

# Perceptron and Logistic Regression

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unless otherwise stated

After this lecture you should be able to

- Think about binary classification using **geometric intuition** and use the **perceptron algorithm**.
- Define the **main concepts of information theory** (entropy, cross-entropy, KL-divergence) and prove their properties.
- Derive training objectives using the **maximum likelihood principle**.
- Implement and use **logistic regression** for binary classification with SGD.

Binary classification is a classification in two classes.

The simplest way to evaluate classification is **accuracy**, which is the ratio of input examples that were classified correctly – i.e., where the predicted class and the target class match.

To extend linear regression to binary classification, we might seek a **threshold** and then classify an input as negative/positive depending on whether  $y(\mathbf{x}; \mathbf{w}) = \mathbf{x}^T \mathbf{w} + b$  is smaller/larger than a given threshold.

Zero value is usually used as the threshold, both because of symmetry and also because the **bias** parameter acts as a trainable threshold anyway.

The set of points with prediction 0 is called a **decision boundary**.

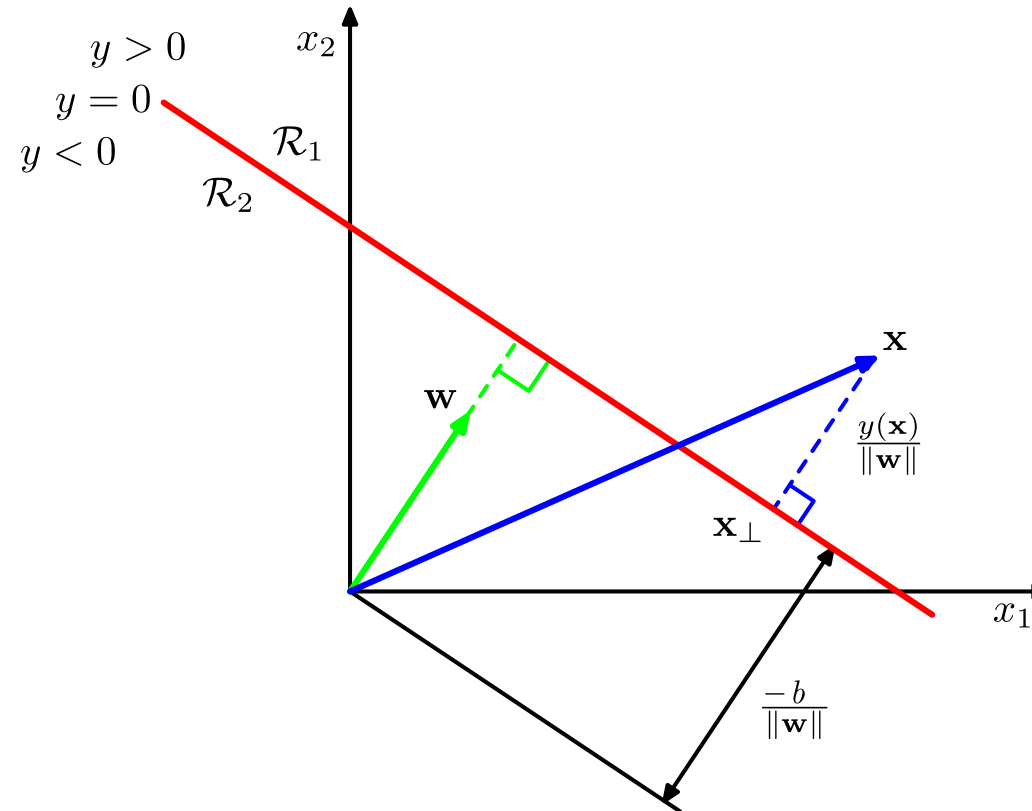


Figure 4.1 of Pattern Recognition and Machine Learning.

The perceptron algorithm is probably the oldest one for training weights of a binary classification. Assuming the target value  $t \in \{-1, +1\}$ , the goal is to find weights  $\mathbf{w}$  such that for all train data,

$$\text{sign}(y(\mathbf{x}_i; \mathbf{w})) = \text{sign}(\mathbf{x}_i^T \mathbf{w}) = t_i,$$

or equivalently,

$$t_i y(\mathbf{x}_i; \mathbf{w}) = t_i \mathbf{x}_i^T \mathbf{w} > 0.$$

Note that a set is called **linearly separable**, if there exists a weight vector  $\mathbf{w}$  such that the above equation holds.

