# 310 Quals Strategy Compendium

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## 1 Conditional Expectation

#### **Definition 1** (Absolute continuous measures)

We say that  $\nu$  is absolutely continuous wrt  $\mu$ , ie  $\nu \ll \mu$  if

$$\mu(A) = 0 \implies \nu(A) = 0$$

### **Theorem 2** (Radon-Nikodym Theorem)

If  $\nu \ll \mu$ , both sigma finite, then there exists  $f \in m\mathcal{F}_+$  finite valued such that  $\nu = f\mu = \int f d\mu$ .

Ie existence of a density wrt dominating measure. f is called the Radon-Nikodym derivative and we write  $f = \frac{d\nu}{d\mu}$ .

We use the above to prove the existence of CE.

### **Definition 3** (Conditional Expectation - CE)

Given  $X \in L^1(\Omega, \mathcal{F}, P)$  and  $\mathcal{G} \subset \mathcal{F}$ , then there exists  $Y \in L^1(\Omega, \mathcal{G}, P)$ :

$$Y := \mathbf{E}[X|\mathcal{G}]$$
 such that  $\mathbf{E}[(X - Y)\mathbf{1}_G] = 0 \quad \forall G \in \mathcal{G}.$ 

Defined uniquely for P-almost every  $\omega$ .

To check that something is a CE, use the following:

### **Theorem 4** (Check CE on $\pi$ system)

Ex 4.1.3 Dembo. If  $\mathcal{G} = \sigma(P)$  and P a  $\pi$  system, then if the above holds for every  $G \in P$ , then  $Y = \mathbf{E}[X|\mathcal{G}]$ .

That is, we can just check for every set in a **generating**  $\pi$ -system.

### Recipe 5 (Prove something is CE)

Do the following

- 1. Check that  $X \in L^1$ .
- 2. Check that candidate  $Y \in L^1$ .
- 3. Show that X and Y integrate  $\mathcal{G}$ -sets the same, possibly using the pi system fact.

A note of caution that  $Y_{\pm} \neq \mathbf{E}[X_{\pm}|\mathcal{G}]$  in general.

### 1.1 Properties of CE

If  $X \in L^1(\Omega, \mathcal{F}, P)$ :

- 1. If also  $X \in L^1(\Omega, \mathcal{G}, P)$ , then  $\mathbf{E}X|\mathcal{G} = X$ .
- 2. Drop conditioning if independent:
  - (a) if  $\mathcal{H} \perp \!\!\!\perp \sigma(\mathcal{X})$ , then  $\mathbf{E}[X|\mathcal{H}] = \mathbf{E}X$ .
  - (b) If  $\mathcal{H} \perp \!\!\!\perp \sigma(\sigma(X), \mathcal{G})$ , then  $\mathbf{E}[X|\sigma(\mathcal{H}, \mathcal{G})] = \mathbf{E}[X|\mathcal{G}]$
- 3.  $X \ge 0 \implies Y \ge 0$  a.s., and  $X > 0 \implies Y > 0$  a.s.
- 4. Linearity
- 5. Monotonicity: if  $X_1 \leqslant X_2$  then  $\mathbf{E}X_1 | \mathcal{G} \leqslant \mathbf{E}X_2 | \mathcal{G}$ .
- 6. If  $\mathbf{E}X|Y = Y$  and  $\mathbf{E}Y|X = X$  then X = Y.
- 7. Tower
  - (a)  $\mathbf{E}X = \mathbf{E}[\mathbf{E}(X|\mathcal{G})]$
  - (b) If  $\mathcal{H} \subset \mathcal{G} \subset \mathcal{F}$ , then  $\mathbf{E}[X|\mathcal{H}] = \mathbf{E}[\mathbf{E}(X|\mathcal{G})|\mathcal{H}]$ . "The small one stays on the outside"
- 8. If  $\mathbf{E}[X|\mathcal{G}] \perp \!\!\! \perp X$ , then  $\mathbf{E}X|\mathcal{G} = \mathbf{E}X$ .
- 9. Take-out: if  $Y \in m\mathcal{G}$ , then  $\mathbf{E}[XY|\mathcal{G}] = Y\mathbf{E}[X|\mathcal{G}]$ .
- 10. Conditional Jensen
- 11. Conditional Markov, Holder
- 12. MCT, Fatou, DCT for CE
- 13. If  $X_n \stackrel{L^q}{\to} X_{\infty}$ , then  $\mathbf{E}[X_n | \mathcal{G}] \stackrel{L^q}{\to} \mathbf{E}[X_{\infty} | \mathcal{G}]$  (apply Jensen)
- 14.  $\{\mathbf{E}[X|\mathcal{H}]: \mathcal{H} \subset \mathcal{F} \text{ is a sigma algebra}\}\$ is UI. (4.2.33).

