

# Week11

November 10, 2021

## 1 Week11

Up until this point in lecture we discussed how to model the conditional distribution of a random variable with a normal distribution—linear regression. Linear regression can be used to understand how a set of covariates  $(x_1, x_2, \dots, x_p)$  relate linearly to a random variable  $Y$ . We assume that  $Y$  is a normally distributed variable. Because we assume  $Y$  can be modeled as a normally distributed random variable we expect that our  $Y$  data is some set of negative, positive decimal data (i.e. -2.12, 3.14, 78.2, etc).

But what if our  $Y$  data is instead a set of 0s and 1s that represent the absence and presence of some phenomena? Our plan is to explore logistic regression.

The goal of logistic regression is to model a set of  $Y$  data with binary responses (Yes/No), (Presence/Absence).

### 1.1 Data setup

Suppose we have a dataset of  $p$  covariates and a single target variable  $y$ . Then our dataset can be represented as

$$\mathcal{D} = [(x_1^1, x_1^2, \dots, x_1^p, y^1), (x_2^1, x_2^2, \dots, x_2^p, y^2), \dots, (x_N^1, x_N^2, \dots, x_N^p, y^N)] \quad (1)$$

where  $x_a^b$  corresponds to the  $a$ th covariate from the  $b$ th datapoint. Note that our setup above is identical to our setup for multivariate linear regression (MLR).

The difference between MLR and logistic regression (LR) is that in LR each  $y_i$  is either the value 0 or 1.

### 1.2 A model for $Y$

Lets start to model  $Y$  by first searching for a random variable that generates the value 0 or 1. We know that a Bernoulli distributed random variable with parameter  $\theta$  will generate the value 1 with probability  $\theta$  and the value 0 with probability  $1 - \theta$ .

That is, if  $Y$  is a Bernoulli distributed random variable then

$$Y \sim \text{Bern}(\theta) \quad (2)$$

$$p(Y = 1) = \theta \quad (3)$$

$$p(Y = 0) = 1 - \theta \quad (4)$$

$$p(Y = y) = \theta^y (1 - \theta)^{1-y} \quad (5)$$

The expected value of  $Y$  is  $\mathbb{E}(Y) = \theta$  and the variance of  $Y$  is  $\mathbb{V}(Y) = \theta(1 - \theta)$ .

When we explore MLR, we started by assuming our  $Y$  had a normal distribution  $\mathcal{N}(\mu, \sigma^2)$  and then modified  $\mu$  so that it was a function that depended on parameters  $\beta$  and  $x$  data. We chose to modify  $\mu$  because the expected value of  $Y$  was  $\mu$ .

Let us take the same approach with our  $Y$  above and model the conditional distribution of  $Y$  given parameters  $\beta$  and  $x$  data as

$$Y_i | \beta_0, \beta_1, x_i \sim \text{Bern}(\theta(x)) \quad (6)$$

$$\theta(x) = \beta_0 + \beta_1 x \quad (7)$$

$$(8)$$

But we have a problem. The parameter  $\theta$  is constrained to be a value between 0 and 1, yet the quantity  $\beta_0 + \beta_1 x$  can take any value from negative to positive infinity. We need a way to constrain our values for  $\theta$  to be between 0 and 1.

### 1.2.1 The logistic function

One method to constrain  $\theta$  to be between 0 and 1 is to use the logistic function. The logistic function  $f$  is

