

Occam's window and Bayesian model averaging for graphical models

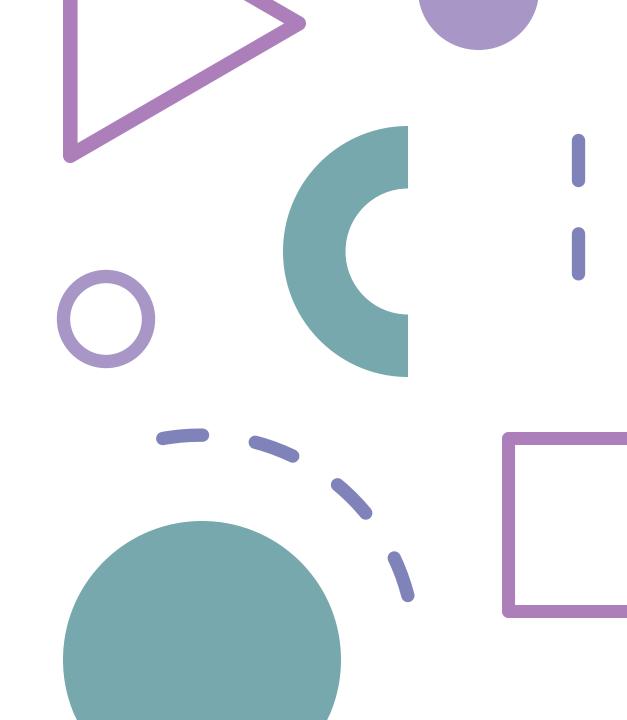
Research Master's Thesis David Coba 22 of July of 2022

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Overview

- The problem of single model inference and Bayesian model averaging
- The idea behind Occam's window
- Occam's window algorithms
 - (Including mine!)
- Simulation study
- Discussion

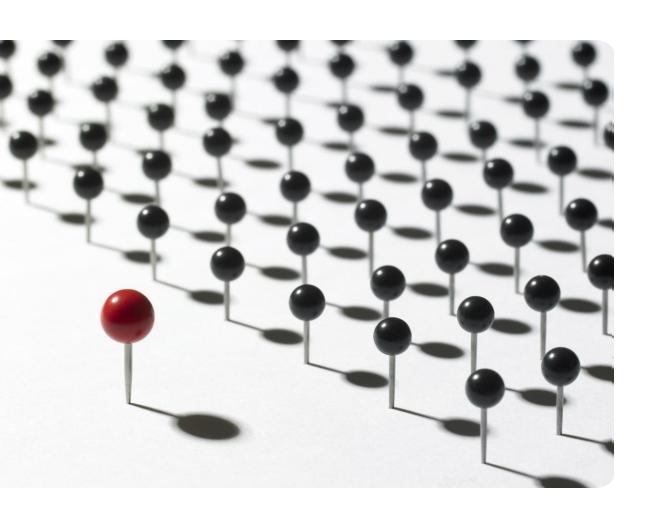
One question.



What is the probability that a parameter is included in a statistical model?

...which is a hard question, particularly when the number of possible models increases exponentially with the number of variables, like with graphical models.

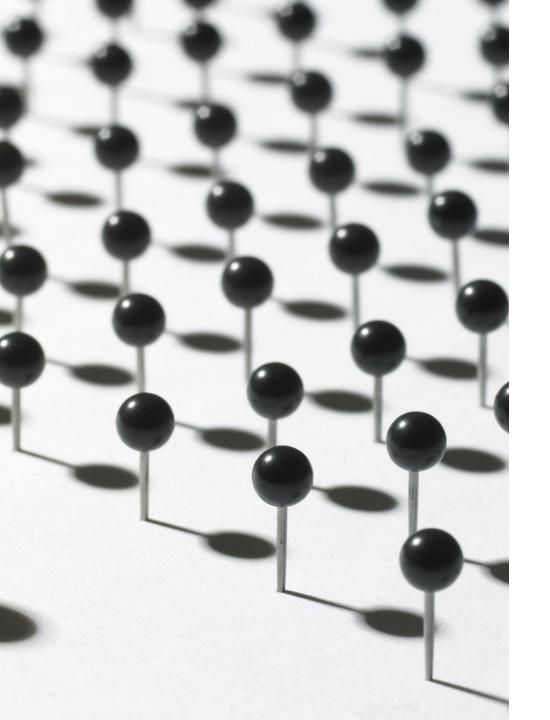




Select a statistical model

Use that model to perform the inference

Is the value of the parameter close to 0? Compare vs null model, Cls, the other Cls, etc.