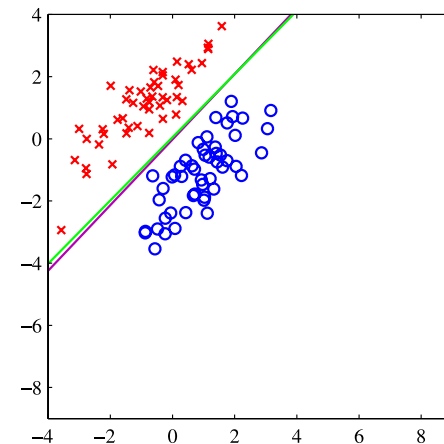


Figure 4.4 of Pattern Recognition and Machine Learning.

The perceptron algorithm was invented by Rosenblatt in 1958.

**Input:** Linearly separable dataset ( $\mathbf{X} \in \mathbb{R}^{N \times D}$ ,  $\mathbf{t} \in \{-1, +1\}^N$ ).

**Output:** Weights  $\mathbf{w} \in \mathbb{R}^D$  such that  $t_i \mathbf{x}_i^T \mathbf{w} > 0$  for all  $i$ .

- $\mathbf{w} \leftarrow \mathbf{0}$
- until all examples are classified correctly, process example  $i$ :
  - $y \leftarrow \mathbf{x}_i^T \mathbf{w}$
  - if  $t_i y \leq 0$  (incorrectly classified example):
    - $\mathbf{w} \leftarrow \mathbf{w} + t_i \mathbf{x}_i$

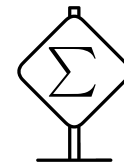
We will prove that the algorithm always arrives at some correct set of weights  $\mathbf{w}$  if the training set is linearly separable.

# Proof of Perceptron Convergence

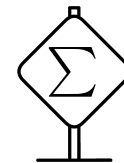
Let  $\mathbf{w}_*$  be some weights correctly classifying (separating) the training data, and let  $\mathbf{w}_k$  be the weights after  $k$  nontrivial updates of the perceptron algorithm, with  $\mathbf{w}_0$  being 0.

We will prove that the angle  $\alpha$  between  $\mathbf{w}_*$  and  $\mathbf{w}_k$  decreases at each step. Note that

$$\cos(\alpha) = \frac{\mathbf{w}_*^T \mathbf{w}_k}{\|\mathbf{w}_*\| \cdot \|\mathbf{w}_k\|}.$$



Assume that the maximum norm of any training example  $\|\mathbf{x}\|$  is bounded by  $R$ , and that  $\gamma$  is the minimum margin of  $\mathbf{w}_*$ , so for each training example  $(\mathbf{x}, t)$ ,  $t\mathbf{x}^T \mathbf{w}_* \geq \gamma$ .



First consider the dot product of  $\mathbf{w}_*$  and  $\mathbf{w}_k$ :

$$\mathbf{w}_*^T \mathbf{w}_k = \mathbf{w}_*^T (\mathbf{w}_{k-1} + t_k \mathbf{x}_k) \geq \mathbf{w}_*^T \mathbf{w}_{k-1} + \gamma.$$

By iteratively applying this equation, we get

$$\mathbf{w}_*^T \mathbf{w}_k \geq k\gamma.$$

Now consider the length of  $\mathbf{w}_k$ :

$$\|\mathbf{w}_k\|^2 = \|\mathbf{w}_{k-1} + t_k \mathbf{x}_k\|^2 = \|\mathbf{w}_{k-1}\|^2 + 2t_k \mathbf{x}_k^T \mathbf{w}_{k-1} + \|\mathbf{x}_k\|^2.$$

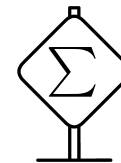
Because  $\mathbf{x}_k$  was misclassified, we know that  $t_k \mathbf{x}_k^T \mathbf{w}_{k-1} \leq 0$ , so  $\|\mathbf{w}_k\|^2 \leq \|\mathbf{w}_{k-1}\|^2 + R^2$ .

When applied iteratively, we get  $\|\mathbf{w}_k\|^2 \leq k \cdot R^2$ .



