ML_PS3

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```
import numpy as np
import pandas as pd
import os
import time
import statsmodels.api as sm
# Sci-kit learn
from sklearn.model_selection import train_test_split
from sklearn.linear_model import LogisticRegression,LinearRegression, Lasso
from sklearn.metrics import accuracy_score
from sklearn.preprocessing import PolynomialFeatures
from sklearn.model_selection import cross_val_score, cross_validate
from sklearn.model_selection import LeaveOneOut
from sklearn.metrics import mean_squared_error
from sklearn.model_selection import KFold
from sklearn.metrics import mean_squared_error, make_scorer
from mlxtend.feature_selection import ExhaustiveFeatureSelector as EFS
# Statsmodels
import statsmodels.api as sm
# Visualizations
import matplotlib.pyplot as plt
import seaborn as sns
# Iterator building blocks
# example: combinations('ABCD', 2) --> AB AC AD BC BD CD
from itertools import combinations
# Many were concerned with warnings. The short answer is that FutureWarning (most common
# appears when a functionality is deprecated and will be replaced. Here's how to ignore
# even if you should find a way to resolve them in a production environment.
import warnings
```

```
warnings.filterwarnings("ignore", category=FutureWarning)
```

Part 1

1.a.

```
# Load the dataset
directory = r"C:\Users\clari\OneDrive\Documents\Machine Learning\ps3"
default_path = os.path.join(directory, "Data-Default.csv")
default_df = pd.read_csv(default_path)
print(default_df.dtypes)
print(default_df.shape)
default_df.head()
```

default object student object balance float64 income float64 dtype: object (10000, 4)

	default	student	balance	income
0	No	No	729.526495	44361.625074
1	No	Yes	817.180407	12106.134700
2	No	No	1073.549164	31767.138947
3	No	No	529.250605	35704.493935
4	No	No	785.655883	38463.495879

```
na_count = pd.DataFrame(np.sum(default_df.isna(), axis = 0), columns = ["Count NAs"])
print(na_count)
```

	Count	NAs
default		0
student		0
balance		0
income		0

```
encoding_dict = {'Yes': 1,'No': 0}
for col in ['default', 'student']:
    default_df[col] = default_df[col].map(encoding_dict)

default_df.head()
```

	default	student	balance	income
0	0	0	729.526495	44361.625074
1	0	1	817.180407	12106.134700
2	0	0	1073.549164	31767.138947
3	0	0	529.250605	35704.493935
4	0	0	785.655883	38463.495879

getting a sense of the estimates of the coefficients as well as the standard errors associated with them.

```
X = default_df[['balance', 'income']]
X = sm.add_constant(X)

y = default_df['default']

display(X.head(), y.head())
print(y.value_counts())
results = sm.Logit(y, X).fit()
print(results.summary())
```

	const	balance	income
0	1.0	729.526495	44361.625074
1	1.0	817.180407	12106.134700
2	1.0	1073.549164	31767.138947
3	1.0	529.250605	35704.493935
4	1.0	785.655883	38463.495879

```
0 0
1 0
2 0
3 0
4 0
```

Name: default, dtype: int64

default

0 9667 1 333

Name: count, dtype: int64

Optimization terminated successfully.

Current function value: 0.078948

Iterations 10

Logit Regression Results

========				=====			
Dep. Variab	ole:	def	ault	No. C)bservations:		10000
Model:		I	ogit	Df Re	esiduals:		9997
Method:			MLE	Df Mo	odel:		2
Date:	•	Thu, 20 Feb	2025	Pseud	lo R-squ.:		0.4594
Time:		22:4	7:52	Log-I	Likelihood:		-789.48
converged:			True	LL-Nu	111:		-1460.3
Covariance	Type:	nonro	bust	LLR p	-value:		4.541e-292
========			=====	=====		=======	========
	coef	std err		Z	P> z	[0.025	0.975]
const	-11.5405	0.435	-26	.544	0.000	-12.393	-10.688
balance	0.0056	0.000	24	.835	0.000	0.005	0.006
income	2.081e-05	4.99e-06	4	.174	0.000	1.1e-05	3.06e-05

Possibly complete quasi-separation: A fraction 0.14 of observations can be perfectly predicted. This might indicate that there is complete quasi-separation. In this case some parameters will not be identified.

The SE for balance is 0 and 4.99e-06 for income

1.b.

