

1 Table of Distributions

Notation	PMF/PDF and Support	Expected Value	Variance	MGF
Bernoulli Be(p)	$P(X = 1) = p$ $P(X = 0) = q$	p	pq	$q + pe^t$
Binomial Bin(n, p)	$P(X = k) = \binom{n}{k} p^k q^{n-k}$ $k = 0, 1, 2, \dots, n$	np	npq	$(q + pe^t)^n$
Geometric Geom(p)	$P(X = k) = pq^{k-1}$ $k = 1, 2, \dots$	$\frac{1}{p}$	$\frac{q}{p^2}$	$\frac{pe^t}{1-qe^t},$ $qe^t < 1$
Negative Binomial NB(r, p)	$P(X = k) = \binom{k-1}{r-1} p^r q^{k-r}$ $k = r, r+1, \dots$	$\frac{r}{p}$	$\frac{rq}{p^2}$	$\left(\frac{pe^t}{1-qe^t}\right)^r,$ $qe^t < 1$
Hypergeometric HGeom(n, N, m)	$P(X = k) = \frac{\binom{m}{k} \binom{N-m}{n-k}}{\binom{N}{n}}$ $k = 0, 1, \dots, n$	$\frac{nm}{N}$	$\frac{nm}{N} \left[\frac{(n-1)(m-1)}{N-1} + 1 - \frac{nm}{N} \right]$	messy
Poisson Poisson(λ)	$P(X = k) = \frac{e^{-\lambda} \lambda^k}{k!}$ $k = 0, 1, 2, \dots$	λ	λ	$e^{\lambda(e^t-1)}$
Uniform U(a, b)	$f(x) = \frac{1}{b-a}$ $x \in (a, b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{bt}-e^{at}}{(b-a)t}$
Normal $\mathcal{N}(\mu, \sigma^2)$	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$ $x \in (-\infty, \infty)$	μ	σ^2	$\exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right)$
Exponential Exp(λ)	$f(x) = \lambda e^{-\lambda x}$ $x \in [0, \infty)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda-t},$ $t < \lambda$
Gamma Gamma(a, λ)	$f(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{a-1}}{\Gamma(a)}$ $x \in [0, \infty)$	$\frac{a}{\lambda}$	$\frac{a}{\lambda^2}$	$\left(\frac{\lambda}{\lambda-t}\right)^a,$ $t < \lambda$
Beta Beta(a, b)	$f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$ $x \in (0, 1)$	$\frac{a}{a+b}$	$\frac{ab}{(a+b)^2(a+b+1)}$	messy
Chi-Square χ_n^2	$f(x) = \frac{1}{2^{n/2} \Gamma(n/2)} x^{n/2-1} e^{-x/2}$ $x \in (0, \infty)$	n	$2n$	$(1-2t)^{-n/2}$ $t < 1/2$

2 Useful Results

2.1 Probability

Sample space is the set of all possible outcomes of an experiment, usually denoted by S . For tossing two dice,

$$S = \{(1, 1), (1, 2), \dots, (1, 6), (2, 1), \dots, (6, 6)\}$$

$$= \{(i, j) : 1 \leq i, j \leq 6\}$$

Event Any subset A of the sample space is an event

Odds The odds of an event A is defined by $\frac{P(A)}{1-P(A)}$

Increasing/Decreasing Events

If $\{E_n\}$ is an increasing sequence of events, then

$$\lim_{n \rightarrow \infty} E_n = \bigcup_{i=1}^{\infty} E_i$$

If $\{E_n\}$ is a decreasing sequence of events, then

$$\lim_{n \rightarrow \infty} E_n = \bigcap_{i=1}^{\infty} E_i$$

If $\{E_n\}$ is either an increasing or decreasing sequence of events, then

$$P\left(\lim_{n \rightarrow \infty} E_n\right) = \lim_{n \rightarrow \infty} P(E_n)$$

