

## ISyE 3044 — Final Exam

Fall 2013

NAME: SOLUTIONS

	Deductions
Problem 1:	
Problem 2:	
Problem 3:	
Problem 4:	
Problem 5:	
Problem 6:	
SCORE:	

The test is designed for 90 minutes. No books and notes are allowed. You can use only the supplied formula sheet and tables, as well as a scientific calculator with single-variable statistical functions. *Show all your work to receive credit.* Signing your name implies strict compliance with the institute's honor code.

1. [6 points] Consider the density  $f(x) = p(1-x)^{p-1}$ ,  $0 < x < 1$ , where  $p > 0$  is an unknown (shape) parameter.

- (a) Derive the maximum likelihood estimator of  $p$  based on a sample  $X_1, X_2, \dots, X_n$ .

$$\hat{p} = -n / \sum_{i=1}^n \ln(1-x_i)$$

ANSWER: \_\_\_\_\_

$$L(p) = \prod_{i=1}^n p(1-x_i)^{p-1} = p^n \left[ \prod_{i=1}^n (1-x_i) \right]^{p-1}$$

$$\ell(p) = n \ln p + (p-1) \sum_{i=1}^n \ln(1-x_i)$$

$$\ell'(p) = \frac{n}{p} + \sum_{i=1}^n \ln(1-x_i) = 0 \Leftrightarrow p = - \frac{n}{\sum_{i=1}^n \ln(1-x_i)}$$

$$\ell''(p) = -\frac{n}{p^2} < 0.$$

(b) Consider the following observations: .03, .64, .34, .25, .53

Use the Kolmogorov-Smirnov test with type I error .15 to assess the fit of the model with  $p = 2$ .

ANSWER: Failure to reject hypothesis

$$F(x) = 1 - (1-x)^2$$

$X_{(i)}$	.03	.25	.34	.53	.64
$F(X_{(i)})$	.059	.438	.564	.779	.870
$\frac{i}{5} - F(X_{(i)})$	.141	—	.036	.021	.130
$F(X_{(i)}) - \frac{i-1}{5}$	.059	.238	.164	.179	.070

$$D_5 = \max \{ .141, .238 \} = .238$$

$$\text{Adjusted test statistic} = \left( \sqrt{5} + .12 + \frac{.11}{\sqrt{5}} \right) (.238) = .571$$

Since the critical value at the 5% level is 1.139, we fail to reject the null hypothesis.

2. [4 points] The random variable  $X$  has the density function

$$f(x) = \begin{cases} e^{2x} & x < 0 \\ e^{-2x} & x \geq 0. \end{cases}$$

(a) Compute the cumulative distribution function of  $X$ .

$$F(x) = \begin{cases} \frac{1}{2} e^{2x} & x < 0 \\ 1 - \frac{1}{2} e^{-2x} & x \geq 0 \end{cases}$$

ANSWER: \_\_\_\_\_

(b) Use the inverse-transform method to obtain formulas for generating realizations of  $X$ .

$$X = \begin{cases} \frac{1}{2} \ln(2U) & 0 \leq U < \frac{1}{2} \\ -\frac{1}{2} \ln[2(1-U)] & \frac{1}{2} \leq U \leq 1 \end{cases}$$

ANSWER: \_\_\_\_\_

Solve  $F(X) = U$  for  $X$ :

4. [5 points] Consider the following list pseudo-random numbers (read from left to right):  
 .13, .06, .25, .34, .89, .72, .61, .99, .58, .15, .64, .37

- (a) Use the first two values to generate two independent realizations from the standard normal distribution.  $Z_1 = [-2\ln(.13)]^{1/2} \cos[2\pi(.06)] = 1.88$

ANSWER:  $Z_2 = [-2\ln(.13)]^{1/2} \sin[2\pi(.06)] = 0.74$

- (b) Generate two realizations from the normal distribution with mean  $-1$  and variance  $4$ .

ANSWER:  $-1 + 2(1.88) = 2.76$  ;  $-1 + 2(0.74) = 0.48$

- (c) Let's use the first two numbers from the list. Name the distribution of the realization

$$\left\lceil \frac{\ln(.87)}{\ln(.75)} \right\rceil + \left\lceil \frac{\ln(.94)}{\ln(.75)} \right\rceil$$

and specify the parameters.

