

# Reversibility and Time Reversals

Detailed Balance Equations:

A distribution  $\pi$  over  $S$  satisfies the detailed balance equations if:

$$\forall x, y \in S \quad \pi(x) P(x, y) = \pi(y) P(y, x).$$

Theorem: Any distribution  $\pi$  that satisfies the detailed balance equations for a Markov chain with transition matrix  $P$ , is a stationary distribution for  $P$ .

$$\text{Detailed Balance Equations} \Rightarrow \pi P = \pi$$

Proof:

$$\pi P(i) = \pi P^{\overset{\text{ith column}}{i}} = \sum_j \pi(j) P(j, i) \stackrel{\text{DBE}}{=} \sum_j \pi(j) P(i, j) = \pi(i) \quad \square$$

→ If  $\pi$  satisfies DBE for  $P$ , then seeing a  $x \rightarrow y$  transition is as likely as seeing a  $y \rightarrow x$  transition, if we start from  $\pi$ .

→ Even more, if a chain  $(X_t)$  has the initial distribution  $\pi$ , then  $(X_0, \dots, X_n)$  has the same distribution as  $(X_n, \dots, X_0)$ . More concretely,

$$\forall x_0, x_1, \dots, x_n, \quad \Pr_{\pi}[X_0 = x_0, \dots, X_n = x_n] = \Pr_{\pi}[X_0 = x_n, \dots, X_n = x_0].$$

→ This means that we can't distinguish the chain going forward or backward in time! Thus, such a chain is called reversible.

①P,

# Time Reversal

Definition: The time reversal of an irreducible Markov chain with transition matrix  $P$  and stationary distribution  $\pi$  is the chain with matrix:

$$\hat{P}(x, y) := \frac{\pi(y)}{\pi(x)} P(y, x)$$

Theorem: Let  $(X_t)_t$  be an irreducible chain with transition matrix  $P$  and stationary distribution  $\pi$ . Write  $(\hat{X}_t)_t$  for the time-reversed chain with transition matrix  $\hat{P}$ . Then a)  $\pi$  is stationary for  $\hat{P}$  and b) for any  $x_0, \dots, x_t \in S$ ,

$$\Pr_{\pi}[X_0 = x_0, \dots, X_t = x_t] = \Pr_{\pi}[\hat{X}_0 = x_t, \dots, \hat{X}_t = x_0].$$

Proof:

$$a) \pi \hat{P}(x) = \sum_y \pi(y) \hat{P}(y, x) = \sum_y \pi(y) \frac{\pi(x)}{\pi(y)} P(x, y) = \pi(x)$$

b) To show that the trajectory probabilities are equal, note that

$$\begin{aligned} \Pr_{\pi}[X_0 = x_0, \dots, X_t = x_t] &= \pi(x_0) P(x_0, x_1) \dots P(x_{t-1}, x_t) \\ &= \pi(x_t) \hat{P}(x_t, x_{t-1}) \dots \hat{P}(x_1, x_0) = \Pr_{\pi}[\hat{X}_0 = x_t, \dots, \hat{X}_t = x_0] \end{aligned}$$

(Because  $\pi(x_i) P(x_i, x_{i+1}) = \pi(x_{i+1}) \hat{P}(x_{i+1}, x_i)$ )

□

→ Observe that if a chain is reversible (i.e., satisfies DBE)

then  $P = \hat{P}$ .

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# Markov Chain Monte Carlo

Let's say we have a distribution  $\pi$  over a set  $S$  that we would like to sample from. One way of doing this is to create a Markov chain,  $P$ , over  $S$  with stationary distribution  $\pi$ . If this chain converges to its stationary distribution, then after running the chain for some initial steps, it will have distribution  $\sim \pi$ .

## Metropolis Chains:

Suppose  $Q$  is a symmetric, irreducible transition matrix over the set of states  $S$ . The uniform distribution satisfies the detailed balance equations for  $Q$ :

$$\pi_Q \sim \text{Uniform}(S) \Rightarrow \forall x, y \in S \quad \pi_Q(x)Q(x, y) = \pi_Q(y)Q(y, x).$$

Thus, the uniform distribution is the unique stationary distribution of  $Q$ .

The Metropolis algorithm suggests a method for "modifying" the dynamics of  $Q$  such that the stationary distribution becomes  $\pi$ .

This is how the modified dynamics are:

→ At each state  $x$ , generate a candidate next state  $y \sim Q(x, \cdot)$

→ Accept the new candidate state with probability  $a(x, y)$ , o.w. stay in  $x$ .

