

# Form 2A - Research Master's Psychology: Thesis Research Proposal

# 1. General Information

#### 1.1 Student information

Student name: David Coba

Student Id card number: 12439665

Address: -

Postal code and residence: -

Telephone number: -

Email address: coba@cobac.eu

Major: Psychological methods



## 1.2 Supervisor information

Supervisor name: Maarten Marsman

Second assesor name: Jonas Haslbeck

Specialization: Psychological Methods

#### 1.3 Other information

Date: 1.04.2022

Status: First draft

Number of ECs for the thesis: 32EC

Ethics Review Board (ERB) code: -

# 2. Title and Summary of the Research Project

2.1 Title: Assessing the performance Occam's window for Bayesian model averaging

## 2.2 Summary of proposal

When we select a statistical model and use it to make inferences about its parameters, we usually ignore the uncertainty derived from the model selection process. There are techniques that address this issue, like Bayesian model averaging. However, it is not evident how to choose the set of candidate models to consider. This is specially the case in contexts where the space of possible models is vast, such as with graphical models.

Graduate School of Psychology

University of Amsterdam Psychology

There are multiple model search algorithms used under a Bayesian framework, one of them being Occam's window.

The goal of this project is to assess in general terms if Occam's window is a suitable method to explore the model space, specifically in the context of graphical models. To this end we will conduct a simulation study exploring how the algorithm performs under different conditions and how it compares to other alternative model search techniques.

Keywords: Bayesian inference, model selection, model search algorithms, Occam's window

Word count: 146/150

## 3. Project description

#### 3.1 Prior research

When we perform statistical inferences, such as hypotheses tests about the inclusion of a parameter in a model or whether a parameter lays within an interval, we typically select a statistical model and then use that model to perform the inference. However, this approach underestimates the total uncertainty in our inferences, since it essentially ignores the uncertainty derived from the model selection process and produces overconfident conclusions (Leamer, 1978; Draper et al., 1987; Hoeting et al., 1999; for a recent review of the issue see Kaplan, 2021).

There are different approaches under a Bayesian framework that allow to model the uncertainty of the model selection process, and address the issue of using a single model for our statistical inferences. These approaches can be categorized into two groups. The first group is using mixture models that encompass all possible models. To estimate the joint posterior distribution of all possible models researchers usually employ samplers

like Markov chain Monte Carlo model composition (MC<sup>3</sup>, Madigan & York, 1995), which is based on reversible jump Markov chain Monte Carlo (Green, 1995). However, these samplers are hard to implement, are computationally demanding and tend to have instability issues; which is why the second group of methods is generally preferred (Yao et al., 2018). The second group of approaches to multiple model inference is to combine the information of a set of candidate models  $\mathcal{A}$ . With these methods, the final inference  $p(\Delta|D)$  is the weighted average of that inference across all candidate models  $p(\Delta|M_k,D)$  shown in Equation ??. This approach allows to separate the use of multiple models into two steps: identifying a set of candidate models  $\mathcal{A}$  and combining the uncertainty from those models.

$$p(\Delta|D) = \sum_{\forall k: M_k \in \mathcal{A}} p(\Delta|\mathcal{M}_k, D) w_k \tag{1}$$

There are two main methods to combine multiple models and not ignore the uncertainty of the model-selection process. The first method is Bayesian model averaging (BMA, Hinne et al., 2020; Hoeting et al., 1999; Leamer, 1978) and uses the posterior probability of candidate models as the model weights of Equation??. This posterior probability

$$p(M_k|D) = \frac{p(D|\mathcal{M}_k)p(M_k)}{\sum_{\forall l:M_l \in \mathcal{A}} p(D|M_l)p(M_l)}$$

depends on the marginal likelihood of the data under each model

$$p(D|M_k) = \int p(D|\theta_k, M_k) p(\theta_k|M_k) d\theta_k$$

and their prior probability  $p(M_k)$ . To calculate the marginal likelihoods we need to integrate the product of the likelihood function of each model  $p(D|\theta_k, M_k)$  and the prior distribution of the model parameters  $p(\theta_k|M_k)$  over the whole parameter space. In most cases it is not possible to calculate the marginal likelihoods analytically, and we require of approximate solutions. At the end of this section we provide an overview of the most common approximations.

