

Mathematical Basics of Computational Psychiatry: Generative Modeling

Klaas Enno Stephan



Translational Neuromodeling Unit



Universität
Zürich^{UZH}

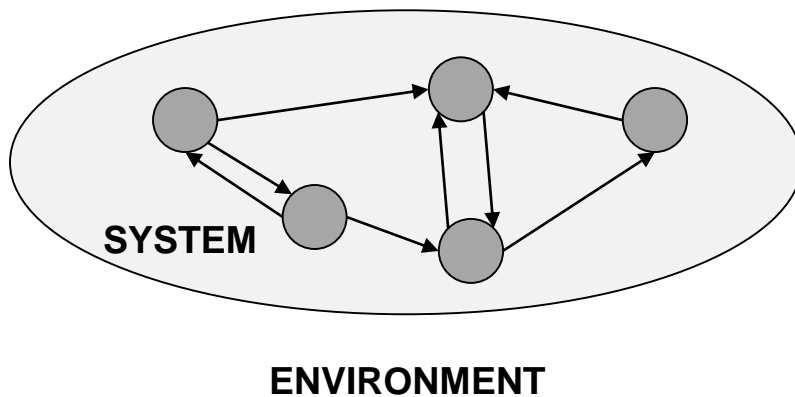


Eidgenössische Technische Hochschule Zürich
Swiss Federal Institute of Technology Zurich

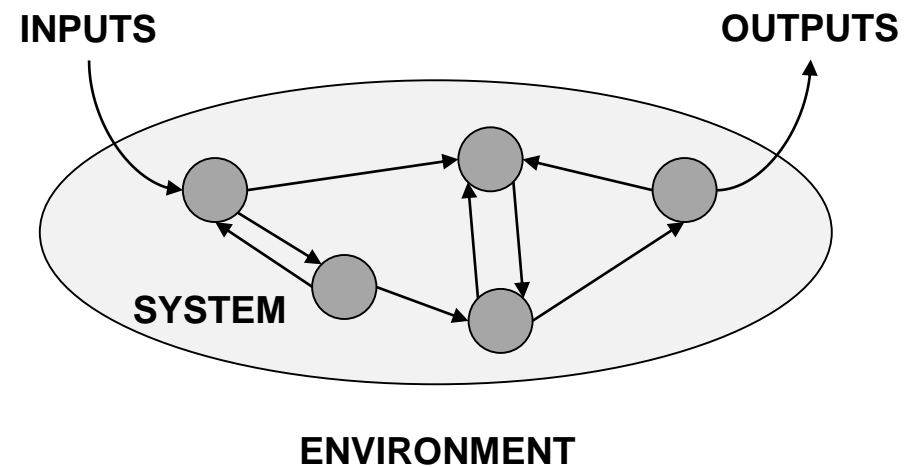
Systems

- system = a set of entities that interact to form a unified whole
- biological systems are open systems: they interact with their environment (exchange of energy, matter, information)

isolated system

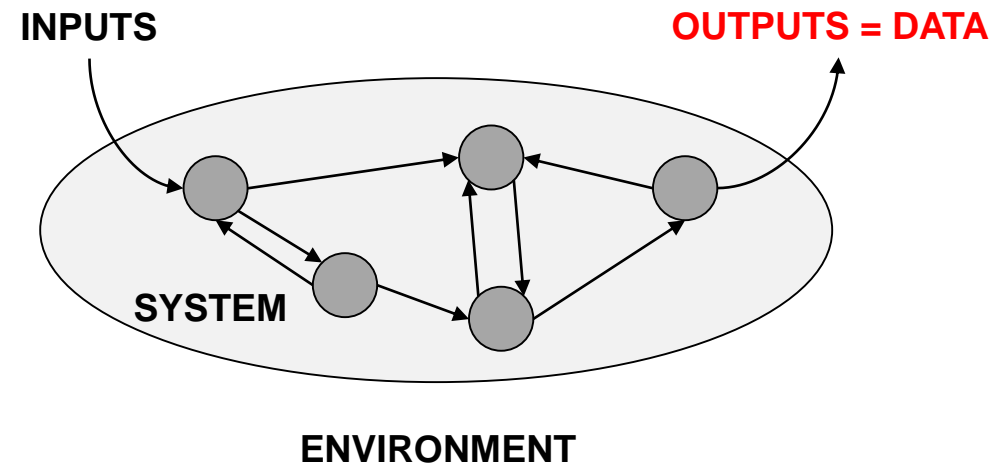


open system



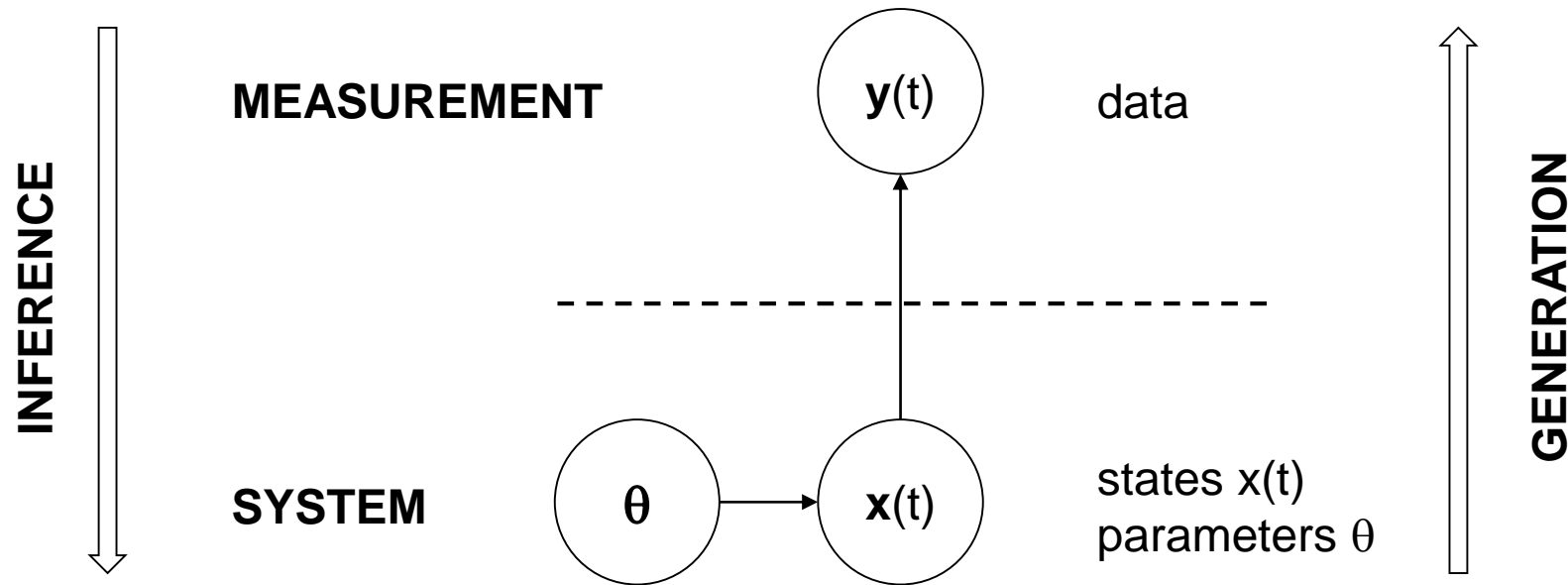
System models (state space models, latent process models)

- mathematically formal description of a system's behavior
(at an algorithmic or biophysical level that cannot be observed directly)
- central concept: hidden (latent) system states cause noisy measurements
- forward models that combine three things:
 - how system states evolve in time
 - how states determine system outputs
 - how outputs are corrupted by measurement noise



Forward modeling

- many ways to categorise modeling approaches
- one possibility: distinguish presence vs. absence of a forward model



States, parameters, inputs

- mandatory system components:
 - what are the relevant variables whose dynamics are of interest? → **states** $\mathbf{x}(t)$
 - what are structural determinants of their interactions? → **parameters** θ
 - what perturbations need to be considered? → **inputs** $\mathbf{u}(t)$
- system states:

state vector

$$\mathbf{x}(t) = \begin{bmatrix} x_1(t) \\ \vdots \\ x_N(t) \end{bmatrix}$$

neurophysiological or
algorithmic variables

state (or evolution) equations, e.g.:

$$\frac{d\mathbf{x}}{dt} = f(\mathbf{x}(t), \theta_f, \mathbf{u}(t)) \quad \text{as differential equation}$$

$$\mathbf{x}(t+1) = f(\mathbf{x}(t), \theta_f, \mathbf{u}(t)) \quad \text{as difference equation}$$

