Applications of spectral methods (ℓ_2 theory)



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What we have learned so far

- Classical ℓ_2 matrix perturbation theory:
 - Davis-Kahan's $\sin \Theta$ theorem
 - Wedin's $\sin \Theta$ theorem
 - Eigenvector perturbation of probability transition matrices

- Matrix concentration inequalities:
 - Matrix Bernstein inequality

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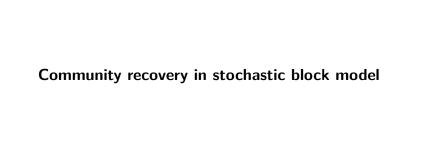
- Matrix concentration inequalities:
 - Matrix Bernstein inequality

— we will check their applications today

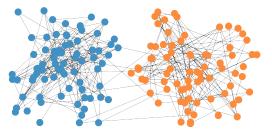
Outline

- Community recovery in stochastic block model
 - application of Davis-Kahan's theorem
- Low-rank matrix completion

- application of Wedin's theorem
- Ranking from pairwise comparisons
 - application of eigenvector perturbation of prob. transition matrix



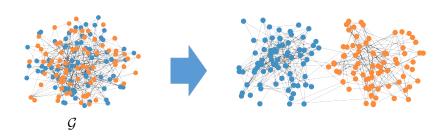
Stochastic block model (SBM)



$$x_i^{\star} = 1$$
: 1st community $x_i^{\star} = -1$: 2nd community

- n nodes $\{1,\ldots,n\}$
- 2 communities
- n unknown variables: $x_1^{\star}, \dots, x_n^{\star} \in \{1, -1\}$
 - encode community memberships

Stochastic block model (SBM)

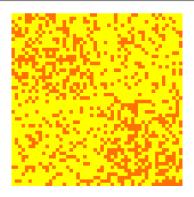


ullet observe a graph $\mathcal G$ $(i,j)\in \mathcal G \mbox{ with prob. } \begin{cases} p, & \mbox{if } i \mbox{ and } j \mbox{ are from same community} \\ q, & \mbox{else} \end{cases}$

Here, p > q

ullet Goal: recover community memberships of all nodes, i.e., $\{x_i^\star\}$

Adjacency matrix

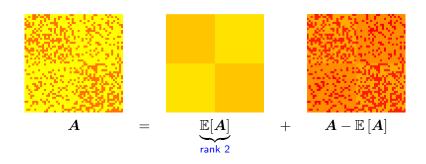


Consider the adjacency matrix $A \in \{0,1\}^{n \times n}$ of \mathcal{G} : (assume $A_{ii} = p$)

$$A_{i,j} = \begin{cases} 1, & \text{if } (i,j) \in \mathcal{G} \\ 0, & \text{else} \end{cases}$$

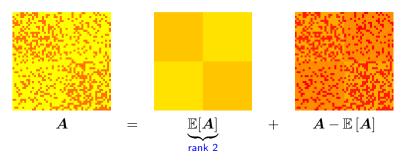
• WLOG, suppose $x_1^\star=\cdots=x_{n/2}^\star=1$; $x_{n/2+1}^\star=\cdots=x_n^\star=-1$

Adjacency matrix



$$\mathbb{E}[\boldsymbol{A}] = \begin{bmatrix} p \mathbf{1} \mathbf{1}^\top & q \mathbf{1} \mathbf{1}^\top \\ q \mathbf{1} \mathbf{1}^\top & p \mathbf{1} \mathbf{1}^\top \end{bmatrix} = \underbrace{\frac{p+q}{2}}_{\text{uninformative bias}} + \underbrace{\frac{p-q}{2}}_{=\boldsymbol{x}^\star = [x_i]_{1 \leq i \leq n}} [\mathbf{1}^\top, -\mathbf{1}^\top]$$

Spectral clustering



- 1. computing the leading eigenvector $m{u} = [u_i]_{1 \leq i \leq n}$ of $m{A} \frac{p+q}{2} \mathbf{1} \mathbf{1}^{ op}$
- 2. rounding: output $x_i = \begin{cases} 1, & \text{if } u_i > 0 \\ -1, & \text{if } u_i < 0 \end{cases}$