```
# Defining a function which performs the `logit` on each subsample
def boot_fn(data, index, constant=True, features=['balance','income'], target='default')

X = data[features].loc[index]
if constant:
    X = sm.add_constant(X)
y = data[target].loc[index]

lr = sm.Logit(y,X).fit(disp=0)
coefficients = [lr.params[0], lr.params[1], lr.params[2]]

return coefficients
```

```
# Fundtion to execute it
def boot(data, func, B):
    1 1 1
   Executing a bootstrap over B subsamples
    to estimate the (mean of) coefficients
    and their associated standard errors.
    Inputs:
        - data (pd.Dataframe): data to sample from
        - func (fn): function executing the regression
        - B (int): number of subsamples
    Ouput:
        - restults (dict of dicts): bootstrapped coefficients
            and the associated standard errors
    1.1.1
    # Step 4
    coef_intercept = []
    coef_balance = []
    coef_income = []
    coefs = ['intercept', 'balance', 'income']
    output = {coef: [] for coef in coefs}
    for i in range(B):
        # set new seed (=i) every time you get a new subsample
        reg_out = func(data, get_indices(data, len(data), i))
        for i, coef in enumerate(coefs):
            output[coef].append(reg_out[i])
    # Step 5
    results = {}
```

```
for coef in coefs:
    results[coef] = {
        'estimate': np.mean(output[coef]),
        'std_err': np.std(output[coef])
    }
return results
```

1.d. the estimated SE's using the bootstrap method and the logit one in a are quite close.

Part 2

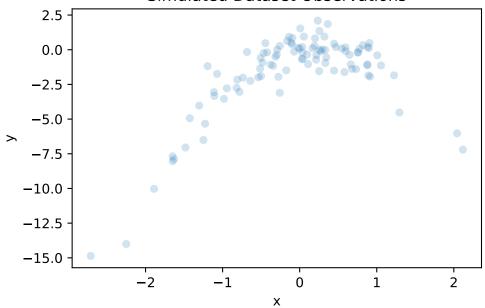
2.a.

```
# Generate a simulated dataset
rng = np.random.default_rng(1)
x = rng.normal(size=100)
y = x - 2 * x**2 + rng.normal(size=100)
```

2.b.

```
# Scatter plot X and Y with fitted line
fig, ax = plt.subplots()
ax.scatter(x, y,
alpha=0.2, s=25)
ax.set_title("Simulated Dataset Observations")
ax.set_xlabel("x")
ax.set_ylabel("y")
plt.show()
```

Simulated Dataset Observations



IN terpretation

2.c.

```
# Set random seed
np.random.seed(10)

simulated_df = pd.DataFrame({"x": x, "y": y})
for i in range(2, 5):
    simulated_df[f'x{i}'] = simulated_df['x']**i

models = [['x'], ['x', 'x2'], ['x', 'x2', 'x3'], ['x', 'x2', 'x3', 'x4']]
for i, model in enumerate(models, 1):
    X = simulated_df[model]
```

```
y = simulated_df['y']

lr = LinearRegression()
scores = cross_val_score(lr, X, y, cv=LeaveOneOut(), scoring='neg_mean_squared_error
loocv_error = -scores.mean()

print(f"Model {i} LOOCV Error: {loocv_error:.6f}")
```

Model 1 LOOCV Error: 6.633030 Model 2 LOOCV Error: 1.122937 Model 3 LOOCV Error: 1.301797 Model 4 LOOCV Error: 1.332394

2.d. There was no change in the results versus those in part c when using a different random seed. This is because inLOOCV, the validation set consists of only one observation, and the process is repeated for all n data points. Since the splitting of data into training and validation sets are without replacement (each observation is used exactly once as the validation set), randomization isn't needed, so changing the random seed din't change the estimates.