State space representation

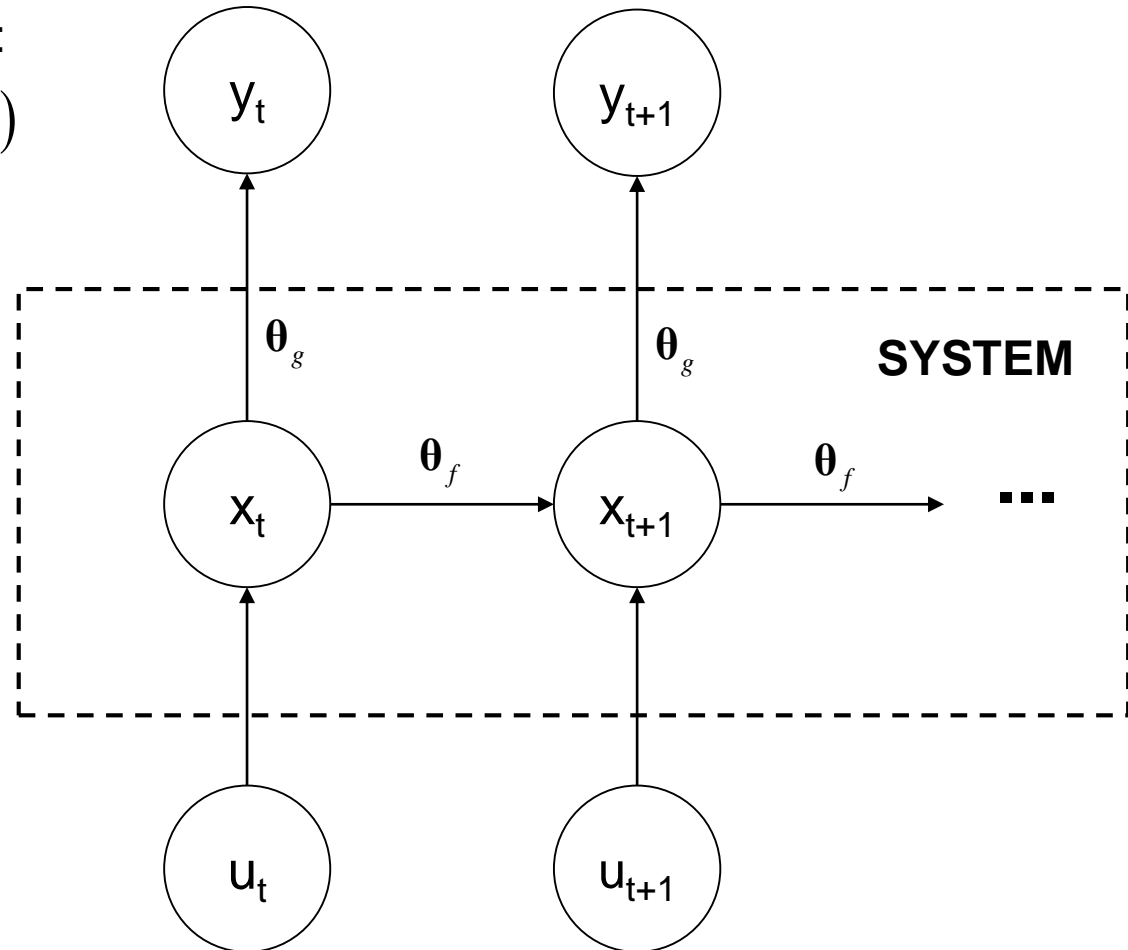
measurement equation:

$$\mathbf{y}(t) = g(\mathbf{x}(t), \boldsymbol{\theta}_g) + \boldsymbol{\varepsilon}(t)$$

ENVIRONMENT

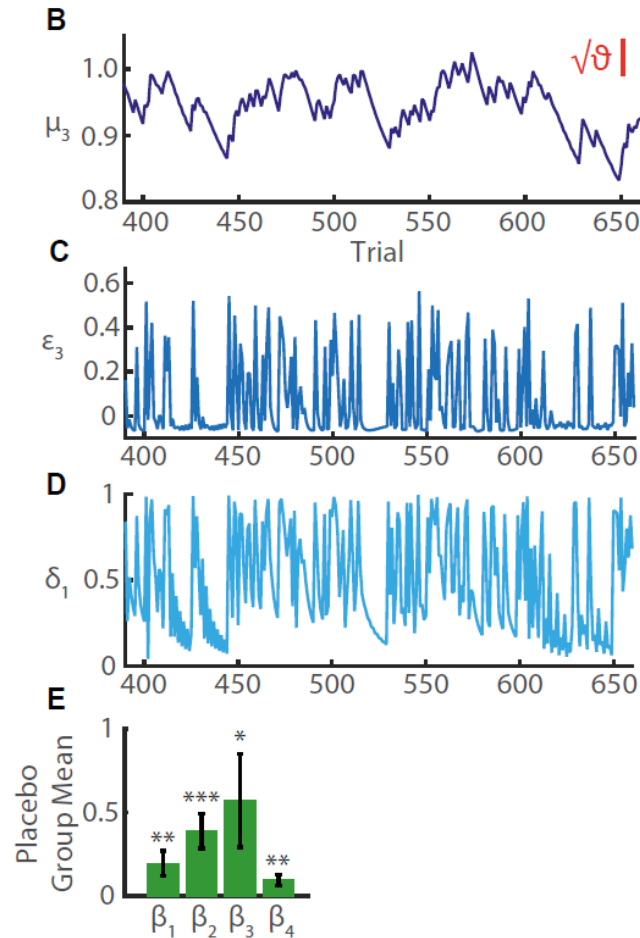
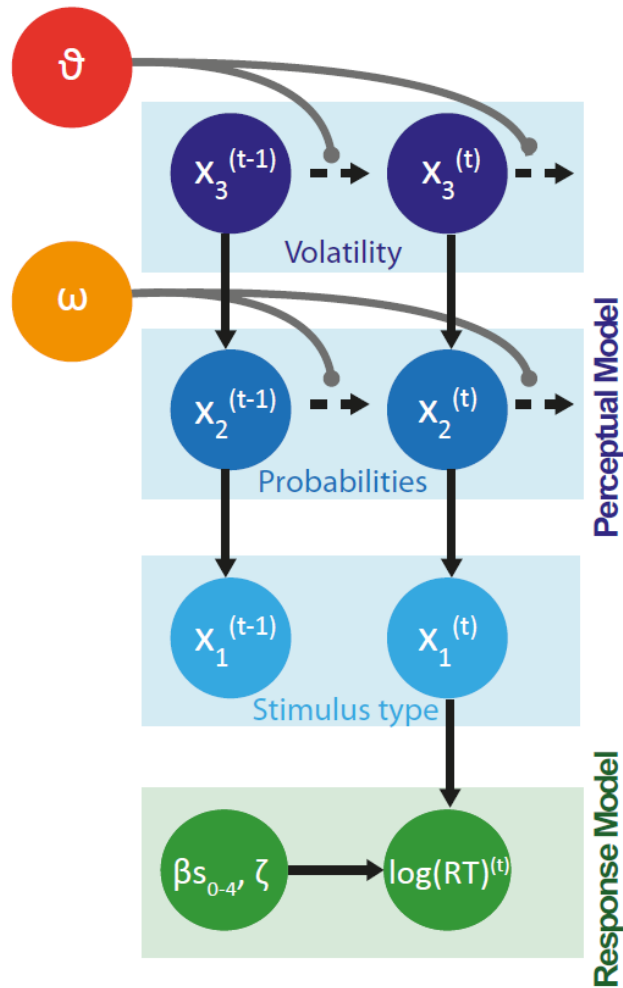
inputs

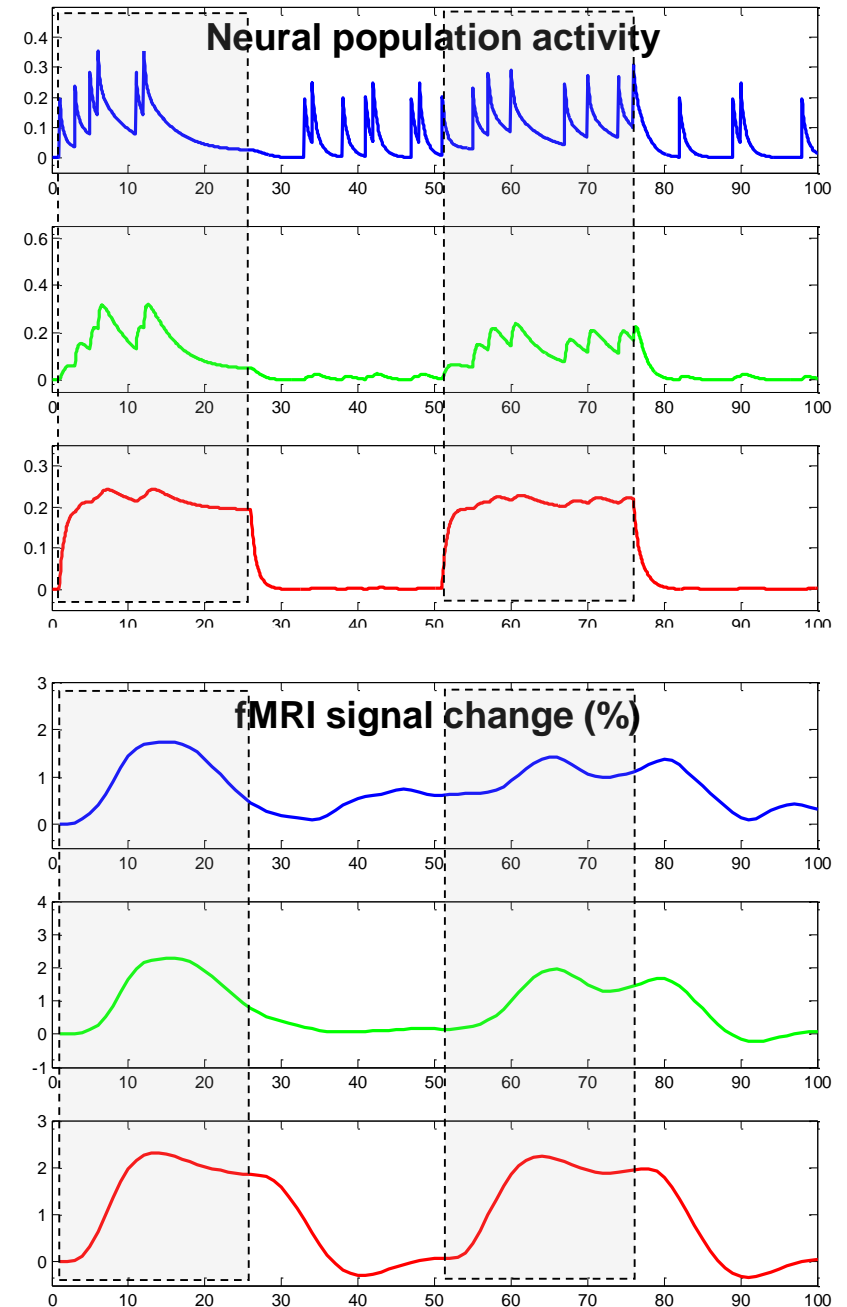
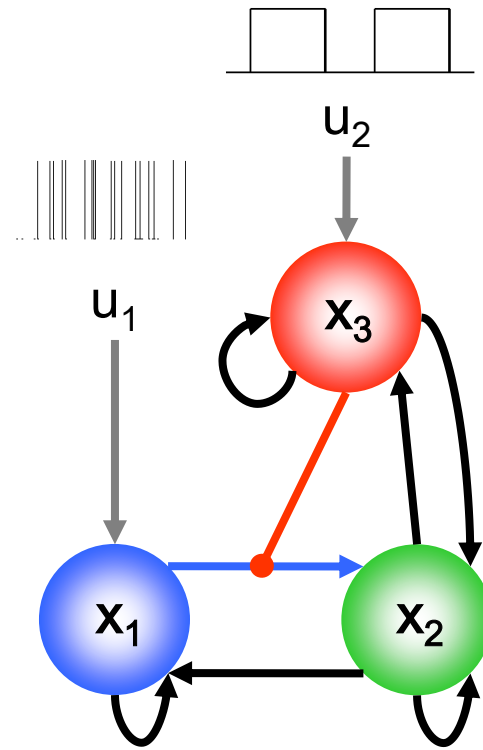
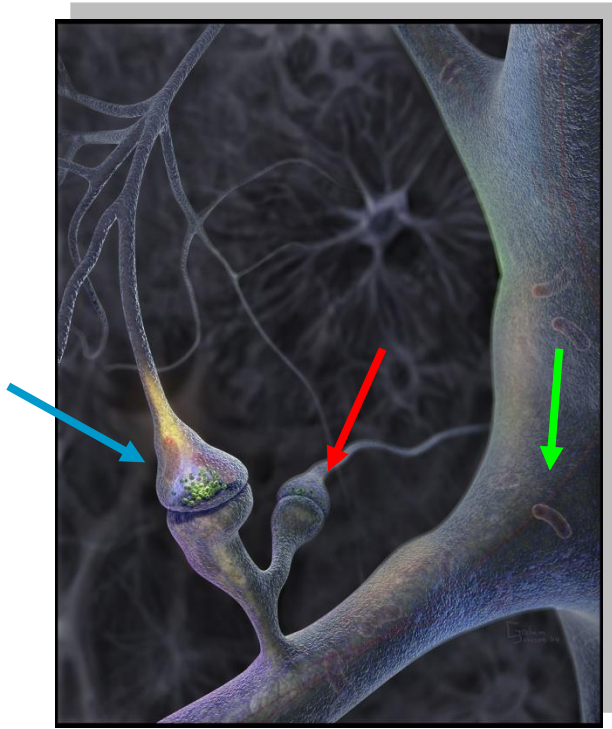
observed system behaviour



