

The Effects of Alphabet and Expertise on Letter Perception

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Long-standing questions in human perception concern the nature of the visual features that underlie letter recognition and the extent to which the visual processing of letters is affected by differences in alphabets and levels of viewer expertise. We examined these issues in a novel approach using a same–different judgment task on pairs of letters from the Arabic alphabet with 2 participant groups: 1 with no prior exposure to Arabic and 1 with reading proficiency. Hierarchical clustering and linear mixed-effects modeling of reaction times and accuracy provide evidence that both the specific characteristics of the alphabet and observers' previous experience with it affect how letters are perceived and visually processed. The findings of this research further our understanding of the multiple factors that affect letter perception and support the view of a visual system that dynamically adjusts its weighting of visual features as expert readers come to more efficiently and effectively discriminate the letters of the specific alphabet they are viewing.

Keywords: letter perception, Arabic alphabet, reading, expertise, visual features

Experimental psychologists have investigated letter perception since the founding days of the discipline (Cattell, 1886; Javal, 1881). Letters have been of such interest because of their critical role in reading and because—as relatively simple visual objects—their study can provide insights into object recognition more generally. According to feature detection theories,¹ a key component of object recognition is the response of visual cortex to specific features of a visual stimulus (for a review, see Palmer, 1999; for letter perception specifically, see Grainger, Rey, & Dufau, 2008). Consistent with this view, Pelli, Burns, Farell, and Moore-Page (2006) presented arguments that human readers employ a feature-detection approach to letter recognition as opposed to other approaches such as template matching. Within this general framework, considerable research has been directed at identifying the set of features used in letter recognition, with recent work by Fiset and colleagues (2008, 2009) arguing for a specific set of 10 visual features for letter recognition, most of which had support from previous research.

The vast majority of the work on visual feature processing in letter recognition has involved proficient adult readers of languages that use the Roman alphabet (but see Changizi, Zhang, Ye,

& Shimojo, 2006; Gibson, Gibson, Pick, & Osner 1962; Pelli et al., 2006; Treiman & Kessler, 2011; Wang & Cavanagh, 1993). Therefore, it is unclear from the literature the extent to which the visual feature set and visual processes used in letter recognition are the same across different alphabets and different levels of expertise or proficiency. In this regard, previous work has shown that performance on tasks involving letter perception may be generally affected by experience and various knowledge-based factors (Egeth & Blecker, 1971; Wang, Cavanagh, & Green, 1994), and, furthermore, there is evidence that literacy can affect activation in occipital cortex (Dehaene et al., 2010). However, the experimental designs and analytic techniques employed have not usually allowed researchers to determine whether these factors affect visual feature processing itself or whether they affect other aspects of performing the particular experimental tasks. In the research reported here, we use the novel approach of applying linear mixed-effects modeling (LMEM; Baayen, Davidson, & Bates, 2008; Barr, Levy, Scheepers, & Tily, 2013) to the analysis of reaction time (RT) and accuracy data from same–different judgments of Arabic letters carried out by naïve and expert readers of Arabic. This allows us to make progress on several fronts. First, the use of the Arabic alphabet allows us to compare visual processing of Arabic and Roman letters (reported in Fiset et al., 2008, 2009). Second, the use of viewers with different levels of expertise allows us to examine the effects of experience on visual processing of letters. Third, the analytic approach adopted allows us to better localize the observed behavioral effects to visual processes themselves. In this way, this research allows us to investigate if and

This article was published Online First February 25, 2016.

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We thank Howard Egeth for guidance with the literature and input on a draft and the support of National Institutes of Health Grant DC012283 to Brenda Rapp. Robert W. Wiley and Brenda Rapp developed the experimental questions and experimental design. Data collection was carried out by R. Wiley, who performed the statistical analyses in collaboration with Colin Wilson. The manuscript was written by B. Rapp and R. Wiley with input from C. Wilson.

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¹ Feature detection theories are distinguished from alternatives such as template-matching or exemplar-based accounts, which propose that objects are recognized by a process that compares the whole object to a prototype or template, rather than making use of local features within the stimulus. Although ideal observer analyses have shown that template matching may be more efficient than feature detection, empirical findings indicate that normal literate adults may not adopt this approach (Fiset et al., 2008; Pelli et al., 2006; Tjan, Braje, Legge, & Kersten, 1995).

how the visual processing of letters is consistent (universal) across stimulus sets and viewer expertise. Addressing these issues is not only important for a deeper understanding of the cognitive machinery of letter identification, it is also relevant for guiding the teaching of first- and second-language reading as well as the rehabilitation of reading in the visually impaired or those with acquired perceptual deficits resulting from neural injury.

On Notions of Universality

It is important to indicate from the outset that there are various positions regarding the universality of visual feature processing across stimulus sets and tasks, and the effects of learning on visual feature representation and processing. Briefly, as indicated by Schyns, Goldstone, and Thibaut (1998), one general position is that the visual feature set is fixed, and another is that it is flexible and can be rapidly affected by learning (e.g., Pelli et al., 2006) and shaped by the specific demands of the learning task and other factors (e.g., alphabet and expertise). However, even if we assume that the feature set is flexible, there are still a variety of ways of thinking about that flexibility. For example, one is that learning may affect the set of visual features per se by adding or subtracting features or otherwise changing the nature of the features (for a review, see Schyns et al., 1998). Another is that the feature set itself is actually fixed, but that the ranking or relative importance of features can be affected by learning and experience (e.g., Haider & Frensch, 1996; Livingston & Andrews, 1995). It is not clear how, or even if, it is possible to distinguish between these two interpretations of the flexibility view. This is especially true if one assumes a fixed feature set with a possible ranking = 0. In that case, evidence of “new” features would be interpreted as moving the ranking of an existing feature to a value greater than zero. Despite these interpretative difficulties, it is nonetheless possible to further our understanding of whether and how alphabet and expertise affect visual feature processing during letter recognition. Evidence that these factors do affect letter feature processing would, at a general level, provide further evidence of the dynamic nature of the visual system and, at a specific level, provide insight into some of the particular characteristics of visual letter processing and learning.

The Challenge of Isolating Visual Feature Processing: Effects of Experience

Because any behavioral task involves numerous processing stages and representational types, the perennial challenge for cognitive research is to identify, among the many processes involved in a task, the one(s) affected by the experimental manipulations or factors of interest. In this article, we make a distinction between “knowledge-based factors” as effects that stem from the activation of knowledge that is associated with the stimuli (such as phonological names and motoric stroke patterns) but unrelated to their visual feature processing, and “visual expertise” as effects of extensive experience with visual processing of the stimuli. The problem is that either of these can affect performance on tasks commonly used to investigate visual perception, and the challenge is to distinguish these types of factors from one another.

In terms of the question of the influence of expertise on visual processing, despite the common assumption that the basic machin-

ery of vision is universal, there is also evidence that experience with a stimulus set can influence performance on perceptual tasks (e.g., for a review, see Schyns et al., 1998; and see also Delorme, Rousselet, Macé, & Fabre-Thorpe, 2004, Grill-Spector & Kanwisher, 2005, and Schyns & Oliva, 1999; and for letters specifically, see Lupyan, Thompson-Schill, & Swingle, 2010). For example, there is considerable evidence that with overlearned stimuli (such as letters, words, or faces; Wolfe, 2000), task-irrelevant information may become active and affect performance. Along these lines, neuroimaging studies involving visually presented letters have shown activity in neural areas that represent the motor programs used to write letters (James & Atwood, 2009) as well as activity related to the abstract identity of letters and their spoken names (Rothlein & Rapp, 2014). However, although effects of familiarity and training on perceptual judgments and processing have been well documented, it still remains unclear exactly how early in visual processing these effects of learning extend (see Wolfe, 2000, for a review); in other words, it is unclear whether these observed effects of learning and expertise are because of changes in visual processing per se.

The literature reveals inconsistent results regarding possible influence of what we are referring to as knowledge-based factors on letter perception. Simpson, Mousikou, Montoya, and Defior (2013), using subjective similarity ratings for letter pairs, concluded that any effects of letter-sound or identity knowledge were “not strong enough to mask the effect of visual similarity” (p. 436). Consistent with this, Courrieu, Farioli, and Grainger (2004) found no evidence of phonological encoding affecting perceptual judgments in simultaneous same–different judgments. Lupyan et al. (2010) also reported no evidence of letter identity knowledge on simultaneous same–different judgments (with stimuli consisting of the pairs Bb, Bp, BB, and bb), but did find such effects with stimulus onset asynchronies greater than zero. In contrast, others have argued for the influence of nonvisual knowledge on various visual tasks. In work following up on Lupyan et al., Chen and Proctor (2012) found the influence of letter identity information both with simultaneous and delayed presentation. Mueller and Weidemann (2012) also found evidence of the influence of abstract letter category in experiments in which participants had to identify target letters between pre- and postmasks while the content of the masks was manipulated. With regard to other knowledge-based factors, Treiman and Kessler (2003) reported that the interaction between the visual similarity of letters and the phonological similarity of their names influences the learning of letters by children. In short, even this brief, nonexhaustive review indicates that findings regarding the influence of factors that are not strictly visual in nature on visual processing of letters are mixed. Given this state of the literature, it cannot be safely assumed that knowledge-based factors will not influence the perceptual judgments that form the basis of researchers’ inferences regarding the effects of expertise and alphabet on visual processing. Clearly, a better understanding of the role of these factors is needed and, if effects of knowledge-based factors are indeed present, it would be helpful to have a means for extracting their contributions from the data obtained in perceptual tasks.

In this research, we adopt the regression analysis approach of LMEM that is well suited to statistically extracting the contribu-

tions of experience-based factors in order to evaluate their significance and more successfully isolate visual feature processing itself. We specifically consider the influence of the following knowledge-based factors: letter name similarity, motor similarity, alphabetic order, abstract identity, and Roman alphabet similarity (the perceived similarity of Arabic letters to Roman letters). Most of these could be relevant only to experienced readers of Arabic (e.g., letter name, motor similarity, alphabetic order), whereas others (e.g., Roman alphabet similarity) may be relevant to both groups of participants. By including these variables in the statistical models, we are able to get a view of the visual feature processing in letter identification that is less biased by the influence of these knowledge-based factors.

