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# Estimation of Non-Stationary, Linear Time Series

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Background

## Definition: ARMA(p,q) process

An Autoregressive-Moving Average (ARMA) process of autoregressive order  $p$  and moving average order  $q$  is defined as:

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + u_t + \theta_1 u_{t-1} + \theta_2 u_{t-2} + \cdots + \theta_q u_{t-q}$$
$$u_t \sim N(0, \sigma^2) \text{ i.i.d.}, \quad t \in \mathbb{Z}, t \leq T$$

This can be equivalently expressed in terms of  $\phi(L), \theta(L)$  the characteristic polynomial of the AR and MA component respectively.

$$\phi(L)y_t = \theta(L)u_t$$
$$\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$$
$$\theta(L) = 1 + \theta_1 L + \theta_2 L^2 + \cdots + \theta_q L^q$$

## Definition: Stationary process

The time series is said to be stationary if its first and second moments are time invariant and the autocovariance function depends only on the difference in time  $h$  and not  $t$ . Formally, a stochastic process  $y_t$  is stationary if

$$E[y_t] = \mu \text{ and } E[(y_t - \mu)(y_{t-h} - \mu)] = \gamma_h = \gamma_{-h} \quad \forall t, h \in \mathbb{Z}$$

A non-stationary process violates one or both of these conditions.

An stable ARMA process has all roots of the AR characteristic polynomial  $\phi(L)$  **outside** the complex unit circle.

An invertible ARMA process has at all roots of the MA characteristic polynomial  $\theta(L)$  **outside** the complex unit circle.

# Spectral Factorisation of Moving Average processes

For each non-invertible MA(q) process  $\theta(L)u_t$ , there exists a unique invertible MA(q) process  $\theta^s(L)u_t^s$  such that the autocovariance functions of each process are equal and the processes are equivalent under this condition.

In the MA(1) case, it can be shown that for the non-invertible process  $\theta(L)u_t$  the equivalent invertible process is given by

$$y_t = \frac{1}{\theta} u_{t-1}^s + u_t^s \quad u_t^s \sim N(0, \theta^2 \sigma^2) \text{ i.i.d}$$

Equivalence of ARMA(p,q) and AR( $\infty$ )

## II Equivalence of ARMA(p,q) and AR( $\infty$ )

### AR( $\infty$ ) of an ARMA(1,1)

Consider an ARMA(1,1) process and isolate  $u_t$

$$\begin{aligned} y_t &= \phi y_{t-1} + \theta u_{t-1} + u_t & (1) \\ \Leftrightarrow u_t &= y_t - \phi y_{t-1} - \theta u_{t-1} \end{aligned}$$

Applying the backshift operator to both sides we get

$$u_{t-1} = y_{t-1} - \phi y_{t-2} - \theta u_{t-2}$$

So, by substituting for  $u_{t-1}$  into equation (1) and similarly for  $u_{t-2}, u_{t-3} \dots$  we derive

$$\begin{aligned} y_t &= \phi y_{t-1} + \theta \left( y_{t-1} - \phi y_{t-2} - \theta (y_{t-2} - \phi y_{t-3} - \theta (\dots)) \right) + u_t \\ &= u_t + \sum_{i=1}^{\infty} (\phi + \theta) \theta^{i-1} y_{t-i} \end{aligned}$$

Which is an AR( $\infty$ ).



## II Equivalence of ARMA(p,q) and AR( $\infty$ )

### Conditions of AR( $\infty$ )

1. Are limited to finite data sets, and so we must truncate the infinite series at some  $p^* \in N$  to form an AR( $p^*$ ) process which can be fit to the time series.
2. We are required to condition the likelihood function on the initial  $p^*$  observations.

### AR( $p^*$ ) approximation of an ARMA(1,1)

Given a set of observations  $\{y_t\}_{t=1}^T$  and an ARMA(1, 1) process with gaussian error process  $u_t \sim N(0, \sigma^2)$  i.i.d, the resulting parameter estimates  $\hat{\phi}, \hat{\theta}, \hat{\sigma}^2$  are those which maximise the conditional maximum likelihood of the AR( $p^*$ ) process.

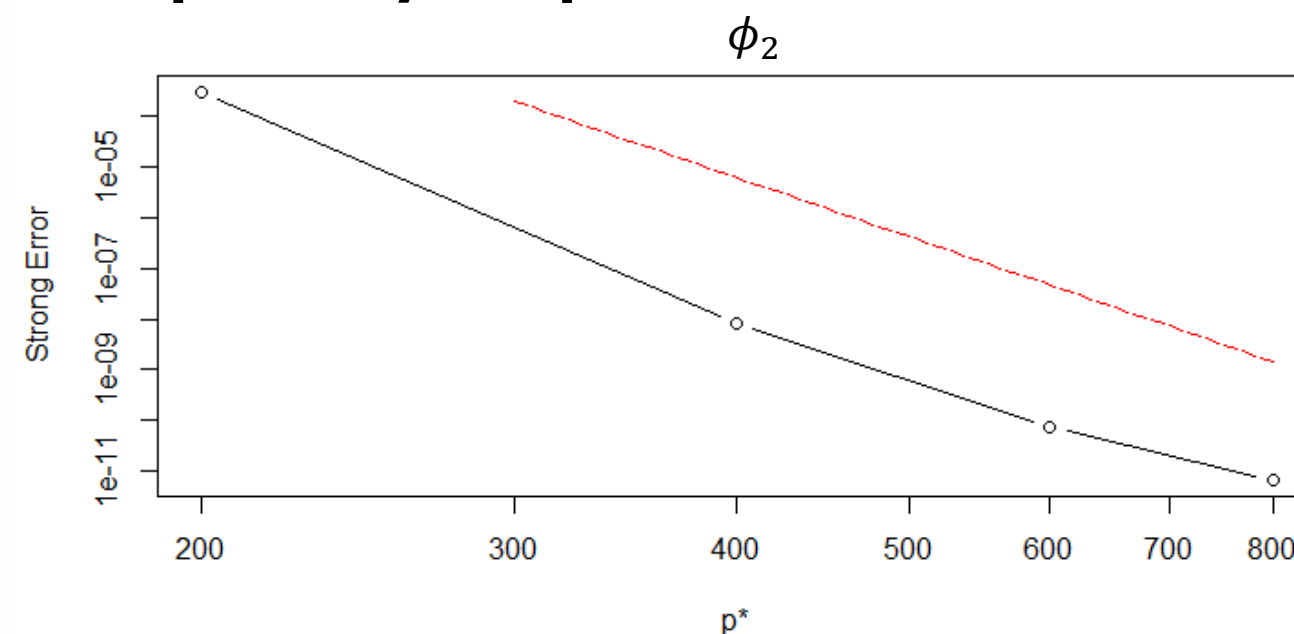
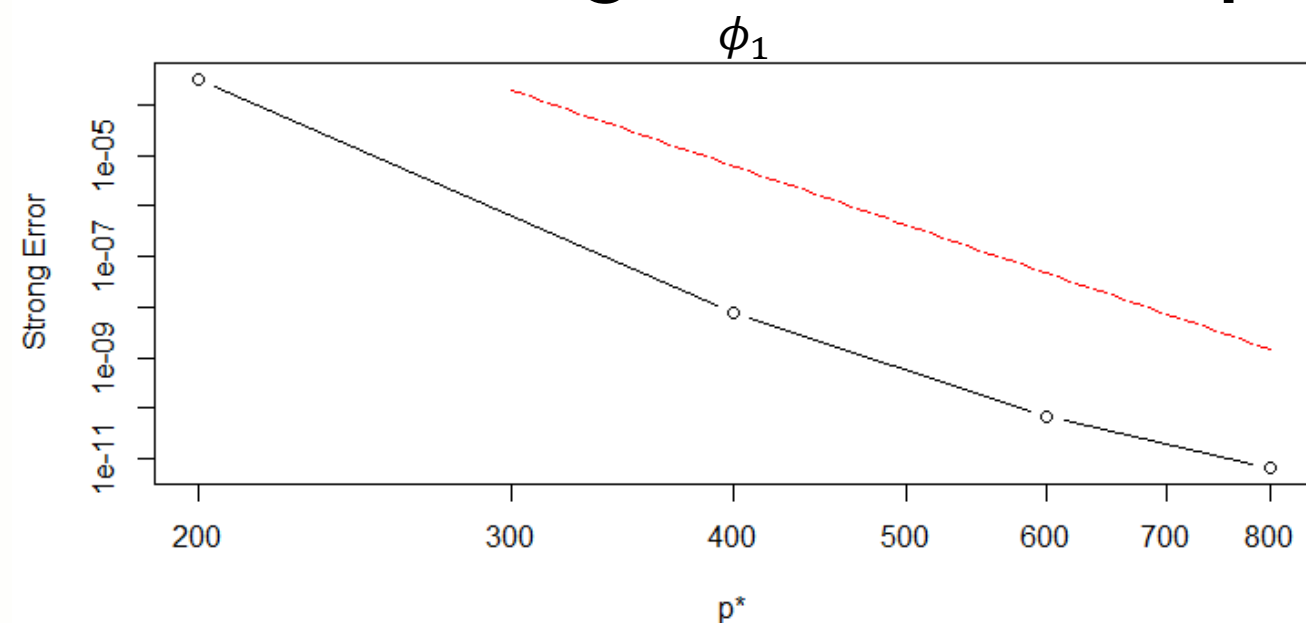
$$y_t = u_t + \sum_{i=1}^{p^*} (\phi + \theta)\theta^{i-1}y_{t-i}$$