The second method is model stacking, which minimizes the leave-one-out cross-validation (LOOCV) estimate of a loss function to assign weights to different models (Wolpert, 1992). Stacking is a common technique to aggregate point estimations from different models, but Yao et al. (2018) extend the method to combine Bayesian predictive distributions, producing combined uncertainty distributions similarly to BMA. It is possible to calculate LOOCV estimates from samples of the posterior distribution (Vehtari et al., 2016), which makes it convenient if one is using methods such as Markov chain Monte Carlo to estimate the posterior distributions in the first place.

The main difference between BMA and model stacking is their asymptotic behavior when the data-generating model is not in the set of candidate models  $\mathcal{A}$ . In this scenario, BMA will select the single model that minimizes the Kullback-Leibler divergence from the data-generating process, while model stacking will select the mixture of models that minimizes the loss function that was used to find the model weights (Yao et al., 2018). The literature is divided between proponents of marginal likelihood based methods, such as Bayes factors and BMA, and proponents of methods based on the posterior predictive distributions, such as LOOCV and model stacking. The disagreements seem to be rooted on differences in philosophical positions and scientific goals (Gronau & Wagenmakers, 2018, 2019; Lotfi et al., 2022; Vehtari et al., 2018).

When we do not have strong theoretical arguments to pre-select a set of candidate models  $\mathcal{A}$ , we can use a model search algorithm. One possible algorithm is Occam's window (Madigan & Raftery, 1994; A. E. Raftery et al., 1997), which is based on Occam's razor principle. Occam's razor (also known as the law of parsimony) states than when one is presented with competing hypotheses that explain equally well a particular phenomena, one should choose the simplest one. In general terms, Occam's window algorithm first selects a set of models that fit the data reasonably well, and then discards all models that have simpler counterparts that fit the data equally well. Formally, the first step equals constructing the set of models

$$\mathcal{A}' = \left\{ M_k : \frac{\max\{p(M_l|D)\}}{p(M_k|D)} \le c \right\}$$

with posterior probabilities  $p(M_k|D)$  not significantly lower than the model with highest posterior probability of all models  $M_l \in \mathcal{H}'$ . The constant c specifies the range of



posterior probabilities—the size of the window—that fit the data reasonably well. For second step the algorithm identifies the set of models

$$\mathcal{B} = \left\{ M_k : \exists M_l \in \mathcal{A}', M_l \subset M_k, \frac{p(M_l|D)}{p(M_k|D)} > 1 \right\}$$

that have at least one submodel  $M_l$  in  $\mathcal{A}'$  with greater posterior probability. The final set of candidate models is  $\mathcal{A} = \mathcal{A}' \setminus \mathcal{B}$ . Computationally, the algorithm is a deterministic greedy search that performs two passes over the model space. The first pass goes from the bottom to the top (i.e. comparing the simplest models with p parameters to models with p+1 parameters and so on), and the second pass starts from the most complex models and compares all the way to the simplest. To calculate posterior model probabilities  $p(M_k|D)$  we need to compute the marginal likelihood  $p(D|M_k)$  of each model, similarly to BMA.

One of the drawbacks of Occam's window is that it overestimates the posterior probability of the selected "best" candidate models and it underestimates —essentially nullifies—the posterior probability of the rest of the models. This is by design and acknowledged by Madigan and Raftery (1994), and it is a trade-off we have to make to avoid having to combine information from the complete model space. Occam's window is implemented for linear regression models using priors that allow to analytically calculate the marginal likelihoods (A. E. Raftery et al., 1997) in the R package BMA (A. Raftery et al., 2015). There is also an extension of Occam's window to allows to model streams of data that become available sequentially (Onorante & Raftery, 2016).

Alternative model search algorithms include Bayesian adaptive sampling (BAS) and birth-death Markov chain Monte Carlo (BDMCMC). BAS samples without replacement from the space of possible models and uses the marginal likelihoods of the sampled models to iteratively estimate the marginal likelihoods of the models that remain unsampled (M. A. Clyde et al., 2011). BAS is available for (generalized) linear models as an R package (M. Clyde, 2021). BDMCMC (A. Mohammadi & Wit, 2015) samples from the joint posterior space of all possible models, and uses a Poisson process to model the rate at which the Markov chains jump from one model to another. BDMCMC is available in the R package BDGraph (R. Mohammadi & Wit, 2019) for graphical models, which uses a pseudo-likelihood function (Pensar et al., 2017) and an analytical

approximation to the ratio of marginal likelihoods (R. Mohammadi et al., 2017).

