Machine Learning Exercise Sheet 05

Linear Classification

Exercise sheets consist o[illegible]k and in-class exercises. You solve the homework exercises on your own or with you[illegible]d upload it to Moodle for a possible grade bonus. The in-class exercises will be sol[illegible]ring the tutorial. You do not have to upload any solutions of the in-class exercises.

In-class Exercises

Multi-Class Classification

Problem 1: Consider a generative classification model for C classes defined by class probabilities $\mathrm{p}(y = c) = \pi_c$ and general class-conditional densities $\mathrm{p}(\boldsymbol{x} \mid y = c, \boldsymbol{\theta}_c)$ where $\boldsymbol{x} \in \mathbb{R}^D$ is the input feature vector and $\boldsymbol{\theta} = \{\boldsymbol{\theta}_c\}_{c=1}^C$ are further model parameters. Suppose we are given a training set $\mathcal{D} = \{(\boldsymbol{x}^{(n)}, y^{(n)})\}_{n=1}^N$ where $y^{(n)}$ is a binary target vector of length C that uses the 1-of-C (one-hot) encoding scheme, so that it has components $y_c^{(n)} = \delta_{ck}$ if pattern n is from class $y = k$. Assuming that the data points are i.i.d., show that the maximum-likelihood solution for the class probabilities $\boldsymbol{\pi}$ is given by

$$\pi_c = \frac{N_c}{N}$$

where N_c is the number of data points assigned to class c.

The data likelihood given the parameters $\{\pi_c, \boldsymbol{\theta}_c\}_{c=1}^C$ is

$$\mathrm{p}(\mathcal{D} | \{\pi_c, \boldsymbol{\theta}_c\}_{c=1}^C) = \prod_{n=1}^N \prod_{c=1}^C (\mathrm{p}(\boldsymbol{x}^{(n)} | \boldsymbol{\theta}_c)\pi_c)^{y_c^{(n)}}$$

and so the data log-likelihood is given by

$$\log \mathrm{p}(\mathcal{D} | \{\pi_c, \boldsymbol{\theta}_c\}_{c=1}^C) = \sum_{n=1}^N \sum_{c=1}^C y_c^{(n)} \log \pi_c + \text{const w.r.t. } \pi_c.$$

In order to maximize the log likelihood with respect to π_c we need to preserve the constraint $\sum_c \pi_c = 1$. For this we use the method of Lagrange multipliers where we introduce λ as an unconstrained additional parameter and find a local extremum of the unconstrained function

$$\sum_{n=1}^N \sum_{c=1}^C y_c^{(n)} \log \pi_c - \lambda \left(\sum_{c=1}^C \pi_c - 1 \right).$$

instead. See wikipedia article on Lagrange multipliers for an intuition of why this works. This function is a sum of concave terms in π_c as well as λ and is therefore itself concave in these variables.

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We can find the extremum by finding the root of the derivatives. Setting the derivative with respect to π_c equal to zero, we obtain

$$\frac{1}{\lambda}\sum_{n=1}^{N} y_c^{(n)} = \frac{N_c}{\lambda}.$$

Setting the derivative [illegible]al to zero, we obtain the original constraint

$$\sum_{c=1}^{C} \pi_c = 1$$

where we can now plug in the previous result $\pi_c = \frac{N_c}{\lambda}$ and obtain $\lambda = \sum_c N_c = N$. Plugging this in turn into the expression for π_c we obtain

$$\pi_c = \frac{N_c}{N}$$

which we wanted to show.

Linear Discriminant Analysis

Problem 2: Using the same classification model as in the previous question, now suppose that the class-conditional densities are given by Gaussian distributions with a *shared* covariance matrix, so that

$$p(\boldsymbol{x} \mid y = c, \boldsymbol{\theta}) = p(\boldsymbol{x} \mid \boldsymbol{\theta}_c) = \mathcal{N}(\boldsymbol{x} \mid \boldsymbol{\mu}_c, \boldsymbol{\Sigma}).$$

Show that the maximum likelihood estimate for the mean of the Gaussian distribution for class c is given by

$$\boldsymbol{\mu}_c = \frac{1}{N_c}\sum_{\substack{n=1\\ y^{(n)}=c}}^{N} \boldsymbol{x}^{(n)}$$

which represents the mean of the observations assigned to class c.