Apply Davis-Kahan's theorem

Consider

$$m{M}^\star \coloneqq \mathbb{E}[m{A}] - rac{p+q}{2} m{1} m{1}^ op = rac{p-q}{2} egin{bmatrix} m{1} \\ -m{1} \end{bmatrix} egin{bmatrix} m{1}^ op & -m{1}^ op \end{bmatrix},$$

which obevs

$$\lambda_1(\boldsymbol{M}^\star) \coloneqq rac{(p-q)n}{2}, \quad ext{and} \quad \boldsymbol{u}^\star \coloneqq rac{1}{\sqrt{n}} \left[egin{array}{c} \mathbf{1}_{n/2} \\ -\mathbf{1}_{n/2} \end{array}
ight].$$

Also, we have

$$oldsymbol{M}\coloneqq oldsymbol{A} - rac{p+q}{2} oldsymbol{1} oldsymbol{1}^ op$$

Then Davis-Kahan $\sin \Theta$ Theorem yields

$$\operatorname{dist}(\boldsymbol{u}, \boldsymbol{u}^{\star}) \leq \frac{\|\boldsymbol{M} - \boldsymbol{M}^{\star}\|}{\lambda_{1}(\boldsymbol{M}^{\star}) - \|\boldsymbol{M} - \boldsymbol{M}\|} = \frac{\|\boldsymbol{A} - \mathbb{E}[\boldsymbol{A}]\|}{\frac{(p-q)n}{2} - \|\boldsymbol{A} - \mathbb{E}[\boldsymbol{A}]\|}$$
(5.1)

as long as
$$\|oldsymbol{A} - \mathbb{E}[oldsymbol{A}]\| < \lambda_1(oldsymbol{M}^\star) = rac{(p-q)n}{2}$$

5 - 10

Bounding $\| \boldsymbol{A} - \mathbb{E}[\boldsymbol{A}] \|$

Matrix concentration inequalities tell us that

Lemma 5.1

Consider SBM with p>q and $p\gtrsim \frac{\log n}{n}$. Then with high prob.

$$\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| \lesssim \sqrt{np \log n}$$
 (5.2)

— better concentration yields \sqrt{np} bound

Statistical accuracy of spectral clustering

Substitute (5.2) into (5.1) to reach

$$\mathsf{dist}(\boldsymbol{u},\boldsymbol{u}^\star) \leq \frac{\|\boldsymbol{A} - \mathbb{E}[\boldsymbol{A}]\|}{\frac{(p-q)n}{2} - \|\boldsymbol{A} - \mathbb{E}[\boldsymbol{A}]\|} \lesssim \frac{\sqrt{np\log n}}{(p-q)n}$$

provided that $(p-q)n \gg \sqrt{np\log n}$

Thus, under condition $\frac{p-q}{\sqrt{p}}\gg\sqrt{\frac{\log n}{n}}$, with high prob. one has

$$\mathsf{dist}(oldsymbol{u},oldsymbol{u}^\star) \ll 1 \qquad \Longrightarrow \qquad \mathsf{nearly perfect clustering}$$

From estimation error to mis-clustering error

WLOG assume that $\|\boldsymbol{u}-\boldsymbol{u}^{\star}\|_{2}=\operatorname{dist}(\boldsymbol{u},\boldsymbol{u}^{\star})$. Consider the set

$$\mathcal{N} \coloneqq \{i \mid |u_i - u_i^{\star}| \ge 1/\sqrt{n}\}.$$

We claim that

$$\frac{1}{n} \sum_{i=1}^{n} \mathbb{1} \{ x_i \neq x_i^{\star} \} \le \frac{1}{n} \sum_{i=1}^{n} \mathbb{1} \{ |u_i - u_i^{\star}| \ge \frac{1}{\sqrt{n}} \} = \frac{|\mathcal{N}|}{n}$$

