Introduction

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1 Some basic aspects of continuous-time Markov Chains

Definition. A sequence of random variables is called a *stochastic process* or *process*. The process $X = (X_n)_{n \geq 1}$ is called a discrete-time Markov Chain with state space I if for all $x_0, x_1, \ldots, x_n \in I$

$$\mathbb{P}(X_n = x_n | X_{n-1} = x_{n-1}, \ldots, X_0 = x_0) = \mathbb{P}(X_n = x_n | X_{n-1} = x_{n-1}).$$

If $\mathbb{P}(X_{n+1} = y | X_n = x)$ is independent of n, the chain is called *time-homogeneous*. We then write $P = (P_{x,y})_{x,y \in I}$ for the *transition matrix* where $P_{x,y} = \mathbb{P}(X_1 = y | X_0 = x)$. The data associated to every time-homogeneous Markov Chain is the transition matrix P and the initial distribution μ, i.e $\mathbb{P}(X_0 = x_0) = \mu(x_0)$.

From now on:

- I denotes a countable (or finite) state space.

- $(\Omega, \mathcal{F}, \mathbb{P})$ is the probability space on which all the relevant random variables are defined.

Definition. $X = (X(t) : t \geq 0)$ is a (right-continuous) continuous-time random process with values in I if

(a) for all $t \geq 0$, $X(t) = X_t$ is a random variable such that $X(t) : \Omega \to I$;

(b) for all $\omega \in \Omega$, $t \mapsto X_t(\omega)$ is right-continuous (right-continuous sample path). In our case this means for all $\omega \in \Omega$, for all $t \geq 0$, there exists $\varepsilon > 0$ (depending on ω, t) such that

$$X_t(\omega) = X_s(\omega) \ \forall s \in [t, t + \varepsilon].$$

Fact. A right-continuous random process is defined by its finite-dimensional distributions

$$\mathbb{P}(X_{t_0 = i}, X_{t_1 = i_1}, \ldots, X_{t_n} = i_{t_n}), \ n \geq 0, \ t_k \geq 0, \ i_k \in I.$$

For every $\omega \in \Omega$, the path $t \mapsto X_t(\omega)$ of a right-continuous process stays constant for a while. So there are 3 possibilities:

(i) The path makes infinitely many jumps overall but only finitely many in a given interval $[0, t]$.

(ii) The path makes finitely many jumps & then gets absorbed in some state.

(iii) The path makes infinitely jumps in a finite time interval. After the 'explosion time' ζ, the process starts up again.

Write $J_0 = 0, J_1, J_2, \ldots$ for the jump times and $S_1, S_2, \ldots$ for the holding times, defined by

$$J_0 = 0, \ J_{n+1} = \inf\{t \geq J_n : X_t \neq X_{J_n}\},$$

$$S_n = \begin{cases} J_n - J_{n-1} & J_{n-1} < \infty \\ \infty & \text{otherwise} \end{cases}.$$

By right-continuity, $S_n > 0$ for all n. If $J_{n+1} = \infty$ for some n, we define $X_\infty = X_{J_n}$ as the final value, otherwise X_∞ is not defined. The explosion time ζ is defined by

$$\zeta = \sup(J_n) = \sum_{n=1}^\infty S_n.$$

We are not going to consider what happens to a chain after explosion. We thus set $X_t = \infty$ for all $t \geq \zeta$ (adjoining a new state '∞'). We call such a chain *minimal*.

Definition. We define the *jump chain* Y_n of $(X_t)_{t \geq 0}$ by setting $Y_n = X_{J_n}$ for all n.

Definition. A right-continuous random process $X = (X_t)_{t \geq 0}$ has the Markov property (and is called a continuous-time markov chain) if for all $i_1, i_2, \ldots, i_n \in I$ and $0 \leq t_1 < t_2 < \ldots < t_n$,

$$\mathbb{P}(X_{t_n} = i_n | X_{t_{n-1}} = i_{n-1}, \ldots, X_{t_0} = i_0) = \mathbb{P}(X_{t_n} = i_n | X_{t_0} = i_0).$$

Remark. For all $h > 0$, $Y_n = X(hn)$ defines a discrete-time Markov Chain.