Issues

- Ignore the uncertainty derived from the model selection process
 - overconfident inferences
- The parameter estimates are always conditioned on the rest of the parameters of the model



Solution

Don't use a single model, but instead model the uncertainty derived from the model selection process

Two approaches under a Bayesian framework

Big mixture models

- They consider the full model space
- Simulation based (Markov chain Monte Carlo) methods (MC³, RJMCMC)
- Their solutions tend to be unstable, hard to implement efficiently, high computational cost

Bayesian model averaging (BMA) Combine the uncertainty from (only) a set of candidate models

P(Parameter) =

P(Parameter | Model) · P(Model | Data)

For all candidate models

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For all candidate models

Posterior probability of a model, P(Model | Data)

- Depends on the marginal likelihood of the data under that model
- Solving an integral over the whole parameter space

 Not tractable, we need to use approximations

10 combine a set of candidate models we can ging (BMA, Hinne et al., 2020; Hoeting et al., 1999; Lewer, proach allows to separate the use of multiple models into two steps: first α a set of candidate models A and then combining the uncertainty from those 4s. With BMA, the posterior probability of our target inference (Δ , e.g. hether a arameter is included in the model or not) given the observed data, $p(\Delta|D)$ the weighted average of that inference across all candidate models $p(\Delta|M_k, D)$, $M_k \in A$. BMA uses the posterior probability of candidate models $p(M_k|D)$ as model weights, and our target inference $p(\Delta|D)$ becomes

$$p(\Delta|D) = \sum_{\forall M_k \in A} p(\Delta|\mathcal{M}_k, D)p(M_k|D).$$

For example, the posterior probability that a specific parameter is included in a model becomes the sum of the posterior probabilities of all models that include that parameter.

From Bayes theorem we know that the posterior probability of a model is the product of the prior probability of that model $p(M_k)$ times the marginal likelihood of data under that model $p(D|M_k)$, divided by the sum of that same product for all

$$\begin{array}{c} \text{Marginal likelihood approximations} \\ \text{Since } O_{\text{Ccam's window uses marginal likelihoods}} \\ \text{Atta under that model } p(D|M_k), \text{ divided by the sum of the sum of the model search, we need efficient ways of approximation is to use the Bayesian information criterion (BIC, Schwarz, models) and the product of the model parameters of the model's parameters, d_{M_k} is the likelihood function evaluated at the maximum likelihood function of the model approximated as d_{M_k} is the miniber of parameters in the model and the likelihood function of the model and the model and the likelihood function of the model and the likelihood function of the model and the likelihood function of the model and the model and the likelihood function of the model and the model and the likelihood function of the model and the likelihood function of the model and the model and the likelihood function of the model and the model and the likelihood function of the model and the likelihood function of the model and the model and the likelihood function of the model and the likelihood function of the model and the model and the likelihood function of the model and the model and the likelihood function of the model and the likelihood function of the model and the model and the likelihood function of t$$

ss the model space, the posterior model narginal likelihoods. And, to calculate the the product of the likelihood function of each ition of the model parameters $p(\theta_k|M_k)$ over

$$\int p(D|\theta_k, M_k) p(\theta_k|M_k) d\theta_k$$

Bayesian model averaging (BMA) Combine the uncertainty from (only) a set of candidate models

P(Parameter) =

P(Parameter | Model) · P(Model | Data)

For all candidate models

How to generate a set of candidate models?

 Ideally, domain knowledge and theoretical arguments.

Not realistic in a lot of scenarios, so...