Convergence of  $AR(p^*)$  coefficient estimates

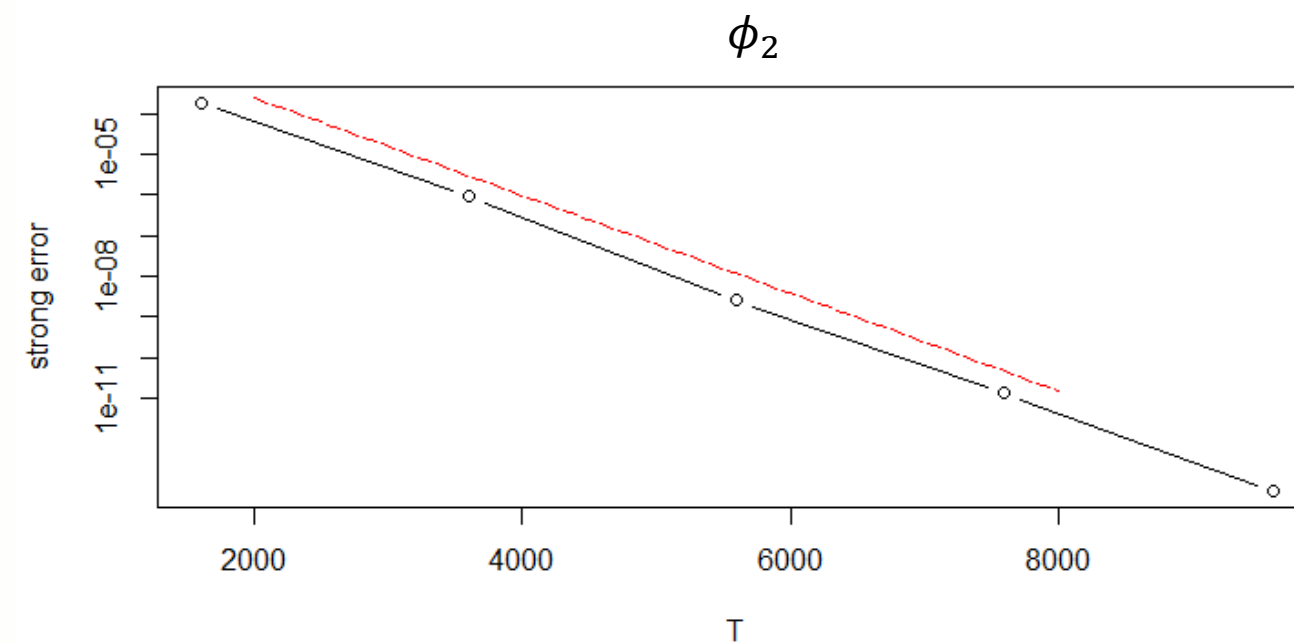
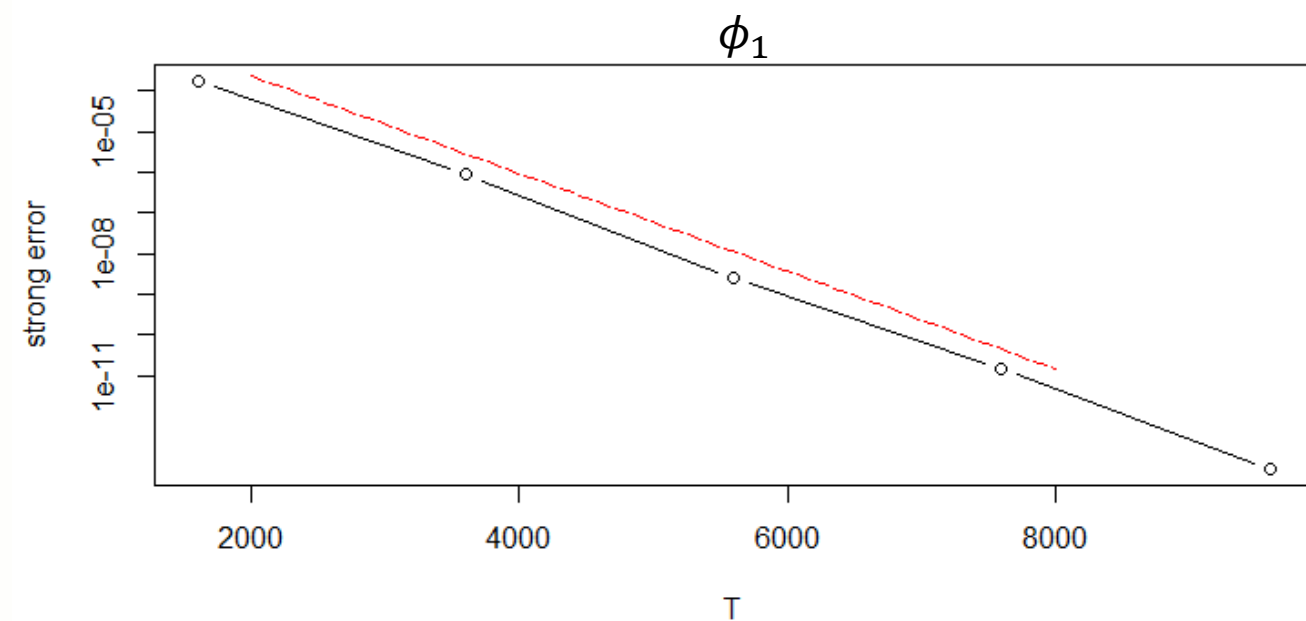
### III Convergence of $AR(p^*)$ coefficient estimates