Definition. The transition probabilities are $P_{ij}(s, t) = \mathbb{P}(X_t = j | X_s = i)$, $s \leq t, i, j \in I$. It is called *time-homogeneous* if it depends on $t - s$ only, i.e

$$P_{ij}(s, t) = P_{i,j}(0, t - s).$$

In this case we just write $P_{ij}(t - s)$. As in the case of discrete time, a (time-homogeneous) Markov process is characterised by

1. Its initial distribution $\lambda_i = \mathbb{P}(X_0 = i)$, $i \in I$;

2. Its *family of transition matrices* $(P(t))_{t \geq 0} = (P_{ij}(t))_{t \geq 0}$.

The family $(P(t))_{t \geq 0}$ is called the *transition subgroup* of the MC.

A (time-homogeneous) Markov process is characterised by

- its initial distribution;

- its transition subgroup $(P(t))_{t \geq 0}$

$$(P(t))_{t \geq 0} = (P(t))_{\substack{i,j \in I \\ t \geq 0}} = (\mathbb{P}(X_t = j | X_0 = i))_{\substack{i,j \in I \\ t \geq 0}}$$

It is easy to see that

- $P(0)$ is the identity

- $P(t)$ is a stochastic matrix for all t (i.e rows sum to 1)

- $P(t + s) = P(t)P(s)$ $\forall s, t$ (Chapman-Kolmogorov equation)

$$\begin{aligned}
P_{xz}(t + s) &= \mathbb{P}(X_{t+s} = z | X_0 = x) \\
&= \sum_{y \in I} \mathbb{P}(X_{t+s} = z | X_0 = x, X_t = y) \mathbb{P}(X_t = y | X_0 = x) \\
&= \sum_{y \in I} \mathbb{P}(X_s = z | X_0 = y) \mathbb{P}(X_t = y | X_0 = x) \\
&= \sum_{y \in I} P_{yz}(s) P_{xy}(t) = P_{x\cdot}(t) P_{\cdot z}(s)
\end{aligned}$$

Holding times

Let X be a (right-continuous continuous-time time-homogeneous) Markov Chain on a countable state-space I.

Suppose X starts from $x \in I$. Question: how long does X stay in the state x?

Definition. We call S_x the *holding time* at state x ($S_x > 0$ by right-continuity).

Let $s, t \geq 0$. Then

$$\begin{aligned}
\mathbb{P}(S_x > t + s | S_x > s) &= \mathbb{P}(X_u = x \; \forall u \in [0, t+s] | X_u = x \; \forall u \in [0, s]) \\
&= \mathbb{P}(X_u = x \; \forall u \in [s, t+s] | X_u = x \; \forall u \in [0, s]) \\
&= \mathbb{P}(X_u = x \; \forall u \in [s, t+s] | X_s = x) \\
&= \mathbb{P}(X_u = x \; \forall u \in [0, t] | X_0 = x) \\
&= \mathbb{P}(S_x > t).
\end{aligned}$$

Thus S_x has the memoryless property.

By the next theorem, we will get that S_x has the exponential distribution, say with parameter q_x.

Theorem 1.1 (Memoryless property). *Let S be a positive random variable. Then S has the memoryless property, i.e $\mathbb{P}(S > t + s | S > s) = \mathbb{P}(S > t)$ for all $s, t \geq 0$ if and only if S has the exponential distribution.*

Proof. It is easy to see the exponential distribution is memoryless. So we prove the other direction. Set $F(t) = \mathbb{P}(S > t)$. Then $F(s + t) = F(s)F(t)$ for all $s, t \geq 0$.

Since S is a positive random variable, there exists $n \in \mathbb{N}$ large such that $F(1/n) = \mathbb{P}(S > 1/n) > 0$. Then $F(1) = F(1/n)^n > 0$. So we can set $F(1) = e^{-\lambda}$ for some $\lambda \geq 0$.

For $k \in \mathbb{N}$, $F(k) = F(1)^k = e^{-\lambda k}$. For p/q rational, $F(p/q) = F(1/q)^p = (F(1/q)^q)^{p/q} = F(1)^{p/q} = e^{-\lambda \frac{p}{q}}$.

