

Chapter 3: Poisson Processes

3.1 Poisson Defn. and Facts

stochastic $\{N(t), t \geq 0\}$ → counting process if $N(t)$ gives # by time t

independent if $N(s) - N(0)$ is independent of $N(10) - N(s)$ events between those times

stationary increments if $N(t) - N(0) = N(t+s) - N(s)$ so only depends on t and not s

counting process $\{N(t), t \geq 0\}$ → Poisson Process w/ rate λ if:

- 1) no events at $t=0$, i.e. $N(0) = 0$
- 2) independent increments, i.e. $N(s) - N(0) \perp\!\!\!\perp N(10) - N(s)$
- 3) stationary increments, i.e. $N(s+t) - N(s) = N(t) - N(0)$
- 4) $P(N(t) = n) = \frac{(\lambda t)^n}{n!} e^{-\lambda t}$, $n = 0, 1, 2, \dots$ $E(N(t)) = \text{Var}(N(t)) = \lambda t$

$\lambda \perp\!\!\!\perp t$ therefore constant and can be referred to as homogenous Poisson Process

interarrival time: time between 2 consecutive occurrences of events $(n-1)^{\text{th}} - n^{\text{th}}$ event → $T_n: \{T_1, T_2, \dots, T_n\}$ w/ $T_1 \sim E_1$

waiting time: $S_n = T_1 + T_2 + \dots + T_n$ → time at event n

Prop 3.1: Interarrival times $T_n, n=1, 2, \dots$ are independent exp r.v. w/ density $f_{T_n}(t) = \lambda e^{-\lambda t}, t \geq 0$

Proof: note $P(T_1 > t) = P(N(t) = 0) = e^{-\lambda t}$ so $T_1 \sim \text{Exp}(\lambda)$

$$\text{so } P(T_2 > t) = \int_0^t P(T_2 > t + T_1 = s) \lambda e^{-\lambda s} ds = \int_0^t P(N(t+s) - N(s) = 0 \mid N(s) = 1) \lambda e^{-\lambda s} ds = \int_0^t P(N(t) = 0) \lambda e^{-\lambda s} ds = e^{-\lambda t} \int_0^t \lambda e^{-\lambda s} ds = e^{-\lambda t} \text{ so } T_2 \sim \text{Exp}(\lambda)$$

therefore by using same logic for T_n or induction will show $P(T_{n+1} > t) = e^{-\lambda t} \rightarrow \prod T_1 \rightarrow T_n \rightarrow f_{T_n}(t) = \lambda e^{-\lambda t}$

Remark 3.1: exp(x) → memoryless and therefore $\text{pois}(\lambda)$ or T_n should inherit that

thought: $(\lambda = 2 \text{ per hour then } E(e) = 1/2 \text{ an hour})$

Prop 3.2: $S_n \sim \text{Gamma}(n, \lambda)$ so $E(S_n) = \frac{n}{\lambda}$, $\text{Var}(S_n) = \frac{n}{\lambda^2}$ density $f_{S_n}(s) = \frac{\lambda^n s^{n-1}}{(n-1)!} e^{-\lambda s}, s \geq 0$

Proof: $S_n = T_1 + \dots + T_n$ $\prod \text{exp}(\lambda) = \text{gamma}(n, \lambda)$

Remark 3.2: Poisson Process → continuous-time Markov Chain

Remark 3.3: Each event can only happen by itself like can't have 2 at once

Example 3.1: Poisson: ppl entering bank Not Poisson: bus time arrivals (schedule)

Example 3.2: Tour bus w/ 50 ppl come every 15 minutes ind. and exp.

a) $E(S_5) = \frac{5}{\lambda} = 1.25 \text{ hours}$ $\text{Var}(S_5) = \frac{5}{\lambda^2} = \frac{5}{16} \text{ hours}$

b) $E(T \text{ between } 11 \rightarrow 1:30) = E(2.5 \text{ hrs}) = 10$

Prop 3.3: can break $\text{Pois}(\lambda)$ into $N(t) = \overset{p}{N_1(t)} + \overset{1-p}{N_2(t)}$ $N_1 \sim \text{Pois}(\lambda p)$ $N_2 \sim \text{Pois}(\lambda(1-p))$

but N_1 and N_2 are independent, with both being thinned from the superposition $N(t)$

Application 3.2: Sports Analytics of Goal Opening

assume points scored by team A $\sim \text{pois}(\lambda_A) \sim \{N_A(t), t \geq 0\}$

and team B $\sim \text{pois}(\lambda_B) \sim \{N_B(t), t \geq 0\}$

have to assume independence

a) superposition $N(t) = N_A(t) + N_B(t)$ w/ rate $\lambda_A + \lambda_B$ for either team to score