## Convergence of autoregressive coefficient estimates

### Strong error with respect to $p^*$ – purely explosive ARMA(2,1)



### Strong error with respect to $T$ – purely explosive ARMA(2,1)



### III Convergence of AR( $p^*$ ) coefficient estimates

## Convergence of autoregressive coefficient estimates

### **Hypothesis 1.**

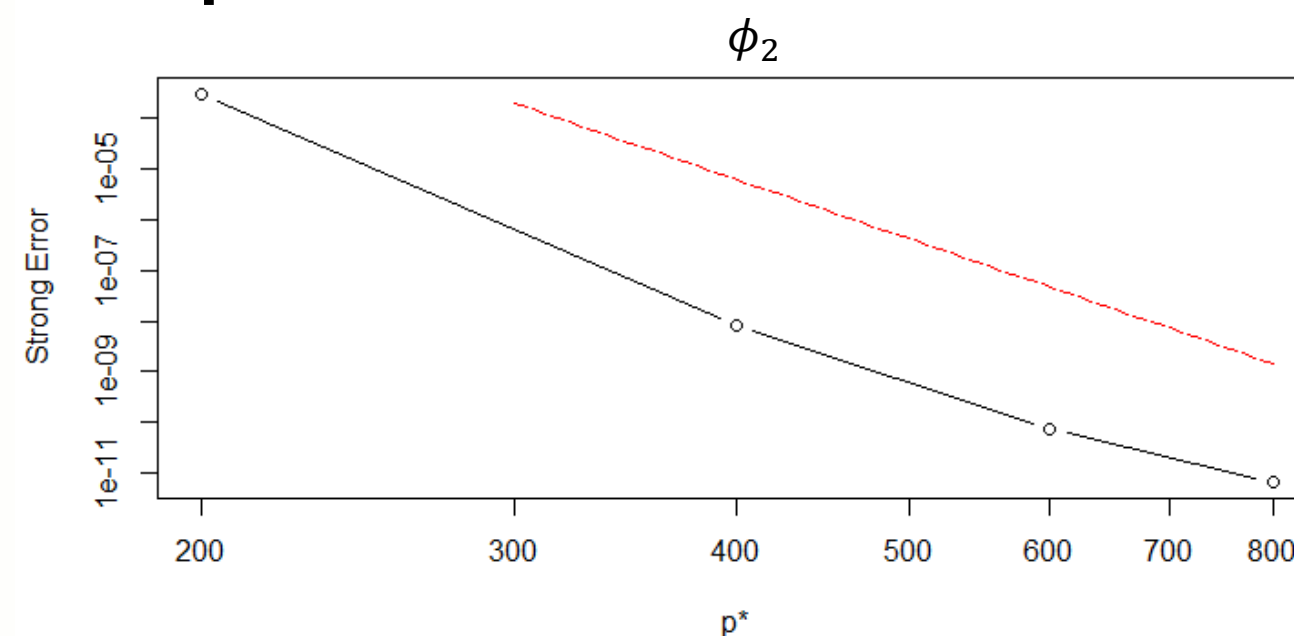
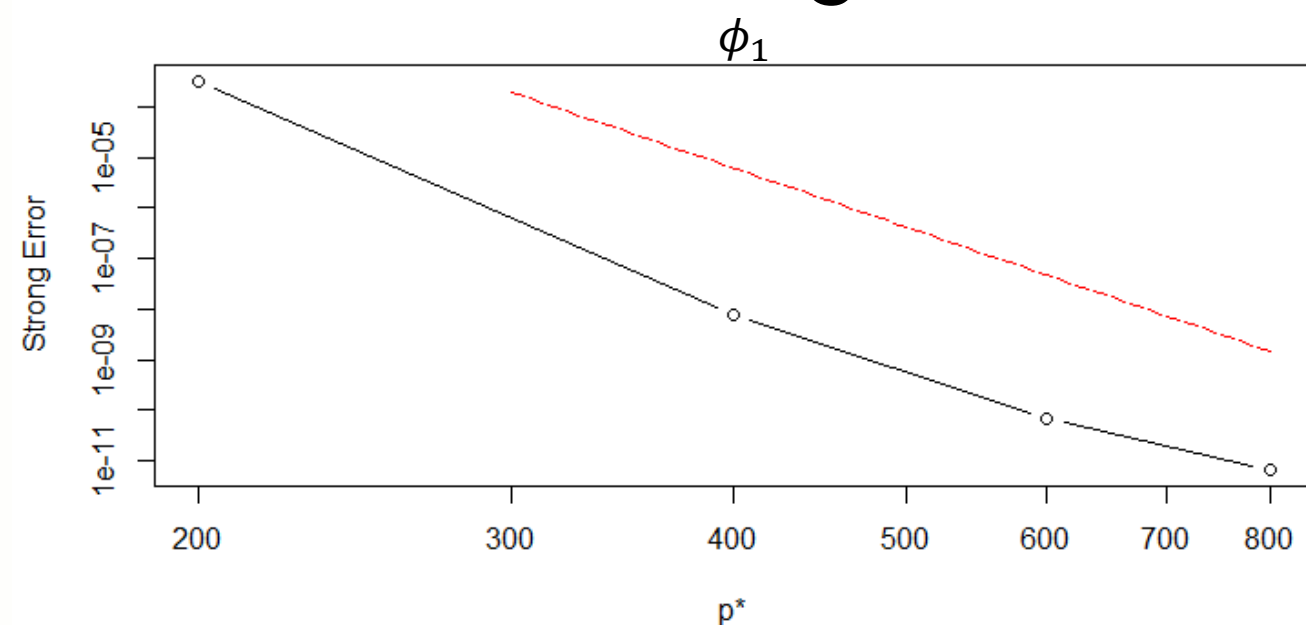
The autoregressive coefficient estimates  $\hat{\phi}(L)$  of an non-stationary ARMA(2,q) process with purely explosive AR component obtained under the AR( $p^*$ ) method converge strongly as  $p^*, T \rightarrow \infty$ , such that

$$\lim_{p^*, T \rightarrow \infty} \hat{\phi}(L) \rightarrow \phi(L) \text{ a.s}$$

