Annual Progression Review

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Literature Review

Gravitational waves (GWs) are perturbations of spacetime generated by mass distributions with a non-vanishing second derivative of the quadrupole mass-moment. Binary systems of stellar objects are exemplary GW emitters and the more compact the object, the larger the wave amplitude, making the system more suitable for detection [1].

The LIGO-Virgo-KAGRA (LVK) collaboration uses a system of interferometers to detect gravitational waves [2–4]. The most recent complete catalogue they published is the third gravitational-wave transient catalogue (GWTC-3), which contains about a hundred signals generated by the coalescence of binaries of neutron stars (NSs) and stellar-mass black holes (BHs) [5].

When LVK detects an astrophysical signal, it analyzes the data to infer the properties of the event, such as the masses of the binary components, their spins, the distance from the observer and the sky localization. The statistical framework used for individual event analysis is Bayesian inference. As the number of detections is growing, population analyses become crucial, as they enable further constraints on the astrophysics of compact objects (COs) [5–8]. To this end, one uses Hierarchical Bayesian statistics [9, 10].

The models used have multi-dimensional parameters, resulting in probability distributions computationally prohibitive. Traditional approaches rely on stochastic sampling methods (e.g. Markov Chain Monte Carlo (MCMC) or Nested sampling) to infer probability distributions [10–13]. However powerful, this approach is computationally expensive, as probability distributions are evaluated iteratively million of times for each analysis. Evaluations require to generate a waveform on the fly, which substantially contributes to the overall computational cost. Machine learning approaches have the advantage of shifting most of the cost to the training stage, dramatically speeding up inference [14–16].

1.1 Bayesian statistics

Let us refer to the vector of binary parameters as θ . It consists of intrinsic parameters, such as the masses and spins, and extrinsic parameters, such as the redshift and binary orientation. The goal of the analysis is to infer the posterior distribution $p(\theta|d)$, which assigns to each point in the binary-parameter space the probability

that binary the corresponding binary generated the observed data d. The observed data is the sum of the astrophysical signal, which is assumed to be deterministic and a function of binary-parameters, and a noise realization in the detector, which is assumed to be stochastic. To infer the posterior one needs a prior $\pi(\theta)$ and a likelihood $\mathcal{L}(d|\theta)$ [17]. The prior is the probability distribution that an event with parameters θ occurs in the Universe. A possible approach is to construct the prior based on theory or previous experiments. Nevertheless, the preferred approach is often conservative, using uninformative priors, such as priors with constant probability inside the binary-parameters domain and zero outside. The likelihood is the probability in the data space, given parameters θ . To evaluate the likelihood one needs to generate waveforms for θ and a noise model [18, 19]. To construct the noise model one often assumes that noise is stationary and Gaussian. Bayes' theorem links the posterior to the prior and the likelihood:

$$p(\theta|d) = \frac{\pi(\theta) \times \mathcal{L}(d|\theta)}{p(d)}.$$
 (1.1)

The term p(d) is called evidence, and it is the probability of observing the data d under the assumed model. The evidence is defined as

$$p(d) = \int d\theta \pi(\theta) \times \mathcal{L}(d|\theta). \tag{1.2}$$

The evidence acts as a normalization factor and does not affect the posterior shape. For this reason, it is typically not computed during inference, except when performing model comparison.

1.2 Hierarchical Bayesian statistics

The goal of population analysis is to infer how binary parameters are distributed at the population level. First, one parametrizes the distribution of binary parameters with the population parameters λ , which determine the shape of the distribution. Examples of such hyperparameters include the slope of the binary black hole (BBH) mass distribution, as well as the minimum and maximum mass of stellar-mass BHs. The prior on binary parameters is conditioned on the population parameters, and we refer to it as the population model $p_{\text{pop}}(\theta|\lambda)$. For the analysis, one also needs the prior on population parameters, called hyperprior, $\pi(\lambda)$.

