

Machine Learning for Physicists: Recitation Notes

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1 Review of Probability

Definition 1 (Conditional Probability)

$$\mathbb{P}[A|B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]} \quad (1.1)$$

Notice that $\mathbb{P}[A \cap B] = \mathbb{P}[B \cap A]$, this allows us to relate $\mathbb{P}[A|B]$ and $\mathbb{P}[B|A]$.

$$\mathbb{P}[A|B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]} \quad (1.2)$$

$$= \frac{\mathbb{P}[B \cap A]}{\mathbb{P}[B]} \quad (1.3)$$

$$\boxed{\mathbb{P}[A|B] = \frac{\mathbb{P}[B|A] \mathbb{P}[A]}{\mathbb{P}[B]}} \quad (1.4)$$

This is **Bayes' Formula**.

Definition 2 (Probability Density Function) A function with the following properties is a **probability density**

- *Positive:* $p : \mathcal{X} \rightarrow \mathbb{R}_{\geq 0}$
- *Normalized:* $\int_{\mathcal{X}} p(x) dx = 1$

It is interpreted as the probability of observing an event $A \subset \mathcal{X}$ as

$$\mathbb{P}[x \in A] = \int_{A \subset \mathcal{X}} p(x) dx \quad (1.5)$$

The nice part of densities is that you can compute statistics with that. I.e. what's the mean, variance.

$$\mathbb{E}_{x \sim p}[f(x)] = \int_{\mathcal{X}} f(x) p(x) dx \quad (1.6)$$

Definition 3 (Characteristic Function) Consider the probability distribution p_X . It has an associated **characteristic function** φ_X which is its Fourier Transform

$$\varphi_X(k) = \int_{\mathbb{R}} e^{ikx} p(x) dx = \mathbb{E}_{x \sim p}[e^{ikx}] \quad (1.7)$$

2 Review of Statistics & Loss Functions

In machine learning, we adjust a model's parameters θ to minimize a loss function $\mathcal{L}(\theta)$. There's a bunch so I think it's nice to hear where they come from. We'll cover

- Mean squared error (MSE)

$$\mathcal{L}(\theta) = \sum_{i=1}^n \|y_i - f_{\theta}(x_i)\|^2$$

- Cross entropy

$$\mathcal{L}(\theta) = \sum_{i=1}^n \|\cdot\|$$

- MSE + L2 Regularization (Ridge)

$$\mathcal{L}(\theta) = \sum_{i=1}^n \|y_i - f_{\theta}(x_i)\|_2^2 + \lambda \|\theta\|_2^2$$

2.1 Maximum Likelihood Inference & Mean Squared Error

Say you have the dataset $\mathcal{D} = \{(y_i, x_i)\}_{i=1}^n$ (which we assume you observed in an iid way). You believe that y_i is a noisy observation of some model $f_\theta(x_i)$. Your objective is to come up with the "best" estimate of the parameter θ which matches the data \mathcal{D} ... You think about it for some while, and realize you maximize the probability of seeing the data for a given θ . This is **maximum likelihood estimation (MLE)**.

To illustrate this method (and all others), we have to assume a particular model. So let's say you believe the noise is additive & gaussian:

$$y_i = f_\theta(x_i) + \epsilon_i, \text{ where } \epsilon_i \sim_{iid} \mathcal{N}(0, \mathbb{I}) \quad (2.1)$$

Since ϵ_i is a random variable, you can interpret y_i as a random variable as well.

$$y_i \sim \mathcal{N}(f_\theta(x_i), \mathbb{I}) \quad (2.2)$$

$$p(y_i|\theta) \propto \exp\left(-\frac{1}{2}(y_i - f_\theta(x_i))^2\right) \quad (2.3)$$

$$\log p(y_i|\theta) = -\frac{1}{2}(y_i - f_\theta(x_i))^2 + \text{Constant w.r.t. } \theta \quad (2.4)$$

I've wrote the log prob for reasons that will become clear in a moment.

Note you have more data $\{(y_i, x_i)\}_{i=1}^n$ (which is all iid), so you actually have a joint distribution.

$$p(y_1, \dots, y_n|\theta) = \prod_i p(y_i|\theta) \quad (2.5)$$

We'll call this the **likelihood** $L(\theta)$ (that is the likeliness / probability of seeing the data given a configuration of model parameters). For MLE, you choose $\hat{\theta}$ which maximizes the likelihood. However arg max of a product of functions is quite difficult, we can compose the function w/ a monontonic function, and that leaves the arg max invariant.

$$\log L(\theta) = \log p(y_1, \dots, y_n|\theta) \quad (2.6)$$

$$= \sum_i \log p(y_i|\theta) \quad (2.7)$$

$$\propto \sum_i (y_i - f_\theta(x_i))^2 \quad (2.8)$$

This recovers the MSE loss.

2.2 Cross Entropy & Another MLE

2.3 L2 Regularization

In Bayesian statistics, instead of asking what's the probability of seeing the data given a model parameter, we ask *what's the probability of seeing a model parameter given the data?* We can formalize the inverse question using Bayes' theorem

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)} \quad (2.9)$$

- $p(\theta)$ is your prior. It encodes your prior beliefs into the distribution.

- $p(y|\theta)$ is the likelihood (from the previous sections)
- $p(\theta|y)$ is the posterior. It accounts for your prior beliefs & what the data says (likelihood).
- $p(y)$ is the evidence. I won't say much about it today.