To see this, observe that for any i obeying $x_i \neq x_i^\star$, one has $\mathrm{sgn}(u_i) \neq \mathrm{sgn}(u_i^\star)$, thus indicating that $|u_i - u_i^\star| \geq |u_i^\star| = 1/\sqrt{n}$ In the end, we have

$$|\mathcal{N}| \le \frac{\|\boldsymbol{u} - \boldsymbol{u}^\star\|_2^2}{(1/\sqrt{n})^2} = o(n)$$

Statistical accuracy of spectral clustering

$$\frac{p-q}{\sqrt{p}}\gg\sqrt{\frac{\log n}{n}}\quad\Longrightarrow\quad \text{nearly perfect clustering}$$

• dense regime: if $p \approx q \approx 1$, then this condition reads

$$p - q \gg \sqrt{\frac{\log n}{n}}$$

• "sparse" regime: if $p=\frac{a\log n}{n}$ and $q=\frac{b\log n}{n}$ for $a,b\asymp 1$, then $a-b\gg \sqrt{a}$

This condition is information-theoretically optimal (up to log factor) — Mossel, Neeman, Sly '15, Abbe '18

Proof of Lemma 5.2

We write $A - \mathbb{E}[A]$ as sum of independent random matrices

$$oldsymbol{A} - \mathbb{E}[oldsymbol{A}] = \sum_{j>i} ig(A_{i,j} - \mathbb{E}[A_{i,j}]ig)ig(oldsymbol{e}_ioldsymbol{e}_j^ op + oldsymbol{e}_joldsymbol{e}_i^ op$$

We only need to consider $m{A}_{\mathsf{upper}} \coloneqq \sum_{i < j} \underbrace{(A_{i,j} - \mathbb{E}[A_{i,j}]) m{e}_i m{e}_j^{ op}}_{=: m{X}_{i,j}}$

- First, $||X_{i,j}|| \le 1 =: B$
- Since $\operatorname{Var}(A_{i,j}) \leq p$, one has $\mathbb{E}\left[\boldsymbol{X}_{i,j} \boldsymbol{X}_{i,j}^{\top} \right] \preceq p \boldsymbol{e}_i \boldsymbol{e}_i^{\top}$, which gives

$$\sum\nolimits_{i < j} \mathbb{E}\left[\boldsymbol{X}_{i,j} \boldsymbol{X}_{i,j}^{\top}\right] \preceq \sum\nolimits_{i < j} p\boldsymbol{e}_{i} \boldsymbol{e}_{i}^{\top} \preceq np \, \boldsymbol{I}_{n}$$

Similarly, $\sum_{i < j} \mathbb{E}\left[\boldsymbol{X}_{i,j}^{\top} \boldsymbol{X}_{i,j} \right] \preceq np \, \boldsymbol{I}_n$. As a result,

$$v \coloneqq \max \left\{ \left\| \sum_{i,j} \mathbb{E} \left[\boldsymbol{X}_{i,j} \boldsymbol{X}_{i,j}^{\top} \right] \right\|, \left\| \sum_{i,j} \mathbb{E} \left[\boldsymbol{X}_{i,j}^{\top} \boldsymbol{X}_{i,j} \right] \right\| \right\} \le np$$

Proof of Lemma 5.2 (cont.)

Take the matrix Bernstein inequality to conclude that with high prob.,

$$\begin{split} \| \boldsymbol{A} - \mathbb{E}[\boldsymbol{A}] \| \lesssim \sqrt{v \log n} + B \log n \lesssim \sqrt{n p \log n} \\ & - \text{since } p \gtrsim \frac{\log n}{n} \end{split}$$



Low-rank matrix completion

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figure credit: Candès

- ullet consider a low-rank matrix $M^\star = U^\star \Sigma^\star V^{\star op}$
- each entry $M_{i,j}^{\star}$ is observed independently with prob. p
- intermediate goal: estimate U^{\star}, V^{\star}

Spectral method for matrix completion

- 1. identify the key matrix M^{\star}
- 2. construct surrogate matrix $M \in \mathbb{R}^{n \times n}$ as

$$M_{i,j} = \begin{cases} \frac{1}{p} M_{i,j}^{\star}, & \text{if } M_{i,j}^{\star} \text{ is observed} \\ 0, & \text{else} \end{cases}$$