The choice of the population model is crucial, and several approaches have been proposed to construct it. The state-of-the-art methods are parametric models, which rely on simple parametric forms motivated by qualitative astrophysical arguments [5, 9, 20]. Other approaches move toward more flexible models, referred to as semi-parametric or non-parametric models, where the hyperparameters are for instance spline knots and perturbations [21–23]. These models enable data-driven fits but their results are often more difficult to interpret physically. Astrophysical approaches often consist of comparing observed data with the output of simulations that synthesize binary populations [24–27].

In population analyses, it is necessary to take selection effects into account [28–30]. Since different events have different probabilities of being detected, the observed

population is biased. An event is considered detectable if the data exceeds a given threshold. The detection probability for an event is defined as

$$p_{\text{det}}(\theta) = \int_{d>treshold} ddp(d|\theta). \tag{1.3}$$

Let us refer to the dataset of observed events as d and to individual events as d_i , with $i = 1, ..., N_{\text{obs}}$, where N_{obs} is the number of observed events. The posterior for the astrophysical distribution, called the hyperposterior, is:

$$p(\lambda|d) = \frac{\pi(\lambda)}{p(d)} \prod_{i=1}^{N_{\text{obs}}} \frac{\int d\theta \mathcal{L}(d_i|\theta) p_{\text{pop}}(\theta|\lambda)}{\int d\theta p_{\text{pop}}(\theta|\lambda) p_{\text{det}}(\theta)},$$
(1.4)

where $\mathcal{L}(d_i|\theta)$ is the likelihood of individual events as defined in Eq. 1.1. The term p(d) is the hyperevidence, and it is the probability of the observed dataset under the assumed population model, i.e.:

$$p(d) = \int d\lambda' \pi(\lambda') \prod_{i=1}^{N_{\text{obs}}} \mathcal{L}(d_i | \lambda').$$
 (1.5)

Here, the likelihood respect to the hyperparameters $\mathcal{L}(d_i|\lambda')$ is related to the individualevent likelihood through

$$\mathcal{L}(d_i|\lambda') = \int d\theta \mathcal{L}(d_i|\theta) p_{\text{pop}}(\theta|\lambda). \tag{1.6}$$

The hyperevidence is often computed in population analysis to compare different population models and assess how well they describe the observed dataset.

Population analyses are typically carried out in two steps. First, individual events are analyzed using Bayesian statistics. Then, one uses the resulting individual-event posteriors and the assumed population model to infer the hyperposterior in Eq. 1.4.

1.3 Simulation based inference

Simulation-based inference (SBI) is a deep learning technique for approximating probability distributions using simulated data [31–34]. I am working with the DINGO code, which applies SBI to model posteriors of gravitational-wave data [15, 16, 35, 36]. To simulate data, we need to draw samples from the prior and the likelihood. Unlike to traditional methods, SBI does not require explicit evaluation of the likelihood. In principle, more complex models that lack an analytical expression can be used. For instance, one can relax the assumptions of stationary and gaussian noise, or use waveforms generated from simulations [37, 38].

Normalizing flows are often used in SBI to approximate complex distributions. A normalizing flow is a mapping from a simple $\pi(u)$ to an approximation of the posterior $q_{\phi}(\theta|d)$, where ϕ are the parameters of the mapping. The mapping $f_{\phi}(u)$ must be invertible and with a simple Jacobian determinant, so that:

$$q_{\phi}(\theta|d) = \pi(f_{\phi}^{-1}(\theta)) \left| \det J_{f_{\phi}}^{-1} \right|.$$
 (1.7)

To evaluate $q_{\phi}(\theta|d)$ one uses Eq. 1.7. Drawing samples from $q_{\phi}(\theta|d)$ is equivalent to drawing $u \sim \pi(u)$ and transforming $\theta = f_{\phi}(u)$. One often chooses the base distribution $\pi(u)$ to be a multivariate standard normal. In the case of DINGO, the flow f_{ϕ} is a sequence of transforms $f_{\phi}^{(i)}$, such that

$$f_{\phi}^{i}(u) = \begin{cases} u_{i} & \text{if } i \leq d/2\\ c_{i}(u_{i}; u_{1:d/2}, d) & \text{if } i > d/2 \end{cases}.$$
 (1.8)

Half of the latent variables u remain unchanged, while the other half are transformed. At each $f_{\phi}^{(i)}$ the indices are permuted. The transformation used is the neural spline coupling transform c_i , whose parameters depend on the unchanged u_i and on data strain d.