On this slide, time is indexed by subscripts.

Examples of models discussed later in the course: HGF...





... and nonlinear DCM for fMRI

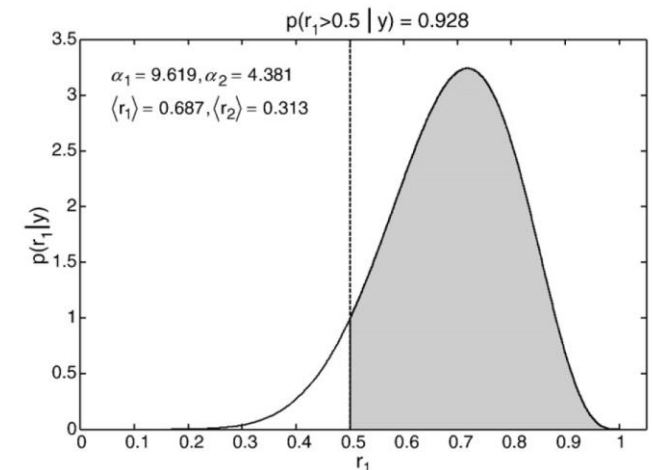
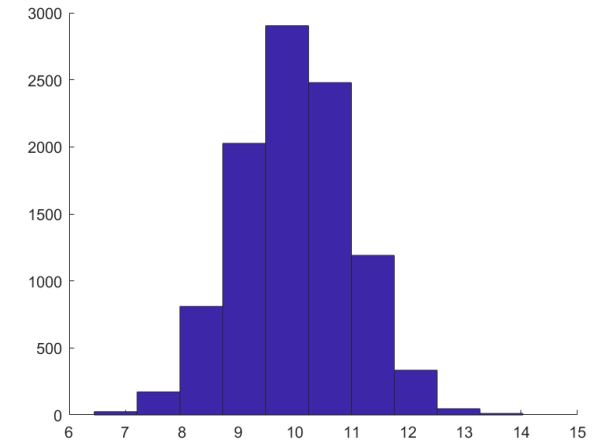
$$\frac{dx}{dt} = \left(A + \sum_{i=1}^m u_i B^{(i)} + \sum_{j=1}^n x_j D^{(j)} \right) x + Cu$$

Statistical interlude: random variables/vectors

- **random variable**: a variable whose possible values are outcomes of a random phenomenon
- **random vector**: a vector of random variables

Statistical interlude: probability distributions and densities

- **probability distribution:**
 - describes the probability that a **discrete** random variable takes on a particular value
- **probability density:**
 - describes the probability of a **continuous** random variable falling within a particular range of values



Statistical interlude: probability distributions and densities

- notation example (Normal densities):
 - for scalars: $p(x) = N(x; \mu, \sigma^2)$ μ = mean; σ^2 = variance
 - for vectors: $p(\mathbf{x}) = N(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$ $\boldsymbol{\Sigma}$ = covariance matrix
= $E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^\top]$

Statistical interlude: probability distributions and densities

- notation example (Normal densities):

- for scalars:

$$p(x) = N(x; \mu, \lambda^{-1})$$

μ = mean; $\lambda = 1/\sigma^2$ = precision

- for vectors:

$$p(\mathbf{x}) = N(\mathbf{x}; \boldsymbol{\mu}, \Lambda^{-1})$$

Λ = precision matrix

same thing, just expressed wrt. precision

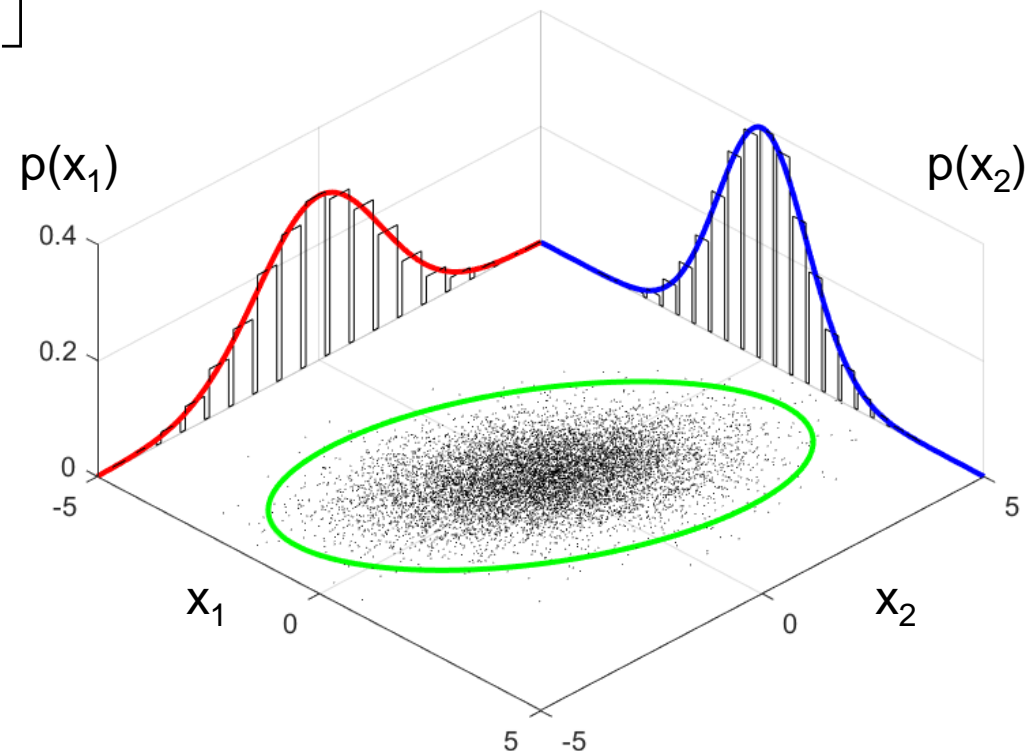
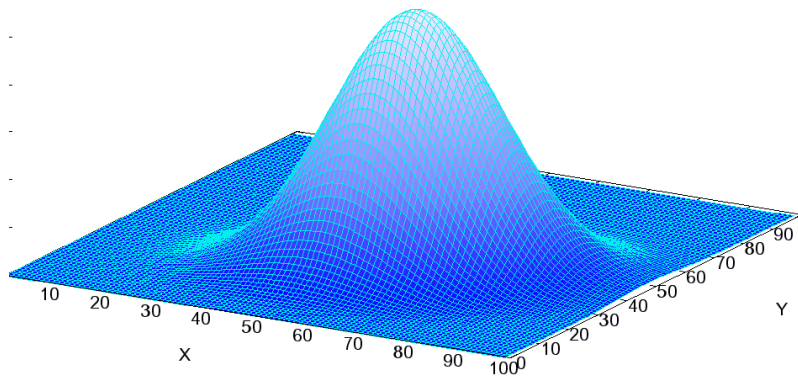
Statistical interlude: multivariate Gaussian/Normal

p-dimensional random vector: $p(\mathbf{x}) = N(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$

PDF:
$$p(\mathbf{x}) = \frac{1}{(2\pi)^{p/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

$$\mathbf{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_p \end{bmatrix}$$

covariance matrix:
$$\boldsymbol{\Sigma} = \mathbf{E}\left[\left((\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T\right)\right]$$



Signal-generating equations (forward model) → likelihood

- State (evolution) equation * $\mathbf{x}(t+1) = f(\mathbf{x}(t), \boldsymbol{\theta}, \mathbf{u}(t))$
- Measurement (observation) equation $\mathbf{y}(t) = g(\mathbf{x}(t), \boldsymbol{\theta}) + \boldsymbol{\varepsilon}(t)$
- Assuming IID Gaussian noise, write the (known) data as a probabilistic function of the (unknown) parameters:
 $\boldsymbol{\varepsilon} = N(\boldsymbol{\varepsilon}; 0, \sigma^2)$
 $p(\mathbf{y} | \boldsymbol{\theta}) = N(\mathbf{y}; g(\mathbf{x}, \boldsymbol{\theta}), \sigma^2 \mathbf{I})$
- This turns our forward model into a probability statement:
the **likelihood** of the observed data \mathbf{y} , given any particular value of $\boldsymbol{\theta}$.