$$f(x) = \frac{e^x}{1 + e^x} \quad (9)$$

For  $x$  values that approach positive infinity the logistic function approaches the value 1. For  $x$  values that approach negative infinity the logistic function approaches the value 0.

```
[8]: def logistic(x):
      import numpy as np
      e = np.exp(x)
      return e/(1+e)

fig, ax = plt.subplots()
domain = np.linspace(-5, 5, 10**3)

ax.plot(domain, logistic(domain))

ax.set_xlabel("x")
```

```

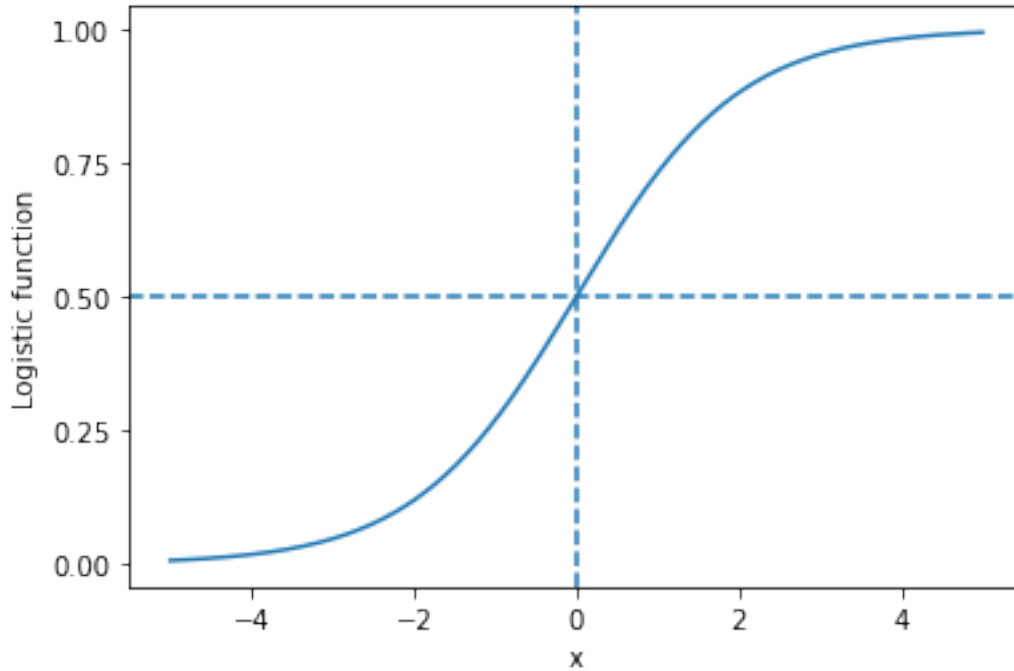
ax.set_ylabel("Logistic function")

ax.axvline(0,ls="--")
ax.axhline(0.5,ls="--")

ax.set_yticks([0,0.25,0.50,0.75,1.0])

plt.show()

```



Because logistic function maps values from negative infinity to positive infinity to numbers between 0 and 1, we can use this function to give us valid  $\theta$  values.

$$Y_i | \beta_0, \beta_1, x_i \sim \text{Bern}(\theta(x)) \quad (10)$$

$$\theta(x) = L(\beta_0 + \beta_1 x) \quad (11)$$

$$(12)$$

where  $L$  is the logistic function. In other words,

$$Y_i | \beta_0, \beta_1, x_i \sim \text{Bern}(\theta(x)) \quad (13)$$

$$\theta(x) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}} \quad (14)$$

$$(15)$$

### 1.3 A way to write logistic regression that looks more familiar

At times it can be convenient to rewrite our logistic regression model above to look more similar to the more familiar multivariate linear regression. For multivariate linear regression, the expected value of our target variable was equal to  $\beta_0 + \beta_1 x$ .

or logistic regression our expected value is

$$\mathbb{E}(Y) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}} \quad (16)$$

Let's rearrange the above equation to isolate  $\beta_0 + \beta_1 x$ .

$$\theta = \mathbb{E}(Y) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}} \quad (17)$$

$$1 + e^{\beta_0 + \beta_1 x}(\theta) = e^{\beta_0 + \beta_1 x} \quad (18)$$

$$\theta + \theta e^{\beta_0 + \beta_1 x} = e^{\beta_0 + \beta_1 x} \quad (19)$$

$$\theta = e^{\beta_0 + \beta_1 x} - \theta e^{\beta_0 + \beta_1 x} \quad (20)$$

$$\theta = e^{\beta_0 + \beta_1 x}(1 - \theta) \quad (21)$$

$$\frac{\theta}{1 - \theta} = e^{\beta_0 + \beta_1 x} \quad (22)$$

$$\log\left(\frac{\theta}{1 - \theta}\right) = \beta_0 + \beta_1 x \quad (23)$$

The ratio  $\frac{\theta}{1 - \theta}$  is called the **odds** and so the quantity  $\log\left(\frac{\theta}{1 - \theta}\right)$  is called the **log odds**.