- \circ rationale for rescaling: ensures $\mathbb{E}[M] = M^\star$
- 3. compute the rank-r SVD $U\Sigma V^{ op}$ of M, and return (U,Σ,V)

Statistical accuracy of spectral estimate

Let's analyze a simple case where $oldsymbol{M}^\star = oldsymbol{u}^\star oldsymbol{v}^{\star op}$ with

$$oldsymbol{u}^\star = rac{1}{\| ilde{oldsymbol{u}}\|_2} ilde{oldsymbol{u}}, \quad oldsymbol{v}^\star = rac{1}{\| ilde{oldsymbol{v}}\|_2} ilde{oldsymbol{v}}, \quad ilde{oldsymbol{u}}, ilde{oldsymbol{v}} \sim \mathcal{N}(oldsymbol{0}, oldsymbol{I}_n)$$

From Wedin's Theorem: if $\| m{M} - m{M}^\star \| \leq \frac{1}{2} \sigma_1(m{M}^\star) = \frac{1}{2}$, then

$$\max \left\{ \mathsf{dist}(\boldsymbol{u}, \boldsymbol{u}^{\star}), \mathsf{dist}(\boldsymbol{v}, \boldsymbol{v}^{\star}) \right\} \lesssim \frac{\|\boldsymbol{M} - \boldsymbol{M}^{\star}\|}{\sigma_{1}(\boldsymbol{M}^{\star})} \asymp \|\boldsymbol{M} - \boldsymbol{M}^{\star}\| \quad (5.3)$$

Bounding $\| oldsymbol{M} - oldsymbol{M}^\star \|$

Matrix concentration inequalities tell us that

Lemma 5.2

Consider matrix completion with $p \gg \frac{\log^3 n}{n}$. Then with high prob.

$$\|\boldsymbol{M} - \boldsymbol{M}^{\star}\| \lesssim \sqrt{\frac{\log^3 n}{n}} = o(1)$$
 (5.4)

Sample complexity

For rank-1 matrix completion,

$$p \gg \frac{\log^3 n}{n} \qquad \Longrightarrow \qquad \text{nearly accurate estimates}$$

Sample complexity needed to yield reliable spectral estimates is

$$\underbrace{n^2p \asymp n\log^3 n}_{\text{optimal up to log factor}}$$

Proof of inequality (5.4)

Write $M-M^\star=\sum_{i,j} X_{i,j}$, where $X_{i,j}=(M_{i,j}-M^\star_{i,j})e_ie_j^{ op}$

• First, based on Gaussianity, we have

$$\|\boldsymbol{X}_{i,j}\| \leq \frac{1}{p} \max_{i,j} |M_{i,j}^{\star}| \lesssim \frac{\log n}{pn} := B \quad (\mathsf{check})$$

• Next, $\mathbb{E}[X_{i,j}X_{i,j}^{\top}] = \mathsf{Var}(M_{i,j})e_ie_i^{\top}$ and hence

$$\mathbb{E}\big[\sum\nolimits_{i,j}\boldsymbol{X}_{i,j}\boldsymbol{X}_{i,j}^{\top}\big] \preceq \Big\{\max_{i,j} \mathsf{Var}\big(M_{i,j}\big)\Big\} n\boldsymbol{I} \preceq \Big\{\frac{n}{p}\max_{i,j}(M_{i,j}^{\star})^2\Big\} \boldsymbol{I}$$

$$\implies \qquad \big\| \mathbb{E} \big[\sum\nolimits_{i,j} \boldsymbol{X}_{i,j} \boldsymbol{X}_{i,j}^{\top} \big] \big\| \leq \frac{n}{p} \max_{i,j} (M_{i,j}^{\star})^2 \lesssim \frac{\log^2 n}{np} \quad (\mathsf{check})$$

