A brief introduction to Bayesian Statistics through Astronomical Applications (Lecture 3)

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More Examples

Coordinate Transformations

- Lets say we have a problem in which we know how to write our posterior P(x) in a variable x, but in reality we are interested in knowing the posterior P(y) for variable y=f(x)
- We know the integral of the probability must be conserved

$$\int P(x)dx = \int P(y)dy \qquad P(x)dx = P(y)dy$$

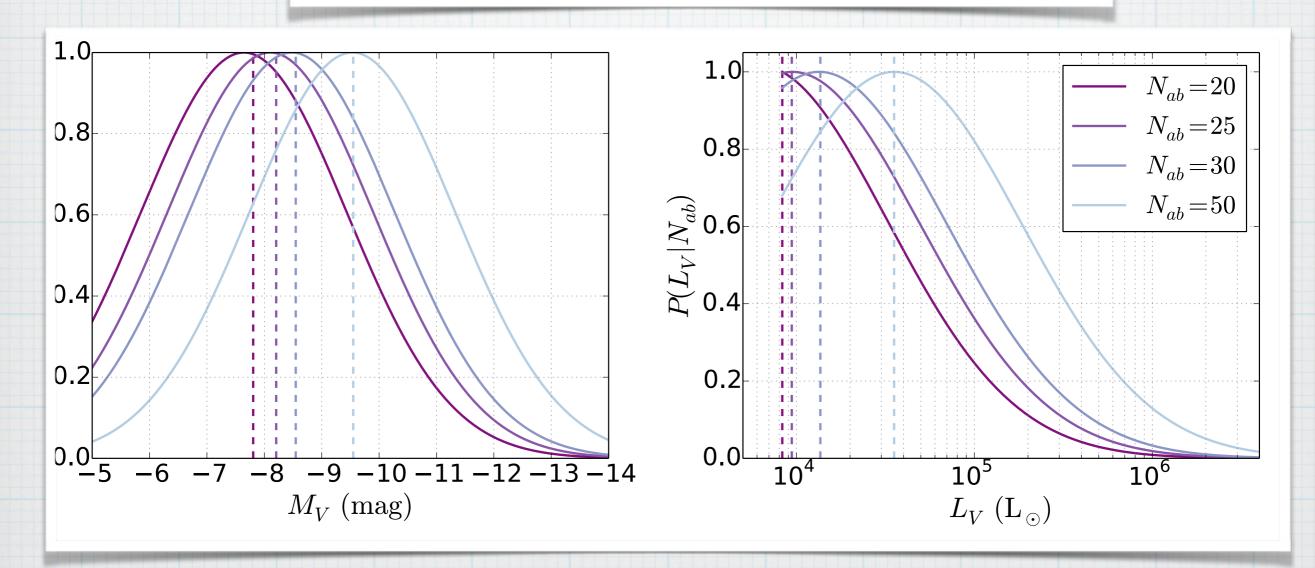
◆ This makes it easy, we know how to change variables inside an integral

$$P(y) = \left| \frac{dx}{dy} \right| P(x)$$

Jacobian of the transformation f(x)

The RR Lyrae problem from L1

 $P(L_V|\log N_{ab}) = P(M_V(L_V)|\log N_{ab})/L_V$



Propagation

Lets say we know the PDF for two variables X and Y

$$P(X)$$
 $p(Y)$ $|data, I|$

◆ Now we'd like to know what is the PDF for Z, where

$$Z = X + Y$$

First, lets look at the case when X and Y are conditionally independent. If this is the case, the joint probability P(X,Y)=P(X)p(Y), from the Marginalization rule this is:

$$P(Z) = \iint P(X)p(Y)dXdY \qquad |data, I|$$
 for all X,Y t.q. Z=X+y

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$$P(Z) = \iint P(X)p(Y)\delta(Z - [X + Y])dXdY$$

$$P(Z) = \int P(X)p(Y = Z - X)dX$$

note this is
the
convolution of
P(X) and p(Y)