Visual Feature Processing in Letter Identification

With regard to the universality of visual feature processing across alphabets, Pelli et al. (2006) compared the performance of human observers in identifying letters (from a number of different alphabets) embedded in Gaussian noise with an ideal observer and concluded that the mental computations subserving object recognition are the same for letters as for any other object, and that, *regardless of the alphabet*, letter identification can be achieved by the detection of approximately seven independent visual features. In the research we report on here, the choice of Arabic letter stimuli will allow us to determine how well the specific features proposed on the basis of the Roman alphabet by Fiset et al. (2008, 2009) generalize to Arabic.

In Wolfe's (2000) review of research on the basic feature set of vision, he identified eight to 10 basic feature dimensions (color, orientation, motion, size, curvature, depth, Vernier offset, gloss, intersection and spatial position; see also Palmer, 1999). In terms of specific features used in letter identification, using an image classification approach with Roman letters (the Bubbles method; Gosselin & Schyns, 2001), Fiset et al. (lowercase [LC] and uppercase [UC] letters: Fiset et al., 2008; UC letters only: Fiset et al., 2009) proposed a set of 10 visual features for letters: lines (horizontal, vertical, oblique right and left), curves (open to the top, bottom, left, and right), intersections, and terminations—effectively a subset of those enumerated by Wolfe. Fiset et al. (2009) especially underscored the role of horizontal lines and terminations, proposing,

On this basis, we are confident that the prime importance of terminations and horizontals, in particular, and the relative importance of the other features for human participants is due to constraints imposed by the human visual system rather than constraints imposed by the stimuli or analyses. (p. 33)

Thus, they seem to propose a fundamental role for these features that is determined by the nature of the visual system itself and might not be expected to vary with alphabet (stimulus set) or expertise. Although methodologically creative and sound, the Bubbles method does require the presentation of degraded stimuli, leaving open the question of how well the Fiset et al. feature set generalizes to normal viewing conditions.

Previous work on visual features and early visual processes has involved a wide range of tasks, including visual search (Malinowski & Hübner, 2001; Treisman & Gelade, 1980), same–different judgments (e.g., Courrieu et al., 2004; Podgorny & Gar-

ner, 1979), letter naming (e.g., Gervais, Harvey, & Roberts, 1984; Keren & Baggen, 1981), and explicit similarity ratings (e.g., Boles & Clifford, 1989; Simpson et al., 2013). The research we report on here shares with many of these approaches the assumption that letter confusability reveals similarity and that, in turn, similarity reveals underlying visual feature representations. In other words, it is assumed that two letters are confused with one another (or are difficult to discriminate; e.g., A and R) because they share visual features. In this approach, letter confusability matrices are typically developed and the similarity structure of the matrices is then evaluated to infer underlying visual features.

However, the various tasks used to generate confusion matrices have specific strengths and weaknesses. For example, although letter-naming paradigms have strong ecological validity, errors rates on these tasks are usually very low, and to increase error rates, single-letter stimuli are typically presented for naming under abnormal visual conditions such as eccentric presentation, visual degradation, brief exposure duration, and so forth (for a review, see Mueller & Weidemann, 2012). This raises the concern that the visual feature sets derived from confusions elicited under abnormal viewing conditions might not correspond to those involved under normal conditions (Fiset et al., 2008; Pelli et al., 2006). On the other hand, a task such as explicit visual similarity ratings of letter pairs allows for normal viewing conditions, but may be especially susceptible to strategies and the influence of knowledge (e.g., letter names) that should be irrelevant to judging if two shapes are physically identical or not. In the research we report on here, we have chosen to examine data from speeded same–different judgments for letter pairs, as this task allows for normal viewing conditions while somewhat limiting the use of strategies and the influence of task irrelevant knowledge. The patterns of RT and accuracy obtained from this task can be analyzed to reveal the underlying visual representations that can then be evaluated for effects of alphabet and expertise. We specifically assume that longer response times and lower accuracy for same–different judgments reflect greater discrimination difficulties and, hence, greater similarity.

With regard to specific visual features, we examined 15 feature types selected largely on the basis of the previous literature: oriented lines (horizontal, vertical, slant right, slant left), oriented curves (open left, right, up, down), intersections, terminations, closed space, cyclicity, and symmetry (the latter three proposed by Gibson, 1969, among others), as well as two novel features related to characteristics of the Arabic alphabet: dots (diacritics) and closed curves.² The choice of alphabet, task, participant groups, and analytic approach allowed us to evaluate whether, and to what extent, the visual processing of these features is consistent across levels of expertise and alphabet.

² In Arabic, many letters are distinguished solely on the basis of the number and location of diacritics. As for “closed curves,” there are a number of curves (e.g., و and ف) that are so small that the approach of Fiset et al. (2008, 2009) to treat closed curves (such as in the letter O) as four separate open, oriented curves did not seem appropriate.

The Arabic Alphabet: A Brief Overview

The Arabic alphabet is used today for writing not only the Arabic language but also several other languages with millions of speakers, including Persian and Urdu. The Arabic alphabet is usually considered to have 28 letters representing consonants and three long vowels. Short vowels are indicated only by diacritic marks above or below the consonant after which they occur phonologically and are not usually included in text written for adult readers. Both individual letters and whole words are written in cursive from right to left, with spaces occurring only between words or after a nonligating letter. The ligation, or attachment, between letters is taught as part of the letter form (much as the ligations are taught in cursive letters in the Roman alphabet for letters such as “p” or “w”). However, six of the 28 letters do not attach to the following letter, creating small gaps within the structure of the words. There are two characteristics of this writing system that are highly distinct from the Roman alphabet. First, each letter is described as having a form that encodes whether it is in the initial, medial, final, or isolated position (see Figure 1). Positions do not refer to absolute position in the whole word but rather position within so-called “subwords” (Lorigo & Govindaraju, 2006) that are formed by the gaps between nonligating letters. Another distinctive feature is the use of diacritic marks (dots) to distinguish one letter from another. Several shapes are distinguished only by the number of dots (zero to three) and their location (above, below, or within the shape). Note that this differs from the dots in “i” and “j,”

which are not distinctive. Figure 1 illustrates key characteristics of the Arabic alphabet that will be relevant to this study.

The Current Study

In sum, there has been long-standing interest in understanding the effects of experience and knowledge on visual processing and, specifically, whether visual feature processing itself is affected by these factors. A fundamental challenge for research on this topic has been to identify the contribution of knowledge-based and other factors on the perceptual judgments that serve as the database from which to infer visual processing characteristics. The study we report on here, through the comparison of the performance of naïve and expert viewers of Arabic and the application of LMEM analyses addresses the following questions, both methodological and theoretical: (a) Is LMEM a useful tool in the analysis of perceptual judgment data allowing for the identification of possible knowledge-based factors and, in so doing, providing a clearer view of visual processing? (b) Are there specific knowledge-based factors that influence perceptual judgments of letters? (c) Is visual feature processing comparable for Arabic and Roman letters? And, if not, how do they differ? (d) Is visual processing itself affected by expertise with a stimulus set (alphabet)? The answers to these questions are relevant to our understanding of letter perception and processing and also have implications for visual object processing more generally.

Method

Participants

A naïve group (NG; $n = 25$) consisted of Johns Hopkins University undergraduates with normal or corrected-to-normal vision who reported no knowledge of Arabic or a language with an Arabic script. The NG participants were compensated with course credit. One NG participant was dropped from the analyses for failing to complete the second testing session. An expert group (EG; $n = 25$) consisted of adults literate in Arabic recruited from the wider Johns Hopkins University community. Eligibility was determined by completion of a screener assessing Arabic reading and writing, and participants were compensated with \$10 per session. The group consisted of individuals who either learned Arabic as their first language, as a second language from one or more parent who was a native speaker, or as an academic subject in school. One EG participant was dropped because of experimenter error.

Stimuli

The Arabic alphabet consists of 28 letter identities that correspond to 54 unique shapes when including the allographic variants. To reduce trial numbers, 45 letter shapes were selected such that the excluded shapes differed from included shapes only in the number or location of diacritic marks. The font Adobe Arabic size 24 was used (letters varied from 0.17° to 0.31° , and from 0.05° to 0.35° of visual angle in height and width, respectively). Pilot testing indicated that naïve observers could easily discriminate these stimulus sizes. Five size-matched symbols (tilde, open pa-

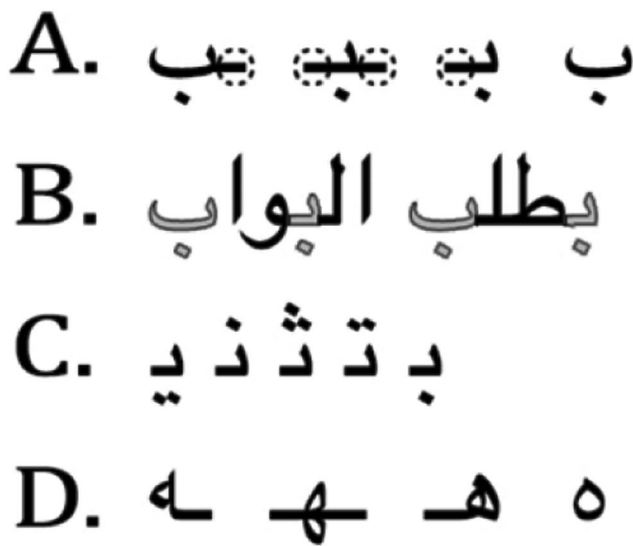


Figure 1. Examples illustrating properties of the Arabic script. (A) From right to left, the isolated, initial, medial, and final forms (allographs) of the letter “ba.” The ligatures are circled, illustrating that—apart from the ligatures—the shapes of the isolated and final forms, and initial and medial forms, are identical. (B) Four forms of the letter “ba” (in gray) embedded within words: initial and final forms on the right (“bitalab”), or medial and isolated forms on the left (“albawwab”). This also illustrates the spacing structure both within and across words. (C) One of the sets of letters (“ya,” “nuun,” “tha,” “ta,” and “ya”), differentiated only by their diacritics. (D) One of the three letters that have four different shapes even after accounting for the ligation.

renthesis, asterisk, closed curly bracket, and less-than sign) were used for practice trials.