Similarly, show that the maximum likelihood estimate for the shared covariance matrix is given by

$$\boldsymbol{\Sigma} = \sum_{c=1}^{C} \frac{N_c}{N}\boldsymbol{S}_c \quad \text{where} \quad \boldsymbol{S}_c = \frac{1}{N_c}\sum_{\substack{n=1\\ y^{(n)}=c}}^{N} (\boldsymbol{x}^{(n)} - \boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)} - \boldsymbol{\mu}_c)^{\mathrm{T}}.$$

Thus $\boldsymbol{\Sigma}$ is given by a weighted average of the sample covariances of the data associated with each class, in which the weighting coefficients N_c/N are the prior probabilities of the classes.

We begin by writing out the data log-likelihood.

$$\begin{aligned}
&\log p(\mathcal{D} \mid \{\pi_c, \boldsymbol{\theta}_c\}_{c=1}^{C}) \\
&= \sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)} \log \pi_c \cdot p(\boldsymbol{x}^{(n)} \mid \boldsymbol{\mu}_c, \boldsymbol{\Sigma})
\end{aligned}$$

Then we plug in the definition of the multivariate Gaussian

$$= \sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)} \log\left((2\pi)^{-\frac{D}{2}}\det(\boldsymbol{\Sigma})^{-\frac{1}{2}}\exp\left(-\frac{1}{2}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)\right)\right) + y^{(n)}\log\pi_c$$

and simplify.

$$= -\frac{1}{2}\sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)} \text{[illegible]}\boldsymbol{\Sigma}) + (\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c) - 2\log\pi_c\Big)$$

This expression is conc[illegible]btain the maximizer by finding the root of the derivative. With the help of the matrix cookbook, we identify the derivative with respect to $\boldsymbol{\mu}_c$ as

$$\sum_{n=1}^{N} y_c^{(n)}\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)$$

which we can set to 0 and solve for $\boldsymbol{\mu}_c$ to obtain

$$\boldsymbol{\mu}_c = \frac{1}{\sum_{n=1}^{N} y_c^{(n)}}\sum_{n=1}^{N} y_c^{(n)}\boldsymbol{x}^{(n)} = \frac{1}{N_c}\sum_{\substack{n=1\\ y^{(n)}=c}}^{N}\boldsymbol{x}^{(n)}$$

To find the optimal $\boldsymbol{\Sigma}$, we need the trace trick

$$a = \mathrm{Tr}(a) \text{ for all } a \in \mathbb{R} \quad \text{and} \quad \mathrm{Tr}(\boldsymbol{ABC}) = \mathrm{Tr}(\boldsymbol{BCA}).$$

With this we can rewrite

$$(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c) = \mathrm{Tr}\left(\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\right)$$

and use the matrix-trace derivative rule $\frac{\partial}{\partial \boldsymbol{A}}\mathrm{Tr}(\boldsymbol{AB}) = \boldsymbol{B}^{\mathrm{T}}$ to find the derivative of the data log-likelihood with respect to $\boldsymbol{\Sigma}$. Because the log-likelihood contains both $\boldsymbol{\Sigma}$ and $\boldsymbol{\Sigma}^{-1}$, we convert one into the other with $\log\det\boldsymbol{A} = -\log\det\boldsymbol{A}^{-1}$ to obtain

$$-\frac{1}{2}\sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)}\left(-\log\det\boldsymbol{\Sigma}^{-1} + \mathrm{Tr}\left(\Sigma^{-1}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\right)\right) + \text{const w.r.t. } \boldsymbol{\Sigma}.$$

Finally, we use rule (57) from the matrix cookbook $\frac{\partial \log|\det \boldsymbol{X}|}{\partial \boldsymbol{X}} = (\boldsymbol{X}^{-1})^{\mathrm{T}}$ and compute the derivative of the log-likelihood with respect to $\boldsymbol{\Sigma}^{-1}$ as

$$-\frac{1}{2}\sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)}\left(-\boldsymbol{\Sigma}^{\mathrm{T}} + (\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\right).$$

We find the root with respect to $\boldsymbol{\Sigma}$ and find

$$\boldsymbol{\Sigma} = \frac{1}{\sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)}}\left(\sum_{n=1}^{N}\sum_{c=1}^{C} y_c^{(n)}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}\right)^{\mathrm{T}} = \frac{1}{N}\sum_{c=1}^{C}\sum_{\substack{n=1\\ y^{(n)}=c}}^{N}(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)(\boldsymbol{x}^{(n)}-\boldsymbol{\mu}_c)^{\mathrm{T}}$$

which we can immediately break apart into the representation in the instructions.