For any $t \geq 0$, for any $r, s \in \mathbb{Q}$ such that $r \leq t \leq s$, since F is decreasing

$$e^{-\lambda s} = F(s) \leq F(t) \leq F(r) = e^{-\lambda r}.$$

So taking sequences of rationals approaching t, we have $F(t) = e^{-\lambda t}$. $\qquad \square$

Poisson Process'

We are now going to look at the simplest (and most important) example of continuous time Markov Chains - the Poisson process.

Definition. Suppose $S_1, S_2, \ldots$ are iid random variables with $S_1 \sim \text{Exp}(\lambda)$. Define the *jump times* $J_0 = 0, J_1 = S_1, J_n = S_1 + \ldots + S_n$ for all n, and set $X_t = i$ if $J_i \leq t < J_{i+1}$. Then $I = \{0, 1, 2, \ldots\}$ and note that X is right-continuous and increasing. X is called a *Poisson process* of parameter/intensity λ. We sometimes refer to the jump times $(J_i)_{i \geq 1}$ as the *points* of the Poisson process, then $X =$ number of points in $[0, t]$.

Theorem 1.2 (Markov property). *Let $(X_t)_{t \geq 0}$ be a Poisson process of intensity λ. Then for all $s \geq 0$, the process $(X_{s+t} - X_s)_{t \geq 0}$ is also a Poisson process of intensity λ, and is independent of $(X_t)_{0 \leq t \leq s}$.*

Proof. Set $Y_t = X_{t+s} - X_s$ for all $t \geq 0$. Let $i \in \{0, 1, 2, \ldots\}$ and condition on $\{X_s = i\}$, Then the jump times for the process Y are $J_{n+1} - s, J_{n+2} - s, \ldots$ and the holding times are

$$\begin{aligned} T_1 &= J_{n+1} - s = S_{i+1} - (s - J_i) \\ T_2 &= S_{i+2} \\ T_3 &= S_{i+3} \\ &\vdots \end{aligned}$$

Since $\{X_s = i\} = \{J_i \leq s\} \cap \{s < J_{i+1}\} = \{J_i \leq s\} \cap \{S_{i+1} > s - J_i\}$, conditional on $\{X_s, i\}$, by the memoryless property of the exponential distribution (and

independence of S_{i+1} and J_i) we see that $T_1 \sim \text{Exp}(\lambda)$. Moreover the times J_j, $j \geq 2$ are independent of S_k, $k \leq i$ and hence independent of $(X_r)_{r \leq s}$, and they have iid $\text{Exp}(\lambda)$ distribution. Thus $((X_{s+t} - X_s))_{t \geq 0}$ is a Poisson process of parameter λ and is independent of $(X_t)_{0 \leq t \leq s}$. $\qquad\square$

Similar to this, one can show the Strong Markov property for a Poisson process of parameter λ. Recall a random variable $T \in [0, \infty]$ is called a *stopping time* if for all t, the event $\{T \leq t\}$ depends only on $(X_s)_{s \leq t}$.

Theorem 1.3 (Strong Markov property). *Let $(X_t)_{t \geq 0}$ be a Poisson process of parameter λ and T a stopping time. Then conditional on $T < \infty$, the process $(X_{T+t} - X_T)_{t \geq 0}$ is a Poisson process of parameter λ and independent of $(X_s)_{s \leq T}$.*

Theorem 1.4. *Let $(X_t)_{t \geq 0}$ be an increasing right-continuous process taking values in $\{0, 1, 2, \ldots\}$ with $X_0 = 0$. Let $\lambda > 0$. Then the following are equivalent*

(a) *The holding times $S_1, S_2, \ldots$ are iid $\mathrm{Exp}(\lambda)$ and the jump chain is given by $Y_n = n$ (i.e X is a poisson process of intensity λ)*

(b) *(Infinitesimal def) X has independent increments and as $h \downarrow 0$ uniformly in t we have*
$$\mathbb{P}(X_{t+h} - X_t = 1) = \lambda h + o(h),$$
$$\mathbb{P}(X_{t+h} - X_t = 0) = 1 - \lambda h + o(h).$$