Putting everything together, we get

$$\cos(\alpha) = \frac{\mathbf{w}_*^T \mathbf{w}_k}{\|\mathbf{w}_*\| \cdot \|\mathbf{w}_k\|} \geq \frac{k\gamma}{\sqrt{k}R^2\|\mathbf{w}_*\|}.$$

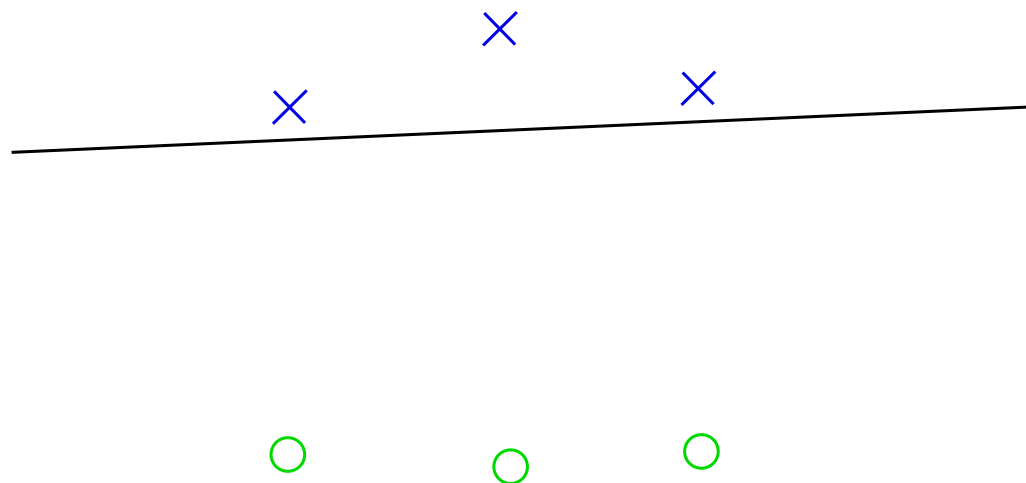


Therefore, the  $\cos(\alpha)$  increases during every update. Because the value of  $\cos(\alpha)$  is at most one, we can compute the upper bound on the number of steps when the algorithm converges as

$$1 \geq \frac{\sqrt{k}\gamma}{\sqrt{R^2}\|\mathbf{w}_*\|} \text{ or } k \leq \frac{R^2\|\mathbf{w}_*\|^2}{\gamma^2}.$$

Perceptron has several drawbacks:

- If the input set is not linearly separable, the algorithm never finishes.
- The algorithm performs only prediction, it is not able to return the probabilities of predictions.
- Most importantly, Perceptron algorithm finds *some* solution, not necessarily a good one, because once it finds some, it cannot perform any more updates.



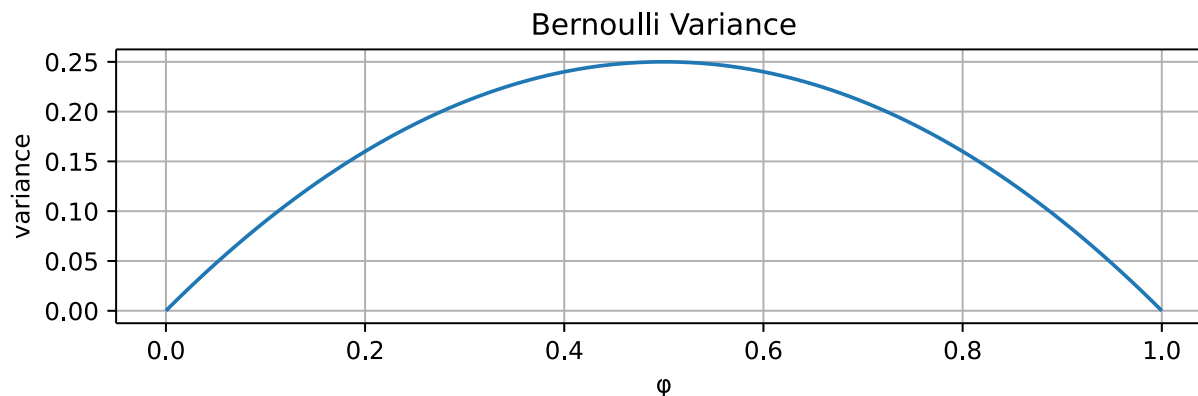
## Bernoulli Distribution

The Bernoulli distribution is a distribution over a binary random variable. It has a single parameter  $\varphi \in [0, 1]$ , which specifies the probability that the random variable is equal to 1.

$$P(x) = \varphi^x (1 - \varphi)^{1-x}$$

$$\mathbb{E}[x] = \varphi$$

$$\text{Var}(x) = \varphi(1 - \varphi)$$



## Categorical Distribution

Extension of the Bernoulli distribution to random variables taking one of  $K$  different discrete outcomes. It is parametrized by  $\mathbf{p} \in [0, 1]^K$  such that  $\sum_{i=0}^{K-1} p_i = 1$ .

