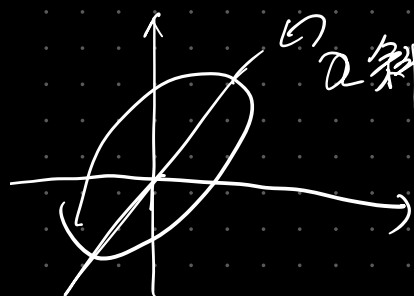


$$Y = aX \Rightarrow E(Y - aX)^2 \text{ (mean square Error)}$$



$$f_{XY}(x,y) = \begin{cases} \frac{1}{2\pi} & (x,y) \in D \\ 0 & \text{others} \end{cases}$$

$$\frac{\partial}{\partial x} = \frac{2f(x)a}{\partial} \Rightarrow a = \frac{Y}{X} = \frac{XY}{X^2}$$

两者相差一个系数

$$\Rightarrow \min_a (Y - aX)^2 \Rightarrow a_{opt} = \frac{E(XY)}{E(X^2)} \rightarrow \frac{E(XY)}{E(X^2)} \xrightarrow{\text{Correlation}} R$$

$$E(X - EX)(Y - EY) = E(XY) - E(X)E(Y) - E(Y)E(X) + E(X)E(Y) = E(XY) - E(X)E(Y)$$

不正负两者相关的表达 (只差一个常数而已)

$$\text{Uncorrelated: } E(XY) = 0$$

↑
Independence:

$$\theta \sim U(0, 2\pi)$$

$$X = \cos(\theta), Y = \sin(\theta)$$

$$X^2 + Y^2 = 1$$

$$\Rightarrow EX = \int_{-\infty}^{\infty} \cos\theta \cdot f_{\theta}(\theta) d\theta = \int_0^{2\pi} \cos\theta \cdot \frac{1}{2\pi} d\theta = 0$$

$$EY = 0, E(XY) = \frac{1}{2\pi} \int_0^{2\pi} \sin\theta \cdot \cos\theta d\theta = 0$$

$$E(XY) = EX \cdot EY \text{ 故 } X \text{ 与 } Y \text{ 不相关}$$

Geometric View

$$E(XY) = \langle X, Y \rangle \text{ Inner Product 内积}$$

$$\langle x, y \rangle: H \times H \rightarrow \mathbb{R} \quad ① \langle x, y \rangle = \langle y, x \rangle \quad ② \langle x, x \rangle \geq 0$$

$$\langle x, x \rangle = 0 \Rightarrow x = 0$$

$$③ \text{Bilinear: } \langle x, \alpha y + \beta z \rangle = \alpha \langle x, y \rangle + \beta \langle x, z \rangle$$

$$\langle \alpha x + \beta y, z \rangle = \alpha \langle x, z \rangle + \beta \langle y, z \rangle$$

$$E|X|^2 = 0 \Rightarrow X = 0$$

$$\Downarrow P(X=0)=1$$

概率为0并不是不发生。反例：数轴上取一个点，取到它的P为0，但不能说有无穷多个点。

把相关看成内积

$$\cos \angle(x, y) = \frac{\langle x, y \rangle}{(\langle x, x \rangle \langle y, y \rangle)^{1/2}} \quad \text{Random Variable} \Leftrightarrow \text{Vector}$$

$$\cos \angle(X, Y) = \frac{E(XY)}{[E X^2 \cdot E Y^2]^{1/2}} \quad \text{Cauchy-Schwarz}$$

保证 \cos 在 $[-1, 1]$ 之间

$$0 \leq \rho(A) = \langle \lambda x + y, \lambda x + y \rangle = \lambda^2 \langle x, x \rangle + 2\lambda \langle x, y \rangle + \langle y, y \rangle \Rightarrow \Delta \leq 0$$

$$|\sum_k x_k y_k| \leq (\sum_k x_k^2)^{1/2} (\sum_k y_k^2)^{1/2} \quad \int f(x)g(x)dx \leq (\int f^2 dx \int g^2 dx)^{1/2}$$

Cauchy-Schwarz 其他形式



$$\|Y\| \cos \theta \frac{X}{\|X\|} = \left(\frac{\|Y\|}{\|X\|} \cos \theta \right) \cdot X$$

$$\frac{\|Y\|}{\|X\|} \cdot \frac{E(XY)}{\|X\| \|Y\|} = \frac{E(XY)}{E X^2} = \rho$$

Correlation Function

自相关 Auto

随机场

$X(t)$: 随机过程是一组随机变量, t : Index | 两维: Random Field.