During training, the neural network learns the mapping by optimizing the parameters ϕ to minimize a loss function. In these problems, the loss functions is typically based on the Kullback-Leibler divergence, defined as:

$$D_{\mathrm{KL}}(p||q_{\phi}) = \int d\theta p(\theta|d) \log \frac{p(\theta|d)}{q_{\phi}(\theta|d)}, \tag{1.9}$$

which vanishes as the approximant becomes closer to the true posterior. Since the approximant should not depend on the individual data, we define the loss function as the expectation value over p(d) of $D_{KL}(p||q_{\phi})$:

$$L = \int ddp(d) \int d\theta p(\theta|d) \log \frac{p(\theta|d)}{q_{\phi}(\theta|d)}.$$
 (1.10)

Using Bayes Theorem one obtains:

$$L \simeq -\int d\theta \pi(\theta) \int ddp(d|\theta) \log q_{\phi}(\theta|d), \qquad (1.11)$$

where the posterior at the numerator of the logarithm can be omitted since the minimization over ϕ does not depend on the true posterior. Simulating events is equivalent to drawing N events from the prior, $\theta_i \sim \pi(\theta)$, and data from the likelihood $d_i \sim p(d|\theta_i)$, with i = 1, ..., N. The loss is then given by

$$L = -\frac{1}{N} \sum_{i} \log q_{\phi}(\theta|d), \qquad (1.12)$$

where Monte Carlo (MC) sums are used to approximate the integrals [39].

DINGO architecture uses an embedding network to compress data strains before passing them to the flow [16]. We train the embedding jointly with the flow. During training the network learns how to compress the data, i.e. it identifies the most important features.

Completed Work

Transfer learning is a set of techniques in machine learning used to transfer knowledge from one neural network to another.

My first projects involves testing transfer learning with DINGO for individual-events analysis, following the multifidelity approach in Ref. [40]. The typical training dataset size required to obtain well-trained NNs in DINGO is of the order of 10^6 simulations. Let us consider two waveform models. A high-fidelity model is accurate but computational expensive, with prohibitive costs that limit the number of simulations to well below 10^6 (e.g. numerical relativity simulations). A low-fidelity model is less accurate and cheap, allowing us to easily simulate 10^6 data. Assuming that the flow architecture is the same in both cases, we fine-tune the high-fidelity network using the parameters optimized during the low-fidelity training run. The NN for the high-fidelity model achieves higher performances when pretrained, partially inheriting knowledge acquired during the low-fidelity training.

2.1 Diagnostics to check the performance of the neural network

There are several diagnostics to assess the quality of the posterior approximant.

The evolution of the loss function provides information on how training is progressing. When training a NN, the dataset is split into a training set and validation set. The training set is used to update the NN parameters in order to minimize the loss. The validation set is used to compute the loss after each update, allowing one to monitor the NN's performance on unseen data. A common sanity check consists in ensure that the validation and training losses decrease together during training. By the end of training, both losses should reach a plateau.

Importance sampling tests how closely $q_{\phi}(\theta|d)$ approximates $p(\theta|d)$ in multiple dimensions. During inference one draws samples θ_i from the approximant of the posterior $\theta_i \sim q_{\phi}(\theta|d)$. Each θ_i can be assigned a weight w_i :

$$w_i \propto \frac{\pi(\theta_i)p(d|\theta_i)}{q_\phi(\theta_i|d)}.$$
 (2.1)

The variance of weights provides a measure of the sampling efficiency

$$\epsilon = \frac{\left(\sum_{i} w_{i}\right)^{2}}{n \sum_{i} w_{i}^{2}} \tag{2.2}$$

[36]. Smaller values of ϵ correspond to poorer posterior approximations.