Procedure

As depicted in Figure 2, each trial consisted of (a) a central fixation cross presented for 250 ms, (b) two horizontally aligned letters (separated by 48 pixels) presented for 2,000 ms or until a response was given (whichever was shorter), and (c) an intertrial interval blank screen for 500 ms.

Half of the trials consisted of same pairs (two identical stimuli; $n = 990$) and half of different pairs ($n = 990$). Each letter was paired with itself 22 times, and with another letter 44 times; each specific different pair was presented once to each participant. Participants were asked to respond if the two shapes were physically identical with a yes–no key press (“a” or “l” on the keyboard). The 1,980 trials were administered in two sessions with two enforced breaks per session, and the second session was completed within 2 weeks of the first. Each session began with 31 same–different practice trials with symbol stimuli.

Two stimulus lists were created to allow different pairs to be presented in both left–right orders across participants. On each session, there were 50% same and different pairs with trial order randomized for each participant (with no more than three consecutive same or different pairs). Yes–no mapping to response keys (and response hands) was randomized for each session.

Data Analysis

Preprocessing: Discrimination times. Following Courrieu et al. (2004), “discrimination times” were calculated by removing from each raw RT the time needed for simple detection and response programming and execution (t_0); t_0 was calculated from a set of detection trials at the beginning of each session on which participants pressed one of the response keys as soon as a visual

stimulus (two tilde symbols) was detected, with 25 trials for each key. Trial structure was identical to experimental trials except that fixation duration varied randomly from 600 to 1,000 ms. For each participant, t_0 was calculated as the average of the 20 fastest RTs for each key and then subtracted from the RT for each experimental trial in which the participant pressed that key.

Analysis 1: Hierarchical clustering analysis (HCA). To visualize the similarity relationships between the 45 shapes, HCA was carried out for both discrimination times (t_0 -adjusted RTs) and accuracies of the different pairs for the NG and EG groups. The data for the HCA consisted of four 45×45 confusion matrices. For accuracy data, each matrix cell value corresponds to the average of the mean cross-participant accuracies for the two horizontal presentation orderings of a pair. Discrimination time matrices were calculated as follows: For each cell, the discrimination times on correct trials for each participant were normalized so that the average value for each participant was equal to 1, and then the cross-participant medians of the normalized discrimination times were entered into the cell. The matrix was symmetrized by taking the mean value of the two presentation orders of each letter pair. Finally, the reciprocal of each normalized discrimination time was used to transform the matrix into a *dissimilarity* matrix, providing a measure of “perceptual distance” between letter pairs (values closer to zero represent letters that are more difficult to distinguish).

The confusion matrices were transformed by a multidimensional scaling procedure (“stats” library in R, version 3.0.1; R Core Team, 2015) into a set of coordinates in a high-dimensional space, and the Euclidean distances between these coordinates were used to develop four dendrograms (via the Ward method of minimum variance; Murtagh & Legendre, 2014). The dendrograms provide a compact overview of the similarity relations among letters and a starting point for understanding similarities of and differences between the two participant groups. However, inferences based on quantitative analyses relied on the results of the LMEM, described in the next section.

Analysis 2: Model-based analyses. To address key questions of interest, we examined four models using LMEM. Model 1 had the normalized discrimination times (log-transformed to address positive skew) as the dependent variable, and included fixed effects for 15 predictors of visual similarity and the interactions between each of these factors and group (NG vs. EG). The normalized discrimination times were used for the LMEM rather than raw RTs so that the dependent measure was the same as what was used in the HCA. Model 2 was identical to Model 1 but additionally included the six predictors for knowledge-based factors; Models 3 and 4 were analogous to 1 and 2 but had accuracy as the dependent variable, and were performed with generalized linear mixed models for binomial data (“lme4” library, version 1.1–7 [Bates, Maechler, Bolker, & Walker, 2015] in R [R Core Team, 2015]). These analyses were conducted only on the different pairs.

The models had maximal random effects structures consistent with the experimental design (Barr et al., 2013) and model convergence. Specifically, by participant, both random intercepts by and random slopes for each of the fixed effects were included. By item, only random intercepts were included. Cor-

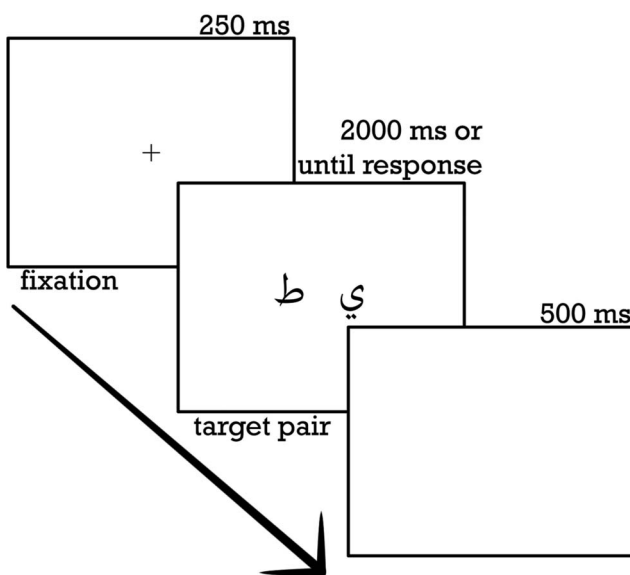


Figure 2. Trial structure for the same–different judgment task.



Figure 3. A representative set of the letter shapes used in the experiment. Each of the features into which they were decomposed for purposes of computing feature overlap is indicated in red. The 15 different visual features were used as predictors in Models 1 to 4.

relations between the random slopes and intercepts by participant led to model nonconvergence and therefore were omitted.

Knowledge-based predictors.

Abstract letter identity. Each letter pair was assigned a value to represent whether or not the letters were allographs of the same letter identity (either 1 or -1, sum-coding). **Motor stroke similarity:** An inventory of the strokes used to write the letters was developed, based on review of videos of the EG writing performance (following methods in Rapp & Caramazza, 1997). Each letter was then described by its set of stroke bigrams (pairs of consecutive strokes), and the proportion of shared stroke bigrams was calculated for each letter pair. **Letter-name similarity:** The proportion of shared distinctive features for the names of each letter pair (e.g., place, manner, voicing, and sonorant or obstruent status for consonants; Bailey & Hahn, 2005). **Alphabetic distance:** To account for spatial effects of alphabetic position (Gevers, Reynvoet, & Fias, 2003; Jonas, Taylor, Hutton, Weiss, & Ward, 2011), analogous to those reported for the mental number line (Dehaene, Bossini, & Giraux, 1993), the difference in the position (1–28) of each letter in the modern Arabic *hija'i* order was calculated for each letter pair (with allographs of the same letter assigned the same position). **Roman alphabet similarity:** The average subjective visual similarity of each of the letters in a pair to Roman letters or to digits. A separate study was conducted with a new set of 24 naïve participants who were asked to respond to individual Arabic letters with the first digit or Roman letter that came to mind as being similar, and to rate the strength of this similarity from 1 to 5. The stimuli and viewing conditions were the same as in the main experiment. Roman alphabet similarity scores were computed for each pair of Arabic letters by calculating the proportion of all the participants ($n = 24$) who judged that both letters of a given pair most resembled the same letter of the Roman alphabet; this proportion was then multiplied by the average rating (for the two letters) of the rated strength of the similarity.

Visual feature predictors. Fifteen visual features were included, based both on the literature and examination of the hierarchical clustering analyses. We used the 10 features reported by Fiset et al. (2008): four line types (horizontal, vertical, slanted right and slanted left), four curve types (open to the top, bottom, left and right), intersections, and terminations—henceforth, the “10 feature set.” Five additional features included two features salient in written Arabic—dots (diacritics) and closed curves—and also three other features proposed previously in the literature on visual recognition of Roman letters: closed space, cyclicity, and symmetry³—the “15 feature set.” The full decomposition of the 45 Arabic letter stimuli is depicted in Figure 3.

The visual feature similarity of each letter pair presented in the different-pair trials was determined for each of the visual feature dimensions by calculating the number of shared features, divided by the total number of features for the pair. To this end, each letter was decomposed into a list of features (see

³ Fiset et al. presumably did not include these features, as these cannot be evaluated with the Bubbles method, precluding comparison with Fiset et al. on these features.

Table 1 for an example). All features (except cyclicity, symmetry, and closed space) had to be visible within a circular “window” with a 4-pixel diameter (proportional to the window within which terminations and intersections were defined in Fiset et al., 2008). Lines were considered to be slanted only if they were $\geq 20^\circ$ off the horizontal or vertical. Curves were categorized as open to the left, right, top, or bottom, except for those that closed within the 4-pixel window. Each termination was assigned to one of the following eight locations: top, bottom, left, right, or any possible conjunction of two of these locations (e.g., top left or bottom right). Intersections were categorized as L, T, or X. A feature was considered to be shared between two letters only if it matched in orientation or position (e.g., a pair of letters shared an intersection only if both letters had an L intersection, and not if one had an L and the other had a T or X). Note that although terminations and intersections were subcategorized in this way for purposes of computing overlap, these subcategories were not assigned individual predictors in the analyses (i.e., there were not eight termination predictors, one for each location).

Model-based analyses. The models were used to address the following issues.

The role of knowledge-based factors. Models 1 and 2 (normalized discrimination time) and Models 3 and 4 (accuracy) were compared to determine (a) whether or not the knowledge-based predictors contribute significantly to model fit (using chi-square tests for nested models), (b) the magnitude of the contribution of the specific knowledge-based factors to the performance of the EG and NG (using R^2 measures), and (c) if (and how) the inclusion of these factors affects the contributions of the visual features.

Effects of expertise on visual feature processing. The beta weights of the 15 visual feature predictors for the NG and EG groups, and corresponding interaction terms, were examined to assess differences between the groups in terms of the relative importance of the features for predicting RT and accuracy.