(c) *X has independent and stationary increments and for all $t \geq 0$, $X_t \sim \mathrm{Poi}(\lambda t)$.*

Proof. First we show (a)$\Rightarrow$(b). If (a) holds, then by the Markov property, the increments are independent and stationary $((X_{t+s} - X_s)_{t \geq 0} =^d (X_t - X_0)_{t \geq 0})$. Using stationarity we have (uniformly in t) as $h \to 0$,

$$\mathbb{P}(X_{t+h} - X_t = 0) = \mathbb{P}(X_h = 0) = \mathbb{P}(S_1 > h) = e^{-\lambda h} = 1 - \lambda h + o(h),$$

$$\mathbb{P}(X_{t+h} - X_t \geq 1) = \mathbb{P}(X_h \geq 1) = \mathbb{P}(S_1 \leq h) = 1 - e^{-\lambda h} = \lambda h + o(h),$$

$$\begin{aligned}
\mathbb{P}(X_{t+h} - X_t \geq 2) = \mathbb{P}(X_h \geq 2) &= \mathbb{P}(S_1 + S_2 \leq h) \\
&\leq \mathbb{P}(S_1 \leq h, S_2 \leq h) \\
&= \mathbb{P}(S_1 \leq h)^2 \\
&= (\lambda h + o(h))^2 = o(h).
\end{aligned}$$

Now we show (b)$\Rightarrow$(c). If X satisfies (b), then $(X_{t+s} - X_s)_{t \geq 0}$ also satisfies (b). So X has independent and stationary increments. Now set $p_j(t) = \mathbb{P}(X_t = j)$. Then since increments are independent and X is increasing,

$$\begin{aligned}
p_j(t+h) = \mathbb{P}(X_{t+h} = j) = \sum_{i=0}^{j} \mathbb{P}(X_t = j - i)\mathbb{P}(X_{t+h} - X_t) \\
= p_j(t)(1 - \lambda h + o(h)) + p_{j-1}(t)(\lambda h + o(h)) + o(h).
\end{aligned}$$

Thus, $\frac{p_j(t+h) - p_j(t)}{h} = -\lambda p_j(t) + \lambda p_{j-1}(t) + o(1)$. Setting $s = t + h$, we get

$$\frac{p_j(s) - p_j(s-h)}{h} = -\lambda p_j(s-h) + \lambda p_{j-1}(s-h) + o(1).$$

In particular, $p_j(t)$ is continuous and differentiable with

$$p_j'(t) = -\lambda p_j(t) + \lambda p_{j-1}(t).$$

Differentiating

$$\left(e^{\lambda t} p(t)\right)' = \lambda e^{\lambda t} p_j(t) + e^{\lambda t} p_j'(t) = \lambda e^{\lambda t} p_{j-1}(t).$$

For $j = 0$ we have $p_0(t + h) = p_0(t)(1 - \lambda h + o(h))$, i.e $p_0'(t) = -\lambda p_0(t)$ so $p_0(t) = e^{-\lambda t}$. Thus

$$p_1'(t) = -\lambda p_1(t) + \lambda e^{-\lambda t}, \text{ i.e } p_1(t) = \lambda t e^{-\lambda t}.$$

And by induction

$$p_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!},$$

i.e $X_t \sim \text{Poi}(\lambda t)$.

Finally we show (c)$\Rightarrow$(a). We know X has independent stationary increments, We have for $t_1 \leq \ldots \leq t_k$, $n_1 \leq \ldots \leq n_k$,

$$\mathbb{P}(X_{t_1} = n_1, \ldots, X_{t_k} = n_k)$$
$$= \mathbb{P}(X_{t_1} = n_1) \underbrace{\mathbb{P}(X_{t_2} - X_{t_1} = n_2 - n_1)}_{\sim \text{Poi}(\lambda t_1)} \ldots \underbrace{\mathbb{P}(X_{t_k} - X_{t_{k-1}} = n_k - n_{k-1})}_{\sim \text{Poi}(\lambda(t_2 - t_1))}.$$

So (c) determines the finite-dimensional distributions (f.d.d) of a right-continuous process X, hence (c) determines X. So (c)$\Rightarrow$(a).

