

2021 Preseason Pink Salmon Forecast

Sara Miller, Rich Brenner, Jim Murphy

October 26, 2020–draft

Contents

1	Objective	1
2	Executive Summary	1
3	Forecast Models (pink_cal_pooled_species vessel calibration coefficient)	2
3.1	Analysis	2
3.2	Model Diagnostics	3
3.3	Results	7
3.4	Conclusion	7
4	References	8

1 Objective

To forecast the Southeast Alaska (SEAK) pink salmon harvest in 2021.

2 Executive Summary

Forecasts were developed using an approach originally described in Wertheimer et al. (2006), and modified in Orsi et al. (2016) and Murphy et al. (2019). We used a similar approach to Murphy et al. (2019) but assumed a log-normal error. This approach is based on a multiple regression model with juvenile pink salmon catch-per-unit-effort (CPUE) and temperature data from the Southeast Alaska Coastal Monitoring Survey (SECM; Murphy et al. 2020). The final model used for the forecast was:

$$E(y) = \alpha + \beta_1 X_1 + \beta_2 X_2 + \epsilon,$$

where y is $\ln(\text{pink salmon harvest in SEAK})$, β_1 is the coefficient for CPUE using the ‘pooled’ vessel calibration coefficient (see *calibration_coefficient_discussion* document), β_2 is the coefficient for the environmental covariate water temperature, and ϵ represents the lognormal error term. The CPUE data are the average log-transformed catches standardized to an effort of a 20 minute trawl set and calibrated to the fishing power of the NOAA Ship *John N. Cobb*. For each year, the standardized catch is either taken from June or July, whichever month had the highest average catches in a given year. Water temperature data is the

average (May through July) temperature in the upper 20 m at the eight SECM stations in Icy Strait. This is similar to what has been identified as the Icy Strait Temperature Index ('ISTI'), but historically this index has included May through August temperature data. Based on performance metrics (Akaike Information Criterion corrected for small sample sizes; AICc values; Burnham and Anderson 2004; mean and median absolute percentage error (MAPE, MEAPE); mean absolute scaled error (MASE) (Hyndman and Kohler 2006)) used to evaluate forecast accuracy of alternative vessel calibration coefficients using the same model, the 'pooled' vessel calibration coefficient was used in the following analysis.

Leave-one-out cross validation (hindcast) and model performance metrics were used to evaluate the forecast accuracy of models. These metrics included MAPE, MEAPE, MASE, and AICc. Statistical analyses were performed with the R Project for Statistical computing version 3.6.3 (R Core Team 2020). Based on the AICc, the MASE metric, and significant coefficients in the models, the preferred model (i.e., the additive model with CPUE and temperature; model m2) predicted that the SEAK pink salmon harvest in 2021 will be in the moderate range with a point estimate of 28.3 million fish (80% prediction interval: 19.1 to 41.9 million fish).

3 Forecast Models (pink_cal_pooled_species vessel calibration coefficient)

3.1 Analysis

Three hierarchical models were investigated. The full model was model m3:

$$E(y) = \alpha + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1 X_2,$$

where X_1 is the average CPUE for catches in either the June or July survey, whichever month had the highest average catches in a given year, and based on the pooled-species vessel calibration coefficient, and X_2 is the average temperature in Icy Strait in May, June, and July at the eight SECM stations sampled in Icy Strait (Icy Strait and Upper Chatham transects; 'ISTI'), and β_3 is the interaction term between CPUE and the temperature index. The CPUE data are log-transformed in the model (Table 5; m3). If temperature is actually altering how CPUE is related to abundance it makes sense to restrict the temperature data to the CPUE months in the forecast model (June and July). The month of May is included as there are important migratory dynamics prior to the time juveniles are actually sampled in Icy Strait. In the past, the 'ISTI' variable was the average temperature in the upper 20 m during May through August at eight stations in Icy Strait. For simplicity, although the definition of the variable has changed, the variable is still called 'ISTI.'

Model m1 only contained the CPUE variable and the model m2 contained CPUE, and a May through July temperature variable ('ISTI'). The regression coefficients CPUE and temperature ('ISTI') were significant in the first two models (m1, m2). The interaction term was not significant in the full model (model m3; Table 5). Therefore, only the first two models will be considered further.