Model search algorithms, like

Occam's window

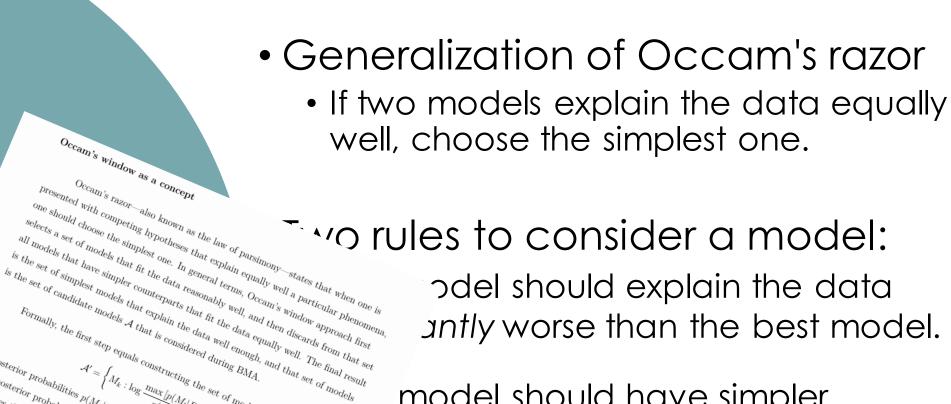
Is Occam's window a potentially useful algorithm to explore the model space of GGMs?

Does it produce OK results?

Does it work in a reasonable timeframe?







Occam's window as a concept

is the set of candidate models A that is considered during BMA.

that have at least one submodel Mi in A with significantly greater posterior That have at least one submodel M_1 in A with significantly greater posteror M_1 in M_2 with significantly greater posteror M_2 in the factor of M_3 in M_4 with significantly greater posteror M_4 in the factor of M_4 in the factor of The final set of candidate models is the set of models in the first set that are not The that set of candidate models is the set of models in the first set that are not that $A = A' \mid B$. That is, the set of models that fit the data reasonably

Both constants, O_L and O_R, determine the size of the window—hence the name

Present in the second A = A

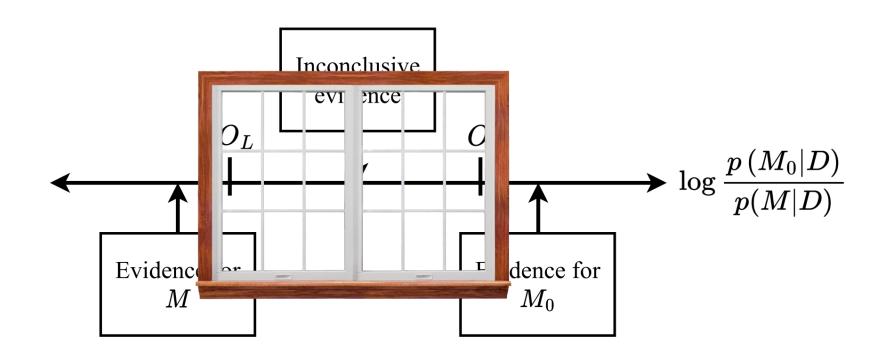
of the algorithm—of acceptable posterior probabilities to consider. denicted in

The

Two rules to consider a model: is the set of simplest models that explain the data equally well. The multiparts with the data well enough, and that set of models

adel should explain the data antly worse than the best model.

model should have simpler nodels that explain the data equally



Occam's window algorithms

Algorithm 1 (original from Raftery & Madigan, 1994)

- Start from a set of initial candidate models
 - Single saturated model
- Deterministic step-wise greedy search testing whether models fall inside the window or not

Algorithm 2 (R package BMA)

- Start from a set of initial candidate models as a proxy of the **full** model space.
 - Super-fast leaps-and-bounds
 - Only linear (and logistic) regression models
- Check which models are in the window
 - Does not require fitting new models!

