Biostatistics 615 - Statistical Computing

Lecture 13 Random Numbers and Monte Carlo Methods

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Random Numbers

True random numbers

- Truly random, non-deterministic numbers
- Easy to imagine conceptually
- Very hard to generate one or test its randomness
- For example, http://www.random.org generates randomness via atmospheric noise

Pseudo random numbers

- A deterministic sequence of random numbers (or bits) from a seed
- Good random numbers should be very hard to guess the next number just based on the observations.

Usage of random numbers in statistical methods

- Resampling procedure
 - Permutation
 - Boostrapping
- Simulation of data for evaluating a statistical method.
- Stochastic processes
 - Markov-Chain Monte-Carlo (MCMC) methods

Usage of random numbers in other areas

Hashing

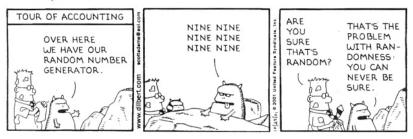
- Good hash function uniformly distribute the keys to the hash space
- Good pseudo-random number generators can replace a good hash function

Cryptography

- Generating pseudo-random numbers given a seed is equivalent to encrypting the seed to a sequence of random bits
- If the pattern of pseudo-random numbers can be predicted, the original seed can also be deciphered.

True random numbers

DILBERT By Scott Adams



- Generate only through physical process
- Hard to generate automatically
- Very hard to provide true randomness

Pseudo-random numbers: Example code

```
#include <iostream>
#include <cstdlib>
int main(int argc, char** argv) {
  int n = (argc > 1) ? atoi(argv[1]) : 1;
  int seed = (argc > 2 ) ? atoi(argv[2]) : 0;
  srand(seed); // set seed -- same seed, same pseudo-random numbers
  for(int i=0; i < n; ++i) {</pre>
    std::cout << (double)rand()/(RAND MAX+1.) << std::endl;</pre>
    // generate value between 0 and 1
  return 0;
```

Pseudo-random numbers : Example run

```
user@host:~/$ ./randExample 3 0
0.242578
0.0134696
0.383139
user@host:~/$ ./randExample 3 0
0.242578
0.0134696
0.383139
user@host:~/$ ./randExample 3 10
7.82637e-05
0.315378
0.556053
```

Properties of pseudo-random numbers

Deterministic given the seed

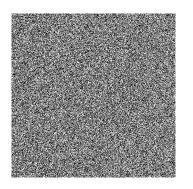
- Given a fixed random seed, the pseudo-random numbers should generate identical sequence of random numbers
- Deterministic feature is useful for debugging a code

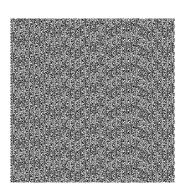
Irregularity and unpredictability without knowing the seed

- Without knowing the seed, the random numbers should be hard to guess
- If you can guess it better than random, it is possible to exploit the weakness to generate random numbers with a skewed distribution.

偏态分布

Good vs. bad random numbers





- Images using true random numbers from random.org vs. rand() function in PHP
- Visible patterns suggest that rand() gives predictable sequence of pseudo-random numbers

Generating uniform random numbers - example in R

Generating uniform random numbers in C++

```
#include <iostream>
#include <boost/random/uniform int.hpp>
#include <boost/random/uniform real.hpp>
#include <boost/random/variate generator.hpp>
#include <boost/random/mersenne twister.hpp>
int main(int argc, char** argv) {
  typedef boost::mt19937 prgType; // Mersenne-twister : a widely used
  prgType rng;
                     // lightweight pseudo-random-number-generator
  boost::uniform int<> six(1,6); // uniform distribution from 1 to 6
  boost::variate generatorrprgTvpe&. boost::uniform int<> > die(rng.six):
  // die maps random numbers from rng to uniform distribution 1..6
        映射
  int x = die();  // generate a random integer between 1 and 6
  std::cout << "Rolled die : " << x << std::endl;</pre>
  boost::uniform real<> uni dist(0,1);
  boost::variate generatorrprgType&, boost::uniform real<> > uni(rng,uni dist);
  double y = uni(); // generate a random number between 0 and 1
  std::cout << "Uniform real : " << y << std::endl;</pre>
  return 0;
```

Running Example

user@host:~/\$./randExample

```
Rolled die : 5
Uniform real : 0.135477

user@host:~/$ ./randExample
Rolled die : 5
Uniform real : 0.135477

The random number does not vary (unlike R)
to get a different sequence, we need to set seed
```

