

## **Applications of spectral methods ( $\ell_2$ theory)**



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# What we have learned so far

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- Classical  $\ell_2$  matrix perturbation theory:
  - Davis-Kahan's  $\sin \Theta$  theorem
  - Wedin's  $\sin \Theta$  theorem
  - Eigenvector perturbation of probability transition matrices
- Matrix concentration inequalities:
  - Matrix Bernstein inequality

# What we have learned so far

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- Classical  $\ell_2$  matrix perturbation theory:
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  - Eigenvector perturbation of probability transition matrices
- Matrix concentration inequalities:
  - Matrix Bernstein inequality

— *we will see their applications today*

# Outline

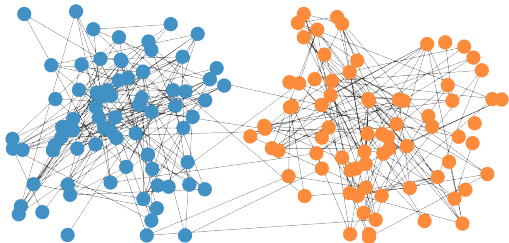
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- Community recovery in stochastic block model
  - *application of Davis-Kahan's theorem*
- Low-rank matrix completion
  - *application of Wedin's theorem*
- Ranking from pairwise comparisons
  - *application of eigenvector perturbation of prob. transition matrix*

## **Community recovery in stochastic block model**

# Stochastic block model (SBM)

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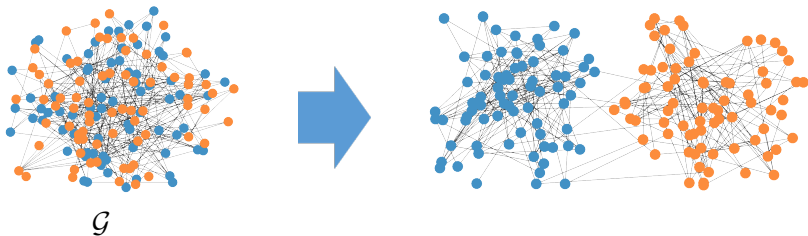
$x_i^* = 1$ : 1<sup>st</sup> community

$x_i^* = -1$ : 2<sup>nd</sup> community

- $n$  nodes  $\{1, \dots, n\}$
- 2 communities
- $n$  unknown variables:  $x_1^*, \dots, x_n^* \in \{1, -1\}$ 
  - encode community memberships

# Stochastic block model (SBM)

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- observe a graph  $\mathcal{G}$

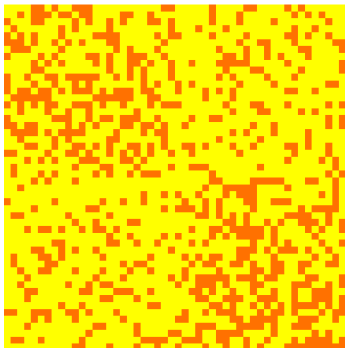
$$(i, j) \in \mathcal{G} \text{ with prob. } \begin{cases} p, & \text{if } i \text{ and } j \text{ are from same community} \\ q, & \text{else} \end{cases}$$

Here,  $p > q$

- **Goal:** recover community memberships of all nodes, i.e.,  $\{x_i^*\}$

# Adjacency matrix

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Consider the adjacency matrix  $\mathbf{A} \in \mathbb{R}^{n \times n}$  of  $\mathcal{G}$ : (assume  $A_{ii} = p$ )

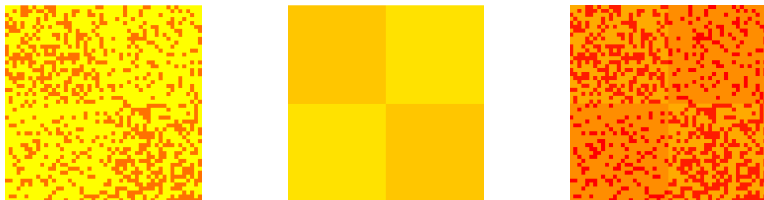
$$A_{i,j} = \begin{cases} 1, & \text{if } (i,j) \in \mathcal{G} \\ 0, & \text{else} \end{cases}$$

