Hidden Markov Models

Tapas Kanungo

Center for Automation Research
University of Maryland

Web: www.cfar.umd.edu/~kanungo

Email: kanungo@cfar.umd.edu

Outline

- 1. Markov models
- 2. Hidden Markov models
- 3. Forward/Backward algorithm
- 4. Viterbi algorithm
- 5. Baum-Welch estimation algorithm

Markov Models

• Observable states:

$$1, 2, \dots, N$$

• Observed sequence:

$$q_1,q_2,\ldots,q_t,\ldots,q_T$$

• First order Markov assumption:

$$P(q_t = j | q_{t-1} = i, q_{t-2} = k, ...) = P(q_t = j | q_{t-1} = i)$$

• Stationarity:

$$P(q_t = j | q_{t-1} = i) = P(q_{t+l} = j | q_{t+l-1} = i)$$

Markov Models

• State transition matrix A:

$$A = egin{bmatrix} a_{11} & a_{12} & \cdots & a_{1j} & \cdots & a_{1N} \ a_{21} & a_{22} & \cdots & a_{2j} & \cdots & a_{2N} \ dots & dots & \cdots & dots & \cdots & dots \ a_{i1} & a_{i2} & \cdots & a_{ij} & \cdots & a_{iN} \ dots & dots & \cdots & dots & \cdots & dots \ a_{N1} & a_{N2} & \cdots & a_{Nj} & \cdots & a_{NN} \ \end{bmatrix}$$

where

$$a_{ij} = P(q_t = j | q_{t-1} = i)$$
 $1 \le i, j, \le N$

 \bullet Constraints on a_{ij} :

$$egin{array}{lll} a_{ij} & \geq & 0, & & orall i,j \ \sum\limits_{i=1}^{N} a_{ij} & = & 1, & & orall i \end{array}$$

- States:
 - 1. Rainy (R)
 - 2. Cloudy (C)
 - 3. Sunny (S)
- State transition probability matrix:

$$A = \begin{bmatrix} 0.4 & 0.3 & 0.3 \\ 0.2 & 0.6 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{bmatrix}$$

• Compute the probability of observing SSRRSCS given that today is S.

Basic conditional probability rule:

$$P(A,B) = P(A|B)P(B)$$

The Markov chain rule:

$$P(q_1, q_2, \dots, q_T)$$

$$= P(q_T | q_1, q_2, \dots, q_{T-1}) P(q_1, q_2, \dots, q_{T-1})$$

$$= P(q_T | q_{T-1}) P(q_1, q_2, \dots, q_{T-1})$$

$$= P(q_T | q_{T-1}) P(q_{T-1} | q_{T-2}) P(q_1, q_2, \dots, q_{T-2})$$

$$= P(q_T | q_{T-1}) P(q_{T-1} | q_{T-2}) \cdots P(q_2 | q_1) P(q_1)$$

ullet Observation sequence O:

$$O = (S, S, S, R, R, S, C, S)$$

• Using the chain rule we get:

$$\begin{split} &P(O|model)\\ &= P(S,S,S,R,R,S,C,S|model)\\ &= P(S)P(S|S)P(S|S)P(R|S)P(R|R)\times\\ &\quad P(S|R)P(C|S)P(S|C)\\ &= \pi_3a_{33}a_{33}a_{31}a_{11}a_{13}a_{32}a_{23}\\ &= (1)(0.8)^2(0.1)(0.4)(0.3)(0.1)(0.2)\\ &= 1.536\times 10^{-4} \end{split}$$

• The prior probability $\pi_i = P(q_1 = i)$

• What is the probability that the sequence remains in state *i* for exactly *d* time units?

$$p_i(d) = P(q_1 = i, q_2 = i, ..., q_d = i, q_{d+1} \neq i, ...)$$

= $\pi_i(a_{ii})^{d-1}(1 - a_{ii})$

- Exponential Markov chain duration density.
- What is the expected value of the duration d in state i?

$$\bar{d}_{i} = \sum_{d=1}^{\infty} dp_{i}(d)$$

$$= \sum_{d=1}^{\infty} d(a_{ii})^{d-1}(1 - a_{ii})$$

$$= (1 - a_{ii}) \sum_{d=1}^{\infty} d(a_{ii})^{d-1}$$

$$= (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \sum_{d=1}^{\infty} (a_{ii})^{d}$$

$$= (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \left(\frac{a_{ii}}{1 - a_{ii}}\right)$$

$$= \frac{1}{1 - a_{ii}}$$

• Avg. number of consecutive sunny days =

$$\frac{1}{1 - a_{33}} = \frac{1}{1 - 0.8} = 5$$

- Avg. number of consecutive cloudy days = 2.5
- Avg. number of consecutive rainy days = 1.67