We represent outcomes as vectors  $\in \{0, 1\}^K$  in the **one-hot encoding**. Therefore, an outcome  $x \in \{0, 1, \dots, K - 1\}$  is represented as a vector

$$\mathbf{1}_x \stackrel{\text{def}}{=} ([i = x])_{i=0}^{K-1} = (\underbrace{0, \dots, 0}_x, 1, \underbrace{0, \dots, 0}_{K-x-1}).$$

The outcome probability, mean, and variance are very similar to the Bernoulli distribution.

$$\begin{aligned} P(\mathbf{x}) &= \prod_{i=0}^{K-1} p_i^{x_i} \\ \mathbb{E}[x_i] &= p_i \\ \text{Var}(x_i) &= p_i(1 - p_i) \end{aligned}$$

## Self Information

Amount of **surprise** when a random variable is sampled.

- Should be zero for events with probability 1.
- Less likely events are more surprising.
- Independent events should have **additive** information.

$$I(x) \stackrel{\text{def}}{=} -\log P(x) = \log \frac{1}{P(x)}$$

## Entropy

Amount of **surprise** in the whole distribution.

$$H(P) \stackrel{\text{def}}{=} \mathbb{E}_{x \sim P}[I(x)] = -\mathbb{E}_{x \sim P}[\log P(x)]$$

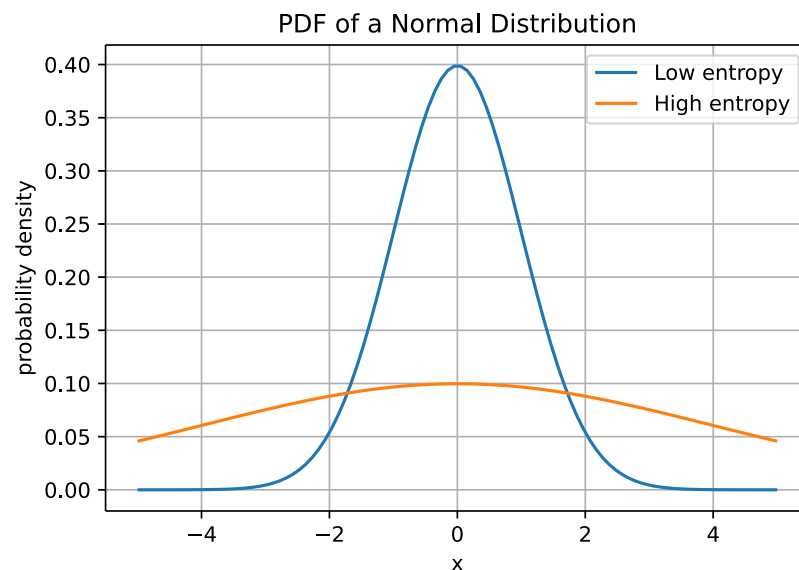
- for discrete  $P$ :  $H(P) = -\sum_x P(x) \log P(x)$
- for continuous  $P$ :  $H(P) = -\int P(x) \log P(x) dx$

Because  $\lim_{x \rightarrow 0} x \log x = 0$ , for  $P(x) = 0$  we consider  $P(x) \log P(x)$  to be zero.

Note that in the continuous case, the continuous entropy (also called *differential entropy*) has slightly different semantics, for example, it can be negative.

For binary logarithms, the entropy is measured in **bits**.

However, from now on, all logarithms are *natural logarithms* with base  $e$  (and then the entropy is measured in units called **nats**).



## Cross-Entropy

$$H(P, Q) \stackrel{\text{def}}{=} -\mathbb{E}_{x \sim P} [\log Q(x)]$$

### Gibbs Inequality

- $H(P, Q) \geq H(P)$
- $H(P) = H(P, Q) \Leftrightarrow P = Q$

Proof: Consider  $H(P) - H(P, Q) = \sum_x P(x) \log \frac{Q(x)}{P(x)}$ .

Using the fact that  $\log x \leq (x - 1)$  with equality only for  $x = 1$ , we get

$$\sum_x P(x) \log \frac{Q(x)}{P(x)} \leq \sum_x P(x) \left( \frac{Q(x)}{P(x)} - 1 \right) = \sum_x Q(x) - \sum_x P(x) = 0.$$

For the equality to hold,  $\frac{Q(x)}{P(x)}$  must be 1 for all  $x$ , i.e.,  $P = Q$ .

