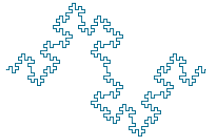


High-Throughput Sequencing Course

Statistical Inference: Sources of Variability

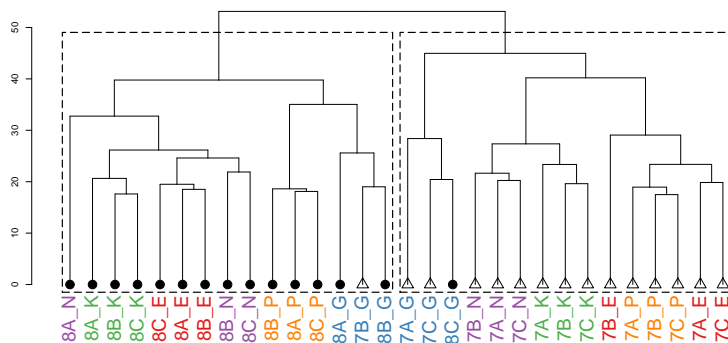
Biostatistics and Bioinformatics



Summer 2019



CLASS DISCOVERY



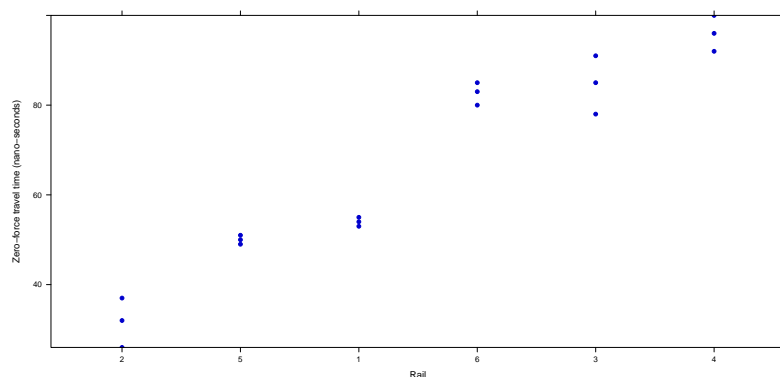
INTRA- AND INTER-SUBJECT VARIABILITY

- ▶ In most experiments, including RNA-Seq, the variability may not be exclusively due to measurement error
- ▶ Another source could be due to repeated measurements
- ▶ or sampling from strains or cell lines
- ▶ or due to batch effects (e.g., team effect)
- ▶ We will motivate these ideas using a classical toy example
- ▶ We will illustrate the caveats of properly accounting for these two sources of variability through two simulation studies

RAILS DATA

- Observation adjusted travel time for ultrasonic head-waves in the rail (nanoseconds).
- Data set: $n = 6$ rails; the travel time is sampled three ($m = 3$) times per rail
- Eighteen ($n \times m = 18$) measurements
- Six ($n = 6$) experimental units
- Implicit assumption: The six rails are randomly selected from a *large* pool of rails
- What is of interest is neither the batch or any of these 6 rails (specifically)
- What is of interest is the population (the huge pool)

RAIL DATA



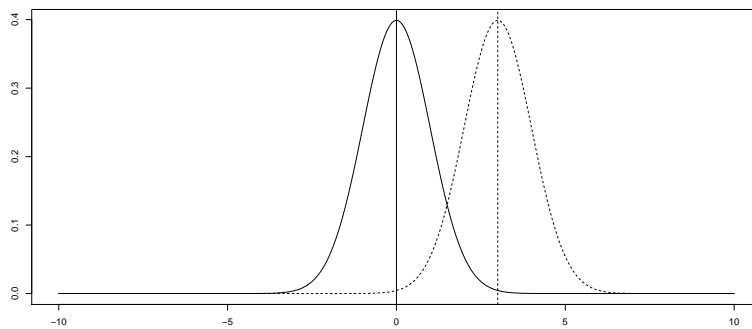
RAIL DATA: MODEL FORMULATION

- μ denotes the *true* travel time
- μ is an unknown fixed quantity
- Y_i denotes the *observed* travel time (for observation $i = 1, \dots, 18$)
- In absence of noise, true value μ is observed
- In other words, $Y_i = \mu$ for $i = 1, \dots, 18$

IMPORTANT FACT ABOUT NORMAL DISTRIBUTION

- ▶ Consider a normal distribution with mean 0 and standard deviation σ
- ▶ If the data are shifted by a constant μ , then
 1. resulting distribution remains normal
 2. The mean of the new distribution is $\mu + 0 = \mu$
 3. Its standard deviation remains unchanged
- ▶ The last two (but not first) property are true for any distribution

