

Analysis of Covariance

Cathy Maugis-Rabusseau

INSA Toulouse / IMT
GMM 116

cathy.maugis@insa-toulouse.fr

2021-2022

- 1 Introduction
- 2 Modelings
- 3 Parameter estimation
- 4 Testing procedures

1 Introduction

Context and notation

- ANCOVA= Analysis of covariance
- We want to explain a **quantitative** response variable Y using **qualitative** and **quantitative** variables together
- Here we only consider one covariate z and one factor T with l levels
- n_i = number of observations for the i -th level of T , $n = \sum_{i=1}^l n_i$.
- Y_{ij} = value of the response Y for $j = 1, \dots, n_i$, $i = 1, \dots, l$
- z_{ij} = value of the covariate z for $j = 1, \dots, n_i$, $i = 1, \dots, l$

Example

We want to find if temperature and oxygenation conditions influence the evolution of oyster weight. We have $n = 20$ bags of 10 oysters. We place, during a month, these 20 bags randomly in $I = 5$ different locations of a channel cooling of a power station at the rate of $n_i = 4$ bags per location. These locations are differentiated by their temperature and oxygenation.

For each bag, we have

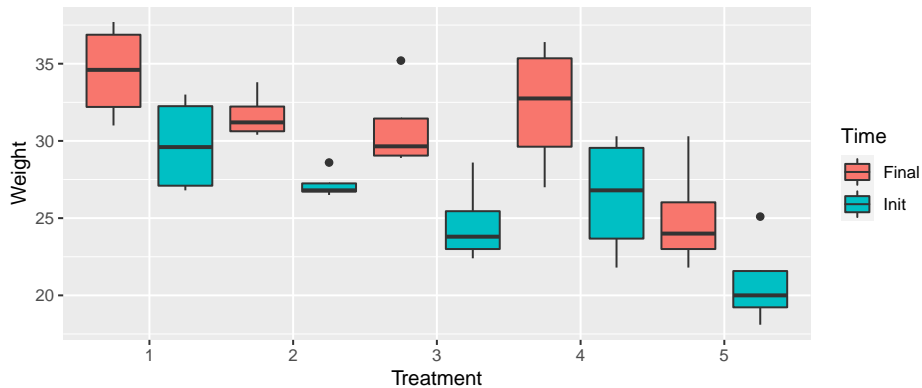
- its weight after the experiment (*Final weight*) = the response Y
- its weight before the experiment (*Init weight*) = the explanatory variable z
- the location (*Treatment* - 1 to 5) = the qualitative variable T

Example

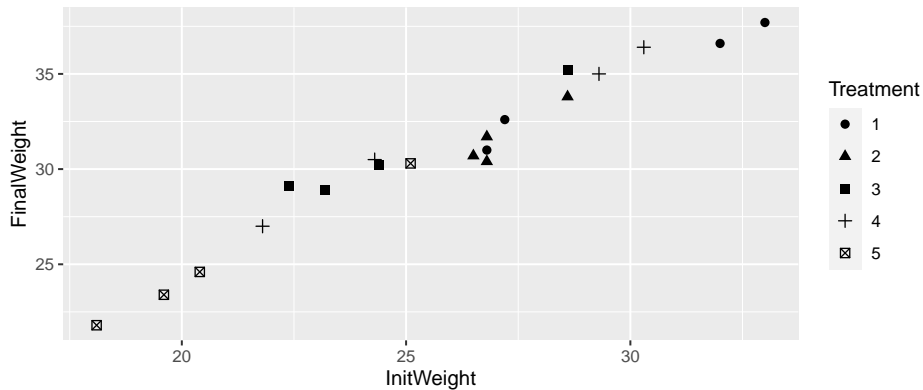
```
print(oyster)
```

	InitWeight	FinalWeight	Treatment
1	27.2	32.6	1
2	32.0	36.6	1
3	33.0	37.7	1
4	26.8	31.0	1
5	28.6	33.8	2
6	26.8	31.7	2
7	26.5	30.7	2
8	26.8	30.4	2
9	28.6	35.2	3
10	22.4	29.1	3
11	23.2	28.9	3
12	24.4	30.2	3
13	29.3	35.0	4
14	21.8	27.0	4
15	30.3	36.4	4
16	24.3	30.5	4
17	20.4	24.6	5
18	19.6	23.4	5
19	25.1	30.3	5
20	18.1	21.8	5