```
# Set new random seed
np.random.seed(20)
# Create df of simulated data set
simulated_df = pd.DataFrame({"y": y,
    "x": x})
# Add features
simulated_df["x2"] = simulated_df["x"] ** 2
simulated_df["x3"] = simulated_df["x"] ** 3
simulated_df["x4"] = simulated_df["x"] ** 4
# Prepare list to hold squared errors
squared_errors = {i: [] for i in range(1, 5)}
# Iterate over models
for model_num in range(1, 5):
    for index in simulated_df.index:
        # Split training and validation set
        train = simulated_df.iloc[simulated_df.index != index, :]
        valid = simulated_df.iloc[simulated_df.index == index, :]
```

```
# Select features
        if model_num == 1:
            features = ["x"]
        elif model_num == 2:
            features = ["x", "x2"]
        elif model_num == 3:
            features = ["x", "x2", "x3"]
        else:
            features = ["x", "x2", "x3", "x4"]
        # Fit model and calculate mse
        ols = LinearRegression().fit(train[features], train[["y"]])
        valid_pred = ols.predict(valid[features])
        valid_actual = valid[["y"]]
        squared_error = np.power(valid_pred - valid_actual, 2)
        # Store squared error
        squared_errors[model_num].append(squared_error.values[0][0])
# Calculate mean squared errors (MSE)
for model_num in range(1, 5):
    print(f'MSE for model {model_num} using LOOCV: {round(np.mean(squared_errors[model_num))
```

```
MSE for model 1 using LOOCV: 6.63303
MSE for model 2 using LOOCV: 1.12294
MSE for model 3 using LOOCV: 1.3018
MSE for model 4 using LOOCV: 1.33239
```

2.e. Model 2 had the smallest LOOCV error of 1.12294, confirming that it most closely approximates the true underlying relationship in the data. This result aligns with our expectations because the data was generated using a quadratic relationship between x and yfrom n 2a. Since the true data-generating process is quadratic, a quadratic model i can better capture this relationship compared to the other models.

2.f.

