

Applications of spectral methods (ℓ_2 theory)



Cong Ma

University of Chicago, Autumn 2021

What we have learned so far

- Classical ℓ_2 matrix perturbation theory:
 - Davis-Kahan's $\sin \Theta$ theorem
 - Wedin's $\sin \Theta$ theorem
 - Eigenvector perturbation of probability transition matrices
- Matrix concentration inequalities:
 - Matrix Bernstein inequality

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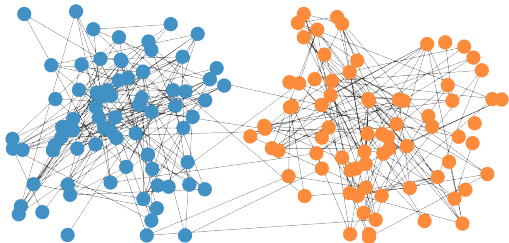
— *we will see their applications today*

Outline

- Community recovery in stochastic block model
 - *application of Davis-Kahan's theorem*
- Low-rank matrix completion
 - *application of Wedin's theorem*
- Ranking from pairwise comparisons
 - *application of eigenvector perturbation of prob. transition matrix*

Community recovery in stochastic block model

Stochastic block model (SBM)

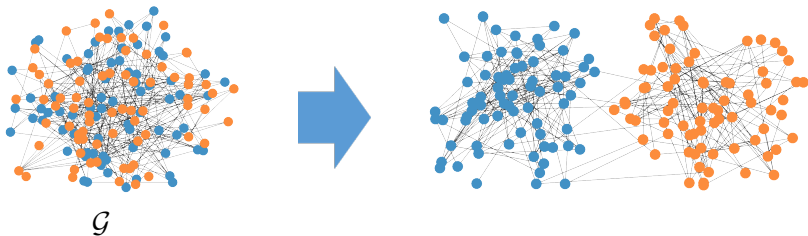


$x_i^* = 1$: 1st community

$x_i^* = -1$: 2nd community

- n nodes $\{1, \dots, n\}$
- 2 communities
- n unknown variables: $x_1^*, \dots, x_n^* \in \{1, -1\}$
 - encode community memberships

Stochastic block model (SBM)



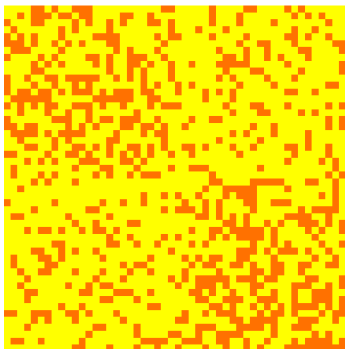
- observe a graph \mathcal{G}

$$(i, j) \in \mathcal{G} \text{ with prob. } \begin{cases} p, & \text{if } i \text{ and } j \text{ are from same community} \\ q, & \text{else} \end{cases}$$

Here, $p > q$

- **Goal:** recover community memberships of all nodes, i.e., $\{x_i^*\}$

Adjacency matrix

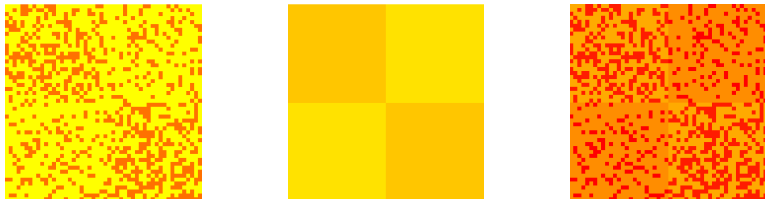


Consider the adjacency matrix $\mathbf{A} \in \mathbb{R}^{n \times n}$ of \mathcal{G} : (assume $A_{ii} = p$)

$$A_{i,j} = \begin{cases} 1, & \text{if } (i,j) \in \mathcal{G} \\ 0, & \text{else} \end{cases}$$

- WLOG, suppose $x_1^* = \dots = x_{n/2}^* = 1$; $x_{n/2+1}^* = \dots = x_n^* = -1$