- WLOG, suppose  $x_1^* = \cdots = x_{n/2}^* = 1$ ;  $x_{n/2+1}^* = \cdots = x_n^* = -1$



# Adjacency matrix

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$$A = \underbrace{\mathbb{E}[A]}_{\text{rank 2}} + A - \mathbb{E}[A]$$

$$\mathbb{E}[A] = \begin{bmatrix} p\mathbf{1}\mathbf{1}^\top & q\mathbf{1}\mathbf{1}^\top \\ q\mathbf{1}\mathbf{1}^\top & p\mathbf{1}\mathbf{1}^\top \end{bmatrix} = \underbrace{\frac{p+q}{2}\mathbf{1}\mathbf{1}^\top}_{\text{uninformative bias}} + \frac{p-q}{2} \underbrace{\begin{bmatrix} \mathbf{1} \\ -\mathbf{1} \end{bmatrix}}_{=\mathbf{x}^*=[x_i]_{1 \leq i \leq n}} [\mathbf{1}^\top, -\mathbf{1}^\top]$$

# Spectral clustering

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The diagram illustrates the decomposition of a matrix  $A$  into its expected value and a residual matrix. On the left is a noisy heatmap representing  $A$ . In the middle is a block matrix representing  $\mathbb{E}[A]$ , which is the expected value of  $A$ . This block matrix is divided into four quadrants: the top-left and bottom-right quadrants are a lighter yellow, while the top-right and bottom-left quadrants are a darker yellow. To the right of the block matrix is a plus sign, followed by another noisy heatmap representing the residual matrix  $A - \mathbb{E}[A]$ . Below the block matrix, the text  $\mathbb{E}[A]$  is written, with a bracket underneath it and the text "rank 2" in blue below the bracket.

$$\mathbf{A} = \underbrace{\mathbb{E}[\mathbf{A}]}_{\text{rank 2}} + \mathbf{A} - \mathbb{E}[\mathbf{A}]$$

1. computing the leading eigenvector  $\mathbf{u} = [u_i]_{1 \leq i \leq n}$  of  $\mathbf{A} - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top$
2. rounding: output  $x_i = \begin{cases} 1, & \text{if } u_i \geq 0 \\ -1, & \text{if } u_i < 0 \end{cases}$

# Analysis of spectral clustering

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Consider “ground-truth” matrix

$$\mathbf{M}^* := \mathbb{E}[\mathbf{A}] - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top = \frac{p-q}{2} \begin{bmatrix} \mathbf{1} \\ -\mathbf{1} \end{bmatrix} \begin{bmatrix} \mathbf{1}^\top & -\mathbf{1}^\top \end{bmatrix},$$

which obeys

$$\lambda_1(\mathbf{M}^*) := \frac{(p-q)n}{2}, \quad \text{and} \quad \mathbf{u}^* := \frac{1}{\sqrt{n}} \begin{bmatrix} \mathbf{1}_{n/2} \\ -\mathbf{1}_{n/2} \end{bmatrix}.$$

Also, we have perturbed matrix

$$\mathbf{M} := \mathbf{A} - \frac{p+q}{2} \mathbf{1}\mathbf{1}^\top$$

Davis-Kahan implies if  $\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| < \lambda_1(\mathbf{M}^*) = \frac{(p-q)n}{2}$ , then

$$\text{dist}(\mathbf{u}, \mathbf{u}^*) \leq \frac{\|\mathbf{M} - \mathbf{M}^*\|}{\lambda_1(\mathbf{M}^*) - \|\mathbf{M} - \mathbf{M}^*\|} = \frac{\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|}{\frac{(p-q)n}{2} - \|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|} \quad (5.1)$$

## Bounding $\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|$

---

Matrix Bernstein inequality tells us that

### Lemma 5.1

Consider SBM with  $p > q$  and  $p \gtrsim \frac{\log n}{n}$ . Then with high prob.

