

TA Session 4  
Dynamic Discrete Choice: CCP  
Microeometrics II with Joan Llull  
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# Using CCP to Represent Differences in Continuation Values

$$\begin{aligned}
 v_j(x_t) &= u_j(x_t) + \beta EV(x_t) \\
 \Rightarrow v_0(x_t) - v_1(x_t) &= u_0(x_t) - u_1(0) + \beta EV(x_t) - \beta EV(0) \\
 &= u_0(x_t) - u_1(0) \\
 &\quad + \beta \ln [e^{v_0(x_{t+1})} + e^{v_1(x_{t+1})}] - \beta \ln [e^{v_0(1)} + e^{v_1(1)}] \\
 &= u_0(x_t) - u_1(0) + \beta \ln \frac{e^{v_0(x_{t+1})} + e^{v_1(x_{t+1})}}{e^{v_0(1)} + e^{v_1(1)}} \\
 &\stackrel{(*)}{=} u_0(x_t) - u_1(0) + \beta \ln \frac{e^{v_0(x_{t+1})} - v_1(x_{t+1}) + 1}{e^{v_0(1)} - v_1(1) + 1} \\
 &= u_0(x_t) - u_1(0) + \beta \ln \frac{p_1(1)}{p_1(x_{t+1})}
 \end{aligned}$$

where  $(*)$  is implied by the renewal property  $v_1(x_{t+1}) = v_1(1)$ , and  $d_t = 1 \Rightarrow x_t = 0, x_{t+1} = (0, 1, 2) \cdot (\varphi_0, \varphi_1, \varphi_2)'$ .

# Invertible Mapping between CCP and Differences in Continuation Values

- Therefore,

$$p_1(x) = \frac{1}{1 + e^{v_0(x) - v_1(x)}} = \frac{1}{1 + e^{u_0(x_t) - u_1(0) + \beta \ln \frac{p_1(1)}{p_1(x_{t+1})}}}$$

- Intuitively, the CCP for the current replacement is the CCP for a static model with an offset term.
- The offset term accounts for differences in continuation values using future CCPs that characterize optimal future replacements.
- An important implication of the invertible mapping is that estimation can be done without computing the value functions.

# CCP Estimation

- CCP estimator in two stages:

- ① Estimate non-parametrically for  $p_1(x)$  and the transitions
  - Estimate  $p_1(x)$  from the relative frequencies:

$$\hat{p}_1(x) \equiv \frac{\sum_i \sum_t \mathbb{1}\{d_{it} = 1\} \cdot \mathbb{1}\{x_{it} = x\}}{\sum_i \sum_t \mathbb{1}\{x_{it} = x\}}$$

- ② Solve for utility parameters  $\theta$

- Substitute  $\hat{p}_1(x)$  into the likelihood as incidental parameters to estimate  $\theta^1$  with a logit:

$$\frac{d_1 + d_0 \cdot e^{u_0(x_t; \theta_0) - u_1(0; \theta_1) + \beta \ln \frac{\hat{p}_1(1)}{\hat{p}_1(x_{t+1})}}}{1 + e^{u_0(x_t; \theta_0) - u_1(0; \theta_1) + \beta \ln \frac{\hat{p}_1(1)}{\hat{p}_1(x_{t+1})}}}$$

where  $x_{t+1} = (0, 1, 2) \cdot (\varphi_0, \varphi_1, \varphi_2)'$  given  $d_t = 1$ .

- Correct the std. err. for  $\theta$  induced by the first stage estimates of  $\hat{p}_1(x)$ .

<sup>1</sup>The term  $\ln \frac{\hat{p}_1(1)}{\hat{p}_1(x_{t+1})}$  enters the logit as an individual specific component of the data, thus  $\beta$  enters the logit in the same way as the utility parameters  $\theta$ . You can also estimate  $\beta$  in this estimation.

# CCP Estimation vs. Full Solution Approach

- The full solution approach
  - The full solution approach computes the optimal decision rule, it's computationally intensive especially if you have a rich state space.
  - While it may be theoretically possible to parametrically identify dynamic models with data sets that only track the first few periods of the decision maker's problem.
- The CCP approach
  - Instead of computing continuation values by solving the dynamic problem for all elements in the state space, the CCP approach "measures" continuation values using a function of CCPs.
  - The CCP estimator is faster but less versatile. It requires samples to be drawn from the population of all possible histories.

## References

- Hotz, V. J., & Miller, R. A. (1993). Conditional choice probabilities and the estimation of dynamic models. *The Review of Economic Studies*, 60(3), 497-529.