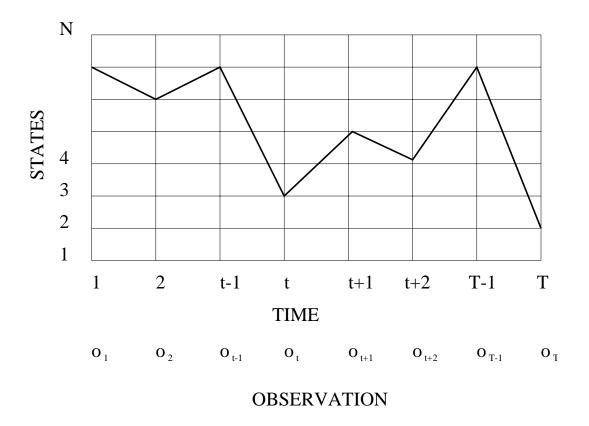
Hidden Markov Models

- States are not observable
- Observations are probabilistic functions of state
- State transitions are still probabilistic

Urn and Ball Model

- N urns containing colored balls
- M distinct colors of balls
- Each urn has a (possibly) different distribution of colors
- Sequence generation algorithm:
 - 1. Pick initial urn according to some random process.
 - 2. Randomly pick a ball from the urn and then replace it
 - 3. Select another urn according a random selection process associated with the urn
 - 4. Repeat steps 2 and 3

The Trellis



Elements of Hidden Markov Models

- \bullet N the number of hidden states
- Q set of states $Q = \{1, 2, \dots, N\}$
- \bullet M the number of symbols
- V set of symbols $V = \{1, 2, \dots, M\}$
- \bullet A the state-transition probability matrix.

$$a_{ij} = P(q_{t+1} = j | q_t = i) \quad 1 \le i, j, \le N$$

 \bullet B – Observation probability distribution:

$$B_i(k) = P(o_t = k | q_t = j) \quad 1 \le k \le M$$

• π – the initial state distribution:

$$\pi_i = P(q_1 = i) \quad 1 \le i \le N$$

• λ – the entire model $\lambda = (A, B, \pi)$

Three Basic Problems

- 1. Given observation $O=(o_1,o_2,\ldots,o_T)$ and model $\lambda=(A,B,\pi),$ efficiently compute $P(O|\lambda).$
 - Hidden states complicate the evaluation
 - Given two models λ_1 and λ_2 , this can be used to choose the better one.
- 2. Given observation $O=(o_1,o_2,\ldots,o_T)$ and model λ find the optimal state sequence $q=(q_1,q_2,\ldots,q_T)$.
 - Optimality criterion has to be decided (e.g. maximum likelihood)
 - "Explanation" for the data.
- 3. Given $O=(o_1,o_2,\ldots,o_T),$ estimate model parameters $\lambda=(A,B,\pi)$ that maximize $P(O|\lambda).$

Solution to Problem 1

- Problem: Compute $P(o_1, o_2, \dots, o_T | \lambda)$
- Algorithm:
 - Let $q = (q_1, q_2, \dots, q_T)$ be a state sequence.
 - Assume the observations are independent:

$$egin{array}{lll} P(O|q,\lambda) &=& \prod\limits_{i=1}^{T} P(o_{t}|q_{t},\lambda) \ &=& b_{q_{1}}(o_{1})b_{q_{2}}(o_{2})\cdots b_{q_{T}}(o_{T}) \end{array}$$

- Probability of a particular state sequence is:

$$P(q|\lambda) = \pi_{q_1} a_{q_1 q_2} a_{q_2 q_3} \cdots a_{q_{T-1} q_T}$$

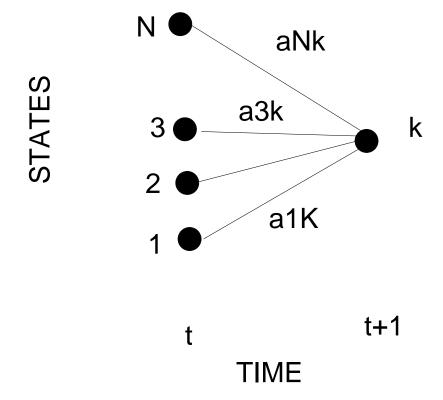
- Also, $P(O, q|\lambda) = P(O|q, \lambda)P(q|\lambda)$
- Enumerate paths and sum probabilities:

$$P(O|\lambda) = \sum_{q} P(O|q,\lambda) P(q|\lambda)$$

ullet N^T state sequences and O(T) calculations.