$$\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| \lesssim \sqrt{np \log n} \quad (5.2)$$

— better concentration yields  $\sqrt{np}$  bound

- with high probability in this course often means “with probability at least  $1 - O(n^{-8})$ ”

# Statistical accuracy of spectral clustering

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Substitute ineq. (5.2) into ineq. (5.1) to reach

$$\text{dist}(\mathbf{u}, \mathbf{u}^*) \leq \frac{\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|}{\frac{(p-q)n}{2} - \|\mathbf{A} - \mathbb{E}[\mathbf{A}]\|} \lesssim \frac{\sqrt{np \log n}}{(p-q)n} = o(1)$$

provided that  $\sqrt{np \log n} = o((p-q)n)$

Now question is

— *how to transfer from estimation error to mis-clustering error*

# From estimation error to mis-clustering error

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WLOG assume that  $\|\mathbf{u} - \mathbf{u}^*\|_2 = \text{dist}(\mathbf{u}, \mathbf{u}^*)$ . Consider the set

$$\mathcal{N} := \{i \mid |u_i - u_i^*| \geq 1/\sqrt{n}\}$$

We claim that

$$\frac{1}{n} \sum_{i=1}^n \mathbb{1}\{x_i \neq x_i^*\} \leq \frac{1}{n} \sum_{i=1}^n \mathbb{1}\left\{|u_i - u_i^*| \geq \frac{1}{\sqrt{n}}\right\} = \frac{|\mathcal{N}|}{n}$$

To see this, observe that for any  $i$  obeying  $x_i \neq x_i^*$ , one has  $\text{sgn}(u_i) \neq \text{sgn}(u_i^*)$ , thus indicating that  $|u_i - u_i^*| \geq |u_i^*| = 1/\sqrt{n}$ . In the end, we have

$$|\mathcal{N}| \leq \frac{\|\mathbf{u} - \mathbf{u}^*\|_2^2}{(1/\sqrt{n})^2} = o(n)$$

# Statistical accuracy of spectral clustering

$$\frac{p - q}{\sqrt{p}} \gg \sqrt{\frac{\log n}{n}} \implies \text{almost exact recovery}$$

- **dense regime:** if  $p \asymp q \asymp 1$ , then this condition reads

$$p - q \gg \sqrt{\frac{\log n}{n}} \quad (\text{extremely small gap})$$

- **“sparse” regime:** if  $p = \frac{a \log n}{n}$  and  $q = \frac{b \log n}{n}$  for  $a, b \asymp 1$ , then

$$a - b \gg \sqrt{a}$$

This condition is information-theoretically optimal (up to log factor)  
— Mossel, Neeman, Sly '15, Abbe '18

## Proof of Lemma 5.2

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We write  $\mathbf{A} - \mathbb{E}[\mathbf{A}]$  as sum of independent random matrices

$$\mathbf{A} - \mathbb{E}[\mathbf{A}] = \sum_{i < j} (A_{i,j} - \mathbb{E}[A_{i,j}]) (\mathbf{e}_i \mathbf{e}_j^\top + \mathbf{e}_j \mathbf{e}_i^\top)$$

We only need to consider  $\mathbf{A}_{\text{upper}} := \underbrace{\sum_{i < j} (A_{i,j} - \mathbb{E}[A_{i,j}]) \mathbf{e}_i \mathbf{e}_j^\top}_{=: \mathbf{X}_{i,j}}$

- First,  $\|\mathbf{X}_{i,j}\| \leq 1 =: B$
- Since  $\text{Var}(A_{i,j}) \leq p$ , one has  $\mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq p \mathbf{e}_i \mathbf{e}_i^\top$ , which gives

$$\sum_{i < j} \mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \sum_{i < j} p \mathbf{e}_i \mathbf{e}_i^\top \preceq np \mathbf{I}_n$$