Specifying the seed

```
int main(int argc, char** argv) {
  typedef boost::mt19937 prgType;
  prgType rng;
  if ( argc > 1 )
    rng.seed(atoi(argv[1])); // set seed if argument is specified

boost::uniform_int<> six(1,6);
  // ... same as before
}
```

Running Example

```
user@host:~/$ ./randExample
Rolled die : 5
Uniform real: 0.135477
user@host:~/$ ./randExample 1
Rolled die : 3
Uniform real: 0.997185
user@host:~/$ ./randExample 3
Rolled die: 4
Uniform real: 0.0707249
user@host:~/$ ./randExample 3
Rolled die: 4
Uniform real: 0.0707249
```

If we don't want the reproducibility

Running Example

```
user@host:~/$ ./randExample
Rolled die : 4
Uniform real: 0.367588
user@host:~/$ ./randExample
Rolled die : 5
Uniform real: 0.0984682
user@host:~/$ ./randExample 3
Rolled die: 4
Uniform real: 0.0707249
user@host:~/$ ./randExample 3
Rolled die: 4
Uniform real: 0.0707249
```

Generating random numbers from non-uniform distribution

Sampling from known distribution using R

```
> x <- rnorm(1)  # x is a random number sampled from N(0,1)
> y <- rnorm(1,3,2)  # y is a random number sampled from N(3,2^2)
> z <- rbinom(1,1,0.3) # z is a Bernoulli random number with p=0.3</pre>
```

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What if runif() was the only random number generator we have?

Generating random numbers from non-uniform distribution

Sampling from known distribution using R

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```

What if runif() was the only random number generator we have?

```
If we know the inverse CDF, it is easy to implement
> x <- qnorm(runif(1))  # x follows N(0,1)
> y <- qnorm(runif(1),3,2)  # equivalent to y <- qnorm(runif(1))*2+3
> z <- qbinom(runif(1),1,0.3)  # z is a Bernoulli random number with p=0.3</pre>
```

Inverse transform sampling

- Goal: Sample from a distribution with a known CDF function F.
- Theorem: Let $U \sim Uniform(0,1)$, and $X = F^{-1}(U)$, then $X \sim F$.
- Example: Sample $X \sim Exp(\lambda)$.
 - Density: $f(x) = \lambda e^{-\lambda x}$.
 - CDF: $F(x) = 1 e^{-\lambda x}$.
 - $\bullet \Rightarrow X = -\frac{1}{\lambda} \ln(1 U).$
- Proof:

$$P(X \le x)$$

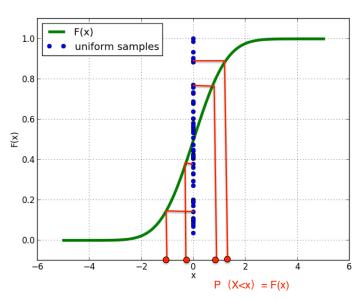
$$= P(F^{-1}(U) \le x)$$

$$= P(U \le F(x))$$

$$= F(x)$$

(http://en.wikipedia.org/wiki/Inverse_transform_sampling)

Inverse transform sampling



(http://kennychowdhary.me/2012/10/

Random number generation in C++

```
#include <iostream>
#include <ctime>
#include <boost/random/normal distribution.hpp>
#include <boost/random/variate generator.hpp>
#include <boost/random/mersenne twister.hpp>
int main(int argc, char** argv) {
 typedef boost::mt19937 prgType;
  prgType rng;
 if ( argc > 1 )
    rng.seed(atoi(argv[1]));
  else.
    rng.seed(std::time(0));
  boost::normal distribution <> norm dist(0,1); // standard normal distribution
  // PRG sampled from standard normal distribution
  boost::variate generator<prgType&, boost::normal distribution<> >
        norm(rng,norm dist);
  double x = norm(); // Generate a random number from the PRG
  std::cout << "Sampled from standard normal distribution : " << x << std::endl;</pre>
  return 0;
```

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Sample from Gaussian distribution

- Inverse CDF ⇒ no closed form inverse CDF function
- Central Limit Theorem ⇒ needs multiple random samples
- The Box-Muller transformation

(http:

//en.wikipedia.org/wiki/Normal_distribution#Generating_values_from_normal_distribution)

Box-Muller Transformation (Box and Muller, 1958)

Let

$$\begin{array}{rcl} U_1,\,U_2 & \sim & Uniform(0,1] \\ R & = & \sqrt{-2\ln U_1} \\ \Theta & = & 2\pi \, U_2 \\ Z_0 & = & R\cos(\Theta) \\ Z_1 & = & R\sin(\Theta) \end{array}$$

Then

$$Z_0, Z_1 \sim N(0,1)$$
, i.i.d.