Finally, we also want to give an overview of possible ways of approximating the marginal likelihoods that are required for BMA and Occam's window. The first and crudest one is to use the Bayesian information criterion (BIC, Schwarz, 1978) as an approximation. The BIC of a model  $M_k$  is defined as

$$BIC(M_k) = -2\log p\left(D|\widehat{\theta}, M_k\right) + d_{M_k}\log n,$$

where  $p\left(D|\widehat{\theta},M_k\right)$  is the likelihood of the maximum likelihood estimate for the parameter values under that model,  $d_{Mi}$  is the number of parameters of the model and n is the sample size. The logarithm of the marginal likelihood of a model can be approximated as

$$\log p(D|M_k) \approx \log p(D|\widehat{\theta}, M_k) - \frac{1}{2}d_{M_k}\log n$$

if we assume an unit information prior, which means that

$$\log p\left(D|M_k\right) \approx \frac{\mathrm{BIC}(M_k)}{-2}$$

and that the ratio of marginal likelihoods—the Bayes factor—between two models is

$$2\log B_{ij} = -\mathrm{BIC}(M_i) + \mathrm{BIC}(M_j).$$

Another method to approximate the marginal likelihood is to use bridge sampling (Bennett, 1976; Gronau et al., 2017). Bridge sampling generally provides accurate approximations of the marginal likelihoods, but is also very computationally demanding since it has to draw samples. A method between BIC and bridge sampling in terms of accuracy and computational demands is the Laplace approximation (Kass & Raftery, 1995; LeCam, 1953). This method approximates the posterior distribution with a normal distribution centered around the posterior mode, which can be estimated using expectation-maximization algorithms. The standard Laplace approximation is accurate to the second moment of the posterior distribution, but it is possible to extend it get more accurate approximations at the cost of more computational resources or further assumptions (Hubin & Storvik, 2016; Rue et al., 2009; Ruli et al., 2016; Tierney & Kadane, 1986; Tierney et al., 1989). Lastly, note that in the context of Occam's window, it is possible to use a faster but less accurate approximation during model search, and use a slower but more accurate approximation during the model combination step.



#### 3.2 Key questions

The main goal of this project is to assess how Occam's window model search algorithm performs in general terms. To our knowledge there are no simulation studies evaluating its performance under different conditions. We want to explore the possible trade-offs between accuracy and computational speed of different marginal likelihood approximations, and also how it compares to alternative model search algorithms. Specifically, we are motivated by the issue of deciding whether to include or not particular edges in graphical models. The number of possible graphical models grows exponentially with the number of variables, and we want to check if it is feasible to use Occam's window in this context or not.

Word count: 1449/1200

## 4. Procedure

## 4.1 Operationalization

To address our research questions we will first implement Occam's window model search algorithm and then conduct a simulation study. We plan on implementing our algorithm and running our simulations in the Julia programming language (Bezanson et al., 2017).

There are more simulation conditions that are potentially interesting than how many we can realistically tackle during this project, and the number of conditions that we can test will depend on how smoothly the project progresses. In general terms we plan on running simulations under the following conditions. First, regarding which models to use during our simulations, linear regression is the obvious simplest choice to start developing the algorithm. Logistic regression is a next step that increases the complexity of the procedure, and the Gaussian graphical model and the Ising model are

the ones that motivate this project. Regarding the choice of model search algorithm we will only implement Occam's window algorithm, and rely on the implementations of BAS (M. Clyde, 2021) for linear models and BDgraph (R. Mohammadi & Wit, 2019) for graphical models as benchmarks. Also, since Occam's window algorithm uses marginal likelihoods during model search, it is most practical to use BMA to combine the candidate models, which is also how the algorithm was originally conceived. We will not use model stacking during our simulations. Next, regarding the choice of approximations of the marginal likelihood, we will start with the BIC since it is the simplest approximation and it will allow us to test our implementation of Occam's window while developing it. We will also implement Laplace approximations since we predict that these approximations will be the most efficient in terms of the trade-off between computational speed and accuracy. We will have to explore which specific implementation of the Laplace approximation is more appropriate for our goals. Lastly, we will also consider different sample sizes and sparsity levels in the covariance matrices of the data-generating models. Taking this into consideration, these are broadly speaking the conditions we will prioritize testing:

- 1. Occam's window with linear regression models and BIC approximation.
- 2. Occam's window with linear regression models and Laplace approximation.
- 3. Occam's window with logistic regression models and Laplace approximation.
- 4. Occam's window with Gaussian graphical models and BIC approximation.
- 5. Occam's window with Gaussian graphical models and Laplace approximation.
- 6. BAS with its current implementation in R.
- 7. BDgraph with its current implementation in R.
- 8. Occam's window with Ising models and BIC approximation.
- 9. Occam's window with Ising models and Laplace approximation.
- 10. Using Occam's window model search with BIC, re-run BMA but using the Laplace approximation.



11. Using Occam's window model search with BIC, re-run BMA but using bridge sampling.

We believe that it is realistic to complete up to condition no. 9 in this project. Evaluating the performance of conditions no. 10 and no. 11 will most likely remain open questions for future research.

#### 4.2 Sample characteristics

We plan on generating data from a set of models and evaluate how well each simulation condition recovers the characteristics of the true data-generating models. However, we do not think it makes sense to commit to specific data-generating processes at this stage of the project.

## 4.4 Data analysis

This project is inherently exploratory and, similarly to the last section, we do not think it makes sense to commit at this stage to a specific analysis plan. In general terms, to assess how well each model-search algorithm performs we will compare the posterior probabilities of the true data-generating model, and the posterior probabilities of including specific edges that are present on the data-generating model. To assess computational costs we will use real runtime in order to not penalize algorithms that benefit from parallel computations. If instead we used CPU time, we would be penalizing all parallelizable algorithms by a factor of the number of parallel processes or threads.



## 4.4 Modifiability of procedure

In section 4.1 we have ordered some possible simulation conditions in order of priority and we have estimated how many are realistic to complete during this thesis project. If our estimations prove to be overconfident, we can choose to exclude additional conditions, starting with the ones with lowest priority. Similarly, if everything goes smoother than planned, we can choose to simulate and analyze additional conditions.

Word count: 703/1000

## 5. Intended results

The main goal of this project is to assess in general terms how Occam's window performs. If our analysis concludes that the algorithm compares favorably against alternative methods, we will show that Occam's window can be a useful tool to supplement the use of BMA to avoid the problem of single model inference. We are motivated specially by the case of graphical models, where the space of possible models grows exponentially with the number of variables. Current approaches to sampling from the complete model space have limitations, and we anticipate that Occam's window can be a useful tool that is currently underused. In case that our results show that the performance of Occam's window does not compensate for its shortcomings, we would have provided an updated assessment of its performance that is currently lacking in the literature. Moreover, we expect to contribute software that implements BMA and Occam's window, and that integrates with the rest of the Julia ecosystem.

Word count: 163/250



# 6. Work plan

#### 6.1 Time schedule

This thesis project consists of 28 EC, excluding the thesis proposal. This is equivalent to approximately 18 weeks working full time. We aim to complete and present the project by the 15th of July 2022. In broad terms we plan to achieve the following milestones each month:

#### April

- Week 1/2: Address feedback on the proposal and implement Occam's window algorithm for linear regression models using BIC as an approximation to the marginal likelihood.
- Week 3: Implement the Laplace approximation to the marginal likelihood and test its performance with linear regression models.
- Week 4: Implement the Laplace approximation for logistic models and buffer time.

#### May • Wee

- Week 1: Buffer time and hopefully enjoy the UvA teaching-free days.
- Week 2: Implement the Laplace approximation for graphical Gausian models and start running simulations.
- Week 3: Buffer time and start running simulations with BAS and BDGraph.
- Week 4: Continue running simulations and buffer time.

#### June

- Week 1: Implement the Laplace approximation for Ising models. Continue running simulations.
- Week 2: Continue running simulations and start analyzing results. Start writing the thesis.
- Week 3/4: Analyze results and thesis writing. Complete a first draft of the full thesis.

• Weeks 1/2: Complete writing the thesis and prepare the presentation.

Graduate School of Psychology

University of Amsterdam Psychology

As detailed in section 4.4 "Modifiability of procedure", the scope of this project is

highly flexible, and we can adapt which conditions to include or exclude in our

simulation study depending on how fast we progress.