### 1.4 A 1-unit change in X

Just like in MLR, we can look at how a one unit change in an x covariate will change our expected value. Let's take the difference between

$$\log\left(\frac{\theta}{1 - \theta}\right) = \beta_0 + \beta_1 x \quad (24)$$

and

$$\log\left(\frac{\theta^*}{1 - \theta^*}\right) = \beta_0 + \beta_1(x + 1) \quad (25)$$

$$\log\left(\frac{\theta^*}{1-\theta^*}\right) - \log\left(\frac{\theta}{1-\theta}\right) = \beta_0 + \beta_1(x+1) - (\beta_0 + \beta_1x) \quad (26)$$

$$\log\left(\frac{\theta^*}{1-\theta^*}\right) - \log\left(\frac{\theta}{1-\theta}\right) = \beta_0 + \beta_1(x+1) - \beta_0 - \beta_1x \quad (27)$$

$$\log\left(\frac{\theta^*}{1-\theta^*}\right) - \log\left(\frac{\theta}{1-\theta}\right) = \beta_1 \quad (28)$$

$$(29)$$

We can simplify the left hand side using the logarithm rule

$$\log\left(\frac{x}{y}\right) = \log(x) - \log(y) \quad (30)$$

$$\log\left(\frac{\theta^*}{1-\theta^*}\right) - \log\left(\frac{\theta}{1-\theta}\right) = \beta_1 \quad (31)$$

$$\log\left(\frac{\frac{\theta^*}{1-\theta^*}}{\frac{\theta}{1-\theta}}\right) = \beta_1 \quad (32)$$

$$(33)$$

The expression inside the logarithm on the left hand side of the equals is called the odds ratio. The logarithm of the odds ratio is often called the log odds ratio. To be clear, the log odds ratio is related to a one unit change in x.

We can also estimate the odds ratio by exponentiating both sides of the above equation

$$\log\left(\frac{\frac{\theta^*}{1-\theta^*}}{\frac{\theta}{1-\theta}}\right) = \beta_1 \quad (34)$$

$$\exp\left(\log\left(\frac{\frac{\theta^*}{1-\theta^*}}{\frac{\theta}{1-\theta}}\right)\right) = e^{\beta_1} \quad (35)$$

$$\frac{\frac{\theta^*}{1-\theta^*}}{\frac{\theta}{1-\theta}} = e^{\beta_1} \quad (36)$$

### 1.4.1 An example

```
[148]: import statsmodels.api as sm
import pandas as pd

d = pd.read_csv("diabetes.csv")
d
```

```

X = d.drop(columns=["Outcome"])
X = sm.add_constant(X)

averages = X.mean()
print(averages)

model = sm.Logit( d.Outcome, X)
model = model.fit()

print(model.summary())

def fromLogOdds2prob(vs,var,model):
    def invlogit(x):
        import numpy as np
        e = np.exp(x)
        return e/(1+e)

    probs = []
    for v in vs:
        x = averages
        x["{:s}".format(var)] = v

        logodds = model.params.dot(x)
        p = invlogit(logodds)
        probs.append(p)
    return probs

glucoseLevels = np.arange(40,210)
probsGlucose = fromLogOdds2prob( glucoseLevels, "Glucose", model )

pregs = np.arange(0,20,1)
probsPregs = fromLogOdds2prob( pregs, "Pregnancies", model )

dpfs = np.linspace(-12,4,100)
probsdps = fromLogOdds2prob( dpfs, "DiabetesPedigreeFunction", model )

fig,axs = plt.subplots(1,3)

axs[0].plot(glucoseLevels, probsGlucose )
axs[1].plot(pregs, probsPregs )
axs[2].plot(dpfs, probsdps )

fig.set_size_inches(12,5)
fig.set_tight_layout(True)
plt.show()