Similarly,  $\sum_{i < j} \mathbb{E} [\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \preceq np \mathbf{I}_n$ . As a result,

$$v := \max \left\{ \left\| \sum_{i,j} \mathbb{E} [\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \sum_{i,j} \mathbb{E} [\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \leq np$$



## Proof of Lemma 5.2 (cont.)

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Take the matrix Bernstein inequality to conclude that with high prob.,

$$\|\mathbf{A} - \mathbb{E}[\mathbf{A}]\| \lesssim \sqrt{v \log n} + B \log n \lesssim \sqrt{np \log n}$$

— as long as  $p \gtrsim \frac{\log n}{n}$

## **Low-rank matrix completion**

# Low-rank matrix completion



*figure credit: Candès*

- consider a low-rank matrix  $M^* = U^* \Sigma^* V^{*\top}$
- each entry  $M_{i,j}^*$  is observed independently with prob.  $p$
- **intermediate goal:** estimate  $U^*, V^*$

# Spectral method for matrix completion

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1. identify the key matrix  $M^\star$
2. construct surrogate matrix  $M \in \mathbb{R}^{n \times n}$  as

$$M_{i,j} = \begin{cases} \frac{1}{p} M_{i,j}^\star, & \text{if } M_{i,j}^\star \text{ is observed} \\ 0, & \text{else} \end{cases}$$

- **rationale for rescaling:** ensures  $\mathbb{E}[M] = M^\star$

3. compute the rank- $r$  SVD  $U\Sigma V^\top$  of  $M$ , and return  $(U, \Sigma, V)$

# Statistical accuracy of spectral estimate

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Let's analyze a simple case where  $\mathbf{M}^\star = \mathbf{u}^\star \mathbf{v}^{\star\top}$  with

$$\mathbf{u}^\star = \frac{1}{\|\tilde{\mathbf{u}}\|_2} \tilde{\mathbf{u}}, \quad \mathbf{v}^\star = \frac{1}{\|\tilde{\mathbf{v}}\|_2} \tilde{\mathbf{v}}, \quad \tilde{\mathbf{u}}, \tilde{\mathbf{v}} \stackrel{\text{indep.}}{\sim} \mathcal{N}(\mathbf{0}, \mathbf{I}_n)$$

From Wedin's Theorem: if  $\|\mathbf{M} - \mathbf{M}^\star\| \leq \frac{1}{2} \sigma_1(\mathbf{M}^\star) = \frac{1}{2}$ , then

$$\max \{ \text{dist}(\mathbf{u}, \mathbf{u}^\star), \text{dist}(\mathbf{v}, \mathbf{v}^\star) \} \lesssim \frac{\|\mathbf{M} - \mathbf{M}^\star\|}{\sigma_1(\mathbf{M}^\star)} \asymp \|\mathbf{M} - \mathbf{M}^\star\| \quad (5.3)$$

## Bounding $\|M - M^\star\|$

---

Matrix Bernstein inequality tells us that

### Lemma 5.2

*Consider matrix completion with  $p \gg \frac{\log^3 n}{n}$ . Then with high prob.*

$$\|M - M^\star\| \lesssim \sqrt{\frac{\log^3 n}{np}} = o(1) \quad (5.4)$$

# Sample complexity

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For rank-1 matrix completion,

$$p \gg \frac{\log^3 n}{n} \implies \text{nearly accurate estimates}$$

Sample complexity needed to yield reliable spectral estimates is

$$\underbrace{n^2 p \asymp n \log^3 n}_{\text{optimal up to log factor}}$$

— *sub-optimal accuracy though*

## Proof of inequality (5.4)

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Write  $M - M^* = \sum_{i,j} \mathbf{X}_{i,j}$ , where  $\mathbf{X}_{i,j} = (M_{i,j} - M_{i,j}^*) \mathbf{e}_i \mathbf{e}_j^\top$

- First, based on Gaussianity, we have

$$\|\mathbf{X}_{i,j}\| \leq \frac{1}{p} \max_{i,j} |M_{i,j}^*| \lesssim \frac{\log n}{pn} := B \quad (\text{check})$$