Effects of expertise, alphabet, and case on visual feature rankings. Using a cross-validation approach, comparisons with the relative rankings of the visual features for the Roman alphabet (as reported in Fiset et al., 2008, 2009) were carried out to determine the extent to which the visual features rankings were comparable: within versus across levels of expertise, across the two alphabets, and across UC and LC for the Roman alphabet. *Across levels of expertise*, participants from a given group (NG or EG) were randomly assigned to two split halves, new feature rankings were obtained by LMEM for each (half) group, and the Spearman correlation between the rankings of the two (half) groups was recorded. This process was repeated 20 times for each group and each measure, and then each of the four sets of 20 correlations were transformed to Fisher's z scores and averaged to provide the overall ranking correlation (mean and confidence interval [CI]) for within-Arabic/within-group. Then, following a similar procedure, the rankings obtained from the split-halves were correlated *across* groups (e.g., NG RT with EG RT, NG accuracy with EG accuracy) and the four sets of 20 correlation values were Fisher's z -transformed and averaged to estimate the within-Arabic/across-groups ranking correlation (mean and CI). Both the within- and across-Arabic-group analyses were carried out for both the 10 and the

Table 1

Example of the Calculation of Shared Features for a Pair of Arabic Letters

	ل	ف
Features	1 vertical line 1 termination-top 1 termination-bottom Symmetric	1 vertical line 1 closed curve 1 termination-bottom 1 T-intersection Closed space
Total	4	5
Shared	1 vertical line, 1 termination-bottom	
Visual feature predictors	Verticals = $2/9 = .222$ Terminations = $2/9 = .222$	

15 feature sets. The resulting within- versus across-group comparison was carried out by 95% CIs on the differences between means.

Across alphabets, the three rankings of the 10 feature set for the Roman alphabet reported by Fiset et al. (LC and UC: Fiset et al., 2008; UC: Fiset et al., 2009) were correlated with the rankings obtained for Arabic from each of the split-half models. These Spearman correlations were Fisher's z -transformed and averaged, with differences assessed again by confidence intervals.

Finally, *across case within the Roman alphabet*, the correlation between the two rankings for UC Roman letters reported by Fiset et al. (2008, 2009), as well as the average pairwise correlation between those two rankings and the LC ranking (Fiset et al., 2008), are reported as a further point of comparison.

Analysis 3: Same pairs and visual complexity. The data from the same pairs was used in the following way to determine whether the visual complexity of letters affects perception and, if so, whether the effect is different for naïve and expert viewers. The visual complexity of each letter was represented by its total number of visual features (calculated as described in the section Visual feature predictors, see Table 1)⁴; over the set of 45 letters, complexity values ranged from four to 15 features ($M = 7.33$). The correlation between visual complexity and both average discrimination time and accuracy was calculated for each participant. Fisher's z transformation was applied to each correlation before averaging the NG and EG separately, and then back-transformed to the Pearson's r ; 95% confidence intervals are provided for each average correlation and for the difference between the groups.

Results

Panoramic Overview of the Similarity Space Provided by HCA

Overall RTs and accuracies. The mean raw RTs and discrimination times ($t0$ -adjusted RTs) on correct trials and accuracy

⁴ The measure of “perimetric complexity” (Pelli et al., 2006), perimeter squared over ink area, was also calculated and used as a measure of complexity. The same pattern of results was obtained with both this measure and the total number of visual features measure; however, we report results only for the latter, as it was found to correlate significantly more highly with both discrimination time and accuracy.

cies by participant group and by pair type are reported in Table 2. With regard to *t0* detection times, the average adjustment made to the raw RTs was 264 ms for the EG and 252 ms for the NG, and the difference of 12 ms was not statistically different from zero, 95% CI [-8, 31]. In terms of discrimination times, the EG was slower than the NG by 35 ms, 95% CI [-7, 78], and mean discrimination time for different pairs was slower than for same pairs by 7 ms, 95% CI [-4, 17]. With regard to accuracy, the EG was more accurate than the NG, on average, by 1.9%, 95% CI [-0.3, 4.2], and overall accuracy was higher on different pairs than same pairs by 1.6%, 95% CI [0.8, 2.4]. However, as indicated by the fact that the CIs of the differences include 0, none of these differences were significant except for the contrast between different and same pairs.

Dendrograms. For comparison purposes, the RT dendrogram for the NG (Figure 4) was color coded, and then the color applied to each letter shape was held constant across the other three dendrograms to facilitate visual comparisons across the figures.

From visual inspection of the dendrograms, one might hypothesize a set of features that accounts for the structure of the trees, an approach taken in various previous studies (e.g., Boles & Clifford, 1989; Courrieu et al., 2004; Gilmore, Hersh, Caramazza, & Griffin, 1979; Loomis, 1982). For example, in the NG RT dendrogram (Figure 4), there are several large clusters of shapes that share salient visual features: The large red cluster contains letters with curves open to the left; the orange cluster consists of short shapes with closed or nearly closed space; the green cluster contains large open curves oriented in the same direction; and so forth. Finer distinctions exist within these clusters, with many of the closest pairs distinguished only by diacritics.

With regard to effects of expertise, there are a few striking differences between the EG and NG RT dendrograms (Figure 4). For example, within the large red NG cluster, for the EG there is a green, a blue, and a red letter (ب, ك, and ح). The considerable visual dissimilarity between these shapes suggests that other factors may have made these letters difficult to discriminate, accounting for their proximity in the dendrogram. In fact, the shape ك is likely grouped with ح because both are allographs of “kaf.” The shape ب shares a similar motor pattern with ك (a vertical stroke down, followed by a horizontal stroke to the left, ending in a curving stroke up to the left,

and then each followed by a pen lift to place a diacritic mark), suggesting that motor similarities may affect same–different judgments. Along similar lines, the EG accuracy dendrogram (Figure 4) also reveals several divergent clusters in which the differences are predictable given the shared name/identity between the shapes.

Likewise, a comparison of the NG’s RT and accuracy (Figure 4) dendrograms reveals similar groupings, with some exceptions. Especially striking is the tri-colored cluster on the right: ا, ب, and و. NG participants reported that these three letters looked similar to the Roman letters o, u, and w, an indication of possible biases derived from knowledge of familiar letter shapes, an issue we pursue in the next analysis.

Inspection of the dendrograms leads to the following conclusions: (a) The highly structured visually based groupings clearly indicate that the same–different judgment task generates both RT and accuracy data that are appropriate for investigating visual feature processing; nonetheless, although visual inspection is a useful tool, a more systematic and quantitative approach is needed; (b) Some of the discrepancies between the NG and EG dendrograms suggest expertise-based differences. However, it will be important to distinguish between-groups differences that arise at the level of visual feature processing from those that do not. The dendrogram results suggest that the latter may include biases arising from various types of knowledge (e.g., such as knowledge of letter names/identities). The dendrograms underscore the point that to evaluate the effects of expertise and alphabet on visual feature processing, it will be necessary to first identify relatively unbiased visual feature sets for the two participant groups.

The Contributions of Knowledge-Based Factors

Chi-square comparisons of Models 1 versus 2, and 3 versus 4, allow for evaluation of the role of knowledge-based factors (on discrimination times and accuracy, respectively) by comparing the fits of models that do not include knowledge-based factors versus those that do. The results of these comparisons reveal a significantly better fit both for Model 2 for discrimination times, $\chi^2(18) = 4505, p < .0001$, and for Model 4 for accuracy, $\chi^2(18) =$

Table 2

*Mean Raw Reaction Time, *t0*-Adjusted Reaction Time (Discrimination Time), and Accuracy Per Group (EG, NG) and Stimulus Type (Same, Different)*

Variable	Same	95% CI	Different	95% CI	Mean	95% CI
Raw reaction time (ms)						
EG	598	[569, 628]	625	[597, 652]	611	[592, 631]
NG	564	[527, 600]	565	[527, 603]	564	[539, 590]
Mean	581	[558, 604]	595	[571, 619]	588	[565, 610]
Discrimination time (ms)						
EG	339	[307, 372]	356	[326, 385]	347	[318, 377]
NG	313	[282, 345]	311	[277, 344]	312	[280, 344]
Mean	325	[305, 348]	333	[311, 356]	330	[309, 351]
Accuracy (%)						
EG	94.9	[93.1, 96.7]	95.8	[94.1, 97.4]	95.3	[93.7, 97.0]
NG	92.3	[90.1, 94.4]	94.6	[93.4, 95.8]	93.4	[91.8, 95.0]
Mean	93.6	[92.2, 95.0]	95.2	[94.2, 96.2]	94.4	[93.3, 95.5]

Note. CI = confidence interval; EG = expert group; NG = naïve group.