Example



Example



2 Modelings

Regular model

- Model:

$$(MR) : \begin{cases} Y_{ij} = a_i + b_i z_{ij} + \varepsilon_{ij}, & \forall i = 1, \dots, I, \forall j = 1, \dots, n_i \\ \varepsilon_{ij} \text{ i.i.d. } \sim \mathcal{N}(0, \sigma^2) \end{cases}$$

\Leftrightarrow Estimating a linear regression of Y on z for each level i of the factor T .

$$\underbrace{\begin{pmatrix} Y_{(1)} \\ \vdots \\ \vdots \\ Y_{(I)} \end{pmatrix}}_Y = \underbrace{\begin{pmatrix} X_{(1)} & & & \\ & X_{(2)} & & \\ & & \ddots & \\ & & & X_{(I)} \end{pmatrix}}_X \underbrace{\begin{pmatrix} a_1 \\ b_1 \\ \vdots \\ a_I \\ b_I \end{pmatrix}}_{\theta} + \underbrace{\begin{pmatrix} \varepsilon_{(1)} \\ \vdots \\ \vdots \\ \varepsilon_{(I)} \end{pmatrix}}_{\varepsilon}$$

with $Y_{(i)} = (Y_{i1}, \dots, Y_{in_i})'$, $X_{(i)} = (\mathbf{1}_{n_i}, z_{(i)})$.

$$(MS) : \begin{cases} Y_{ij} = (\mu + \alpha_i) + (\beta + \gamma_i)z_{ij} + \varepsilon_{ij}, & \forall i = 1, \dots, I, \forall j = 1, \dots, n_i. \\ \varepsilon_{ij} \text{ i.i.d. } \sim \mathcal{N}(0, \sigma^2) \end{cases}$$

- In this parametrization,
 - interaction effect between the covariate z and the factor T : γ_i
 - differential effect of the factor T on Y : α_i
 - differential effect of the covariate z on Y : β
- $2I + 2$ parameters \Rightarrow 2 constraints are required to model identifiability

3 Parameter estimation

Estimation in regular model (MR)

In a regular model, $\hat{\theta} = (X'X)^{-1}X'Y$.

Since $X = \text{diag}(X_{(1)}, \dots, X_{(I)})$, we have

$$(X'X)^{-1} = \text{diag}((X'_{(1)}X_{(1)})^{-1}, \dots, (X'_{(I)}X_{(I)})^{-1})$$

and

$$X'Y = \text{diag}(X'_{(1)}Y_{(1)}, \dots, X'_{(I)}Y_{(I)})$$

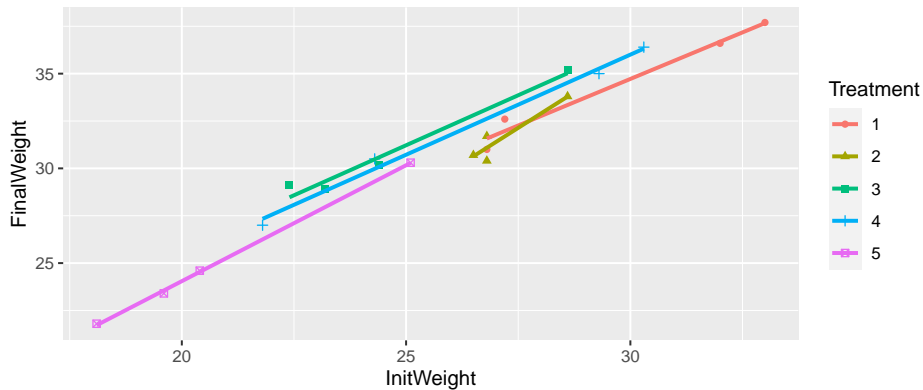
Thus

$$\hat{\theta} = \begin{pmatrix} (X'_{(1)}X_{(1)})^{-1}X'_{(1)}Y_{(1)} \\ \vdots \\ (X'_{(I)}X_{(I)})^{-1}X'_{(I)}Y_{(I)} \end{pmatrix}$$

