Assignment 3 STAT321

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Problem 1

a) Compute probability that no one enters the store from 12:00 to 12:03.

Let $\lambda = 1$ be the typical number of people to enter the store every three minutes and let X be the discrete random variable for the number of people to enter the store from 12:00 to 12:03. Then we have that

$$X \sim Poisson(\lambda)$$
.

So we can compute P(X=0) directly as

$$P(X = 0) = \frac{\lambda^0 e^{-\lambda}}{0!}$$

= e^{-1}
= 0.3678794412

Solution: P(X = 0) = 0.3678794412.

b) Compute probability that no one enters the store from 12:00 to 12:15

Let λ_r be the typical number of people to enter the store every fifteen minutes, and let Y be the discrete random variable for the number of people to enter the store from 12:00 to 12:15. Then we have that

$$\lambda_r = 5 \times \lambda = 5(1) = 5$$

and

$$Y \sim Poisson(\lambda_r = 5).$$

So we compute P(Y=0) directly as

$$P(Y = 0) = \frac{\lambda_r^0 e^{-\lambda_r}}{0!}$$

$$= e^{-5}$$

$$= 0.006737946999$$

Solution: P(Y = 0) = 0.006737946999

c) Determine the probability that ≥ 3 people enter the store from 12 : 00 to 12 : 15 given that > 1 person enter the store during the same interval.

Let Y and λ_r be as defined in 1.b. Then we have that

$$P(Y \ge 3|Y > 1) = \frac{P(Y \ge 3 \cap Y > 1)}{P(Y > 1)}.$$

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Now, by the law of total probability we can see that

$$P(Y \ge 3) = P(Y \ge 3 \cap Y > 1) + P(Y \ge 3 \cap Y \le 1).$$

However, $Y \leq 1$ is mutually exclusive with $Y \geq 3$, so $P(Y \geq 3 \cap Y \leq 1) = 0$, so we have that

$$P(Y \ge 3) = P(Y \ge 3 \cap Y > 1).$$

Which now allows us to compute $P(Y \ge 3|Y > 1)$ directly.

$$\begin{split} P(Y \ge 3|Y > 1) &= \frac{P(Y \ge 3 \cap Y > 1)}{P(Y > 1)} \\ &= \frac{P(Y \ge 3)}{P(Y > 1)} \\ &= \frac{1 - P(Y \le 2)}{1 - P(\le 1)} \\ &= \frac{1 - 0.124652}{1 - 0.04042768} \quad \text{(Computed using the R code found below)} \\ &= \frac{0.875348}{0.95957232} \\ &= 0.9122272306 \end{split}$$

With $P(Y \le 2)$ and $P(Y \le 1)$ computed as P₋₁ and P₋₂ respectively in the following R code:

Solution: $P(Y \ge 3|Y > 1) = 0.9122272306$.

Problem 2

a) Create a probability distribution table for X.

Let r=2 be the number of red balls, n=2 be the number of balls drawn, and N=9 be the number of balls in the urn. Then we have that

$$X \sim hypergeometric(2, 9, 2).$$

Furthermore, we have that

$$P(X = 0) = \frac{\binom{2}{0}\binom{7}{2}}{\binom{9}{2}} = \frac{21}{36} = \frac{7}{12}$$

$$P(X = 1) = \frac{\binom{2}{1}\binom{7}{1}}{\binom{9}{2}} = \frac{14}{36} = \frac{7}{18}$$

$$P(X = 2) = \frac{\binom{2}{2}\binom{7}{0}}{\binom{9}{2}} = \frac{1}{36}$$

Solution: So a probability distribution table for X could be

$$\begin{array}{c|c|c} x & 0 & 1 & 2 \\ \hline P(X=x) & \frac{7}{12} & \frac{7}{18} & \frac{1}{36} \end{array}$$

b) Compute $E_{[X]}$.

Since $X \sim hypergeometric(2,9,2)$, we can compute $E_{[X]}$ directly as

$$E_{[X]} = \frac{(2)(2)}{9} = \frac{4}{9}$$

Solution: one would expect to draw 0.44444... red balls from the urn on average (under the conditions specified).

Problem 3

a)Compute $P(X \leq \frac{1}{2})$.

$$\int_0^{\frac{1}{2}} f(x)dx = \int_0^{\frac{1}{2}} xdx$$
$$= \frac{x^2}{2} \Big|_0^{\frac{1}{2}} = \frac{1}{8}.$$

Solution: $P(X \le \frac{1}{2}) = \frac{1}{8}$.

b) Compute $P(0 \le X \le 2)$.

$$\int_{0}^{2} f(x)dx = \int_{0}^{1} f(x)dx + \int_{1}^{2} f(x)dx$$

$$= \int_{0}^{1} xdx + \int_{1}^{2} \frac{1}{4}dx$$

$$= \frac{x^{2}}{2} \Big|_{0}^{1} + \frac{x}{4} \Big|_{1}^{2}$$

$$= \frac{1}{2} + \frac{1}{2} - \frac{1}{4}$$

$$= \frac{3}{4}$$

Solution: $P(0 \le X \le 2) = \frac{3}{4}$.

Problem 4

a) Find c such that f(x) is a probability density function.