## Kullback-Leibler Divergence (KL Divergence)

Sometimes also called **relative entropy**.

$$D_{\text{KL}}(P\|Q) \stackrel{\text{def}}{=} H(P, Q) - H(P) = \mathbb{E}_{x \sim P}[\log P(x) - \log Q(x)]$$

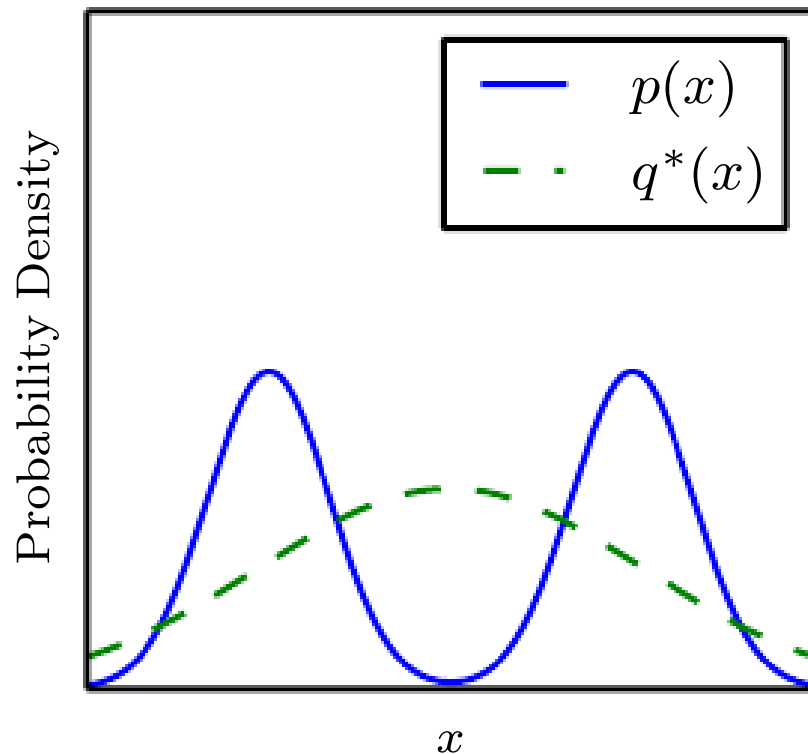
- consequence of Gibbs inequality:  $D_{\text{KL}}(P\|Q) \geq 0$ ,  $D_{\text{KL}}(P\|Q) = 0$  iff  $P = Q$
- generally  $D_{\text{KL}}(P\|Q) \neq D_{\text{KL}}(Q\|P)$



# Nonsymmetry of KL Divergence

Assume we want find the best unimodal distribution  $q$ .

$$q^* = \operatorname{argmin}_q D_{\text{KL}}(p||q)$$



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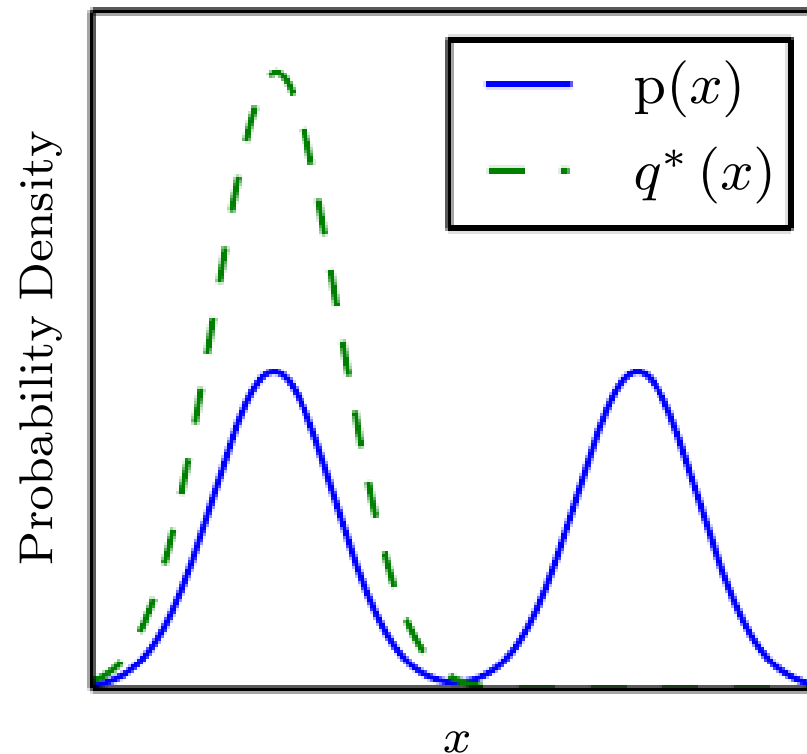