Using results in simple linear regression, we deduce

$$\begin{cases} \hat{b}_i = \text{cov}(Y_{(i)}, z_{(i)})/\text{var}(z_{(i)}) \\ \hat{a}_i = \bar{Y}_{(i)} - \bar{z}_{(i)}\hat{b}_i \end{cases}$$

Example



Estimation in singular model (MS)

- Identifiability constraints: by default in R $\alpha_1 = \gamma_1 = 0$
- Using the link between the parameters in (MR) and (MS), we can easily deduce

$$\left\{ \begin{array}{l} \hat{\mu} = \hat{a}_1 \\ \hat{\alpha}_i = \hat{a}_i - \hat{a}_1 \\ \hat{\beta} = \hat{b}_1 \\ \hat{\gamma}_i = \hat{b}_i - \hat{b}_1 \end{array} \right.$$

```
complet<-lm(FinalWeight~InitWeight * Treatment,data=oyster)
summary(complet)
```

Call:

```
lm(formula = FinalWeight ~ InitWeight * Treatment, data = oyster)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.68699	-0.28193	0.02184	0.10425	0.63075

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	5.24126	2.86473	1.830	0.0972 .
InitWeight	0.98265	0.09588	10.249	1.27e-06 ***
Treatment2	-14.39058	9.15971	-1.571	0.1472
Treatment3	-0.42330	3.97747	-0.106	0.9174
Treatment4	-0.94550	3.50725	-0.270	0.7930
Treatment5	-5.67309	3.57150	-1.588	0.1433
InitWeight:Treatment2	0.51871	0.33406	1.553	0.1515
InitWeight:Treatment3	0.07342	0.14699	0.499	0.6282
InitWeight:Treatment4	0.07428	0.12229	0.607	0.5571
InitWeight:Treatment5	0.24124	0.13980	1.726	0.1151

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.5324 on 10 degrees of freedom

Multiple R-squared: 0.9921, Adjusted R-squared: 0.985

F-statistic: 139.5 on 9 and 10 DF, p-value: 2.572e-09

Example



```
import statsmodels.api as sm
from statsmodels.formula.api import ols
oysterpy=r.oyster;
completpy = ols('FinalWeight ~ InitWeight * Treatment', data=oysterpy).fit();
completpy.summary()
```

