Structural Equation Modeling with R and lavaan

Day 3: Multilevel Structural Equation Models

Constantin Manuel Bosancianu

WZB Berlin Social Science Center *Institutions and Political Inequality* manuel.bosancianu@gmail.com

September 22, 2021

Today we join some of the insights from the January workshop on MLM with what we covered in the past 2 days:

✓ the need for multilevel specifications

Today we join some of the insights from the January workshop on MLM with what we covered in the past 2 days:

- ✓ the need for multilevel specifications
- ✓ the logic of MSEM

Today we join some of the insights from the January workshop on MLM with what we covered in the past 2 days:

- ✓ the need for multilevel specifications
- ✓ the logic of MSEM
- multilevel measurement models
- multilevel path models

Today we join some of the insights from the January workshop on MLM with what we covered in the past 2 days:

- the need for multilevel specifications
- ✓ the logic of MSEM
- multilevel measurement models
- multilevel path models

From the software perspective, **lavaan** is still adding some more advanced capabilities, but the project is making constant advances in this.

The Multilevel Perspective

Value of MI M

MLMs are uniquely suited to capturing one type of social complexity: the way individuals/firms/NGOs act or think may be context-dependent.

Value of MLM

MLMs are uniquely suited to capturing one type of social complexity: the way individuals/firms/NGOs act or think may be context-dependent.

An example which I focused on in January (and continue to do so in my research) are the cross-country differences in the likelihood that lower-income people participate in politics.

Value of MLM

MLMs are uniquely suited to capturing one type of social complexity: the way individuals/firms/NGOs act or think may be context-dependent.

An example which I focused on in January (and continue to do so in my research) are the cross-country differences in the likelihood that lower-income people participate in politics.

Many similar examples related to educational research, e.g. differences between schools in how much progress students make over a 4-year cycle.

Cross-national variance

	Micro	Macro
Political participation	income	party polarization
	efficacy	welfare state institutions
Trust	education	post-communist country
Religiosity	age	income inequality

Cross-national variance

	Micro	Macro
Political participation	income	party polarization
	efficacy	welfare state institutions
Trust	education	post-communist country
Religiosity	age	income inequality

Trying to see the world like this trains your mind: how individual actions shape context, and how context, in turn, shapes individual action.

✓ obtain accurate SEs for estimates in instances of clustered data;

- ✓ obtain accurate SEs for estimates in instances of clustered data;
- ✓ model the heteroskedasticity (unobserved heterogeneity) in the data.

- ✓ obtain accurate SEs for estimates in instances of clustered data;
- ✓ model the heteroskedasticity (unobserved heterogeneity) in the data.

Because we're typically dealing with complex sample designs, formulas for standard errors are imprecise.

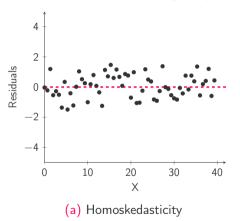
- ✓ obtain accurate SEs for estimates in instances of clustered data;
- ✓ model the heteroskedasticity (unobserved heterogeneity) in the data.

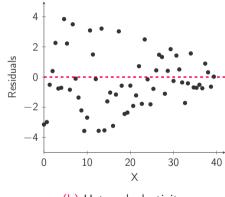
Because we're typically dealing with complex sample designs, formulas for standard errors are imprecise.

As SEs are incorporated into significance tests, we risk rejecting the null hypothesis more often than we should.

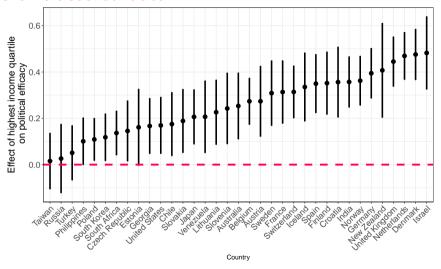
One OLS assumption

Homoskedasticity: $\epsilon \sim \mathcal{N}(0, \sigma^2)$.





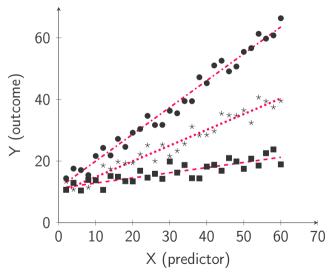
The case of clustered data



Variation in the effect of income on political efficacy (ISSP Citizenship II, 2016)

8 / 31

Consequences of heterogeneity



In this instance, applying an overall slope ("naive pooling") to the data will generate heteroskedasticity.

This can be addressed with country dummies (fixed effects), but these won't explain why you're seeing a specific pattern.

MLM & SEM: substantive reasons

Systematically account for how outcomes (or *effects*) vary across groups, beyond what can be explained by unit-level factors.

MLM & SEM: substantive reasons

Systematically account for how outcomes (or *effects*) vary across groups, beyond what can be explained by unit-level factors.

