

# 2. Probability and Statistical Inference

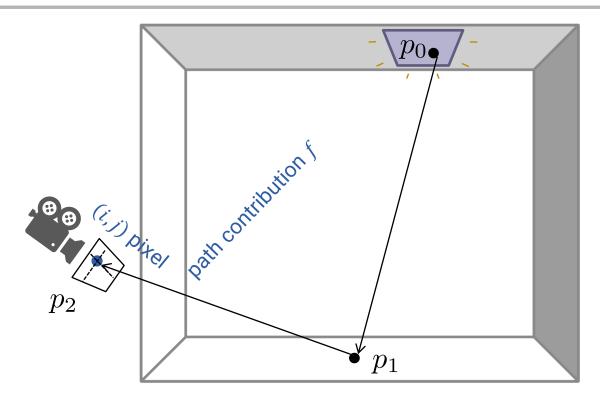
Physically Based Rendering

Shinyoung Yi (이신영)



## Preview: a simple path tracing





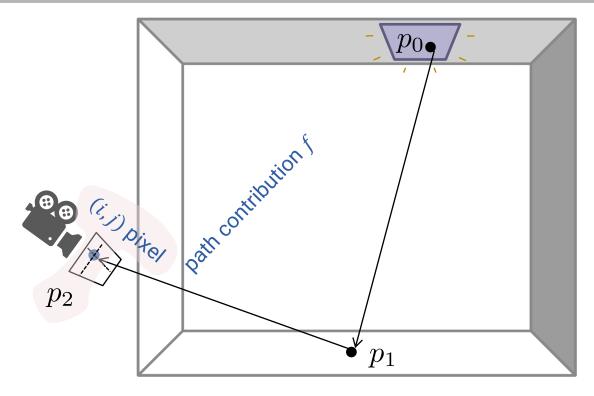
### **Terminology and convention:**

- Notation / actual light:  $p_0 \rightarrow p_1 \rightarrow p_2$
- Computation:  $p_2 \rightarrow p_1 \rightarrow p_0$
- # of bounces < depth < # of vertices</li>

2

3



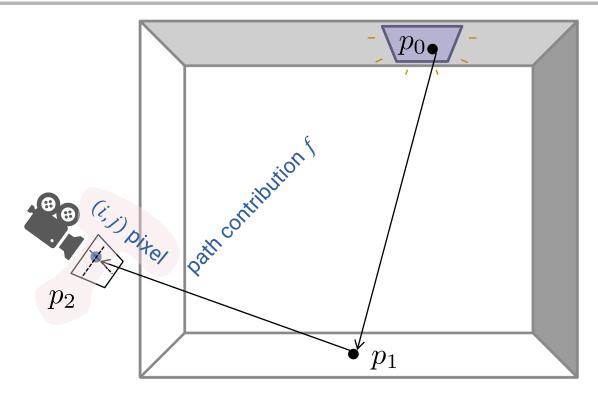


### **Notation:**

$$\bullet \quad \widehat{\omega}_{p_0 p_1} \coloneqq \frac{p_1 - p_0}{\|p_1 - p_0\|}$$

$$I^{(i,j)} = \int_{\mathbb{S}} W_e^{(i,j)} (\hat{\omega}) L(p_2, \hat{\omega}) d\hat{\omega}$$



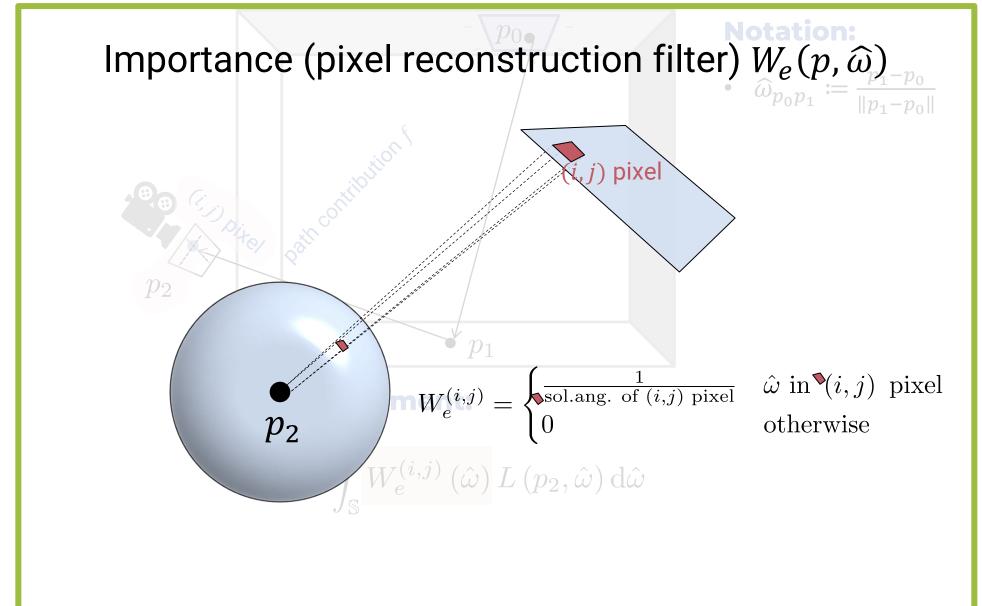


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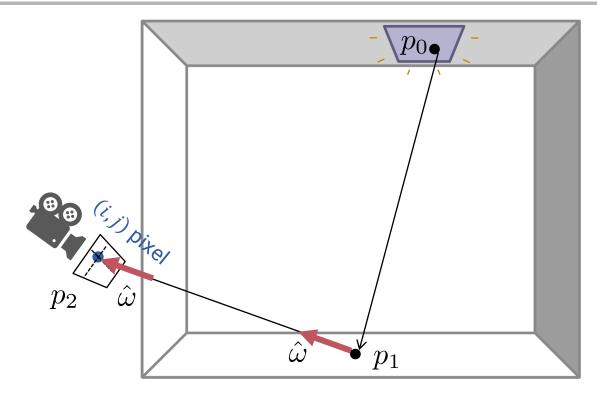
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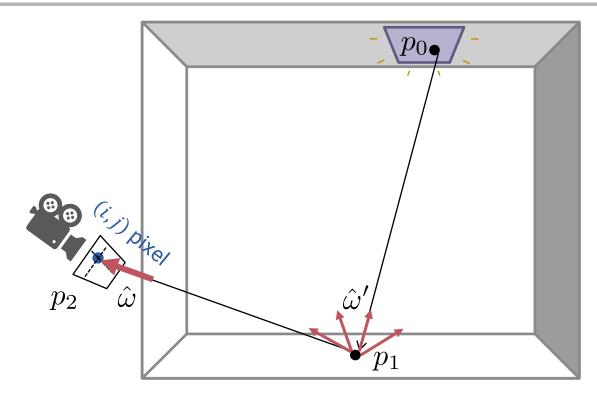
- $\bullet \quad \widehat{\omega}_{p_0 p_1} \coloneqq \frac{p_1 p_0}{\|p_1 p_0\|}$
- $r: \mathbb{R}^3 \times \mathbb{S}^2 \to \mathbb{R}^3$ 
  - ray intersection