```
<class 'statsmodels.iolib.summary.Summary'>
"""
```

OLS Regression Results

Dep. Variable:	FinalWeight	R-squared:	0.992			
Model:	OLS	Adj. R-squared:	0.985			
Method:	Least Squares	F-statistic:	139.5			
Date:	Jeu, 14 oct 2021	Prob (F-statistic):	2.57e-09			
Time:	21:57:22	Log-Likelihood:	-8.8384			
No. Observations:	20	AIC:	37.68			
Df Residuals:	10	BIC:	47.63			
Df Model:	9					
Covariance Type:	nonrobust					
=====						
	coef	std err	t	P> t	[0.025	0.975]

Intercept	5.2413	2.865	1.830	0.097	-1.142	11.624
Treatment[T.2]	-14.3906	9.160	-1.571	0.147	-34.800	6.019
Treatment[T.3]	-0.4233	3.977	-0.106	0.917	-9.286	8.439
Treatment[T.4]	-0.9455	3.507	-0.270	0.793	-8.760	6.869
Treatment[T.5]	-5.6731	3.572	-1.588	0.143	-13.631	2.285
InitWeight	0.9826	0.096	10.249	0.000	0.769	1.196
InitWeight:Treatment[T.2]	0.5187	0.334	1.553	0.152	-0.226	1.263
InitWeight:Treatment[T.3]	0.0734	0.147	0.499	0.628	-0.254	0.401
InitWeight:Treatment[T.4]	0.0743	0.122	0.607	0.557	-0.198	0.347
InitWeight:Treatment[T.5]	0.2412	0.140	1.726	0.115	-0.070	0.553

4 Testing procedures

Absence of any effect

- We want to compare the "null model"

$$(M0) : Y_{ij} = \mu + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

against the full model (MS)

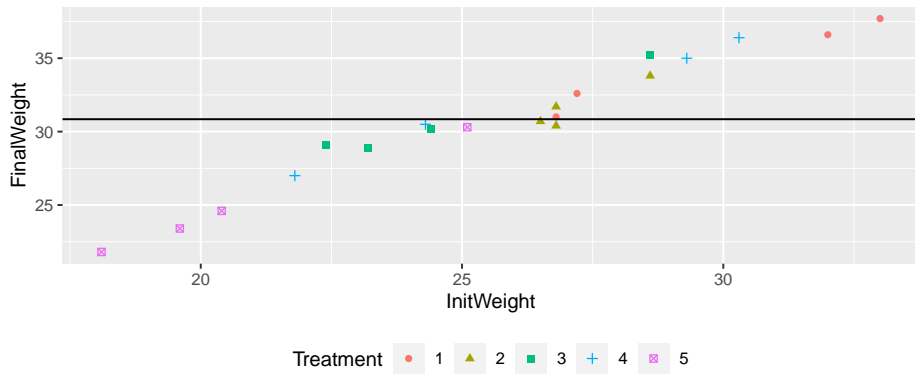
$$(MS) : Y_{ij} = (\mu + \alpha_i) + (\beta + \gamma_i)z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

- Fisher's test statistics:

$$F = \frac{SSE/(2I - 1)}{SSR/n - 2I} \underset{\mathcal{H}_0}{\sim} \mathcal{F}(2I - 1, n - 2I)$$

with $SSR = \|Y - \hat{Y}\|^2$ and $SSE = \|\hat{Y} - \bar{Y}\mathbf{1}_n\|^2$

Example





• With R:

```
M0<-lm(FinalWeight~1,data=oyster)
anova(M0,complet)
```

Analysis of Variance Table

Model 1: FinalWeight ~ 1

Model 2: FinalWeight ~ InitWeight * Treatment

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	19	358.67				
2	10	2.83	9	355.84	139.51	2.572e-09 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

• With Python:

```
from statsmodels.stats.anova import anova_lm
M0py = ols('FinalWeight~1', data=oysterpy).fit()
anova_lm(M0py,completpy)
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	19.0	358.669500	0.0	NaN	NaN	NaN
1	10.0	2.834009	9.0	355.835491	139.510053	2.572066e-09

Test of non-interaction between factor and covariate

- We want to test the null hypothesis:

$$\mathcal{H}_0^{(SI)} : b_1 = b_2 = \cdots = b_I \iff \gamma_1 = \gamma_2 = \cdots = \gamma_I = 0$$

- Fisher's test to compare

- the full model

$$(MS) : Y_{ij} = (\mu + \alpha_i) + (\beta + \gamma_i)z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

$$(MR) : Y_{ij} = a_i + b_i z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

- the sub-model with non-interaction

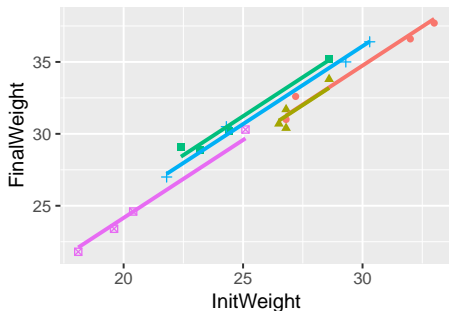
$$(MS_{nonI}) : Y_{ij} = \mu + \alpha_i + \beta z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

$$(MR_{nonI}) : Y_{ij} = a_i + b z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

