Linear Algebra / Statistics Review

Stats 503

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Vectors and Matrices

Vector

$$x = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix} \qquad y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}$$

Matrix

$$A_{[n imes m]} = \left(egin{array}{ccc} A_{11} & \cdots & A_{1m} \ dots & \ddots & dots \ A_{n1} & \cdots & A_{nm} \end{array}
ight)$$

The identity matrix

$$I_{[n \times n]} = \left(\begin{array}{cccc} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & 0 \\ & & \ddots & \\ 0 & \cdots & 0 & 1 \end{array} \right)$$

Matrix multiplication

$$A_{[n imes m]} B_{[m imes k]} = \left(egin{array}{ccc} A_{11} & \cdots & A_{1m} \ dots & \ddots & dots \ A_{n1} & \cdots & A_{nm} \end{array}
ight) \cdot \left(egin{array}{ccc} B_{11} & \cdots & B_{1k} \ dots & \ddots & dots \ B_{m1} & \cdots & B_{mk} \end{array}
ight)$$
 $= \left(egin{array}{ccc} \cdots & \cdots & \cdots & \cdots \ \cdots & \sum_{\ell=1}^m A_{i\ell} B_{\ell j} & \cdots \ \cdots & \cdots & \cdots \end{array}
ight)_{[n imes k]}$

Matrix transpose

$$A' = A_{[n \times m]}^T = \begin{pmatrix} A_{11} & A_{21} & \cdots & A_{n1} \\ A_{12} & A_{22} & \cdots & A_{n2} \\ \vdots & \ddots & \ddots & \vdots \\ A_{1m} & A_{2m} & \cdots & A_{nm} \end{pmatrix}_{[m \times n]}$$

Properties:

$$(A^{T})^{T} = A$$
$$(A+B)^{T} = A^{T} + B^{T}$$
$$(AB)^{T} = B^{T}A^{T}$$

Norm and inner product

- Vector norm (length): $||x|| = \sqrt{x_1^2 + \dots + x_n^2} = \sqrt{x^T x}$
- Inner product: $x^T y = x_1 y_1 + x_2 y_2 + \cdots + x_n y_n$
- Geometric interpretation: $\cos \theta = \frac{x^T y}{\|x\| \cdot \|y\|}$
- x and y are orthogonal if $x^Ty = 0$
- A set of vectors x_1, \ldots, x_m is orthonormal if $x_i^T x_j = 0$ if $i \neq j$ and $||x_i|| = 1$ for all i.

Vector spaces

• Vectors $x_1, ... x_m$ are linearly dependent if there exist scalars $a_1, ... a_m$ such that at least one $a_j \neq 0$ and

$$a_1x_1 + \dots + a_mx_m = 0$$

• A vector space spanned by vectors $x_1, ... x_m$ is

$$\mathscr{X} = \operatorname{span}(x_1, \dots, x_m)$$

= $\{x : x = a_1x_1 + a_2x_2 + \dots + a_mx_m, a_j \in \mathbb{R}\}$

 A minimal set of such vectors is called a basis (all linearly independent)

Basis examples

Give a basis of the space spanned by

$$x_1 = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \quad x_2 = \begin{pmatrix} 1 \\ -1 \\ 1 \end{pmatrix} \quad x_3 = \begin{pmatrix} 1 \\ 0 \\ 1 \end{pmatrix}$$

Give a basis of \mathbb{R}^2 :

Functions of matrices: determinant, trace, norms

Determinant (square matrix): det(A) (complicated formula)

- For diagonal matrices, $det(A) = A_{11} \times \cdots \times A_{nn}$
- det(AB) = det(A) det(B)

Trace (square matrix): $trace(A) = A_{11} + \cdots + A_{nn}$

- trace(A + B) = trace(A) + trace(B)
- trace(AB) = trace(BA)

Matrix norms: many choices. For example, the Frobenius matrix norm

$$||A||_F^2 = \sum_{i=1}^n \sum_{j=1}^m A_{ij}^2 = \operatorname{trace}(AA^T)$$

Matrix inverse (square matrix)

- Defined by $A^{-1}A = AA^{-1} = I$
- If the columns (rows) of A are linearly independent, then A is invertible, i.e. A^{-1} exists; otherwise, A is singular, and det(A) = 0.