$$I^{(i,j)} = \int_{\mathbb{S}} W_e^{(i,j)} \left(\hat{\omega}\right) L\left(p_2, \hat{\omega}\right) d\hat{\omega}$$

$$L\left(p_2, \hat{\omega}\right) = L\left(p_1, \hat{\omega}\right)$$

$$p_1 = r\left(p_2, -\hat{\omega}\right)$$





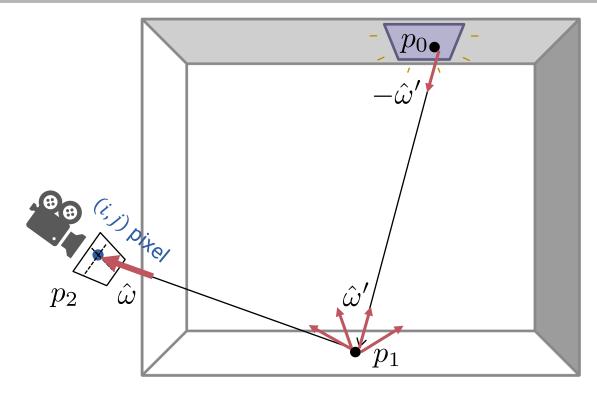
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$$I^{(i,j)} = \int_{\mathbb{S}} W_e^{(i,j)}(\hat{\omega}) L(p_1, \hat{\omega}) d\hat{\omega}$$

$$L(p_1, \hat{\omega}) = L_e(p_1, \hat{\omega}) + \int_{\mathbb{S}} L^{(\text{in})}(p_1, \hat{\omega}') \rho(p_1, \hat{\omega}', \hat{\omega}) |\hat{n} \cdot \hat{\omega}'| d\hat{\omega}'$$





#### **Notation:**

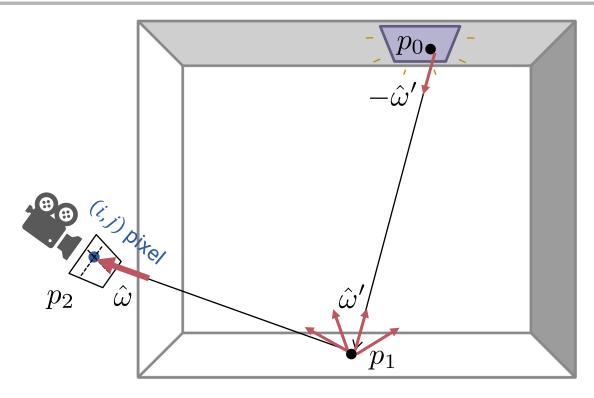
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#### **Notation:**

- $\bullet \quad \widehat{\omega}_{p_0 p_1} \coloneqq \frac{p_1 p_0}{\|p_1 p_0\|}$
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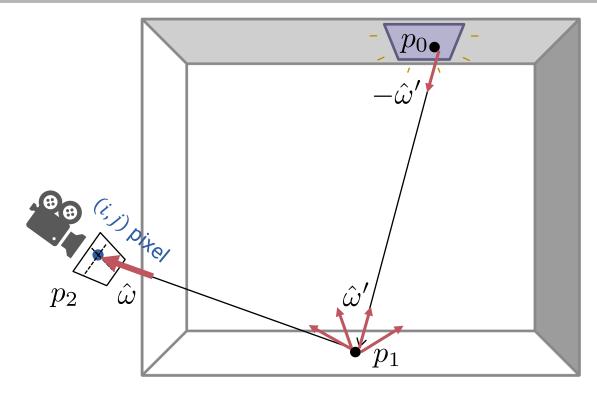
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$$L\left(p_1,\hat{\omega}\right) = L_e\left(p_1,\hat{\omega}\right) + \int_{\mathbb{S}} L^{(\mathrm{in})}\left(p_1,\hat{\omega}'\right) \rho\left(p_1,\hat{\omega}',\hat{\omega}\right) \left|\hat{n}\cdot\hat{\omega}'\right| d\hat{\omega}'$$

$$L\left(p_0,-\hat{\omega}'\right) = L_e\left(p_0,-\hat{\omega}'\right) + \int_{\mathbb{S}} L^{(\mathrm{in})}\left(p_0,\hat{\omega}''\right) \rho\left(p_0,\hat{\omega}'',-\hat{\omega}'\right) \left|\hat{n}\cdot\hat{\omega}''\right| d\hat{\omega}''$$

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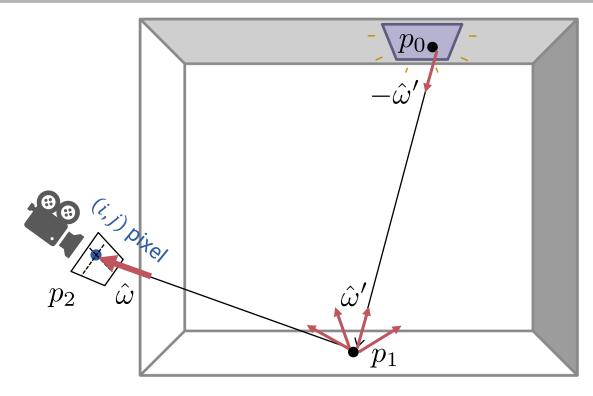
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$$L(p_0, -\hat{\omega}') = L_e(p_0, -\hat{\omega}')$$





