# An Efficient Algorithm for PAC Mode Estimation

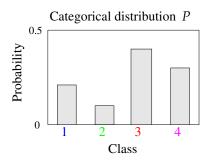
Shubham Anand Jain <sup>2</sup> Rohan Shah <sup>1</sup> Sanit Gupta <sup>1</sup>
Denil Mehta <sup>1</sup> Inderjeet Nair <sup>1</sup> Jian Vora <sup>2</sup> Sushil Khyalia <sup>1</sup>
Sourav Das <sup>3</sup> Vinay J. Ribeiro <sup>1</sup> Shivaram Kalyanakrishnan <sup>1</sup>

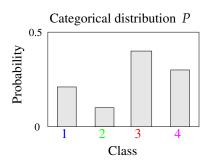
<sup>1</sup>Indian Institute of Technology Bombay

<sup>2</sup>Stanford University

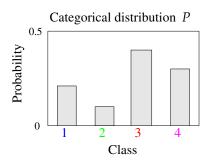
<sup>3</sup>University of Illinois Urbana-Champaign

August 2023

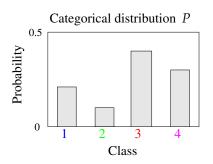




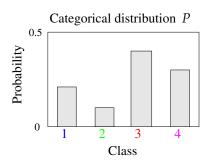
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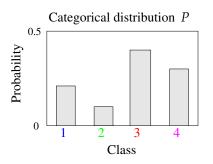
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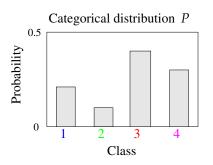
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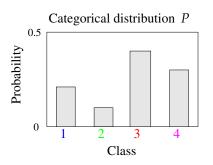
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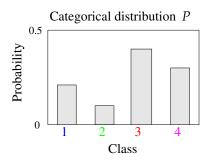
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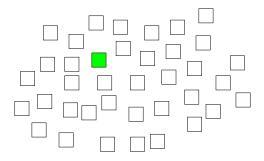
Sequence of samples from  $\mathcal{P}$ : 3, 1, 3, 4, 2, 1, 3, ...

When can we confidently stop sampling and declare the mode?

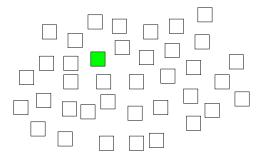
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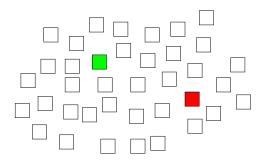
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- Node returns "Result = ■".



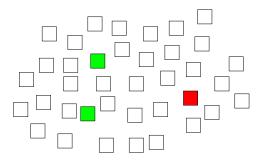
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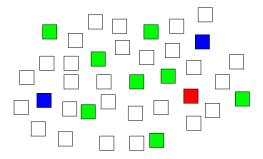
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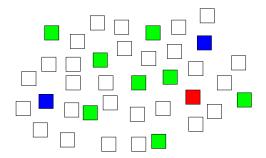
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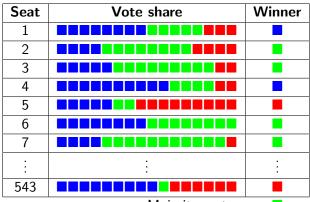


- Assume that most frequently returned result is correct.
- Determine most frequently returned result with sufficient confidence, while minimising number of queries.

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# **Election Forecasting**



Majority party:

### **Election Forecasting**

Seat	Vote share	Winner
1		
2		
3		
4		
5		
6		
7		
:	:	:
543		
Majority party:		

- Confidently identify winner of each seat by sampling.
- Aggregate wins to predict overall winner.

### Overview of Talk

- 1. PAC Formulation
- 2. Solutions for K = 2 classes
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- Categorical distribution  $\mathcal{P} = (p, K)$ :
  - $K \geq 2$  classes,
  - probability vector  $p = \{p_1, p_2, \dots, p_K\}.$

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- Require algorithm  $\mathcal L$  that
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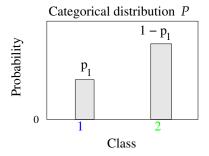
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• **Prefer**  $\delta$ -correct algorithm with low sample complexity.