- Next,  $\mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] = \text{Var}(M_{i,j}) \mathbf{e}_i \mathbf{e}_i^\top$  and hence

$$\mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \left\{ \max_{i,j} \text{Var}(M_{i,j}) \right\} n \mathbf{I} \preceq \left\{ \frac{n}{p} \max_{i,j} (M_{i,j}^*)^2 \right\} \mathbf{I}$$

$$\implies \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\| \leq \frac{n}{p} \max_{i,j} (M_{i,j}^*)^2 \lesssim \frac{\log^2 n}{np} \quad (\text{check})$$

Similar bounds hold for  $\|\mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}]\|$ . Therefore,

$$v := \max \left\{ \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \mathbb{E}[\sum_{i,j} \mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \lesssim \frac{\log^2 n}{np}$$



## Proof of inequality (5.4) (cont.)

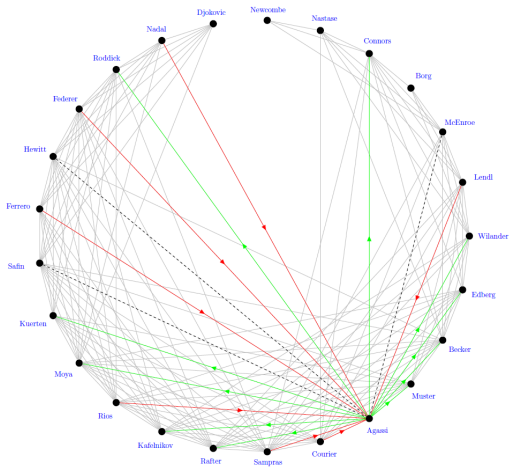
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Take the matrix Bernstein inequality to yield: if  $p \gg (\log^3 n)/n$ , then

$$\|\mathbf{M} - \mathbf{M}^*\| \lesssim \sqrt{v \log n} + B \log n \asymp \sqrt{\frac{\log^3 n}{np}} \ll 1$$

## **Ranking from pairwise comparisons**

# Ranking from pairwise comparisons

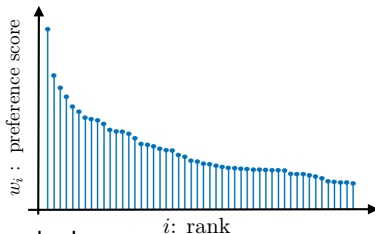


pairwise comparisons for ranking tennis players

figure credit: Bozóki, Csató, Temesi

# Bradley-Terry-Luce (logistic) model

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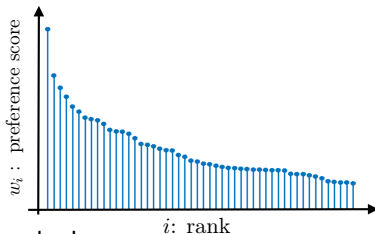


- $n$  items to be ranked
- assign a latent positive score  $\{w_i^*\}_{1 \leq i \leq n}$  to each item, so that  
item  $i \succ$  item  $j$  if  $w_i^* > w_j^*$
- each pair of items  $(i, j)$  is compared independently

$$\mathbb{P}\{\text{item } j \text{ beats item } i\} = \frac{w_j^*}{w_i^* + w_j^*}$$

# Bradley-Terry-Luce (logistic) model

---



- $n$  items to be ranked
- assign a latent positive score  $\{w_i^*\}_{1 \leq i \leq n}$  to each item, so that

$$\text{item } i \succ \text{item } j \quad \text{if} \quad w_i^* > w_j^*$$

- each pair of items  $(i, j)$  is compared independently

$$y_{i,j} \stackrel{\text{ind.}}{=} \begin{cases} 1, & \text{with prob. } \frac{w_j^*}{w_i^* + w_j^*} \\ 0, & \text{else} \end{cases}$$

- **intermediate goal:** estimate score vector  $w^*$  (up to scaling)