Table 1: Parameter estimates for the three potential models.

model	term	estimate	std.error	statistic	p.value
m1	(Intercept)	2.275	0.207	10.967	0.000
m1	CPUE	0.441	0.071	6.232	0.000
m2	(Intercept)	6.893	0.970	7.109	0.000
m2	CPUE	0.503	0.051	9.860	0.000
m2	ISTI	-0.526	0.109	-4.817	0.000
m3	(Intercept)	3.917	2.464	1.590	0.128
m3	CPUE	1.653	0.880	1.879	0.076
m3	ISTI	-0.201	0.270	-0.746	0.465

model	term	estimate	std.error	statistic	p.value
m3	CPUE:ISTI	-0.125	0.095	-1.310	0.206

The model summary results using the metrics AICc, MAPE, MEAPE, and MASE (Hyndman and Kohler 2006) are shown in Table 6. For all of these metrics, the smallest value is the preferred model. The difference (Δ_i) between a given model and the model with the lowest AICc value and the metric MASE were the primary statistics for choosing appropriate models in this analysis. Models with $\text{AICc}\Delta_i \leq 2$ have substantial support, those in which $4 \leq \text{AICc}\Delta_i \leq 7$ have considerably less support, and models with $\text{AICc}\Delta_i > 10$ have essentially no support (Burnham and Anderson 2004). These two metrics (AICc, MASE) suggest that model m2 is the preferred models. Model m2 (based on CPUE and average temperature in May through July) was used to forecast the 2021 pink salmon harvest (Figure 1).

Table 2: Summary of model outputs and forecast error measures.

model	terms	AdjR2	AICc	MAPE	MEAPE	MASE
m1	CPUE	0.632	30	0.115	0.106	0.395
m2	CPUE+ISTI	0.821	15	0.076	0.063	0.257

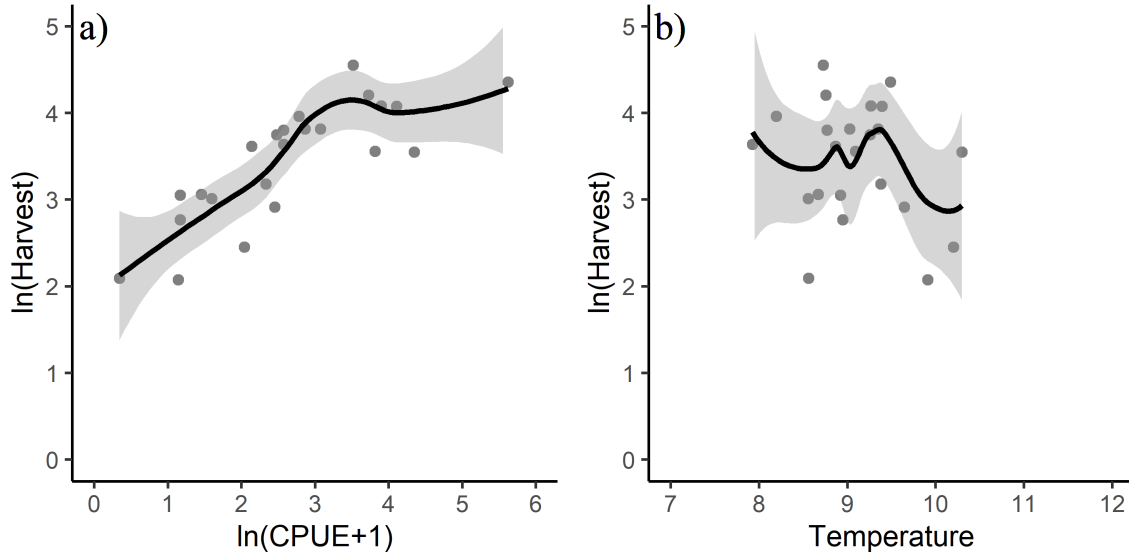


Figure 1: Relationship between a) $\ln(\text{CPUE}+1)$ and $\ln(\text{harvest})$ and b) temperature in May through July (ISTI) and $\ln(\text{harvest})$. The line is a smoothing function applied to the relationship with a 95% confidence level interval.