Conditional Probability

$$P(A|B) = P(A|BC)P(C|B) + P(A|BC^C)P(C^C|B)$$

Inclusion/Exclusion Principle Let E_1, E_2, \dots, E_n be any events, then

$$\begin{aligned} P(E_1 \cup E_2 \cup \dots \cup E_n) &= \sum_{i=1}^n P(E_i) - \sum_{1 \leq i_1 < i_2 \leq n} P(E_{i_1} \cap E_{i_2}) + \dots \\ &\quad (-1)^{r+1} \sum_{1 \leq i_1 < \dots < i_r \leq n} P(E_{i_1} \cap \dots \cap E_{i_r}) \\ &\quad + \dots + (-1)^{n+1} P(E_1 \cap \dots \cap E_n) \end{aligned}$$

Derangement/Matching Problem

$$\begin{aligned} P(\text{at least one match}) &= 1 - \frac{1}{2!} + \frac{1}{3!} - \dots (-1)^{n+1} \frac{1}{n!} \\ P(k \text{ matches}) &= \binom{n}{k} \frac{1}{n(n-1) \dots (n-k+1)} \cdot \left(1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \dots (-1)^{n-k} \frac{1}{(n-k)!} \right) \end{aligned}$$

2.2 Discrete RV

Suppose $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Bin}(X, p)$, then $X - Y$ and Y are independent and

$$X - Y \sim \text{Poisson}(\lambda(1-p)), Y \sim \text{Poisson}(\lambda p)$$

2.3 Continuous RV

Exponential Distribution

$$P(X > s + t | X > s) = P(X > t)$$

Gamma Distribution

$$\Gamma(\alpha) = \int_0^\infty e^{-t} t^{\alpha-1} dt$$

- $\Gamma(1) = 1$ and $\Gamma(\alpha) = (\alpha-1)\Gamma(\alpha-1)$
- $\Gamma(n) = (n-1)!$ for $n = 1, 2, 3, \dots$
- If $X_i \sim \text{Exp}(\lambda)$ independently, then $\sum_{i=1}^n X_i \sim \text{Gamma}(n, \lambda)$

Beta Distribution

$$\begin{aligned} B(a, b) &= \int_0^1 t^{a-1} (1-t)^{b-1} dt \\ &= \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} \\ \text{Beta}(1, 1) &\equiv \text{U}(0, 1) \end{aligned}$$

If $X \sim \text{Gamma}(\alpha, \lambda)$ and $Y \sim \text{Gamma}(\beta, \lambda)$, then $U = X+Y \sim \text{Gamma}(\alpha+\beta, \lambda)$ and $V = X/(X+Y) \sim \text{Beta}(\alpha, \beta)$ and are independent

Bivariate Normal Distribution

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \times \exp\left(-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x}\right)^2 + \left(\frac{y-\mu_y}{\sigma_y}\right)^2 - 2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} \right]\right)$$

Joint Transformation

$$f_{Y_1, Y_2, \dots, Y_n}(y_1, y_2, \dots, y_n) = f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) \frac{1}{|J(x_1, x_2, \dots, x_n)|}$$

Chi-Square Distribution

$$\chi_n^2 \equiv \text{Gamma}\left(\frac{n}{2}, \frac{1}{2}\right)$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

Monotonic Transformation

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| & \text{if } y \in \mathcal{R}(g) \\ 0 & \text{otherwise} \end{cases}$$

2.4 Joint Distribution

$$\begin{aligned} P(a_1 < X \leq a_2, b_1 < Y \leq b_2) &= F_{X,Y}(a_2, b_2) \\ &\quad - F_{X,Y}(a_1, b_2) + F_{X,Y}(a_1, b_1) - F_{X,Y}(a_2, b_1) \end{aligned}$$

$$f_{X,Y}(x, y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x, y)$$

$$f_{X,Y}(x, y) = h(x)g(y) \Leftrightarrow X \text{ and } Y \text{ are independent}$$