# 1.2 CE as Orthogonal Projection

If  $X \in L^2$ , then  $Y = \mathbf{E}[X|\mathcal{G}]$  is the unique  $Y \in L^2$  such that  $||X - Y||_2 = \inf\{||X - W||_2 : W \in L^2(\Omega, \mathcal{G}, P)\}$ . Ie, the conditional expectation is a projection onto the subspace with respect to the  $\langle X, Y \rangle = \int XYdP$  inner product. (Since  $L^2(\Omega, \mathcal{G}, P)$  is a *Hilbert Subspace*).

There exists a unique projection in Hilbert Spaces onto subspaces, ie  $\langle h - \hat{h}, f \rangle = 0$  where  $\hat{h} = \operatorname{Proj}_{L^2(\Omega, \mathcal{G}, P)} h$  and  $f \in L^2(\Omega, \mathcal{F}, P)$ .

**Theorem 6** (Cauchy-Schwarz in  $L^2$ )

$$|\mathbf{E}XY| \leqslant \sqrt{\mathbf{E}X^2\mathbf{E}Y^2}$$

# 1.3 Regular conditional probabbility distributions (RCPD)

# Theorem 7 (Can take conditional expectation wrt conditional density)

If X, Z have a joint density and g(X) is integrable, then

$$\mathbf{E}[g(X)|Z] = \int_{\mathbb{R}} g(x) f_{X|Z}(x|z) dx$$

as in elementary probability.

### **Definition 8** (RCPD (Regular conditional probability distribution))

Let Y be  $\mathbb{S}$ -valued random variable then

$$\hat{P}_{Y|\mathcal{G}}(\cdot,\cdot): \mathcal{S} \times \Omega \mapsto [0,1]$$

is the RCPD of Y given  $\mathcal{G}$  if:

- 1.  $\hat{P}_{Y|\mathcal{G}}(A,\cdot)$  is a version of the CE  $\mathbf{E}[\mathbf{1}[Y \in A|\mathcal{G}]]$  for  $A \in \mathcal{S}$ .
- 2. For any fixed  $\omega$ ,  $\hat{P}_{Y|\mathcal{G}}(\cdot,\omega)$  is a probability measure on  $(\mathbb{S},\mathcal{S})$ .

Analogue to Markov Kernel in 310a.

# **Theorem 9** (RCPD existence)

If X real and  $\mathcal{G}$  a  $\sigma$ -algebra, then RCPD exists.

Also true for any  $\mathcal{B}$ -isomorphic rv X.

Also used to show existence of transition probability- see Ex 4.4.5.

A helpful exercises 4.4.6 shows we can calculate expectations using the RCPD:

$$\mathbf{E}[h(X,Y)|\mathcal{G}](\omega) = \int_{\mathbb{R}} h(x,Y(\omega))d\hat{P}_{X|\mathcal{G}}(x,\omega)$$

ie fix  $\omega$  and integrate over  $x \in \mathbb{R}$ .

### 2 Martingales

To check a martingale:

- 1.  $X_n$  is integrable for all n
- 2. Adapted
- 3.  $\mathbf{E}[X_{n+1}|\mathcal{F}_n] = X_n$  for all n

Alternatively, if  $X_n = \sum_{k=0}^n D_k$ , then check that  $\mathbf{E}[D_{n+1}|\mathcal{F}_n] = 0$ .

#### **Example 10** (Quadratic martingale)

If  $\mathbf{E}\xi_i = 0$  and  $\mathbf{Var}\xi_i = \sigma^2 < \infty$  and the  $\xi_i$  are independent, then

$$S_n^2 - n\sigma^2$$
 is a martingale

#### **Example 11** (Exponential martingale)

If  $S_n$  random walk with independent, iid increments,

$$M_n = \prod_{i=1}^n \exp(\theta \xi_i) / \phi(\theta) = \exp(\theta S_n) / \phi(\theta)^n$$

Is a special case of product martingale

A predictable sequence of random variables  $A_n$  gives the amount of money you'd be willing to bet at time n- must be based on information from previous time points, up through n-1, to make the n-th bet. Think of  $A_n$  as your n-th bet- given by information from time  $\{1, \ldots, n-1\}$ . Can think of winnings in the following decomposition:

$$\sum_{m=1}^{n} H_m(X_m - X_{m-1})$$

where  $H_m$  is the amount you wager between days m and m+1  $X_m$  is the stock price. So our profit is the difference in prices times the amount that we bet/number of shares we hold.