### **Notation:**

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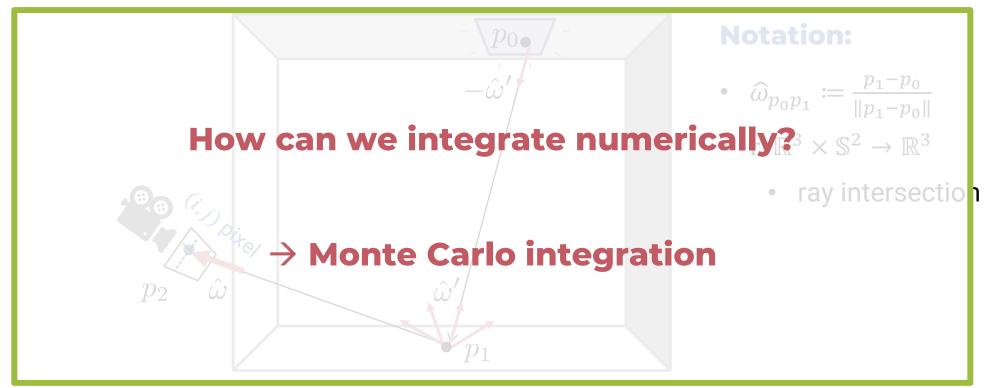
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$$p_1 = r(p_2, -\hat{\omega})$$

$$p_0 = r(p_1, \hat{\omega}')$$





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## **Basic Math: Sets and Functions**

### Sets and functions



Suppose  $A, B, X, \cdots$  are sets and  $f, g, \cdots$  are functions

- $f: X \to Y$ ,  $A \subset X$  then  $f(A) := \{f(x) | x \in A\}$  is the **image of** A **under** f
- $B \subset Y$  then  $f^{-1}(B) := \{x | f(x) \in B\}$  is the **preimage of** B **under** f
  - Even if an inverse function does not exists, preimages are well defined.



## Probability

## What is probability?

## random sample

probability

observation

distribution

random variable

## **Probability overview**



- Sample space (probability space)  $\Omega$
- Event *E*
- Probability P

- Random variable X
- CDF of *X*
- PDF of *X*

## Probability overview



- Sample space (probability space)  $\Omega$ : a set of all things that the outcome can be.
- Event  $E \subset \Omega$
- Probability  $P(E) \in \mathbb{R}_{\geq 0}$

$$- P(E_1 \cup E_2) = P(E_1) + P(E_2) \text{ if } E_1 \cap E_2 = \phi$$

- $-P(\Omega)=1$
- Random variable  $X: \Omega \to \mathbb{R}$  (codomain  $\mathbb{R}^n$  or  $\mathbb{S}^2$  is also okay)
  - We can write an event in a way as  $\{X \le x\} := X^{-1}(\{y \in \mathbb{R} | y \le x\})$
- CDF of X:  $F_X(x) = P(\{X \le x\})$
- PDF of X: some function  $p_X: \mathbb{R} \to \mathbb{R}_{\geq 0}$  s.t.  $P(a \leq x \leq b) = \int_a^b p_X(x) dx$ 
  - $p_X(x) = \frac{\mathrm{d}F_X(x)}{\mathrm{d}x}$  for continuous random variable

## Probability overview: summary



Event 
$$\subseteq \Omega$$
 sample space

Probability:  $0 \le P(E) \le 1$ 

Random variable  $X: \Omega \to \mathbb{R}$ 

## Probability properties



- For a random variable  $X: \Omega \to \mathbb{R}$ 
  - Expectation:  $\mathbb{E}[X] := \int_{\mathbb{R}} x p_X(x) dx$
  - Variance:  $Var(X) := \mathbb{E}[(X \mathbb{E}[X])^2] = \mathbb{E}[X^2] \mathbb{E}[X]^2$

- For a function  $f: \mathbb{R} \to \mathbb{R}$ 
  - $f(X) = f \circ X$ : Ω →  $\mathbb{R}$  is also a random variable

$$- p_{f(X)}(y) = \sum_{x \in f^{-1}(y)} \frac{p_{X(x)}}{|f'(x)|}$$

## Probability properties



- Independence: definition
  - Events  $A, B \subset \Omega$  are called independent if  $P(A \cap B) = P(A)P(B)$
  - Random variables  $X, Y: \Omega \to \mathbb{R}$  are call independent if for any  $x, y \in \mathbb{R}$ ,  $\{X \le x\}$  and  $\{Y \le y\}$  are independent events

- Independence: properties
  - For independent random variables X and Y,

$$F_{X,Y}(x,y) = F_X(x)F_Y(y)$$

$$p_{X,Y}(x,y) = p_X(x)p_Y(y)$$

$$\mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)]\mathbb{E}[g(Y)]$$

## Why is probability so confusing?





### Choice 1

Sample space:  $\{h, t\}$ 

$$H = \{h\}, T = \{t\}$$

$$P[{h}] = \frac{1}{2}, P[{t}] = \frac{1}{2}$$

### Choice 2

Sample space:  $\{\theta | 0 \le \theta \le 2\pi\}$ 

$$H = \{\theta | 0 \le \theta \le \pi\}, T = \{\theta | \pi < \theta \le 2\pi\}$$

$$P[\{\theta | 0 \le \theta_1 < \theta \le \theta_2 \le 2\pi\}] = \frac{\theta_2 - \theta_1}{2\pi}$$

### Event: *H*, *T*

Probability:  $P[H] = \frac{1}{2}$ ,  $P[T] = \frac{1}{2}$ 

### Choice 3

Sample space:  $\{\mathbf{R} | \mathbf{R} \in SO(3)\}$ 

$$H = \left\{ \mathbf{R} \middle| 0 \le \mathbf{R}\hat{z} \le \frac{\pi}{2} \right\}, T = \left\{ \mathbf{R}\hat{z} \middle| \frac{\pi}{2} < \mathbf{R}\hat{z} \le 1 \right\}$$

$$P[A \subset SO(3)] = \frac{1}{8\pi^2} \int_A \sin\beta \, d\alpha d\beta d\gamma$$



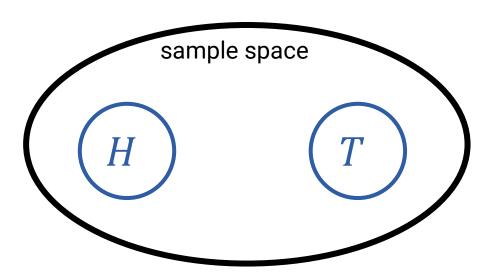


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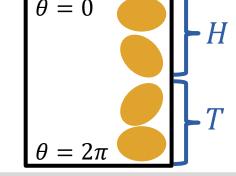


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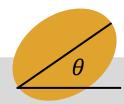
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Choice 2

Sample space: 
$$\{\theta | 0 \le \theta \le \pi\}$$
  
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sample space







Choice 1

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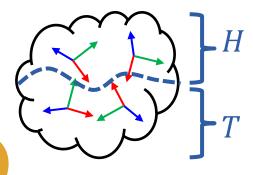
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sample space = SO(3)