These outcomes can be either observed variables or, even latent variables.

MLM & SEM: substantive reasons

Systematically account for how outcomes (or *effects*) vary across groups, beyond what can be explained by unit-level factors.

These outcomes can be either observed variables or, even latent variables.

Tremendously helpful if we want to understand why and a measurement structure varies over (many) groups, or to explain variation in a structural component over the same groups.

Multilevel Path Models

The defining feature of the MLM framework was its ability to allow parameters to vary across groups: $\beta_1 \to \beta_{1j}$.

The defining feature of the MLM framework was its ability to allow parameters to vary across groups: $\beta_1 \to \beta_{1j}$.

These varying parameters got a statistical model of their own, with upper-level variables: $\beta_{1j} = \gamma_{10} + \gamma_{11} * Z_{1j} + \dots$

The defining feature of the MLM framework was its ability to allow parameters to vary across groups: $\beta_1 \to \beta_{1j}$.

These varying parameters got a statistical model of their own, with upper-level variables: $\beta_{1j} = \gamma_{10} + \gamma_{11} * Z_{1j} + \dots$

The same approach is used when moving from the SEM to the MSEM framework: allowing parameters, e.g. factor loadings, path coefficients, intercepts, to vary across groups.

The defining feature of the MLM framework was its ability to allow parameters to vary across groups: $\beta_1 \to \beta_{1j}$.

These varying parameters got a statistical model of their own, with upper-level variables: $\beta_{1j} = \gamma_{10} + \gamma_{11} * Z_{1j} + \dots$

The same approach is used when moving from the SEM to the MSEM framework: allowing parameters, e.g. factor loadings, path coefficients, intercepts, to vary across groups.

The notation becomes cumbersome very quickly (there can now be multiple β_1 s), so I focus on graphical representations.

Estimated covariances can be either fixed () or varying () (different than free or constrained).

Estimated covariances can be either fixed () or varying () (different than free or constrained).

We haven't discussed yet about modeling a mean structure in a SEM, but the same can be done for intercepts



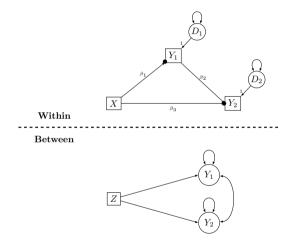
Estimated covariances can be either fixed () or varying () (different than free or constrained).

We haven't discussed yet about modeling a mean structure in a SEM, but the same can be done for intercepts



For modeling, all varying parameters at a lower level are considered latents at the higher level.

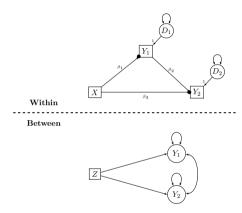
Multilevel path specification



Level-1 intercepts are allowed to vary across level-2 groups.

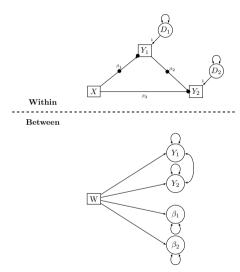
14 / 31

Varying intercepts



Both intercepts (for Y_1 and Y_2) are assumed to follow a Gaussian distribution, with the means explained by Z_i (group-level factor).

Varying intercepts and slopes



Similar to standard SEM, makes heavy use of ML and variations of it that are more robust to non-normality (robust ML, and the Bollen–Stine bootstrap).

17 / 31

Similar to standard SEM, makes heavy use of ML and variations of it that are more robust to non-normality (robust ML, and the Bollen–Stine bootstrap).

Model fit assessment is conducted using same indicators (model χ^2 , RMSEA, CFI, SRMR) as presented yesterday.

Similar to standard SEM, makes heavy use of ML and variations of it that are more robust to non-normality (robust ML, and the Bollen–Stine bootstrap).

Model fit assessment is conducted using same indicators (model χ^2 , RMSEA, CFI, SRMR) as presented yesterday.

The first 3 provide a single number assessment; for multilevel models, that means sometimes the level-1 sample size dominates.

Similar to standard SEM, makes heavy use of ML and variations of it that are more robust to non-normality (robust ML, and the Bollen–Stine bootstrap).

Model fit assessment is conducted using same indicators (model χ^2 , RMSEA, CFI, SRMR) as presented yesterday.

The first 3 provide a single number assessment; for multilevel models, that means sometimes the level-1 sample size dominates.

The SRMR is the only index that provides a model fit summary for the within-level and the between-level.

Partially-saturated model test

A specific approach to MSEM: Ryu and West's (2009) partially-saturated model test; unfortunately, currently not implemented as a standard command.

18 / 31

Partially-saturated model test

A specific approach to MSEM: Ryu and West's (2009) partially-saturated model test; unfortunately, currently not implemented as a standard command.