Figure 3.6 of "Deep Learning" book, <https://www.deeplearningbook.org>

## Normal (or Gaussian) Distribution

Distribution over real numbers, parametrized by a mean  $\mu$  and variance  $\sigma^2$ :

$$\mathcal{N}(x; \mu, \sigma^2) = \sqrt{\frac{1}{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

For standard values  $\mu = 0$  and  $\sigma^2 = 1$  we get  $\mathcal{N}(x; 0, 1) = \sqrt{\frac{1}{2\pi}} e^{-\frac{x^2}{2}}$ .

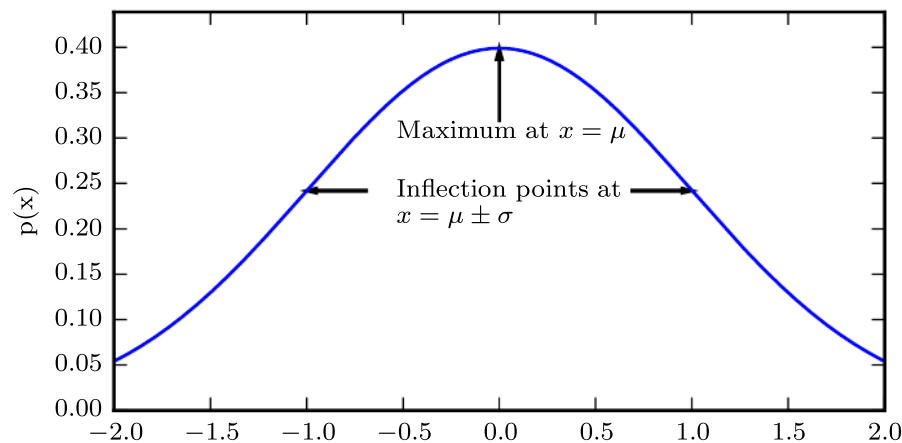


Figure 3.1 of "Deep Learning" book, <https://www.deeplearningbook.org>.

## Central Limit Theorem

The sum of independent identically distributed random variables with finite variance converges to normal distribution.

## Principle of Maximum Entropy

Given a set of constraints, a distribution with maximal entropy fulfilling the constraints can be considered the most general one, containing as little additional assumptions as possible.

Considering distributions with a **given mean and variance**, it can be proven (using variational inference) that such a distribution with **maximum entropy** is exactly the normal distribution.

Let  $\mathbf{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$  be training data drawn independently from the data-generating distribution  $p_{\text{data}}$ .

We denote the **empirical data distribution** as  $\hat{p}_{\text{data}}$ , where

$$\hat{p}_{\text{data}}(\mathbf{x}) \stackrel{\text{def}}{=} \frac{|\{i : \mathbf{x}_i = \mathbf{x}\}|}{N}.$$

Let  $p_{\text{model}}(\mathbf{x}; \mathbf{w})$  be a family of distributions.

- If the weights are fixed,  $p_{\text{model}}(\mathbf{x}; \mathbf{w})$  is a probability distribution.
- If we instead consider the fixed training data  $\mathbf{X}$ , then

$$L(\mathbf{w}) = p_{\text{model}}(\mathbf{X}; \mathbf{w}) = \prod_{i=1}^N p_{\text{model}}(\mathbf{x}_i; \mathbf{w})$$

is called the **likelihood**. Note that even if the value of the likelihood is in range  $[0, 1]$ , it is not a probability, because the likelihood is not a probability distribution.

# Maximum Likelihood Estimation

Let  $\mathbf{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$  be training data drawn independently from the data-generating distribution  $p_{\text{data}}$ . We denote the empirical data distribution as  $\hat{p}_{\text{data}}$  and let  $p_{\text{model}}(\mathbf{x}; \mathbf{w})$  be a family of distributions.