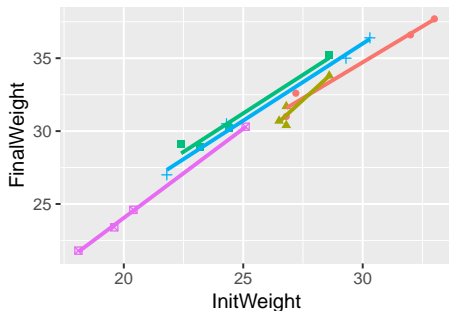
- Test statistics:

$$F = \frac{SSR_{nonI} - SSR / (I - 1)}{SSR / (n - 2I)} \underset{\mathcal{H}_0}{\sim} \mathcal{F}(I - 1, n - 2I)$$

Test of non-interaction between factor and covariate



Treatment 1 2 3 4 5



Treatment 1 2 3 4 5

Model with non-interaction



```
nonI<-lm(FinalWeight~InitWeight+Treatment)
summary(nonI)
```

Call:

```
lm(formula = FinalWeight ~ InitWeight + Treatment)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.8438	-0.3154	-0.2171	0.4863	0.8871

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.25040	1.44308	1.559	0.141205
InitWeight	1.08318	0.04762	22.746	1.87e-12 ***
Treatment2	-0.03581	0.40723	-0.088	0.931169
Treatment3	1.89922	0.45802	4.147	0.000988 ***
Treatment4	1.35157	0.41937	3.223	0.006135 **
Treatment5	0.24446	0.57658	0.424	0.678022

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.5492 on 14 degrees of freedom

Multiple R-squared: 0.9882, Adjusted R-squared: 0.984

F-statistic: 235 on 5 and 14 DF, p-value: 5.493e-13



• With R:

```
anova(nonI,complet)
```

Analysis of Variance Table

Model 1: FinalWeight ~ InitWeight + Treatment

Model 2: FinalWeight ~ InitWeight * Treatment

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	14	4.2223				
2	10	2.8340	4	1.3883	1.2247	0.3602

• With Python:

```
nonIpy = ols('FinalWeight ~ InitWeight + Treatment', data=oysterpy).fit()
from statsmodels.stats.anova import anova_lm
anova_lm(nonIpy,completpy)
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	14.0	4.222323	0.0	NaN	NaN	NaN
1	10.0	2.834009	4.0	1.388314	1.224691	0.360175

ANCOVA with non-interaction

- If the model with non-interaction between the factor and the covariate is retained

- Singular model:

$$(MS_{nonI}) : Y_{ij} = \mu + \alpha_i + \beta z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

- Regular model:

$$(MR_{nonI}) : Y_{ij} = a_i + b z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

- We may test the effect of the factor or the effect of the covariate on the response.

Effect of the covariate z on Y

- Fisher's test to compare
 - the model with non-interaction

$$(MS_{nonI}) : Y_{ij} = \mu + \alpha_i + \beta z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

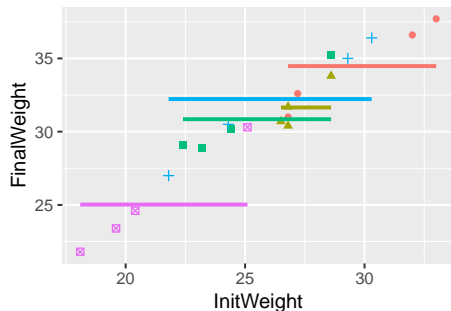
- the one-way ANOVA

$$(MT) : Y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i.$$

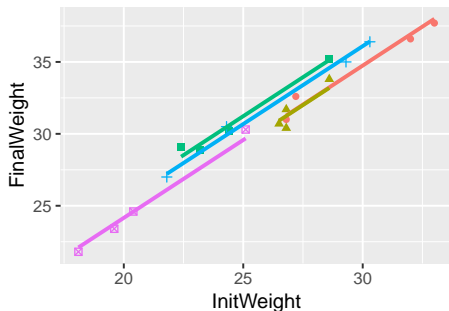
- Test statistics:

$$F = \frac{SSR_T - SSR_{nonI}/1}{SSR_{nonI}/(n - (I + 1))} \underset{\mathcal{H}_0}{\sim} \mathcal{F}(1, n - (I + 1))$$