Adjacency matrix



$$\mathbf{A} = \underbrace{\mathbb{E}[\mathbf{A}]}_{\text{rank 2}} + \mathbf{A} - \mathbb{E}[\mathbf{A}]$$

$$\mathbb{E}[\mathbf{A}] = \begin{bmatrix} p\mathbf{1}\mathbf{1}^\top & q\mathbf{1}\mathbf{1}^\top \\ q\mathbf{1}\mathbf{1}^\top & p\mathbf{1}\mathbf{1}^\top \end{bmatrix} = \underbrace{\frac{p+q}{2}\mathbf{1}\mathbf{1}^\top}_{\text{uninformative bias}} + \frac{p-q}{2} \underbrace{\begin{bmatrix} \mathbf{1} \\ -\mathbf{1} \end{bmatrix}}_{=\mathbf{x}^*=[x_i]_{1 \leq i \leq n}} [\mathbf{1}^\top, -\mathbf{1}^\top]$$

Spectral clustering



The diagram illustrates the decomposition of a matrix A into its expected value and a residual. On the left is a noisy heatmap representing A . In the middle is a block matrix representing $\mathbb{E}[A]$, which is the expected value of A . This matrix is divided into four quadrants: the top-left and bottom-right quadrants are a lighter yellow, while the top-right and bottom-left quadrants are a darker yellow. To the right of the block matrix is a plus sign, followed by another noisy heatmap representing the residual matrix $A - \mathbb{E}[A]$. Below the block matrix, the text $\mathbb{E}[A]$ is written, with a bracket underneath it and the text "rank 2" in blue below the bracket.

$$\mathbf{A} = \underbrace{\mathbb{E}[\mathbf{A}]}_{\text{rank 2}} + \mathbf{A} - \mathbb{E}[\mathbf{A}]$$

1. computing the leading eigenvector $\mathbf{u} = [u_i]_{1 \leq i \leq n}$ of $\mathbf{A} - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top$
2. rounding: output $x_i = \begin{cases} 1, & \text{if } u_i \geq 0 \\ -1, & \text{if } u_i < 0 \end{cases}$

Analysis of spectral clustering

Consider “ground-truth” matrix

$$\mathbf{M}^* := \mathbb{E}[\mathbf{A}] - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top = \frac{p-q}{2} \begin{bmatrix} \mathbf{1} \\ -\mathbf{1} \end{bmatrix} \begin{bmatrix} \mathbf{1}^\top & -\mathbf{1}^\top \end{bmatrix},$$

which obeys

$$\lambda_1(\mathbf{M}^*) := \frac{(p-q)n}{2}, \quad \text{and} \quad \mathbf{u}^* := \frac{1}{\sqrt{n}} \begin{bmatrix} \mathbf{1}_{n/2} \\ -\mathbf{1}_{n/2} \end{bmatrix}.$$

Also, we have perturbed matrix

$$\mathbf{M} := \mathbf{A} - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top$$

Davis-Kahan implies if $\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| < \lambda_1(\mathbf{M}^*) = \frac{(p-q)n}{2}$, then

$$\text{dist}(\mathbf{u}, \mathbf{u}^*) \leq \frac{\|\mathbf{M} - \mathbf{M}^*\|}{\lambda_1(\mathbf{M}^*) - \|\mathbf{M} - \mathbf{M}^*\|} = \frac{\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|}{\frac{(p-q)n}{2} - \|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|} \quad (5.1)$$

Bounding $\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|$

Matrix Bernstein inequality tells us that

Lemma 5.1

Consider SBM with $p > q$ and $p \gtrsim \frac{\log n}{n}$. Then with high prob.

$$\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| \lesssim \sqrt{np \log n} \quad (5.2)$$

— better concentration yields \sqrt{np} bound

- with high probability in this course often means “with probability at least $1 - O(n^{-8})$ ”

Statistical accuracy of spectral clustering

Substitute ineq. (5.2) into ineq. (5.1) to reach

$$\text{dist}(\mathbf{u}, \mathbf{u}^*) \leq \frac{\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|}{\frac{(p-q)n}{2} - \|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|} \lesssim \frac{\sqrt{np \log n}}{(p-q)n} = o(1)$$

provided that $\sqrt{np \log n} = o((p-q)n)$

Now question is

— *how to transfer from estimation error to mis-clustering error*

From estimation error to mis-clustering error

WLOG assume that $\|\mathbf{u} - \mathbf{u}^*\|_2 = \text{dist}(\mathbf{u}, \mathbf{u}^*)$. Consider the set

$$\mathcal{N} := \{i \mid |u_i - u_i^*| \geq 1/\sqrt{n}\}$$

We claim that

$$\frac{1}{n} \sum_{i=1}^n \mathbb{1}\{x_i \neq x_i^*\} \leq \frac{1}{n} \sum_{i=1}^n \mathbb{1}\left\{|u_i - u_i^*| \geq \frac{1}{\sqrt{n}}\right\} = \frac{|\mathcal{N}|}{n}$$

