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①  $KL(P_{\text{emp}} \| q) = \int P_{\text{emp}}(x) (\log P_{\text{emp}}(x) - \log q(x; \hat{\theta})) dx$

$$= \int P_{\text{emp}}(x) \log P_{\text{emp}}(x) dx - \int P_{\text{emp}}(x) \log q(x; \hat{\theta}) dx \quad \text{--- ①}$$

$$\int P_{\text{emp}}(x) \log q(x; \hat{\theta}) dx = \int \frac{1}{n} \sum_{i=1}^n \delta(x, x_i) \log q(x; \hat{\theta}) dx = \frac{1}{n} \sum_{i=1}^n \log q(x_i; \hat{\theta})$$

$$= \frac{1}{n} \sum_{i=1}^n \log(q(x_i; \hat{\theta})) \triangleq \hat{\theta}$$

so when  $\hat{\theta}$  is the maximum likelihood estimator, we get the minimum  $KL(P_{\text{emp}} \| q)$ , which means  $\arg \min_q KL(P_{\text{emp}} \| q)$  is obtained by  $q(x)$ .

②  $J(w)$  to solve

$$J(w) = -\frac{1}{|D|} \sum_{i \in D} \log \sigma(\pm x_i^T w) + \lambda \|w\|_2^2, \text{ label } \log \sigma(x_i^T w) \text{ as } \textcircled{I}$$

$$\frac{\partial (\log \sigma(x_i^T w))}{\partial w_j} = \frac{\partial (\log(1 + e^{-x_i^T w}))}{\partial w_j} = \frac{-e^{-x_i^T w}}{1 + e^{-x_i^T w}} \cdot (-x_{ij}) = (1 - \sigma(x_i^T w)) x_{ij}$$

$$\frac{\partial^2 (\log \sigma(x_i^T w))}{\partial w_j \partial w_k} = \frac{\partial}{\partial w_k} ((1 - \sigma(x_i^T w)) x_{ij}) = -\sigma(x_i^T w)(1 - \sigma(x_i^T w)) x_{ij} x_{ik}$$

$\because \lambda \|w\|_2^2$  is convex,  $\frac{\partial (\log \sigma(x_i^T w))}{\partial w} = \frac{\sigma(x_i^T w)(1 - \sigma(x_i^T w)) x_i x_i^T}{\sigma(x_i^T w)}$

$$= (1 - \sigma(x_i^T w)) x_i x_i^T$$

$$\therefore \frac{\partial ((1 - \sigma(x_i^T w)) x_i x_i^T)}{\partial w} = -\sigma(x_i^T w)(1 - \sigma(x_i^T w)) (x_i x_i^T)^2$$

because  $\sigma(x) \in (0, 1)$ ,  $(\mathbf{I})'' \neq 1 - \sigma(x_i^T \mathbf{w}) \in (0, 1)$   
 $(x_i x_i^T)^2 > 0$ , so  $(\mathbf{I})''$  always  $< 0$

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when it add the coefficient  $-\frac{1}{|\mathcal{D}|}$ , it will become always  $> 0$

so the first subject is also convex [the hessian matrix  $\geq 0$ ]

convex + convex  $\Rightarrow$  convex  $\Rightarrow$  local optimal = global optimal  
 $\Rightarrow$  False

False

L1 norm prefer a sparse  $\hat{\mathbf{w}}$ , but L2 norm prefer a average value  $\hat{\mathbf{w}}$

if we consider  $\bar{J}(\mathbf{w})$  as a loss function, when we optimize and get the  $\text{argmin}_{\mathbf{w}} \bar{J}(\mathbf{w}) = \hat{\mathbf{w}}$ , it's not sparse.

$$\begin{aligned} \textcircled{3} \quad l(\theta) &= \sum_{i=1}^m \log P(x; \theta) = \sum_{i=1}^m \log \sum_z P(x, z; \theta) \\ &= \sum_i \log \sum_{z^{(i)}} Q_i(z^{(i)}) \frac{P(x^{(i)}, z^{(i)}; \theta)}{Q_i(z^{(i)})} \\ &\geq \sum_i \sum_{z^{(i)}} Q_i(z^{(i)}) \log \frac{P(x^{(i)}, z^{(i)}; \theta)}{Q_i(z^{(i)})} \\ &= l(\theta^{(t)}) \end{aligned}$$