SHIFT NORMAL DISTRIBUTION



RAIL DATA: SIMPLE MODEL

- ▶ What is observed is a distorted version of μ

$$Y_i = \mu + \epsilon_i$$

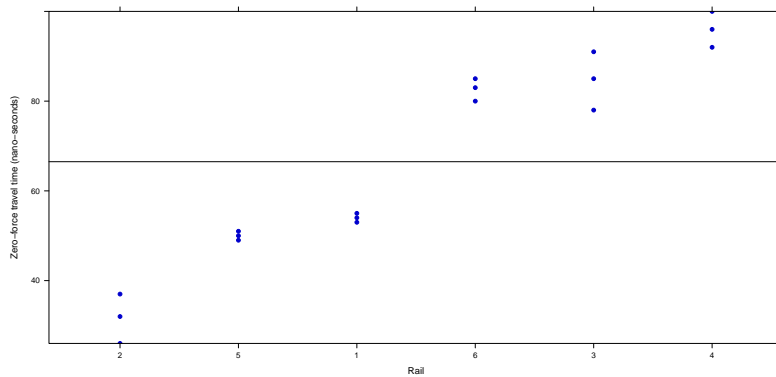
- ▶ Notes:
 - ▶ Y_i is observable
 - ▶ ϵ_i is *not* observable
 - ▶ μ is an unknown parameter
- ▶ The variability observed here is exclusively attributed to the measurement error ϵ_i

LINEAR MODEL

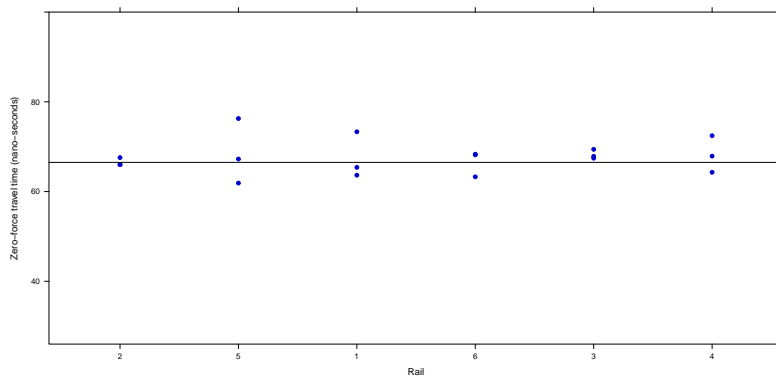
```
mod0 <- summary(lm(travel ~ 1, data = Rail))
mod0

##
## Call:
## lm(formula = travel ~ 1, data = Rail)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -40.50 -16.25   0.00  18.50  33.50
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   66.500      5.573   11.93 1.1e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 23.65 on 17 degrees of freedom
```

LINEAR MODEL: IS THIS A REASONABLE MODEL?



LINEAR MODEL: MAY BE MORE REASONABLE IN THIS CASE?



RAIL DATA: ACCOUNT FOR TWO SOURCE OF VARIABILITY

- What is observed is a distorted version of μ
- It is distorted by a ra
- Y_{ij} : Index the rail by $i = 1, \dots, 6$ and the replicate by $j = 1, 2, 3$
- Y_{23} : The observation for the third replicate for rail 2
- Model

$$Y_{ij} = \mu + b_i + \epsilon_{ij}$$

- Notes:
 - Y_{ij} is observable
 - b_i is *not* observable
 - ϵ_{ij} is *not* observable
 - μ is an unknown parameter

LINEAR MIXED EFFECTS MODEL

```
lme(travel ~ 1, random = ~1 | Rail, data = Rail)

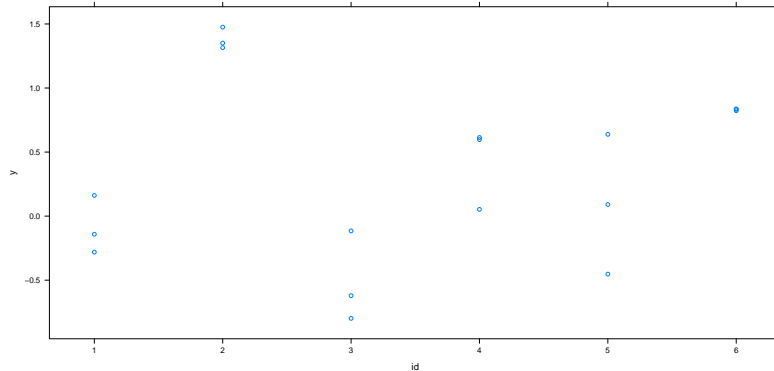
## Linear mixed-effects model fit by REML
## Data: Rail
## Log-restricted-likelihood: -61.0885
## Fixed: travel ~ 1
## (Intercept)
## 66.5
##
## Random effects:
## Formula: ~1 | Rail
## (Intercept) Residual
## StdDev: 24.80547 4.020779
##
## Number of Observations: 18
## Number of Groups: 6
```

IS THE MIXED MODEL ADEQUATE?