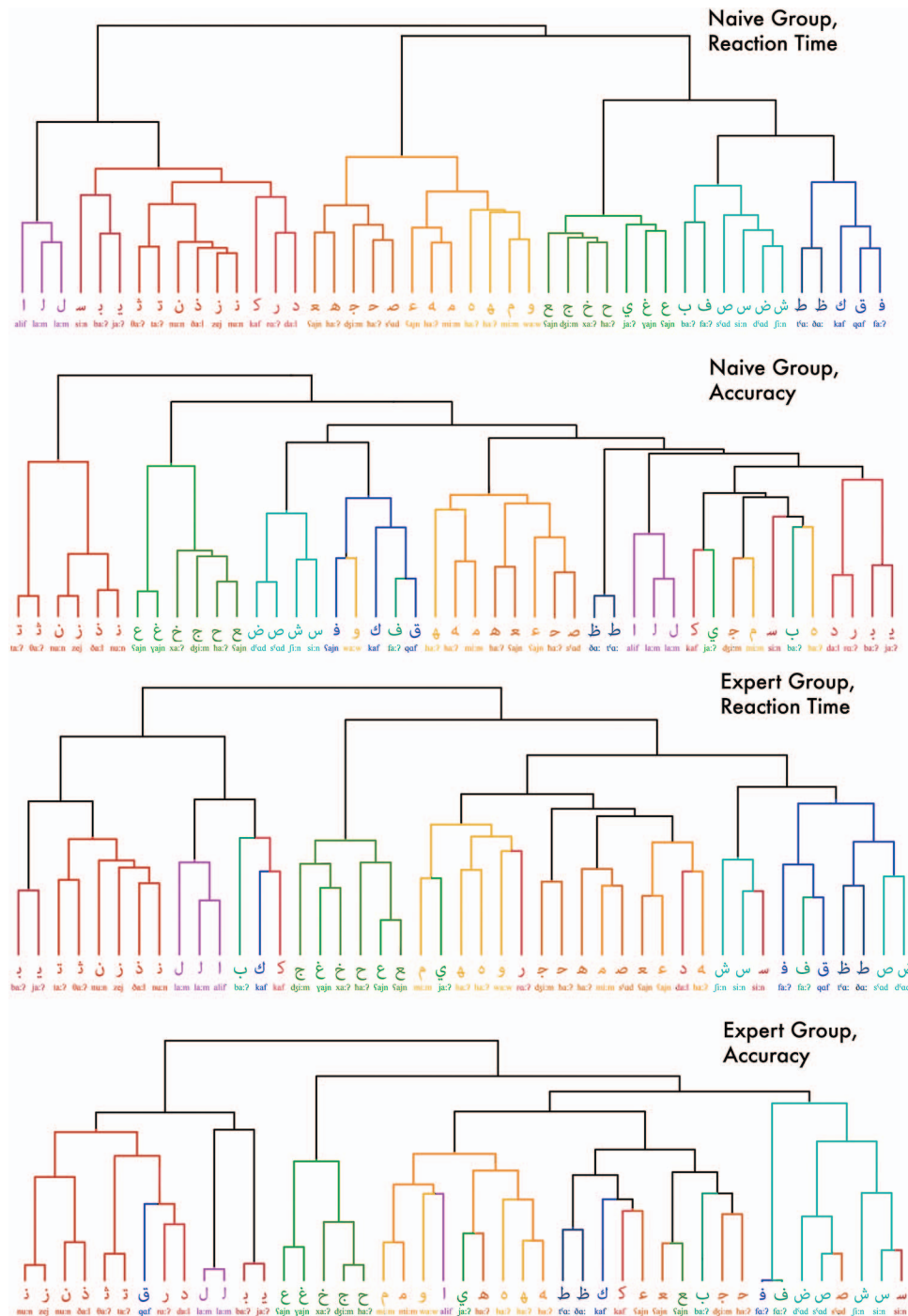


Figure 4. Hierarchical clustering of 45 Arabic letter shapes based on reaction time (RT) and accuracy for the naïve group (NG) and expert group (EG). Color gradient is linked to the NG RT clustering and held constant across all other dendrograms. The height at which a cluster branches reflects the relative similarity of all of the letters in the cluster, such that items in lower branches are more similar; however, the left-to-right ordering within a cluster is arbitrary (e.g., for the purple cluster in the NG RT dendrogram containing “alif” and two allographs of “lam,” one cannot know which of the allographs is more similar to “alif”).

Table 3

Standardized Beta Weights for the Predictors Corresponding to Knowledge-Based Factors for RT (Model 2, Top) and Accuracy (Model 4, Bottom)

	RT					
	EG		NG		Interaction	
	Std. beta	<i>t</i> value	Std. beta	<i>t</i> value	Std. beta	<i>t</i> value
Knowledge-based factor						
Shared identity	.091	3.14	.049	1.77	-.041	-1.20
Trial order	-.116	-6.26	-.124	-8.51	-.005	-.41
Roman alphabet similarity	.021	4.77	.045	9.95	.025	4.64
Alphabetic order	-.003	-.61	.001	.15	.004	.65
Phonological similarity	.007	1.23	.001	.19	-.006	-.91
Motoric stroke similarity	.017	3.86	.004	.09	-.013	-2.53
	Accuracy					
	EG		NG		Interaction	
	Std. beta	<i>z</i> value	Std. beta	<i>z</i> value	Std. beta	<i>z</i> value
Shared identity	-1.225	-3.67	.115	.44	1.429	3.77
Trial order	-.276	-5.10	-.127	-2.28	.141	1.81
Roman alphabet similarity	-.173	-4.53	-3.220	-7.26	-.181	-3.41
Alphabetic order	.027	.58	.036	.83	-.009	-.16
Phonological similarity	-.026	-.61	-.067	-1.55	-.051	-.97
Motoric stroke similarity	-.099	-2.26	-.060	-1.54	.072	1.48

Note. Bold indicates values considered statistically significant at $p < .05$. RT = reaction time; EG = expert group; NG = naïve group; Std. = standardized.

347, $p < .0001$. The superior fits of Models 2 and 4 provide clear evidence of the contribution of knowledge-based factors to same-different judgments. Another measure of the influence of knowledge-based factors is a comparison of the variance accounted for by Model 2 versus Model 1, and Model 4 versus Model 3. The R^2 for generalized linear mixed-effect models, proposed by Nakagawa and Schielzeth (2013), provides a measure of variance explained by the fixed- and random-effects combined (conditional R^2 , or R^2_c). For discrimination times, Model 1 has $R^2_c = 0.077$, which more than doubles in Model 2, $R^2_c = 0.162$. For accuracy, Model 3 has $R^2_c = 0.278$, which increases in Model 4 $R^2_c = 0.312$. This indicates that knowledge-based factors uniquely explain 8.5% of the variance in discrimination times and 3.4% in accuracy (52% and 11% of the total variance explained by Models 2 and 4, respectively).⁵

Consideration of the beta weights and the t and z values⁶ of the specific knowledge-based factors (see Table 3) indicates that the two groups are similar (with no significant interaction between groups), in that (a) trial order significantly (at a threshold of t or $z > 2$, equivalent to an alpha of 0.05) predicts faster and less accurate responses with increasing trial order indicating increasing speed-accuracy trade-offs during the course of the experiment, and (b) there are no significant effects or interactions for either alphabetic order or phonological similarity.

However, factors that interact with group (NG and EG) reveal the different knowledge types that the two groups bring to the task: (a) Although Roman alphabet similarity predicts slower and less accurate responses for both the EG and the NG, the effect is significantly larger for the NG compared with the EG (differences: $\beta = 0.025$, $t = 4.64$, for RT, and $\beta = -0.181$, $z = -3.41$, for accuracy); (b) Shared identity (see Figure 5) and motoric stroke

similarity are better predictors for the EG, such that greater similarity for letter pairs on these factors predicts slower response times (differences: $\beta = -0.041$, $t = -1.20$, for shared identity, and $\beta = -0.013$, $t = -2.53$, for motoric stroke similarity) and lower accuracy (differences: $\beta = 1.429$, $z = 3.77$, for shared identity, and $\beta = 0.072$, $z = 1.48$, for motoric stroke similarity). These differences indicate that the EG is less influenced than the NG by knowledge of Roman letter shapes, and, furthermore, that the EG is affected by knowledge of how letters are written and whether or not letters are allographs—two knowledge types that depend on expertise with the Arabic alphabet.

In addition to establishing the contribution of knowledge-based factors to predicting overall discrimination times and accuracies, it is important to determine whether the inclusion of these factors in the models has an impact on the rankings of the visual feature sets. Differences in visual feature beta magnitudes and rankings for

⁵ When considering magnitude of these R^2 measures, it should be kept in mind that the LMEMs have the ambitious goal of modeling individual trials. One consequence of this is that noise inherent to the data is not “averaged out,” as no averaging by participants or items is carried out. It is likely that much of the remaining unexplained variance comes from intraindividual variability: For example, basic t0 detection times varied widely within participants (standard deviations ranging from 75 ms to 216 ms). Further support for this explanation is the fact that for simplified models in which participants and items are averaged together, R^2 measures for this data set reached as high as 0.7.

⁶ All models determined interactions with “group” as a sum-coded variable, 1 indicating NG and -1 indicating EG. To obtain within-group significance of t and z values, the models were rerun separately for the NG and EG, dropping the “group” variable. Beta-weights were made comparable across predictors by standardizing each predictor to have a variance of 1.

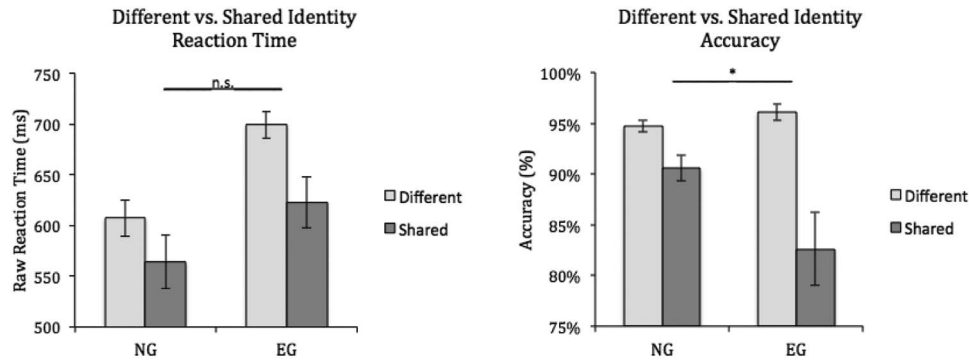


Figure 5. Effect of same-identity on different-pair judgments for the naïve group and expert group, raw reaction time on correct trials (left) and accuracy (right). Error bars reflect standard error of the mean by participants. * Significant interaction between groups, $p < .05$.

Model 1 versus 2, and Model 3 versus 4, for the two groups indicate that “erroneous”⁷ conclusions regarding feature values would have been drawn if the knowledge-based factors had not been included (see Figure 6). The four comparisons reveal that the average ranking change is 1.4 (out of 15 ranks), with changes ranging from 0 to 7 ranks. In addition to changes in the rankings of the visual features, there are also changes in the specific features that constitute statistically significant predictors of discrimination times and accuracy. Specifically, slants-right becomes a significant predictor for both EG and NG discrimination time, whereas seven features are no longer significant predictors in one or more of the models that include knowledge-based factors.

Effects of Expertise on Visual Feature Processing

An examination of the visual feature rankings and beta weights in Models 2 and 4 for the EG and NG reveals similarities and differences across the two groups. First, as can be seen in Figure 7, there are a number of important similarities between the NG and the EG, namely, (a) the most important feature for both groups (ranking first) for both discrimination time and accuracy is the one that is largely unique to Arabic—the diacritic marks; (b) terminations and slants-right are all also significant predictors for both groups for both measures, and closed space, curves-right, and verticals are significant for both groups for at least one of the dependent measures; and (c) curves-top, closed curves, slants-left, and symmetry are not significant for either group.