2.2 From non-spinning to aligned-spins model

To test the multifidelity procedure, we use a non-spinning waveform as the low-fidelity model, while the high-fidelity model includes aligned spins. We use IMR-PhenomXPHM as the approximant to generate waveforms is . For the low-fidelity model, we generate $\sim 10^6$ simulations. For the high-fidelity model, we generate multiple datasets of sizes $\sim 10^5$, $\sim 10^4$ and $\sim 10^3$. For the datasets of $\sim 10^6$ and $\sim 10^3$ simulations, we also perform additional training where we pretrain the embedding only. Results for pretraining only the embedding are preliminary. We plan to explore this approach further, as it could improve the flexibility of the high-fidelity model, allowing us to use pretraining while modifying the flow architecture.

Figure 2.1 shows the evolution of the validation and set loss functions during training. In the two cases with largest number of simulations ($\sim 10^6$ and $\sim 10^5$), pretraining both the embedding and the flow does not significantly affect the final loss plateau. The main advantage of pretraining is the reduced time to reach the plateau, resulting in lower computational costs. For limited simulation budgets ($\sim 10^4$ and $\sim 10^3$), pretraining reduces the loss to lower values than traditional training. Pretraining the embedding alone with $\sim 10^6$ simulations allows a slightly lower final loss compared to training from scratch. For $\sim 10^3$ simulations, this strategy improves results over traditional training but does not achieve the performance of pretraining both the embedding and the flow.

The cornerplot in Fig. 2.2 shows the analysis of an injection for different simulation budgets when pretraining both the flow and the embedding. The marginal posteriors are consistent across budgets. However, the marginal posterior for χ_1 deviates respect to the injected value, while the marginal posterior for χ_2 is close to the prior.

The sampling efficiencies for 1000 injections (Fig. 2.3) confirm that pretraining with $\sim 10^4$ and $\sim 10^3$ simulations improves performance. The histogram in Fig. 2.3 shows that even with pretraining, the $\sim 10^3$ simulation case has considerably lower efficiencies than the other cases, with most injections having $\epsilon < 1\%$. Contrarily to expectations, Fig. 2.3 shows no correlation between sampling efficiency and the higher spin magnitude.

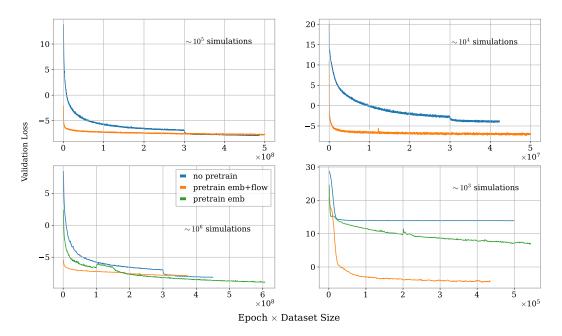


Figure 2.1: Evolution of the validation loss during training. The number of epochs corresponds to the number of times the full training set passes through the neural network. The top two panels show the losses for traditional training (blue) and pretraining both the embedding and the flow (orange) for $\sim 10^5$ (left) and $\sim 10^4$ (right) simulations. The bottom two panels show the validation losses for $\sim 10^6$ (left) and $\sim 10^3$ (right) simulations. In addition to the previous cases, results for pretraining the embedding only (green) are also shown.