Homework

Linear classification

Problem 3: We want [illegible]e binary classification model for classifying *non-negative* one-dimensional data. T[illegible]bels are binary ($y \in \{0, 1\}$) and the samples are $x \in [0, \infty)$.

We assume uniform clas[illegible]

$$[\text{illegible}]0) = \mathrm{p}(y = 1) = \frac{1}{2}.$$

As our samples x are non-negative, we use exponential distributions (and not Gaussians) as class conditionals:

$$\mathrm{p}(x \mid y = 0) = \mathrm{Expo}(x \mid \lambda_0) \qquad \text{and} \qquad \mathrm{p}(x \mid y = 1) = \mathrm{Expo}(x \mid \lambda_1),$$

where $\lambda_0 \neq \lambda_1$. Assume, that the parameters λ_0 and λ_1 are known and fixed.

a) Suppose you are given an observation x. What is the name of the posterior distribution $\mathrm{p}(y \mid x)$? You only need to provide the name of the distribution (e.g., "normal", "gamma" etc.), not estimate its parameters.

Bernoulli.

Remark: y can only take values in $\{0, 1\}$, so obviously Bernoulli is the only possible answer.

b) What values of x are classified as class 1? (As usual, we assume that the classification decision is $\hat{y} = \arg\max_k \mathrm{p}(y = k \mid x)$)

Sample x is classified as class 1 if $\mathrm{p}(y = 1 \mid x) > \mathrm{p}(y = 0 \mid x)$. This is the same as saying

$$\frac{\mathrm{p}(y = 1 \mid x)}{\mathrm{p}(y = 0 \mid x)} \overset{!}{>} 1 \quad \text{or equivalently,} \quad \log \frac{\mathrm{p}(y = 1 \mid x)}{\mathrm{p}(y = 0 \mid x)} \overset{!}{>} 0.$$

We begin by simplifying the left hand side.

$$\begin{aligned} \log \frac{\mathrm{p}(y = 1 \mid x)}{\mathrm{p}(y = 0 \mid x)} &= \log \frac{\mathrm{p}(x \mid y = 1)\,\mathrm{p}(y = 1)}{\mathrm{p}(x \mid y = 0)\,\mathrm{p}(y = 0)} \\ &= \log \frac{\mathrm{p}(x \mid y = 1)}{\mathrm{p}(x \mid y = 0)} \\ &= \log \frac{\lambda_1 \exp(-\lambda_1 x)}{\lambda_0 \exp(-\lambda_0 x)} \\ &= \log \frac{\lambda_1}{\lambda_0} + \lambda_0 x - \lambda_1 x = \log \frac{\lambda_1}{\lambda_0} + (\lambda_0 - \lambda_1) x \end{aligned}$$

To figure out which x are classified as class 1, we need to solve for x.

$$\log \frac{\lambda_1}{\lambda_0} + (\lambda_0 - \lambda_1) x > \log 1 \quad \Leftrightarrow \quad (\lambda_0 - \lambda_1) x > -\log \frac{\lambda_1}{\lambda_0} = \log \lambda_0 - \log \lambda_1$$

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We have to be careful, because if $(\lambda_0 - \lambda_1) < 0$, dividing by it will flip the inequality sign. Hence the answer is

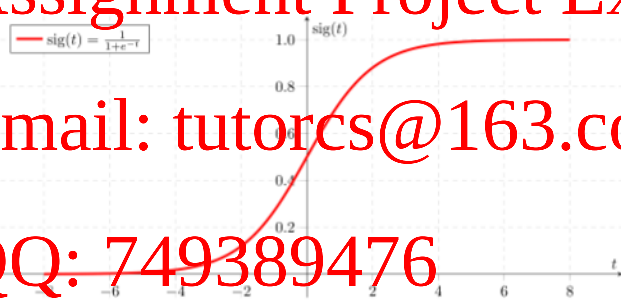
$$\begin{cases} [illegible] \left(\frac{\log\lambda_0 - \log\lambda_1}{\lambda_0 - \lambda_1}, \infty\right) & \text{if } \lambda_0 > \lambda_1 \\ [illegible] \frac{\log\lambda_0 - \log\lambda_1}{\lambda_0 - \lambda_1}\right) & \text{otherwise.} \end{cases}$$

Problem 4: Let $\mathcal{D} = \{($ [illegible] separable dataset for 2-class classification, i.e. there exists a vector $\boldsymbol{w}$ such that sig [illegible] e classes. Show that the maximum likelihood parameter $\boldsymbol{w}$ of a logistic regression model has $\|\boldsymbol{w}\| \to \infty$. Assume that $\boldsymbol{w}$ contains the bias term.