Importantly, notable indicators of the effect of expertise on visual feature processing are provided by the interactions of specific visual feature predictors with the group predictor: (a) There is a trend for the EG to be slower relative to the NG to respond to pairs with matching intersections and vertical lines (differences: $\beta = 0.008$, $t = -1.73$, and $\beta = -0.008$, $t = -1.78$, respectively); (b) The NG is significantly slower relative to the EG to respond to pairs sharing horizontals, closed curves, and cyclicity (differences: $\beta = -0.010$, $t = 2.25$; $\beta = 0.011$, $t = 2.06$; and $\beta = 0.009$, $t = 2.07$, respectively), and shows a trend to be slower than the EG with pairs sharing diacritic marks (difference: $\beta = 0.009$, $t = 1.66$); (c) The NG is significantly less accurate in responding to pairs sharing curves-right (difference: $\beta = -0.093$, $z = -1.97$) and shows a trend to lower accuracy in responding to pairs of

symmetric letters (difference: $\beta = -0.103$, $z = -1.66$). It should be noted that significant interactions between EG and NG exist in Models 1 and 3 as well, indicating that differences in how the two groups process the visual features that were apparent in Models 2 and 4 are not simply an artifact of the inclusion of the knowledge-based factors.

The same pairs provide a valuable opportunity to evaluate the effect of visual complexity—computed as number of features—on discrimination time and accuracy. The average correlation between discrimination time on correct same pair trials and the total number of visual features in a letter is $r = 0.655$ for the NG, 95% CI [0.543, 0.745], and for the EG, $r = -0.295$, 95% CI [-0.492, -0.070]; the difference in Fisher’s z scores of 1.09 is significant, 95% CI [0.80, 1.37]. Accuracy patterns in precisely the same way: the correlation between the number of visual features and average accuracy was -0.143 , 95% CI [-0.331, 0.078], for the NG, and 0.350 , 95% CI [0.165, 0.512], for the EG; the difference in z scores of 0.50 is significant, 95% CI [0.22, 0.78]. In summary, more complex letters are correlated with significantly faster and more accurate “same” responses for the EG relative to the NG.

Effects of Expertise, Alphabet, and Case on Visual Feature Rankings

These analyses consider whether visual feature processing is affected by expertise, alphabet, and/or case, comparing the feature rankings for EG and NG with one another as well as comparing them with those reported by Fiset et al. for the Roman alphabet (LC and UC: Fiset et al., 2008; UC only: Fiset et al., 2009; Tables 4 and 5).

Feature rankings. Fiset et al. (2008, 2009) argued that terminations and horizontals are the most important features for letter detection. This claim was based on their finding that these two features were most highly ranked for both UC (for the 2008 and

⁷ To be clear, conclusions would be erroneous with regard to visual feature processing, but not necessarily with regard to other conclusions. In other words, the importance of identifying and “excluding” the contribution of knowledge-based factors is necessarily dependent on the empirical question one is addressing.

EG RT	Model 1		Model 2		EG Accuracy	Model 3		Model 4	
	Std. beta	Rank	Std. beta	Rank		Std. beta	Rank	Std. beta	Rank
Diacritics	0.0264	2 *	0.0210	1 *	Diacritics	-0.2276	1 *	-0.1740	1 *
Closed Space	0.0339	1 *	0.0203	2 *	Closed Space	-0.1768	3 *	-0.1253	2 *
Terminations	0.0251	3 *	0.0136	3 *	Terminations	-0.2276	2 *	-0.1008	3 *
Intersections	0.0195	4 *	0.0124	4 *	Horizontals	-0.1323	4 *	-0.0983	4 *
Curves-Right	0.0168	5 *	0.0118	5 *	Slants Right	-0.0812	7 *	-0.0919	5 *
Verticals	0.0137	6 *	0.0098	6 *	Curves-Left	-0.1032	5 *	-0.0598	6
Slants Right	0.0119	7 †	0.0079	7 *	Curves-Right	-0.0329	14	-0.0543	7
Curves-Bottom	0.0063	8	0.0055	8	Verticals	-0.0692	9 *	-0.0509	8
Slants Left	0.0037	11	0.0033	9	Intersections	-0.0700	8	-0.0426	9
Curves-Left	0.0061	9	0.0030	10	Cyclicity	-0.0842	6 *	-0.0360	10
Horizontals	0.0012	13	0.0028	11	Curves-Bottom	-0.0623	10	-0.0242	11
Cyclicity	0.0052	10	-0.0002	12	Closed Curves	-0.0374	13	-0.0147	12
Symmetry	-0.0024	14	-0.0016	13	Curves-Top	-0.0392	12	-0.0092	13
Curves-Top	0.0035	12	-0.0034	14	Slants Left	-0.0455	11	0.0001	14
Closed Curves	-0.0078	15	-0.0076	15	Symmetry	0.0748	15	0.0669	15

NG RT	Model 1		Model 2		NG Accuracy	Model 3		Model 4	
	Std. beta	Rank	Std. beta	Rank		Std. beta	Rank	Std. beta	Rank
Diacritics	0.0361	1 *	0.0301	1 *	Diacritics	-0.2586	1 *	-0.2054	1 *
Curves-Right	0.0207	4 *	0.0163	2 *	Curves-Right	-0.1450	3 *	-0.1409	2 *
Closed Space	0.0222	3 *	0.0152	3 *	Slants Right	-0.1136	5 *	-0.1345	3 *
Horizontals	0.0116	8 *	0.0132	4 *	Terminations	-0.2500	2 *	-0.1156	4 *
Terminations	0.0247	2 *	0.0132	5 *	Verticals	-0.0964	7 *	-0.0793	5 *
Curves-Left	0.0192	5 *	0.0121	6 *	Closed Curves	-0.0900	8 *	-0.0734	6 †
Cyclicity	0.0139	6 *	0.0088	7 *	Closed Space	-0.0878	9	-0.0692	7
Curves-Bottom	0.0122	7 *	0.0086	8 *	Intersections	-0.0732	10	-0.0607	8
Slants Right	0.0075	10	0.0078	9 *	Curves-Left	-0.1360	4 *	-0.0576	9
Intersections	0.0063	11	0.0039	10	Curves-Top	-0.1127	6 *	-0.0486	10
Closed Curves	0.0053	12	0.0038	11	Horizontals	-0.0523	12	-0.0327	11
Slants Left	0.0019	14	0.0027	12	Curves-Bottom	-0.0684	11 †	-0.0314	12
Verticals	0.0031	13	0.0019	13	Symmetry	-0.0180	14	-0.0270	13
Curves-Top	0.0100	9 *	0.0017	14	Cyclicity	-0.0487	13	0.0019	14
Symmetry	-0.0006	15	-0.0004	15	Slants Left	-0.0011	15	0.0072	15

Figure 6. The 15 visual features as ranked by the absolute value of the standardized beta weights in models without (Model 1 and 3) and with (Model 2 and 4) predictors for knowledge-based factors. EG = expert group; NG = naïve group; RT = reaction time; Std. = standardized. Significant variable at * $p < .05$. † $p < .10$.

2009 participant groups) and LC Roman letters. As indicated in Table 4 (reporting rankings for only the 10 features evaluated by Fiset et al.), the results we observed with Arabic letters (based on LMEM, Models 2 and 4) replicate the importance of terminations, which is a significant factor that ranks first for the EG and third for the NG for both discrimination time and accuracy. However, horizontals present a more mixed picture, as they are a significant factor ranking second for NG discrimination time and EG accuracy, but are a nonsignificant factor for NG accuracy (ranking eighth) and EG discrimination time (ranking ninth). In contrast to the high ranking of horizontals, Fiset and colleagues found verticals to be particularly unimportant to human observers (ranking eight out of 10 for both studies of UC, and sixth for LC Roman alphabet); for Arabic, we found that they are comparable with horizontals in that they were a significant factor for the EG for discrimination time (ranking fourth) and for the NG for accuracy (also ranking fourth). Strikingly, although curves-right were the lowest ranked factor for Fiset et al. for both studies of UC letters and ninth of 10 for LC, they are the top ranked factor for the NG

in both discrimination and accuracy and are highly ranked for the EG (third and fifth).

Statistical comparison of feature rankings within and across alphabets and case. Table 5 presents the average correlations between the rankings of the features within and across expertise groups, alphabets, and case, based on the split-half cross-validation procedures described in the Data Analysis section. The average within-Arabic/within-groups (i.e., cross-validated rankings of the EG and NG for both discrimination time and accuracy) is 0.517, 95% CI [0.454, 0.575], based on the 10-feature set, and 0.629, 95% CI [0.589, 0.665], based on the full 15-feature set. These correlations are larger than the within-Arabic/across-groups average ranking correlations of 0.396, 95% CI [0.390, 0.402], for the 10-feature set and 0.545, 95% CI [0.541, 0.548], for the 15-feature set. The difference between Fisher's z scores for the two within- versus across-Arabic comparisons is significant, $z = 0.153$, 95% CI [0.069, 0.236], and 0.128 95% CI [0.065, 0.190], for the 10- and 15-feature sets, respectively, indicating the feature rankings were significantly different for the NG and EG.

RT (Model 2)	Std. Betas		Interaction	
	EG	NG		
Diacritics	0.021 *	0.030 *	†	
Closed Space	0.020 *	0.015 *		
Terminations	0.014 *	0.013 *		
Intersections	0.012 *	0.004	†	
Curves-Right	0.012 *	0.016 *		
Verticals	0.010 *	0.002	†	Beta
Slants Right	0.008 *	0.008 *		0.03
Curves-Bottom	0.006	0.009 *		0.015
Slants Left	0.003	0.003		0
Curves-Left	0.003	0.012 *	†	-0.004
Horizontals	0.003	0.013 *	*	-0.008
Cyclicity	0.000	0.009 *	*	
Symmetry	-0.002	0.000		
Curves-Top	-0.003	0.002		
Closed Curves	-0.008	0.004	*	

Accuracy (Model 4)	Std. Betas		Interaction	
	EG	NG		
Diacritics	-0.174 *	-0.205 *		
Closed Space	-0.125 *	-0.069		
Terminations	-0.101 *	-0.116 *		
Horizontals	-0.098 *	-0.033		
Slants Right	-0.092 *	-0.135 *		Beta
Curves-Left	-0.060	-0.058		0.067
Curves-Right	-0.054	-0.141 *	*	0.034
Verticals	-0.051	-0.079 *		0
Intersections	-0.043	-0.061		-0.103
Cyclicity	-0.036	-0.082		-0.205
Curves-Bottom	-0.024	0.002		
Closed Curves	-0.015	-0.073 †		
Curves-Top	-0.009	-0.049		
Slants Left	0.000	0.007		
Symmetry	0.067	-0.027 †		

Figure 7. Standardized beta weights of the visual feature predictors for reaction time (RT; Model 2) and accuracy (Model 4). Features are sorted from most to least important based on the EG (left columns). Note that negative correlations for accuracy indicate that accuracy decreases with increased visual similarity. EG = expert group; NG = naïve group; Std. = standardized. Significant variable at † $p < .10$.