3.2 Model Diagnostics

Model diagnostics for model m2 included residual plots, the curvature test, and influential observation diagnostics using Cook's distance (Cook 1977), the Bonferroni outlier test, and leverage plots. Model diagnostics were used to identify observations that were potential outliers, had high leverage, or were influential (Zhang 2016). These observations may have significant impact on model fitting and may need to be excluded.

3.2.1 Cook's Distance

Cook's distance is a measure of influence, or the product of both leverage and outlier. Cook's distance,

$$D_i = \frac{e_{PSi}^2}{k+1} * \frac{h_i}{1-h_i},$$

where e_{PSi}^2 is the standardized Pearson residuals, h_i are the hat values (measure of leverage), and k is the number of predictor variables in the model, is a measure of overall influence of the i_{th} data point on all n fitted values (Fox and Weisburg 2019). A large value of Cook's distance indicates that the data point is an influential observation. Cook's distance values greater than $4/(n-k-1)$, where n is the number of observations (i.e., 23), was used as a benchmark for identifying the subset of influential observations (Ren et al. 2016). Therefore, a Cook's distance cut-off of 0.20 was used; observations with a Cook's distance greater than 0.20 were investigated further.

3.2.2 Leverage

An observation that is distant from the average covariate pattern is considered to have high leverage. If an individual observation has a leverage value h_i greater than 2 or 3 times p/n (Ren et al. 2016), it may be a concern (where p is the number of parameters in the model including the intercept (i.e., 3), and n is the number of observations in the model (i.e., 23); $p/n = 3/23 = 0.13$ for this study). Therefore, a leverage cut-off of 0.26 was used; observations with a leverage value greater than 0.26 were investigated further.

3.2.3 Residuals vs. Fitted Plot

The characteristics of an unbiased residual vs. fitted plot and what they suggest about the appropriateness of the simple linear regression model include:

- 1) The residuals “bounce randomly” around the 0 line. This suggests that the assumption that the relationship is linear is reasonable;
- 2) The residuals roughly form a “horizontal band” around the 0 line. This suggests that the variances of the error terms are equal; and
- 3) No one residual “stands out” from the basic random pattern of residuals. This suggests that there are no outliers.

The above paragraph was taken almost directly from the source: <https://newonlinecourses.science.psu.edu/stat462/node/117/>.

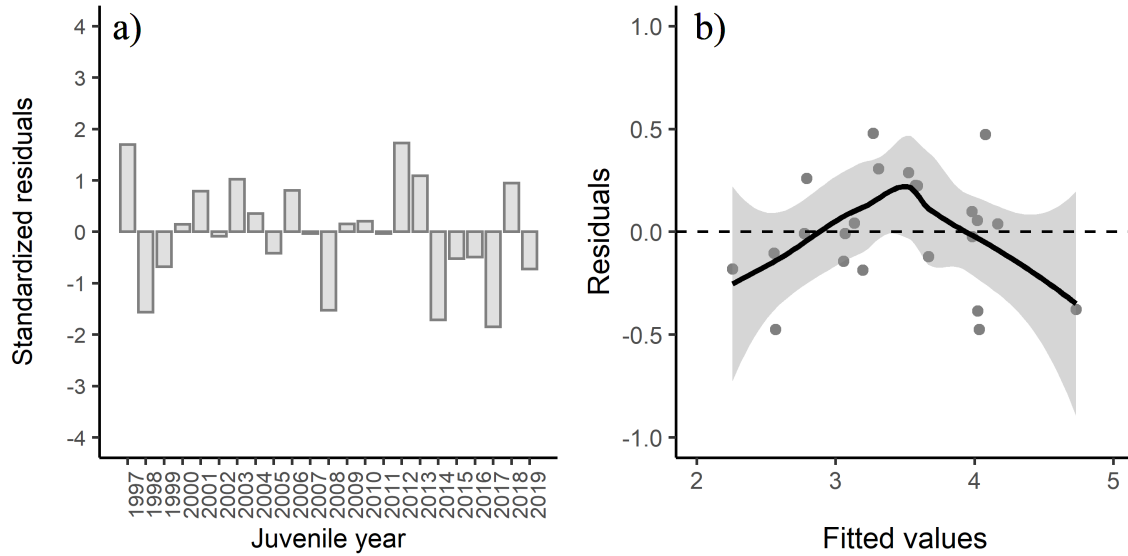


Figure 2: a) Standardized residuals versus juvenile year and b) residuals versus fitted values for model m2. Positive residuals indicate that the observed harvest was larger than predicted by the model.