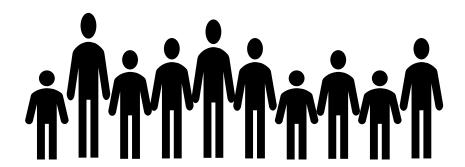
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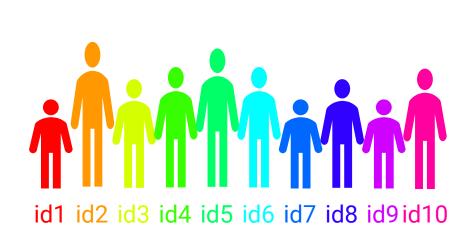


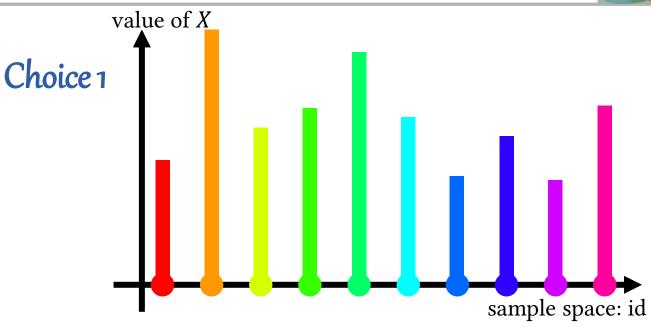


Event: each person

Probability:  $\frac{1}{N}$ 



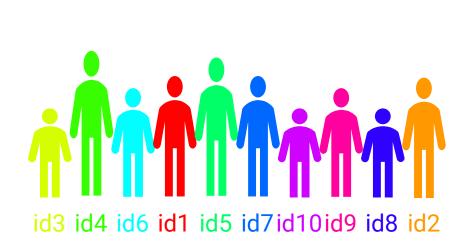




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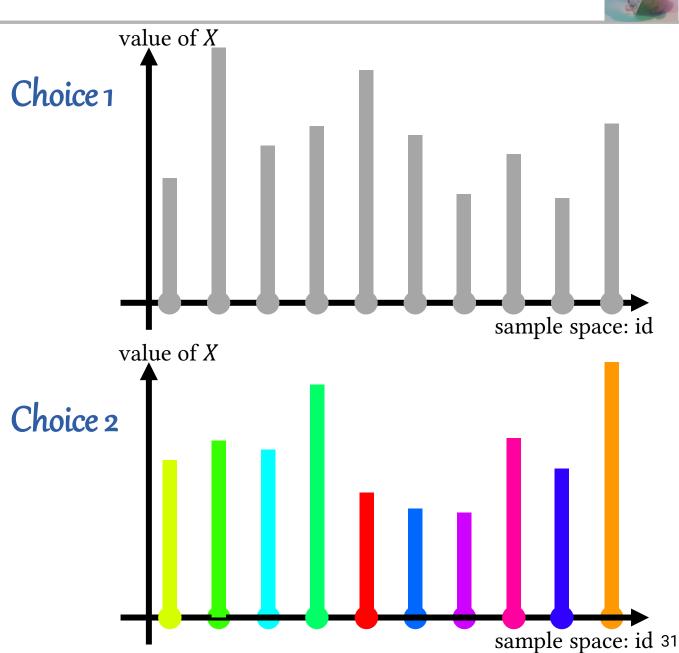
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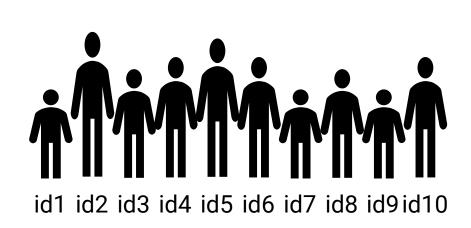
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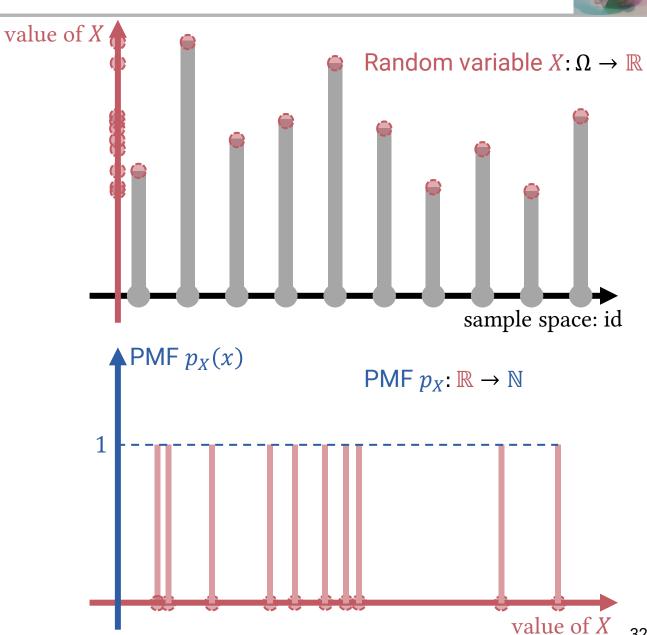
## Two ways for viewing r.v. as a "function"





Event: each person

Probability:  $\frac{1}{N}$ 



## Underlying sample space is often implicit



Suppose there is a random variable  $X \sim U([0,1])$ .

• • •

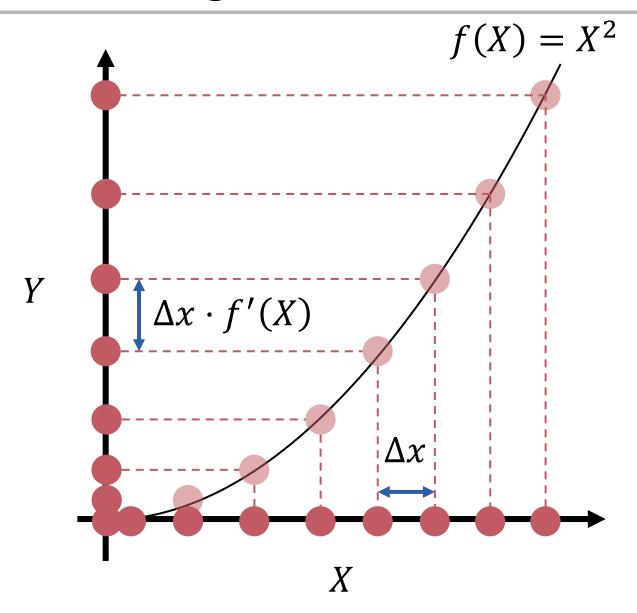


We can consider the underlying sample space is

We have a r.v. X:

## PDF change of variables



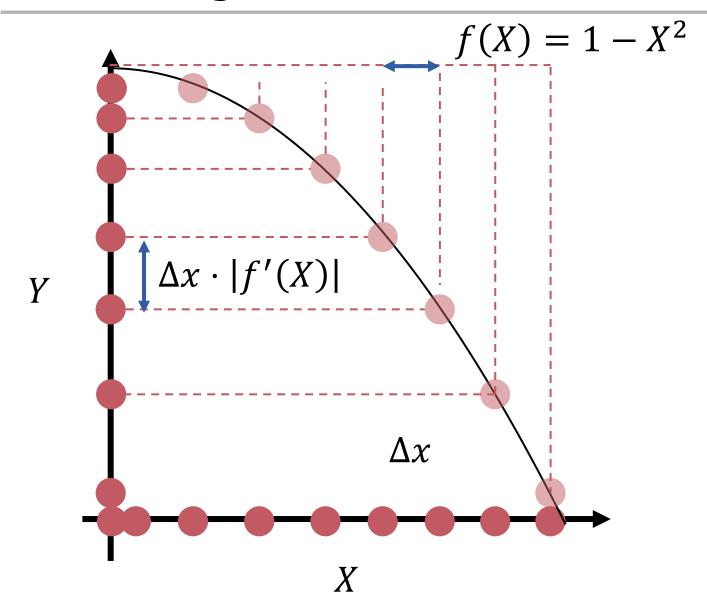


$$Y = f(X)$$

$$p_Y(y) = \frac{p_X(f^{-1}(y))}{f'(f^{-1}(y))}$$

### PDF change of variables





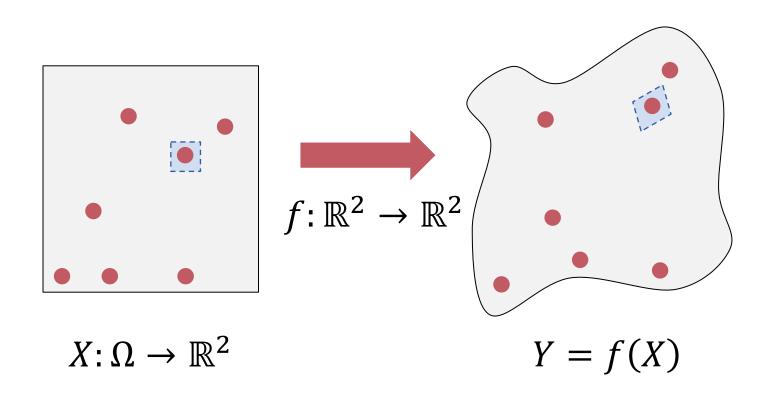
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PDFs should be nonnegative!

## PDF change of variables

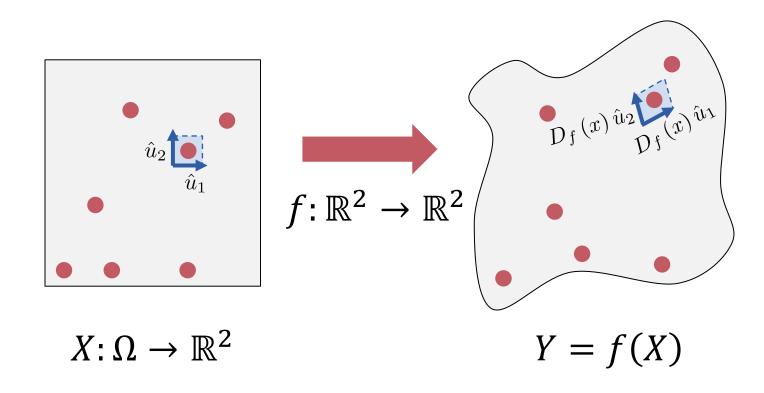




$$Y = f(X)$$

$$p_Y(y) = \frac{p_X(f^{-1}(y))}{|f|/|f|} *$$





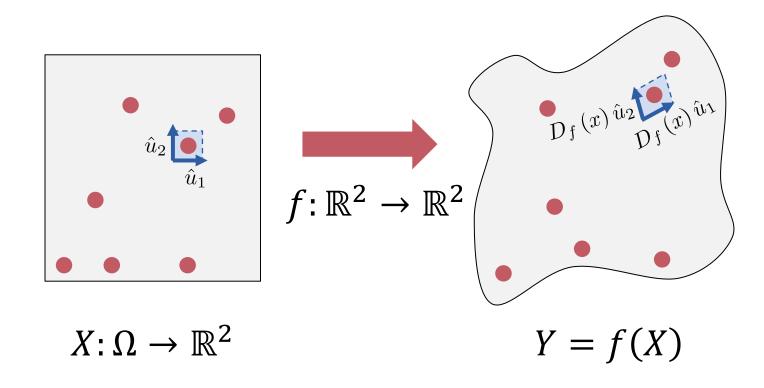
$$Y = f(X)$$

$$p_Y(y) = \frac{p_X(f^{-1}(y))}{|y|/|y|}$$

$$| D_f(x) \hat{u}_1 \times D_f(x) \hat{u}_2 |$$

$$= |\det D_f(x)| ||\hat{u}_1|| ||\hat{u}_2||$$





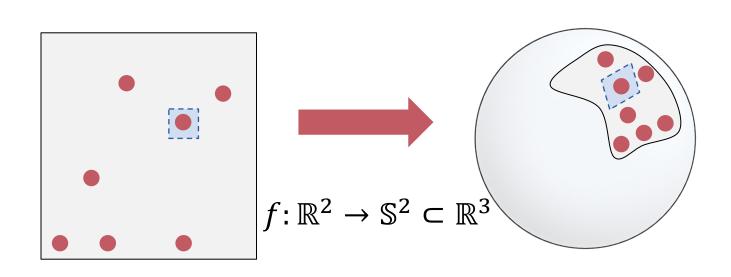
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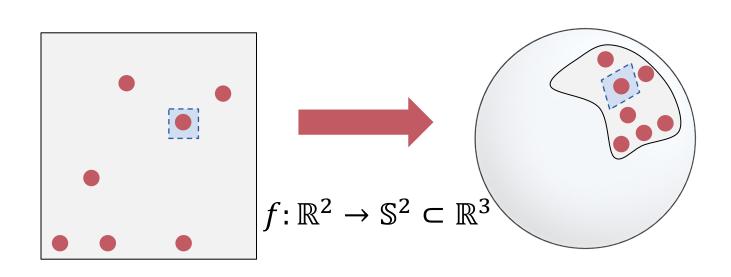
$$Y = f(X)$$

$$p_Y(y) = \frac{p_X(f^{-1}(y))}{|\det D_f(f^{-1}(y))|}^*$$

No determinant for  $3 \times 2$  matrices

 $X:\Omega\to\mathbb{R}^2$ 





Y = f(X)