```

```
predictions = model.predict(X)
```

/usr/local/lib/python3.9/site-packages/statsmodels/tsa/tsatools.py:142:

FutureWarning: In a future version of pandas all arguments of concat except for the argument 'objs' will be keyword-only

```
x = pd.concat(x[:, :order], 1)
```

```
const                1.000000
Pregnancies          3.845052
Glucose              120.894531
BloodPressure        69.105469
SkinThickness        20.536458
Insulin              79.799479
BMI                  31.992578
DiabetesPedigreeFunction 0.471876
Age                  33.240885
```

dtype: float64

Optimization terminated successfully.

Current function value: 0.470993

Iterations 6

#### Logit Regression Results

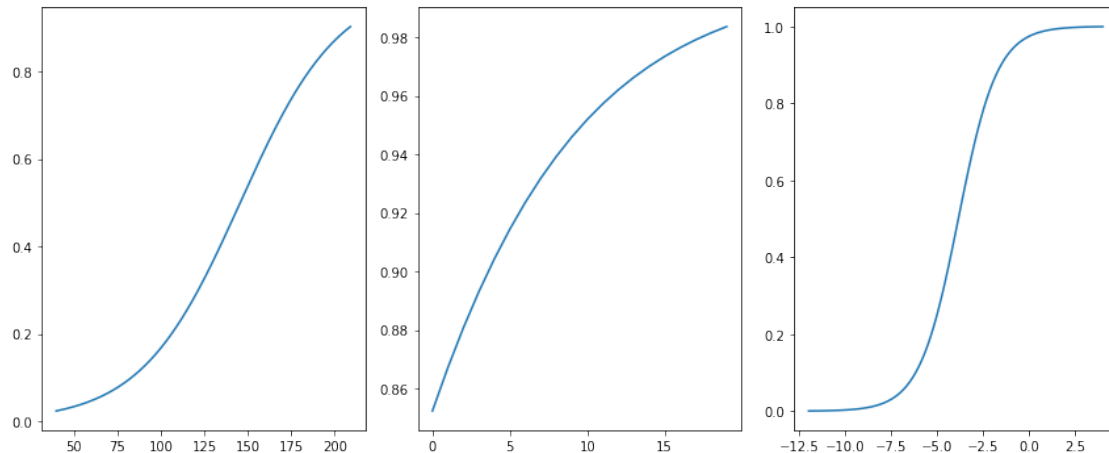
```
=====
Dep. Variable:                Outcome    No. Observations:                768
Model:                        Logit      Df Residuals:                759
Method:                        MLE       Df Model:                  8
Date:                          Wed, 10 Nov 2021    Pseudo R-squ.:                0.2718
Time:                          22:56:30    Log-Likelihood:               -361.72
converged:                      True    LL-Null:                   -496.74
Covariance Type:                nonrobust    LLR p-value:                9.652e-54
=====
```

```
=====
                                coef    std err          z      P>|z|      [0.025
0.975]
-----
const                -8.4047      0.717    -11.728    0.000    -9.809
-7.000
Pregnancies           0.1232      0.032     3.840    0.000     0.060
0.186
Glucose               0.0352      0.004     9.481    0.000     0.028
0.042
BloodPressure        -0.0133      0.005     -2.540    0.011    -0.024
-0.003
SkinThickness         0.0006      0.007     0.090    0.929    -0.013
0.014
Insulin              -0.0012      0.001     -1.322    0.186    -0.003
0.001
BMI                   0.0897      0.015     5.945    0.000     0.060
=====
```

```

0.119
DiabetesPedigreeFunction    0.9452    0.299    3.160    0.002    0.359
1.531
Age                        0.0149    0.009    1.593    0.111    -0.003
0.033
=====
=====