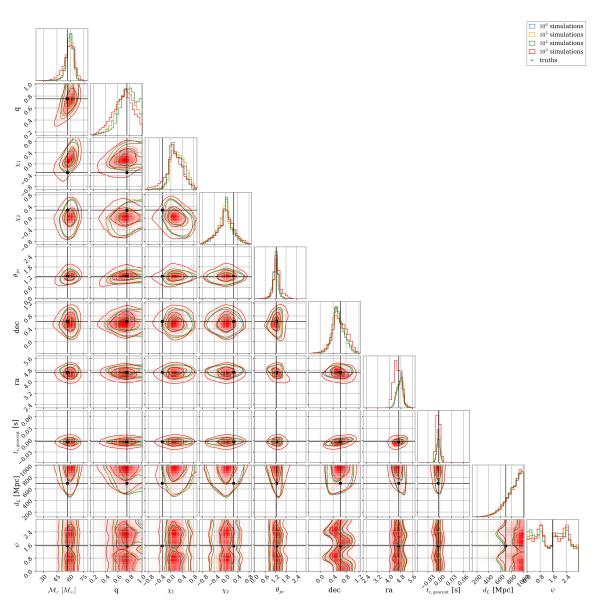


Figure 2.2: Cornerplots showing the one-dimensional (diagonal) and two-dimensional (off-diagonal) posteriors for binary parameters given an injected signal (black lines and markers). Results are obtained using DINGO with pretraining of both the embedding and the flow, with training datasets of order of magnitude $\sim 10^6$ (blue), $\sim 10^5$ (orange), $\sim 10^4$ (green), $\sim 10^3$ (red).

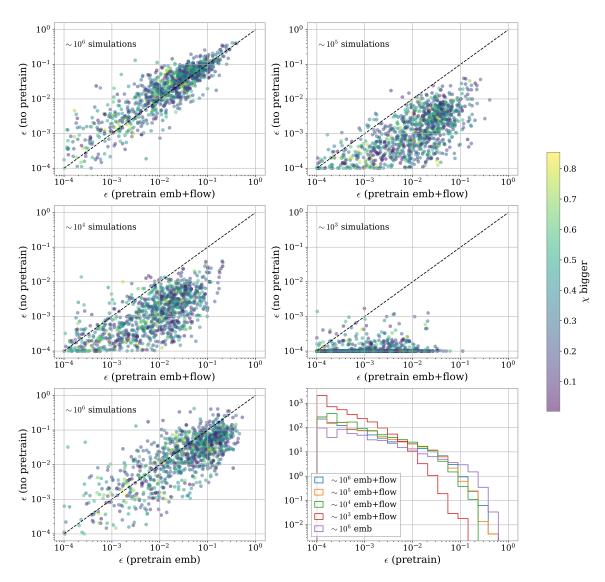


Figure 2.3: Scatter plots of sampling efficiecies for 10^3 injected signals, comparing pretraining (x axis) and no pretraining (y axis). The black dotted line corresponds to y=x; injections below the line have higher efficiency with pretraining. Markers are colored based on the higher-spin value. The top four plots show results when pretraining both the embedding and the flow for different simulation budgets, while the bottom-left plot shows results when pretraining the embedding only ($\sim 10^6$ simulations). The bottom-right panel displays histograms of the sampling efficiencies for pretraining cases.

Work Plan

In the coming months I will continue working on the pretraining project, focusing on whether spin marginals are properly learned. With pretraining we can analyze events with numerical relativity waveform models, using, for instance, beyond general relativity (GR) models, or precessing and eccentric waveforms.

Some attempts are being to apply machine learning to populations [14, 41, 42]. One question we aim to address is whether SBI could solve the "growing pains" problem [43]. This issue arises because MC integrals have an intrinsic variance. In traditional methods likelihoods are estimated through MC integrals, which introduces an uncertainty that grows with the dataset size, and can bias future analyses. Another possible direction is to apply pretraining in population analyses, using population-synthesis codes to generate the dataset. This would allow us to perform inference directly on astrophysical processes, since population parameters would be the input parameters of the codes, e.g. the supernova kick strength, common-envelope efficiency, metallicity, and winds strength.

Personal Development Plan

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