Where

$$R^2 \sim \chi_2^2 = Exp(\frac{1}{2})$$

(http://en.wikipedia.org/wiki/Box%E2%80%93Muller_transform)



Generating random numbers from complex distributions

Problem

- When the distribution is complex, the inverse CDF may not be easily obtainable
- Need to implement your own function to generate the random numbers

A simple example - mixture of two normal distributions

$$f(x; \mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \alpha) = \alpha f_{\mathcal{N}}(x; \mu_1, \sigma_1^2) + (1 - \alpha) f_{\mathcal{N}}(x; \mu_2, \sigma_2^2)$$

How to generate random numbers from this distribution?

Jian Kang Bio

Sample from Gaussian mixture

Key idea

- lacktriangle Introduce a Bernoulli random variable $w \sim \operatorname{Bernoulli}(lpha)$
- Sample $y \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $z \sim \mathcal{N}(\mu_2, \sigma_2^2)$
- Let x = wy + (1 w)z.

Sample from Gaussian mixture

Key idea

- Introduce a Bernoulli random variable $w \sim \text{Bernoulli}(\alpha)$
- Sample $y \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $z \sim \mathcal{N}(\mu_2, \sigma_2^2)$
- Let x = wy + (1 w)z.

An R implementation

```
w <- rbinom(1,1,alpha)
y <- rnorm(1,mu1,sigma1)
z <- rnorm(1,mu2,sigma2)
x <- w*y + (1-w)*z</pre>
```

Sampling from bivariate normal distribution

Bivariate normal distribution

$$\left(\begin{array}{c} x \\ y \end{array}\right) \sim \mathcal{N} \left(\begin{array}{cc} \mu_x \\ \mu_y \end{array}, \left[\begin{array}{cc} \sigma_x^2 & \sigma_{xy} \\ \sigma_{xy} & \sigma_y^2 \end{array}\right]\right)$$

Sampling from bivariate normal distribution

Bivariate normal distribution

$$\left(\begin{array}{c} x \\ y \end{array}\right) \sim \mathcal{N} \left(\begin{array}{c} \mu_x \\ \mu_y \end{array}, \left[\begin{array}{ccc} \sigma_x^2 & \sigma_{xy} \\ \sigma_{xy} & \sigma_y^2 \end{array}\right]\right)$$

Sampling from bivariate normal distribution

```
x <- rnorm(1,mu.x,sigma.x)
y <- rnorm(1,mu.y,sigma.y) # WRONG. Valid only when sigma.xy = 0</pre>
```

How can we sample from a joint distribution?

Possible approaches

Use known packages

- mvtnorm package provides rmvnorm() function for sampling from a multivariate-normal distribution
- Without using it, how to implement it?

Possible approaches

Use known packages

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Use conditional distribution

$$y|x \sim \mathcal{N}\left(\mu_y + \frac{\sigma_{xy}}{\sigma_x^2}(x - \mu_x), \sigma_y^2\left(1 - \frac{\sigma_{xy}^2}{\sigma_x^2\sigma_y^2}\right)\right)$$

Sampling from multivariate normal distribution

Problem

- Randomly sample from $\mathbf{x} \sim \mathcal{N}(\mathbf{m}, V)$
- ullet The covariance matrix V is positive definite

Sampling from multivariate normal distribution

Problem

- Randomly sample from $\mathbf{x} \sim \mathcal{N}(\mathbf{m}, V)$
- ullet The covariance matrix V is positive definite

Using conditional distribution

- Sample $x_1 \sim \mathcal{N}(m_1, V_{11})$
- Sample $x_2 \sim \mathcal{N}(m_2 + V_{12} V_{22}^{-1}(x_1 m_1), V_{22} V_{12}^T V_{11}^{-1} V_{12})$
- lacktriangleq Repetitively sample x_i from subsequent conditional distributions.

This approach would require excessive amount of computational time

Using Cholesky decomposition for sampling from MVN

Key idea

- If $\mathbf{x} \sim \mathcal{N}(\mathbf{m}, V)$, $A\mathbf{x} \sim \mathcal{N}(A\mathbf{m}, AVA^T)$.
- Sample $\mathbf{z} \sim \mathcal{N}(0, I_n)$ from standard normal distribution
- Find A such that

$$\mathbf{x} = A\mathbf{z} + \mathbf{m} \sim \mathcal{N}(\mathbf{m}, AA^T) = \mathcal{N}(\mathbf{m}, V)$$

ullet Cholesky decomposition $V=U^TU$ generates an example $A=U^T$.