◆ If P(X) and p(y) are gaussians, e.g. lets say like yesterday we have a model with gaussian uncertainties and a uniform prior,

$$P(X) = \frac{1}{\sqrt{2\pi}\sigma_X} e^{-\frac{(X - X_o)^2}{2\sigma_X^2}}$$

$$P(Y) = \frac{1}{\sqrt{2\pi}\sigma_Y} e^{-\frac{(Y-Y_o)^2}{2\sigma_Y^2}}$$

then

$$P(Z) = \int P(X)p(Y = Z - X)dX$$

|data, I|

is the convolution of these two gaussians

$$P(Z) = \frac{1}{2\pi\sigma_X\sigma_Y} \int_{-\infty}^{+\infty} e^{-\frac{(X - X_o)^2}{2\sigma_X^2}} e^{-\frac{(Z - X - Y_o)^2}{2\sigma_Y^2}} dX$$

After completing squares and simplifying we get

$$P(Z) = \frac{1}{2\pi\sigma_z} e^{-\frac{(Z-Z_o)^2}{2\sigma_z^2}}$$

where

$$Zo = X_o + Y_o$$
 and

$$\sigma_Z^2 = \sigma_X^2 + \sigma_Y^2$$

so, the sum in quadrature rule is derived

◆ Lets go back to the general problem, where Z=f(X,Y)

◆ In general, i.e. without assuming cond. independence of X,Y we have

$$P(Z) = \iint P(X, Y) dX dY$$
for all X,Y
t.q.
$$Z = f(X,Y)$$

◆ Since Z=f(X,Y) we have

$$P(Z) = \iint P(X, Y)\delta(Z - f(X, Y))dXdY$$

→ ... what follows will depend on the exact form of P(X,Y)...

More Examples

Fitting a density profile to data

- Another example. Lets say you have a catalogue of a particular stellar tracer, e.g. RR Lyrae stars, Red Clump stars, Cepheids or whatever.
- We'd like to use these tracers to study how stars are distributed in space in the different components in the Galaxy, i.e. their density profile ρ
- How do we compare and fit this function to a sample of stars?

Thick Disc and Halo RRL Density Profiles

◆ Thick Disc density profile

$$\rho_{tkd} = e^{-\frac{R - R_{\odot}}{h_R}} e^{-\frac{|z|}{h_z}}$$

Halo density profile

$$\rho_H = \frac{1}{R_{\odot}^n} \left[R^2 + z^2 \right]^{n/2}$$

◆ Total density profile:

$$\rho = N_{mod} \Big(\rho_{tkd} + C_{hd} \rho_H \Big) = N_{mod} \hat{\rho}$$

$$N_{RR} = \iiint_{V_S} \rho(\vec{r}) dV$$

... but the density ρ is not a directly observable quantity! all we have are the total number N_{OBS} of RRLs in the survey volume V_S and their observed positions

Density Profiles: A Bayesian approach

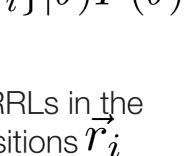
Our free parameters are:

$$\vec{\theta} = \{(h_z, h_R, n, \hat{C}_{hd}), N_{mod}\}$$

◆ Bayes's Theorem is

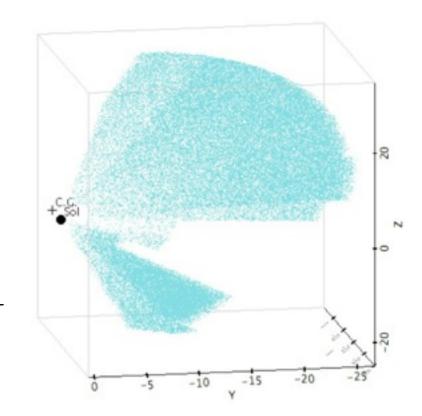
$$P(\vec{\theta}|N_{obs}, \{\vec{r}_i\}) = L(N_{obs}, \{\vec{r}_i\}|\vec{\theta})P(\vec{\theta})$$

lacktriangle Our data is the total number $oldsymbol{N_{obs}}$ of RRLs in the survey volume, and the individual positions $\overrightarrow{r_i}$





$$L(N_{obs}, \{\vec{r}_i\}|\vec{\theta}) = P(N_{obs}|\vec{\theta})P(\{\vec{r}_i\}|\vec{\theta})$$