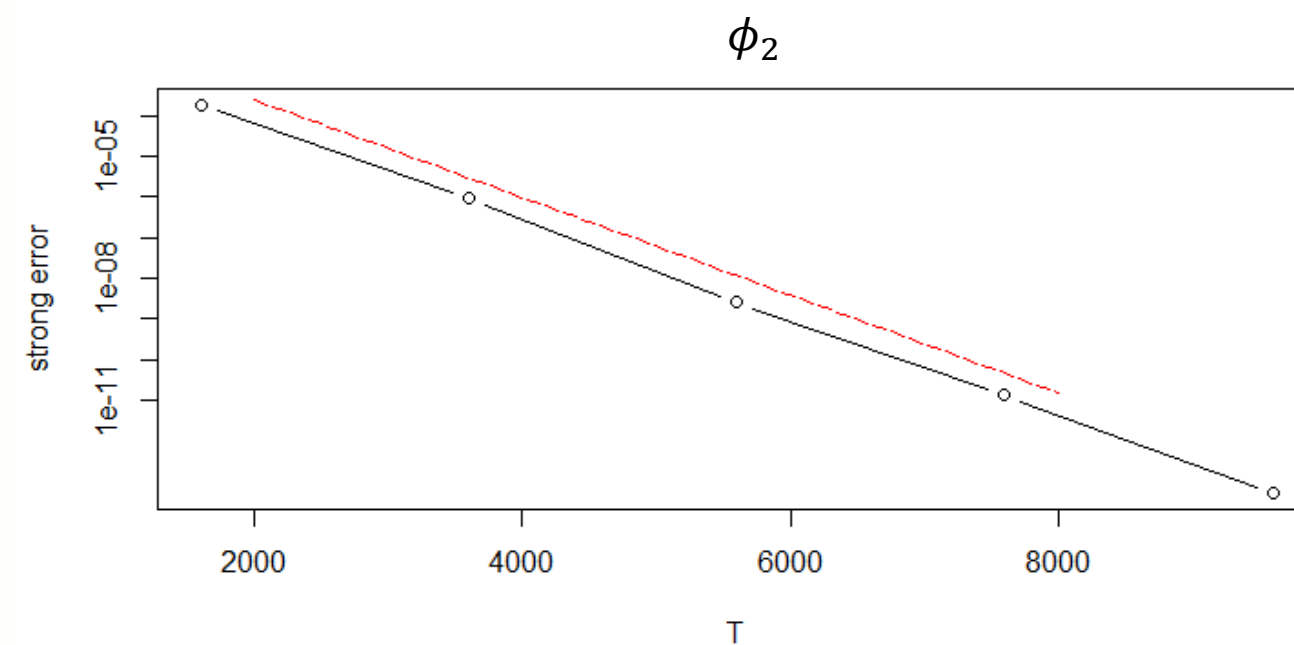
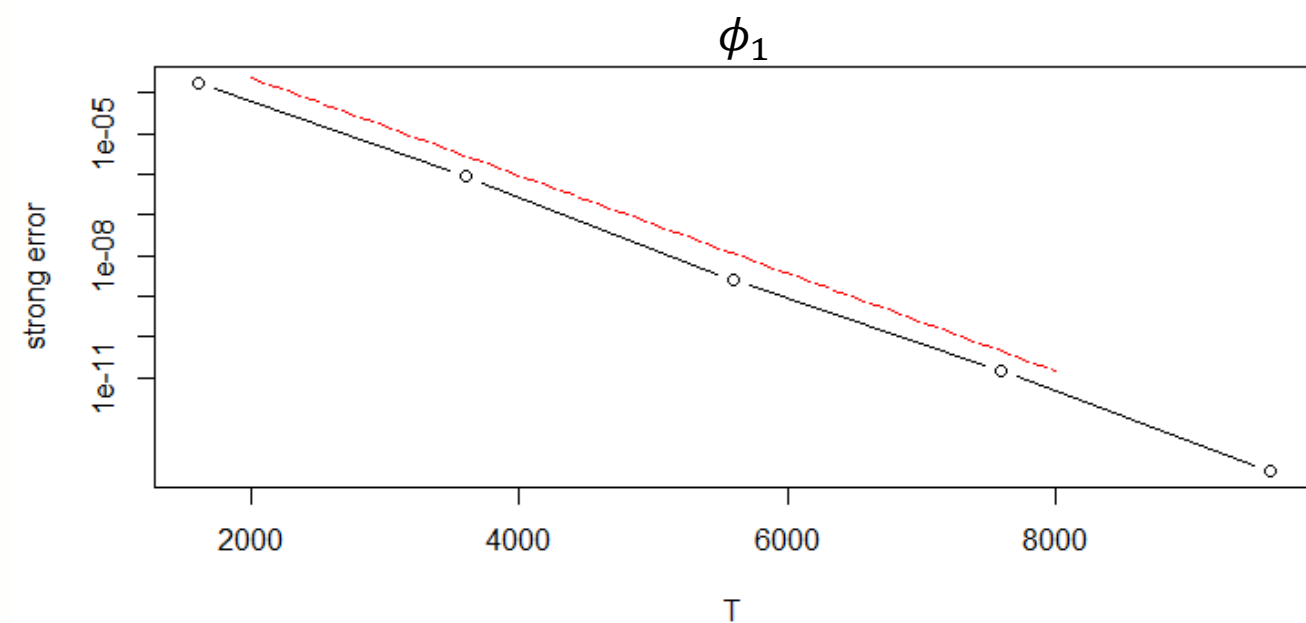
### III Convergence of $AR(p^*)$ coefficient estimates

## Convergence of autoregressive coefficient estimates

### Strong error with respect to $p^*$ – ARMA(2,1)



### Strong error with respect to $T$ – ARMA(2,1)



### III Convergence of $AR(p^*)$ coefficient estimates

## Convergence of autoregressive coefficient estimates

### **Hypothesis 2.**

For  $T$  sufficiently large, the strong order of convergence of the  $AR(p^*)$  autoregressive coefficient estimates of an  $ARMA(2, q)$  with purely explosive AR component under the  $AR(p^*)$  method with respect to  $p^*$

$$O(p^{-12})$$

### **Hypothesis 3.**

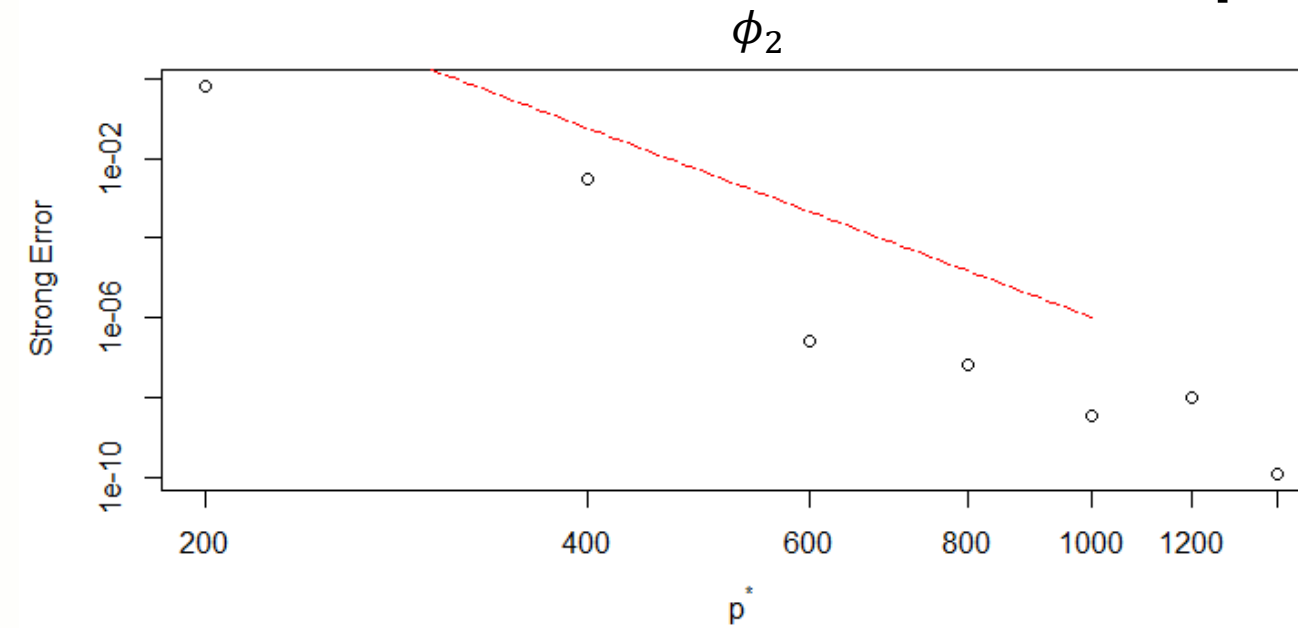
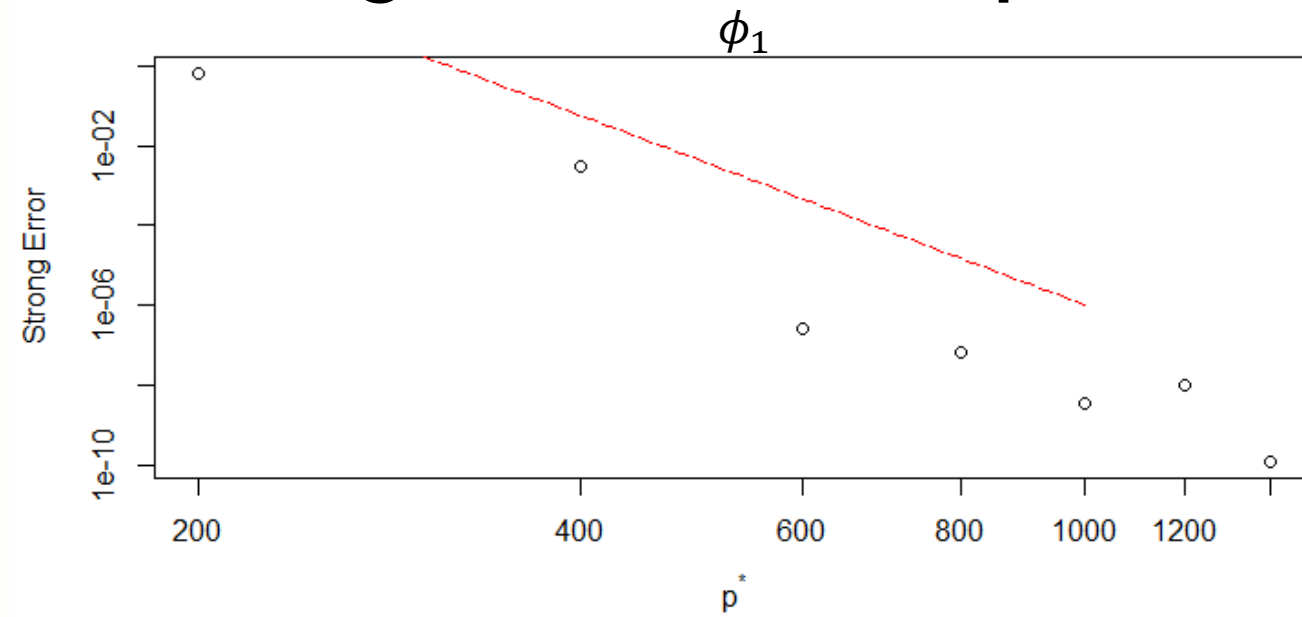
For  $p^*$  sufficiently large, the strong order of convergence of the  $AR(p^*)$  autoregressive coefficient estimates of an  $ARMA(2, q)$  process with purely explosive AR component under the  $AR(p^*)$  method with respect to  $T$  is