6.2 Infrastructure

No special infrastructure is required to complete this project.

6.3 Data storage

We will keep the results of all our simulations under version control and with remote

backups. We do not plan on collecting any data, and in the case we end up deciding to

use empirical data we would use publicly available datasets.

6.3 Budget

In principle we will not require extra funds to complete this project. In the case that

the computational resources that we have access to prove insufficient to conduct the

simulations, we might consider using cloud computing services. In any case, such costs

would not exceed the maximum budget.

Word count: 333/500

7. References

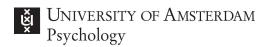
Bennett, C. H. (1976). Efficient estimation of free energy differences from monte carlo

data. Journal of Computational Physics, 22(2), 245–268.

- Bezanson, J., Edelman, A., Karpinski, S., & Shah, V. B. (2017). Julia: A fresh approach to numerical computing. SIAM Review, 59(1), 65–98. https://doi.org/10.1137/141000671
- Clyde, M. (2021). BAS: Bayesian variable selection and model averaging using bayesian adaptive sampling. https://cran.r-project.org/package=BAS
- Clyde, M. A., Ghosh, J., & Littman, M. L. (2011). Bayesian adaptive sampling for variable selection and model averaging. Journal of Computational and Graphical Statistics, 20(1), 80–101.
- Draper, D., Hodges, J. S., Leamer, E. E., Morris, C. N., & Rubin, D. B. (1987). A research agenda for assessment and propagation of model uncertainty. Rand Corporation, Report N-2683-RC.
- Green, P. J. (1995). Reversible jump Markov chain Monte Carlo computation and bayesian model determination. Biometrika, 82(4), 711–732. https://doi.org/10.1093/biomet/82.4.711
- Gronau, Q. F., Sarafoglou, A., Matzke, D., Ly, A., Boehm, U., Marsman, M., Leslie, D. S., Forster, J. J., Wagenmakers, E.-J., & Steingroever, H. (2017). A tutorial on bridge sampling. Journal of Mathematical Psychology, 81(nil), 80–97. https://doi.org/10.1016/j.jmp.2017.09.005
- Gronau, Q. F., & Wagenmakers, E.-J. (2018). Limitations of bayesian leave-one-out cross-validation for model selection. Computational Brain & Behavior, 2(1), 1–11. https://doi.org/10.1007/s42113-018-0011-7
- Gronau, Q. F., & Wagenmakers, E.-J. (2019). Rejoinder: More limitations of bayesian leave-one-out cross-validation. Computational Brain & Behavior, 2(1), 35-47. https://doi.org/10.1007/s42113-018-0022-4
- Hinne, M., Gronau, Q. F., van den Bergh, D., & Wagenmakers, E.-J. (2020). A conceptual introduction to bayesian model averaging. Advances in Methods and Practices in Psychological Science, 3(2), 200–215. https://doi.org/10.1177/2515245919898657
- Hoeting, J. A., Madigan, D., Raftery, A. E., & Volinsky, C. T. (1999). Bayesian model averaging: A tutorial [with comments by M. Clyde, David Draper and EI George, and a rejoinder by the authors]. Statistical science, 14(4), 382–417.

- Hubin, A., & Storvik, G. (2016). Estimating the marginal likelihood with integrated nested laplace approximation (inla). CoRR. http://arxiv.org/abs/1611.01450v1
- Kaplan, D. (2021). On the quantification of model uncertainty: A bayesian perspective. Psychometrika, 86(1), 215–238. https://doi.org/10.1007/s11336-021-09754-5
- Kass, R. E., & Raftery, A. E. (1995). Bayes factors. Journal of the american statistical association, 90(430), 773–795.
- Leamer, E. E. (1978). Specification searches: Ad hoc inference with nonexperimental data. Wiley.
- LeCam, L. (1953). On some asymptotic properties of maximum likelihood estimates and related bayes estimates. University of California Publication in Statististics, 1(11), 277–330.
- Lotfi, S., Izmailov, P., Benton, G., Goldblum, M., & Wilson, A. G. (2022). Bayesian model selection, the marginal likelihood, and generalization. CoRR. http://arxiv.org/abs/2202.11678v1
- Madigan, D., & Raftery, A. E. (1994). Model selection and accounting for model uncertainty in graphical models using occam's window. Journal of the American Statistical Association, 89(428), 1535–1546. https://doi.org/10.1080/01621459.1994.10476894
- Madigan, D., & York, J. (1995). Bayesian graphical models for discrete data. International Statistical Review, 63(2), 215–232.
- Mohammadi, A., & Wit, E. C. (2015). Bayesian Structure Learning in Sparse Gaussian Graphical Models. Bayesian Analysis, 10(1), 109–138. https://doi.org/10.1214/14-BA889
- Mohammadi, R., Massam, H., & Letac, G. (2017). Accelerating bayesian structure learning in sparse gaussian graphical models. CoRR. http://arxiv.org/abs/1706.04416v3
- Mohammadi, R., & Wit, E. C. (2019). BDgraph: An R package for Bayesian structure learning in graphical models. Journal of Statistical Software, 89(3), 1–30. https://doi.org/10.18637/jss.v089.i03
- Onorante, L., & Raftery, A. E. (2016). Dynamic model averaging in large model spaces using dynamic occam s window. European Economic Review, 81, 2–14. https://doi.org/https://doi.org/10.1016/j.euroecorev.2015.07.013