```
Output:
       - Summary of the model
    # Split the data into target and predictors
    X = simulated_df[predictors]
    X = sm.add_constant(X)
    y = simulated_df["y"]
    # Fit the model to the data
    ols = sm.OLS(y, X).fit()
    # Display summary
    print(ols.summary())
# Define models as a dictionary
models = {
   1: ["x"],
   2: ["x", "x2"],
   3: ["x", "x2", "x3"],
   4: ["x", "x2", "x3", "x4"]
}
# Fit models and display summaries
print("\nModel Summaries:")
for model_num, predictors in models.items():
   print(f"\nModel {model_num}:")
   model_fit(predictors)
```

Model Summaries:

Model 1:

OLS Regression Results

Dep. Variable: R-squared: 0.318 У Model: OLS Adj. R-squared: 0.311 Least Squares F-statistic: Method: 45.60 1.04e-09 Date: Thu, 20 Feb 2025 Prob (F-statistic): -230.83 Time: 22:48:18 Log-Likelihood: No. Observations: 100 AIC: 465.7 Df Residuals: 98 BIC: 470.9 Df Model: 1

(oef						
	,001	std err		t P>	t	[0.025	0.975]
const -1.4	650	0.247	-5.93	7 0.0	000	-1.955	-0.975
x 1.9	494	0.289	6.75	2 0.0	000	1.376	2.522
=======================================	======		======	=======			======
Omnibus:		52.	788 Du	rbin-Watso	on:		1.972
<pre>Prob(Omnibus):</pre>		0.	000 Ja	rque-Bera	(JB):		149.089
Skew:		-1.	953 Pr	ob(JB):			4.22e-33
Kurtosis:		7.	530 Cc	nd. No.			1.20
=======================================			======	=======			======

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified

Model 2:

OLS Regression Results

=========		========			=========		
Dep. Variable	:		У	R-sq	uared:		0.887
Model:			OLS	Adj.	R-squared:		0.884
Method:		Least Sqı	ıares	F-st	atistic:		379.5
Date:		Thu, 20 Feb	2025	Prob	(F-statistic)	:	1.36e-46
Time:		22:4	18:18	Log-	Likelihood:		-141.06
No. Observati	ons:		100	AIC:			288.1
Df Residuals:			97	BIC:			295.9
Df Model:			2				
Covariance Ty	pe:	nonro	bust				
=========	======	========					
	coef	std err		t	P> t	[0.025	0.975]
const	-0.0728	0.119	-0	 .611	0.543	-0.309	0.164
x	0.9663	0.126	7	. 647	0.000	0.715	1.217
x2	-2.0047	0.091	-22	.072	0.000	-2.185	-1.824
Omnibus:	======		: L.338	 Durb	========= in-Watson:		2.197
Prob(Omnibus)	:		0.512		ue-Bera (JB):		0.814
Skew:	-		0.119	-	(JB):		0.666
Kurtosis:			3.372		. No.		2.23
========				=====	=========		=

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

Model 3:

OLS Regression Results

ULS Regression Results								
=========		=======	=====	=====	=========	=======		
Dep. Variable	:		У	R-sq	uared:		0.888	
Model:			OLS	Adj.	R-squared:		0.885	
Method:		Least Sq	uares	F-st	atistic:		253.8	
Date:		Thu, 20 Feb	2025	Prob	(F-statistic):		1.70e-45	
Time:		22:	48:18	Log-	Likelihood:		-140.47	
No. Observation	ons:		100	AIC:			288.9	
Df Residuals:			96	BIC:			299.4	
Df Model:			3					
Covariance Typ	ne:	nonr	obust					
==========	, =======	=========	======			=======		
	coef	std err		t	P> t	[0.025	0.975]	
const	-0.0572	0.120	-0	 .477	0.635	-0.295	0.181	
x	1.1146	0.187	5	.945	0.000	0.742	1.487	
x2	-2.0471	0.099	-20	.673	0.000	-2.244	-1.851	
х3	-0.0643	0.060	-1	.070	0.287			
Omnibus:		=======	====== 0.845	===== Durb	========= in-Watson:	======	2.199	
Prob(Omnibus)	:		0.655	Jarq	ue-Bera (JB):		0.392	
Skew:			0.052	-	(JB):		0.822	
Kurtosis:			3.289		. No.		5.95	

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

Model 4:

OLS Regression Results

			=======================================
Dep. Variable:	у	R-squared:	0.894
Model:	OLS	Adj. R-squared:	0.890
Method:	Least Squares	F-statistic:	200.2
Date:	Thu, 20 Feb 2025	Prob (F-statistic):	2.22e-45
Time:	22:48:18	Log-Likelihood:	-137.74
No. Observations:	100	AIC:	285.5
Df Residuals:	95	BIC:	298.5
Df Model:	4		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
const	0.1008	0.136	0.743	0.460	-0.169	0.370
x	0.9050	0.205	4.423	0.000	0.499	1.311
x2	-2.5059	0.221	-11.336	0.000	-2.945	-2.067
x3	0.0338	0.073	0.466	0.642	-0.110	0.178
x4	0.1042	0.045	2.309	0.023	0.015	0.194
========			.=======		========	
Omnibus:		2.	476 Durbin	n-Watson:		2.163
Prob(Omnibu	ıs):	0.	290 Jarque	e-Bera (JB):		2.097
Skew:		0.	118 Prob(.	JB):		0.351
Kurtosis:		3.	669 Cond.	No.		19.9
========			========			=======

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

```
# Display the results
print("model i:\n")
model_fit(["x"])
```

model i:

OLS Regression Results

Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type	3:	y OLS Least Squares Thu, 20 Feb 2025 22:48:18 100 98 1 nonrobust	Adj. F-st Prob Log- AIC: BIC:	uared: R-squared: atistic: (F-statistic): Likelihood:		0.318 0.311 45.60 1.04e-09 -230.83 465.7 470.9
	coef	std err	====== t 	P> t	[0.025	0.975]
	1.4650 1.9494		-5.937 6.752	0.000	-1.955 1.376	-0.975 2.522
Omnibus: Prob(Omnibus): Skew:		52.788 0.000 -1.953	Jarq	======== in-Watson: ue-Bera (JB): (JB):		1.972 149.089 4.22e-33

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