Complexity: $O(TN^T)$ calculations.

Forward Procedure: Intuition



Forward Algorithm

• Define forward variable $\alpha_t(i)$ as:

$$\alpha_t(i) = P(o_1, o_2, \dots, o_t, q_t = i | \lambda)$$

- ullet $lpha_t(i)$ is the probability of observing the partial sequence $(o_1,o_2\ldots,o_t)$ such that the state q_t is i.
- Induction:
 - 1. Initialization: $\alpha_1(i) = \pi_i b_i(o_1)$
 - 2. Induction:

$$lpha_{t+1}(j) = \left[\sum\limits_{i=1}^{N} lpha_t(i) a_{ij}
ight] b_j(o_{t+1})$$

3. Termination:

$$P(O|\lambda) = \sum\limits_{i=1}^{N} lpha_T(i)$$

• Complexity: $O(N^2T)$.

Example

Consider the following coin-tossing experiment:

	State 1	State 2	State 3
P(H)	0.5	0.75	0.25
P(T)	0.5	0.25	0.75

- state-transition probabilities equal to 1/3
- initial state probabilities equal to 1/3
 - 1. You observe O=(H,H,H,H,T,H,T,T,T). What state sequence, q, is most likely? What is the joint probability, $P(O,q|\lambda)$, of the observation sequence and the state sequence?
 - 2. What is the probability that the observation sequence came entirely of state 1?

3. Consider the observation sequence

$$\tilde{O} = (H,T,T,H,T,H,H,T,T,H).$$

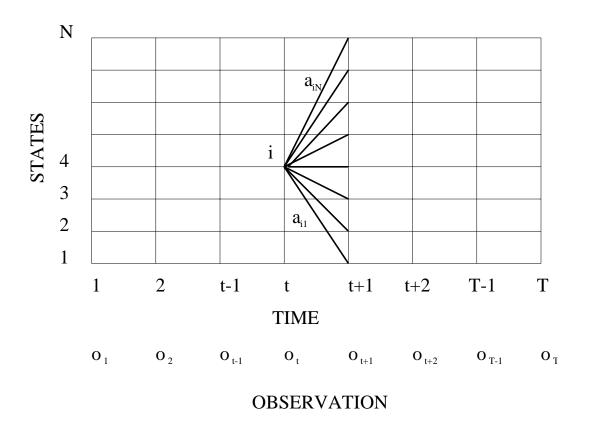
How would your answers to parts 1 and 2 change?

4. If the state transition probabilities were:

$$A' = \begin{bmatrix} 0.9 & 0.45 & 0.45 \\ 0.05 & 0.1 & 0.45 \\ 0.05 & 0.45 & 0.1 \end{bmatrix},$$

how would the new model λ' change your answers to parts 1-3?

Backward Algorithm



Backward Algorithm

ullet Define backward variable $eta_t(i)$ as:

$$eta_t(i) = P(o_{t+1}, o_{t+2}, \ldots, o_T | q_t = i, \lambda)$$

- ullet $eta_t(i)$ is the probability of observing the partial sequence $(o_{t+1},o_{t+2}\dots,o_T)$ such that the state q_t is i.
- Induction:
 - 1. Initialization: $\beta_T(i) = 1$
 - 2. Induction:

$$egin{array}{lll} eta_t(i) &= \sum\limits_{j=1}^N a_{ij} b_j(o_{t+1}) eta_{t+1}(j), \ &1 \leq i \leq N, \ &t = T-1, \ldots, 1 \end{array}$$