Occam's window algorithms

Algorithm 1 (original from Raftery &

Madigan, 1994)

Algorithm 1 Occam's window algorithm from Madigan and Raftery (IIIII). An immediate submodel M_0 or supermodel M_1 means that the models differ from the original model M by a single parameter.

```
Require: C
1: A \leftarrow \emptyset
2: repeat {Down pass}

 Select a model M from C.

 4: C \leftarrow C \setminus \{M\}; A \leftarrow A \cup \{M\}
       for inmediate submodel M_0 of M do
         Compute B = \log \left[ p\left(M_0|D\right) / p(M|D) \right]
          if B > O_B then
             A \leftarrow A \setminus \{M\}
             if M_0 \notin C then
                C \leftarrow C \cup \{M_0\}
           else if O_L \leq B \leq O_R then
             if M_0 \notin C then
                C \leftarrow C \cup \{M\}
14: until C is empty
15: \mathcal{C} \leftarrow \mathcal{A}; \mathcal{A} \leftarrow \emptyset

 Select a model M from C.

       C \leftarrow C \setminus \{M\}; A \leftarrow A \cup \{M\}
       for inmediate supermodel M_1 of M do
           Compute B = \log \left[ p\left( M|D \right) / p(M_1|D) \right]
             A \leftarrow A \setminus \{M\}
             if M_1 \notin C then
                C \leftarrow C \cup \{M_1\}
           else if O_L \leq B \leq O_R then
              if M_1 \notin C then
                C \leftarrow C \cup \{M_1\}
28: until C is empty
29: return A
```

Algorithm 2 (R package BMA)

```
Algorithm 2 Occam's window as implemented in BMA.

Require: C

1: for M|M \in C do

2: Compute B_{max} = \log \frac{\max[p(M_t|D) | \forall M_t \in C]}{p(M|D)}

3: if B_{max} > O_L then

4: C \leftarrow C \setminus \{M\}

5: else

6: for inmediate submodel M_0 of M \mid M_0 \in C do

7: Compute B = \log[p(M_0|D)/p(M|D)]

8: if B > O_R then

9: C \leftarrow C \setminus \{M\}

10: else if B < O_L then

11: C \leftarrow C \setminus \{M_0\}

12: return A \leftarrow C
```

Our Occam's window implementation

Based on Algorithm 1



- Because of Julia's multiple dispatch
 - Works with any statistical model
 - The model parameters can be represented as a vector of bits
 - The user implements a marginal likelihood calculation
- Because of Julia's virtually-zero-cost abstractions
 - Potentially as fast as an implementation in a compiled language
 - If the marginal approximation is implemented in an efficient way
- BIC as an approximation to the marginal likelihood, like BMA.

Simulation study

4-way design

Linear regression

- No. of variables = {5, 10, 20}
- No. of observations per variable= {10, 20, 100}
- Variables used in the datagenerating model = {1/4, 1/2, 1}
- Predictors generated from N(0, 1), parameters from N(0, 10) and noise from N(0, 1)
- 20 datasets per condition.

GGM

- No. of variables = {5, 10}
- No. of observations = {500, 2000}
- Sparsity = $\{25\%, 75\%\}$
- Precision matrices generated like Epskamp et al. (2017) & Yin and Li (2011) de.
 - 15 datasets per condition.