```
print("\nmodel ii:\n")
model_fit(["x", "x2"])
```

model ii:

OLS Regression Results

===========		========	=====		=========	=======	
Dep. Variable:			У	R-sq	uared:		0.887
Model:			OLS	Adj.	R-squared:		0.884
Method:		Least Squ	ares	F-st	atistic:		379.5
Date:		Thu, 20 Feb	2025	Prob	(F-statistic)	:	1.36e-46
Time:		22:4	8:18	Log-	Likelihood:		-141.06
No. Observation	ns:		100	AIC:			288.1
Df Residuals:			97	BIC:			295.9
Df Model:			2				
Covariance Type	e:	nonro	bust				
===========		========			========	=======	
	coef	std err		t	P> t	[0.025	0.975]
const	 -0.0728	0.119	-0	 .611	0.543	-0.309	0.164
x	0.9663	0.126	7	.647	0.000	0.715	1.217
x2	-2.0047	0.091	-22	.072	0.000	-2.185	-1.824
Omnibus:	======	 1	.338	===== Durb	========= in-Watson:	=======	2.197
<pre>Prob(Omnibus):</pre>		C	.512	Jarq	ue-Bera (JB):		0.814
Skew:		C	.119	Prob	(JB):		0.666
Kurtosis:		3	.372	Cond	. No.		2.23

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

```
print("\nmodel iii:\n")
model_fit(["x", "x2", "x3"])
```

model iii:

OLS Regression Results

=======================================									
Dep. Variabl	le:		У	R-sq	uared:		0.888		
Model:		OLS		Adj. R-squared:			0.885		
Method:		Least Squares		F-statistic:			253.8		
Date:		Thu, 20 Feb 2025		Prob (F-statistic):		:	1.70e-45		
Time:		22:48:18		Log-Likelihood:			-140.47		
No. Observations:		100		AIC:		288.9			
Df Residuals:			96	BIC:			299.4		
Df Model:			3						
Covariance Type:		nonrobust							
=======================================									
	coei	std err		t	P> t	[0.025	0.975]		
const	-0.0572	2 0.120	-0	 .477	0.635	-0.295	0.181		
x	1.1146	0.187	5	.945	0.000	0.742	1.487		
x2	-2.0471	0.099	-20	.673	0.000	-2.244	-1.851		
х3	-0.0643	0.060	-1	.070	0.287	-0.184	0.055		
======================================			:=====).845	Durb	======== in-Watson:	=======	2.199		
Prob(Omnibus	z)·		0.655		ue-Bera (JB):		0.392		
Skew:			0.052	Prob(JB):			0.822		
Kurtosis:			3.289		. No.		5.95		
			. 200	55114			0.00		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

```
print("\nmodel iv:\n")
model_fit(["x", "x2", "x3", "x4"])
```

model iv:

OLS Regression Results

Dep. Variable:	у	R-squared:	0.894
Model:	OLS	Adj. R-squared:	0.890
Method:	Least Squares	F-statistic:	200.2
Date:	Thu. 20 Feb 2025	Prob (F-statistic):	2.22e-45

Time: No. Observations: Df Residuals: Df Model: Covariance Type:		22:48	100 AIC: 95 BIC: 4			-137.74 285.5 298.5	
=======	coef	std err	t	P> t	[0.025	0.975]	
const	0.1008	0.136	0.743	0.460	-0.169	0.370	
x	0.9050	0.205	4.423	0.000	0.499	1.311	
x2	-2.5059	0.221	-11.336	0.000	-2.945	-2.067	
x3	0.0338	0.073	0.466	0.642	-0.110	0.178	
x4	0.1042	0.045	2.309	0.023	0.015	0.194	
Omnibus:		2.	476 Durbin	 ı-Watson:		2.163	
Prob(Omnibus):		0.290 Jarque-Bera (JB):				2.097	
Skew:		0.	118 Prob(J	IB):		0.351	
Kurtosis:		3.	669 Cond.	No.		19.9	

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specific

The p-values in each models show that, while linear () and quadratic (2) associations are statistically significant at 1% significance level, predictors of higher degree (3, 4) are not, which agree with the conclusions drawn based on the CV results.

Part 3

3.a.