### Proposition

$$Y = f(X)$$

$$p_Y(y) = \frac{p_X(f^{-1}(y))}{J_f(f^{-1}(y))}$$

$$\star J_f(x) = \sqrt{\det\left(D_f(x)^T D_f(x)\right)}$$



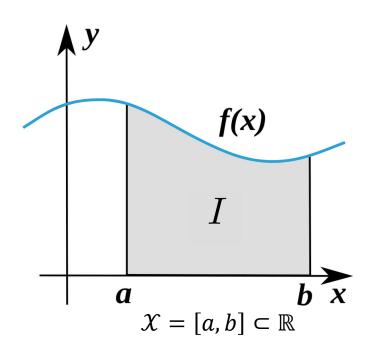
# **Numerical Integration**

## Integration



We have a function f on a domain  $\mathcal{X}$ .

Q. How can we compute  $I = \int_{\mathcal{X}} f(x) dx$ ?



source: User 4C / Wikipedia CC BY-SA 3.0

[Adam Celarek]

### 1. Analytic (symbolic) integration

- We can integrate only few functions analytically.
- Hardly predict difficulty of given problem

Try analytic integration for:

$$\int e^{-x^2} dx$$

$$\int xe^{-x^2} dx$$

$$\int x^2 e^{-x^2} dx$$

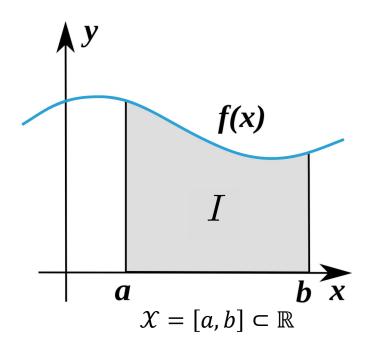
$$\int x^3 e^{-x^2} dx$$

### Integration



We have a function f on a domain X.

Q. How can we compute  $I = \int_{\mathcal{X}} f(x) dx$ ?



source: User 4C / Wikipedia CC BY-SA 3.0

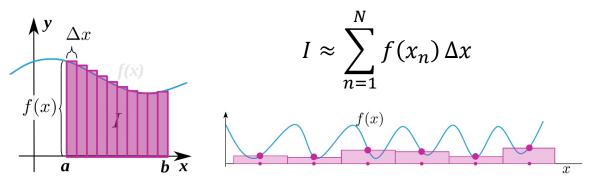
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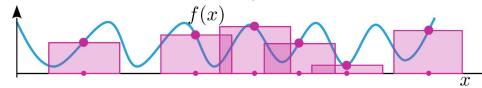
### 2. Numerical approximation by uniform grids

- For a fixed error level, we need  $N^{\dim(\mathcal{X})}$  samples!
- Aliasing!



**3. Monte Carlo estimation**: for a random variable  $X \sim \mathcal{U}(X)$ 

$$I \approx \hat{I} = \frac{|\mathcal{X}|}{N} \sum_{n=1}^{N} f(X_n)$$

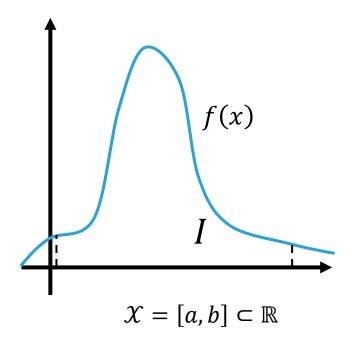


### **Monte Carlo Integration**



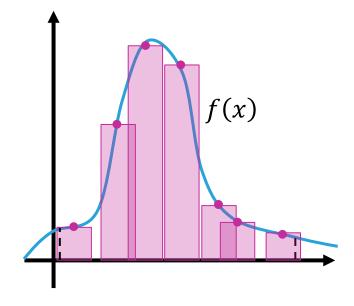
We have a function f on a domain X.

Q. How can we compute  $I = \int_{\mathcal{X}} f(x) dx$ ?



for a random variable  $X \sim \mathcal{U}(X)$ 

$$I \approx \hat{I} = \frac{|\mathcal{X}|}{N} \sum_{n=1}^{N} f(X_n)$$



$$Var(\hat{I}) \propto Var(\square)$$

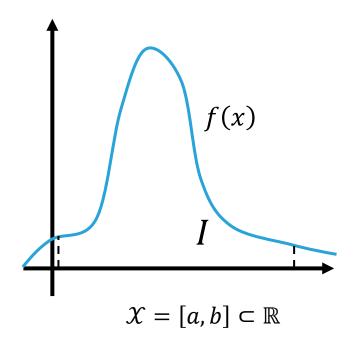
# Monte Carlo Integration: importance sampling



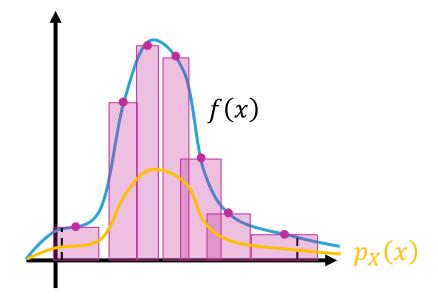
for a random variable X with PDF hopefully  $\propto f$ 

We have a function f on a domain  $\mathcal{X}$ .

Q. How can we compute  $I = \int_{\mathcal{X}} f(x) dx$ ?



$$I \approx \hat{I} = \frac{1}{N} \sum_{n=1}^{N} \frac{f(X_n)}{p_X(X_n)}$$



 $Var(\hat{I}) \propto Var($ 

### **Monte Carlo Integration**

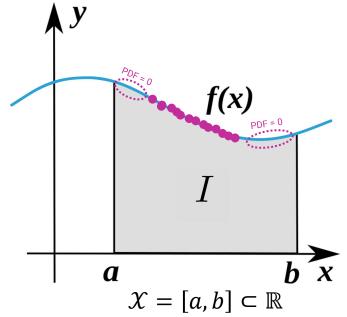


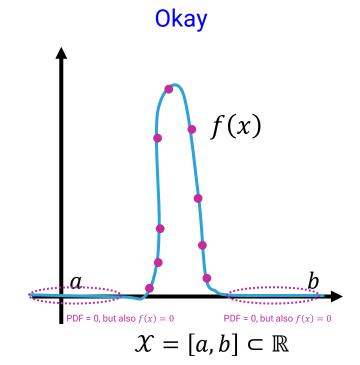
When is the MC integration correct (unbiased)?

$$\int_{\mathcal{X}} f(x) dx = I = \mathbb{E} \left| \hat{I} = \frac{1}{N} \sum_{n=1}^{N} \frac{f(X_n)}{p_X(X_n)} \right|$$

For any  $x \in \mathcal{X}$ , if  $f(x) \neq 0$  then  $p_X(x) > 0$ 

#### Failure case





## Language of Monte Carlo Integration



Sampling: generate a realization of a random variable
Which PDF do we choose?
How draw a sample of a r.v. with such PDF?