An example R code

```
z <- rnorm(length(m))
U <- chol(V)
x <- m + t(U) %*% z</pre>
```

Summary - Random Number Generation

Random Number Generator

- True Random Number Generator
- Pseudo-random Number Generator

Generating Pseudo random Numbers in C++

- Use built-in rand() for toy examples
- Use boost library (e.g. Mersenne-twister) for more serious stuff
- Use inverse CDF for sampling from a known distribution
- For complex distributions, use generative procedure considering computational efficiency.

Monte-Carlo Methods

Informal definition

- Approximation by random sampling.
- Randomized algorithms to solve deterministic problems approximately.

Goals

- Integration: E[f(x)]
- Probability: $P(X \in A) = E[1_{X \in A}]$
- Bayesian inference: $P(\theta|Data) \propto P(\theta)P(Data|\theta)$
- Especially useful when analytic solution is not available or in high dimensional parameter space.

An example problem

Calculating

$$\theta = \int_0^1 f(x) dx$$

where f(x) is a function with $0 \le f(x) \le 1$

The problem is equivalent to computing E[f(u)] where $u \sim U(0,1)$.



The crude Monte-Carlo method

Algorithm

- Generate u_1, u_2, \cdots, u_B uniformly from U(0,1).
- Take their average to estimate θ

$$\hat{\theta} = \frac{1}{B} \sum_{i=1}^{B} f(u_i)$$

The crude Monte-Carlo method

Algorithm

- Generate u_1, u_2, \cdots, u_B uniformly from U(0, 1).
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$$\hat{\theta} = \frac{1}{B} \sum_{i=1}^{B} f(u_i)$$

Desirable properties of Monte-Carlo methods

- ullet Consistency: estimates converges to true answer as B increases
- Unbiasedness: $E[\hat{\theta}] = \theta$
- Minimal Variance

Analysis of crude Monte-Carlo method

Bias

$$E[\hat{\theta}] = \frac{1}{B} \sum_{i=1}^{B} E[f(u_i)] = \frac{1}{B} \sum_{i=1}^{B} \theta = \theta$$

Analysis of crude Monte-Carlo method

Bias

$$E[\hat{\theta}] = \frac{1}{B} \sum_{i=1}^{B} E[f(u_i)] = \frac{1}{B} \sum_{i=1}^{B} \theta = \theta$$

Variance

$$\operatorname{Var}[\hat{\theta}] = \frac{1}{B} \int_0^1 (f(u) - \theta)^2 du$$
$$= \frac{1}{B} E[f(u)^2] - \frac{\theta^2}{B}$$

Analysis of crude Monte-Carlo method

Bias

$$E[\hat{\theta}] = \frac{1}{B} \sum_{i=1}^{B} E[f(u_i)] = \frac{1}{B} \sum_{i=1}^{B} \theta = \theta$$

Variance

$$Var[\hat{\theta}] = \frac{1}{B} \int_0^1 (f(u) - \theta)^2 du$$
$$= \frac{1}{B} E[f(u)^2] - \frac{\theta^2}{B}$$

Consistency

$$\lim_{B\to\infty} \hat{\theta} = \theta$$



Accept-reject (or hit-and-miss) Monte Carlo method

Algorithm

- ① Define a rectangle R between (0,0) and (1,1)
 - Or more generally, between (x_m, x_M) and (y_m, y_M) .
- ② Set h = 0 (hit), m = 0 (miss).
- **3** Sample a random point $(x, y) \in R$.
- 4 If y < f(x), then increase h. Otherwise, increase m
- \odot Repeat step 3 and 4 for B times
- $\hat{\theta} = \frac{h}{h+m}.$

Analysis of accept-reject Monte Carlo method

Bias

Let u_i, v_i follow $\mathit{U}(0,1)$, then $\Pr(v_i < \mathit{f}(u_i)) = \theta$

$$E[\hat{\theta}] = E\left[\frac{h}{h+m}\right] = \frac{\sum_{i=1}^{B} I(v_i < f(u_i))}{B} = \theta$$

Analysis of accept-reject Monte Carlo method

Bias

Let u_i, v_i follow U(0, 1), then $\Pr(v_i < f(u_i)) = \theta$

$$E[\hat{\theta}] = E\left[\frac{h}{h+m}\right] = \frac{\sum_{i=1}^{B} I(v_i < f(u_i))}{B} = \theta$$

Variance

 $h \sim \text{Binom}(B, \theta)$.

$$\operatorname{Var}[\hat{\theta}] = \frac{\theta(1-\theta)}{B}$$

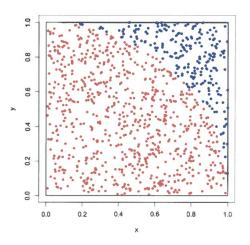
Which method is better?