$R_X(t, s) = E(X(t)X(s))$. Binary \rightarrow Univariate Assumption

① $R_X(t, s) = R_X(s, t)$

② $|R_X(t, s)| \leq (R_X(t, t) R_X(s, s))^{1/2}$ \leftarrow 相关运算带来的

Stationary

平稳性

\Leftarrow Invariance

不变性: 某类统计性质随着时间变化保持不变的特性

令 D 满足 $X(t) \sim m(t)$

Wide-Sense

宽平稳

① 均值 $E(X(t)) = m(t) \equiv m$ (常数)

② $R_X(t, s) = R_X(t+D, s+D), \forall D \in \mathbb{R}$

$\hookrightarrow R_X(t, s) = R_X(t-s) = R_X(\tau)$

相关系数简化

$\tau = t - s$

例子:

① Modulated Signal

二元 \rightarrow 一元
 R_X 变为只与 $t-s$ 有关
 $\tau = t - s$

$x(t) = A(t) \cdot \cos(2\pi f_0 t + \theta)$, $A(t)$, i.i.d. $\theta \sim U(0, 2\pi)$ - Independent
 $E(x(t)) = E(A(t)) \cdot E(\cos(2\pi f_0 t + \theta))$

$$= E(A(t)) \int_0^{2\pi} \frac{1}{2\pi} \cos(2\pi f_0 t + \theta) d\theta = 0$$

$$\begin{aligned}
 R_x(t, s) &= E(A(t) \cdot A(s)) E(\cos(2\pi f_0 t + \theta) \cdot \cos(2\pi f_0 s + \theta)) \\
 &= E(A(t) \cdot A(s)) \left[\frac{1}{2} (E \cos(2\pi f_0 (t+s) + 2\theta) + E(\cos(2\pi f_0 (t-s))) \right] \\
 &= E(A(t) \cdot A(s)) E(\cos(2\pi f_0 (t-s))) \\
 &= \frac{1}{2} R_A(t-s) \cdot \cos(2\pi f_0 (t-s))
 \end{aligned}$$

振幅调制是宽平稳, 那么总的就是宽平稳的

且随机噪声与振幅独立

② Random Telegraph Signal



$[s, t]$ 内, k 切片, 式-1 具有随机性

$$P = \frac{\lambda(t-s)^k}{k!} \exp(-\lambda(t-s)) \quad \text{Poisson Distribution}$$

$$E(x(t) \cdot x(s)) = R_x(t, s) = 1 \cdot P_1 + (-1) \cdot P_{-1}$$

$x(t)$ 与 $x(s)$ 两者乘积
 只能是 1 或 -1

$$P_1 = P([s, t], \text{even}) = \sum_{k \text{ even}} \frac{(\lambda(t-s))^k}{k!} \exp(-\lambda(t-s))$$

$$P_{-1} = P([s, t], \text{odd}) = \sum_{k \text{ odd}} \frac{(\lambda(t-s))^k}{k!} \exp(-\lambda(t-s))$$

$$\sum_{k=0}^{\infty} \frac{\lambda(t-s)^k}{k!} = \exp(\lambda(t-s)), \quad \sum_{k \text{ even}} \frac{(\lambda(t-s))^k}{k!} = \frac{1}{2} (\exp(\lambda(t-s)) + \exp(-\lambda(t-s)))$$

$$P_1 = \frac{1}{2} (1 + \exp(-2\lambda(t-s))), \quad P_{-1} = \frac{1}{2} (1 - \exp(-2\lambda(t-s)))$$

$$\begin{aligned}
 R_x(t, s) &= \frac{1}{2} (1 + \exp(-2\lambda(t-s))) - \frac{1}{2} (1 - \exp(-2\lambda(t-s))) \\
 &= \exp(-2\lambda(t-s))
 \end{aligned}$$