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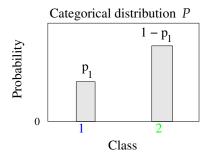
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# Solving the K = 2 Special Case



• Have to ascertain whether  $p_1 > 0.5$  or  $p_1 < 0.5$  based on samples 2, 1, 2, 1, 1, 2, 2, 1, 2, 1, 1 ...

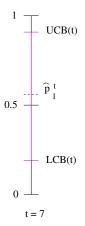
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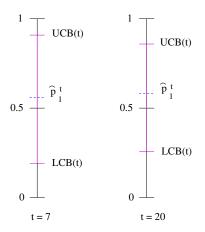


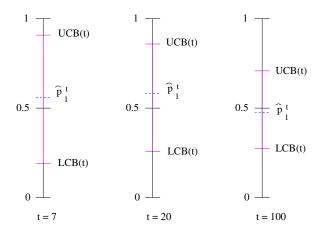
- Have to ascertain whether  $p_1 > 0.5$  or  $p_1 < 0.5$  based on samples 2, 1, 2, 1, 1, 2, 2, 1, 2, 1, 1 . . .
- Maintain lower confidence bound LCB and upper confidence bound UCB on  $p_1$ , stop if LCB > 0.5 or UCB < 0.5.

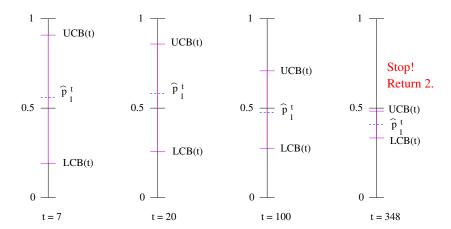
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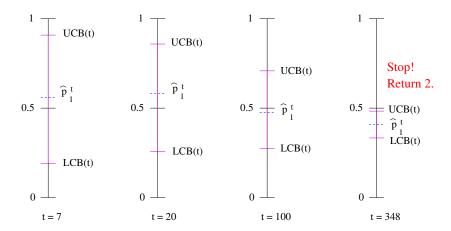








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• Note that the stopping time is random.

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### Some Baselines

Using Hoeffding's Inequality.

$$LCB(t), UCB(t) = \hat{\rho}_1^t \pm \beta(t, \delta),$$
 
$$\beta(t, \delta) = \sqrt{\frac{1}{2t} \ln\left(\frac{kt^{\alpha}}{\delta}\right)}.$$

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• Using Empirical Bernstein bound ( $A_1$  algorithm (Shah et al., 2020)).

$$\begin{aligned} \textit{LCB}(t), \textit{UCB}(t) &= \hat{\rho}_1^t \pm \beta(t, \delta), \\ \beta(t, \delta) &= \sqrt{\frac{2V^t \ln(4t^2/\delta)}{t}} + \frac{7\ln(4t^2/\delta)}{3(t-1)}, \end{aligned}$$

where  $V^t$  is the empirical variance.



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• "Self-normalised" inequality to avoid naïve union bound over t. KL-SN, Garivier (2013).

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• Introduced by Waudby-Smith and Ramdas (2020).

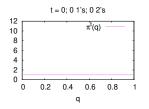
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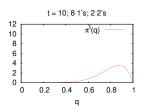
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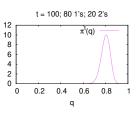
$$\pi^{t}(q) = \frac{\pi^{t-1}(q) \cdot (q)^{\mathbf{1}[x^{t}=1]} \cdot (1-q)^{\mathbf{1}[x^{t}=2]}}{\int_{\rho=0}^{1} \pi^{t-1}(\rho) \cdot (\rho)^{\mathbf{1}[x^{t}=1]} \cdot (1-\rho)^{\mathbf{1}[x^{t}=2]} d\rho}.$$

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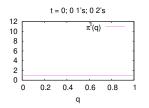
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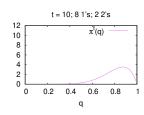
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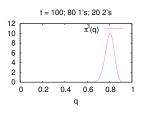
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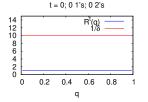
- Proof by establishing that  $R^t(p_1)$  is a martingale, then applying Ville's inequality on nonnegative supermartingales.
- Method more generally applicable, even for estimating multiple parameters simultaneously.