How can we modify the training process to prefer a $\boldsymbol{w}$ of finite magnitude?

In logistic regression, we model the posterior distribution as

$$y_i \mid \boldsymbol{x} \sim \text{Bernoulli}(\sigma(\boldsymbol{w}^{\mathrm{T}}\boldsymbol{x}_i)) \quad \text{where} \quad \sigma(a) = \frac{1}{1 + \exp(-a)}.$$



We fit the logistic regression model by choosing the parameter $\boldsymbol{w}$ that maximizes the data log-likelihood or alternatively minimizes the negative log-likelihood which expands to

$$E(\boldsymbol{w}) = -\log \mathrm{p}(\boldsymbol{y} \mid \boldsymbol{w}, \boldsymbol{X}) = -\sum_{i=1}^{N} y_i \log \sigma(\boldsymbol{w}^{\mathrm{T}}\boldsymbol{x}_i) + (1 - y_i)\log(1 - \sigma(\boldsymbol{w}^{\mathrm{T}}\boldsymbol{x}_i)).$$

We assumed that the data-set is linearly separable, so by definition there is a $\tilde{\boldsymbol{w}}$ such that

$$\tilde{\boldsymbol{w}}^{\mathrm{T}}\boldsymbol{x}_i > 0 \text{ if } y_i = 1 \quad \text{and} \quad \tilde{\boldsymbol{w}}^{\mathrm{T}}\boldsymbol{x}_i < 0 \text{ if } y_i = 0.$$

Scaling this separator $\tilde{\boldsymbol{w}}$ by a factor $\lambda \gg 0$ makes the negative log-likelihood smaller and smaller. To see this, we compute the limit

$$\lim_{\lambda\to\infty} E(\lambda\tilde{\boldsymbol{w}}) = -\left(\sum_{\substack{i=1\\y_i=1}}^{N} \log \lim_{\lambda\to\infty} \sigma(\lambda\,\overbrace{\tilde{\boldsymbol{w}}^{\mathrm{T}}\boldsymbol{x}_i}^{>0}) + \sum_{\substack{i=1\\y_i=0}}^{N} \log\left(1 - \lim_{\lambda\to\infty}\sigma(\lambda\,\overbrace{\tilde{\boldsymbol{w}}^{\mathrm{T}}\boldsymbol{x}_i}^{<0})\right)\right) = 0$$

which equals the smallest achievable value (E is the negative log of a probability, so $E(\boldsymbol{w}) \in [0, \infty)$ and thus $E(\boldsymbol{w}) \geq 0$).

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We can see that E is a convex function because $\log$ is concave and [illegible] is convex if $a < 0$ and concave if $a > 0$. So $\log \sigma(a)$ is concave if $a > 0$ and $\log(1 - \sigma(a))$ is concave if $a < 0$. It follows that E is a convex function because E is the negative sum of concave functions.

A convex function has [illegible] if it attains its minimum value. We know that E tends towards its minimum [illegible] not have a finite minimizer and all its minima are only achieved in the limit. [illegible]ution to the loss minimization problem has infinite norm.

Because E is convex a[illegible]imit of 0 in some directions, we can move the minimum into the space of finite [illegible]y convex term that achieves its minimum such as $\boldsymbol{w}^{\mathrm{T}}\boldsymbol{w}$ or similar forms of wei[illegible]

Problem 5: Show that the softmax function is equivalent to a sigmoid in the 2-class case.

$$
\begin{aligned}
\frac{\exp(\boldsymbol{w}_1^T \boldsymbol{x})}{\exp(\boldsymbol{w}_1^T \boldsymbol{x}) + \exp(\boldsymbol{w}_0^T \boldsymbol{x})} &= \frac{1}{1 + \exp(\boldsymbol{w}_0^T \boldsymbol{x}) / \exp(\boldsymbol{w}_1^T \boldsymbol{x})} \\
&= \frac{1}{1 + \exp(\boldsymbol{w}_0^T \boldsymbol{x} - \boldsymbol{w}_1^T \boldsymbol{x})} \\
&= \frac{1}{1 + \exp(-(\boldsymbol{w}_1 - \boldsymbol{w}_0)^T \boldsymbol{x})} \\
&= \sigma(\hat{\boldsymbol{w}}^T \boldsymbol{x})
\end{aligned}
$$

where $\hat{\boldsymbol{w}} = \boldsymbol{w}_1 - \boldsymbol{w}_0$.