Density Profiles: A Bayesian approach

Our free parameters are:

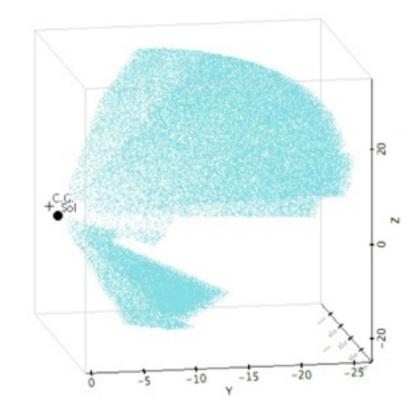
$$\vec{\theta} = \{(h_z, h_R, n, \hat{C}_{hd}), N_{mod}\}$$

- lacktriangle Our data is the total number $oldsymbol{N_{obs}}$ of RRLs in the survey volume, and the individual positions $\overrightarrow{r_i}$
- The Likelihood function can be expressed as

$$L(N_{obs}, \{\vec{r}_i\} | \vec{\theta}) = P(N_{obs} | \vec{\theta}) \prod_{i=1}^{N_{obs}} P(\vec{r}_i | \vec{\theta})$$

$$e^{-N_{mod}} (N_{mod})^{N_{obs}} \qquad \hat{\rho}(\vec{r}_i | \vec{\theta})$$

Poisson probability of observing N_{obs} given predicted number N_{mod}



Density profile (normalized to integrate 1 over the survey volume)

Density Profiles: A Bayesian approach

So we have

$$L(N_{obs}, \{\vec{r}_i\} | \vec{\theta}) = \left(e^{-N_{mod}}(N_{mod})^{N_{obs}}\right) \left(\prod_{i=1}^{N_{obs}} \hat{\rho}(\vec{r}_i | \vec{\theta})\right)$$

which we can write as a log-likelihood for a more convenient computation as:

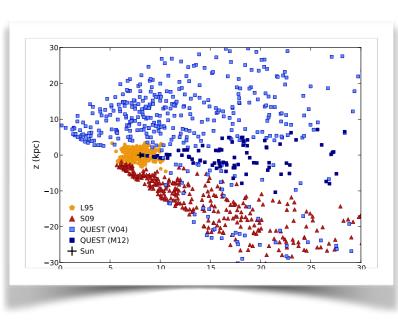
$$\ln L(N_{obs}, \{\vec{r}_i\} | \vec{\theta}) = \left(N_{obs} \ln N_{mod} - N_{mod}\right) + \sum_{i=1}^{N_{obs}} \ln \hat{\rho}(\vec{r}_i | \vec{\theta})$$