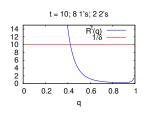
# PPR Martingale Confidence Sequences: Illustration

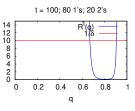












# PPR-Bernoulli Stopping Rule

• For our K=2 case, the belief distribution  $\pi^t$  is Beta.

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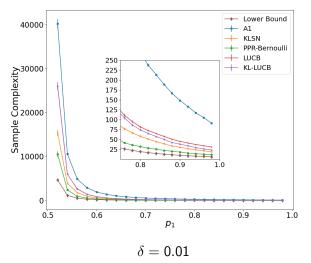
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- Let  $\operatorname{first}(t) \in \{1,2\}$  denote the more frequent class in the data, and  $\operatorname{second}(t) \in \{1,2\}$  the other class. Let  $s_i^t$  be the number of occurrences of class  $i \in \{1,2\}$ . We obtain this stopping rule.

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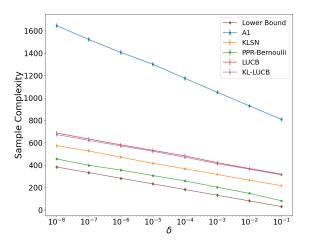
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**PPR-Bernoulli**: Stop, declare first(t) as mode iff Beta  $\left(\frac{1}{2}; s_{\mathsf{first}(t)}^t + 1, s_{\mathsf{second}(t)}^t + 1\right) \leq \delta$ .

# Comparison of Stopping Rules for K = 2



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$$p_1 = 0.65$$

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#### Overview of Talk

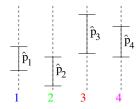
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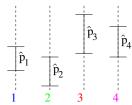
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#### 1-versus-rest approach



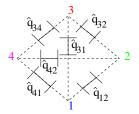
- For each class i, estimate underlying probability p<sub>i</sub>.
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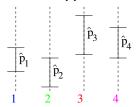
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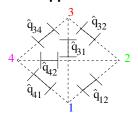
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1-versus-1 appears to waste samples, but ...

# Empirical Comparison: 1vr and 1v1

#### Sample complexity

Distribution	K	Туре	$\mathcal{A}_1$	KL-SN	PPR
$\mathcal{P}_1$ : .5, .25 $ imes$ 2	3	1vr 1v1	$1344{\pm}20$ $1158{\pm}19$	418±14 346±13	$262{\pm}12$ <b>218</b> $\pm$ <b>11</b>
$P_2$ : .4, .2 × 3	4	1vr 1v1	$^{1919\pm29}_{1516\pm24}$	$632{\pm}18$ $468{\pm}15$	$397{\pm}15$ <b>298</b> ${\pm}13$
$\mathcal{P}_3$ : .2, .1 $ imes$ 8	9	1vr 1v1	5082±51 3340±43	1900±42 1138±31	1201±29 <b>789</b> ± <b>28</b>
$\mathcal{P}_4$ : .1, .05 $ imes$ 18	19	1vr 1v1	12015±129 7352±88	$4686{\pm}81$ 2554 ${\pm}57$	$2850{\pm}55$ $1840{\pm}53$
$\mathcal{P}_5$ : .35, .33, .12, .1 × 2	5	1vr 1v1	155277±2356 117988±2078	63739±2238 47205±1291	38001±1311 <b>33660</b> ± <b>1125</b>
$\mathcal{P}_6$ : .35, .33, .04 × 8	10	1vr 1v1	$\begin{array}{c} 158254{\pm}2442 \\ 121150{\pm}2183 \end{array}$	$66939{\pm}2241$ $49576{\pm}1341$	41963±1330 <b>36693</b> ± <b>1185</b>

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**PPR-1v1**: Stop and declare first(t) as mode iff Beta  $\left(\frac{1}{2}; s_{\mathsf{first}(t)}^t + 1, s_{\mathsf{second}(t)}^t + 1\right) \leq \frac{\delta}{K-1}$ .