Logic:

- \checkmark model is specified with a saturated "between" part \Rightarrow misfit must come from "within" part
- \checkmark model is then specified with a saturated "within" part \Rightarrow misfit must come from "between" part

Mediation in MSEM

Mediation poses problem in a standard multilevel setting, as it's hard to isolate cross-level dynamics $(2 \to 1 \to 1)$: we can't disentangle within-group and between-group effects (Zhang, Zyphur, & Preacher, 2009).

Mediation in MSEM

Mediation poses problem in a standard multilevel setting, as it's hard to isolate cross-level dynamics $(2 \to 1 \to 1)$: we can't disentangle within-group and between-group effects (Zhang et al., 2009).

MSEM allows for tests of quite diverse linkages: $1 \to 1 \to 2$, $2 \to 1 \to 2$, or $1 \to 2 \to 2$ (level-2 constructs can be outcomes in SEM).

Multilevel Factor Models

Multiple-group CFA (MG-CFA)

A precursor to multilevel CFA: checking for measurement *invariance* by checking factor loadings across multiple groups, e.g. women and men, East and West Germany.

Multiple-group CFA (MG-CFA)

A precursor to multilevel CFA: checking for measurement *invariance* by checking factor loadings across multiple groups, e.g. women and men, East and West Germany.

Invariance: individuals from two different populations with the same level on the latent construct have the same scores on the measured indicators.

Multiple-group CFA (MG-CFA)

A precursor to multilevel CFA: checking for measurement *invariance* by checking factor loadings across multiple groups, e.g. women and men, East and West Germany.

Invariance: individuals from two different populations with the same level on the latent construct have the same scores on the measured indicators.

This property is important if we want to apply the same measurement instrument (political efficacy, political trust, populist attitudes) across contexts.

Without it, cross-context differences might be just due to measurement error.

Multilevel CFA

Multilevel CFA aims at the same insights as MG-CFA, but has the toolbox to pursue these questions across many more groups:

✓ it can test whether the within-group structure of measurement matches the between-group one

22 / 31

Multilevel CFA

Multilevel CFA aims at the same insights as MG-CFA, but has the toolbox to pursue these questions across many more groups:

- it can test whether the within-group structure of measurement matches the between-group one
- ✓ level-2 covariates can be used to explain variance in level-2 constructs (becoming a full structural regression model)

Multilevel CFA

Multilevel CFA aims at the same insights as MG-CFA, but has the toolbox to pursue these questions across many more groups:

- it can test whether the within-group structure of measurement matches the between-group one
- ✓ level-2 covariates can be used to explain variance in level-2 constructs (becoming a full structural regression model)
- allows for testing of measurement invariance across many groups at the same time

2-level CFA: random intercepts (I)

The variance of an observed indicator is split into 2 components, within- and between-, which are additive and orthogonal:

$$Var(Y) = Var(Y)_B + Var(Y)_W \tag{1}$$

2-level CFA: random intercepts (I)

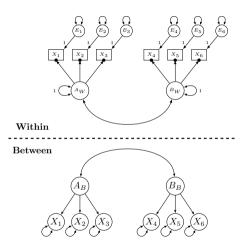
The variance of an observed indicator is split into 2 components, within- and between-, which are additive and orthogonal:

$$Var(Y) = Var(Y)_B + Var(Y)_W \tag{1}$$

The "within" variance is explained by the level-1 latent factors. The "between" part is explained with the level-2 latent factors.

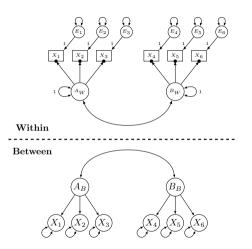
These two sets of specifications together comprise the 2-level measurement structure.

2-level CFA: random intercepts (II)



At the "between" part we're explaining the intercepts from the "within" part using 2 factors.

2-level CFA: random intercepts (II)



At the "between" part we're explaining the intercepts from the "within" part using 2 factors.

We have to assume that the $g_1, g_2, \dots g_j$ covariance matrices for the groups are identical \Rightarrow measurement invariance.

As an initial stage, just as in a standard MLM, we would check that there is sufficient between-group variance in the indicators.

As an initial stage, just as in a standard MLM, we would check that there is sufficient between-group variance in the indicators.

The standard tool for this is the ICC (intra-class correlation coefficient):

$$\rho = \frac{\sigma_{\alpha}^2}{\sigma_{\alpha}^2 + \sigma_{y}^2} \tag{2}$$

As an initial stage, just as in a standard MLM, we would check that there is sufficient between-group variance in the indicators.

The standard tool for this is the ICC (intra-class correlation coefficient):

$$\rho = \frac{\sigma_{\alpha}^2}{\sigma_{\alpha}^2 + \sigma_{y}^2} \tag{2}$$

where $\sigma_{total}^2 = \sigma_{\alpha}^2 + \sigma_{y}^2$

As an initial stage, just as in a standard MLM, we would check that there is sufficient between-group variance in the indicators.