$$w_j^{(i)} = Q_i(z^{(i)} = j) = P(z^{(i)} = j | x^{(i)}; \phi, \mu, \Sigma)$$

$$\begin{aligned} &\sum_{i=1}^m \sum_{z^{(i)}} Q_i(z^{(i)}) \log \frac{P(x^{(i)}, z^{(i)}; \phi, \mu, \Sigma)}{Q_i(z^{(i)})} \quad \text{----- I} \\ &= \sum_{i=1}^m \sum_{j=1}^m w_j^{(i)} \log \frac{\frac{1}{(2\pi)^{\frac{K}{2}} |\Sigma_j|^{\frac{1}{2}}} \exp(-\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)) \cdot \phi_j}{w_j^{(i)}} \end{aligned}$$

$$\nabla_{\mu_1} \text{I} = - \nabla_{\mu_1} \sum_{i=1}^m \sum_{j=1}^K w_j^{(i)} (x^{(i)} - \mu_j)^T \cdot \frac{1}{2} \Sigma_j^{-1} (x^{(i)} - \mu_j)$$

$$= \frac{1}{2} \sum_{i=1}^m w_1^{(i)} \nabla_{\mu_1} \Sigma_1^T \Sigma_1^{-1} x^{(i)} - \mu_1^T \Sigma_1^{-1} \mu_1 = \sum_{i=1}^m w_1^{(i)} (\Sigma_1^{-1} x^{(i)} - \Sigma_1^{-1} \mu_1)$$

令其為0得  $\mu_i = \frac{\sum_1^m w_l^{(i)} x^{(i)}}{\sum_1^m w_l^{(i)}} \Rightarrow \sum_{i=1}^m \sum_{j=1}^k w_j^{(i)} \log \phi_j$

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又:  $\sum_{j=1}^k \phi_j = 1$

$\therefore \text{let } L(\phi) = \sum_{i=1}^m \sum_{j=1}^k w_j^{(i)} \log \phi_j + \beta (\sum_{j=1}^k \phi_j - 1)$

$\frac{\partial L(\phi)}{\partial \phi_j} = \sum_{i=1}^m \frac{w_j^{(i)}}{\phi_j} + \beta \stackrel{\Delta}{=} 0$

$\Rightarrow \phi_j = \frac{\sum_{i=1}^m w_j^{(i)}}{-\beta} \Rightarrow -\beta = \sum_{i=1}^m \sum_{j=1}^k w_j^{(i)} = \sum_{i=1}^m 1 = m$

$\therefore \phi_j = \frac{1}{m} \sum_{i=1}^m w_j^{(i)}$

$\nabla_{\mu_k} \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk} \log \frac{\frac{1}{(2\pi)^{\frac{N}{2}} |\bar{\Sigma}_k|^{\frac{1}{2}}} \exp(-\frac{1}{2}(x_n - \mu_k)^T \bar{\Sigma}_k^{-1} (x_n - \mu_k)) \bar{\pi}_k}{\gamma_{nk}}$

$= (\sum_{n=1}^N \gamma_{nk}) (x_n - \mu_k)^T \bar{\Sigma}_k^{-1}$

$= \sum_{n=1}^N \gamma_{nk} \bar{\Sigma}_k^{-1} (x_n - \mu_k)$

□□  $\sum_{k=1}^K \bar{\pi}_k = 1$ ,  $\bar{\pi}$  is the above  $\phi$ , we have already proved that

$\bar{\pi}_k = \frac{1}{N} \sum_{n=1}^N \gamma_{nk}$

□□□  $\bar{\Sigma}_k$  can be the  $\mu$  above, so we get  $\bar{\Sigma}_k = \frac{\sum_{i=1}^m \gamma_{nk} x_i}{\sum_{n=1}^m \gamma_n}$

which is  $\frac{\partial \ln p(x|\mu, \bar{\Sigma})}{\partial \mu} = -\frac{N}{\bar{\Sigma}} \bar{\Sigma}^{-1} + \frac{1}{\bar{\Sigma}} \bar{\Sigma}^{-1} \left[ \sum_{n=1}^N (x_n - \mu)(x_n - \mu)^T \right] \bar{\Sigma}^{-1}$   
 $\stackrel{\Delta}{=} 0$

$\Rightarrow \bar{\Sigma}_k = \frac{1}{N} \sum_{n=1}^N (x_n - \mu)(x_n - \mu)^T$