# Spectral ranking

---

1. identify key matrix  $P^*$ —probability transition matrix

$$P_{i,j}^* = \begin{cases} \frac{1}{n} \cdot \frac{w_j^*}{w_i^* + w_j^*}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}^*, & \text{if } i = j \end{cases}$$

Rationale:

- $P^*$  obeys

$$w_i^* P_{i,j}^* = w_j^* P_{j,i}^* \quad (\text{detailed balance})$$

- Thus, the stationary distribution  $\pi^*$  of  $P^*$  obeys

$$\pi^* = \frac{1}{\sum_l w_l^*} w^* \quad (\text{reveals true scores})$$

# Spectral ranking

---

2. construct a surrogate matrix  $P$  obeying

$$P_{i,j} = \begin{cases} \frac{1}{n} y_{i,j}, & \text{if } i \neq j \\ 1 - \sum_{l:l \neq i} P_{i,l}, & \text{if } i = j \end{cases}$$

3. return leading left eigenvector  $\pi$  of  $P$  as score estimate

— closely related to PageRank

# Analysis of spectral ranking

---

Apply our perturbation bound to see

$$\|\pi - \pi^*\|_{\pi^*} \leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}}$$

provided that

$$1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*} > 0 \quad (5.5)$$



# Analysis of spectral ranking

---

Apply our perturbation bound to see

$$\|\pi - \pi^*\|_{\pi^*} \leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}}$$

provided that

$$1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*} > 0 \quad (5.5)$$

— *need to understand spectral gap and noise size*

# Spectral gap of Markov chain

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Define condition number

$$\kappa := \frac{\max_{1 \leq i \leq n} w_i^\star}{\min_{1 \leq i \leq n} w_i^\star}$$

## Lemma 5.3

*It follows that*

$$1 - \max \{ \lambda_2(\mathbf{P}^\star), -\lambda_n(\mathbf{P}^\star) \} \geq \frac{1}{2\kappa^2}.$$

- We omit the proof; it's based on comparison between two reversible Markov chains

## Bound $\|E\|_{\pi^\star}$

---

Recall that  $E := P - P^\star$

### Lemma 5.4

*With probability at least  $1 - O(n^{-8})$ ,*

$$\|E\|_{\pi^\star} \leq \sqrt{\kappa} \|E\| \lesssim \sqrt{\frac{\kappa \log n}{n}}.$$

## Analysis of spectral ranking (cont.)

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Recall perturbation bound

$$\begin{aligned}\|\pi - \pi^*\|_{\pi^*} &\leq \frac{\|\pi^{*\top} E\|_{\pi^*}}{1 - \max\{\lambda_2(P^*), -\lambda_n(P^*)\} - \|E\|_{\pi^*}} \\ &\leq 4\kappa^2 \|\pi^{*\top} E\|_{\pi^*} \quad (\text{provided that } n \gg \kappa^5 \log n)\end{aligned}$$

Note that for any  $v$ , one has

$$\|v\|_{\pi^*} \leq \sqrt{\pi_{\max}^*} \|v\|_2, \quad \text{and} \quad \|v\|_2 \leq \frac{1}{\sqrt{\pi_{\min}^*}} \|v\|_{\pi^*}$$

As a result, one has

$$\begin{aligned}\|\pi - \pi^*\|_2 &\leq \frac{1}{\sqrt{\pi_{\min}^*}} \|\pi - \pi^*\|_{\pi^*} \leq \frac{4\kappa^2}{\sqrt{\pi_{\min}^*}} \|\pi^{*\top} E\|_{\pi^*} \\ &\leq 4\kappa^{2.5} \|\pi^{*\top} E\|_2 \leq 4\kappa^{2.5} \|E\| \|\pi^*\|_2\end{aligned}$$