- "This sampling strategy is good for given integrand f."
- "In path tracing, BSDF sampling and emitter samplings are used."

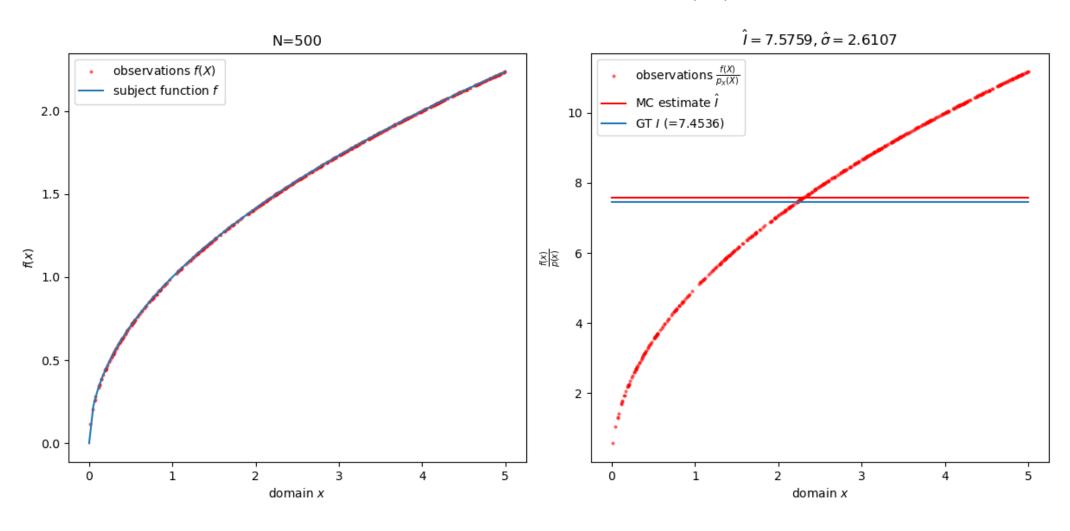
Correct → unbiased (or consistent, at least)

Efficient -> low variance

## Monte Carlo Integration: HW2 Problem 2.(a)



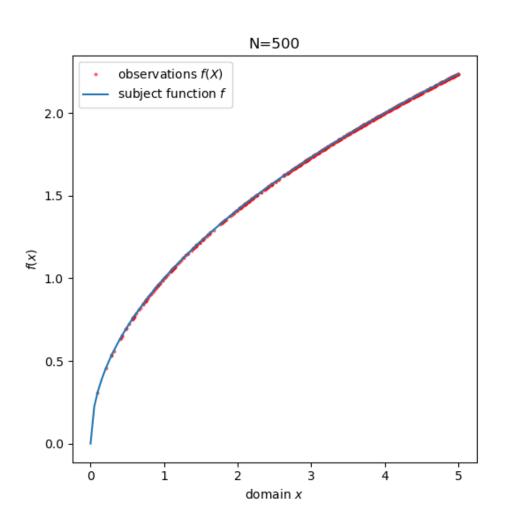
### random variable $X \sim \mathcal{U}(X)$

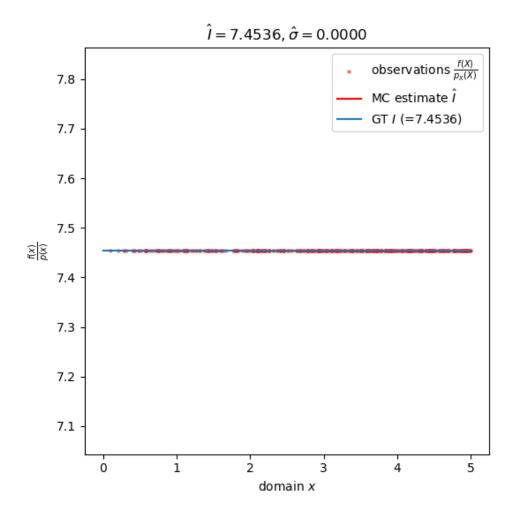


## Monte Carlo Integration: HW2 Problem 2.(a)



### random variable *X* with PDF $\propto f$





## **Monte Carlo Integration**



## How is importance sampling useful in practice?

- Performing importance sampling with a PDF exactly  $\propto f$  is as hard as the integration itself
  - However, using a PDF which roughly  $\propto f$  and analytically integrable is still useful

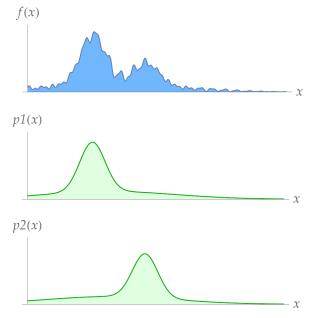
• We have good importance sampling for f and g. How about  $\int_{\mathcal{X}} f(x)g(x)dx$ ?

## Multiple Importance Sampling



# Multiple Importance Sampling (MIS)

- Many possible importance sampling strategies
- Which one should we use for a given integrand?
- MIS: *combine* strategies to preserve strengths of all of them
- *Balance heuristic* is (provably!) about as good as anything:



$$\frac{1}{N}\sum_{i=1}^{n}\sum_{j=1}^{n_i}\frac{f(x_{ij})}{\sum_{k}c_kp_k(x_{ij})}$$
 total # of samples fraction of samples taken w/ kth strategy

Still, several improvements possible (cutoff, power, max)—see Veach & Guibas.