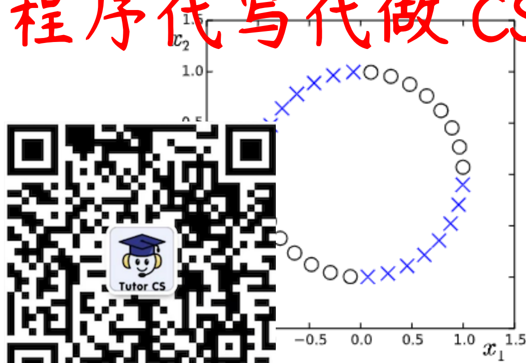
One conclusion we can draw from this is that if we have C parameter vectors $\boldsymbol{w}_c$ for C classes, the logistic regression model is unidentifiable. This means that adding a constant $\boldsymbol{\tau} \in \mathbb{R}^D$ to each vector $\boldsymbol{w}_c := \boldsymbol{w}_c + \boldsymbol{\tau}$ would lead to the same logistic regression model. We can fix this issue by adding a constraint $\boldsymbol{w}_1 = \boldsymbol{0}$, which is what is done implicitly when we use sigmoid (instead of 2-class softmax) in binary classification.

Problem 6: Show that the derivative of the sigmoid function $\sigma(a) = (1 + e^{-a})^{-1}$ can be written as

$$
\frac{\partial \sigma(a)}{\partial a} = \sigma(a)\,(1 - \sigma(a))\,.
$$

$$
\frac{\partial \sigma(a)}{\partial a} = -\frac{1}{(1+e^{-a})^2} \cdot e^{-a} \cdot (-1) = \frac{1}{1+e^{-a}} \frac{e^{-a}}{1+e^{-a}} = \sigma(a) \frac{1 + e^{-a} - 1}{1 + e^{-a}} = \sigma(a)\,(1 - \sigma(a))
$$

Problem 7: Give a basis function $\phi(x_1, x_2)$ that makes the data in the example below linearly separable (crosses in one class, circles in the other).



One example is $\phi(\boldsymbol{x}) = x_1 x_2$ which makes the data separable by the hyperplane $\boldsymbol{w} = (1)$ because the circles will be mapped to the positive real numbers while the crosses go to the negative numbers, i.e. $\boldsymbol{w}^{\mathrm{T}}\boldsymbol{x} > 0$ if $\boldsymbol{x}$ is a circle and $\boldsymbol{w}^{\mathrm{T}}\boldsymbol{x} < 0$ otherwise.

Naive Bayes

Problem 8: In 2-class classification the decision boundary Γ is the set of points where both classes are assigned equal probability,

$$\Gamma = \{\boldsymbol{x} \mid \mathrm{p}(y = 1 \mid \boldsymbol{x}) = \mathrm{p}(y = 0 \mid \boldsymbol{x})\}.$$

Show that Naive Bayes with Gaussian class likelihoods produces a quadratic decision boundary in the 2-class case, i.e. that Γ can be written with a quadratic equation of $\boldsymbol{x}$,

$$\Gamma = \{\boldsymbol{x} \mid \boldsymbol{x}^{\mathrm{T}}\boldsymbol{A}\boldsymbol{x} + \boldsymbol{b}^{\mathrm{T}}\boldsymbol{x} + c = 0\},$$

for some $\boldsymbol{A}$, $\boldsymbol{b}$ and c.

As a reminder, in Naive Bayes we assume class prior probabilities

$$\mathrm{p}(y = 0) = \pi_0 \quad \text{and} \quad \mathrm{p}(y = 1) = \pi_1$$

and class likelihoods

$$\mathrm{p}(\boldsymbol{x} \mid y = c) = \mathcal{N}(\boldsymbol{x} \mid \boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c)$$

with per-class means $\boldsymbol{\mu}_c$ and *diagonal* (because of the feature independence) covariances $\boldsymbol{\Sigma}_c$.