To see this, observe that for any i obeying $x_i \neq x_i^*$, one has $\text{sgn}(u_i) \neq \text{sgn}(u_i^*)$, thus indicating that $|u_i - u_i^*| \geq |u_i^*| = 1/\sqrt{n}$. In the end, we have

$$|\mathcal{N}| \leq \frac{\|\mathbf{u} - \mathbf{u}^*\|_2^2}{(1/\sqrt{n})^2} = o(n)$$

Statistical accuracy of spectral clustering

$$\frac{p - q}{\sqrt{p}} \gg \sqrt{\frac{\log n}{n}} \implies \text{almost exact recovery}$$

- **dense regime:** if $p \asymp q \asymp 1$, then this condition reads

$$p - q \gg \sqrt{\frac{\log n}{n}} \quad (\text{extremely small gap})$$

- **“sparse” regime:** if $p = \frac{a \log n}{n}$ and $q = \frac{b \log n}{n}$ for $a, b \asymp 1$, then

$$a - b \gg \sqrt{a}$$

This condition is information-theoretically optimal (up to log factor)
— Mossel, Neeman, Sly '15, Abbe '18

Proof of Lemma 5.2

We write $\mathbf{A} - \mathbb{E}[\mathbf{A}]$ as sum of independent random matrices

$$\mathbf{A} - \mathbb{E}[\mathbf{A}] = \sum_{i < j} (A_{i,j} - \mathbb{E}[A_{i,j}]) (\mathbf{e}_i \mathbf{e}_j^\top + \mathbf{e}_j \mathbf{e}_i^\top)$$

We only need to consider $\mathbf{A}_{\text{upper}} := \underbrace{\sum_{i < j} (A_{i,j} - \mathbb{E}[A_{i,j}]) \mathbf{e}_i \mathbf{e}_j^\top}_{=: \mathbf{X}_{i,j}}$

- First, $\|\mathbf{X}_{i,j}\| \leq 1 =: B$
- Since $\text{Var}(A_{i,j}) \leq p$, one has $\mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq p \mathbf{e}_i \mathbf{e}_i^\top$, which gives

$$\sum_{i < j} \mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \sum_{i < j} p \mathbf{e}_i \mathbf{e}_i^\top \preceq np \mathbf{I}_n$$

Similarly, $\sum_{i < j} \mathbb{E} [\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \preceq np \mathbf{I}_n$. As a result,

$$v := \max \left\{ \left\| \sum_{i,j} \mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \sum_{i,j} \mathbb{E} [\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \leq np$$

Proof of Lemma 5.2 (cont.)

Take the matrix Bernstein inequality to conclude that with high prob.,

$$\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| \lesssim \sqrt{v \log n} + B \log n \lesssim \sqrt{np \log n}$$

— as long as $p \gtrsim \frac{\log n}{n}$

Low-rank matrix completion

Low-rank matrix completion



figure credit: Candès

- consider a low-rank matrix $M^* = U^* \Sigma^* V^{*\top}$
- each entry $M_{i,j}^*$ is observed independently with prob. p
- **intermediate goal:** estimate U^*, V^*

Spectral method for matrix completion

1. identify the key matrix M^\star
2. construct surrogate matrix $M \in \mathbb{R}^{n \times n}$ as

$$M_{i,j} = \begin{cases} \frac{1}{p} M_{i,j}^\star, & \text{if } M_{i,j}^\star \text{ is observed} \\ 0, & \text{else} \end{cases}$$

- **rationale for rescaling:** ensures $\mathbb{E}[M] = M^\star$

3. compute the rank- r SVD $U\Sigma V^\top$ of M , and return (U, Σ, V)

Statistical accuracy of spectral estimate

Let's analyze a simple case where $\mathbf{M}^\star = \mathbf{u}^\star \mathbf{v}^{\star\top}$ with

$$\mathbf{u}^\star = \frac{1}{\|\tilde{\mathbf{u}}\|_2} \tilde{\mathbf{u}}, \quad \mathbf{v}^\star = \frac{1}{\|\tilde{\mathbf{v}}\|_2} \tilde{\mathbf{v}}, \quad \tilde{\mathbf{u}}, \tilde{\mathbf{v}} \stackrel{\text{indep.}}{\sim} \mathcal{N}(\mathbf{0}, \mathbf{I}_n)$$