$$O(\exp(-12T))$$

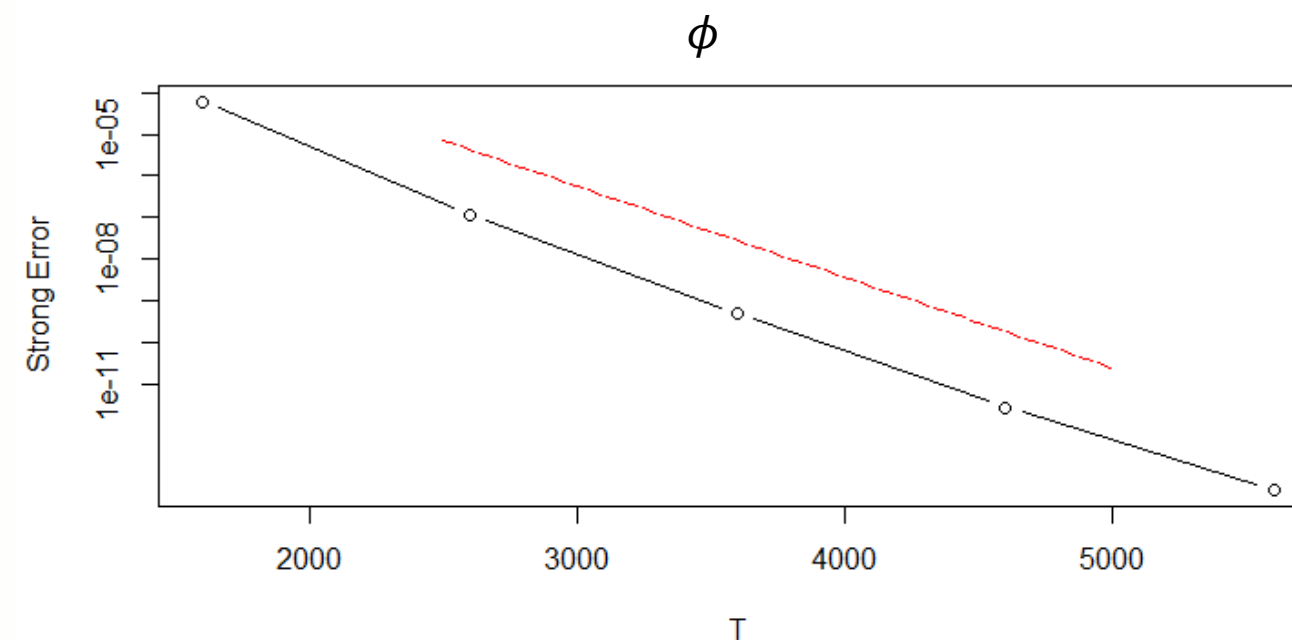
### III Convergence of $AR(p^*)$ coefficient estimates

## Convergence of autoregressive coefficient estimates

Strong error with respect to  $p^*$  – ARMA(2,1) mixed root AR component



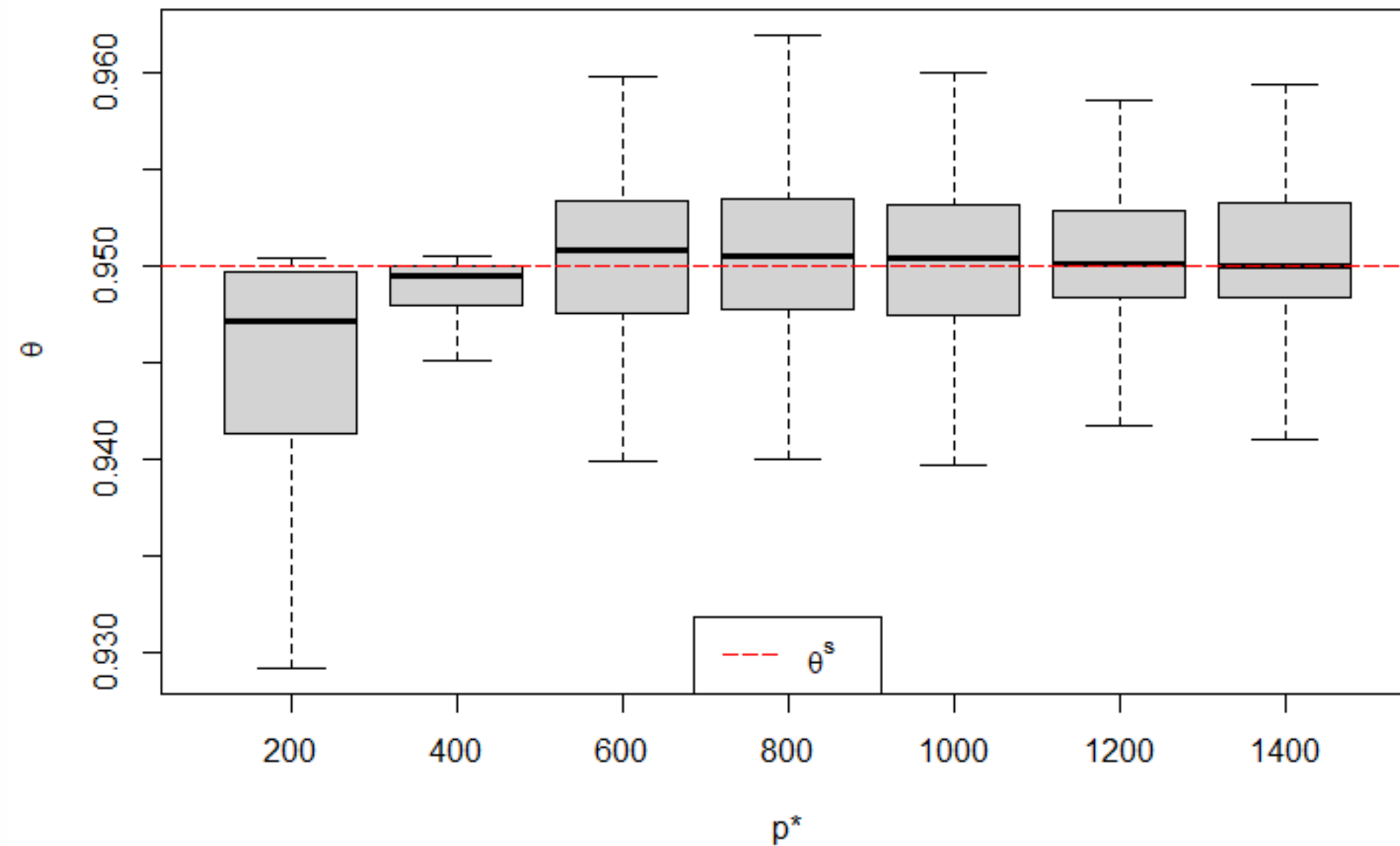
Strong error with respect to  $T$  – ARMA(1,1)