Properties

- $(AB)^{-1} = B^{-1}A^{-1}$
- $(A^{-1})^T = (A^T)^{-1} = A^{-T}$
- $(AB)^{-T} = A^{-T}B^{-T}$
- $\det(A^{-1}) = (\det(A))^{-1}$

Orthogonal matrix: $A^{-1} = A^{T}$. Orthogonal matrices have orthonormal columns.

Eigendecomposition

• Eigenvectors and eigenvalues: there exists a vector $x \neq 0$

$$A_{[n\times n]}x_{[n\times 1]}=\lambda x_{[n\times 1]}$$

All eigenvalues are roots (real or complex) of the equation

$$\det(A - \lambda I) = 0$$

There are exactly *n* eigenvalues (not necessarily distinct)

- If A is real-valued and symmetric $(A^T = A)$, all eigenvalues are real
- A is singular iff at least one of the $\lambda = 0$
- A is positive definite $(x^T A x > 0 \text{ for all } x \neq 0) \Leftrightarrow \text{all eigenvalues } \lambda > 0.$ Non-negative definite: all $\lambda \geq 0$.
- $\det(A) = \lambda_1 \times \cdots \times \lambda_n$
- trace(A) = $\lambda_1 + \cdots + \lambda_n$

• If x is an eigenvector, so is cx for any $c \neq 0$, so require ||x|| = 1 Check:

• Collecting all eigenvectors into a $n \times n$ matrix Q (each column is an eigenvector) and writing $\Lambda = \operatorname{diag}(\lambda_1, \dots, \lambda_n)$, we have

$$AQ = Q\Lambda$$

• If Q is invertible, we say A is diagonalizable. For real symmetric matrices, this is always true, and the eigenvectors form a basis of \mathbb{R}^n . The canonical eigenvectors are orthonormal, and $\mathbb{Q}^{-1} = \mathbb{Q}^T$. Then we have the eigendecomposition

$$A = Q\Lambda Q^T$$

Examples

Find the eigendecomposition of

$$1. \left(\begin{array}{cc} 2 & 0 \\ 0 & 1 \end{array}\right)$$

$$2. \left(\begin{array}{cc} 1 & 1 \\ 1 & 1 \end{array}\right)$$

Singular value decomposition

For any real rectangular $m \times n$ matrix A, there exist orthogonal square matrices $U_{[m \times m]}$ and $V_{[n \times n]}$ such that $A = U \Sigma V^T$, where

$$\Sigma = \begin{pmatrix} \operatorname{diag}(d_1, \dots, d_r) & 0_{[r \times (n-r)]} \\ 0_{[(m-r) \times r]} & 0_{[(m-r) \times (n-r)]} \end{pmatrix},$$

- By convention, $d_1 \ge d_2 \ge \cdots \ge d_r > 0$.
- The rank of A is r
- Diagonal elements of Σ are square roots of the eigenvalues of A^TA .
- Columns of V (right singular vectors) are the eigenvectors of A^TA .
- Columns of U (left singular vectors) are the eigenvectors of AA^T .
- The best rank k approximation to A in Frobenius norm can be obtained by keeping the first k singular values and replacing the rest by 0 in Σ .

Example: the Procrustes problem

- Goal: apply translation and rotation to data matrix Y to make it as similar as possible to the "target" data matrix X
- Applications: the registration problem (imaging, brain scans, etc)
- After translation so that X and Y have the same "center" (i.e. mean), the problem is to minimize

$$||X - YR||_F$$
 subject to $R^TR = I$

- R is a rotation matrix
 - Rx is the rotated vector x
 - ► $R^{-1} = R^T$, or $RR^T = I$ (rotating there and back leaves you in the same place)
 - Any orthonormal matrix represents a rotation
- Turns out the Procrustes solution is given by $R = UV^T$, where $U\Sigma V^T$ is the SVD of Y^TX (see Michailidis "Linear algebra review" chapter for derivation, p. 129).