From Wedin's Theorem: if $\|\mathbf{M} - \mathbf{M}^\star\| \leq \frac{1}{2} \sigma_1(\mathbf{M}^\star) = \frac{1}{2}$, then

$$\max \{ \text{dist}(\mathbf{u}, \mathbf{u}^\star), \text{dist}(\mathbf{v}, \mathbf{v}^\star) \} \lesssim \frac{\|\mathbf{M} - \mathbf{M}^\star\|}{\sigma_1(\mathbf{M}^\star)} \asymp \|\mathbf{M} - \mathbf{M}^\star\| \quad (5.3)$$

Bounding $\|M - M^*\|$

Matrix Bernstein inequality tells us that

Lemma 5.2

Consider matrix completion with $p \gg \frac{\log^3 n}{n}$. Then with high prob.

$$\|M - M^*\| \lesssim \sqrt{\frac{\log^3 n}{n}} = o(1) \quad (5.4)$$

Sample complexity

For rank-1 matrix completion,

$$p \gg \frac{\log^3 n}{n} \quad \implies \quad \text{nearly accurate estimates}$$

Sample complexity needed to yield reliable spectral estimates is

$$\underbrace{n^2 p \asymp n \log^3 n}_{\text{optimal up to log factor}}$$

— *sub-optimal accuracy though*

Proof of inequality (5.4)

Write $M - M^* = \sum_{i,j} \mathbf{X}_{i,j}$, where $\mathbf{X}_{i,j} = (M_{i,j} - M_{i,j}^*) \mathbf{e}_i \mathbf{e}_j^\top$

- First, based on Gaussianity, we have

$$\|\mathbf{X}_{i,j}\| \leq \frac{1}{p} \max_{i,j} |M_{i,j}^*| \lesssim \frac{\log n}{pn} := B \quad (\text{check})$$

- Next, $\mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] = \text{Var}(M_{i,j}) \mathbf{e}_i \mathbf{e}_i^\top$ and hence

$$\mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \left\{ \max_{i,j} \text{Var}(M_{i,j}) \right\} n \mathbf{I} \preceq \left\{ \frac{n}{p} \max_{i,j} (M_{i,j}^*)^2 \right\} \mathbf{I}$$

$$\implies \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\| \leq \frac{n}{p} \max_{i,j} (M_{i,j}^*)^2 \lesssim \frac{\log^2 n}{np} \quad (\text{check})$$

Similar bounds hold for $\|\mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}]\|$. Therefore,

$$v := \max \left\{ \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \lesssim \frac{\log^2 n}{np}$$

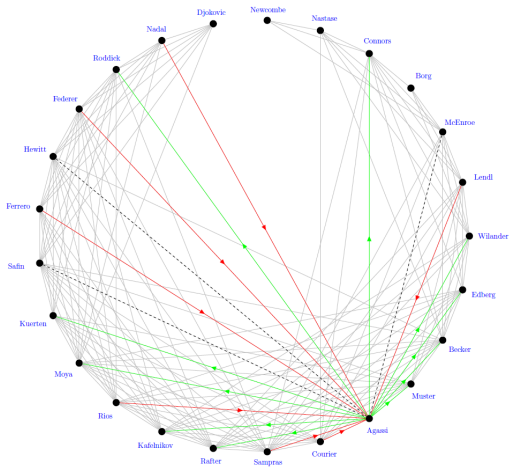
Proof of inequality (5.4) (cont.)

Take the matrix Bernstein inequality to yield: if $p \gg (\log^3 n)/n$, then

$$\|\mathbf{M} - \mathbf{M}^\star\| \lesssim \sqrt{v \log n} + B \log n \asymp \sqrt{\frac{\log^3 n}{n}} \ll 1$$