Question: can we show (a)$\Rightarrow$(c) directly? Indeed note

$$\mathbb{P}(X_t = n) = \mathbb{P}(S_1 + \ldots + S_n \leq t < S_1 + \ldots + S_{n+1})$$
$$= \mathbb{P}(S_1 + \ldots + S_n \leq t) - \mathbb{P}(S_1 + \ldots + S_{n+1} \leq t)$$
$$= \int_0^t \lambda e^{-\lambda x} \frac{(\lambda x)^{n-1}}{(n-1)!} \mathrm{d}x - \int_0^t \lambda e^{-\lambda x} \frac{(\lambda x)^n}{n!} \mathrm{d}x$$
$$= e^{-\lambda t} \frac{(\lambda t)^n}{n!} \text{ (integration by parts).}$$

$\square$

Theorem 1.5 (Superposition). *Let X and Y be two independent Poisson processes with parameters λ and μ respectively. Then $(Z_t)_{t \geq 0} = (X_t + Y_t)_{t \geq 0}$ is a Poisson process with parameter $\lambda + \mu$.*

Proof. We use (c) from the previous theorem. So Z has stationary independent increments. Also $Z_t \sim \text{Poi}(\lambda t + \mu t)$. $\square$

Theorem 1.6 (Thinning). *Let X be a Poisson process with parameter λ. Let $(Z_i)_{i \geq 1}$ be a sequence of iid Bernouilli(p) random variables. Let Y be a Posisson prcess with values in $\{0, \ldots, \}$ which jumps at time t if and only if X_t jumps at time t and $Z_{X_t} = 1$.*

In other words, we keep every point of X with probability p independently. Then Y is another Poisson process, with parameter λp and $X - Y$ is an independent Poisson process with parameter $\lambda(1 - p)$.

Proof. We shall use the infinitesimal definition. The independence of increments for Y is clear. Since $\mathbb{P}(X_{t+h} - X_t \geq 2) = o(h)$, we have

$$\mathbb{P}(Y_{t+h} - Y_t = 1) = p\mathbb{P}(X_{t+h} - X_t = 1) + o(h) = p\lambda h + o(h),$$

$$\begin{aligned}
\mathbb{P}(Y_{t+h} - Y_t = 0) &= \mathbb{P}(X_{t+h} - X_t = 0) + (1-p)\mathbb{P}(X_{t+h} - X_t = 1) + o(h) \\
&= 1 - \lambda h + (1-p)(\lambda h + o(h)) + o(h) \\
&= 1 - \lambda p h + o(h).
\end{aligned}$$

Hence Y is Poisson of paremeter λp. Clearly $X - Y$ is a thinning of X with Bernouilli parameter $1 - p$, so $X - Y$ is Poisson of parameter $\lambda(1-p)$.

Now we show Y and $X - Y$ are independent. It is enough to show that the f.d.d of Y and $X - Y$ are independent, i.e if $0 \leq t_1 \leq t_2 \leq \ldots \leq t_k$, $n_1 \leq \ldots \leq n_k$ and $m_1 \leq \ldots \leq m_k$, then we want to prove

$$\begin{aligned}
&\mathbb{P}(Y_{t_1} = n_1, \ldots, Y_{t_k} = n_k, X_{t_1} - Y_{t_1} = m_1, \ldots X_{t_k} - Y_{t_k} = m_k) \\
&= \mathbb{P}(X_{t_1} = n_1, \ldots, Y_{t_k} = n_k)\mathbb{P}(X_{t_1} - Y_{t_1} = m_1, \ldots, X_{t_k} - Y_{t_k} = m_K).
\end{aligned}$$