Mean and variance

For a random variable X,

$$E(X) = \int x dP(x) \left[= \int x f(x) dx \text{ or } \sum_{i} x_{i} P(X = x_{i}) \right]$$

$$Var(X) = E(X - E(X))^{2}$$

$$Cov(X_{1}, X_{2}) = E[(X_{1} - E(X_{1}))(X_{2} - E(X_{2}))]$$

$$Corr(X_{1}, X_{2}) = \frac{Cov(X_{1}, X_{2})}{\sqrt{Var(X_{1})Var(X_{2})}}$$

If X_1 and X_2 are independent, $Cov(X_1, X_2) = Corr(X_1, X_2) = 0$.

$$E(aX) = aE(X)$$

$$Var(aX) = a^{2}Var(X)$$

$$Cov(aX_{1},bX_{2}) = abCov(X_{1},X_{2})$$

Mean and variance of vectors

Vector of r.v.s: $X = (X_1, ... X_p)'$ Mean: $E(X) = (E(X_1), ... E(X_p))'$ Variance-covariance matrix:

$$Var(X) = Cov(X) = E[(X - E(X))(X - E(X))'] =$$

$$= \begin{pmatrix} Var(X_1) & Cov(X_1, X_2) & \cdots & Cov(X_1, X_p) \\ Cov(X_2, X_1) & Var(X_2) & \cdots & Cov(X_2, X_p) \\ \vdots & & \ddots & \vdots \\ Cov(X_p, X_1) & \cdots & \cdots & Var(X_p) \end{pmatrix}$$

If X_j 's are independent, Var(X) is diagonal. If A is a constant (non-random) matrix, then

$$E(AX) = AE(X)$$

 $Var(AX) = AVar(X)A^{T}$

Derivatives of vector-valued functions

- Argument $x_{[m \times 1]} = (x_1, \dots x_m)'$
- Function $y_{[n\times 1]} = f(x_{[m\times 1]})$;
- Derivative $D_{[m \times n]} = \partial y / \partial x$ is defined by

$$D_{ij} = \frac{\partial y_j}{\partial x_i}$$

• For a rectangular matrix A,

$$\frac{\partial (Ax)}{\partial x} = A^T$$

• For a square matrix B,

$$\frac{\partial (x^T B x)}{\partial x} = (B + B^T) x$$

Multivariate normal distribution

A vector $X = (X_1, ..., X_p)'$ is multivariate normal if any of the below holds:

- Every linear combination $\sum_i a_i X_i$ has a univariate normal distribution
- $X = \mu + AZ$, where Z is a vector of i.i.d. univariate normals N(0,1). Then $E(X) = \mu$, $Var(X) = AA^T$.
- The joint density of X is given by

$$f(x) = \frac{1}{(2\pi)^{p/2} \det(\Sigma)^{1/2}} \exp\left(-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)\right)$$

where μ is any p-vector and Σ is symmetric and non-negative definite. Then we write $X \sim N_p(\mu, \Sigma)$, and $E(X) = \mu$, $Var(X) = \Sigma$.

Example: bivariate normal

$$X \sim N_2 \left(\left(\begin{array}{c} 1 \\ -1 \end{array} \right), \left(\begin{array}{cc} 1 & -1 \\ -1 & 4 \end{array} \right) \right)$$

Properties of the multivariate normal

- Contours of the multivariate density are ellipsoids
- $\Sigma = I$ correponds to a spherical distribution
- $\Sigma_{ij} = 0$ iff X_i and X_j are independent
- The multivariate normal is the only distribution where correlation 0 is equivalent to independence
- $\Sigma_{ij}^{-1} = 0$ iff X_i and X_j are *conditionally* independent given all other variables.
- For any vector a and matrix $B \neq 0$, a + BX is also multivariate normal.

Conditional distribution for the multivariate normal

• Partition $X_{[p\times 1]}$ into two parts $(X_{1_{[p_1\times 1]}},X_{2_{[p_2\times 1]}})$, corresponding to

$$\mu = \left(\begin{array}{c} \mu_1 \\ \mu_2 \end{array} \right) \; , \; \; \Sigma = \left(\begin{array}{cc} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{array} \right) \; . \label{eq:mu}$$

- Then the conditional distribution of X_1 given $X_2 = x$ is also multivariate normal
- The conditional mean is

$$E(X_1|X_2=x) = \mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(x-\mu_2)$$

• The conditional variance-covariance matrix is

$$Var(X_1|X_2 = x) = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$$

Example

Let X be bivariate normal where both variables have mean 1, variance 2, and their correlation is 0.5. Find (a) The distribution of X_1+X_2 Scalar solution:

Vector solution:

(b) Find the conditional distribution of X_1 at $X_2 = 0$

Practice

- Pair up with a neighbor
- Make up a value for the mean (2 × 1 vector) and the covariance matrix (2 × 2 symmetric positive definite matrix) of a bivariate normal.
- Swap the parameters you made up with your neighbor. Sketch an ellipsoid representing a contour of their distribution. Make sure you show the scale of the axes.