Because $\mathrm{p}(y = 1 \mid \boldsymbol{x}) + \mathrm{p}(y = 0 \mid \boldsymbol{x}) = 1$ and we want them to be equal, we can assume that $\mathrm{p}(y = 0 \mid \boldsymbol{x}) > 0$ and rewrite the defining equation as

$$\frac{\mathrm{p}(y = 1 \mid \boldsymbol{x})}{\mathrm{p}(y = 0 \mid \boldsymbol{x})} = 1.$$

Now apply the logarithm to both sides and simplify.

$$
\begin{aligned}
\log \frac{p(y=1 \mid \boldsymbol{x})}{p(y=0 \mid \boldsymbol{x})} &= \log \left(\frac{p(\boldsymbol{x} \mid y=1)\, p(y=1)}{p(\boldsymbol{x})} \cdot \frac{p(\boldsymbol{x})}{p(\boldsymbol{x} \mid y=0)\, p(y=0)} \right) \\
&= \log \left(p(\boldsymbol{x} \mid y=1)\, p(y=1)\right) - \log \left(p(\boldsymbol{x} \mid y=0)\, p(y=0)\right) \\
&= \log \mathcal{N}(\boldsymbol{x} \mid \boldsymbol{\mu}_1, \boldsymbol{S}_1) - \log \mathcal{N}(\boldsymbol{x} \mid \boldsymbol{\mu}_0, \boldsymbol{S}_0) + \log \frac{\pi_1}{\pi_0} \\
&= -\frac{1}{2}\log|\boldsymbol{S}_1| - \frac{1}{2}(\boldsymbol{x}-\boldsymbol{\mu}_1)^T \boldsymbol{S}_1^{-1}(\boldsymbol{x}-\boldsymbol{\mu}_1) \\
&\quad + \frac{1}{2}\log|\boldsymbol{S}_0| + \frac{1}{2}(\boldsymbol{x}-\boldsymbol{\mu}_0)^T \boldsymbol{S}_0^{-1}(\boldsymbol{x}-\boldsymbol{\mu}_0) + \log \frac{\pi_1}{\pi_0} \\
&= -\frac{1}{2}\boldsymbol{x}^T \boldsymbol{S}_1^{-1}\boldsymbol{x} + \boldsymbol{x}^T \boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 - \frac{1}{2}\boldsymbol{\mu}_1^T \boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 \\
&\quad + \frac{1}{2}\boldsymbol{x}^T \boldsymbol{S}_0^{-1}\boldsymbol{x} - \boldsymbol{x}^T \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0 + \frac{1}{2}\boldsymbol{\mu}_0^T \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0 + \frac{1}{2}\log\frac{|\boldsymbol{S}_0|}{|\boldsymbol{S}_1|} + \log \frac{\pi_1}{\pi_0} \\
&= \frac{1}{2}\boldsymbol{x}^T[\boldsymbol{S}_0^{-1} - \boldsymbol{S}_1^{-1}]\boldsymbol{x} + \boldsymbol{x}^T[\boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 - \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0] \\
&\quad - \frac{1}{2}\boldsymbol{\mu}_1^T \boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_0^T \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0 + \log \frac{\pi_1}{\pi_0} + \frac{1}{2}\log\frac{|\boldsymbol{S}_0|}{|\boldsymbol{S}_1|}
\end{aligned}
$$

This shows that Γ is quadratic and can alternatively be written as

$$\Gamma = \{\boldsymbol{x} \mid \boldsymbol{x}^T \boldsymbol{A}\boldsymbol{x} + \boldsymbol{b}^T\boldsymbol{x} + c = 0\}$$

where

$$\boldsymbol{A} = \frac{1}{2}[\boldsymbol{S}_0^{-1} - \boldsymbol{S}_1^{-1}] \qquad \boldsymbol{b} = \boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 - \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0$$

$$c = -\frac{1}{2}\boldsymbol{\mu}_1^T \boldsymbol{S}_1^{-1}\boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_0^T \boldsymbol{S}_0^{-1}\boldsymbol{\mu}_0 + \log\frac{\pi_1}{\pi_0} + \frac{1}{2}\log\frac{|\boldsymbol{S}_0|}{|\boldsymbol{S}_1|}.$$

If both classes had the same covariance matrix ($\boldsymbol{S}_0 = \boldsymbol{S}_1$), $\boldsymbol{A}$ would be the zero matrix and we would obtain a linear decision boundary as we did in the lecture (also, $\log \frac{|\boldsymbol{S}_0|}{|\boldsymbol{S}_1|} = 0$).