- Pensar, J., Nyman, H., Niiranen, J., & Corander, J. (2017). Marginal pseudo-likelihood learning of discrete markov network structures. Bayesian Analysis, 12(4), nil. https://doi.org/10.1214/16-ba1032
- Raftery, A., Hoeting, J., Volinsky, C., Painter, I., & Yeung, K. Y. Y. (2015). Bma:
  Bayesian model averaging. https://cran.r-project.org/package=BMA
- Raftery, A. E., Madigan, D., & Hoeting, J. A. (1997). Bayesian model averaging for linear regression models. Journal of the American Statistical Association, 92(437), 179–191. https://doi.org/10.1080/01621459.1997.10473615
- Rue, H., Martino, S., & Chopin, N. (2009). Approximate bayesian inference for latent gaussian models by using integrated nested laplace approximations. Journal of the royal statistical society: Series b (statistical methodology), 71(2), 319–392.
- Ruli, E., Sartori, N., & Ventura, L. (2016). Improved laplace approximation for marginal likelihoods. Electronic Journal of Statistics, 10(2), 3986–4009.
- Schwarz, G. (1978). Estimating the dimension of a model. The annals of statistics, 461-464.
- Tierney, L., & Kadane, J. B. (1986). Accurate approximations for posterior moments and marginal densities. Journal of the american statistical association, 81(393), 82–86.
- Tierney, L., Kass, R. E., & Kadane, J. B. (1989). Fully exponential laplace approximations to expectations and variances of nonpositive functions. Journal of the american statistical association, 84(407), 710–716.
- Vehtari, A., Gelman, A., & Gabry, J. (2016). Practical bayesian model evaluation using leave-one-out cross-validation and waic. Statistics and Computing, 27(5), 1413–1432. https://doi.org/10.1007/s11222-016-9696-4
- Vehtari, A., Simpson, D. P., Yao, Y., & Gelman, A. (2018). Limitations of "limitations of bayesian leave-one-out cross-validation for model selection". Computational Brain & Behavior, 2(1), 22–27. https://doi.org/10.1007/s42113-018-0020-6
- Wolpert, D. H. (1992). Stacked generalization. Neural networks, 5(2), 241–259.
- Yao, Y., Vehtari, A., Simpson, D., & Gelman, A. (2018). Using Stacking to Average Bayesian Predictive Distributions (with Discussion). Bayesian Analysis, 13(3), 917–1007. https://doi.org/10.1214/17-BA1091



# 8. Further steps

Make sure your supervisor submits an Ethics Checklist for your intended research to the Ethics Review Board of the Department of Psychology at https://www.lab.uva.nl/lab/ethics/

# 7. Signatures

- □ I hereby declare that both this proposal, and its resulting thesis, will only contain original material and is free of plagiarism (cf. Teaching and Examination Regulation in the research master's course catalogue).
- □ I hereby declare that the result section of the thesis will consist of two subsections, one entitled "confirmatory analyses" and one entitled "exploratory analyses" (one of the two subsections may be empty):
  - 1. The confirmatory analysis section reports exactly the analyses proposed in Section 4 of this proposal.
  - 2. The exploratory analysis section contains not previously specified, and thus exploratory, proposal analyses.

Location: Student's signature: Supervisor's signature:

Amsterdam