* For simplicity, we assume deterministic state equations (no state noise) and absorb all parameters into a single vector $\boldsymbol{\theta} = \{\theta_f, \theta_g\}$.

Maximum likelihood estimation (MLE)

- For any particular value of θ , we can refer to the definition of a multivariate Gaussian to compute the **likelihood of the entire dataset \mathbf{Y}** (all system nodes, all time points):

$$p(\mathbf{y} | \theta) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{y} - g(\mathbf{x}, \theta))^T \Sigma^{-1}(\mathbf{y} - g(\mathbf{x}, \theta))\right)$$

$$p(\mathbf{Y} | \theta) = p(\mathbf{y}(1), \dots, \mathbf{y}(T) | \theta) = \prod_{t=1}^T p(\mathbf{y}(t) | \theta)$$

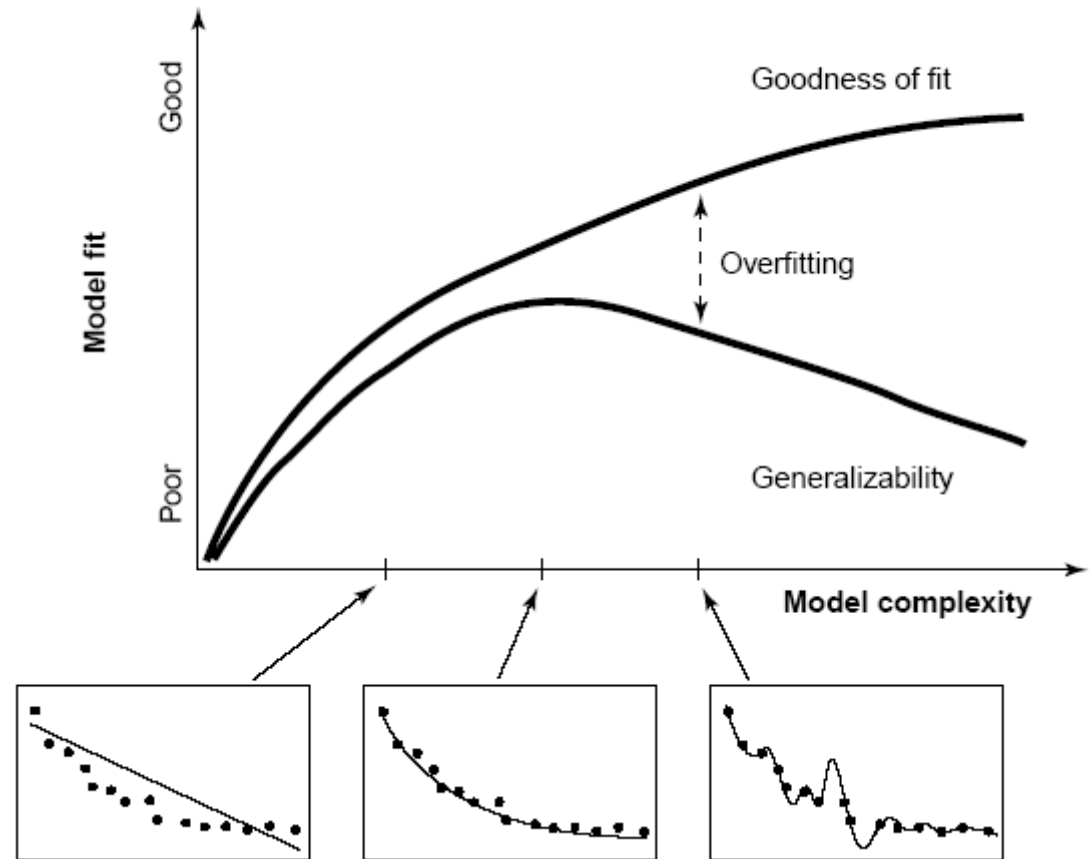
- We could now search for the parameter value that maximises the likelihood (or, for numerical reasons, the log likelihood), or put simply: the parameter value for which the model fits the data best.

This is known as **maximum likelihood estimation (MLE)**:

$$\hat{\theta}_{ML} = \arg \max_{\theta} \ln p(\mathbf{Y} | \theta)$$

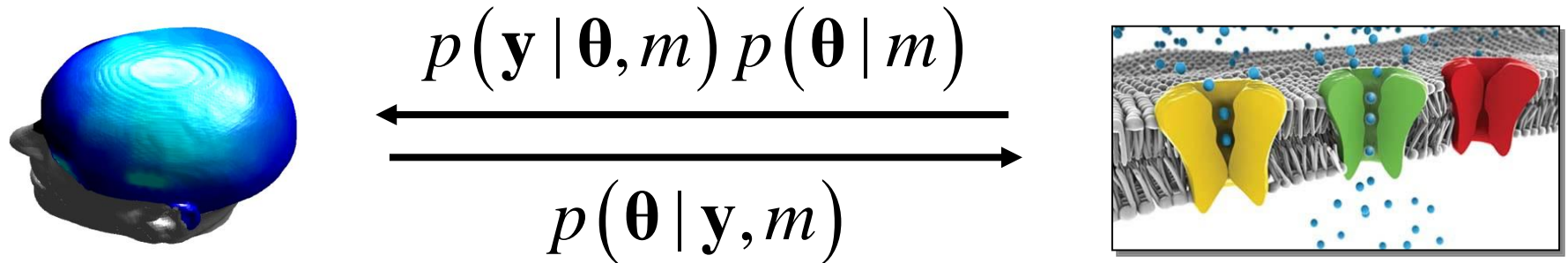
Overfitting

- MLE has various limitations. For example, for complex models and limited data, **overfitting** is a severe problem.
- For more robust inference, we turn to Bayesian methods
→ need to define a prior distribution of parameters
- Together, likelihood and prior define a **generative model**.



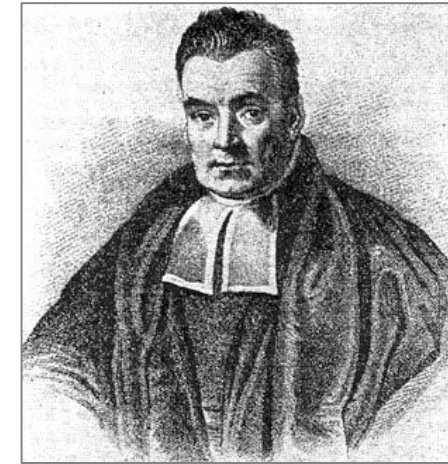
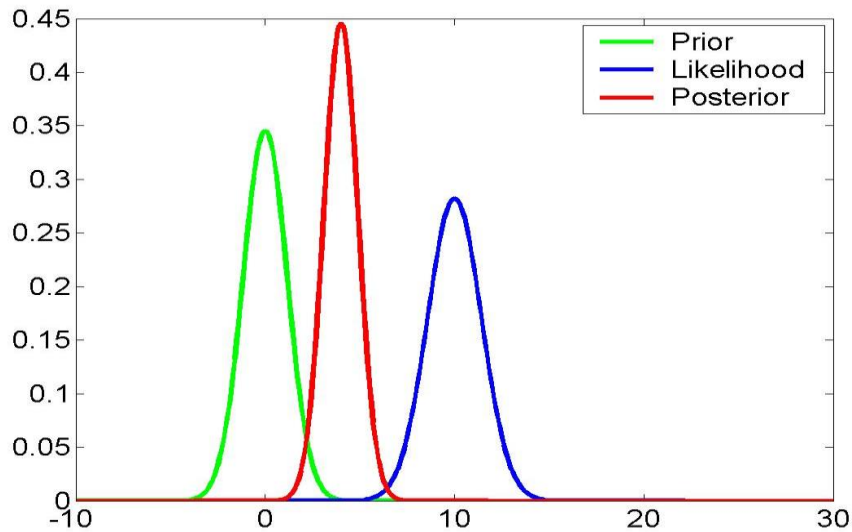
Pitt & Myung (2002) *TICS*

Generative models



1. a probabilistic forward mapping from parameters to data, defined by likelihood and prior
2. provide the joint probability of parameters and data
3. enforce mechanistic thinking: how could the data have been caused?
4. generate synthetic data (observations) by sampling from the prior – can model explain certain phenomena at all?
5. model inversion = inference about parameters $\rightarrow p(\boldsymbol{\theta} | \mathbf{y})$

Bayes' theorem

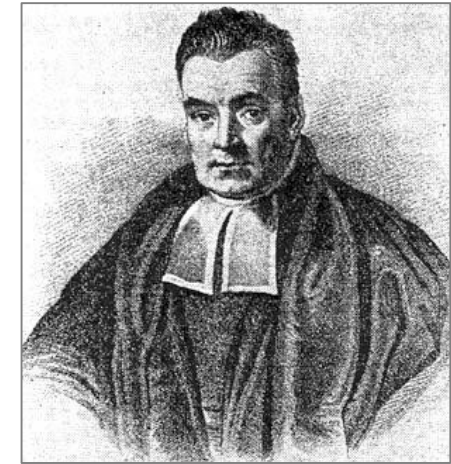
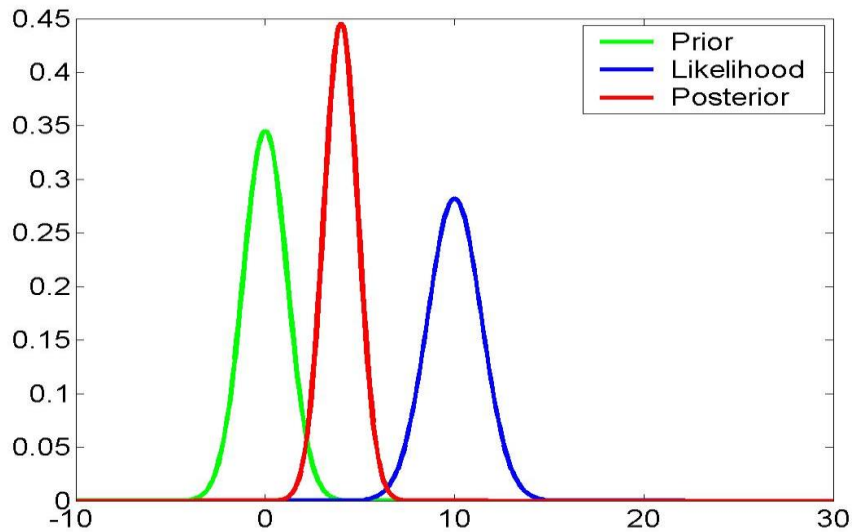