The above is known as the martingale transform of  $\{X_n\}$  by the predictable process  $\{A_n\}$ .

### Fact 12 (Martingale transforms)

MG transforms of  $mgs \implies mg$ 

 $Mg transforms of sub/sup mgs \implies sub/sup mg$ 

**Example 13** (Conditions of a.s. convergence of (sub/sup) mg don't necessarily give  $L^1$  convergence)

Durrett 193 - example with  $S_n$  simple random walk starting at  $S_0 = 1$ ., let  $X_n = S_{\tau \wedge n}$  where  $\tau$  is first time  $S_n = 0$ . Then can  $X_n$  is a non-negative martingale so a.s. limit exists, must be  $X_n \xrightarrow{\text{a.s.}} 0$ . But  $\mathbf{E}X_n = 1$ . So  $L^1$  convergence can't occur.

### **Example 14** (Martingale converging as to $-\infty$ )

Construct something such that the positive event happens only finitely often.  $X_n \xrightarrow{\text{a.s.}} -\infty$  even though  $X_n$  is a fair bet. There's a remote chance of a big reward.

### **Theorem 15** (Martingale with bounded increments converges or oscillates between $\pm \infty$ )

(Durrett 4.3.1)

If  $X_n$  mg with  $|X_{n+1} - X_n| \leq M < \infty$ , then if

$$C = \{\lim X_n < \infty \text{ exists}\}$$
  $D = \{\lim \sup X_n = \infty, \lim \inf X_n = -\infty\}$ 

Then:

$$P(C \cup D) = 1.$$

So given a martingale, any  $\omega$  in a set of prob 1 must be do one of these two things.

### Example 16 (Biased random walk)

If positive step is  $p \neq 1/2$ , then

$$X_n = \left(\frac{q}{p}\right)^{S_n}$$
 is a martingale

### Theorem 17 (Wald's Equation)

### Theorem 18 (Doob Decomposition)

For any  $\{X_n\}$  stochastic process adapted to  $\{\mathcal{F}_n\}$ , write:

$$M_n = \sum_{k=0}^{n-1} (X_{k+1} - \mathbf{E}(X_{k+1}|\mathcal{F}_k)) \quad A_n = \sum_{k=0}^{n-1} (\mathbf{E}(X_{k+1}|\mathcal{F}_k) - X_k)$$

so that

$$X_n = X_0 + M_n + A_n,$$

where  $M_n$  is a martingale and  $A_n \in m\mathcal{F}_{n-1}$  ie is adapted. In summary, an adapted (discrete) stochastic process can be written as the sum of a martingale and a predictable process.

In the case that  $X_n$  is a submartingale,  $A_n$  is non-negative and increasing.

#### Theorem 19 (BC2 version 2)

If  $\{B_n\}$  sequence of events, then:

$$\{B_n \text{ i.o.}\} = \{\sum_{n=1}^{\infty} P(B_n | \mathcal{F}_{n-1}) = \infty\}.$$

*Proof.* Idea is that  $X_n = \sum_{m \leq n} \mathbf{1}[B_m]$  is a submartingale, apply the Doob's decomposition and then note that:

$$|M_n - M_{n-1}| \leqslant 1$$

so we can apply the theorem with bounded martingale differences to get that  $P(C \cup D) = 1$ , show that this means that the two events are the same.

Note that Dembo has an extra comment about the rate at which  $X_n$  goes to infinity.

### Example 20 (Polya's Urn)

(Durret section 4.3.2)

Contains r red and g green balls— each time we draw a ball, we replace it and add c more of the same color. Let  $G_n$  is # of green balls,  $X_n$  is the fraction of green balls after the n-th draw, ie  $G_n/N_n$ .

- 1.  $X_n$  is a non-negative martingale
- 2.  $X_n \xrightarrow{\text{a.s.}} X_{\infty}$  as
- 3. If b = g = 1, then  $P(G_n = m + 1) = \frac{1}{n+1}$ , ie uniform on  $\{1, \dots, m+1\}$ .
- 4.  $X_{\infty}$  then has a uniform distribution on (0,1).
- 5. In general,  $X_{\infty} \stackrel{d}{=} \text{Beta}(g/c, r/c)$ .

