phi(B)Y_t$, then $\{M_t\}$ is M(q) process, which means $\gamma_M(h) = 0, \forall h > q$. Then

$$E(Y_t \cdots)(Y_{t+h} \cdots) = 0, \forall h > q,$$

which also implies $EY_tY_{t+h} = 0$ for $h > q$.

(b)

From (a), $\{W_t\}$ is zero-mean stationary process with autocovariance $\gamma(h) = 0$ for $|h| > q$ and $\gamma(q) \neq 0$. By proposition 3.3.1, $\{W_t\}$ is a MA(q) process, which means $\{X_t\}$ is ARMA(p,q) process.

Problem 13

(a)

By definition, $\hat{\gamma}(\cdot)$ is non-negative definite if and only if

$$\sum_{i,j=1}^n a_i \hat{\gamma}(t_i - t_j) a_j \geq 0,$$

for any $a = (a_1, \dots, a_n)^\top$ and $\{t_1, \dots, t_n\} \subset Z$ for given n. We can rewrite

$$\begin{aligned}
&\sum_{i,j=1}^n a_i \hat{\gamma}(t_i - t_j) a_j \\
&= a^\top \Gamma a,
\end{aligned}$$

where Γ is a $n \times n$ matrix with $\Gamma_{ij} = \hat{\gamma}(|i - j|)$. Then

$$a^\top \Gamma a = \text{var}(a^\top Z_t) \geq 0, \quad Z_t = (X_1 - EX_1, \dots, X_n - EX_n)^\top.$$

Therefore Γ is non-negative definite, which means $\hat{\gamma}(h)$ is non-negative definite.

Assume $\widehat{\gamma}(h) > 0$ and Γ_n is singular for some fixed n . This means the matrix is not a full rank and there exists $1 \leq r \leq n$ and (a_1, \dots, a_r) such that

W.L.O.G, we can assume $EX_t = 0$, then

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(b)

When $h > \max(p, q + 1)$, acf behaves like AR(p),

$$\psi_j = \psi_{j-1}\phi_1 + \cdots \psi_{j-p}\phi_p.$$

This implies that

$$\gamma(h) - \phi_1\gamma(h-1) - \cdots - \phi_p\gamma(h-p) = 0.$$

By using Yule-Walker equations, this has a solution of the form

$$\gamma(h) = A_1\alpha_1^h + \cdots + A_p\alpha_p^h \leq C * \max\{\alpha_1, \dots, \alpha_p\}^h,$$

where $\alpha_1, \dots, \alpha_p$ are the roots of $x^p - \phi_1x^{p-1} - \cdots - \phi_p = 0$, with $|\alpha_i| < 1$.

Therefore $\gamma(h)$ exponentially decays.