Ranking from pairwise comparisons

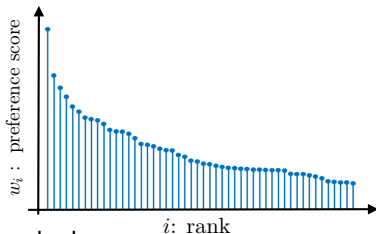
Ranking from pairwise comparisons



pairwise comparisons for ranking tennis players

figure credit: Bozók, Csátó, Temesi

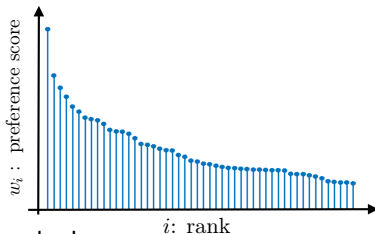
Bradley-Terry-Luce (logistic) model



- n items to be ranked
- assign a latent positive score $\{w_i^*\}_{1 \leq i \leq n}$ to each item, so that
item $i \succ$ item j if $w_i^* > w_j^*$
- each pair of items (i, j) is compared independently

$$\mathbb{P}\{\text{item } j \text{ beats item } i\} = \frac{w_j^*}{w_i^* + w_j^*}$$

Bradley-Terry-Luce (logistic) model



- n items to be ranked
- assign a latent positive score $\{w_i^*\}_{1 \leq i \leq n}$ to each item, so that

$$\text{item } i \succ \text{item } j \quad \text{if} \quad w_i^* > w_j^*$$

- each pair of items (i, j) is compared independently

$$y_{i,j} \stackrel{\text{ind.}}{=} \begin{cases} 1, & \text{with prob. } \frac{w_j^*}{w_i^* + w_j^*} \\ 0, & \text{else} \end{cases}$$

- **intermediate goal:** estimate score vector w^* (up to scaling)

Spectral ranking

1. identify key matrix P^* —probability transition matrix

$$P_{i,j}^* = \begin{cases} \frac{1}{n} \cdot \frac{w_j^*}{w_i^* + w_j^*}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}^*, & \text{if } i = j \end{cases}$$

Rationale:

- P^* obeys

$$w_i^* P_{i,j}^* = w_j^* P_{j,i}^* \quad (\text{detailed balance})$$

- Thus, the stationary distribution π^* of P^* obeys

$$\pi^* = \frac{1}{\sum_l w_l^*} w^* \quad (\text{reveals true scores})$$

Spectral ranking

2. construct a surrogate matrix P obeying

$$P_{i,j} = \begin{cases} \frac{1}{n} y_{i,j}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}, & \text{if } i = j \end{cases}$$

3. return leading left eigenvector π of P as score estimate

— closely related to PageRank

Analysis of spectral ranking

Apply our perturbation bound to see

$$\|\pi - \pi^*\|_{\pi^*} \leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}}$$

provided that

$$1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*} > 0 \quad (5.5)$$

Analysis of spectral ranking

Apply our perturbation bound to see

$$\|\pi - \pi^*\|_{\pi^*} \leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}}$$

provided that

$$1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*} > 0 \quad (5.5)$$

— *need to understand spectral gap and noise size*

Spectral gap of Markov chain

Define condition number

$$\kappa := \frac{\max_{1 \leq i \leq n} w_i^*}{\min_{1 \leq i \leq n} w_i^*}$$

Lemma 5.3

It follows that

$$1 - \max \{ \lambda_2(\mathbf{P}^*), -\lambda_n(\mathbf{P}^*) \} \geq \frac{1}{2\kappa^2}.$$

- We omit the proof; it's based on comparison between two reversible Markov chains

Bound $\|E\|_{\pi^\star}$

Recall that $E := P - P^\star$

Lemma 5.4

With probability at least $1 - O(n^{-8})$,

$$\|E\|_{\pi^\star} \leq \sqrt{\kappa} \|E\| \lesssim \sqrt{\frac{\kappa \log n}{n}}.$$

Analysis of spectral ranking (cont.)

Recall perturbation bound

$$\begin{aligned}\|\pi - \pi^*\|_{\pi^*} &\leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}} \\ &\leq 4\kappa^2 \|\pi^{*\top} E\|_{\pi^*} \quad (\text{provided that } n \gg \kappa^5 \log n)\end{aligned}$$

Note that for any v , one has

$$\|v\|_{\pi^*} \leq \sqrt{\pi_{\max}^*} \|v\|_2, \quad \text{and} \quad \|v\|_2 \leq \frac{1}{\sqrt{\pi_{\min}^*}} \|v\|_{\pi^*}$$