ANSWER: negative binomial with  $r=2$  and  $p=0.25$

- (d) Use the third number from the list at the top and an approximation method to generate a realization from the standard normal distribution.

ANSWER:  $[\frac{.25^{.135}}{.75^{.135}}] / .1475 = -0.67$

- (e) Use the standard normal sample from part (d) to generate a realization from the Poisson distribution with mean  $36$ .

ANSWER:  $\lceil 36 + (-0.67)\sqrt{36} - 0.5 \rceil = 32$

5. [6 points] A simulation project was used to compare two configurations for a production line by means of the mean monthly cost (in thousands of dollars). Ten independent replications for each configuration produced the following monthly averages ( $Y_{ij}$  is the average from replication  $j$  for configuration  $i$ ):

Run	1	2	3	4	5	6	7	8	9	10
Config. 1 ( $Y_{1j}$ )	52	42	50.1	49.9	42.5	49.6	45.5	46.9	55.5	44.5
Config. 2 ( $Y_{2j}$ )	51.3	35.3	48.2	47.8	36	47.4	40.8	43.1	56.8	39.2

- (a) The experiments used common random numbers (CRN) in an attempt to induce (circle the correct term)

positive negative

correlation between the  $Y_{1j}$  and  $Y_{2j}$ . Evaluate the effectiveness of CRN by computing the following estimate for this correlation:

$$\frac{\sum_{j=1}^{10} (Y_{1j} - \bar{Y}_1)(Y_{2j} - \bar{Y}_2)}{9S_1S_2} = \frac{\sum_{j=1}^{10} Y_{1j}Y_{2j} - 10\bar{Y}_1\bar{Y}_2}{9S_1S_2},$$

where  $\bar{Y}_1$  and  $\bar{Y}_2$  are the sample means of the  $Y_{1j}$  and  $Y_{2j}$ , and  $S_1$  and  $S_2$  are the sample standard deviations of the  $Y_{1j}$  and  $Y_{2j}$ , respectively.

ANSWER: estimate = 0.99998  $\approx$  1  $\rightarrow$  very effective!

- (b) Conduct a paired- $t$  test with type I error .05 to assess the equality of the mean monthly costs under the two configurations.

ANSWER: There is strong evidence that  $\mu_1 > \mu_2$ .

$$D_j = Y_{1j} - Y_{2j}$$

$$\bar{D} = 3.26$$

$$S = 2.59$$

$$t_{9, .025} = 2.26$$

$$\left. \begin{array}{l} S = 2.59 \\ t_{9, .025} = 2.26 \end{array} \right\} \text{half length of 95\% CI} = 2.26 \frac{2.59}{\sqrt{10}} = 1.85$$

$$\text{CI: } 3.26 \pm 1.85 = (1.41, 5.11)$$

## 6. [5 points] Short Questions.

- (a) The “traditional” negative binomial distribution models the sum of  $r$  i.i.d. geometric random variables; therefore  $r$  is a nonnegative integer. There is a negative binomial distribution with a non-integer parameter  $r$ .

☒ True ☐ False

- (b) The Kolmogorov-Smirnov test can be used to assess the fit of discrete distributions.

True ☒ False

- (c) Use the uniform random number .20 to generate a realization (sample) from the geometric distribution that counts the number of failures before the first success in sequence of Bernoulli( $p = 0.75$ ) trials.

ANSWER: 0

$$\left\lceil \frac{\ln(1-.20)}{\ln(1-.75)} \right\rceil - 1 = 0$$

- (d) If a simulation's run length is substantially long, Simio can report confidence intervals for a steady-state means (e.g., the mean size of a buffer). What method is it using?

ANSWER: batch means

- (e) The sequential procedure of Kim and Nelson (2001), implemented in Simio, uses an indifference zone  $\delta > 0$  and an error  $\alpha \in (0, 1)$  to compare the means  $\mu_i$  of  $m$  systems. Suppose that smaller is better and that (unknown to us)  $\mu_1 < \mu_2 < \dots < \mu_m$ . The procedure makes the following guarantee: the probability of selecting system 1 as the “best” when  $\mu_2 - \mu_1 \geq \delta$  is  $\geq 1 - \alpha$ .

☒ True ☐ False