### III Convergence of AR( $p^*$ ) coefficient estimates

## Convergence of Moving Average coefficient estimates

### Distribution of MA coefficient estimates – ARMA(2,1)



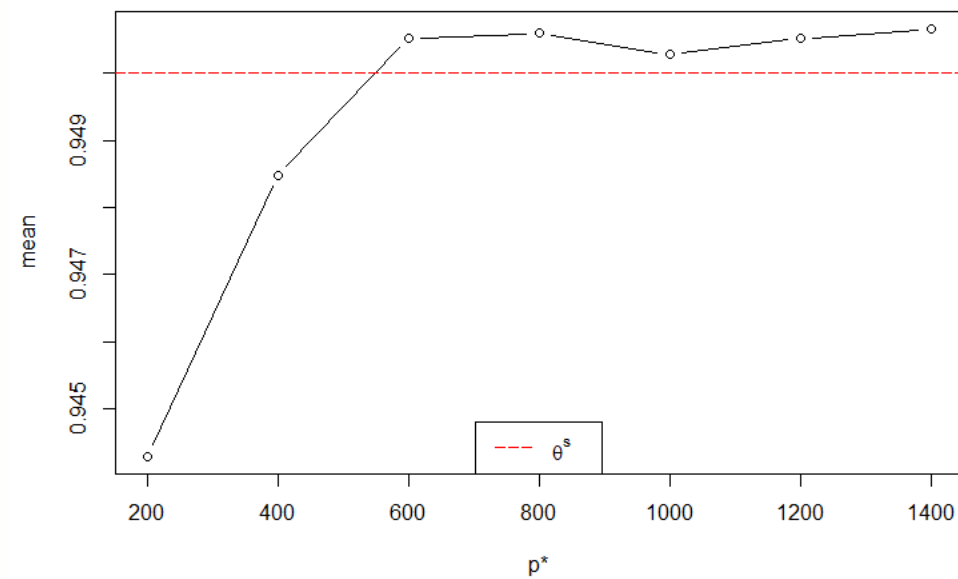


### III Convergence of AR( $p^*$ ) coefficient estimates

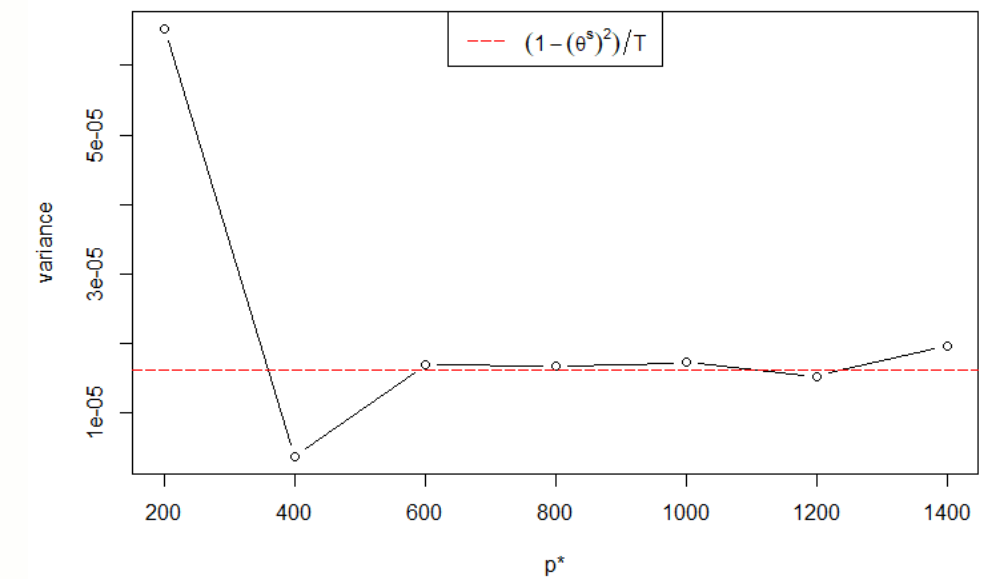
# Convergence of Moving Average coefficient estimates

