

1 Probability Review

Multinomial Distribution

$$\Pr(X_1 = x_1, \dots, X_r = x_r) = \binom{n}{x_1, \dots, x_r} \prod_{i=1}^r p_i^{x_i}$$

Mean Square Error (MSE)

$$E\{(Y - c)^2\} = \text{var}(Y) + \{E(Y) - c\}^2$$

$$E\{(Y - c)^2|x\} = \text{var}[Y|x] + \{E[Y|x] - c\}^2$$

which are special cases of $E(Y^2) = \text{var}(Y) + [E(Y)]^2$. MSE is minimized if and only if $c = E(Y)$ or $E[Y|x]$.

Usually the formula for $E[Y|x] = f(x)$ is determined from observations/data and x can be a vector of realisations from covariates.

$$\text{MSE}_{\text{empirical}} = \frac{1}{n} \sum_{i=1}^n \{E[Y|x_i] - y_i\}^2$$

In the real world, we have different realisations x_i of the random variable X , hence the mean MSE is

$$\frac{1}{n} \sum_{i=1}^n \text{var}[Y|x_i] \approx E(\text{var}[Y|X]) \leq \text{var}(Y)$$

Analysis of Variance (ANOVA)

involves breaking of variance into components

$$\text{var}(Y) = E(\text{var}[Y|X]) + \text{var}(E[Y|X])$$

1.1 Distributions

χ_1^2 distribution

Let $Z \sim \mathcal{N}(0, 1)$. $V = Z^2$ has a χ^2 distribution with 1 degree of freedom

$$f(v) = \frac{1}{\sqrt{2\pi}} v^{-1/2} e^{-v/2}$$

Gamma distribution

$$f(t) = \frac{\lambda^\alpha}{\Gamma(\alpha)} t^{\alpha-1} e^{-\lambda t}, t \geq 0$$

$$\Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx$$

χ_n^2 distribution

Let V_1, \dots, V_n be IID χ_1^2

$$V = \sum_{i=1}^n V_i$$

has a χ_n^2 distribution with n degrees of freedom

t distribution

Let $Z \sim \mathcal{N}(0, 1)$ and $V \sim \chi_n^2$ be independent

$$t_n = \frac{Z}{\sqrt{V/n}}$$

has a t distribution with n degrees of freedom

F distribution

Let $V \sim \chi_m^2$ and $W \sim \chi_n^2$ be independent

$$F_{m,n} = \frac{V/m}{W/n}$$

has an F distribution with (m, n) degrees of freedom

*Note: $t_n^2 = F_{1,n}$

1.2 Sample Variance

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

\bar{X} and S^2 are independent

$$\bar{X} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

$$\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1}$$

2 Survey and Random Sampling

Let X_1, \dots, X_N be random draws without replacement from a population of size N with mean μ and variance σ^2 .

$$\text{cov}(X_i, X_j) = -\frac{\sigma^2}{N-1} \forall i \neq j$$

$$\text{var}(\bar{X}) = \left(\frac{N-n}{N-1}\right) \frac{\sigma^2}{n}$$

2.1 Exchangeable

RV's Y_1, \dots, Y_k are exchangeable if all reordered vectors have the same distribution as (Y_1, \dots, Y_k) . i.e. for any permutation π on $\{1, \dots, K\}$,

$$(Y_{\pi(1)}, \dots, Y_{\pi(k)}) \stackrel{d}{=} (Y_1, \dots, Y_k)$$

2.2 Estimate and Estimator

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

- μ, σ, σ^2 are **parameters**
- \bar{x} is an **estimate** of μ
- \bar{x} is a realisation of the **estimator** \bar{X}
- **Standard Error (SE)** is defined as the SD of the estimator

$$\text{SE} = \text{SD}(\bar{X}) = \frac{\sigma}{\sqrt{n}}$$

which is how much \bar{X} fluctuates around μ

- Estimate of σ

- Biased estimate of σ^2

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$E(\hat{\sigma}^2) = \frac{n-1}{n} \sigma^2$$

- Unbiased estimate of σ^2 (preferred)

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$E(s^2) = \sigma^2$$

How to estimate μ ?