# Multiple Importance Sampling: in context of rendering





- Many possible importance sampling strategies
- Which one should we use for a given integrand?
- MIS: combine strategies to preserve strengths of all of them

rendering eq. (for direct illumination): BRDF \* incident light (source emission)



generate random direction proportional to light source



## Multiple Importance Sampling: statement



### For integration problem

$$I = \int_{\mathcal{X}} f(x) \mathrm{d}x$$

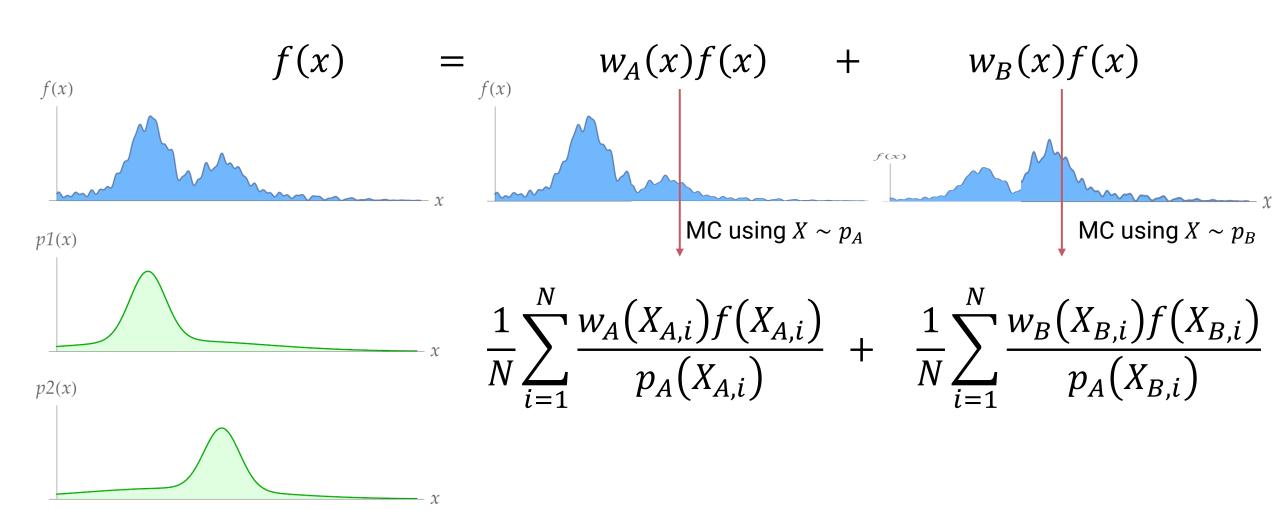
We have 2N samples  $X_{A,i} \sim p_A$  and  $X_{B,i} \sim p_B$  with  $i=1,\cdots,N$ 

$$\hat{I}_{\text{MIS}(A,B)} = \frac{1}{N} \sum_{S \in \{A,B\}} \sum_{i=1}^{N} w_{S}(X_{S,i}) \frac{f(X_{S,i})}{p_{S}(X_{S,i})}$$
, where

- $\sum_{S \in \{A,B\}} w_S(x) = 1$  whenever  $f(x) \neq 0$ ,
- $w_s(x) = 0$  whenever  $p_s(x) = 0$

### A way to understand MIS





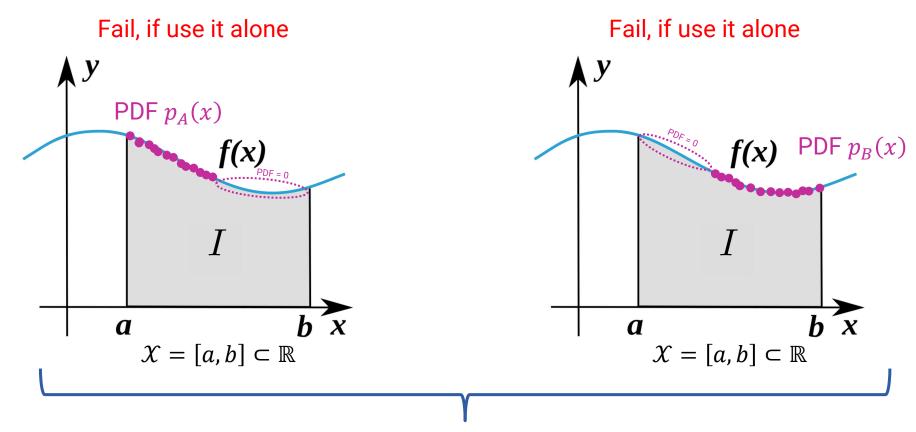
[Keenan Crane]

## Multiple Importance Sampling: unbiasedness



When is the MIS correct (unbiased)?

For any  $x \in \mathcal{X}$ , if  $f(x) \neq 0$  then either  $p_A(x) > 0$  or  $p_B(x) > 0$ 



But  $\hat{I}_{MIS(A,B)}$  is a correct estimator

### MIS strategies



Q. Then how can we compute such MIS weightes  $w_A(x)$  and  $w_B(x)$ ?

- 1. Very naively, just  $w_A = w_B = \frac{1}{2}$ 
  - only works if both of each strategy A and B are unbiased.
  - Do not use in practice
- 2. Balance heuristic:  $w_A(x) = \frac{p_A(x)}{p_A(x) + p_B(x)}, w_B(x) = \frac{p_B(x)}{p_A(x) + p_B(x)}$ 
  - A certain upper bound of difference between it and an optimum is known
    - [Veach 1997] Theotem 9.2
- 3. Power heuristic:  $w_A(x) = \frac{p_A(x)^{\alpha}}{p_A(x)^{\alpha} + p_B(x)^{\alpha}}, w_B(x) = \frac{p_B(x)^{\alpha}}{p_A(x)^{\alpha} + p_B(x)^{\alpha}}$ 
  - Usually  $\alpha = 2$

### MIS strategies



Q. Then how can we compute such MIS weightes  $w_A(x)$  and  $w_B(x)$ ?

- 1. Very naively, just  $w_A = w_B = \frac{1}{2}$ 
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- 2. Balance heuristic:  $w_A(x) = \frac{p_A(x)}{p_A(x) + p_B(x)}, w_B(x) = \frac{p_B(x)}{p_A(x) + p_B(x)}$ 
  - A certain upper bound of difference between it and an optimum is known
    - [Veach 1997] Theotem 9.2

**Theorem 9.2.** Let f,  $n_i$ , and  $p_i$  be given, for i = 1, ..., n. Let F be any unbiased estimator of the form (9.4), and let  $\hat{F}$  be the estimator that uses the weighting functions  $\hat{w}_i$  (the balance heuristic). Then

$$V[\hat{F}] - V[F] \le \left(\frac{1}{\min_i n_i} - \frac{1}{\sum_i n_i}\right) \mu^2, \tag{9.9}$$

where  $\mu=E[F]=E[\hat{F}]$  is the quantity to be estimated. (A proof is given in Appendix 9.A.)