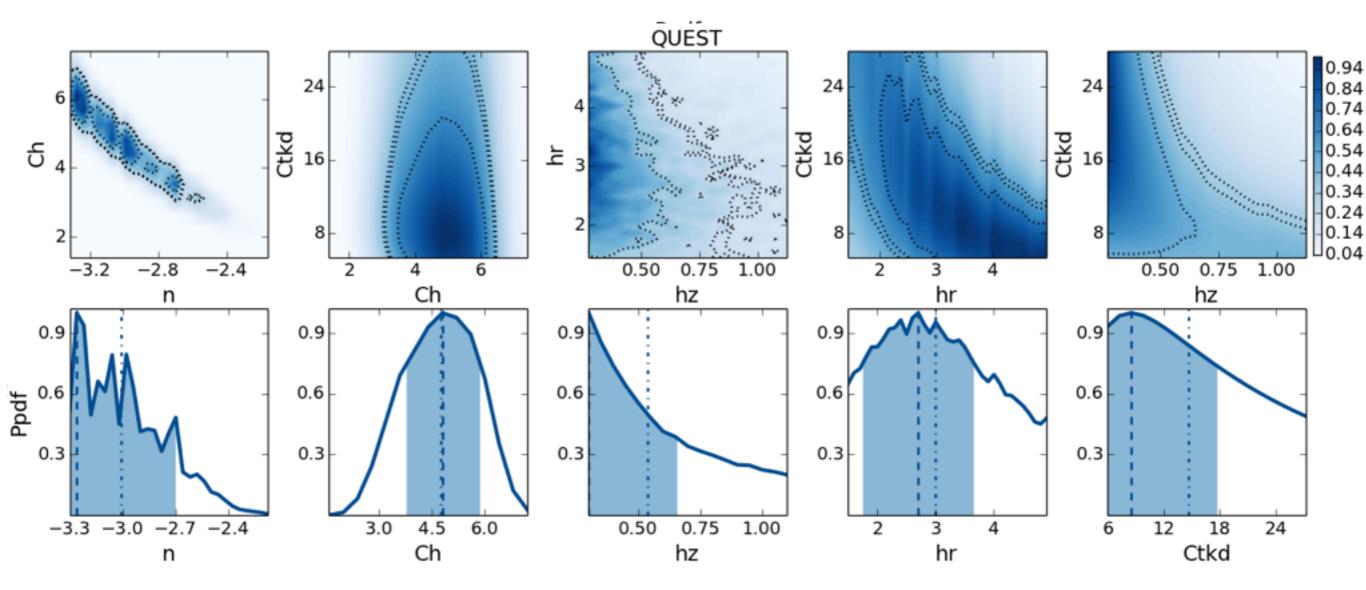
This is the computationally intensive part

$$N_{RR} = \iiint_{V_S} \rho(\vec{r}) dV$$

Density Profiles: One sample at a time

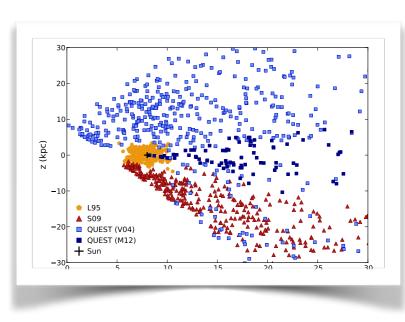
The marginal posteriors taking only the **QUEST** sample are as follows:

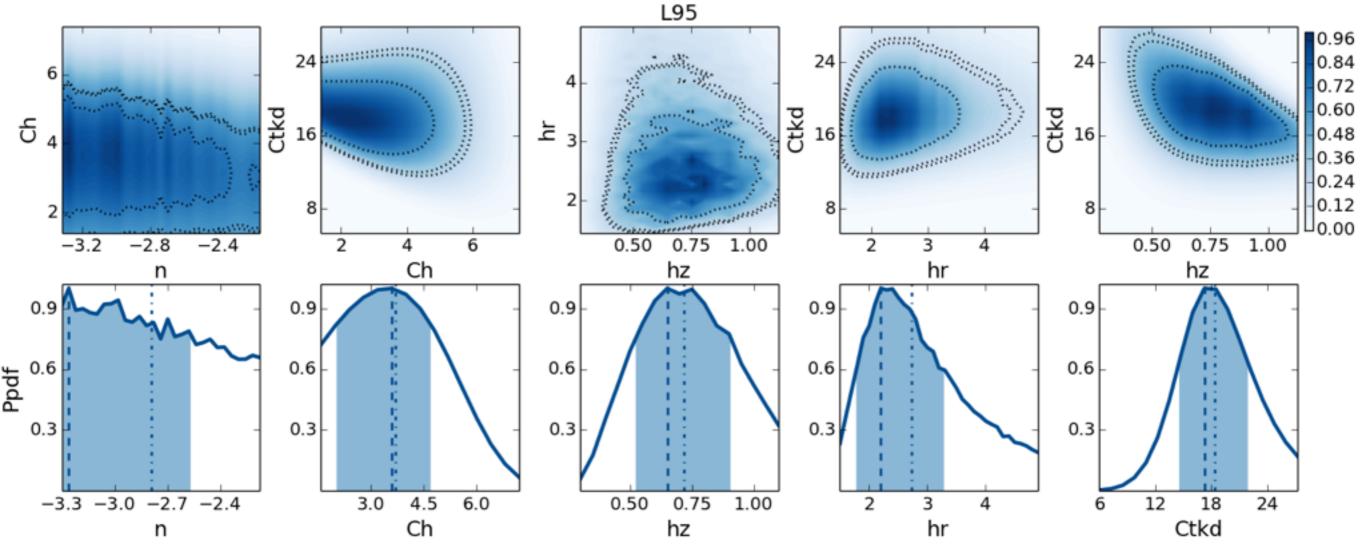




Density Profiles: One sample at a time

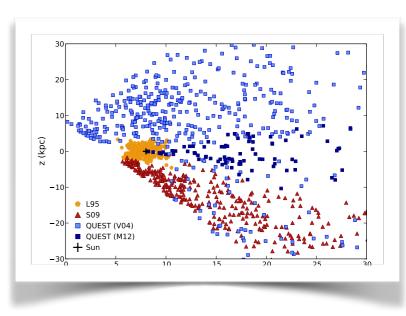
The marginal posteriors taking only the Layden sample are as follows:

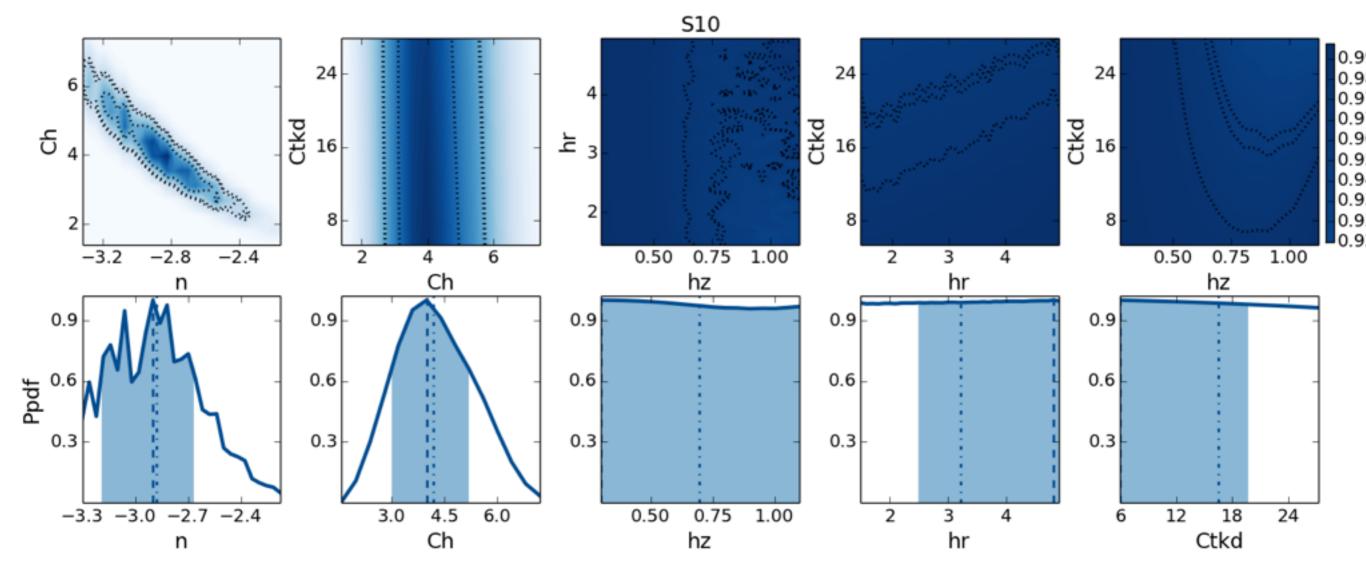




Density Profiles: One sample at a time

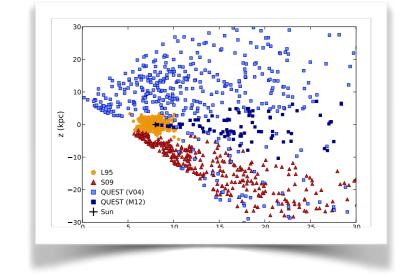
The marginal posteriors taking only the **Sesar** sample are as follows:

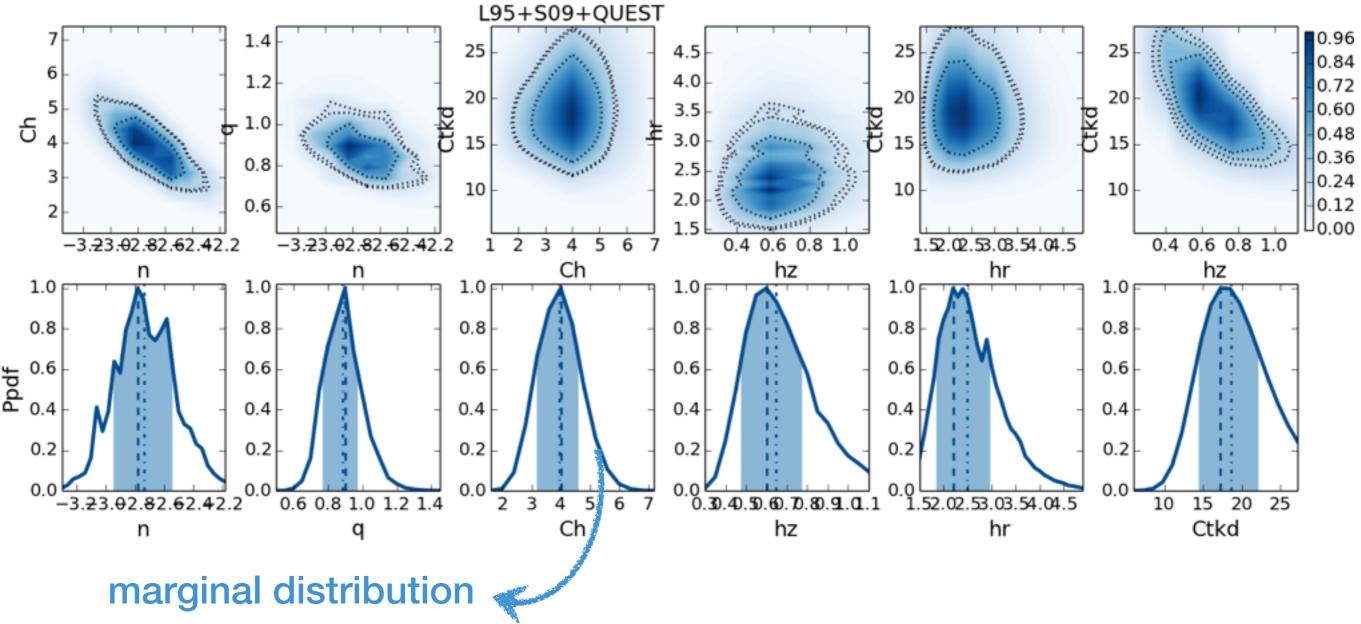




Density Profiles: Combined samples

Combining three different samples we find these marginal posteriors (remember its the product of the individual pdfs in 5-D space, and then marginalizing):





Density Profiles: Combined samples

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