The Reverend Thomas Bayes
(1702-1761)

$$p(\boldsymbol{\theta} | \mathbf{y}) = \frac{p(\mathbf{y} | \boldsymbol{\theta}) p(\boldsymbol{\theta})}{p(\mathbf{y})}$$

posterior = likelihood • prior / evidence

Bayes' theorem



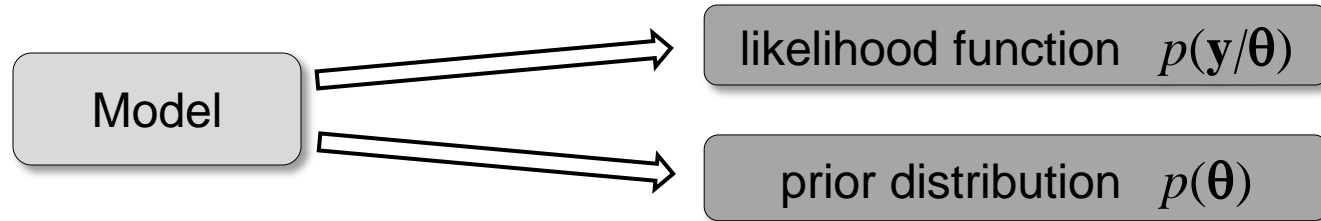
The Reverend Thomas Bayes
(1702-1761)

$$p(\boldsymbol{\theta} | \mathbf{y}) = \frac{p(\mathbf{y} | \boldsymbol{\theta}) p(\boldsymbol{\theta})}{\int p(\mathbf{y} | \boldsymbol{\theta}) p(\boldsymbol{\theta})}$$

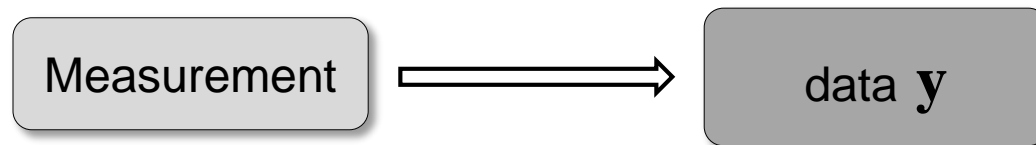
posterior = likelihood • prior / evidence

Principles of generative modeling

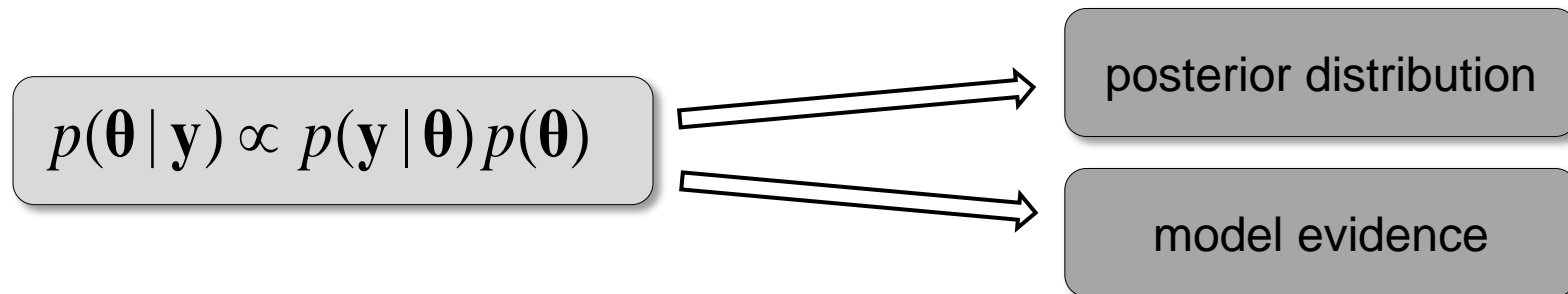
⇒ Specifying a **generative model**



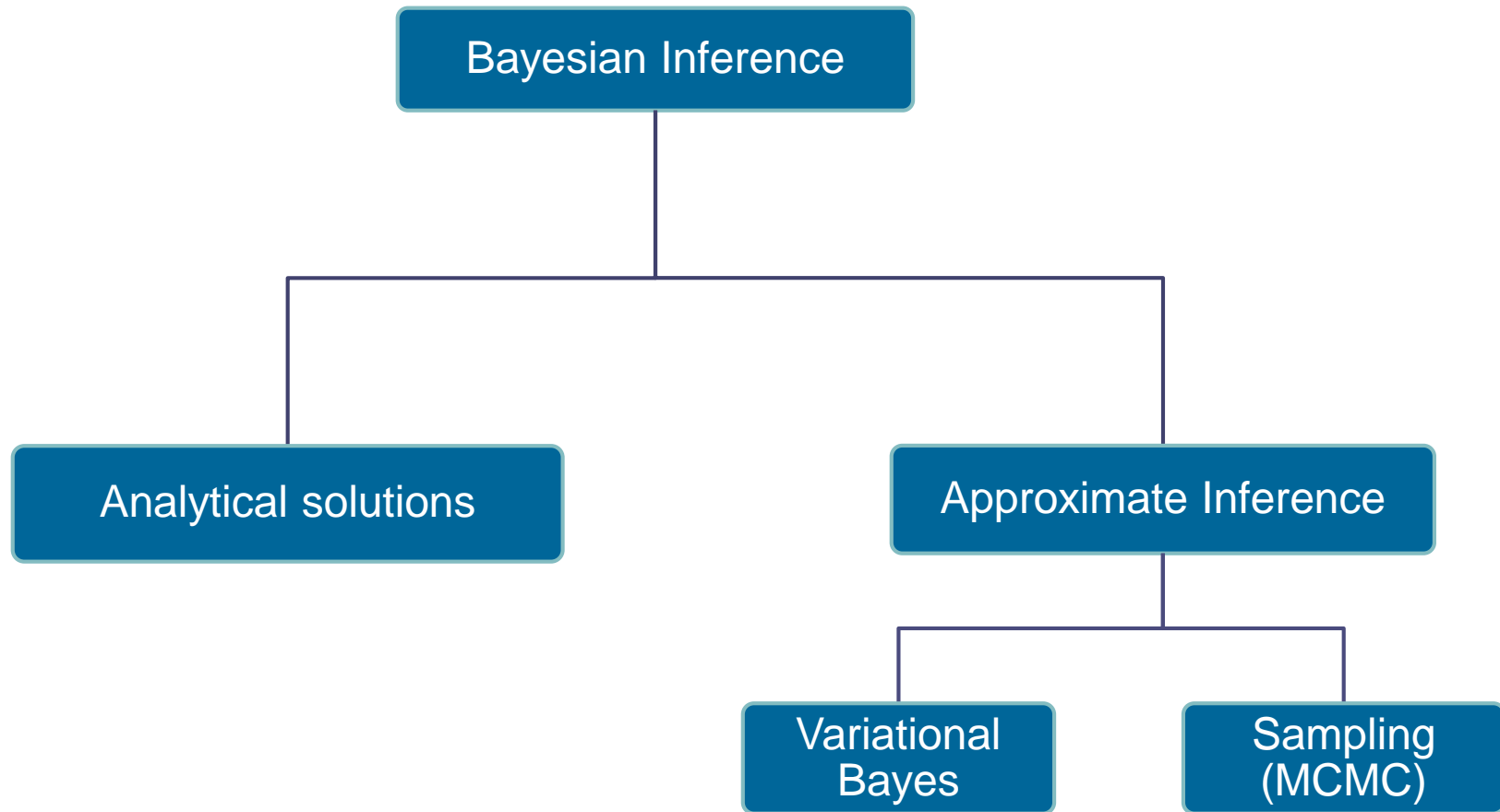
⇒ Observation of **data**



⇒ **Model inversion**



Methods for model inversion




How is the posterior computed = how is a generative model inverted?