If you ask, what's the parameter maximizes the posterior (probability of seeing a parameter given the data), this is called **maximum a posteriori estimation (MAP)**.

$$\hat{\theta}_{MAP} = \arg \max_{\theta} p(\theta|y) \quad (2.10)$$

For an example, let's assume we have the additive noise model

$$y_i = f_{\theta}(x_i) + \epsilon_i \quad (2.11)$$

and that you believe the weights should look distributed according to a Gaussian

$$p(\theta) = \mathcal{N}(0, \lambda^{-1}\mathbb{I}) \quad (2.12)$$

You can see that log posterior has the form

$$\log p(\theta|y) = \sum_i \|y_i - f_{\theta}(x_i)\|^2 + \lambda \|\theta\|^2 \quad (2.13)$$

2.4 Minimizing the loss function

- The value of the MSE, in a traditional statistics setting, tells you about the uncertainty quantification of the model. However ML models tend to not obey this.
- Difficulty of optimizing via oracle access.
- However! Do you even want to perfectly minimize the loss function? Memorization.

3 Linear Regression

Consider making iid noisy observations of data $\{(x_i, y_i)\}_{i=1}^n$, where $x_i \in \mathbb{R}^p$ and $y_i \in \mathbb{R}$. We'll assume that the noise is additive, that is

$$y_i = f(x_i) + \epsilon_i \quad (3.1)$$

where $f(x) = \beta^T x$ is the model (we've assumed it's linear for this discussion) and the noise is gaussian $\epsilon_i = \mathcal{N}(0, \sigma_i^2)$ (which is another assumption for this discussion). Since ϵ_i is a random variable, this implies that y_i is also a random variable

$$y_i|\beta = \mathcal{N}(y_i; \beta^T x_i, \sigma_i^2) \quad (3.2)$$

This is just one observation, but in fact, we have a joint distribution $p(y|x) \equiv p(y_1, \dots, y_N|x_1, \dots, x_N)$ over all observations, which we'll call the **likelihood**. Since observations are iid, it factorizes.

$$p(y|\beta) = \prod_i p(y_i|x_i) \quad (3.3)$$

Your task is to find the β which "best" describes the data. I'll note that "best" is subjective and we'll discuss consequences of this later.

3.1 Frequentist, Maximum Likelihood Estimator

One method is **maximum likelihood estimation**, that is you select the parameters which is the global maximizer of the likelihood. Why? Just read off what you're doing: adjust β s.t. the probability of having this combination of y 's (given x 's) is highest.

Apart from being very intuitive, there are also strong theoretical guarantees (which I won't have time to prove) (Notation: when I generically talk about model parameters, we use θ)

- Consistency: $\lim_{n \rightarrow \infty} \hat{\theta}_n = \theta$
- Normality: $\hat{\theta}_n \sim \mathcal{N}(0, \mathcal{I})$ (where \mathcal{I} is the fisher information matrix)
- Efficiency: $\text{Var}(\hat{\theta}) \geq 1/\mathcal{I}(\theta)$.

Since $\arg \max$ is invariant under compositions of monotonic functions, we can maximize the log-likelihood which emits a nicer function

$$\log p(y|x) = \sum_i \log p(y_i|x_i) \quad (3.4)$$

$$= \sum_i \log \mathcal{N}(y_i; \beta^T x_i, \sigma_i^2) \quad (3.5)$$

$$= \sum_i -\frac{1}{2} \frac{(y_i - \beta^T x_i)^2}{\sigma_i^2} + \text{Constant} \quad (3.6)$$

A small comment, this is why you "minimize the squared error" when fitting straight lines in lab, you have been secretly doing maximum likelihood inference this whole time. Notice this is a quadratic form, so you can rewrite it using matrix multiplication

$$\sum_{i=1}^n (y_i - \beta^T x_i)^2 / \sigma_i^2 = (y - X\beta)^T \Sigma^{-1} (y - X\beta) \quad (3.7)$$

$$\text{where: } y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \in \mathbb{R}^n, \quad (3.8)$$

$$\beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_p \end{pmatrix} \in \mathbb{R}^p \quad (3.9)$$

$$\Sigma = \text{diag}(\sigma_1^2, \dots, \sigma_n^2) \in \mathbb{R}^{n \times n} \quad (3.10)$$

$$X = \begin{pmatrix} - & x_1^T & - \\ - & x_2^T & - \\ & \vdots & \\ - & x_n^T & - \end{pmatrix} \in \mathbb{R}^{n \times p} \quad (3.11)$$

From here we can find the argmax of the quantity

$$0 = \left. \frac{\partial \log p(y|x)}{\partial \beta} \right|_{\beta=\hat{\beta}} = X^T \Sigma^{-1} (y - X\beta) \quad (3.12)$$

$$\implies X^T \Sigma^{-1} y = X^T \Sigma^{-1} X \hat{\beta} \quad (3.13)$$

$$\boxed{\hat{\beta}_{MLE} = (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1} y} \quad (3.14)$$

and find the maximum likelihood estimate for β .

Now we can talk about inference. Say your boss gives you new data X_* , and you're asked what is the corresponding \hat{y}_* . You'll report back

$$\boxed{\hat{y}_* = X_* \hat{\beta}_{MLE}} \tag{3.15}$$

3.2 Bayesian Linear Regression

In the Bayesian framework, you're asked what is the probability of seeing the model parameters *given* the data $p(\beta|y)$. You can calculate this using Bayes's formula

$$p(\beta|y) = \frac{p(y|\beta)p(\beta)}{p(y)} \tag{3.16}$$

4 Double Descent

4.1 Soft Inductive Biases

Another way to conceptualize this is **soft inductive biases** (see Andrew Gordon Willson's paper <https://arxiv.org/pdf/2503.02113>).

5 Training Large Models

5.1 Transformers

5.2 μ P Optimizer

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6 Geometric Deep Learning

7 DDPM & Score Based Diffusion

8 Stochastic Interpolants