```
directory = r"C:\Users\clari\OneDrive\Documents\Machine Learning\ps3"
boston_path = os.path.join(directory, "Boston.csv")
boston_df = pd.read_csv(boston_path)
print(boston_df.dtypes)
print(boston_df.shape)
boston_df.head()
X = boston_df.drop("CRIM", axis=1)
y = boston_df["CRIM"]
```

CRIM float64

```
ZN
           float64
INDUS
           float64
CHAS
           float64
NOX
           float64
RM
           float64
AGE
           float64
DIS
           float64
RAD
           float64
TAX
           float64
PTRATIO
           float64
           float64
LSTAT
           float64
MDEV
           float64
dtype: object
(506, 14)
# Set up cross-validation
kfold = KFold(n_splits=5, shuffle=True, random_state=24)
# Make RMSE scorer
def rmse(y_true, y_pred):
    return sqrt(mean_squared_error(y_true, y_pred))
rmse_scorer = make_scorer(
    rmse,
    greater_is_better=False
```

BEST SUBSET

```
# Create a custom RMSE scorer
def rmse(y_true, y_pred):
    return np.sqrt(mean_squared_error(y_true, y_pred))

rmse_scorer = make_scorer(rmse, greater_is_better=False)

# Set up cross-validation
kfcv = KFold(n_splits=5, shuffle=True, random_state=24)

# Set up the model
model = LinearRegression()

# Start timing
```

```
start_time = time.time()
# Perform best subset selection
efs = EFS(
   model,
   min features=1,
   max_features=13, # Boston dataset has 13 features excluding CRIM
   scoring=rmse_scorer,
   cv=kfcv,
   n_jobs=-1
)
efs.fit(X, y)
# Find the best subset
best_subset = max(efs.subsets_.values(), key=lambda x: x["avg_score"])
best_features = best_subset["feature_names"]
best_rmse = -best_subset["avg_score"]
# End timing
end_time = time.time()
print(f"Best Model Features: {best_features}")
print(f"Best Model RMSE: {round(best_rmse, 5)}")
print(f"Execution Time: {round(end_time - start_time, 2)} seconds")
Features: 1/8191Features: 2/8191Features: 3/8191Features: 4/8191Features: 5/8191Features
Best Model Features: ('ZN', 'INDUS', 'CHAS', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT')
Best Model RMSE: 6.04235
Execution Time: 69.92 seconds
print("\nBest Model from Best Subset Approach:")
print(best_features)
print(f"\nBest Model RMSE: {round(best_rmse, 5)}")
end_time = time.time()
print(f"Running Time: {end_time - start_time}")
Best Model from Best Subset Approach:
('ZN', 'INDUS', 'CHAS', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT')
```

Best Model RMSE: 6.04235

Running Time: 71.97150897979736

FORWARD STEPWISE

```
from sklearn.feature_selection import SequentialFeatureSelector
start_time = time.time()
sfs_forward = SequentialFeatureSelector(
    ols,
    n_features_to_select="auto",
    direction="forward",
    scoring=rmse_scorer, # Use RMSE instead of R2
    cv=kfcv
)
# Fit the feature selector
sfs_forward.fit(X, y)
# Extract selected features
forward_selected_features = X.columns[sfs_forward.get_support()].tolist()
print("\nThe best model from the Forward Stepwise Approach is:")
print(forward_selected_features)
best_model_rmse = -cross_val_score(
    ols,
    X[forward_selected_features],
    cv=kfcv,
    scoring=rmse_scorer
).mean()
print(f"\nThe best model RMSE: {round(best_model_rmse, 5)}")
end_time = time.time()
print(f"Running Time: {end_time - start_time}")
The best model from the Forward Stepwise Approach is:
['ZN', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT']
The best model RMSE: 6.04652
Running Time: 0.9751005172729492
```

BACKWARD STEPWISE

```
# Backward Stepwise
start_time = time.time()
sfs_backward = SequentialFeatureSelector(
ols,
n_features_to_select="auto",
direction="backward",
scoring=rmse_scorer,
cv=kfcv
# Fit the feature selector
sfs_backward.fit(X, y)
# Extract selected features
backward_selected_features = X.columns[sfs_backward.get_support()].tolist()
print("\nBest Model from Backward Stepwise Approach:")
print(backward_selected_features)
# Evaluate RMSE of the best model
best_model_rmse = -cross_val_score(
X[backward_selected_features],
у,
cv=kfcv,
scoring=rmse_scorer
).mean()
print(f"\nBest Model RMSE: {round(best_model_rmse, 5)}")
end_time = time.time()
print(f"Running Time: {end_time - start_time}")
Best Model from Backward Stepwise Approach:
['ZN', 'INDUS', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT']
Best Model RMSE: 6.04247
```

RMSE: Based on the RMSE, the best predictive power to wrost is: Best Subset > Backward Stepwise > Forward Stepwise. The variation in predictive power is liekly influenced by which predictors are included or excluded. Forward Stepwise model excludes predictors like "INDUS," while both stepwise approaches exclude "CHAS."

Running Time: 1.0645697116851807

Running time: Because of how thorough the Best Subset approach is, it takes much longer to run—a whole minute more than the stepwise methods, even with only 13 predictors. The Stepwise models are more efficient, making them preferable when resources or time are limited.

3.b.