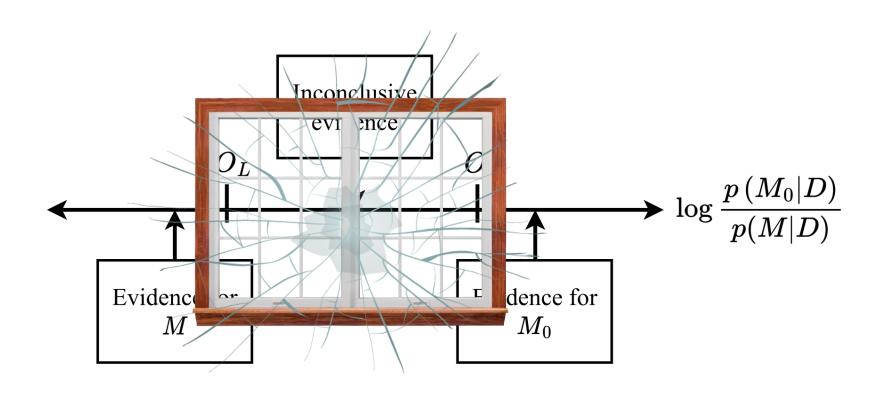
Simulation study

4-way design

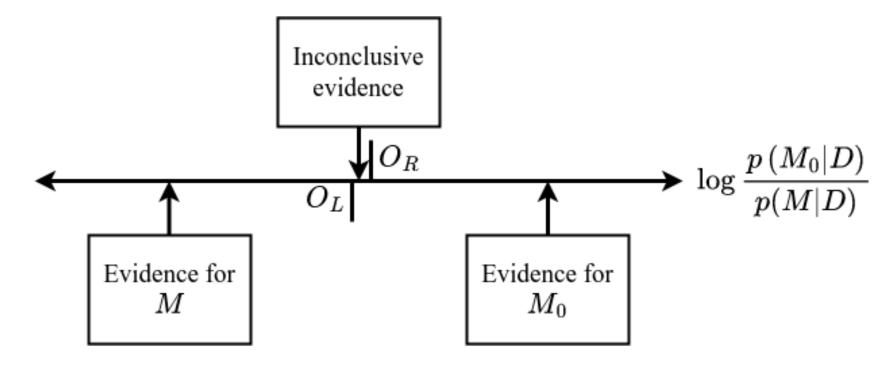
Linear regression

Model	Algorithm	Starting models	Constants
BAS	BAS	-	-
ВМА	Occam's window 2	Leaps & Bounds	Normal
Ours	Occam's window 1	Leaps & Bounds	Normal
11	***	11	Collapsed
"	"	Saturated model	Normal
tt	"	"	Collapsed

Collapsed constants?



Collapsed constants?



Just a step-wise search that selects the model with higher posterior probability at every iteration.