```



```

[146]: model = sm.Logit( d.Outcome, sm.add_constant(X.loc[:,["Glucose","BMI"]]) )
model = model.fit()
print(model.summary())

fig,ax = plt.subplots()
ax.scatter( X.Glucose, X.BMI, color=["red" if _ else "blue" for _ in d.Outcome.
→values] )

def f(x):
    return (7.51-x*0.035)/(0.0763)
ax.plot([40,200],[ f(40), f(200) ], color="black")

ax.set_xlim(40,210)
ax.set_ylim(10,70)

ax.set_xlabel("Glucose")
ax.set_ylabel("BMI")

plt.show()

```

```

Optimization terminated successfully.
Current function value: 0.502215
Iterations 6

```



# Logit Regression Results

```

=====
Dep. Variable:          Outcome    No. Observations:          768
Model:                  Logit      Df Residuals:              765
Method:                 MLE        Df Model:                  2
Date:                   Wed, 10 Nov 2021    Pseudo R-squ.:          0.2235
Time:                   22:56:02    Log-Likelihood:         -385.70
converged:              True        LL-Null:                 -496.74
Covariance Type:        nonrobust    LLR p-value:            5.967e-49
=====

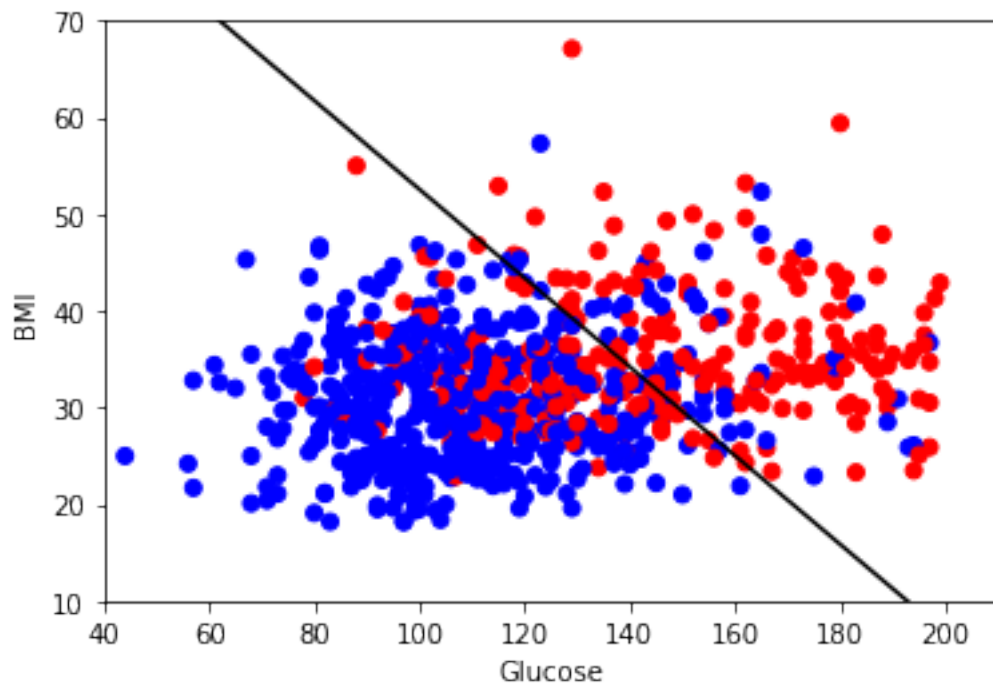
```

	coef	std err	z	P> z	[0.025	0.975]
const	-7.5156	0.605	-12.416	0.000	-8.702	-6.329
Glucose	0.0352	0.003	10.693	0.000	0.029	0.042
BMI	0.0763	0.013	5.723	0.000	0.050	0.102

```

/usr/local/lib/python3.9/site-packages/statsmodels/tsa/tsatools.py:142:
FutureWarning: In a future version of pandas all arguments of concat except for
the argument 'objs' will be keyword-only
  x = pd.concat(x[:, :order], 1)

```



**1.4.2 Interpreting  $\beta$**

**1.4.3 Interpreting the standard deviation**

**1.4.4 Interpreting the confidence intervals**

**1.4.5 Interpreting the pvalue**

**1.4.6 Decision boundary**

We can use logistic regression