The **maximum likelihood estimation** of  $\mathbf{w}$  is:

$$\begin{aligned}
 \mathbf{w}_{\text{MLE}} &= \arg \max_{\mathbf{w}} p_{\text{model}}(\mathbf{X}; \mathbf{w}) = \arg \max_{\mathbf{w}} \prod_{i=1}^N p_{\text{model}}(\mathbf{x}_i; \mathbf{w}) \\
 &= \arg \min_{\mathbf{w}} \sum_{i=1}^N -\log p_{\text{model}}(\mathbf{x}_i; \mathbf{w}) \\
 &= \arg \min_{\mathbf{w}} \mathbb{E}_{\mathbf{x} \sim \hat{p}_{\text{data}}} [-\log p_{\text{model}}(\mathbf{x}; \mathbf{w})] \\
 &= \arg \min_{\mathbf{w}} H(\hat{p}_{\text{data}}(\mathbf{x}), p_{\text{model}}(\mathbf{x}; \mathbf{w})) \\
 &= \arg \min_{\mathbf{w}} D_{\text{KL}}(\hat{p}_{\text{data}}(\mathbf{x}) \| p_{\text{model}}(\mathbf{x}; \mathbf{w})) + H(\hat{p}_{\text{data}}(\mathbf{x}))
 \end{aligned}$$

# Maximum Likelihood Estimation

MLE can be easily generalized to the conditional case, where our goal is to predict  $t$  given  $\mathbf{x}$ :

$$\begin{aligned}
 \mathbf{w}_{\text{MLE}} &= \arg \max_{\mathbf{w}} p_{\text{model}}(\mathbf{t}|\mathbf{X}; \mathbf{w}) = \arg \max_{\mathbf{w}} \prod_{i=1}^N p_{\text{model}}(t_i|\mathbf{x}_i; \mathbf{w}) \\
 &= \arg \min_{\mathbf{w}} \sum_{i=1}^N -\log p_{\text{model}}(t_i|\mathbf{x}_i; \mathbf{w}) \\
 &= \arg \min_{\mathbf{w}} \mathbb{E}_{(\mathbf{x}, t) \sim \hat{p}_{\text{data}}} [-\log p_{\text{model}}(t|\mathbf{x}; \mathbf{w})] \\
 &= \arg \min_{\mathbf{w}} H(\hat{p}_{\text{data}}(t|\mathbf{x}), p_{\text{model}}(t|\mathbf{x}; \mathbf{w})) \\
 &= \arg \min_{\mathbf{w}} D_{\text{KL}}(\hat{p}_{\text{data}}(t|\mathbf{x}) || p_{\text{model}}(t|\mathbf{x}; \mathbf{w})) + H(\hat{p}_{\text{data}}(t|\mathbf{x}))
 \end{aligned}$$

where the conditional entropy is defined as  $H(\hat{p}_{\text{data}}) = \mathbb{E}_{(\mathbf{x}, t) \sim \hat{p}_{\text{data}}} [-\log(\hat{p}_{\text{data}}(t|\mathbf{x}; \mathbf{w}))]$  and the conditional cross-entropy as  $H(\hat{p}_{\text{data}}, p_{\text{model}}) = \mathbb{E}_{(\mathbf{x}, t) \sim \hat{p}_{\text{data}}} [-\log(p_{\text{model}}(t|\mathbf{x}; \mathbf{w}))]$ .

The resulting *loss function* is called **negative log-likelihood (NLL)**, or **cross-entropy**, or **Kullback-Leibler divergence**.

An extension of perceptron, which models the conditional probabilities of  $p(C_0|\mathbf{x})$  and of  $p(C_1|\mathbf{x})$ . Logistic regression can in fact handle also more than two classes, which we will see in the next lecture.

Logistic regression employs the following parametrization of the conditional class probabilities:

$$\begin{aligned} p(C_1|\mathbf{x}) &= \sigma(\mathbf{x}^T \mathbf{w} + b) \\ p(C_0|\mathbf{x}) &= 1 - p(C_1|\mathbf{x}), \end{aligned}$$

where  $\sigma$  is a **sigmoid function**

$$\sigma(x) = \frac{1}{1 + e^{-x}}.$$

It can be trained using the SGD algorithm.