Solution to Problem 2

- Choose the most likely path
- ullet Find the path (q_1,q_t,\ldots,q_T) that maximizes the likelihood:

$$P(q_1,q_2,\ldots,q_T|O,\lambda)$$

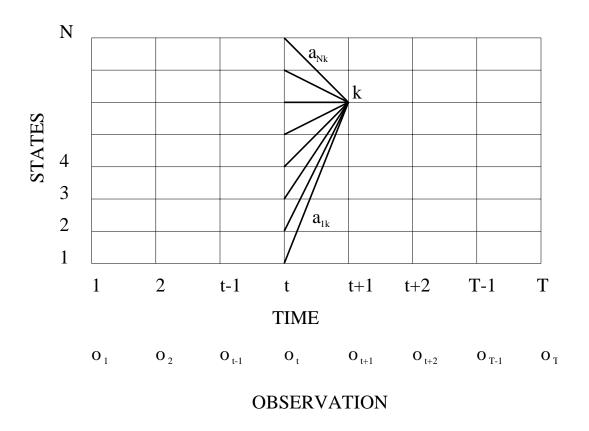
- Solution by Dynamic Programming
- Define:

$$\delta_t(i) = \max_{q_1, q_2, \dots, q_{t-1}} P(q_1, q_2, \dots, q_t = i, o_1, o_2, \dots, o_t | \lambda)$$

- ullet $\delta_t(i)$ is the highest prob. path ending in state i
- By induction we have:

$$\delta_{t+1}(j) = \max_{i} [\delta_t(i)a_{ij}] \cdot b_j(o_{t+1})$$

Viterbi Algorithm



Viterbi Algorithm

• Initialization:

$$\delta_1(i) = \pi_i b_i(o_1), \quad 1 \le i \le N$$

$$\psi_1(i) = 0$$

• Recursion:

$$\delta_t(j) = \max_{1 \le i \le N} [\delta_{t-1}(i)a_{ij}]b_j(o_t)$$

$$\psi_t(j) = \arg\max_{1 \le i \le N} [\delta_{t-1}(i)a_{ij}]$$

$$2 \le t \le T, 1 \le j \le N$$

• Termination:

$$P^* = \max_{1 \le i \le N} [\delta_T(i)]$$
 $q_T^* = \arg\max_{1 \le i \le N} [\delta_T(i)]$

• Path (state sequence) backtracking:

$$q_t^* = \psi_{t+1}(q_{t+1}^*), \quad t = T - 1, T - 2, \dots, 1$$

Solution to Problem 3

- \bullet Estimate $\lambda = (A,B,\pi)$ to maximize $P(O|\lambda)$
- No analytic method because of complexity iterative solution.
- Baum-Welch Algorithm:
 - 1. Let initial model be λ_0 .
 - 2. Compute new λ based on λ_0 and observation O.
 - 3. If $\log P(O|\lambda) \log P(O|\lambda_0) < DELTA$ stop.
 - **4.** Else set $\lambda_0 \leftarrow \lambda$ and goto step **2.**

Baum-Welch: Preliminaries

• Define $\xi(i,j)$ as the probability of being in state i at time t and in state j at time t+1.

$$egin{array}{lll} \xi(i,j) &= rac{lpha_t(i)a_{ij}b_j(o_{t+1})eta_{t+1}(j)}{P(O|\lambda)} \ &= rac{lpha_t(i)a_{ij}b_j(o_{t+1})eta_{t+1}(j)}{\sum_{i=1}^N\sum_{j=1}^Nlpha_t(i)a_{ij}b_j(o_{t+1})eta_{t+1}(j)} \end{array}$$

• Define $\gamma_t(i)$ as probability of being in state i at time t, given the observation sequence.

$$\gamma_t(i) = \sum\limits_{i=1}^N \xi_t(i,j)$$

- $\sum_{t=1}^{T} \gamma_t(i)$ is the expected number of times state i is visited.
- $\sum_{t=1}^{T-1} \xi_t(i,j)$ is the expected number of transitions from state i to state j.

Baum-Welch: Update Rules

- $\bar{\pi}_i$ = expected frequency in state i at time (t = 1) = $\gamma_1(i)$.
- $\bar{a}_{ij} = (\text{expected number of transition from state } i$ to state j)/ (expected nubmer of transitions from state i):

$$ar{a}_{ij} = rac{\sum oldsymbol{\xi}_t(i,j)}{\sum \gamma_t(i)}$$

• $\bar{b}_j(k) =$ (expected number of times in state j and observing symbol k) / (expected number of times in state j:

$$ar{b}_j(k) = rac{\sum_{t,o_t=k} \gamma_t(j)}{\sum_t \gamma(j)}$$

Properties

- ullet Covariance of the estimated parameters
- Convergence rates

Types of HMM

- Continuous density
- Ergodic
- State duration

Implementation Issues

- Scaling
- Initial parameters
- ullet Multiple observation

Comparison of HMMs

- What is a natural distance function?
- If $\rho(\lambda_1, \lambda_2)$ is large, does it mean that the models are really different?