$$\sigma_{AR}^2 - \sigma_{crude}^2 = \frac{\theta(1-\theta)}{B} - \frac{1}{B}E[f(u)^2] + \frac{\theta^2}{B}$$
$$= \frac{\theta - E[f(u)]^2}{B}$$
$$= \frac{1}{B} \int_0^1 f(u)(1-f(u)) du \ge 0$$

The crude Monte-Carlo method has less variance then accept-rejection method

Example

- Let $X, Y \sim Uniform(0, 1)$
- What is $P(X^2 + Y^2 \ge 1)$?
- Accept-reject Monte Carlo in 1D is equivalent to crude Monte Carlo in 2D.



Summary

- Crude Monte Carlo method
 - Use uniform distribution (or other original generative model) to calculate the integration
 - Every random sample is equally weighted.
 - Straightforward to understand
- Rejection sampling
 - Estimation from discrete count of random variables
 - Larger variance than crude Monte-Carlo method
 - Typically easy to implement
 - Can be used to sample from any shape

General rejection sampling (von Neumann, 1951)

- Goal: sample from a target distribution $\pi(x)$ whose PDF function is known up to a constant $f(x) = c\pi(x)$.
- Rejection sampling:
 - Construct an envelope function g(x) with a constant M such that $Mg(x) \ge f(x)$ for all x.
 - ② Sample x from $g(\cdot)$ and u from Uniform(0,1)
 - **3** Compute the ratio $r = \frac{f(x)}{Mg(x)}$.
 - If u < r, accept x.
 - \bullet Otherwise, discard x.
 - Go back to Step 2.
- Theorem: the accepted sample x follows the target distribution π .

(http://en.wikipedia.org/wiki/Rejection_sampling)



Proof of rejection sampling

$$\begin{array}{ll} P(x \text{ is accepted}) & = & \int P(u < r | X = x) g(x) dx \\ \\ & = & \int \frac{f(x)}{Mg(x)} g(x) dx \\ \\ & = & \int \frac{c\pi(x)}{Mg(x)} g(x) dx \\ \\ & = & \frac{c}{M} \int \pi(x) dx \\ \\ & = & \frac{c}{M} \end{array}$$

Therefore

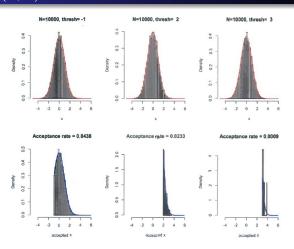
$$\begin{array}{ll} P(X=x|x \text{ is accepted}) & = & \frac{P(X=x,x \text{ is accepted})}{P(x \text{ is accepted})} \\ & = & \frac{\frac{c\pi(x)}{Mg(x)}g(x)}{\frac{c}{M}} \\ & = & \pi(x) \end{array}$$

Example

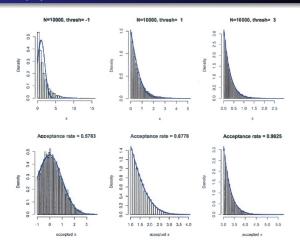
- Target: truncated Gaussian distribution $\pi(x) \propto \phi(x) I_{x>c}$, where $\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$ is the standard Gaussian density function.
- Envelope 1: $g(x) \sim N(0,1)$, i.e., $g(x) = \phi(x)$.
 - $Mg(x) \ge \phi(x)I_{x>c} \Rightarrow M=1$ is ok.
 - $r = \frac{\phi(x)I_{x>c}}{\phi(x)} = I_{x>c} \Rightarrow$ acceptance rate is $1 \Phi(c)$.
- Envelope 2: $g(z) = \lambda e^{-\lambda z}$ and x = z + c.
 - $Mg(z) \ge \phi(z+c)I_{z+c>c} \Rightarrow M\lambda e^{-\lambda z} \ge \frac{1}{\sqrt{2\pi}}e^{-\frac{(z+c)^2}{2}}$ for all z.
 - First let λ be fixed, $M \ge \max_z \frac{1}{\sqrt{2\pi\lambda}} e^{-\frac{(z+c)^2}{2} - \lambda z} = \frac{1}{\sqrt{2\pi\lambda}} e^{-\frac{\lambda^2}{2} - \lambda c}.$
 - How to choose λ to maximize acceptance rate?



Envelope 1: N(0, 1)



Envelope 2: $Exp(\lambda)$



Good envelope function

- Easy to construct.
- Easy to sample from.
- Close to the target function ⇒ low rejection rate.