3.2.4 Residuals vs. Predictor Plots

The interpretation of a “residuals vs. predictor plot” is identical to that for a “residuals vs. fits plot.” That is, a “well-behaved” plot will bounce randomly and form a roughly horizontal band around the residual = 0 line. In addition, no data points will stand out from the basic random pattern of the other residuals. The above paragraph was taken directly from the source: <https://newonlinecourses.science.psu.edu/stat462/node/117/>.

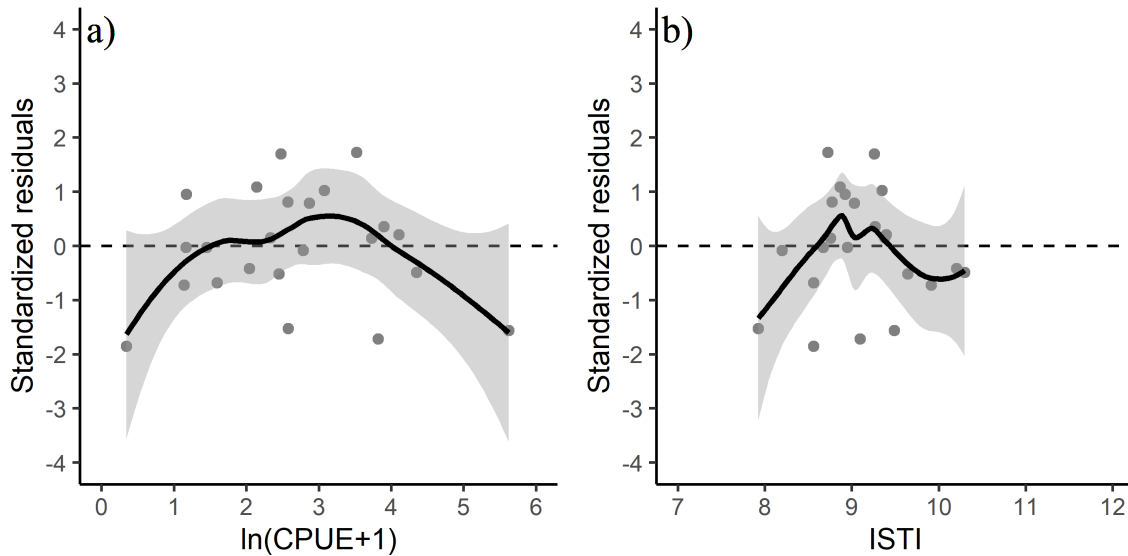


Figure 3: Standardized residuals versus predicted plots for a) CPUE and b) temperature.

3.2.5 Influential Datapoints

To determine if a variable has a relationship with residuals, a lack-of fit curvature test was performed. In this test, terms that are non-significant suggest a properly specified model. The CPUE term was significant

in the lack-of-fit curvature test ($P < 0.05$), suggesting some lack of fit for this term (Figure 3a). Diagnostics indicated that three of the data points were above the cut-off value for the Cook's distance (Figure 4a). Two observations had a high leverage value (Figure 4b), but none of the observations affected model fitting. Based on the Bonferroni outlier test, none of the data points had a studentized residual with a significant Bonferroni P -value suggesting that none of the data points impacted the model fitting; although observations 16, 18 and 21 and were the most extreme (juvenile years 2012, 2014, and 2017 corresponding to years 2013, 2015, and 2018) based on standardized residuals (Figure 2a; Table 7). Based on the lightly curved fitted lines in the residual versus fitted plot (Figure 2b), the fitted plot shows some lack of fit of the model.