- μ is estimated by \bar{x}
- Error in \bar{x} is measured by the SE:

$$SD(\bar{x}) = \frac{\sigma}{\sqrt{n}}$$

which is **estimated** by $\frac{s}{\sqrt{n}}$ since σ is unknown

- **Conclusion:** μ is estimated as \bar{x} , give or take $\frac{s}{\sqrt{n}}$

$$\text{SE estimated by } \frac{s}{\sqrt{n}} = \frac{\sqrt{\frac{n}{n-1}} \times SD}{\sqrt{n}}$$

where $SD = \hat{\sigma}$

How to estimate p ?

- \hat{p} is the estimator of p

$$E(\hat{p}) = p$$

$$\text{var}(\hat{p}) = \frac{\sigma^2}{n} = \frac{p(1-p)}{n}$$

$$SE = SD(\hat{p}) = \sqrt{\frac{p(1-p)}{n}}$$

which is **estimated** by realisations of \hat{p}

2.3 Interval estimation

2.3.1 Definitions

- For sufficiently large n ,

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim \mathcal{N}(0, 1)$$

- The p -quantile of $Z \sim \mathcal{N}(0, 1)$ is the number q such that

$$\Phi(q) = \Pr(Z \leq q) = p$$

$$q = \Phi^{-1}(p)$$

```
1 q <- qnorm(p)
```

```
2 p <- pnorm(q)
```

- For $0 < p < 0.5$, let z_p be such that

$$\Pr(Z > z_p) = p$$

$$z_p = \Phi^{-1}(1-p)$$

In other words, $z_p = (1-p)$ -quantile of Z

2.3.2 CI Estimation

- For large n ,

$$\bar{X} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\Pr\left(-z_{\frac{\alpha}{2}} \leq \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \leq z_{\frac{\alpha}{2}}\right) \approx 1 - \alpha$$

$$\Pr\left(\bar{X} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right) \approx 1 - \alpha$$

where the above, $\left(\bar{X} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$ is a random interval. Realisation \bar{x} of \bar{X} gives the realised interval

- $(1 - \alpha)$ -CI for μ is of the form

$$(\text{estimate} - z_{\frac{\alpha}{2}} \text{SE}, \text{estimate} + z_{\frac{\alpha}{2}} \text{SE})$$

2.3.3 Exact CI

- Let $t_{\frac{\alpha}{2}, n-1}$ be the number such that

$$\Pr(t_{n-1} > t_{\frac{\alpha}{2}, n-1}) = \alpha/2$$

- **[Important]** Exact CI only works if $X \sim \mathcal{N}(\mu, \sigma^2)$ and x_i 's are realisations from IID Normal Distribution
- * CI is exact means that $\Pr(\mu \text{ is within the interval})$ is exactly $1 - \alpha$
- $(1 - \alpha)$ -CI for μ is

$$\left(\bar{x} - t_{\frac{\alpha}{2}, n-1} \frac{s}{\sqrt{n}}, \bar{x} + t_{\frac{\alpha}{2}, n-1} \frac{s}{\sqrt{n}}\right)$$

2.4 Bias in Survey

Famous example: US presidential election survey conducted by *Literary Digest* in 1936

2.4.1 Bias in Measurement

- x_1, \dots, x_n are realisations of random draws X_1, \dots, X_n from a population with mean $\mu + b$ and variance σ^2
- $SE = \sigma/\sqrt{n}$ measures how far \bar{x} is from $E(\bar{X}) = \mu + b$
- How to remove bias?
- MSE

$$E(\bar{X} - \mu)^2 = \text{var}(\bar{X}) + \{E(\bar{X}) - \mu\}^2$$

$$MSE = SE^2 + \text{bias}^2$$

However μ is unknowable, hence it is not possible to remove bias unless we make very careful observations

3 Useful Results

$$\sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n X_i^2 - n\bar{X}^2$$