Effect of the covariate z on Y



Treatment 1 2 3 4 5



Treatment 1 2 3 4 5

Example



```
MT<-lm(FinalWeight~Treatment)
anova(MT,nonI)
```

Analysis of Variance Table

Model 1: FinalWeight ~ Treatment

Model 2: FinalWeight ~ InitWeight + Treatment

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	15	160.263				
2	14	4.222	1	156.04	517.38	1.867e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
MTpy = ols('FinalWeight ~ Treatment', data=oysterpy).fit()
anova_lm(MTpy,nonIpy)
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	15.0	160.262500	0.0	NaN	NaN	NaN
1	14.0	4.222323	1.0	156.040177	517.383995	1.867369e-12

Effect of the factor T on Y

- Fisher's test to compare
 - the model with non-interaction

$$(MS_{nonI}) : Y_{ij} = \mu + \alpha_i + \beta z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

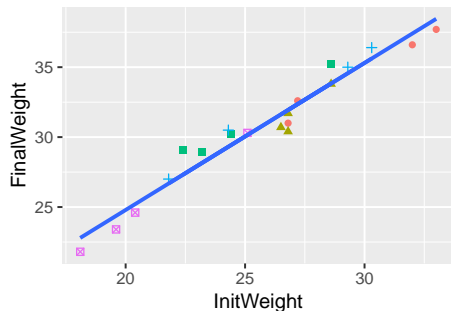
- the linear regression

$$(Mz) : Y_{ij} = \mu + \beta z_{ij} + \varepsilon_{ij}, \forall i = 1, \dots, I, \forall j = 1, \dots, n_i$$

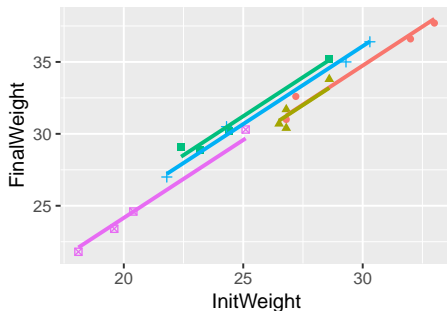
- Test statistics:

$$F = \frac{SSR_z - SSR_{nonI}/(I - 1)}{SSR_{nonI}/(n - (I + 1))} \underset{\mathcal{H}_0}{\sim} \mathcal{F}(I - 1, n - (I + 1))$$

Effect of the factor T on Y



Treatment 1 2 3 4 5



Treatment 1 2 3 4 5



• With R:

```
Mz<-lm(FinalWeight~InitWeight)
anova(Mz,nonI)
```

Analysis of Variance Table

Model 1: FinalWeight ~ InitWeight

Model 2: FinalWeight ~ InitWeight + Treatment

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	18	16.3117				
2	14	4.2223	4	12.089	10.021	0.0004819 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

• With Python:

```
Mzpy = ols('FinalWeight ~ InitWeight', data=oysterpy).fit()
anova_lm(Mzpy,nonIpy)
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	18.0	16.311683	0.0	NaN	NaN	NaN
1	14.0	4.222323	4.0	12.089359	10.021203	0.000482

Summary

- Know how to write an ANCOVA model (individually and matricially), regular and singular
- Know how to distinguish a regular model from a singular model
- Know how to estimate the parameters of the ANCOVA model in the regular case and in the singular case (by adapting to the chosen constraint(s))
- Know how to construct a confidence interval for a parameter of the ANCOVA model
- Know how to construct a test to test the effect of the factor, the interaction effect, ... and know how to organize these tests
- Know how to associate a graphic representation with a sub-model of ANCOVA