- We stop when some class i beats all classes  $j \neq i$ .
- When we stop, clearly i must be first(t).
- If first(t) beats all  $j \neq \text{first}(t)$ , then first(t) must beat second(t), and vice versa).
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• Computationally light compared to PPR-1vr, and other KL-divergence-based rules (which have to perform a numerical computation at each step).

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# Asymptotic Optimality

• Lower bound [Shah et al., 2020]: Any  $\delta$ -correct algorithm requires at least LB( $\mathcal{P}, \delta$ ) samples in expectation, where

$$\mathsf{LB}(\mathcal{P}, \delta) \stackrel{\mathsf{def}}{=} \sup_{\mathcal{P}' : \mathsf{mode}(\mathcal{P}') \neq \mathsf{mode}(\mathcal{P})} \frac{1}{\mathsf{KL}(\mathcal{P}||\mathcal{P}')} \ln \left( \frac{1}{2.4\delta} \right).$$

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• **PPR-1v1 upper bound**: Let  $\tau(\mathcal{P}, \delta)$  be the expected stopping time of PPR-1v1 on  $\mathcal{P}$  when run with mistake probability  $\delta$ . Then,

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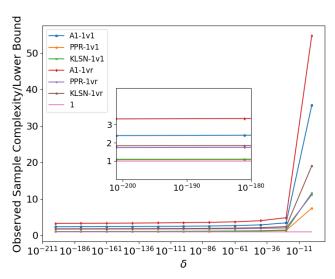
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• Proof uses probability, functions, some results from Garivier and Kaufmann (2016) (who obtain similar result for PAC bandits).

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# Asymptotic Optimality (continued)

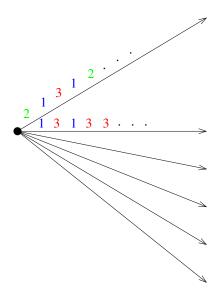


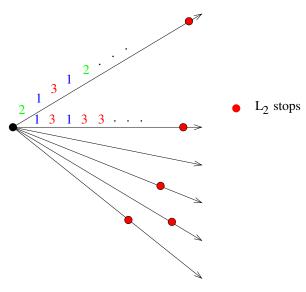
Result from  $\mathcal{P}_3$ 

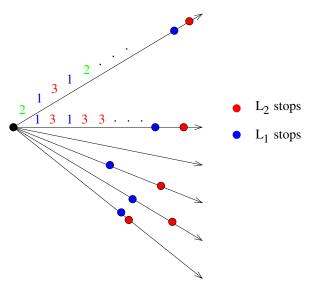


- Let  $X = x^1, x^2, \dots$  be an infinite sequence of samples from  $\mathcal{P}$ .
- For algorithm L, let T(L,X) denote the stopping time of L on X.
- Define algorithm  $L_1$  to uniformly dominate algorithm  $L_2$  if for every sequence X,  $\delta \in (0,1)$ ,

$$T(L_2,X)<\infty \implies T(L_1,X)\leq T(L_2,X).$$







• Hoeffding-1v1 uniformly dominates Hoeffding-1vr.

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- These proofs done by showing

 $L_2$  terminates on  $X \implies L_1$  terminates on X.

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### **Concluding Notes**

- Have addressed PAC mode estimation problem.
- PPR an elegant approach to handle random stopping time.
- 1v1 outperforms 1vr to generalise to  $K \ge 2$  classes.
- PPR-1v1 simple, asymptotically optimal, efficient in practice.
- Flexible to apply in practice, validated on two distinct applications.
- Code base released along with AISTATS 2022 paper.
- Conjecture on PPR-1v1 termination preceding PPR-1vr termination remains unresolved.
- Can PPR help in bandits, too?
- Do you have a practical application of mode estimation?

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