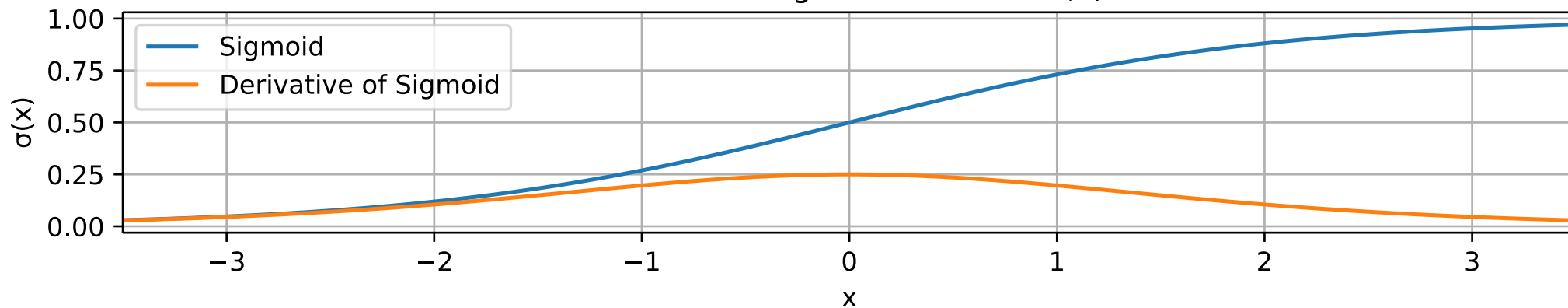
# Sigmoid Function

The sigmoid function has values in range  $(0, 1)$ , is monotonically increasing and it has a derivative of  $\frac{1}{4}$  at  $x = 0$ .

$$\sigma(x) = \frac{1}{1 + e^{-x}}$$

$$\sigma'(x) = \sigma(x)(1 - \sigma(x))$$

Plot of the Sigmoid Function  $\sigma(x)$





We denote the output of the “linear part” of the logistic regression as

$$\bar{y}(\mathbf{x}; \mathbf{w}) = \mathbf{x}^T \mathbf{w},$$

and the overall prediction as

$$y(\mathbf{x}; \mathbf{w}) = \sigma(\bar{y}(\mathbf{x}; \mathbf{w})) = \sigma(\mathbf{x}^T \mathbf{w}).$$

The logistic regression output  $y(\mathbf{x}; \mathbf{w})$  models the probability of class  $C_1$ ,  $p(C_1|\mathbf{x})$ .

To give some meaning to the output of the linear part  $\bar{y}(\mathbf{x}; \mathbf{w})$ , starting with

$$p(C_1|\mathbf{x}) = \sigma(\bar{y}(\mathbf{x}; \mathbf{w})) = \frac{1}{1 + e^{-\bar{y}(\mathbf{x}; \mathbf{w})}},$$

we arrive at

$$\bar{y}(\mathbf{x}; \mathbf{w}) = \log \left( \frac{p(C_1|\mathbf{x})}{1 - p(C_1|\mathbf{x})} \right) = \log \left( \frac{p(C_1|\mathbf{x})}{p(C_0|\mathbf{x})} \right),$$

which is called a **logit** and it is a logarithm of odds of the probabilities of the two classes.

To train the logistic regression, we use MLE (the maximum likelihood estimation). Its application is straightforward, given that  $p(C_1|\mathbf{x}; \mathbf{w})$  is directly the model output  $y(\mathbf{x}; \mathbf{w})$ .

Therefore, the loss for a minibatch  $\mathbb{X} = \{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_N, t_N)\}$  is

$$E(\mathbf{w}) = \frac{1}{N} \sum_i -\log(p(C_{t_i}|\mathbf{x}_i; \mathbf{w})).$$

**Input:** Input dataset  $(\mathbf{X} \in \mathbb{R}^{N \times D}, \mathbf{t} \in \{0, +1\}^N)$ , learning rate  $\alpha \in \mathbb{R}^+$ .

- $\mathbf{w} \leftarrow \mathbf{0}$  or we initialize  $\mathbf{w}$  randomly
- until convergence (or patience runs out), process a minibatch of examples  $\mathbb{B}$ :
  - $\mathbf{g} \leftarrow \frac{1}{|\mathbb{B}|} \sum_{i \in \mathbb{B}} \nabla_{\mathbf{w}} \left( -\log(p(C_{t_i}|\mathbf{x}_i; \mathbf{w})) \right)$
  - $\mathbf{w} \leftarrow \mathbf{w} - \alpha \mathbf{g}$

Everything we learned about **features** and  $L^2$  **regularization** holds for logistic regression too.



After this lecture you should be able to

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