Convolution of Independent Distributions

$$F_{X+Y}(a) = \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy$$

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where

$$J(x_1, x_2, \dots, x_n) = \begin{vmatrix} \frac{\partial g_1}{\partial x_1} & \frac{\partial g_1}{\partial x_2} & \dots & \frac{\partial g_1}{\partial x_n} \\ \frac{\partial g_2}{\partial x_1} & \frac{\partial g_2}{\partial x_2} & \dots & \frac{\partial g_2}{\partial x_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial g_n}{\partial x_1} & \frac{\partial g_n}{\partial x_2} & \dots & \frac{\partial g_n}{\partial x_n} \end{vmatrix},$$

$Y_i = g_i(X_1, X_2, \dots, X_n)$, and $x_i = h_i(y_1, y_2, \dots, y_n)$ for $i = 1, 2, \dots, n$

2.5 Expectations and Variance

Expectations

- If X and Y are independent, then $E[g(X)h(Y)] = E[g(x)]E[h(y)]$
- $E(X) = E[E[X|Y]]$

Covariance

$$\begin{aligned} Cov(X, Y) &= E[(X - \mu_X)(Y - \mu_Y)] \\ &= E(XY) - E(X)E(Y) \end{aligned}$$

$$Cov\left(\sum_{i=1}^n a_i X_i, \sum_{j=1}^m b_j Y_j\right) = \sum_{i=1}^n \sum_{j=1}^m a_i b_j Cov(X_i, Y_j)$$

$$Var\left(\sum_{k=1}^n X_k\right) = \sum_{k=1}^n Var(X_k) + 2 \sum_{1 \leq i < j \leq n} Cov(X_i, X_j)$$

$$Var(X|Y) = E[X^2|Y] - [E(X|Y)]^2$$

$$Var(X) = E[Var(X|Y)] + Var[E(X|Y)]$$

Correlation

$$\rho(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}}$$

$$\rho(X, Y) = \pm 1 \text{ iff } Y = \pm aX + b \text{ where } a = \frac{\sigma_Y}{\sigma_X}$$

Moment Generating Functions

$$M_X(t) = E(e^{tX})$$

$$E(X^n) = \left. \frac{d^n}{dt^n} M_X(t) \right|_{t=0}$$

Taylor Series Expansion

$$g(x) = g(\mu) + \frac{g'(\mu)}{1!}(x - \mu) + \frac{g''(\mu)}{2!}(x - \mu)^2 + \dots$$

Multiplicative Property If X and Y are independent, then

$$M_{X+Y}(t) = M_X(t)M_Y(t)$$

Uniqueness Property If $\exists t$ such that $M_X(t) = M_Y(t) \forall t \in (-h, h)$ then X and Y have the same distributions

Joint Moment Generating Functions

$$M_{X_1, X_2, \dots, X_n}(t_1, t_2, \dots, t_n) = E[e^{t_1 X_1 + t_2 X_2 + \dots + t_n X_n}]$$

$$M_{X_i}(t) = E[e^{tX_i}] = M_{X_1, X_2, \dots, X_n}(0, \dots, 0, t, 0, \dots, 0)$$

2.6 Inequalities

$$a \leq X \leq b \Rightarrow a \leq E(X) \leq b$$

Monotone Property $X \leq Y \Rightarrow E(X) \leq E(Y)$

Boole's inequality

$$P\left(\bigcup_{k=1}^n A_k\right) \leq \sum_{k=1}^n P(A_k)$$

Markov's inequality Let X be a nonnegative random variable. For $a > 0$, we have

$$P(X \geq a) \leq \frac{E(X)}{a}$$

Chebyshev's inequality For $a > 0$, we have

$$P(|X - \mu| \geq a) \leq \frac{Var(X)}{a^2}$$

One-sided Chebyshev's inequality If X is a r.v. with mean 0 and finite variance σ^2 , then, for any $a > 0$,

$$P(X \geq a) \leq \frac{\sigma^2}{\sigma^2 + a^2}$$

Jensen's inequality If $g(x)$ is a convex function, then

$$E[g(x)] \geq g[E(X)]$$