In addition, we see much lower *across*-alphabet ranking correlations: The average ranking correlation of Arabic with Roman LC is just 0.183, 95% CI [0.128, 0.237], and with Roman UC is 0.193, 95% CI [0.160, 0.225]. Fisher's z scores for these average correlations are both significantly lower than the within-Arabic/across-groups correlations, 0.234, 95% CI [0.178, 0.291], and 0.224, 95% CI [0.189, 0.249]. As a point of comparison, the correlation between the two rankings for the Roman UC presented in [Fiset et al., 2008](#) and 2009, is 0.927, and the average correlation within-Roman, across-case is 0.464. Taken as a whole, these results indicate that not only are the feature ranking differences between the EG and NG reliable, but they are also more similar to each other than they are to the feature rankings of the Roman alphabet.

Discussion

We examined the response times and accuracies of participants who were asked to judge whether pairs of visually presented letters were physically identical or not. The innovative aspects of the research were the choice of alphabet, Arabic, and the participants, expert or inexperienced readers of Arabic. The combination of

stimuli, participants, task, and analytic methods (LMEM) allowed us to specifically evaluate visual feature processing of Arabic letters in naïve (NG) and expert (EG) viewers. Based on this approach, the work produced a number of novel findings: (a) We find clear evidence that knowledge-based factors (such as letter identity) influence perceptual judgments; (b) We show that analytic methods such as LMEM can provide a means for removing the contributions of these knowledge-based factors to more clearly identify the characteristics of the visual processing itself; (c) We demonstrate that when this approach is applied, the evidence indicates that both the nature of the stimulus set (the alphabet) and also the level of expertise with it influence visual processing itself; (d) We are able to identify specific differences and similarities in the visual feature processing of Arabic and Roman letters.

Knowledge-Based Factors: Postperceptual Effects

A key finding of this research is that even when given the same visual shapes and asked to make “low-level” judgments of physical identity, observers behave differently depending on at least some nonvisual knowledge that they bring to task. Specifically, we

Table 4

Rankings of 10 Visual Feature Types for the Identification of Roman UC and LC Letters From [Fiset et al. \(2008, 2009\)](#) and for Arabic Letters for the Naïve and Expert Groups From Model 2 (DT) and Model 4 (Accuracy) in the Current Study

Feature	Roman UC (2008)	Roman UC (2009)	Roman LC (2008)	NG – DT	NG – Accuracy	EG – DT	EG – Accuracy
Terminations	1	1	1	3	3	1	1
Horizontals	2	2	2	2	8	9	2
Intersections	3	6	10	7	5	2	7
Curves-left	4	3	7	4	6	8	4
Slants right	5	4	3	6	2	5	3
Slants left	6	5	8	8	10	7	10
Curves-top	7	7	4	10	7	10	9
Verticals	8	8	6	9	4	4	6
Curves-bottom	9	9	5	5	9	6	8
Curves-right	10	10	9	1	1	3	5

Note. UC = uppercase; LC = lowercase; NG = naïve group; DT = discrimination time; EG = expert group.

Table 5
Estimated Mean Correlations (Spearman's) of Visual Feature Rankings

Feature Rankings	Mean	95% CI
Within-Arabic/within-measures-and-groups (10)	.517	[.454, .575]
Within-Arabic/within-measures-and-groups (15)	.629	[.589, .665]
Within-Arabic/across-measures-and-groups (10)	.396	[.390, .402]
Within-Arabic/across-measures-and-groups (15)	.545	[.541, .548]
Across alphabet (Arabic/Roman LC)	.183	[.128, .237]
Across alphabet (Arabic/Roman UC)	.193	[.160, .225]
Within-Roman/within-uppercase ^a	.927	
Within-Roman/across-case ^a	.464	

Note. 95% confidence interval (CI) based on empirical percentiles from split-half cross-validation. UC = upper case; LC = lower case; Groups = EG, NG (Expert and Naive); measures = Discrimination time, accuracy; 10 = the 10 feature types in Fiset et al. (see Table 4); 15 = the 15 feature types examined in this study (Figures 5, 6).

^a Based on rankings reported by Fiset et al. (2008, 2009).

found that models that included task-irrelevant “knowledge-based” predictors in addition to the visual feature predictors provided significantly better fits to overall RTs and accuracies, and, in fact, the knowledge-based predictors accounted for 52% and 11% of the total variance explained by the models for discrimination time and accuracy, respectively.

The perceptual decisions of experienced readers of Arabic (but not the NG) were affected by their knowledge of the motoric patterns used to produce letters as well as their knowledge of the shared identity/name of allographs (see Figure 5). The latter result is comparable with findings from other paradigms involving Roman alphabet stimuli that have reported that judgments of physical identity are slower for visually dissimilar cross-case letters pairs (e.g., A–a) than for other visually dissimilar pairs (V–a; Lupyán et al., 2010). This knowledge is clearly irrelevant for shape-matching judgments. Although one cannot rule out that this knowledge affects the visual feature space itself, it would seem prudent and conservative to assume that these factors affect judgment times and accuracies at later processing stages (e.g., abstract letter identification or decision processes). Although these effects are interesting in their own right, more important is that they reveal that consideration of the influence of these factors is important for work on visual feature processing. This is because correlations between similarity on knowledge-based dimensions and visual feature dimensions can “contaminate” inferences about visual feature processing if they are not taken into account. In our findings, although the overall feature rankings change only an average of 1.4 ranks when knowledge-based factors are included, specific features change by as many as seven ranks, and there are also changes in the specific features that cross the statistical significance threshold. Therefore, it is likely that inclusion of such factors in an LMEM approach provides a better characterization of visual processing.

Learning Effects: Early Versus Late and Bottom-Up Versus Top-Down

In addition to the effects of knowledge-based factors, this study identified effects of learning (alphabet and expertise) on visual processing itself. We discuss these findings in more detail in

subsequent sections; however, we thought it would be useful to first deal with the difficult issues relating to drawing inferences about whether the effects of visual learning that we report arise early or late or are the result of bottom-up or top-down processing.

The learning effects we report naturally raise a number of important questions regarding the manner in which the visual system adjusts to the visual attributes of different stimulus sets and the locus of changes that arise with expertise. In the introduction, this was discussed in the context of the universality of feature sets and whether these are fixed or flexible. The results of this study clearly indicate that learning affects visual processing and, therefore, a natural question concerns how “early” visual flexibility and these learning effects arise. One possibility is that there is “early” flexibility and readers proficient in multiple alphabets have feature sets and configurations that are unlike those of monolingual readers of either script. Also consistent with early flexibility is that feature (re)weighting occurs dynamically online based on the current context and task demands. These notions of early flexibility contrast with “late flexibility,” according to which the differences in feature processing across alphabets and case arise after basic feature processing has occurred with a fixed set of features, and, at these later stages, flexibility affects the way in which the basic features are combined, attended to, and so forth. One potentially relevant source of information might come from bicultural individuals and a comparison of their feature rankings from both alphabets. Such a comparison could reveal rankings that differ flexibly depending on the alphabet they are viewing, or rankings that are stable but different from those of monoscriptal individuals. Such data would certainly provide further insights into the characteristics of visual learning; however, it is important to underscore that distinguishing between early and later sources of flexibility is really not possible without a detailed theory that allows for an independent means for establishing whether effects arise early or later.

Similar difficulties arise when considering the orthogonal issue of whether the effects that we report occur during feedforward, bottom-up visual processing (early or late) or during top-down processing in which “higher” levels of processing—or even other modalities of information—influence the state of “lower” visual processing levels. Given the many arguments for top-down processing based on behavioral and neuroimaging findings (e.g., in the context of letter perception specifically, see Lupyán et al., 2010; but also see Chen & Proctor, 2012 and Klemfuss, Prinzmetal, & Ivry, 2012, for alternate accounts of findings that have been used to argue for top-down effects on perception), we cannot rule out that some or all of the effects we report are not the result of top-down processes that attentionally (or otherwise) bias visual feature processing. Although we acknowledge that these questions are both interesting and important, there is much controversy regarding criteria for establishing these effects (Awh, Belopolsky, & Theeuwes, 2012; Leonard & Egeth, 2008; Lyons, 2011; Pylyshyn, 1999), and this study was not designed to tackle these issues. Therefore, although we believe that our findings provide clear evidence of learning and flexibility in visual feature processing itself (rather than postperceptual decision processing), we are not able to assign the effects to specific levels or stages of visual processing.

Visual learning: Effects of alphabet. Pelli et al. (2006) proposed that letter identification, for any alphabet, requires 7 ± 2

visual features. Our findings for the Arabic alphabet are consistent with this claim, as we found that of the 15 feature types we evaluated there were seven different features with statistically significant beta values for either RT or accuracy (or both) for the EG and 10 features for the NG.

In terms of the specific features themselves, we compared our findings with Arabic to those reported by Fiset et al. (2008, 2009) from two studies involving Roman letters. Differences in experimental paradigms, such as the smaller size of our stimuli and the larger set of possibilities (45 vs. 26 letter shapes) should be kept in mind when comparing the results. Nonetheless, there were a number of noteworthy similarities and differences. In terms of similarities, we found support for Fiset et al.'s proposal regarding the importance of terminations, which they consistently found to be the single most important feature for letter identification of both UC and LC Roman letters. In our study, terminations were highly ranked for both the NG and EG, for both RT and accuracy, with rankings ranging from one to three. Fiset et al. (2008) suggested that the fact that there are neurons in primary monkey cortex that are highly responsive to terminations (e.g., Hubel & Wiesel, 1968) may indicate that these features are detected very early in the course of visual processing. This would be consistent with their prominent role across alphabets and case.