Ken $\lambda_A \rightarrow$ Team A scores, $\lambda_B \rightarrow$ Team B scores, $(\lambda_A + \lambda_B)$ either team score

assume team A every 10 min and team B every 12 minute $E(\text{either team score}) = \frac{1}{10} + \frac{1}{12} = 5.45$

b) Probability of first team scoring: $E(T_A) = \frac{1}{\lambda_A}$ $E(T_B) = \frac{1}{\lambda_B}$

$$P(\text{team A scores before team B}) = P(T_A < T_B) = \int_0^\infty P(T_B > t) \lambda_A(t) dt = \int_0^\infty e^{-\lambda_B t} \lambda_A e^{-\lambda_A t} dt = \frac{\lambda_A}{\lambda_A + \lambda_B} = \frac{10}{10+12} = .545$$

$$P(\text{B before A}) = P(T_B < T_A) = \frac{\lambda_B}{\lambda_A + \lambda_B} = 1 - .545 = .455$$

c) tie, team A wins, team B wins

$T \rightarrow$ end of game

$$P(\text{game ties}) = \sum_{n=0}^{\infty} P(N_A(T) = n, N_B(T) = n) = e^{-(\lambda_A + \lambda_B)T} \sum_{n=0}^{\infty} \frac{(\lambda_A \lambda_B T^2)^n}{(n!)^2}$$

$$P(\text{team A wins}) = P(N_A(T) > N_B(T)) = 2 \sum_{n=0}^{\infty} P(N_A(T) = n, N_B(T) = n) = e^{-(\lambda_A + \lambda_B)T} \left[\sum_{n=1}^{\infty} \frac{(\lambda_A \lambda_B T^2)^n}{n!} \cdot \sum_{k=0}^{\infty} \frac{(\lambda_A T)^k}{(n+k)!} \right]$$

substitute for team B or 1 - ties - A = B

$$\text{if } T=60, P(\text{ties}) = e^{-(10+12)(60)} \sum_{n=0}^{\infty} \frac{((10)(12)(60)^2)^n}{(n!)^2} = e^{-11} \sum_{n=0}^{\infty} \frac{30^n}{(n!)^2} = 0.1166$$

$$P(A) = e^{-11} \left[\sum_{n=1}^{\infty} \frac{30^n}{n!} \cdot \sum_{k=0}^{\infty} \frac{10^k}{(n+k)!} \right] = 0.5590$$

$$P(B) = 1 - \text{ties} - A = 0.5244$$

Application 3.3: Poisson pedestrian vs. traffic flow

Exercise 3.1: $\{N(s), t \geq 0\}$ $\text{Pois}(\lambda)$, joint pdf $P(N(s)=m, N(t)=n), t \geq s \geq 0, n \geq m \geq 0$

$$P(N(s)=m, N(t)=n) = P(N(t)-N(s) = n-m, N(s)=m) = P(N(t)-N(s) = n-m) P(N(s)=m) =$$

$$P(N(t-s) = n-m) P(N(s) = m) = \frac{(\lambda(t-s))^{n-m}}{(n-m)!} e^{-\lambda(t-s)} \cdot \frac{(\lambda s)^m}{m!} e^{-\lambda s} = \frac{(t-s)^{n-m} e^{-\lambda(t-s)}}{(n-m)! m!} \lambda^m e^{-\lambda t}$$

$$\text{Exercise 3.2: } \text{Cov}(N(s), N(t)) = E[N(s)N(t)] - E[N(s)]E[N(t)]$$

$$= \text{through terrible properties} \rightarrow E[N(t-s)]E[N(s)] + \text{Var}[N(s)] + [E(N(s))]^2$$

$$\rightarrow \lambda(t-s)\lambda + \lambda + (\lambda s)^2 \quad (1)$$

$$= (1) - (2) = \lambda s$$

Exercise 3.4: 5 min per call, .15 for buying

a) 2 hours $\rightarrow 5 \text{ min} \rightarrow 24 \cdot .15 = 3.6$

b) within 60 $\rightarrow 15$ calls $\rightarrow 5$ successes $P(N(1)=15, N_s(1)=5) \rightarrow P(N_s=5, N_{ns}=10) = \frac{(1.5)^5}{5!} e^{-1.5} \frac{(10.5)^{10}}{10!} e^{-10.5} = .003$

c) $P_4=10, P_1=3 \rightarrow P_3=7 \rightarrow \frac{(1.5)^7}{7!} e^{-1.5} = 0.119$