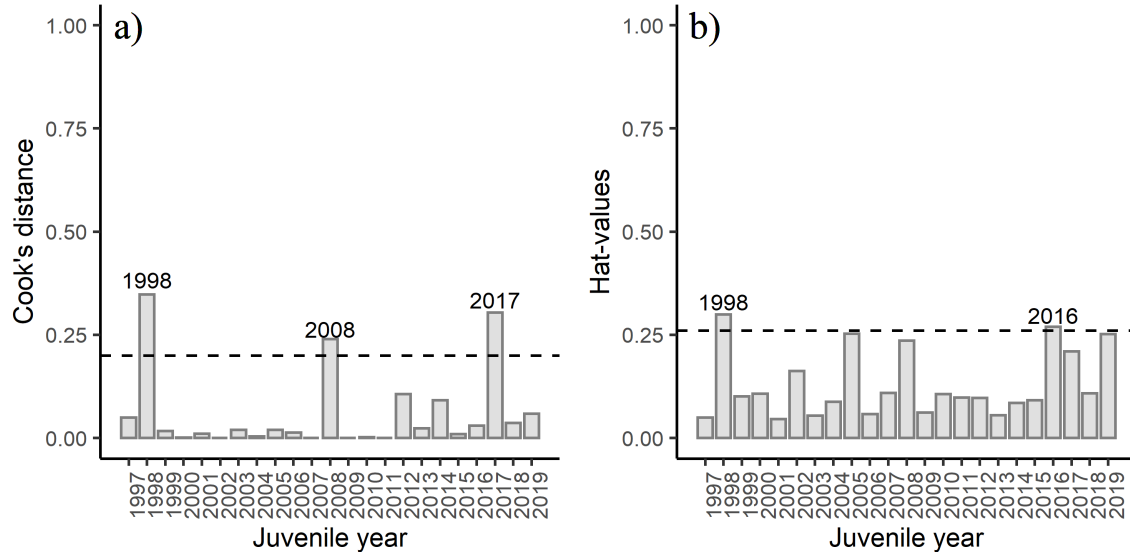


Figure 4: Diagnostics plots of influential observations including a) Cook's Distance (with a cut-off value of 0.20), and b) leverage values (with a cut-off value of 0.26) from model m2.

Table 3: Detailed output for model m2. Juvenile year 2012, 2014, and 2017 (year 2013, 2015, and 2018) show the largest standardized residual. The variable SEAKCatch is commercial harvest of adult fish in millions and the variable CPUE is $\ln(\text{CPUE}+1)$ of outmigrating juvenile pink salmon. Year refers to the forecast year.

year	juvenile_year	SEAKCatch	CPUE	ISTI	resid	hat_values	Cooks_distance	std_resid	fitted
1998	1997	42.5	2.478	9.259	0.478	0.049	0.050	1.697	3.271
1999	1998	77.8	5.622	9.489	-0.378	0.299	0.348	-1.562	4.732
2000	1999	20.3	1.598	8.560	-0.186	0.101	0.017	-0.679	3.197
2001	2000	67.0	3.730	8.756	0.039	0.108	0.001	0.142	4.166
2002	2001	45.3	2.869	9.028	0.224	0.046	0.010	0.792	3.590
2003	2002	52.5	2.785	8.198	-0.023	0.163	0.000	-0.087	3.984
2004	2003	45.3	3.078	9.349	0.288	0.054	0.020	1.023	3.526
2005	2004	59.1	3.899	9.269	0.098	0.088	0.004	0.355	3.981
2006	2005	11.6	2.040	10.203	-0.104	0.253	0.020	-0.418	2.555
2007	2006	44.8	2.573	8.771	0.227	0.058	0.013	0.808	3.576
2008	2007	15.9	1.168	8.951	-0.008	0.109	0.000	-0.031	2.775
2009	2008	38.0	2.577	7.925	-0.385	0.236	0.240	-1.525	4.023
2010	2009	24.0	2.333	9.378	0.042	0.061	0.000	0.150	3.136
2011	2010	58.9	4.108	9.395	0.056	0.107	0.002	0.205	4.020
2012	2011	21.3	1.455	8.669	-0.008	0.098	0.000	-0.031	3.067
2013	2012	94.7	3.522	8.725	0.474	0.097	0.106	1.725	4.077