- Assumptions:
 - b_i is normally distributed $N[0, \sigma_b^2]$
 - σ_b^2 does *not* depend on i (homoscedastic)
 - ϵ_{ij} is normally distributed $N[0, \sigma_e^2]$
 - σ_e^2 does *not* depend on i or j (homoscedastic)
 - Error model is additive (could be multiplicative)

EXAMPLE 1: SETUP

- What are the ramifications for ignoring the clustering?
- We will sample 6 experimental units each with three replicates
- $\mu = 0, \sigma_e = 0.25, \sigma_b = 0.5$



EXAMPLE 1: SIMULATION

- Simulation outline
 1. Simulate a data set
 2. Test $H_0 : \mu = 0$ ignoring the random effect (save P -value)
 3. Test $H_0 : \mu = 0$ accounting for the random effect (save P -value)
- Repeat the three steps 999 additional times
- Given that the *true* $\mu = 0$ (by design), we would expect 50 of these P -values to be less than 0.05
- Why?

EXAMPLE 1: RESULTS

```
set.seed(210)
res = replicate(83, sim.ranef(3, 6, 0.25, 0.5, verbose = FALSE))
mean(res[1, ] < 0.05)

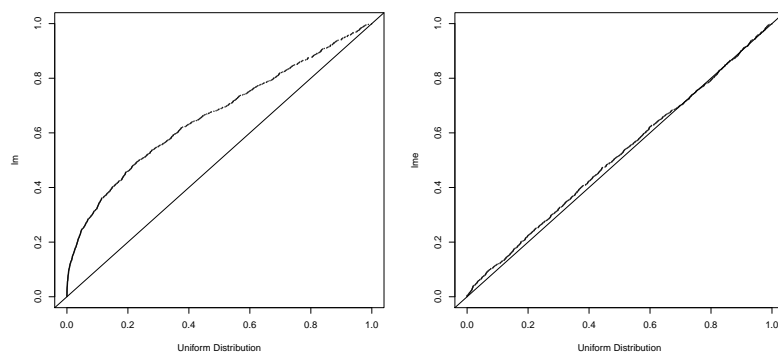
## [1] 0.247

mean(res[2, ] < 0.05)

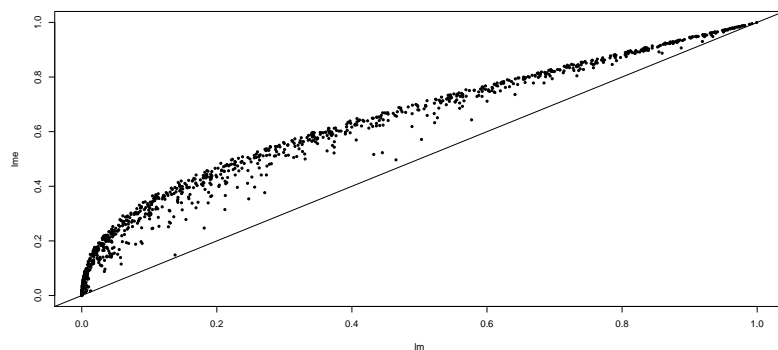
## [1] 0.072
```

- The empirical type I error rate when not accounting for the random effect is 0.25.
- This inflated by a factor of 4.9.
- The empirical error rate when accounting for the random effect is slightly inflated
- This is due to the small sample size ($n = 6$)
- More on this later.

EXAMPLE 1: RESULTS



EXAMPLE 1: RESULTS



EXAMPLE 1: RESULTS

- Now, we repeat the simulation with a larger sample size

```
res <- replicate(B3, sim.ranef(3, 50, 0.25, 0.5, verbose = FALSE))
mean(res[1, ] < 0.05)

## [1] 0.215

mean(res[2, ] < 0.05)

## [1] 0.052
```

- The empirical type I error when not accounting for the random effect remains inflated by a factor of 4.3.
- The empirical type I error when accounting for the random effect is now right about the nominal level of 0.05

EXAMPLE 2: SETUP

- ▶ Now consider the two-sample problem we have previously considered with a twist
- ▶ Question: Does treatment alter the distribution of the RNA level of a given gene?
- ▶ Assumptions:
 - ▶ the RNA level for the untreated group follows a normal distribution with mean μ_0 and variance σ^2
 - ▶ The RNA level for the treated group follows a normal distribution with mean μ_1 and variance σ^2
- ▶ Sample n units from each treatments in replicates of 3
- ▶ Apply the two-sample t-test which does not account for the clustering

EXAMPLE 2: SIMULATION

```
set.seed(2314)
# Simulate with no clustering effect (sb=0)
pval0 = replicate(B3, sim.twosample.clustered(3, 10, 0.25, 0))
# Simulate with no clustering effect (sb>0)
pval1 = replicate(B3, sim.twosample.clustered(3, 10, 0.25, 0.5))
mean(pval0 < 0.05)

## [1] 0.049

mean(pval1 < 0.05)

## [1] 0.252
```

- ▶ The empirical type I error when there is no clustering effect is 0.049
- ▶ The empirical type I error when there is a clustering effect is 0.25
- ▶ This off by a factor of 5!