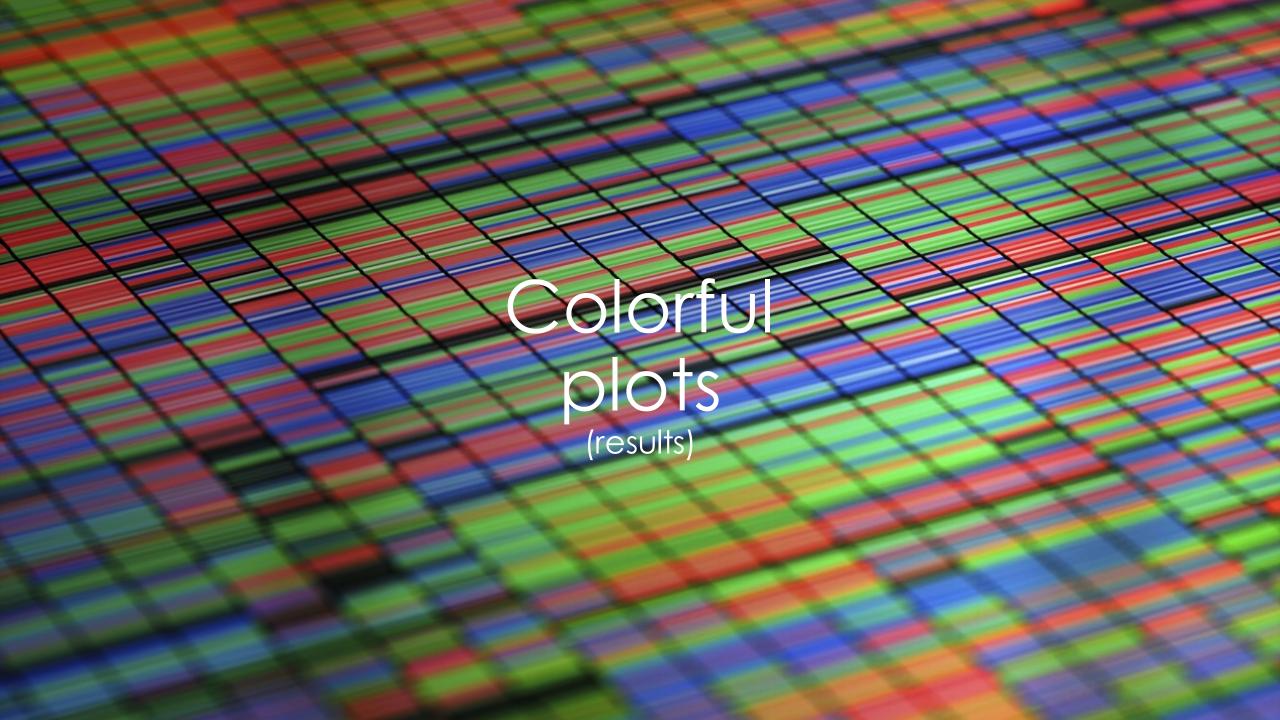
Simulation study

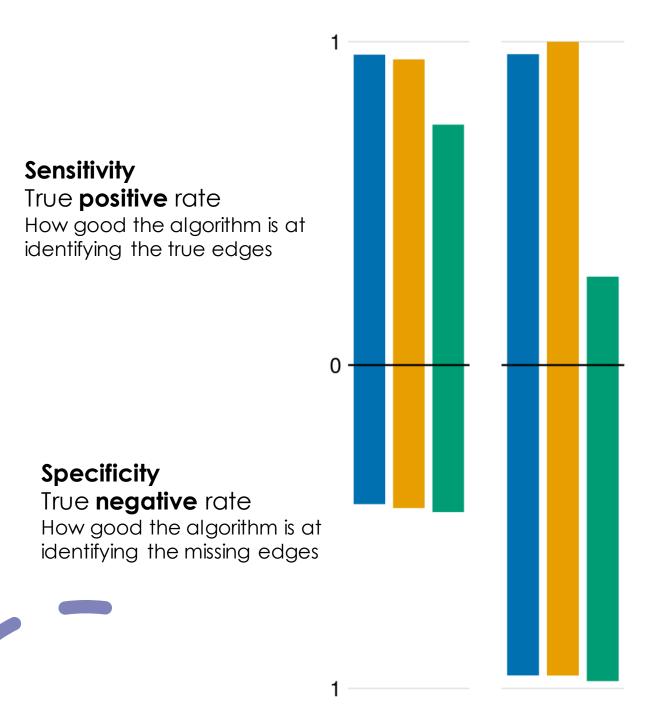
4-way design

GGM

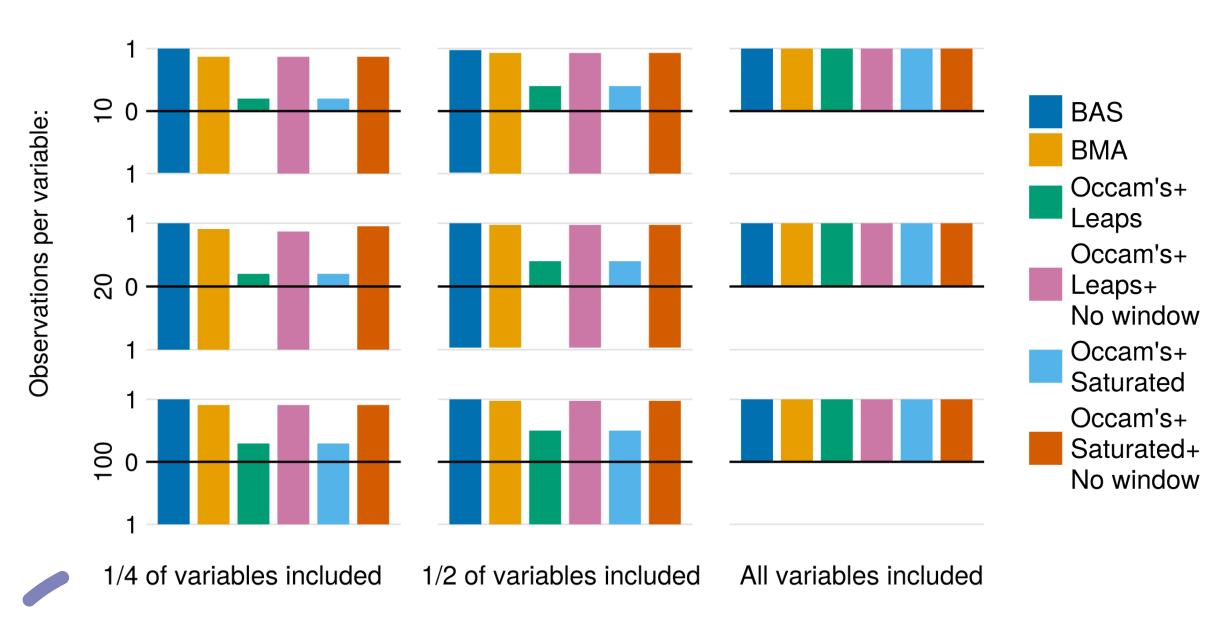
Model	Algorithm	Starting models	Constants
BGGM	Pairwise BF	-	-
BDgraph	(BD)MCMC	-	-
Ours	Occam's window 1	Saturated	Normal