year	juvenile_year	SEAKCatch	CPUE	ISTI	resid	hat_values	Cooks_distance	std_resid	fitted
2014	2013	37.2	2.143	8.865	0.306	0.056	0.023	1.091	3.310
2015	2014	35.1	3.817	9.093	-0.474	0.085	0.091	-1.716	4.033
2016	2015	18.4	2.453	9.645	-0.144	0.092	0.009	-0.521	3.056
2017	2016	34.7	4.351	10.297	-0.121	0.270	0.030	-0.491	3.668
2018	2017	8.1	0.346	8.560	-0.475	0.210	0.304	-1.850	2.567
2019	2018	21.1	1.172	8.925	0.259	0.108	0.037	0.949	2.790
2020	2019	8.0	1.142	9.911	-0.182	0.252	0.059	-0.727	2.257

3.3 Results

The best regression model based on the AICc value, the MASE metric, and significant coefficients in the model was model m2 (i.e., the model containing CPUE, and a May through July temperature variable). The adjusted R^2 value for model m2 was 0.82 (Table 6) indicating overall a good model fit.

3.4 Conclusion

Based upon a model that includes juvenile pink salmon CPUE and May-July temperature (model m2), the 2021 SEAK pink salmon harvest in 2021 is predicted to be in the moderate range with a point estimate (model mean) of 28.3 million fish (80% prediction interval: 19.1 to 41.9 million fish).

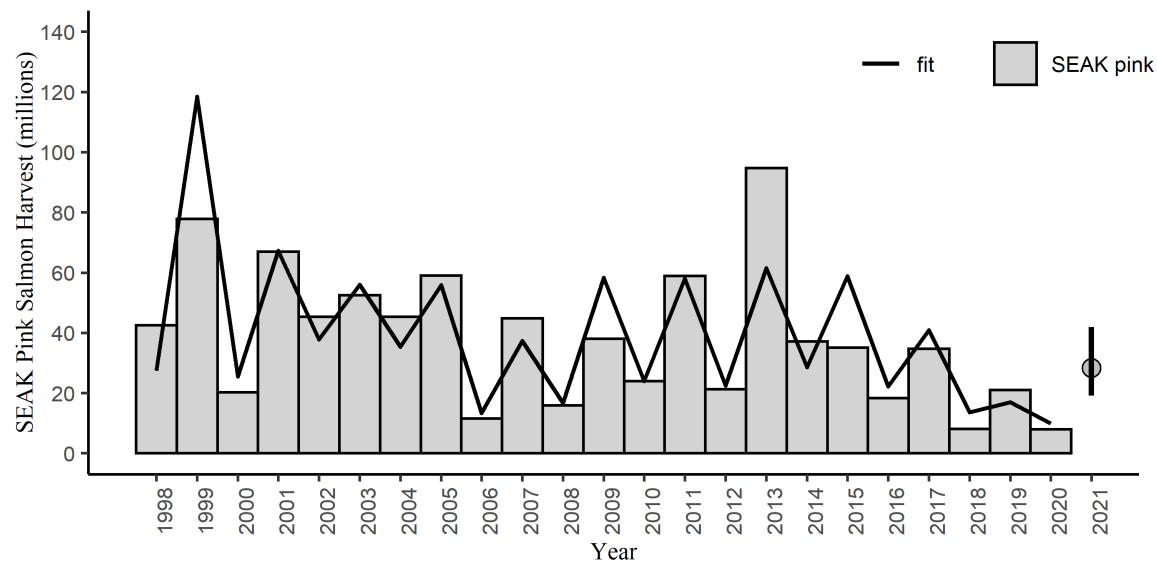


Figure 5: SEAK pink salmon harvest (millions) by year with the model fit (line). The predicted 2021 forecast is symbolized as a grey circle with an 80% prediction interval (19.1 to 41.9 million fish).