## Proof of Lemma 5.4

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By construction of  $\mathbf{P}$  and  $\mathbf{P}^\star$ , we see that

$$E_{i,j} = P_{i,j} - P_{i,j}^\star = \frac{1}{n} (y_{i,j} - \mathbb{E}[y_{i,j}]) \quad (5.6)$$

for any  $i \neq j$ . In addition, for all  $1 \leq i \leq n$ , it follows that

$$E_{i,i} = P_{i,i} - P_{i,i}^\star = - \sum_{j:j \neq i} E_{i,j} = -\frac{1}{n} \sum_{j:j \neq i} (y_{i,j} - \mathbb{E}[y_{i,j}]). \quad (5.7)$$

We shall decompose the matrix  $\mathbf{E}$  into three parts: upper triangular part, diagonal part, and lower triangular part:

$$\|\mathbf{E}\| \leq \|\mathbf{E}_{\text{upper}}\| + \|\mathbf{E}_{\text{diag}}\| + \|\mathbf{E}_{\text{lower}}\| \quad (5.8)$$

— we will upper bound  $\|\mathbf{E}_{\text{upper}}\|$

## Control $\|E_{\text{diag}}\|$

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Note that

$$\|E_{\text{diag}}\| = \max_{1 \leq i \leq n} |E_{i,i}| = \max_{1 \leq i \leq n} \frac{1}{n} \left| \underbrace{\sum_{j:j \neq i} (y_{i,j} - \mathbb{E}[y_{i,j}])}_{=: X_j} \right|$$

- First, we have  $|X_j| \leq 1 =: B$
- Second, one has

$$\sum_{j:j \neq i} \mathbb{E}[X_j^2] = \sum_{j:j \neq i} \text{Var}(y_{i,j}) \leq n =: v$$

By Bernstein's inequality and union bound, we have w.h.p.

$$\max_i |E_{i,i}| \lesssim \frac{1}{n} \cdot (\sqrt{v \log n} + B \log n) \asymp \sqrt{\frac{\log n}{n}}$$

## Control $\|E_{\text{upper}}\|$

---

First of all, we have

$$E_{\text{upper}} = \sum_{i < j} E_{i,j} \mathbf{e}_i \mathbf{e}_j^\top = \sum_{i < j} \underbrace{\frac{1}{n} (y_{i,j} - \mathbb{E}[y_{i,j}]) \mathbf{e}_i \mathbf{e}_j^\top}_{=: \mathbf{X}_{i,j}}$$

Then

- $\|\mathbf{X}_{i,j}\| \leq \frac{1}{n} =: B$
- Since  $\text{Var}(y_{i,j}) \leq 1$ , one has  $\mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \frac{1}{n^2} \mathbf{e}_i \mathbf{e}_i^\top$ , which gives

$$\sum_{i < j} \mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \preceq \sum_{i < j} \frac{1}{n^2} \mathbf{e}_i \mathbf{e}_i^\top \preceq \frac{1}{n} \mathbf{I}_n$$

Similarly,  $\sum_{i < j} \mathbb{E}[\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \preceq \frac{1}{n} \mathbf{I}_n$ . As a result,

$$v := \max \left\{ \left\| \sum_{i,j} \mathbb{E}[\mathbf{X}_{i,j} \mathbf{X}_{i,j}^\top] \right\|, \left\| \sum_{i,j} \mathbb{E}[\mathbf{X}_{i,j}^\top \mathbf{X}_{i,j}] \right\| \right\} \leq \frac{1}{n}$$

## Control $\|E_{\text{upper}}\|$ (cont.)

---

Invoke matrix Bernstein to obtain

$$\|E_{\text{upper}}\| \lesssim \sqrt{v \log n} + B \log n \asymp \sqrt{\frac{\log n}{n}}$$

— *same bound holds for  $\|E_{\text{lower}}\|$*



# Putting pieces together

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Assuming  $\kappa = O(1)$ , we have

$$\|\boldsymbol{\pi} - \boldsymbol{\pi}^*\|_2 \lesssim \sqrt{\frac{\log n}{n}} \|\boldsymbol{\pi}^*\|_2$$

- vanishing relative error when  $n$  goes to infinity
- optimal error up to a log factor

— Negahban, Oh, Shah '16, Chen, Fan, Ma, Wang '19