We will only show this for fixed t ($k = 1$) the general case follows similarly using independence of increments. We have

$$\begin{aligned}
\mathbb{P}(Y_t = n, X_t - Y_t = m) &= \mathbb{P}(X_t = m + n, Y_t = n) \\
&= \mathbb{P}(X_t = m + n)\mathbb{P}(Y_t = n | X_t = m + n) \\
&= e^{-\lambda t}\frac{(\lambda t)^{m+n}}{(m+n)!}\binom{m+n}{n}p^n(1-p)^m \\
&= e^{-\lambda t p}\frac{(\lambda t p)^n}{n!}e^{-\lambda t(1-p)}\frac{(\lambda t(1-p))^m}{m!} \\
&= \mathbb{P}(X_t = n)\mathbb{P}(X_t - Y_t = m),
\end{aligned}$$

as required. $\square$

Theorem 1.7. *Let X be a Poisson Process. Conditional on the event $(X_t = n)$, the jump times $J_1, J_2, \ldots, J_n$ are distributed as the order statistics of n iid $U[0,t]$ random variables. That is, they have joint density*

$$f(t_1, \ldots, t_n) = \frac{n!}{t^n} \mathbb{1}(0 \leq t_1 \leq \ldots \leq t_n \leq t).$$

Proof. Since $S_1, S_2, \ldots$ are iid $\text{Exp}(\lambda)$, the joint density of $(S_1, \ldots, S_{n+1})$ is

$$\lambda^{n+1} e^{-\lambda(S_1 + \ldots + S_{n+1})} \mathbb{1}(S_i \geq 0 \; for\, all\, i).$$

Then the jump times $J_1 = S_1, J_2 = S_1 + S_2, \ldots, J_{n+1} = S_1 + \ldots + S_{n+1}$ have joint density

$$g(t_1, \ldots, t_{n+1}) = \lambda^{n+1} e^{-\lambda t_{n+1}} \mathbb{1}(0 \leq t_1 \leq t_2 \leq \ldots t_{n+1}).$$

(Noting the Jacobian of the transformation is 1.) Now take $A \subseteq \mathbb{R}^n$ so

$$\mathbb{P}((J_1, \ldots, J_n) \in A | X_t = n) = \frac{\mathbb{P}((J_1, \ldots, J_n) \in A | X_t = n)}{\mathbb{P}(X_t = n)}.$$

Note

$$
\begin{aligned}
&\mathbb{P}((J_1, \ldots, J_n) \in A, X_t = n) \\
&= \mathbb{P}((J_1, \ldots, J_n) \in A, J_n \leq t < J_{n+1}) \\
&= \int_{(t_1, \ldots, t_{n+1}) \in A \times \mathbb{R}} g(t_1, \ldots, t_n) \mathbb{1}(t_{n+1} \geq t \geq t_n) dt_1 \ldots dt_{n+1} \\
&= \int_A \int_t^\infty \lambda^{n+1} e^{-\lambda t_{n+1}} \mathbb{1}(0 \leq t_1 \leq \ldots \leq t_n \leq t) dt_{n+1} dt_1 \ldots dt_n \\
&= \int_A \lambda^n e^{-\lambda t} \mathbb{1}(0 \leq t_1 \leq \ldots \leq t_n \leq t) dt_1 \ldots dt_n.
\end{aligned}
$$

Then we get

$$\mathbb{P}((J_1, \ldots, J_n) \in A | X_t = n) = \int_A \frac{n!}{t^n} \mathbb{1}(0 \leq t_1 \leq \ldots \leq t_n \leq t) dt_1 \ldots dt_n.$$

As required. $\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad$ $\square$

Now we look at a generalisation of a Poisson Process: called a Birth Process. For a Poisson Process, the rate of going from i to $i+1$ is λ. For a Birth Process, this is q_i (can depend on i). More precisely:

Definition (Birth Process)**.** For each i, let $S_i = \text{Exp}(q_i)$ with $S_1, S_2, \ldots$ independent. Set $J_i = S_1 + \ldots + S_i$ and $X_t = i$ if $J_i \leq t < J_{i+1}$. Then X is called a *Birth Process*.

We have some special cases:

1. <u>Simple birth process</u>: when $q_i = \lambda i$ for $i = 1, 2, \ldots$;

2. <u>Poisson Proces</u> $q_i = \lambda$ for all i.

Motivation for Simple Birth Process (SBP): at time 0 there is only one 'individual' i.e $X_0 = 1$. Each individual has an exponential clock of parameter λ independently. Then if there are i individuals, the first clock rings after $\mathrm{Exp}(\lambda i)$ time, and we jump from i to $i+1$ individuals. Indeed, by the memoryless property, the process begins afresh after each jump.