```
boston_df = boston_df.reset_index(drop=True)
X = boston_df.drop("CRIM", axis=1)
y = boston_df["CRIM"].copy()
```

Best Subset Model

```
best_subset_features = ['ZN', 'INDUS', 'CHAS', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT']
X_bs = sm.add_constant(boston_df[best_subset_features])
model_bs = sm.OLS(y, X_bs).fit()

# Forward Stepwise Model
forward_features = ['ZN', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT']
X_fs = sm.add_constant(boston_df[forward_features])
model_fs = sm.OLS(y, X_fs).fit()

# Backward Stepwise Model
backward_features = ['ZN', 'INDUS', 'NOX', 'DIS', 'RAD', 'B', 'LSTAT']
X_bw = sm.add_constant(boston_df[backward_features])
model_bw = sm.OLS(y, X_bw).fit()
```

5fCV

Evaluating

```
five_fcv("Best Subset", best_subset_features)
five_fcv("Forward Stepwise", forward_features)
five_fcv("Backward Stepwise", backward_features)
```

Best Subset Approach:

- 5FCV MSE: 42.92 - 5FCV RMSE: 6.04

Forward Stepwise Approach:

- 5FCV MSE: 43.04 - 5FCV RMSE: 6.05

Backward Stepwise Approach:

- 5FCV MSE: 42.97 - 5FCV RMSE: 6.04

AIC- asked phoenix how to do AIC calculation in python

```
print("AIC Values")
print(f"Best Subset: {model_bs.aic:.1f}")
print(f"Forward Stepwise: {model_fs.aic:.1f}")
print(f"Backward Stepwise: {model_bw.aic:.1f}")

print(" Key Findings")
print("Best model by 5FCV: Best Subset (lowest MSE)")
print("Best model by AIC: Forward Stepwise (lowest AIC)")
```

AIC Values

Best Subset: 3341.0 Forward Stepwise: 3339.3 Backward Stepwise: 3340.2 Key Findings

Best model by 5FCV: Best Subset (lowest MSE)
Best model by AIC: Forward Stepwise (lowest AIC)

The Best Subset approach had the lowest 5fCV MSE (42.92), demonstrating superior predictive performance compared to Forward Stepwise (43.04) and Backward Stepwise (42.97) methods. But the Forward Stepwise model showed the best AIC value (3339.3), indicating better balance between model fit and complexity.

According to the book (after asking Phoenix: why would one model be better in terms of MSE but another at AIC? what is the reasoning behind it?), a model may have a lower MSE due to its ability to fit the training data very well, but it may not necessarily perform better on unseen data. Meanwhile, the AIC adjusts for this by incorporating a penalty for the number of parameters to avoid overfitting, thus sometimes selecting simpler models that generalize better despite potentially higher MSE on training data.

Why they select different models: Because the Best Subset method searched through all predictor combinations (8 features), it is liekly that it has a stronger predictive performance, though at greater computational cost. Forward Stepwise builds models sequentially by adding features, potentially missing important combos, while Backward Stepwise removes features from a full model, preserving more predictors.

5 FCV is preferred over training error since it prevents overfitting through repeated data splitting, providing more reliable estimates of real-world performance. AIC remains valuable for model comparison with limited data as it penalizes complexity (2k/n penalty term). The metrics may not align because 5 FCV focuses on predictive accuracy while AIC estimates information loss relative to the true data-generating process.

3.c. NO, the selected models don't use all 13 features. Excluded predictors were AGE, TAX, and PTRATIO.

This is beceeause the model tries to reduce overvitting by removing variables that don't add much to the model in terms of predictive power, which in itself helps us avoid the problem of multicollinearity. This thus makes the model outcome easier to interpret.