- **compute the posterior analytically**
 - requires conjugate priors
- **variational Bayes (VB)**
 - often hard work to derive, but fast to compute
 - uses approximations (approx. posterior, mean field)
 - problems: local minima, potentially inaccurate approximations
- **Sampling: Markov Chain Monte Carlo (MCMC)**
 - theoretically guaranteed to be accurate (for infinite computation time)
 - problems: may require very long run time in practice, convergence difficult to prove

Conjugate priors

- for a given likelihood function, the choice of prior determines the mathematical form of the posterior
- for some probability distributions a prior can be found such that the posterior has the same mathematical form as the prior
- such a prior is called “conjugate” to the likelihood
- examples (prior \times likelihood \propto posterior):
 - Normal \times Normal \propto Normal
 - Beta \times Binomial \propto Beta
 - Dirichlet \times Multinomial \propto Dirichlet

$$p(\boldsymbol{\theta} | \mathbf{y}) \propto p(\mathbf{y} | \boldsymbol{\theta}) p(\boldsymbol{\theta})$$


same form

A simple example: univariate Gaussian belief update

Likelihood & prior

$$p(\mathbf{y} | \theta) = N(\theta, \lambda_e^{-1})$$

$$p(\theta) = N(\mu_p, \lambda_p^{-1})$$

Posterior: $p(\theta | y) = N(\mu, \lambda^{-1})$

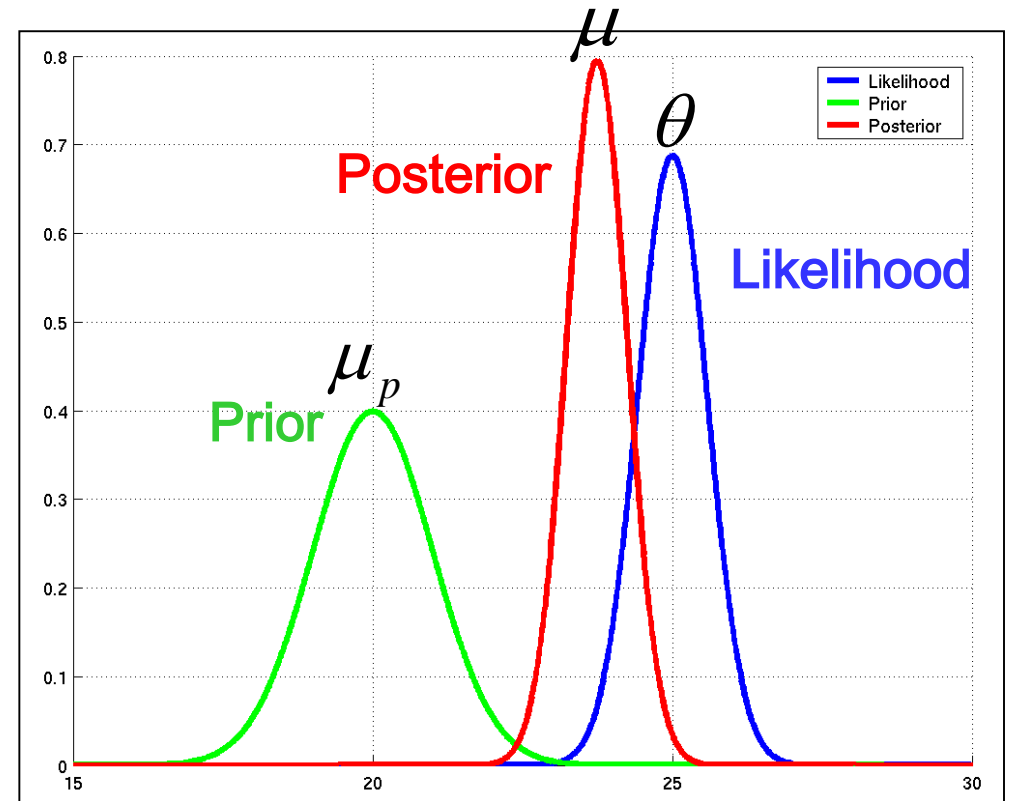
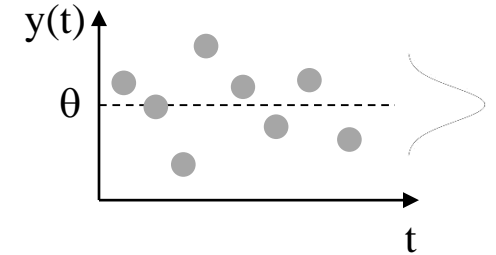
$$\lambda = \lambda_e + \lambda_p$$

$$\mu = \frac{\lambda_e}{\lambda} \bar{y} + \frac{\lambda_p}{\lambda} \mu_p$$

relative precision weighting:

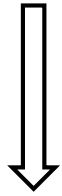
posterior mean = precision-weighted
combination of prior mean and data mean

$$\mathbf{y} = \theta + \boldsymbol{\varepsilon}$$

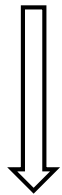


Model comparison and selection

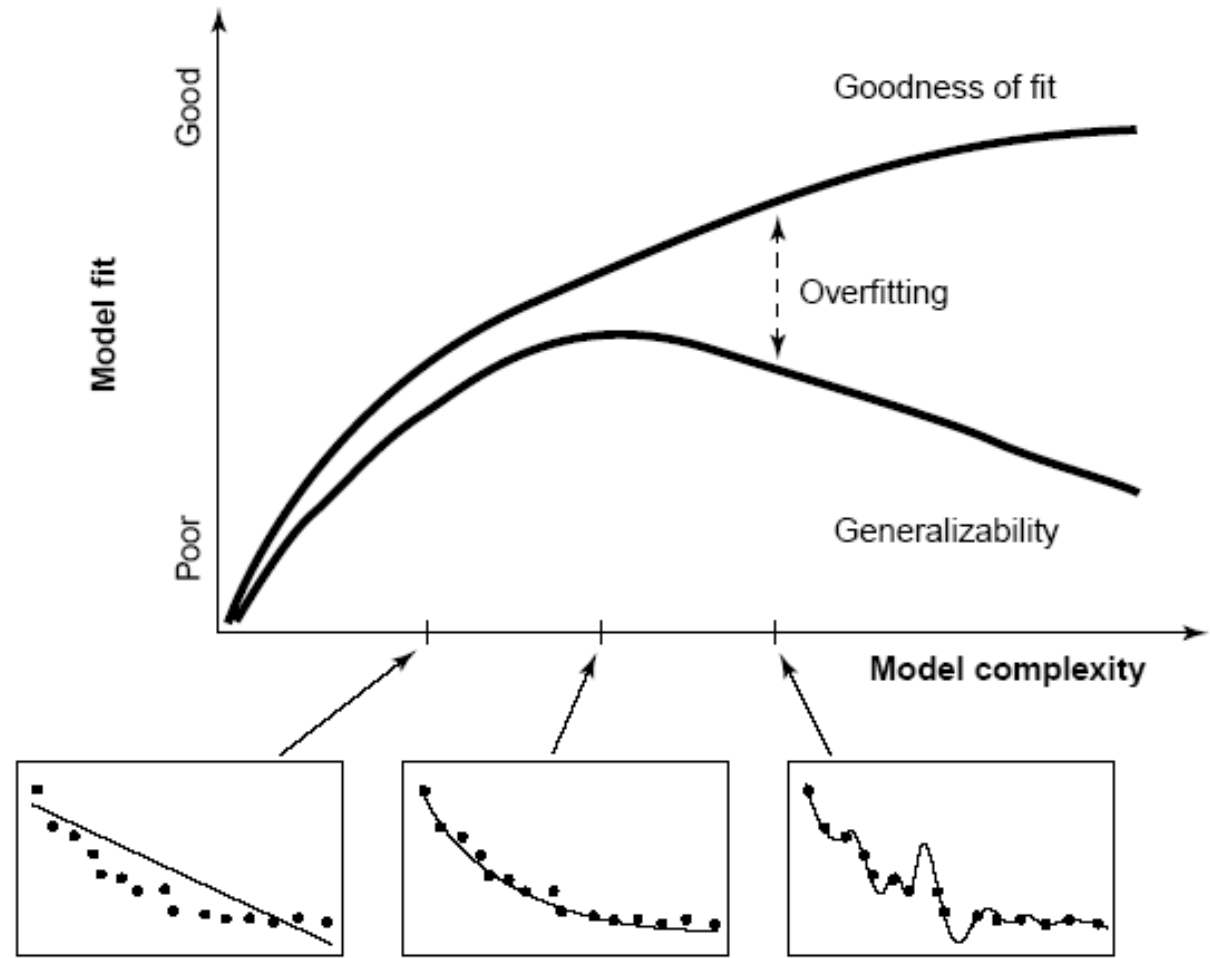
Given competing hypotheses on structure & functional mechanisms of a system, which model is the best?



Which model represents the best balance between model fit and model complexity?



For which model m does $p(y|m)$ become maximal?



Bayesian model selection (BMS)

- First step of inference: define model space M

$$|M| \in [1, \infty[$$

- Inference on model structure m :

Posterior model probability

$$\begin{aligned} p(m | y) &= \frac{p(y | m) p(m)}{p(y)} \\ &= \frac{p(y | m) p(m)}{\sum_m p(y | m) p(m)} \end{aligned}$$

- For a uniform prior on m , model evidence sufficient for model selection

Model evidence:

$$p(y | m) = \int p(y | \theta, m) p(\theta | m) d\theta$$

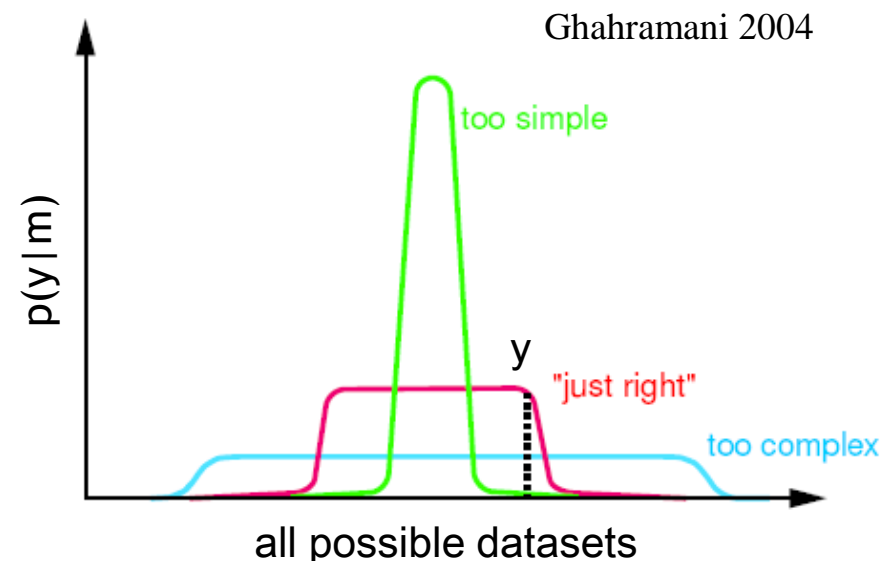
Bayesian model selection (BMS)

Model evidence:

$$p(y | m) = \int p(y | \theta, m) p(\theta | m) d\theta$$

⇒ probability that data were generated by model m , averaging over all possible parameter values (as specified by the prior)