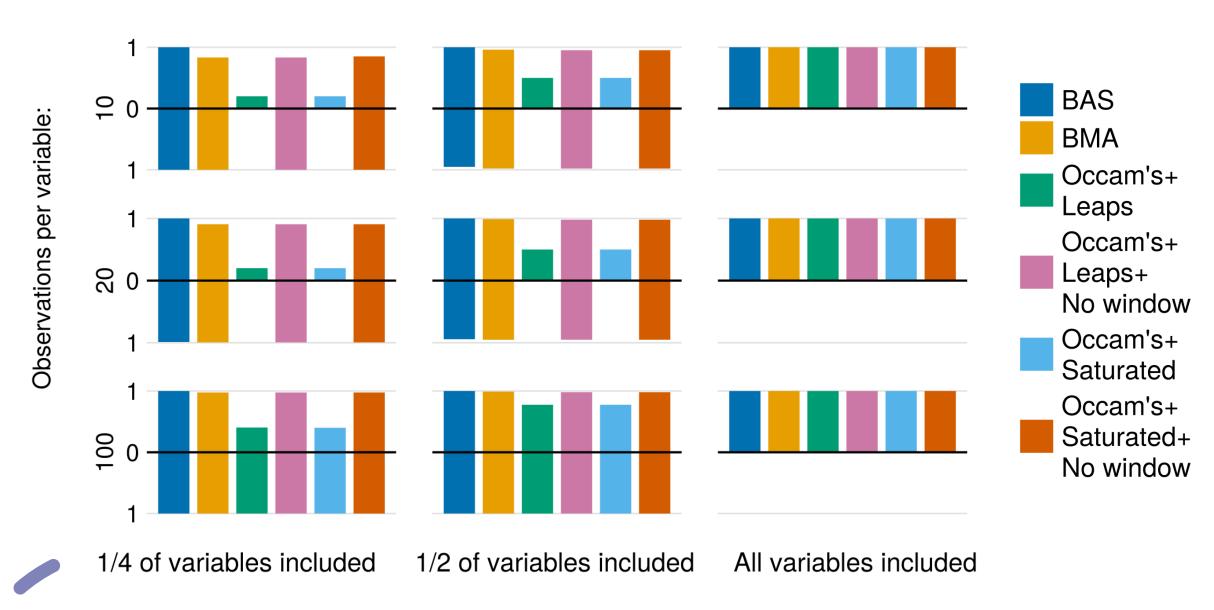


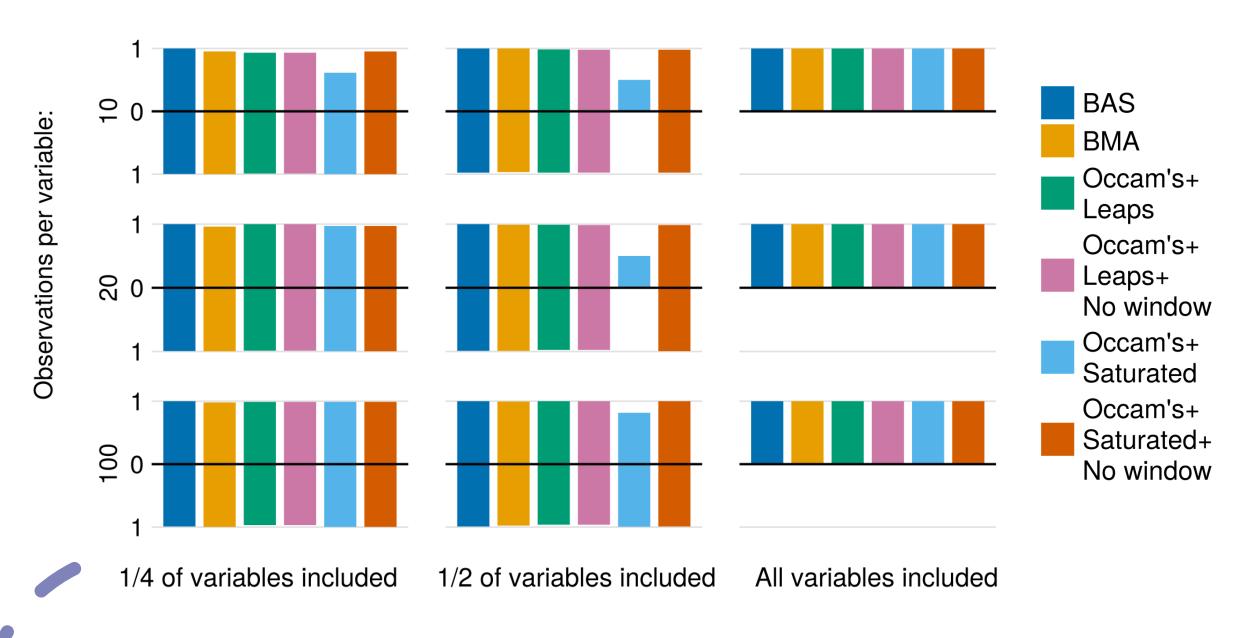






Total no. of variables = 10

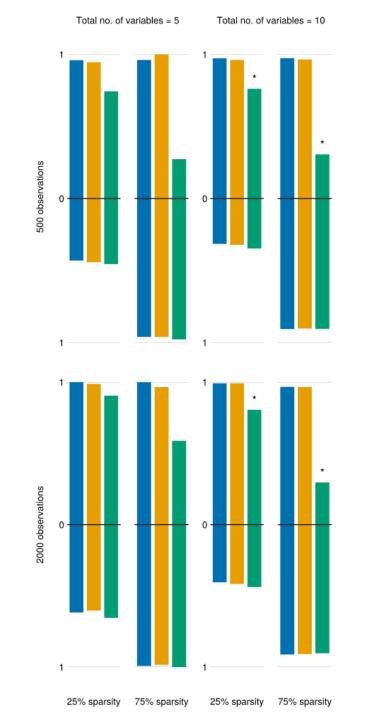




Results linear regression

- BMA & BAS just work great
- Occam's window Algorithm 1 performs worse
 - Especially worse when the algorithm is used as indented (with the window)
 - Tendency to over-include parameters
- Run-time
 - BAS took ~1.5 s
 - BMA a fraction of a second
 - Occam's a few seconds, less than 5







Results GGM

- BGGM & BDgraph work better
- All struggle on low sparsity conditions & all tend to overinclude parameters
- Less differences than in the linear regression case
- Run-time
 - BGGM runs in a fraction of a second
 - BDgraph a few seconds
 - Occam's +30 min, over 1h sometimes



Is Occam's window a potentially useful algorithm to explore the model space of GGMs?

- Does it produce OK results?
- Does it work in a reasonable timeframe?



Discussion

- BGGM and BDgraph work fine in the simple (continuous normal data) case
 - We know their performance is much worse in other cases
- All other algorithms use very optimized calculations, ours is a bare-bones implementation

Discussion

- How to improve speed?
 - Use some sort of sequential calculation for the marginal likelihood
 - Possibly a pseudo-likelihood like Pensar et al. (2017) or Mohammadi et al. (2017).
- How to improve the results?
 - Occam's window Algorithm 2 is the key for BMA's performance
 - Requires a good reduced set of candidate models as a proxy for the whole model space
 - Maybe Marsman et al. (2022), only exploring models with promising edges

Is Occam's window a potentially useful algorithm to explore the model space of GGMs?