In terms of differences between our findings and those of Fiset et al. (2008, 2009), perhaps the most salient one concerns horizontals and verticals. Fiset et al. found a dissociation between these, such that although horizontals were the second most important predictor, verticals were only ranked eight (sixth for LC). The difference between horizontals and verticals led Fiset et al. (2008) to propose that the human visual system may not be well equipped to process verticals. In contrast, we did not find a difference between them, as both features were significant predictors for both groups for either RT or accuracy, and were also similarly ranked. Another salient divergence with the Fiset et al. findings concerns curves open to the right, which was the lowest ranked factor for the Fiset et al. participants for UC, and ninth of 10 for LC, but was highly ranked for the NG for both RT and accuracy (first) and highly ranked for the EG in RT (third). Finally, it is important to mention that of the five feature types we considered that were not evaluated by Fiset et al. (diacritics, closed curves, symmetry, cyclicity, and closed space), diacritics and closed space were two of the most highly ranked predictors for the NG and EG for both RT and accuracy (see Table 5).

What can be made of these similarities and differences? As we have indicated both in the introduction and just above, it is a profound challenge for experimental studies of vision to identify the locus of effects within the visual system. Although we generally assume that properties of early visual processing should manifest themselves across stimulus types, it does not follow that effects observed across stimulus types necessarily arise early in the system. Thus, although the convergence across alphabets and empirical approaches regarding the importance of terminations is certainly consistent with the claim that they correspond to features that are detected early in visual processing, one should be cautious in drawing strong conclusions in this regard. At a minimum, however, this convergence highlights the important role of these features at some relatively basic, shared level of visual processing.

The divergence across studies regarding horizontals, vertical, and curves-right points to important differences in the visual

processing of the different alphabets (stimulus sets). The strong convergence across methods (same–different judgments vs. Bubbles image classification) with regard to terminations suggests that the observed differences are not likely because of the differences in methods. Instead, these divergences suggest a visual system that flexibly weights features in a way that is, to some considerable extent, driven by the similarity structure of the stimulus set. This conclusion is also supported by the findings regarding ranking correlations within and across alphabet and case. These correlations indicate the following: (a) Visual feature processing is most similar for the same letters of a given alphabet, for different measures (discrimination time and accuracy for the Arabic alphabet) and even when evaluated with different groups of participants with differences in expertise (EG and NG); (b) This stands in stark contrast with the across-alphabet (Arabic/Roman) feature rankings in which the correlations are much lower; (c) Across case (UC and LC within the Roman alphabet) represents an intermediate situation. The difference between cross-case and within-case correlations for the Roman alphabet indicates a clear effect of case, but, additionally, the fact that the cross-case correlations are still higher than the cross-alphabet correlations is an indication that there is greater similarity in visual feature processing within the Roman alphabet (despite case differences) than between the Roman and Arabic alphabets.

Visual learning: Effects of expertise. A primary goal of this research was to investigate whether and how expertise affects visual feature processing. Before discussing differences between the groups, it is important to first note that there were a great many similarities between the naïve and expert participants. The average ranking correlation across the two groups and measures viewing the Arabic alphabet was 0.545, notably higher than the average across-alphabet correlations of 0.183 and 0.193 for LC and UC Roman letters, respectively. In addition to the overall ranking similarity between the NG and EG, salient specific similarities included the high ranking of diacritic marks, closed space, curves-right, and terminations.

However, there were important differences between the groups that support the conclusion that expertise does in fact affect visual processing—a finding that is consistent with many findings in the literature regarding visual learning (e.g., Pelli et al., 2006). First, with regard to specific features, intersections were significant predictors of discrimination time for the EG, but not the NG, and there were several features that were significant predictors of discrimination time for the NG, but not the EG: horizontals, curves-left, cyclicity, and curves-bottom. Importantly, two of those features (horizontals and cyclicity) also interacted statistically significantly with the group factor, and thus are the most associated with expertise. We hypothesize that horizontals and cyclicity have a “dispreferred” status for the EG because in text reading, unlike the situation with the single-letter same–different task, Arabic letters are most often ligated to one another to form words, altering the diagnosticity of these two features. Figure 8 illustrates that horizontal lines and cyclicity are ubiquitous in Arabic words, as the ligation of letters to one another forms contiguous horizontal lines in the subword, and the number of letters which have “teeth” (small vertical peaks) align to create stretches of the cyclicity feature. Thus, horizontal lines and cyclicity are not especially useful features for identifying individual Arabic letters, because in the context of whole words, they occur as features that span

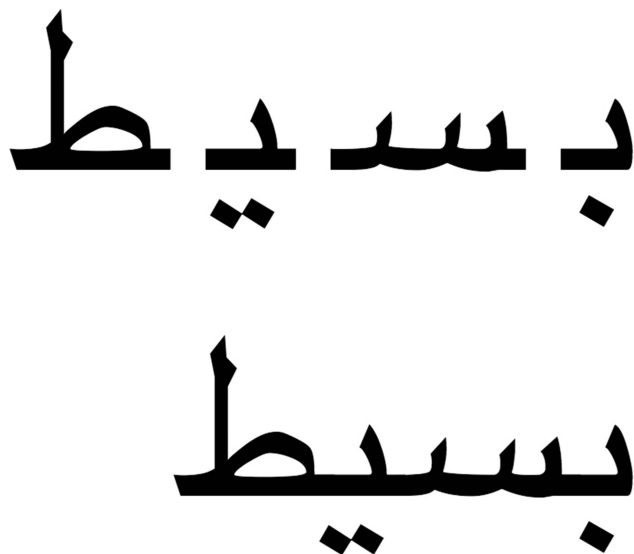


Figure 8. Example of four Arabic letters (top, from right to left, “ba,” “seen,” “ya,” and “taw”) written in isolated forms as appearing in the same–different task, and the same four letters written as a word (bottom line), which creates one continuous horizontal line through the whole word as well as a long cyclic feature.

multiple letters. This highlights the point that expertise in letter identification largely involves the tuning of visual processes and representations for word reading.

Second, analysis of performance on same pairs revealed that, for the NG, the greater the visual complexity of the letter, the slower and less accurate the responses. In striking contrast, for the EG, increasing complexity led to faster RTs and greater accuracy. These results reveal not only that the naïve observer has more difficulty processing more complex letters, but that for the expert observer, more complex letters may be easier to identify, perhaps because they are more distinctive. For expert observers, one possibility is that with expertise, features may become “bundled” together to form fewer features that are each more internally complex. This type of learning would be consistent with findings regarding certain types of experience-dependent perceptual learning (Goldstone, 1998; Kellman & Garrigan, 2009; Sireteanu & Rettenbach, 2000) and visual crowding, in which more complex targets have been shown to be more readily identified among less complex flankers (Bernard & Chung, 2011; Chanceaux, Mathôt, & Grainger, 2014).

Implications for Applications

For teachers of written Arabic as a first or second language, it may be helpful to understand that becoming proficient in identifying Arabic letters involves learning to process or attend to fewer visual features overall. Furthermore, the features that show a significant interaction between the two groups (see Figure 7) may provide some specific instructional guidance. The interactions reveal that there are certain features—intersections and verticals—that tend to be weighted more strongly by expert compared with naïve readers, and it may be useful to explicitly draw learners’ attention to these features. In contrast, there are a number of

features that novices weight more heavily than the experts. These may need to be deemphasized, including horizontals, curves-left, curves-right, and symmetry. The feature “horizontals” is a particularly good example in this regard. Fiset et al. (2008, 2009) emphasized the significance of horizontals in the perception of Roman letters; however, for Arabic letters, most of the horizontal segments occur in the ligatures between letters and do not, therefore, play an important role in distinguishing between letter identities. Second-language learners with a first language that uses the Roman alphabet may bring the high weighting of horizontals to their learning of Arabic, something that will not serve them in good stead. The methodology used in this study could potentially be used in diagnostic testing for identifying the feature weighting for individual students and then tailors teaching accordingly, encouraging the deployment of attention to the critical features. Of course, whether or not explicit instruction based on these findings can actually accelerate the learning process would need to be empirically evaluated.

The finding that motor stroke similarity significantly affected both the time and accuracy of same–different judgments for the EG may be relevant to current debates regarding the role of handwriting in literacy (e.g., Deardorff, 2011; Guan, Liu, Chan, Ye, & Perfetti, 2011; James & Engelhardt, 2012). With the rise of keyboarding relative to handwriting, there is great interest from teachers, curriculum developers, parents, and students to understand the relevance of handwriting for literacy development. The findings from this study are consistent with others in the behavioral (Babcock & Freyd, 1988; James & Gauthier, 2009; Longcamp, Zerbato-Poudou, & Velay, 2005) and neuroimaging (James & Atwood, 2009; James & Gauthier, 2006; Longcamp, Anton, Roth, & Velay, 2003) literatures that have reported that handwriting codes are active when viewing letters or reading words. Clearly, handwriting is not necessary for literacy, as children who are paralyzed learn to read and adults who become paralyzed do not lose their reading abilities. What needs to be addressed empirically is whether handwriting confers any advantage to the learning process, adds any efficiency to the experienced reader, or provides any redundancy that is beneficial in cases of neural injury.

Conclusions

This study reveals a visual system that can be understood as dynamically adjusting its weighting of visual features in ways that, at least in part, reflect the characteristics of the stimulus set (the alphabet) that it is faced with. We also find clear evidence that visual feature processing is susceptible to learning as expert readers come to more efficiently and effectively discriminate the letters of the alphabet they are viewing, relying on fewer features and different feature weightings. Furthermore, the findings indicate that although visual complexity is an obstacle to be overcome for early learners, the expert viewer is able to exploit complexity for improved performance.

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Received April 21, 2015

Revision received December 20, 2015

Accepted December 31, 2015 ■