Proposition 1.8. *Let $(T_k)_{k \geq 1}$ be a sequence of independent random variables with $T_K \sim \mathrm{Exp}(q_k)$ and $\sum_k q_k < \infty$. Let $T = \inf_k T_k$. Then*

(a) $T \sim \mathrm{Exp}\left(\sum_k q_k\right)$

(b) The infimum is attained at a point T_K almost surely, and

$$\mathbb{P}(K = n) = \frac{q_n}{\sum_k q_k}.$$

(c) T and K are independent.

Proof. See example sheet. □

The main difference between a Poisson Process and a Birth Process is that there is the possibility of explosion in the Birth Process. Recall explosion occurs when $\zeta := \sum_n S_n < \infty$.

Proposition 1.9. *Let X be a Birth Process with rates q_i and $X_0 = 1$. Then*

1. *If $\sum_{i=1}^{\infty} \frac{1}{q_i} < \infty$, then X is explosive, i.e $\mathbb{P}(\zeta < \infty) = 1$;*

2. *If $\sum_{i=1}^{\infty} \frac{1}{q_i} = \infty$, then X is non-explosive, i.e $\mathbb{P}(\zeta = \infty) = 1$.*

Remark. This shows the SBP (as well as the PP) is non-explosive.

Proof.

1. If $\sum_n \frac{1}{q_n} < \infty$, then

$$\mathbb{E}[\zeta] = \mathbb{E}\left[\sum_n S_n\right] = \sum_n \mathbb{E}S_n = \sum_n \frac{1}{q_n} < \infty.$$

Where we have swapped summation and expectation by the MCT (monotone convergence theorem). Thus $\zeta = \sum_n S_n < \infty$ almost surely.

2. If $\sum_n \frac{1}{q_n} = \infty$, then $\prod_n \left(1 + \frac{1}{q_n}\right) \geq 1 + \sum_n \frac{1}{q_n} = \infty$. Then

$$
\begin{aligned}
\mathbb{E}[e^{-\zeta}] &= \mathbb{E}\left[e^{-\sum_{n=1}^{\infty} S_n}\right] \\
&= \lim_{n\to\infty}\left[e^{-\sum_{i=1}^{n} S_i}\right] && \text{(MCT)} \\
&= \lim_{n\to\infty}\prod_{i=1}^{n}\mathbb{E}[e^{-S_i}] && \text{(independence)} \\
&\leq \lim_{n\to\infty}\prod_{i=1}^{n}\frac{1}{1+1/q_i} = 0.
\end{aligned}
$$

Since $e^{-\zeta} \geq 0$, since $\mathbb{E}(e^{-\zeta}) = 0$ we have $e^{-\zeta} = 0$ almsot surely, i.e $\mathbb{P}(\zeta = \infty) = 1$.

$\square$

Theorem 1.10 (Markov Property)**.** *Let X be a BP with parameters (q_i). Conditional on $X_s = i$, the process $(X_{s+t})_{t \geq 0}$ is a birth process with rates $(q_j)_{j \geq i}$ starting from i, and independent of $(X_r)_{r \leq s}$.*

Proof. As in the Poisson Process case. $\qquad\qquad\qquad\qquad\qquad\qquad\qquad$ $\square$

Theorem 1.11. *Let X be an increasing right-continuous process with values in $\{1, 2, \ldots\} \cup \{\infty\}$. Let $0 \leq q_j < \infty$ for all $j \geq 0$. Then the following are equivalent:*

1. *(jump chain/holding time definition) conditional on $X_s = i$, the holding times $S_1, S_2, \ldots$ are independent exponentials with rates $q_i, q_{i+1}, \ldots$ respectively and the jump chain is given $Y_n = i + n$ for all n.*