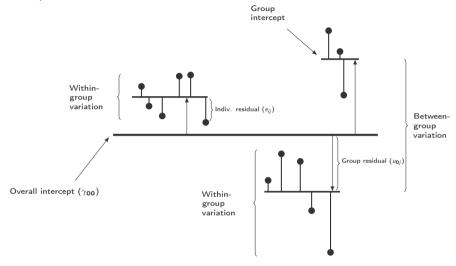
The standard tool for this is the ICC (intra-class correlation coefficient):

$$\rho = \frac{\sigma_{\alpha}^2}{\sigma_{\alpha}^2 + \sigma_{y}^2} \tag{2}$$

where $\sigma_{total}^2 = \sigma_{\alpha}^2 + \sigma_{y}^2$

With insufficient variance, there is not much for group-level factors to explain.

ICC decomposition



Adapted from Merlo et al. (2005).

The rule presented in the beginning is that there need to be fewer free parameter than the number of observations: $\frac{p(p+1)}{2}$.

The rule presented in the beginning is that there need to be fewer free parameter than the number of observations: $\frac{p(p+1)}{2}$.

These are also the off-diagonal elements in the variance-covariance matrix of observed indicators.

The rule presented in the beginning is that there need to be fewer free parameter than the number of observations: $\frac{p(p+1)}{2}$.

These are also the off-diagonal elements in the variance-covariance matrix of observed indicators.

In a multilevel setting, we have 2 such matrices: for the "within" and the "between" components.

The rule presented in the beginning is that there need to be fewer free parameter than the number of observations: $\frac{p(p+1)}{2}$.

These are also the off-diagonal elements in the variance-covariance matrix of observed indicators.

In a multilevel setting, we have 2 such matrices: for the "within" and the "between" components.

Therefore, the maximum number of estimable parameters is p(p+1) + k, where k is the number of indicator intercepts.

The assumption of measurement invariance is limiting. What if specific groups have a different measurement structure than others, e.g. employees from certain companies have different patterns in their scores?

The assumption of measurement invariance is limiting. What if specific groups have a different measurement structure than others, e.g. employees from certain companies have different patterns in their scores?

To check for this, we would allow both intercepts and factor loadings to vary across groups.

The assumption of measurement invariance is limiting. What if specific groups have a different measurement structure than others, e.g. employees from certain companies have different patterns in their scores?

To check for this, we would allow both intercepts and factor loadings to vary across groups.

The implication is that we now consider the $g_1, g_2, \dots g_j$ group covariance matrices to be different.

The assumption of measurement invariance is limiting. What if specific groups have a different measurement structure than others, e.g. employees from certain companies have different patterns in their scores?

To check for this, we would allow both intercepts and factor loadings to vary across groups.

The implication is that we now consider the $g_1, g_2, \dots g_j$ group covariance matrices to be different.

As before, we can add group-level indicators to explain the cross-group variance in loadings.

Final point on software

The capabilities of lavaan are still limited here, though growing fast.

Currently, lavaan can only run multilevel CFA with random intercepts.

Final point on software

The capabilities of lavaan are still limited here, though growing fast.

Currently, lavaan can only run multilevel CFA with random intercepts.

Furthermore, only with continuous indicators!

Final point on software

The capabilities of lavaan are still limited here, though growing fast.

Currently, lavaan can only run multilevel CFA with random intercepts.

Furthermore, only with continuous indicators!

For more advanced models, **Mplus** is, by and large, the most capable software. It also has a **R** package to send data for estimation: **MplusAutomation** (Hallquist & Wiley, 2018).

Thank you for the kind attention!

References I

- Hallquist, M. N., & Wiley, J. F. (2018). MplusAutomation: An R Package for Facilitating Large-Scale Latent Variable Analyses in Mplus. Structural Equation Modeling: A Multidisciplinary Journal, 25(4), 621–638.
- Merlo, J., Chaix, B., Yang, M., Lynch, J., & Råstam, L. (2005). A Brief Conceptual Tutorial of Multilevel Analysis in Social Epidemiology: Linking the Statistical Concept of Clustering to the Idea of Contextual Phenomenon. *Journal of Epidemiology and Community Health*, 59(6), 443-449.
- Ryu, E., & West, S. G. (2009). Level-Specific Evaluation of Model Fit in Multilevel Structural Equation Modeling. *Structural Equation Modeling: A Multidisciplinary Journal*, 16(4), 583–601.
- Zhang, Z., Zyphur, M. J., & Preacher, K. J. (2009). Testing Multilevel Mediation Using Hierarchical Linear Models: Problems and Solutions. *Organizational Research Methods*, 12(4), 695–719.