Similar bounds hold for $\|\mathbb{E}[\sum_{i,j} X_{i,j}^{\top} X_{i,j}]\|$. Therefore,

$$v := \max \left\{ \left\| \mathbb{E} \left[\sum_{i,j} \boldsymbol{X}_{i,j} \boldsymbol{X}_{i,j}^{\top} \right] \right\|, \left\| \mathbb{E} \left[\sum_{i,j} \boldsymbol{X}_{i,j}^{\top} \boldsymbol{X}_{i,j} \right] \right\| \right\} \lesssim \frac{\log^2 n}{np}$$

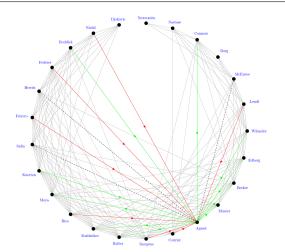
Proof of inequality (5.4)

Take the matrix Bernstein inequality to yield: if $p \gg \log^3 n/n$, then

$$\|\hat{\boldsymbol{M}} - \boldsymbol{M}\| \lesssim \sqrt{v \log n} + B \log n \approx \sqrt{\frac{\log^3 n}{n}} \ll 1$$



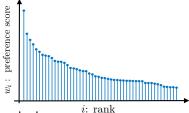
Ranking from pairwise comparisons



pairwise comparisons for ranking tennis players

figure credit: Bozóki, Csató, Temesi

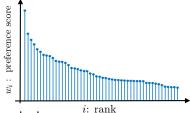
Bradley-Terry-Luce (logistic) model



- ullet n items to be ranked
- assign a latent score $\{w_i^\star\}_{1\leq i\leq n}$ to each item, so that item $i\succ$ item j if $w_i^\star>w_j^\star$
- ullet each pair of items (i,j) is compared independently

$$\mathbb{P}\left\{\text{item } j \text{ beats item } i\right\} = \frac{w_j^\star}{w_i^\star + w_j^\star}$$

Bradley-Terry-Luce (logistic) model



- \bullet *n* items to be ranked
- \bullet assign a latent score $\{w_i^{\star}\}_{1\leq i\leq n}$ to each item, so that

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ullet each pair of items (i,j) is compared independently

$$y_{i,j} \stackrel{\text{ind.}}{=} \begin{cases} 1, & \text{with prob. } \frac{w_j^\star}{w_i^\star + w_j^\star} \\ 0, & \text{else} \end{cases}$$

• intermediate goal: estimate score vector w^* (up to scaling)

Spectral ranking

1. identify key matrix P^* —probability transition matrix

$$P_{i,j}^{\star} = \begin{cases} \frac{1}{n} \cdot \frac{w_j^{\star}}{w_i^{\star} + w_j^{\star}}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}^{\star}, & \text{if } i = j \end{cases}$$

Rationale:

 \circ $oldsymbol{P}^{\star}$ obeys

$$w_i^{\star} P_{i,j}^{\star} = w_j^{\star} P_{j,i}^{\star}$$
 (detailed balance)

 \circ Thus, the stationary distribution π^\star of P^\star obeys

$$\pi^{\star} = \frac{1}{\sum_{l} w_{l}^{\star}} w^{\star}$$
 (reveals true scores)

Spectral ranking

2. construct a surrogate matrix \boldsymbol{P} obeying

$$P_{i,j} = \begin{cases} \frac{1}{n} y_{i,j}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}, & \text{if } i = j \end{cases}$$

3. return leading left eigenvector π of P as score estimate

— closely related to PageRank

Spectral ranking

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Key: stability of eigenspace against perturbation $M-M^{\star}$?

Statistical guarantees for spectral ranking

— Negahban, Oh, Shah'16, Chen, Fan, Ma, Wang'19

Suppose $\max_{i,j} \frac{w_i}{w_j} \lesssim 1$. Then with high prob.

$$\frac{\|\hat{\pi} - \pi\|_2}{\|\pi\|_2} \asymp \frac{\|\hat{\pi} - \pi\|_\pi}{\|\pi\|_2} \lesssim \underbrace{\frac{1}{\sqrt{n}}}_{\text{nearly perfect estimate}} \to 0$$

• a consequence of Theorem ?? and matrix Bernstein (exercise)