This new chain has the following transition dynamics  $P$ :

$$P(x, y) = \begin{cases} Q(x, y) a(x, y), & y \neq x \\ 1 - \sum_{z \neq x} Q(x, z) a(x, z), & y = x \end{cases}$$

## Metropolis Chains

We would like  $a(x,y)$ 's be such that the stationary distribution of  $P$  be  $\pi$ . So we find them so that  $\pi$  satisfies the detailed balance equations:

$$\begin{aligned}\pi(x)P(x,y) &= \pi(y)P(y,x) \Rightarrow \pi(x)Q(x,y)a(x,y) \\ &= \pi(y)Q(y,x)a(y,x)\end{aligned}$$

$$\Rightarrow \underbrace{\pi(x)a(x,y)}_{\alpha} = \underbrace{\pi(y)a(y,x)}_{\beta}$$

Because  $a(\cdot, \cdot)$  is a probability, we must have  $\alpha = \beta \leq \pi(x), \pi(y)$ .

Thus,  $\alpha \leq \min\{\pi(x), \pi(y)\}$ . We would like to reject transitions as hardly as possible; so we maximize the acceptance probability by setting:

$$\alpha = \min\{\pi(x), \pi(y)\} \Rightarrow \pi(x)a(x,y) = \min\{\pi(x), \pi(y)\}$$

$$\Rightarrow a(x,y) = \min\left\{1, \frac{\pi(y)}{\pi(x)}\right\}$$

So the new dynamics is as follows:

$$P(x,y) = \begin{cases} Q(x,y) \min\left\{1, \frac{\pi(y)}{\pi(x)}\right\}, & \text{if } x \neq y \\ 1 - \sum_{z \neq x} Q(x,z) \min\left\{1, \frac{\pi(z)}{\pi(x)}\right\}, & \text{if } x = y \end{cases}$$

→ Basically, if  $\pi(y) \geq \pi(x)$  we always accept the transition from  $x$  to  $y$ .  
However, if  $\pi(y) < \pi(x)$  we reject the transition with probability  $1 - \frac{\pi(y)}{\pi(x)}$ .



## Metropolis - Hastings Chains

Hastings extended Metropolis chains so that the initial transition dynamics

(Q) need not be symmetric, only irreducible.

The general idea is the same: we sample a transition from Q and accept it with probability  $\alpha(x, y)$ . However, here for the new chain to be time-reversible with stationary distribution  $\pi$ ,  $\alpha(x, y)$  must be different.

Writing the detailed balance equations for the new chain, P, and  $\pi$  we get:

$$\begin{aligned}\pi(x) P(x, y) &= \pi(y) P(y, x) \Rightarrow \pi(x) Q(x, y) \alpha(x, y) \\ &= \pi(y) Q(y, x) \alpha(y, x)\end{aligned}$$

$$\Rightarrow \alpha(x, y) \leq \frac{\pi(y) Q(y, x)}{\pi(x) Q(x, y)}$$

This gives rise to the following chain:

$$P(x, y) = \begin{cases} Q(x, y) \min \left\{ \frac{\pi(y) Q(y, x)}{\pi(x) Q(x, y)}, 1 \right\} & \text{if } y \neq x \\ 1 - \sum_{z \neq x} Q(x, z) \min \left\{ \frac{\pi(z) Q(z, x)}{\pi(x) Q(x, z)}, 1 \right\} & \end{cases}$$

# Glauber Dynamics

Often, we study chains whose state space is contained in a set of the form  $S^V$  where  $S$  is a finite label set and  $V$  is the vertex set of a graph. We can think of  $S^V$  as the set of labelings of the vertices of the graph with  $S$ . Elements of  $S^V$  are called configurations.

## Definition: [Glauber Dynamics]

Given a distribution  $\pi$  on a space of configurations, the Glauber dynamics for  $\pi$  is a Markov chain with stationary distribution  $\pi$ .

In statistics, it is often referred to as the Gibbs sampler.

## Intuition:

At a high level, the Glauber chain moves as follows:

when at a state  $x$ , select a vertex  $v$  uniformly from  $V$ . The new state is chosen according to the measure  $\pi$ , conditioned on the set of states that are equal to  $x$  except possibly at  $v$ .

# Glauber Dynamics (Cont.)

## Definition: [Glauber Dynamics]

Let  $S$  and  $V$  be finite sets and let  $X \subseteq S^V$ . Let  $\pi$  be a probability distribution over  $X$ . The Glauber dynamics for  $\pi$  is a reversible Markov chain with state space  $X$ , stationary distribution  $\pi$  and transition probabilities as described below:

For a configuration  $x \in X$  and vertex  $v \in V$ , let

$$X(x, v) = \{y \in X : y(w) = x(w) \text{ for all } w \neq v\}$$

be the set of states agreeing with  $x$  everywhere except possibly at  $v$ .

Define

$$\pi^{x,v}(y) = \pi(y | X(x, v)) = \begin{cases} \frac{\pi(y)}{\pi(X(x, v))} & \text{if } y \in X(x, v) \\ 0 & \text{if } y \notin X(x, v) \end{cases}$$

to be the distribution of  $\pi$  conditioned on  $X(x, v)$ . The rule for updating a configuration is as follows: pick a vertex  $v$  uniformly from  $V$ , and choose a new configuration according to  $\pi^{x,v}$ .

→  $\pi$  is always stationary and reversible for Glauber dynamics.

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