2. *(infinitesimal definition) for all $t, h \geq 0$, conditional on $X_t = i$, the process $(X_{t+h})_{h \geq 0}$ is independent of $(X_s)_{s \leq t}$ and as $h \to 0$, uniformly in t we have*

$$\mathbb{P}(X_{t+h} = i | X_t = i) = 1 - q_i h + o(h),$$

$$\mathbb{P}(X_{t+h} = i + 1 | X_t = i) = q_i h + o(h).$$

3. *(transition probability definition) for all $n = 0, 1, 2, \ldots$ and all times $0 \leq t_0 \leq t_1 \leq \ldots \leq t_{n+1}$, and all states $i_0, i_1, \ldots, i_{n+1}$,*

$$\mathbb{P}(X_{t_{n+1}} = i_{n+1} | X_0 = i_0, \ldots, X_{t_n} = i_n) = p_{i_n, i_{n+1}}(t_{n+1} - t_n),$$

where $(p_{ij}(t) : i, j = 0, 1, 2, \ldots)$ is the unique solution to the equation (called Kolmogorov's forward equation)

$$p'_{ij}(t) = q_{j-1} p_{i,j-1}(t) - q_j p_{i,j}(t). \qquad\qquad (*)$$

(as in the Poisson Process, $p_{ij}(t+h) = p_{i,j-1}(t)q_j h + p_{i,j}(t)(1 - q_j h) + o(h)$.)

Existence and uniqueness of a solution in (3) gollow since for $i = j$ $p'_{i,i}(t) = -q_i p_{i,i}(t)$ and $p_{i,i}(0) = 1$, so $p_{i,i}(t) = e^{-q_i t}$. Then by induction, if the unique solution for $p_{i,j}(t)$ exists, then plug into $(*)$ to see there exists a unique solution for $p_{i,j+1}(t)$.

Also note that we can write the equation in matrix form:

$$P'(t) = P(t)Q, \text{ where } Q = \begin{pmatrix} -q_1 & q_1 & 0 & \ldots \\ 0 & -q_2 & q_2 & \ldots \\ \vdots & \ddots & \ddots & \ddots \end{pmatrix}.$$

Q-matrix and construction of Markov Processes

Definition. $Q = (q_{ij})_{i,j \in I}$ is called a Q-matrix if

(a) $-\infty < q_{ii} \leq 0$ for all $i \in I$;

(b) $0 \leq q_{ij} < \infty$ for all $i, j \in I$ with $i \neq j$;

(c) $\sum_{j \in I} q_{ij} = 0$ for all $i \in I$.

Write $q_i = -q_{ii} = \sum_{j \notin I} q_{ij}$ for all $i \in I$.

Given a Q-matrix Q, we define a jump matrix P as follows. For $x \neq y$ with $q_x \neq 0$, set $p_{xy} = \frac{q_{xy}}{q_x}$ and $p_{xx} = 0$. If $q_x = 0$, set $p_{xy} = \mathbb{1}(x = y)$.

Example.

$$
Q = \begin{pmatrix} -1 & 1 & 0 \\ 1 & -2 & 1 \\ 2 & 1 & -3 \end{pmatrix} \implies P = \begin{pmatrix} 0 & 1 & 0 \\ 1/2 & 0 & 1/2 \\ 2/3 & 1/3 & 0 \end{pmatrix}.
$$

Definition. Let Q be a Q-matrix and λ a probability measure on the state space I. Then a (minimal) random process X is a *Markov process* with initial distribution λ and infinitesimal generator Q if

(a) The jump chain $Y_n = X_{J_n}$ is a discrete time Markov chain starting from $Y_0 \sim \lambda$ with transition matrix P.

(b) Conditional on $Y_0, Y_1, \ldots, Y_n$, the holding times $S_1, \ldots, S_{n+1}$ are independent with $S_i \sim \mathrm{Exp}(q_{Y_{i-1}})$ for $i = 1, \ldots, n+1$.

We write $X \sim \mathrm{Markov}(\lambda, Q)$.

Example. Birth-Processes are $\mathrm{Markov}(\lambda, Q)$ with $I = \mathbb{N}$ and

$$
Q = \begin{pmatrix} -q_1 & q_1 & 0 & \cdots \\ 0 & -q_2 & q_2 & \cdots \\ \vdots & \ddots & \ddots & \ddots \end{pmatrix} \text{ and } P = \begin{pmatrix} 0 & 1 & 0 & \cdots \\ 0 & 0 & 1 & \cdots \\ \vdots & \ddots & \ddots & \ddots \end{pmatrix}.
$$

And jump chain $Y_n = Y_0 + n$.