⇒ accounts for both accuracy and complexity of the model



Various approximations:

- negative free energy (F)
- Akaike Information Criterion (AIC)
- Bayesian Information Criterion (BIC)

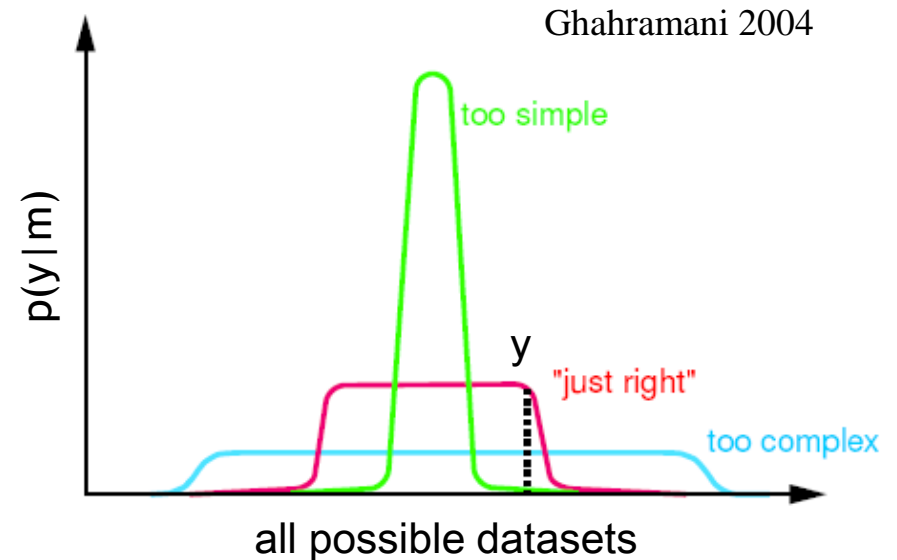
Bayesian model selection (BMS)

Model evidence:

$$p(y | m) = \int p(y | \theta, m) p(\theta | m) d\theta$$

⇒ “If I randomly sampled from my prior and plugged the resulting value into the likelihood function, how close would the predicted data be – on average – to my observed data?”

⇒ accounts for both accuracy and complexity of the model



Various approximations:

- negative free energy (F)
- Akaike Information Criterion (AIC)
- Bayesian Information Criterion (BIC)

Generative models as computational assays for addressing key clinical questions

SYMPTOMS

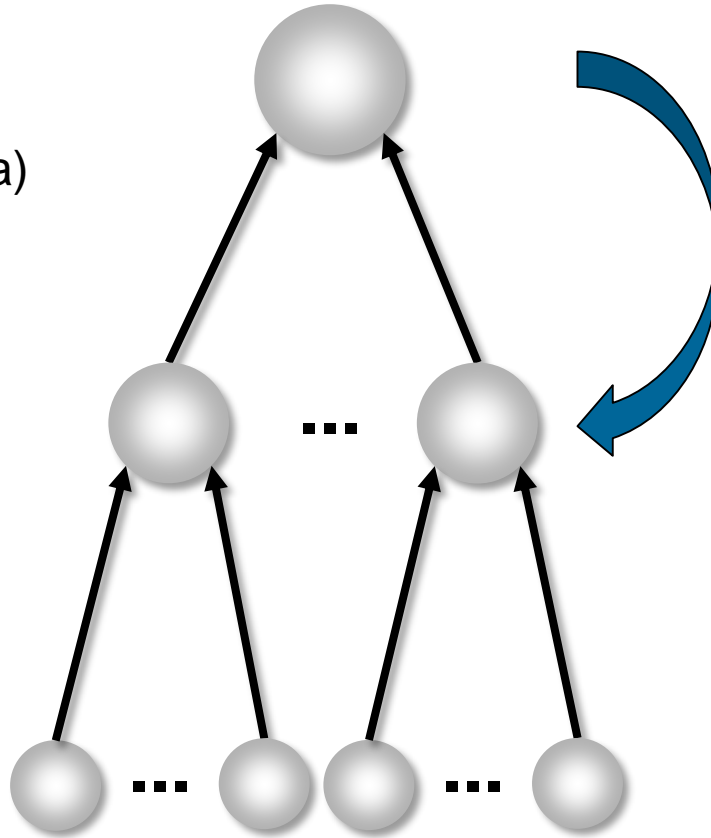
(behavioural or physiological data)

MECHANISMS

(computational, physiological)

CAUSES

(aetiology)

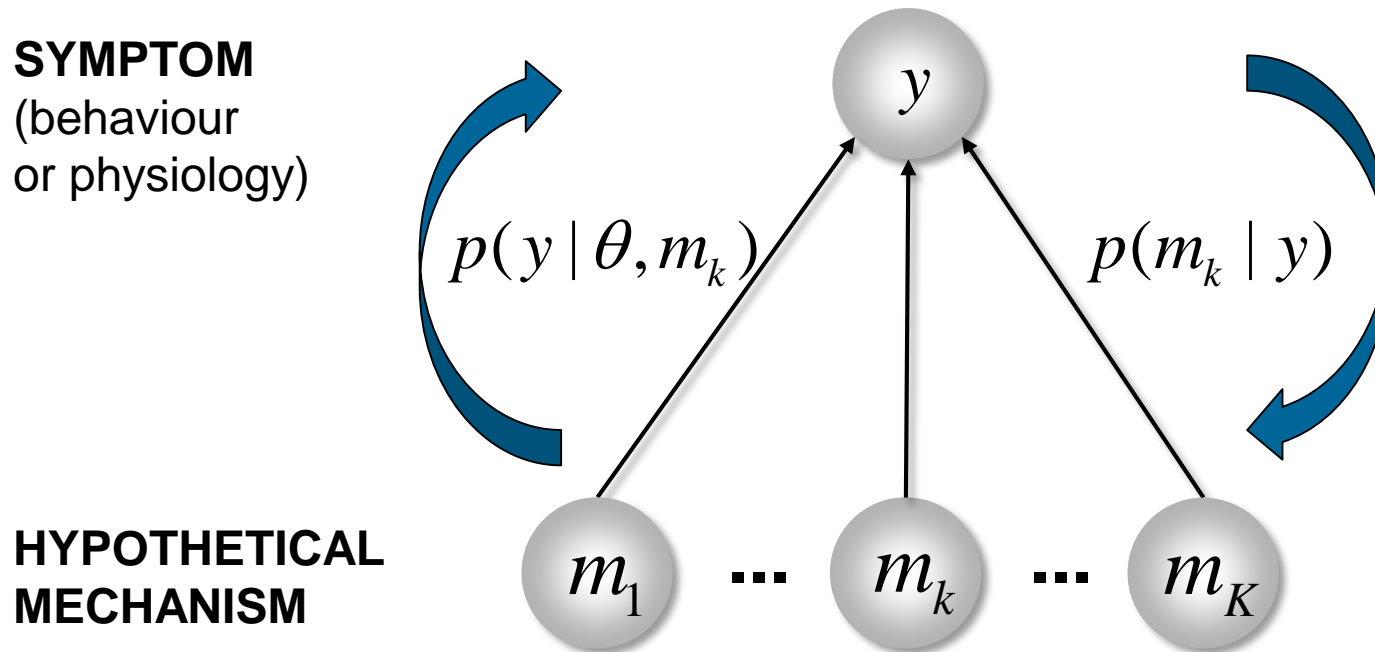


① differential diagnosis of alternative disease mechanisms

② stratification / subgroup detection into mechanistically distinct subgroups

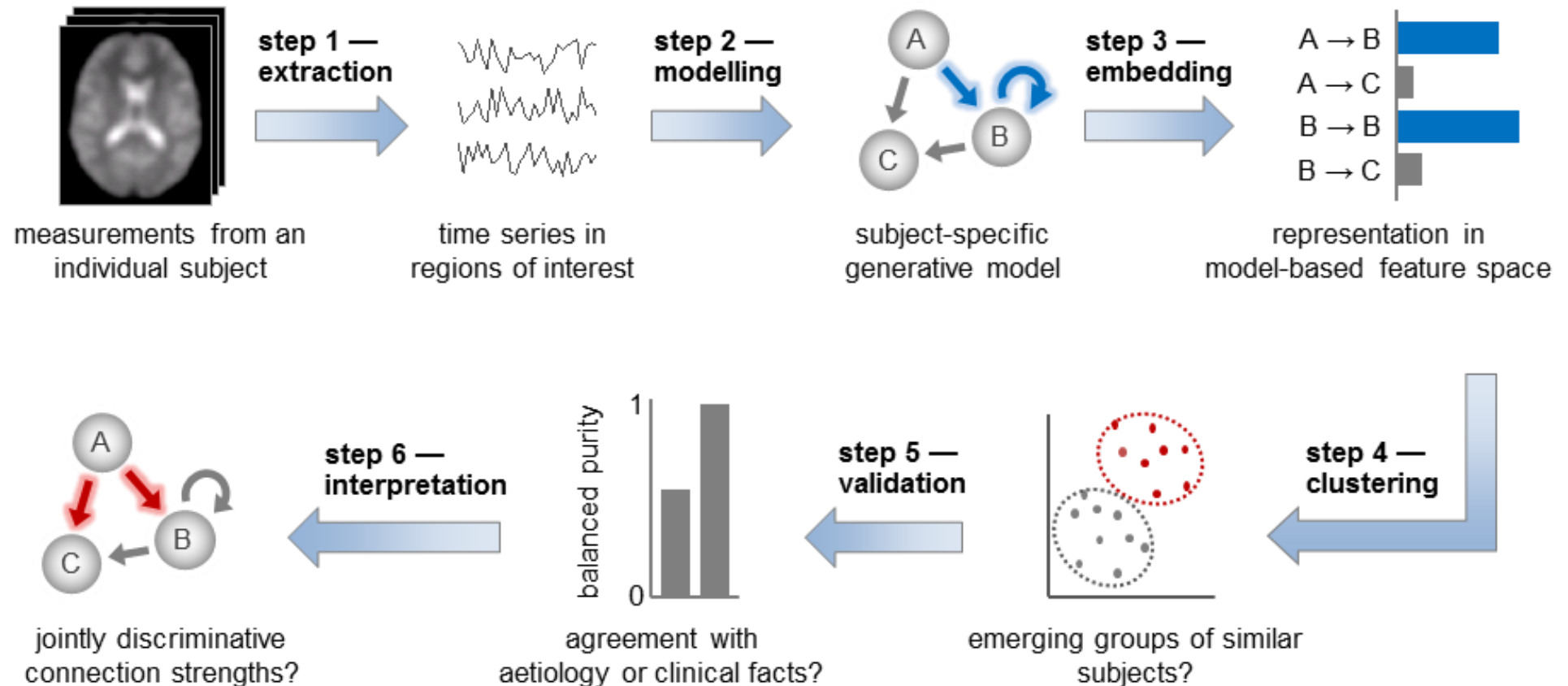
③ prediction of clinical trajectories and treatment response