We require that

$$\int_0^2 f(x)dx = 1.$$

This allows us to compute c directly by

$$1 = \int_0^2 f(x)dx = \int_0^2 c(2x - x^2)dx$$
$$= c \left[\int_0^2 2xdx - \int_0^2 x^2dx \right]$$
$$= c \left[x^2 \Big|_0^2 - \frac{x^3}{3} \Big|_0^2 \right]$$

$$= c \left[4 - \frac{8}{3} \right]$$
$$\frac{4c}{3} = 1 \Rightarrow c = \frac{3}{4}.$$

Solution: For f(x) to be a valid probability density function, we require that $c = \frac{3}{4}$.

Problem 5

a) Compute probability that the region will see exactly 10 hurricanes over 2 years.

Let X be the discrete random variable for the number of hurricanes the region experiences in 2 years and $\lambda = 4.04$ be the number of hurricanes the region typically experiences per year. Finally, let λ_r be the typical number of hurricanes over a 2 year period. Then we have that

$$\lambda_r = 2 \times \lambda = 8.08$$

and

$$X \sim Poisson(\lambda_r = 8.08).$$

Therefore, we compute the desired probability directly as

$$P(X = 10) = \frac{\lambda_r^{1} 0^{-\lambda_r}}{e} 10!$$
$$= \frac{8.08^{10} e^{-8.08}}{10!}$$
$$= 0.101216464.$$

Solution: P(X = 10) = 0.101216464.

b) Compute $E_{[X]}$.

With X and λ_r defined as in 5.a, we have that since $X \sim Poisson(\lambda_r)$, we have

$$E_{[X]} = \lambda_r \qquad = 8.08$$

Solution: we'd expect the region to experience 8.08 hurricanes over a 2 year period.

Problem 6

a) Show that f(x) is not a valid probability density function.

$$\int_{1}^{2} f(x)dx = \int_{1}^{2} 2x - 1dx$$
$$= \left[x^{2} - x\right] \Big|_{1}^{2}$$
$$= 2 - 0 = 2$$

Solution: since the area under f(x) is 2 > 1 over the real numbers, f(x) is not a probability density function.

Problem 7

a) Compute the probability that no defective screws appear in the display bin.

Let N = 500 be the total number of screws in the case, r = 11 be the known number of defective screws and n = 125 be the number selected from N without replacement for display. Then let X be the random variable for the number of defective screws in the display bin. We have that

$$X \sim hypergeometric(11, 500, 125).$$

So we can compute P(X=0) as

$$P(X = 0) = \frac{\binom{11}{0}\binom{489}{125}}{\binom{500}{125}}$$
$$= \frac{2.155861 \times 10^{119}}{5.298292 \times 10^{120}}$$
$$= 0.04068972683$$

Solution: P(X = 0) = 0.04068972683.

b) Compute $P(X \ge 2)$.

Let N, n, r and X be as defined in 7.a. Then we seek to compute directly $P(X \ge 2)$ as

$$P(X \ge 2) = 1 - P(X \le 1)$$

$$= 1 - (P(X = 0) + P(X = 1))$$

$$= 1 - \left(0.04068972683 + \frac{\binom{11}{1}\binom{489}{124}}{\binom{500}{125}}\right) \quad \text{(from 7.a)}$$

$$= 1 - (0.04068972683 + 0.15328321751)$$

$$= 1 - (0.1939729443)$$

$$= 0.8060270557$$

Solution: $P(X \ge 2) = 0.8060270557$.

c) Compute P(X = 2) + P(X = 3).

Let N, n, r and X be as defined in 7.a. Then we compute P(X=2) + P(X=3) by first computing each probability separately, and then adding the results.

$$P(X = 2) = \frac{\binom{11}{2}\binom{489}{123}}{\binom{500}{125}}$$

$$= 0.25966009524$$

$$P(X = 3) = \frac{\binom{11}{3}\binom{489}{122}}{\binom{500}{125}}$$

$$= 0.26107513663$$

$$P(X = 2) + P(X = 3) = 0.25966009524 + 0.26107513663 = 0.5207352319$$

Solution: P(X = 2) + P(X = 3) = 0.5207352319.

Problem 8

a) Show that $E_{[X]} = \lambda$ for $X \sim Poisson(\lambda)$ using the provided $M_X(t)$ for the Poisson distribution.

We have that

$$M_X(t) = e^{\lambda(e^t - 1)}$$

And we know that $E_{[X]} = \frac{d}{dt} M_X(t) \Big|_{t=0}$, so we proceed to differentiate $M_X(t)$ with respect to t.

$$E_{[X]} = \frac{d}{dt} M_X(t)$$

$$= \frac{d}{dt} e^{\lambda(e^t - 1)}$$

$$= \lambda e^t e^{\lambda(e^t - 1)}$$

$$= \lambda e^{\lambda e^t - \lambda + t}$$

Then we simply evaluate the equation at t = 0

$$E_{[X]} = \lambda e^{\lambda e^t - \lambda + t} \bigg|_{t=0} = \lambda e^{\lambda - \lambda} = \lambda e^0 = \lambda.$$

Solution: thus, by the above we have verifed that $E_{[X]} = \lambda$ for $X \sim Poisson(\lambda)$.