#### Example 21 (Likelihood ratios)

Durrett..

### 2.1 Branching process

If  $\mu = \mathbf{E}\xi_i^{(m)}$  where the  $\xi_i^{(m)}$  are the number of offspring of the *i*-th member of the *m*-th generation, then:

$$Z_n/\mu^n$$
 is a martingale

If  $\mu < 1$  or  $\mu \le 1$  and  $P(\xi_i = 1) < 1$ , then the population dies out, ie  $Z_n/\mu^n \xrightarrow{\text{a.s.}} 0$ . "If the average number of offspring is fewer than 1, then the population dies out. For  $\mu > 1$ , we use *generating functions* and we can prove things about the limiting random variables.

### 2.2 Doob's maximal inequality and $L^p$ convergence

Think of Doob's inequality as an improvement on Markov's for (sub)-martingales.

#### Theorem 22 (Doob's Inequality)

For submartingale  $\{X_n\}$ , we have:

$$P(\max_{k \leqslant n} X_k \geqslant \lambda) \leqslant \lambda^{-1} \mathbf{E} X_n^+.$$

Kolmogorov's Maximal inequality is a special case. If  $Z_n = \sum_{i=1}^n Y_i$  with  $\mathbf{E}Y_i = 0$ ,  $\mathbf{Var}Y_i < \infty$ , then  $Z_n$  is obviously a mg and

$$P(\max_{k \leqslant n} |Z_n| > \lambda) = P(\max_{k \leqslant n} Z_n^2 > \lambda^2) \leqslant \frac{1}{\lambda^2} \mathbf{Var} Z_n.$$

### **Theorem 23** ( $L^p$ maximal inequality for p > 1)

If  $\{X_n\}$  submg,

$$\mathbf{E}(\max_{k \le n} (X_n^p)_+) \le \left(\frac{p}{p-1}\right)^p \mathbf{E}(X_n^+)^p.$$

So if  $\{X_n\}$  is a mg,

$$\mathbf{E}(\max_{k \leqslant n} |X_n|)^p \leqslant \left(\frac{p}{p-1}\right)^p \mathbf{E}|X_n|^p.$$

The previous  $L^p$  inequality leads to:

### **Theorem 24** ( $L^p$ convergence theorem)

If  $X_n$  is a martingale with  $\sup \mathbf{E}|X_n|^p < \infty$  with p > 1, then  $X_n \xrightarrow{\text{a.s.}} X$  and  $X_n \xrightarrow{\text{L}^p} X$ .

*Proof.* Almost sure convergence comes from the Doob's convergence theorem and Markov.  $L^p$  convergence comes from  $L^p$  maximal inequality, see Durrett 4.4.6.

Note that this theorem:

- 1. Is only for martingales
- 2. Does not have a  $L^1$  analog

### 2.3 Square Integrable Martingales

#### Fact 25 (Square Integrable martingales (Dembo Ex 5.1.8))

Have uncorrelated differences  $D_n$ . Durrett 4.4.7.

Suppose that  $X_n$  is an  $L^2$  martingale with  $X_0 = 0$ . Then:

- 1.  $X_n^2$  is a submartingale
- 2. Apply Doob's Decomposition theorem  $X_n^2 = M_n + A_n$  with  $M_n$  mg and  $A_n$  is a **predictable**, increasing sequence, called the **predictable compensator**.
- 3.  $\langle X \rangle_n := A_n = X_0^2 + \sum_{m=1}^n \mathbf{E}(X_m^2 | \mathcal{F}_{m-1}) X_{m-1}^2 = \sum_{m=1}^n \mathbf{E}((X_m X_{m-1})^2 | \mathcal{F}_{m-1})$
- 4. Think of the predictable compensator as the total variance at time n of the path made by  $\{X_n\}$ .

Theorem 26 (Finite predictable compensator limit means finite limit of the original martingale)

If  $\langle X \rangle_{\infty} < \infty$  then we have that  $X_n \xrightarrow{\text{a.s.}} X_{\infty} < \infty$ .

#### Recipe 27 (Convergence of random series)

Dembo 5.3.37. Want to show that  $\sum_{n=1}^{\infty} X_n(\omega)$  converges. Do we need symmetric distribution of  $X_n$ 

#### Example 28 (Quals second q this week)

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