❶ Differential diagnosis: model selection

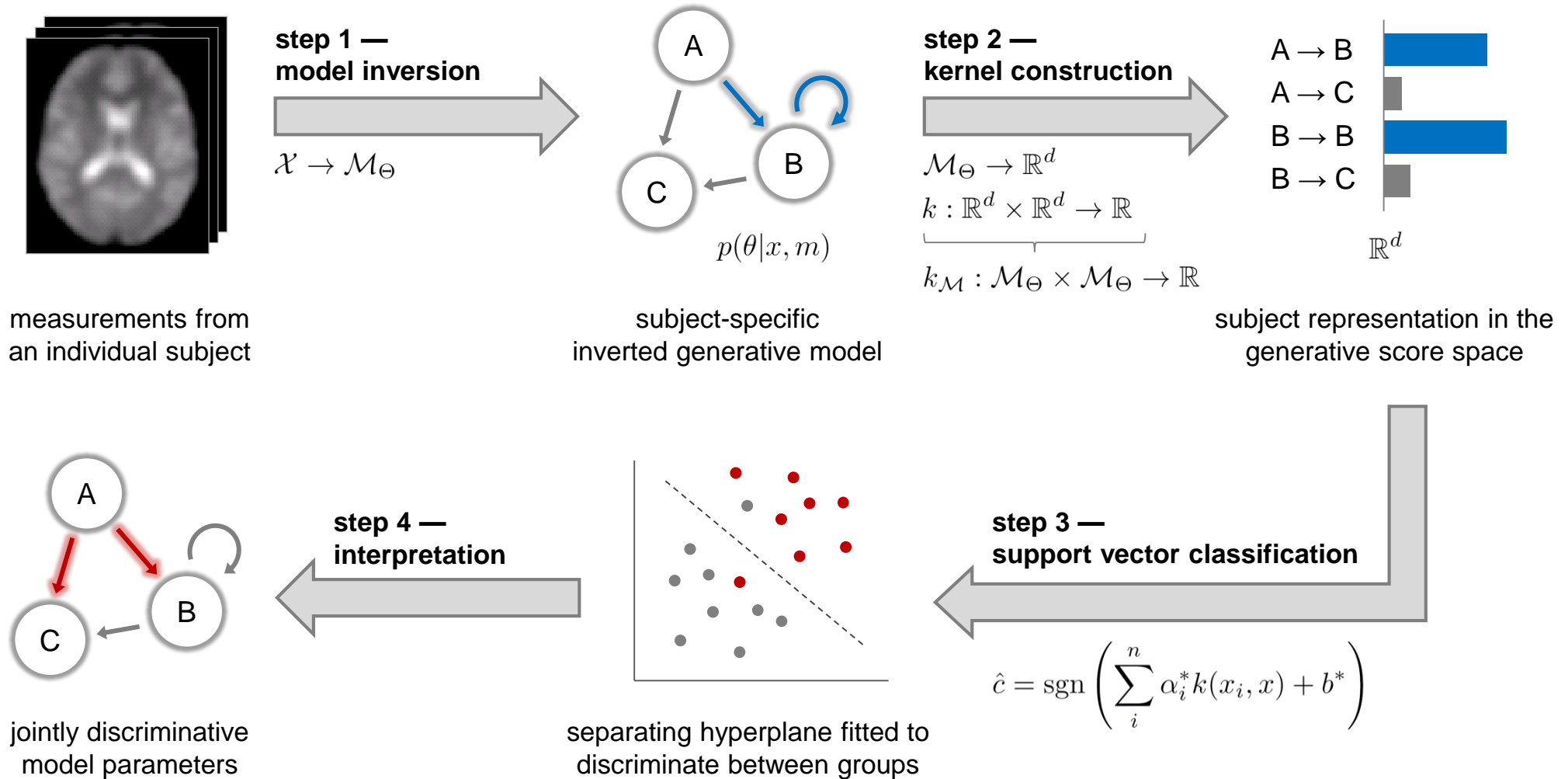


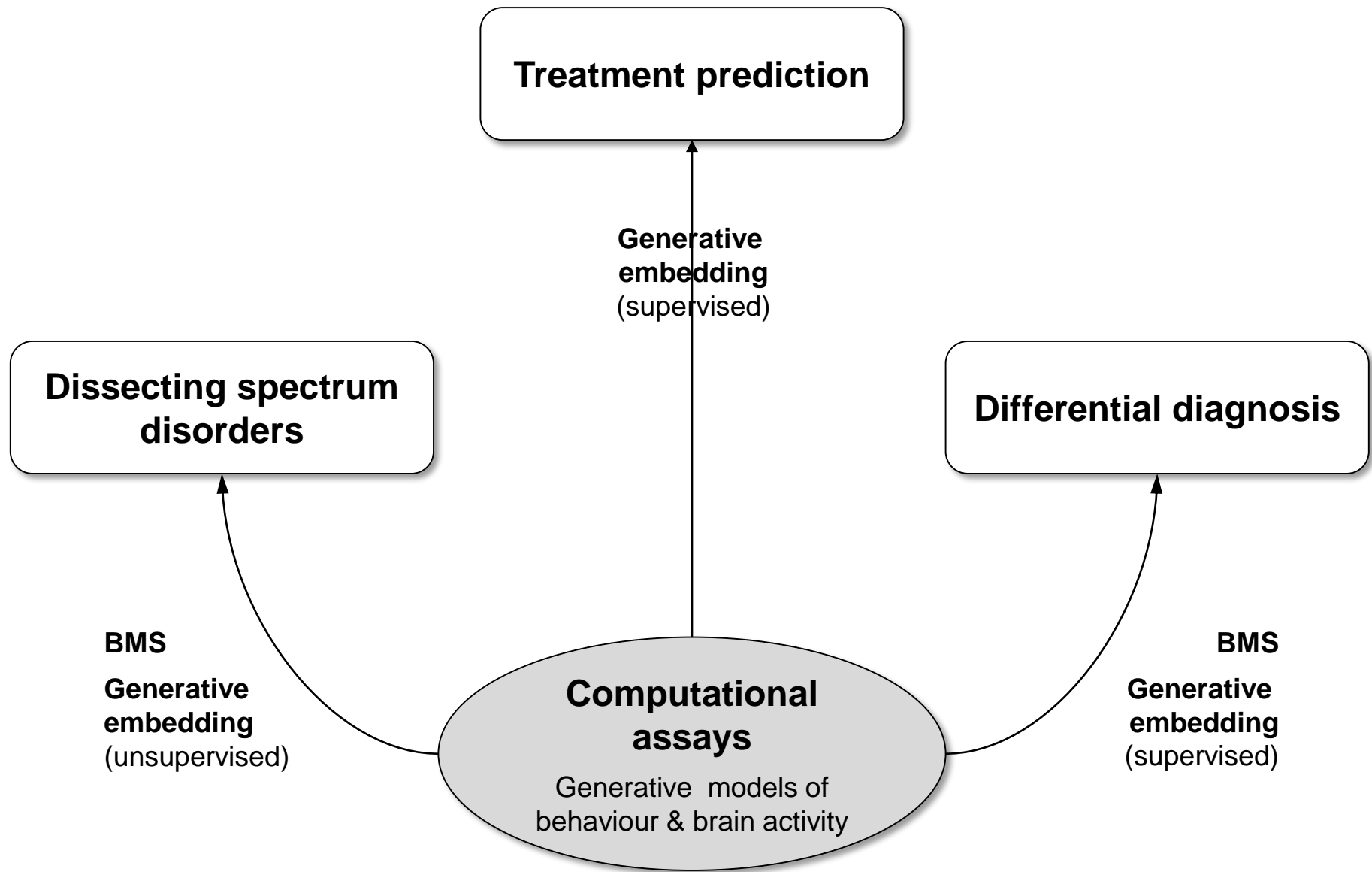
$$p(m_k | y) = \frac{p(y | m_k) p(m_k)}{\sum_k p(y | m_k) p(m_k)}$$

② Stratification / subgroup detection: Generative embedding (unsupervised)



③ Prediction: Generative embedding (supervised)





Further reading

- Brodersen, K.H., Schofield, T.M., Leff, A.P., Ong, C.S., Lomakina, E.I., Buhmann, J.M., Stephan, K.E., 2011. Generative embedding for model-based classification of fMRI data. PLoS Comput. Biol. 7, e1002079
- Brodersen, K.H., Deserno, L., Schlagenhauf, F., Lin, Z., Penny, W.D., Buhmann, J.M., Stephan, K.E., 2014. Dissecting psychiatric spectrum disorders by generative embedding. Neuroimage Clin. 4, 98–111.
- Stephan KE (2004) On the role of general system theory for functional neuroimaging. Journal of Anatomy 205: 443-470.
- Stephan KE, Schlagenhauf F, Huys QJM, Raman S, Aponte EA, Brodersen KH, Rigoux L, Moran RJ, Daunizeau J, Dolan RJ, Friston KJ, Heinz A (2017) Computational Neuroimaging Strategies for Single Patient Predictions. NeuroImage 145:180-199

Thank you