As a result, one has

$$\begin{aligned}\|\pi - \pi^*\|_2 &\leq \frac{1}{\sqrt{\pi_{\min}^*}} \|\pi - \pi^*\|_{\pi^*} \leq \frac{4\kappa^2}{\sqrt{\pi_{\min}^*}} \|\pi^{*\top} E\|_{\pi^*} \\ &\leq 4\kappa^{2.5} \|\pi^{*\top} E\|_2 \leq 4\kappa^{2.5} \|E\| \|\pi^*\|_2\end{aligned}$$

Proof of Lemma 5.4

By construction of \mathbf{P} and \mathbf{P}^\star , we see that

$$E_{i,j} = P_{i,j} - P_{i,j}^\star = \frac{1}{n} (y_{i,j} - \mathbb{E}[y_{i,j}]) \quad (5.6)$$

for any $i \neq j$. In addition, for all $1 \leq i \leq n$, it follows that

$$E_{i,i} = P_{i,i} - P_{i,i}^\star = - \sum_{j:j \neq i} E_{i,j} = -\frac{1}{n} \sum_{j:j \neq i} (y_{i,j} - \mathbb{E}[y_{i,j}]). \quad (5.7)$$

We shall decompose the matrix \mathbf{E} into three parts: upper triangular part, diagonal part, and lower triangular part:

$$\|\mathbf{E}\| \leq \|\mathbf{E}_{\text{upper}}\| + \|\mathbf{E}_{\text{diag}}\| + \|\mathbf{E}_{\text{lower}}\| \quad (5.8)$$

— we will upper bound $\|\mathbf{E}_{\text{upper}}\|$

Control $\|E_{\text{diag}}\|$

Note that

$$\|E_{\text{diag}}\| = \max_{1 \leq i \leq n} |E_{i,i}| = \max_{1 \leq i \leq n} \frac{1}{n} \left| \underbrace{\sum_{j:j \neq i} (y_{i,j} - \mathbb{E}[y_{i,j}])}_{=: X_j} \right|$$

- First, we have $|X_j| \leq 1 =: B$
- Second, one has

$$\sum_{j:j \neq i} \mathbb{E}[X_j^2] = \sum_{j:j \neq i} \text{Var}(y_{i,j}) \leq n =: v$$

By Bernstein's inequality and union bound, we have w.h.p.

$$\max_i |E_{i,i}| \lesssim \frac{1}{n} \cdot (\sqrt{v \log n} + B \log n) \asymp \sqrt{\frac{\log n}{n}}$$

Control $\|E_{\text{upper}}\|$

First of all, we have

$$E_{\text{upper}} = \sum_{i < j} E_{i,j} \mathbf{e}_i \mathbf{e}_j^\top = \sum_{i < j} \underbrace{\frac{1}{n} (y_{i,j} - \mathbb{E}[y_{i,j}]) \mathbf{e}_i \mathbf{e}_j^\top}_{=: \mathbf{X}_{i,j}}$$

Then

- $\|\mathbf{X}_{i,j}\| \leq \frac{1}{n} =: B$
- Since $\text{Var}(y_{i,j}) \leq 1$, one has $\mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \frac{1}{n^2} \mathbf{e}_i \mathbf{e}_i^\top$, which gives

$$\sum_{i < j} \mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \sum_{i < j} \frac{1}{n^2} \mathbf{e}_i \mathbf{e}_i^\top \preceq \frac{1}{n} \mathbf{I}_n$$

Similarly, $\sum_{i < j} \mathbb{E}[\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \preceq \frac{1}{n} \mathbf{I}_n$. As a result,

$$v := \max \left\{ \left\| \sum_{i,j} \mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \sum_{i,j} \mathbb{E}[\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \leq \frac{1}{n}$$

Control $\|E_{\text{upper}}\|$ (cont.)

Invoke matrix Bernstein to obtain

$$\|E_{\text{upper}}\| \lesssim \sqrt{v \log n} + B \log n \asymp \sqrt{\frac{\log n}{n}}$$

— *same bound holds for $\|E_{\text{lower}}\|$*

Putting pieces together

Assuming $\kappa = O(1)$, we have

$$\|\boldsymbol{\pi} - \boldsymbol{\pi}^*\|_2 \lesssim \sqrt{\frac{\log n}{n}} \|\boldsymbol{\pi}^*\|_2$$

- vanishing relative error when n goes to infinity
- optimal error up to a log factor

— Negahban, Oh, Shah '16, Chen, Fan, Ma, Wang '19