## Moments of MA coefficient estimate Distributions – ARMA(2,1)

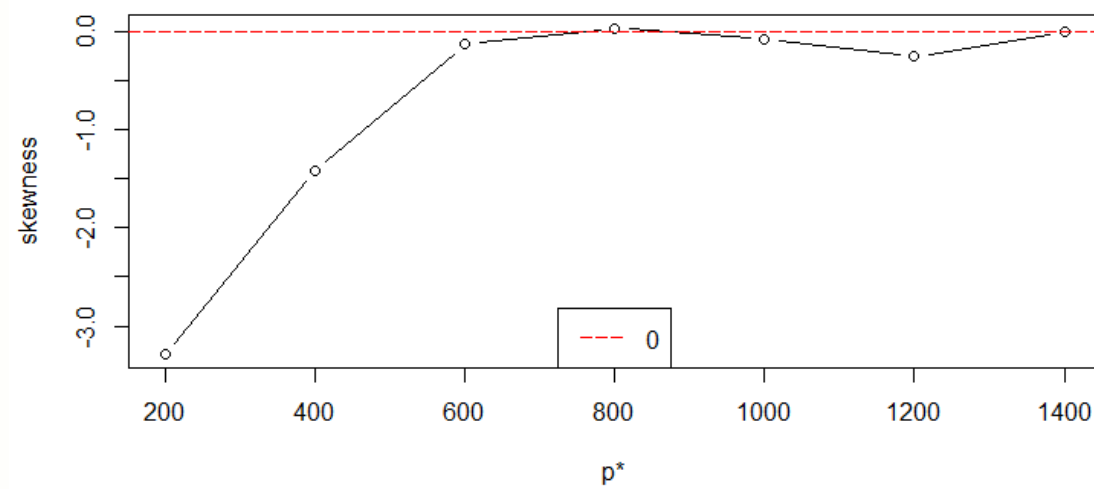
Mean



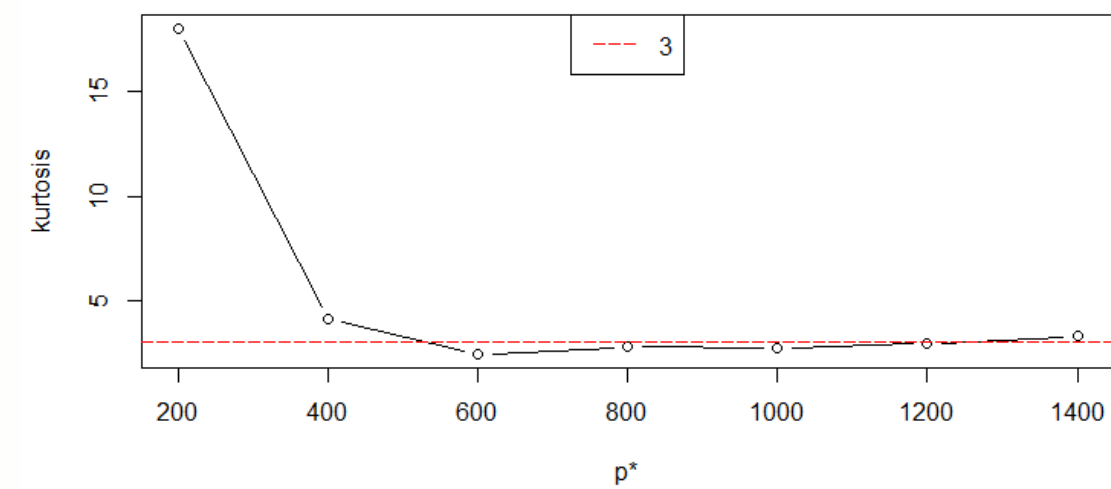
Variance



Skewness



Kurtosis



### III Convergence of AR( $p^*$ ) coefficient estimates

## Convergence of Moving Average coefficient estimates

### **Hypothesis 4.**

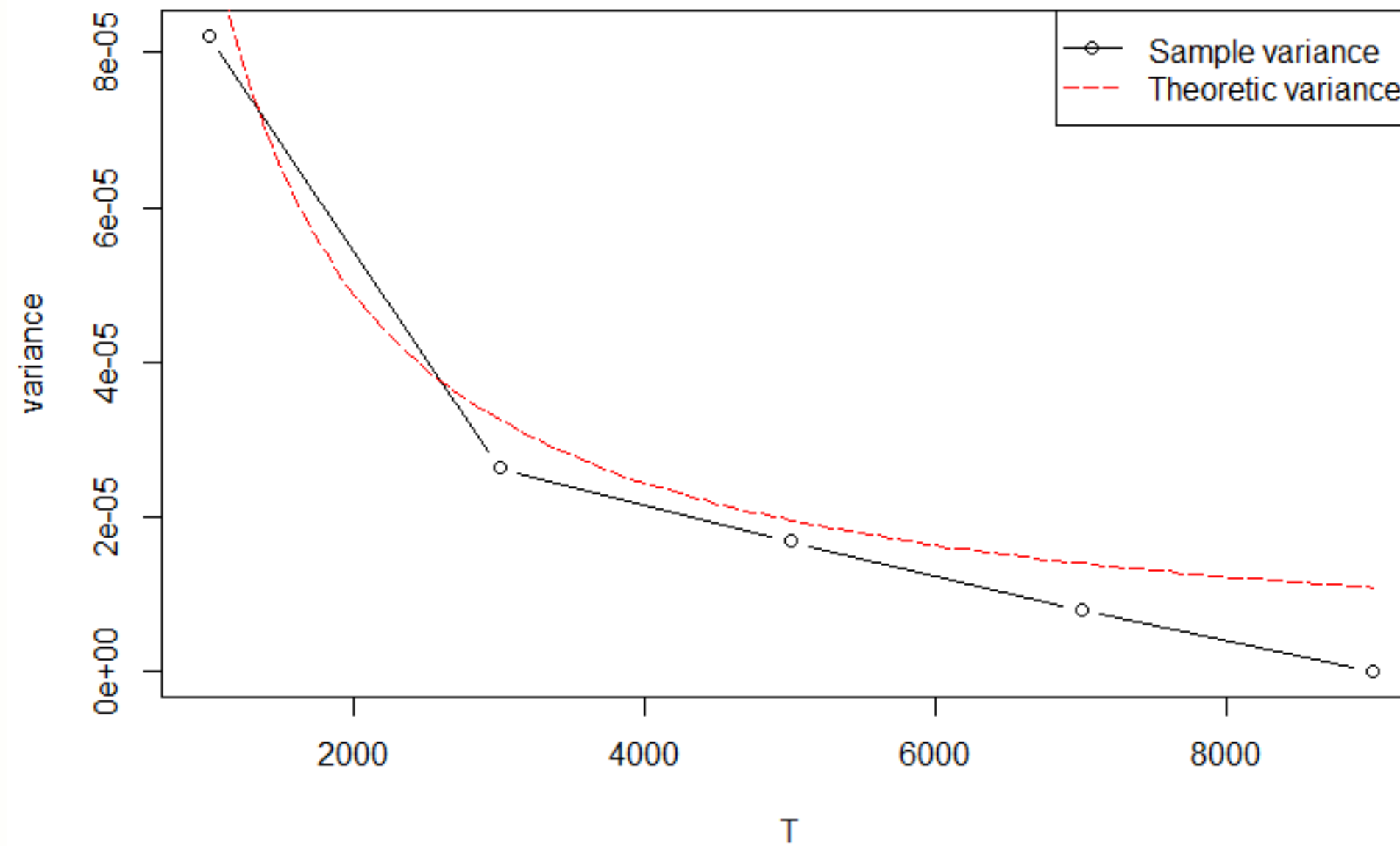
The moving average coefficient estimates  $\hat{\theta}$  of an non-stationary ARMA( $p,1$ ) process obtained under the AR( $p^*$ ) method converge in distribution as  $p^* \rightarrow \infty$  such that

$$\lim_{p^* \rightarrow \infty} \hat{\theta} \xrightarrow{d} N \left( \theta^s, \frac{1 - \theta^{s^2}}{T - p^*} \right)$$

### III Convergence of $AR(p^*)$ coefficient estimates

## Convergence of Moving Average coefficient estimates

### Variance of MA coefficient estimates – ARMA(2,1)



### III Convergence of AR( $p^*$ ) coefficient estimates

## Convergence of Moving Average coefficient estimates

### **Hypothesis 5.**

The moving average coefficient estimates  $\hat{\theta}$  of a non-stationary ARMA( $p,1$ ) process obtained under the AR( $p^*$ ) method converge in probability as  $T, p^* \rightarrow \infty$  such that

$$\lim_{p^*, T \rightarrow \infty} \hat{\theta} \xrightarrow{p} \theta^s$$

Using parameter estimates to create  
stationary time series

## Purely Explosive ARMA process

Given a non-stationary process  $\phi(L)y_t = \theta(L)u_t$  and set of observations  $\{y_t\}_{t=1}^T$ , define  $z_t = \phi(L)y_t$  such that  $z_t = \theta(L)u_t$  is a stationary process.

The filtering procedure is as follows:

### Step 1:

Apply the AR(p\*) method to the series  $\{y_t\}_{t=1}^T$ , and store estimates  $\{\hat{\phi}_1, \dots, \hat{\phi}_p\}$

### Step 2:

Generate series  $\{\hat{z}_t\}_{t=1}^{T-p}$  where  $\hat{z}_t = y_{t+p+1} - \hat{\phi}_1 y_{t+p} - \dots - \hat{\phi}_p y_t$

### Step 3:

Estimate the series  $\{\hat{z}_t\}_{t=1}^{T-p}$  as a stationary MA(q) process to obtain the moving average coefficient estimates  $\{\hat{\theta}_1, \dots, \hat{\theta}_q\}$ .