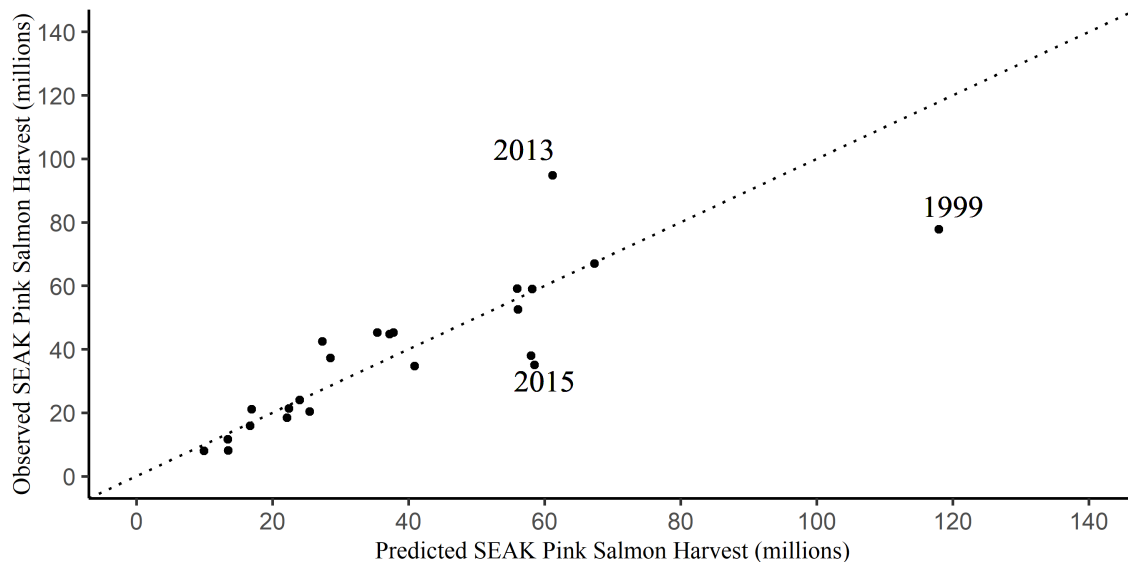


Figure 6: SEAK pink salmon harvest (millions) against the fitted values from model m2 by year. The dotted line is a one to one line.

4 References

- Burnham, K. P., and D. R. Anderson. 2004. Multimodel inference: Understanding AIC and BIC in model selection. *Sociological Methods and Research* 33: 261-304.
- Cook, R. D. 1977. Detection of influential observations in linear regression. *Technometrics* 19: 15-18.
- Fox, J. and S. Weisburg. 2019. *An R Companion to Applied Regression*, Third Edition. Thousand Oaks CA: Sage Publications, Inc.
- Hyndman, R. J. and A. B. Koehler. 2006. Another look at measures of forecast accuracy. *International Journal of Forecasting* 22: 679-688.
- Murphy, J. M., E. A. Fergusson, A. Piston, A. Gray, and E. Farley. 2019. Southeast Alaska pink salmon growth and harvest forecast models. *North Pacific Anadromous Fish Commission Technical Report No. 15*: 75-81.
- Murphy et al. 2020
- Orsi, J. A., E. A. Fergusson, A. C. Wertheimer, E. V. Farley, and P. R. Mundy. 2016. Forecasting pink salmon production in Southeast Alaska using ecosystem indicators in times of climate change. *N. Pac. Anadr. Fish Comm. Bull.* 6: 483-499. (Available at <https://npafc.org>)
- R Core Team. 2020. *R: A language and environment for statistical computing*. R Foundation for Statistical Computing, Vienna, Austria. URL: <http://www.r-project.org/index.html>
- Ren, Y. Y., L. C. Zhou, L. Yang, P. Y. Liu, B. W. Zhao, and H. X. Liu. 2016. Predicting the aquatic toxicity mode of action using logistic regression and linear discriminant analysis. *SAR and QSAR in Environmental Research* 27(9). DOI: 10.1080/1062936X.2016.1229691
- Wertheimer A. C., J. A. Orsi, M. V. Sturdevant, and E. A. Fergusson. 2006. Forecasting pink salmon harvest in Southeast Alaska from juvenile salmon abundance and associated environmental parameters. In *Proceedings of the 22nd Northeast Pacific Pink and Chum Workshop*. Edited by H. Geiger (Rapporteur). Pac. Salmon Comm. Vancouver, British Columbia. pp. 65-72.
- Zhang, Z. 2016. Residuals and regression diagnostics: focusing on logistic regression. *Annals of Translational Medicine* 4: 195.