The final estimated model is:

$$y_t = \hat{\phi}_1 y_{t-1} - \dots - \hat{\phi}_p y_{t-p} + u_t + \hat{\theta}_1 u_{t-1} + \hat{\theta}_q u_{t-q}$$

#### IV Using parameter estimates to create stationary time series

## Mixed Root ARMA process

Given a non-stationary process  $\phi(L)y_t = \theta(L)u_t$  with mixed root AR component and set of observations  $\{y_t\}_{t=1}^T$ , define  $\phi^s(L), \phi^u(L)$  such that

$$\phi^s(L) = \prod_{i=1}^{p^s} \left( \frac{1}{r_i^s} L - 1 \right) ; r_i^s > 1 \text{ for each } i$$

$$\phi^u(L) = \prod_{j=1}^{p^u} \left( \frac{1}{r_j^u} L - 1 \right) ; r_j^u \leq 1 \text{ for each } j$$

$$p^u + p^s = p$$

$$\phi(L) = \phi^s(L)\phi^u(L)$$

From here, define  $z_t = \phi^u(L)y_t$  such that  $\phi^s(L)z_t = \theta(L)u_t$  is a stationary process.

## Mixed Root ARMA process

The filtering procedure for mixed root AR component is as follows:

**Step 1:**

Apply the AR( $p^*$ ) method to the series  $\{y_t\}_{t=1}^T$ , and store estimates  $\{\widehat{\phi}_1, \dots, \widehat{\phi}_p\}$

**Step 2:**

Calculate roots  $\widehat{r}_i^u, \widehat{r}_j^s$  of the characteristic polynomial of coefficient estimates.

**Step 3:**

Generate series  $\{\widehat{z}_t\}_{t=1}^{T-p^u}$  where  $\widehat{z}_t = \widehat{\phi^u(L)} y_t$

**Step 4:**

Estimate the series  $\{\widehat{z}_t\}_{t=1}^{T-p^u}$  as a stationary ARMA( $p^s, q$ ) process to obtain the remaining coefficient estimates  $\{\widehat{\phi^s(L)}, \widehat{\theta}_1, \dots, \widehat{\theta}_q\}$ .

The final estimated model is hence

$$\widehat{\phi^s(L)} \widehat{\phi^u(L)} y_t = u_t + \widehat{\theta}_1 u_{t-1} + \widehat{\theta}_q u_{t-q}$$



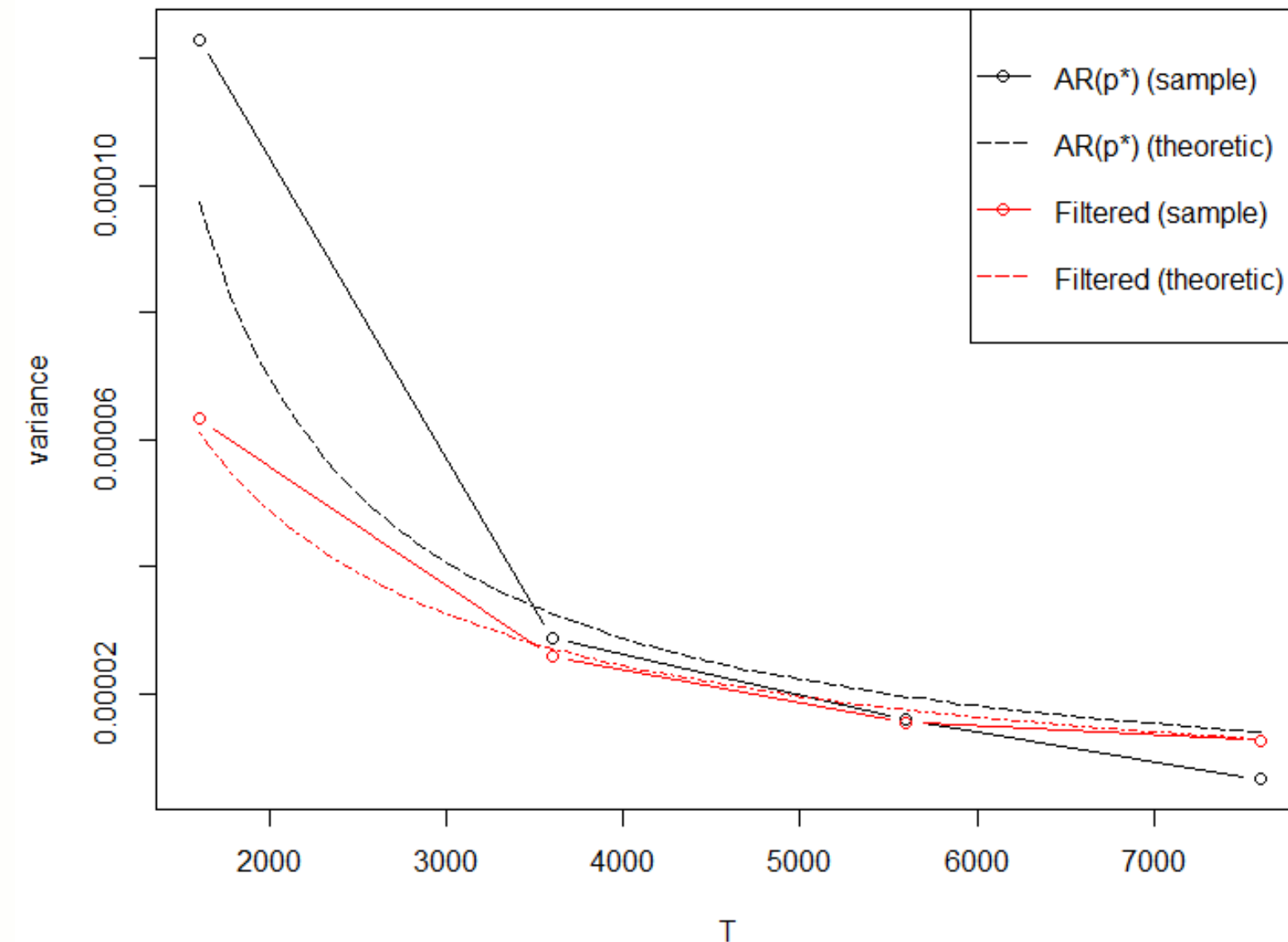
## On $AR(p^*)$ filtering methods

- Common stationary time series methods can be applied to stationary series generated from filtering methods.
- Initial tests suggests that “Filtering method for purely explosive processes” results in lowest variance of coefficient estimates in purely explosive and mixed root cases!

#### IV Using parameter estimates to create stationary time series

## On $AR(p^*)$ filtering methods

- Filtering increases fitting points available to stationary series than  $AR(p^*)$ , reducing standard errors particularly when  $T$  is small.



Variance of Moving Average coefficient estimates of  $ARMA(2,1)$  process with purely explosive AR component ( $p^* = 600$ )

# Computational Errors and inaccuracies

## Non-stationary data and computational accuracy

Observations from non-stationary processes are explosive in nature. Such processes return observations with large absolute values and which differ largely from initial observations. Numerical overflow and round off error may occur as a result.

## Optimization errors

The conditional negative log likelihood of a non-stationary ARMA process is known to have local minima apart from the global minimum. This remains the case in the  $AR(p^*)$  estimation approximation method.

BFGS (Broyden, Fletcher, Goldfarb and Shanno) method was utilised throughout this project. Unusual values were investigated via multidimensional grid search.

# Conclusions & Further Research

## Theoretical Findings

This project is motivated by the work on Hanzon & Scherrer (2019) and serves to forming conjectures that may build upon their work. The results of this project are being considered with conjectures and further research being considered.

## Applications to forecasting

The  $AR(p^*)$  and resulting filtering methods may prove useful in forecasting time series. Optimization is known to converge to local minima. In forecasting, more rigorous optimization methods may be necessary.

## Generalization to higher order ARMA processes

Investigate whether higher order processes and SARMA models display the same behavior under the  $AR(p^*)$  method and if not, in what ways do they differ. Further progression of this would be to vector ARMA processes and linear scalar ARMA processes with non-Gaussian errors.

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**Thank you  
for listening!**

**Questions are  
welcome!**

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